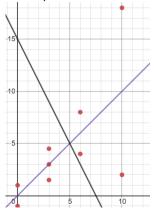
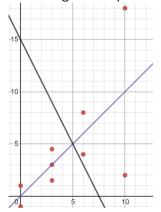
CSCI 3022 Intro to Data Science Regression Inference

Consider the graph below. Do either of the candidate "best fit" lines violate the 4 big assumptions?

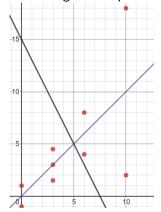


Do either of the candidate "best fit" lines violate the 4 big assumptions? One thing to do: plot the errors as a function of X:



Errors are the values given by "estimated line minus data."

Do either of the candidate "best fit" lines violate the 4 big assumptions? One thing to do: plot the errors as a function of X:



Errors are the values given by "estimated line minus data." Black line the errors clump and move up/down as X moves left-right.

Blue line the errors increase in magnitude as X goes right.

We've looked at the following test statistics for hypothesis testing.

1. To compare proportions against a baseline or against each other, we use Z-statistics.

$$\frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1 - p_0)}{n}}} \ \mathbf{OR} \ \frac{(\hat{p_1} - \hat{p_2}) - \Delta_0}{\sqrt{\frac{\hat{p}(1 - \hat{p})}{n_1} + \frac{\hat{p}(1 - \hat{p})}{n_2}}}$$

2. To compare means when the samples are large **or** underlying normal with *known* variances, we also use Z-statistics.

$$\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \ \ \mathbf{OR} \ \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \ \ \mathbf{OR} \ \frac{\left(\bar{X} - \bar{Y}\right) - \left(\mu_1 - \mu_2\right)}{\sqrt{\frac{\sigma_1^2}{n} + \frac{\sigma_2^2}{m}}} \ \ \mathbf{OR} \ \frac{\left(\bar{X} - \bar{Y}\right) - \left(\mu_1 - \mu_2\right)}{\sqrt{\frac{s_1^2}{n} + \frac{s_2^2}{m}}}$$

3. To compare means when the samples are small ${\bf and}$ underlying normal, we use t-statistics.

$$t = \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \ \ \mathbf{OR} \ \frac{\left(\bar{X} - \bar{Y}\right) - \left(\mu_1 - \mu_2\right)}{\sqrt{\frac{s_1^2}{n} + \frac{s_2^2}{m}}}$$

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

With 3 assumptions on ε :

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$\operatorname{Cov}[\varepsilon_i, \varepsilon_j] = 0 \qquad \forall i, j$$

Independence of errors

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$\operatorname{Cov}[\varepsilon_i, \varepsilon_j] = 0 \qquad \forall i, j$$

Independence of errors

3.

$$Var(\varepsilon_i) = \sigma^2 \qquad \forall i$$

Homoskedasticity of errors

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$\mathsf{Cov}[\varepsilon_i, \varepsilon_j] = 0 \qquad \forall i, j$$

Independence of errors

3.

$$Var(\varepsilon_i) = \sigma^2 \quad \forall i$$

Homoskedasticity of errors

4.

Simple Linear Regression Model

The β estimators in the model are:

1.
$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

2.
$$\hat{\beta}_1 = \frac{Cov[X,Y]}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Important Terminology:

- ightharpoonup x: the independent variable, predictor, or explanatory variable (usually known). x is not random.
- \triangleright Y: The dependent variable or response variable. For fixed x, Y is random.
- \triangleright ε : The random deviation or random error term. For fixed x, ε is random. Has variance σ^2 .
- \triangleright β : the regression coefficients.
- ightharpoonup r: the *residuals* or observed errors. Used to estimate σ^2 .

