

# K-Means Clustering and Gaussian Mixture Model

Il-Chul Moon  
Dept. of Industrial and Systems Engineering  
KAIST

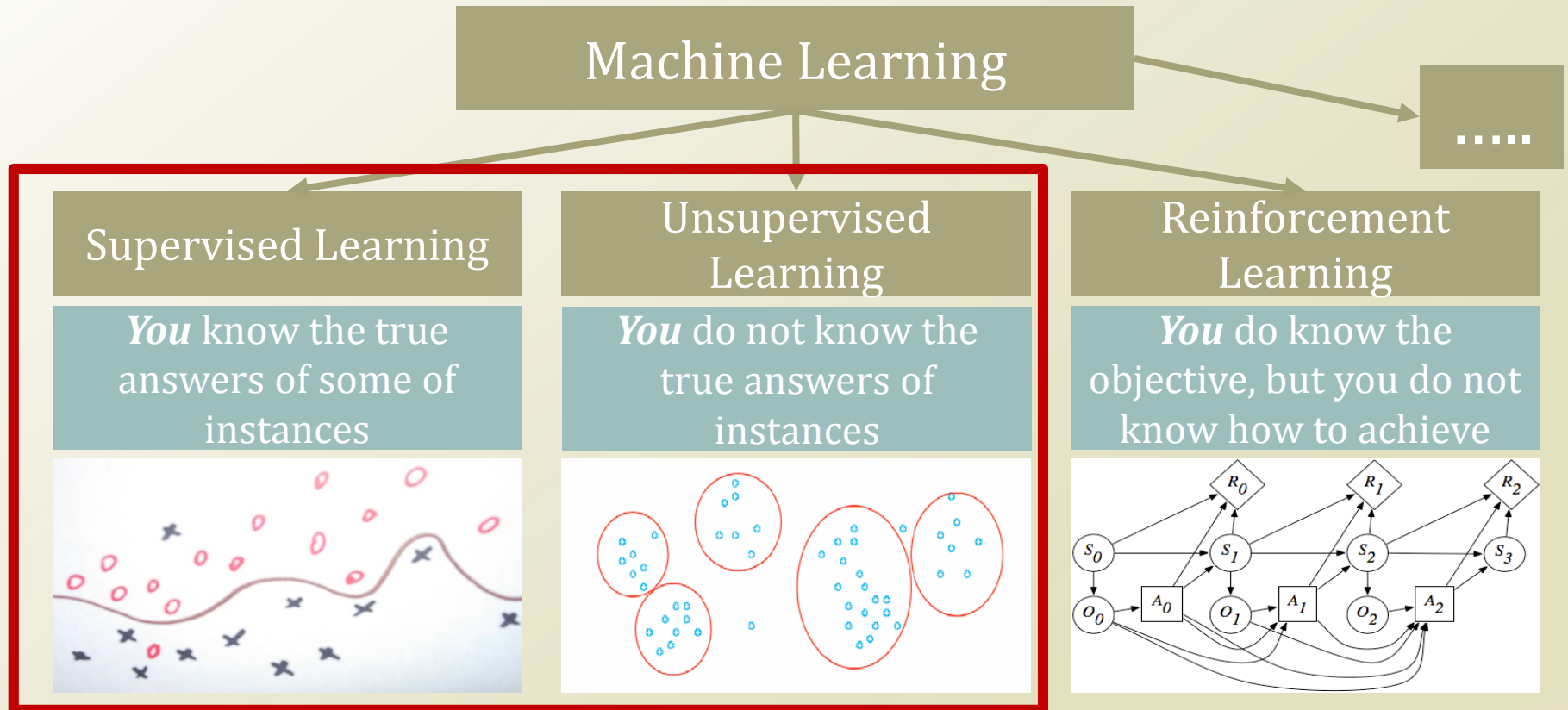
[icmoon@kaist.ac.kr](mailto:icmoon@kaist.ac.kr)

# Weekly Objectives

- Understand the clustering task and the K-means algorithm
  - Know what the unsupervised learning is
  - Understand the K-means iterative process
  - Know the limitation of the K-means algorithm
- Understand the Gaussian mixture model
  - Know the multinomial distribution and the multivariate Gaussian distribution
  - Know why mixture models are useful
  - Understand how the parameter updates are derived from the Gaussian mixture model
- Understand the EM algorithm
  - Know the fundamentals of the EM algorithm
  - Know how to derive the EM updates of a model

# K-MEANS ALGORITHM

# Types of Machine Learning

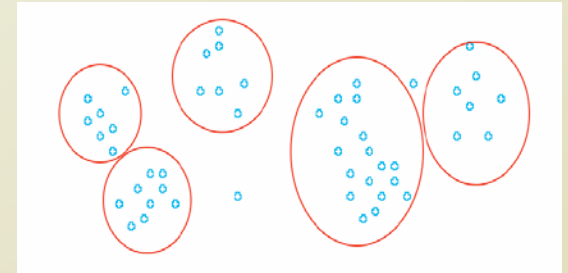


- **You** can
  - Machine learning
  - Dataset provider
  - Machine learning users
  - etc
- Various classifications by different professors
  - Purpose, data types, etc
- Other learning classifications also exist

*You* do not know the true answers of instances

# Unsupervised Learning

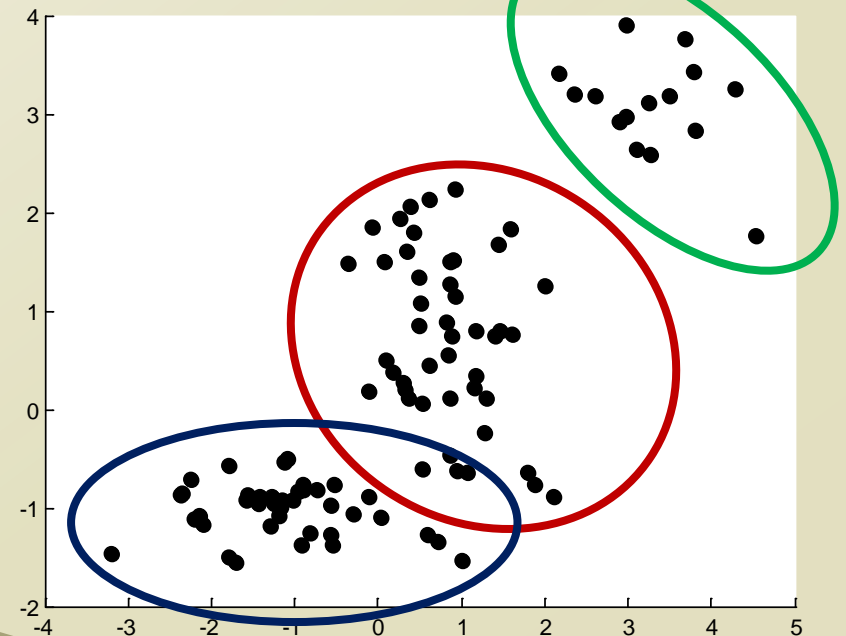
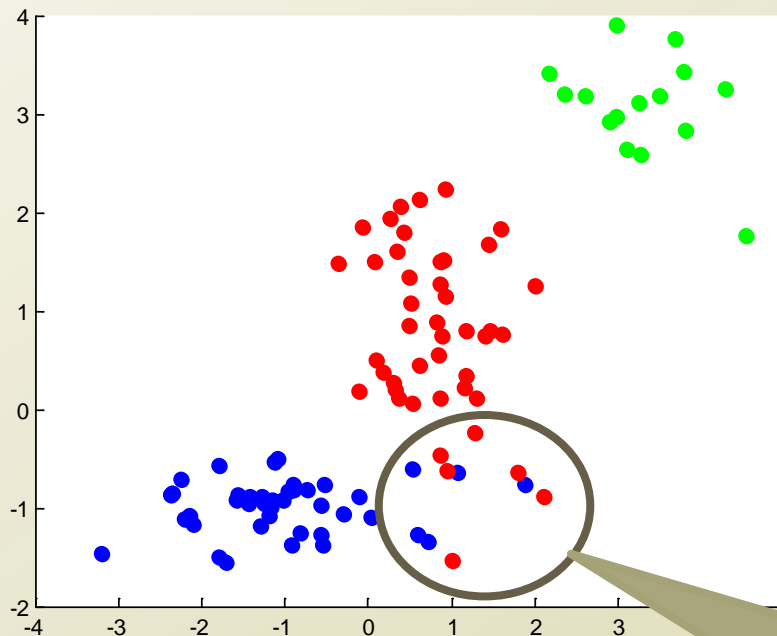
- **You don't know the true value, and you cannot provide examples of the true value.**
- Cases, such as
  - Discovering clusters
  - Discovering latent factors
  - Discovering graph structures
- Clustering or filtering or completing of
  - Finding **the representative topic words from text data**
  - Finding **the latent image from facial data**
  - Completing the incomplete **matrix of product-review scores**
  - Filtering the **noise from the trajectory data**
- Methodologies
  - Clustering: estimating sets and affiliations of instances to the sets
  - Filtering: estimating underlying and fundamental signals from the mixture of signals and noises