Definitions:

1. The *fitted (or predicted) values* __ are obtained by plugging in __ to the equation of the estimated regression line:

2. The *residuals* are the differences between the observed and fitted y values:

Definitions:

1. The fitted (or predicted) values $\underline{\hat{Y}_i}$ are obtained by plugging in $\underline{\hat{X}_i}$ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

2. The *residuals* are the differences between the observed and fitted y values:

Definitions:

1. The *fitted* (or predicted) values __ are obtained by plugging in __ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

2. The *residuals* are the differences between the observed and fitted y values:

$$\hat{\varepsilon_i} = r_i = \hat{e_I} = Y_i - \hat{Y_i} = Y_i - \hat{\beta_0} + \hat{\beta_1} X_i$$

Definitions:

1. The *fitted* (or predicted) values __ are obtained by plugging in __ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

2. The *residuals* are the differences between the observed and fitted y values:

$$\hat{\varepsilon_i} = r_i = \hat{e_I} = Y_i - \hat{Y_i} = Y_i - \hat{\beta_0} + \hat{\beta_1} X_i$$

Residuals are estimates of the true error. Why?

We don't have the true values of β_0, β_1 , so when we estimate them we get variance and error in our estimates.

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

1.
$$\hat{\beta_0} =$$

2.
$$\hat{\beta_1} =$$

What happens if $\beta_0 \approx 0$? If $\beta_1 \approx 0$?

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

1.
$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

2.
$$\hat{\beta}_1 = \frac{Cov[X,Y]}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

What happens if $\beta_0 \approx 0$? If $\beta_1 \approx 0$?

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

1.
$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

2.
$$\hat{\beta}_1 = \frac{Cov[X,Y]}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

What happens if $\beta_0 \approx 0$? If $\beta_1 \approx 0$?

One result: the regression line goes through $(0, \beta_0)$. It also goes through $(\bar{X}, \bar{Y})!$

Definitions:

1. The *fitted (or predicted) values* __ are obtained by plugging in __ to the equation of the estimated regression line:

2. The *residuals* are the differences between the observed and fitted y values:

Definitions:

1. The fitted (or predicted) values $\underline{\hat{Y}_i}$ are obtained by plugging in $\underline{\hat{X}_i}$ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

2. The *residuals* are the differences between the observed and fitted y values:

Definitions:

1. The *fitted (or predicted) values* __ are obtained by plugging in __ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

2. The *residuals* are the differences between the observed and fitted y values:

$$\hat{\varepsilon_i} = r_i = \hat{e_I} = Y_i - \hat{Y_i} = Y_i - \hat{\beta_0} + \hat{\beta_1} X_i$$

Definitions:

1. The *fitted (or predicted) values* __ are obtained by plugging in __ to the equation of the estimated regression line:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

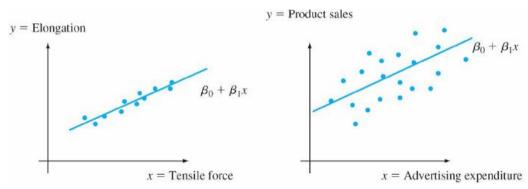
2. The *residuals* are the differences between the observed and fitted y values:

$$\hat{\varepsilon_i} = r_i = \hat{e_I} = Y_i - \hat{Y_i} = Y_i - \hat{\beta_0} + \hat{\beta_1} X_i$$

Residuals are estimates of the true error. Why?

We don't have the true values of β_0, β_1 , so when we estimate them we get variance and error in our estimates.

The parameter σ^2 determines the amount of spread about the true regression line. Two separate examples:



An estimate of σ^2 will be used in confidence interval formulas and hypothesis testing procedures presented in the next days. Recall that the residual sum of squares or sum of squared errors (SSE) is:

SSE =

So, our estimate of the variance of the model is like a measure for an average of this summand:

An estimate of σ^2 will be used in confidence interval formulas and hypothesis testing procedures presented in the next days. Recall that the residual sum of squares or sum of squared errors (SSE) is:

$$SSE = \sum (errors)^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}X_{i})^{2}$$

So, our estimate of the variance of the model is like a measure for an average of this summand:

An estimate of σ^2 will be used in confidence interval formulas and hypothesis testing procedures presented in the next days. Recall that the residual sum of squares or sum of squared errors (SSE) is:

$$SSE = \sum (errors)^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}X_{i})^{2}$$

So, our estimate of the variance of the model is like a measure for an average of this summand:

$$\hat{\sigma^2} = \frac{SSE}{n-2}$$

Wait, what? Why the n-2??

An estimate of σ^2 will be used in confidence interval formulas and hypothesis testing procedures presented in the next days. Recall that the residual sum of squares or sum of squared errors (SSE) is:

$$SSE = \sum (errors)^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}X_{i})^{2}$$

So, our estimate of the variance of the model is like a measure for an average of this summand:

$$\hat{\sigma^2} = \frac{SSE}{n-2}$$

Wait, what? Why the n-2?? These are again *degrees of freedom*.

Degrees of Freedom Intuition

Suppose you have 3 (random) points on the XY plane.

- 1. Can you draw a line through them?
- 2. Can you draw a parabola through them?
- 3. Can you draw a cubic function through them?

4. Can you draw a quartic function through them?

Degrees of Freedom Intuition

Suppose you have 3 (random) points on the XY plane.

- Can you draw a line through them?
 It's very unlikely. In fact, for truly random (normal) points, this result has probability zero!
- Can you draw a parabola through them? Yes, but there's only one such parabola.
- 3. Can you draw a cubic function through them? Yes. Not only that, you could choose any one of a,b,c,d in the $ax^3+bx^2+cx+d=0$ and then solve for the others. You have **one degree of freedom**.
- 4. Can you draw a quartic function through them? Yes. Not only that, you could choose any two of a,b,c,d,e in the $ax^4+bx^3+cx^2+dx+e=0$ and then solve for the others. You have **two degrees of freedom**.

Degrees of Freedom

The takeaway?

One property of mathematical estimation: the more you estimate, the more you risk overfitting. In this model we've estimated **2** "means" $(\hat{\beta}_0, \hat{\beta}_1)$ before we got to σ , which "costs" us two degrees of freedom.

The more we estimate, the less options - degrees of freedom - we get for the remaining terms.

Some properties of our estimate:

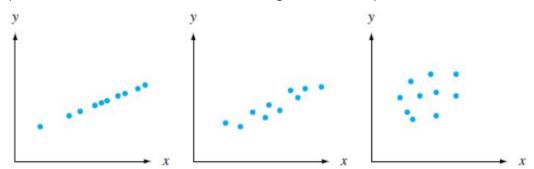
1. The divisor n-2 in is the number of degrees of freedom (df) associated with SSE and $\hat{\sigma}^2$.

2. This is because to obtain $\hat{\sigma}^2$, two parameters must first be estimated, which results in a loss of 2 df.

3. Replacing each y_i in the formula for $\hat{\sigma}^2$ by the r.v. Y_i gives a random variable.

4. It can be shown that the r.v. $\hat{\sigma}^2$ is an unbiased estimator for σ^2 .

The residual sum of squares SSR can be interpreted as a measure of how much variation in y is left unexplained by the model—that is, how much cannot be attributed to a linear relationship. In the first plot, SSE=0, and there is no unexplained variation, whereas unexplained variation is small for second, and large for the third plot.



Picturing Sums of Squares

The goodness-of-fit of a regressive model is often decomposed into three components based on squared deviations. These are:

1. SSE: Sum of squared errors: (vertical) distances from the regression line to the data values.

2. SST: Sum of squares, total: total deviation in Y. Looks like Var[Y].

3. SSR: Sum of squares of regression line: the amount of variability tied to the model.

Picturing Sums of Squares

The goodness-of-fit of a regressive model is often decomposed into three components based on squared deviations. These are:

1. SSE: Sum of squared errors: (vertical) distances from the regression line to the data values.

$$\sum_{i} \left(\hat{Y} - Y_i \right)^2$$

2. SST: Sum of squares, total: total deviation in Y. Looks like Var[Y].

$$\sum_{i} (Y_i - \bar{Y})^2$$

3. SSR: Sum of squares of regression line: the amount of variability tied to the model.

$$\sum_{i} \left(\hat{Y}_{i} - \bar{Y} \right)^{2}$$

Mullen: OLS-SLR Theory

Picturing Sums of Squares

The sum of squared deviations about the least squares line is smaller than the sum of squared deviations about any other line, i.e. SSE < SST unless the horizontal line itself is the least squares line.

The ratio SSE/SST is the proportion of total variation that cannot be explained by the simple linear regression model. The coefficient of determination is:

This coefficient is a number between 0 and 1 and is the proportion of observed y variation explained by the model.

The sum of squared deviations about the least squares line is smaller than the sum of squared deviations about any other line, i.e. SSE < SST unless the horizontal line itself is the least squares line.

The ratio SSE/SST is the proportion of total variation that cannot be explained by the simple linear regression model. The coefficient of determination is:

$$R^2 = 1 - \frac{SSE}{SST} = \frac{SSR}{SST}$$

This coefficient is a number between 0 and 1 and is the proportion of observed y variation explained by the model.

Mullen: OLS-SLR Theory

Again, \mathbb{R}^2 is the proportion of observed y variation explained by the model.

The higher the value of \mathbb{R}^2 , the more successful is the simple linear regression model in explaining y variation, assuming the linear model is correct.

Again, \mathbb{R}^2 is the proportion of observed y variation explained by the model.

The higher the value of \mathbb{R}^2 , the more successful is the simple linear regression model in explaining y variation, assuming the linear model is correct.

Crucially, R^2 is a measure of *linear* dependence between X and Y. If $R^2=0$, X and Y may still be related! Ex: $Y=X^2(+\varepsilon)$.

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., _____), compute confidence intervals, etc.

Distributions:

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., $H_0: \beta_1=0$), compute confidence intervals, etc.

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}; \quad \hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Distributions:

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., $\underline{H}_0: \beta_1=0$), compute confidence intervals, etc.

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}; \quad \hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Distributions:

$$\hat{\beta}_0 \sim N \left(\beta_0, \frac{\sigma^2}{n} + \frac{\sigma^2 \bar{X}^2}{\left(X_i - \bar{X} \right)^2} \right)$$

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{\left(X_i - \bar{X}\right)^2}\right)$$

... but of course, we don't know σ^2 , so we estimate with SSE/(n-2).

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta}_0) = \sigma \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{(X_i - \bar{X})^2}}; \qquad s.e.(\hat{\beta}_1) = \sigma \sqrt{\frac{1}{(X_i - \bar{X})^2}}$$

These lead to CIs of

where we replace σ with the estimate $s = \frac{SSE}{n-2}$

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta}_0) = \sigma \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{(X_i - \bar{X})^2}}; \qquad s.e.(\hat{\beta}_1) = \sigma \sqrt{\frac{1}{(X_i - \bar{X})^2}}$$

These lead to CIs of

$$\beta_i \in (\hat{\beta}_i \pm t_{\alpha/2, n-2} \cdot s.e.(\hat{\beta}_i))$$

where we replace σ with the estimate $s = \frac{SSE}{n-2}$

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta}_0) = \sigma \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{(X_i - \bar{X})^2}}; \qquad s.e.(\hat{\beta}_1) = \sigma \sqrt{\frac{1}{(X_i - \bar{X})^2}}$$

These lead to CIs of

$$\beta_i \in (\hat{\beta}_i \pm t_{\alpha/2, n-2} \cdot s.e.(\hat{\beta}_i))$$

where we replace σ with the estimate $s=\frac{SSE}{n-2}$

Tests then result from comparing $t=\frac{\hat{\beta_i}}{s.e.(\hat{\beta_i})}$ to the corresponding critical t values for a one or two-tailed test.

Inferences about Y

There are more types on confidence intervals we may care about!

- 1. Last slide was how to perform inference on the **parameters** of the *line* β . We also might care about inference on values of Y!
- 2. A **confidence band** is how sure we are about the mean of Y at specific values of X, or E[Y|X].
- 3. A **prediction band** is how we estimate the distribution of new Y observations at specific values of X. It's the same as the confidence band, but also includes our estimate for ε .

See: nb accompanying lecture: SLR Inference

Daily Recap

Today we learned

1. Regression Inference!

Moving forward:

- nb day Friday

Next time in lecture:

- More Regression! More predictor!