# Clustering Problem

- How to cluster the unlabeled data points?
  - No concrete knowledge of their classes
  - Latent (hidden) variable of classes
  - Optimal assignment to the latent classes

How to assign data points to classes?  
→ Clustering  
(here classes == clusters)

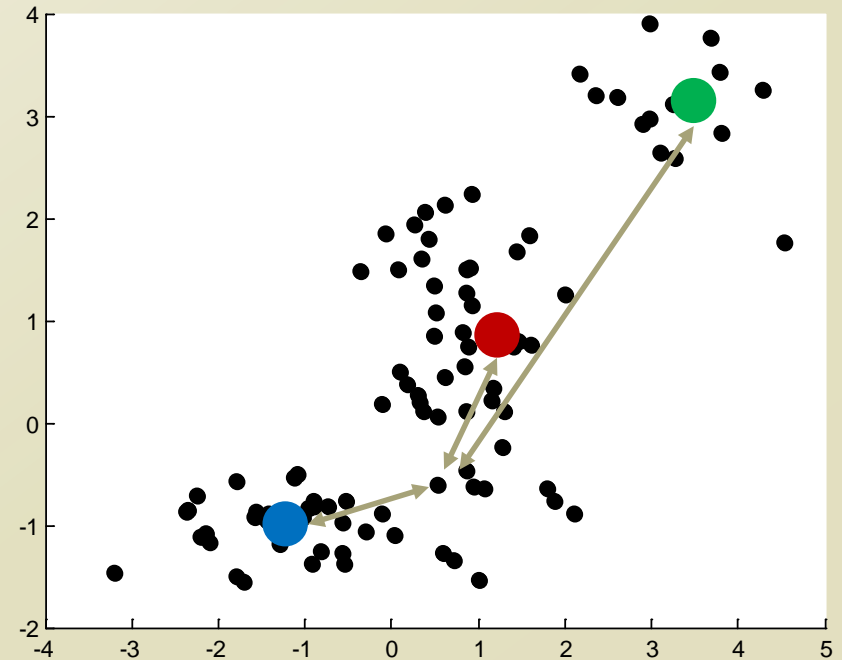
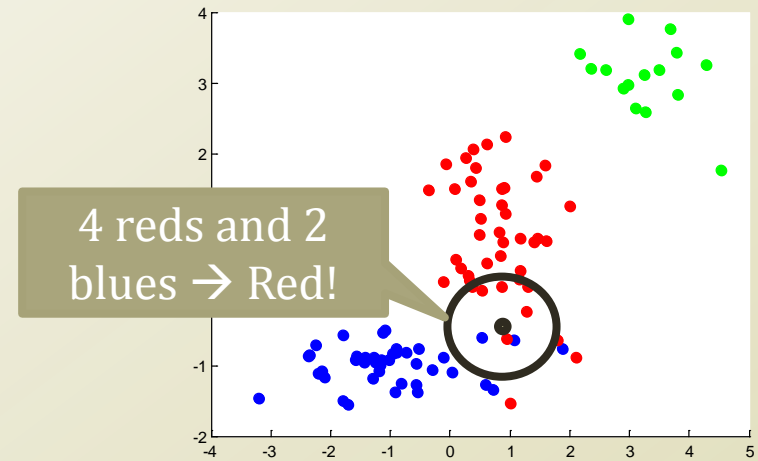


Uncertain area of clustering

# K-Means Algorithm

- K-Means algorithm
  - Setup K number of centroids (or prototypes) and cluster data points by the distance from the points to the nearest centroid
- Formally,
  - $J = \sum_{n=1}^N \sum_{k=1}^K r_{nk} ||x_n - \mu_k||^2$
  - Minimize J by optimizing
    - $r_{nk}$ : the assignment of data points to clusters
    - $\mu_k$ : the location of centroids
  - Iterative optimization
    - Why?
    - Two variables are interacting

K-Means! = K-Nearest Neighbor



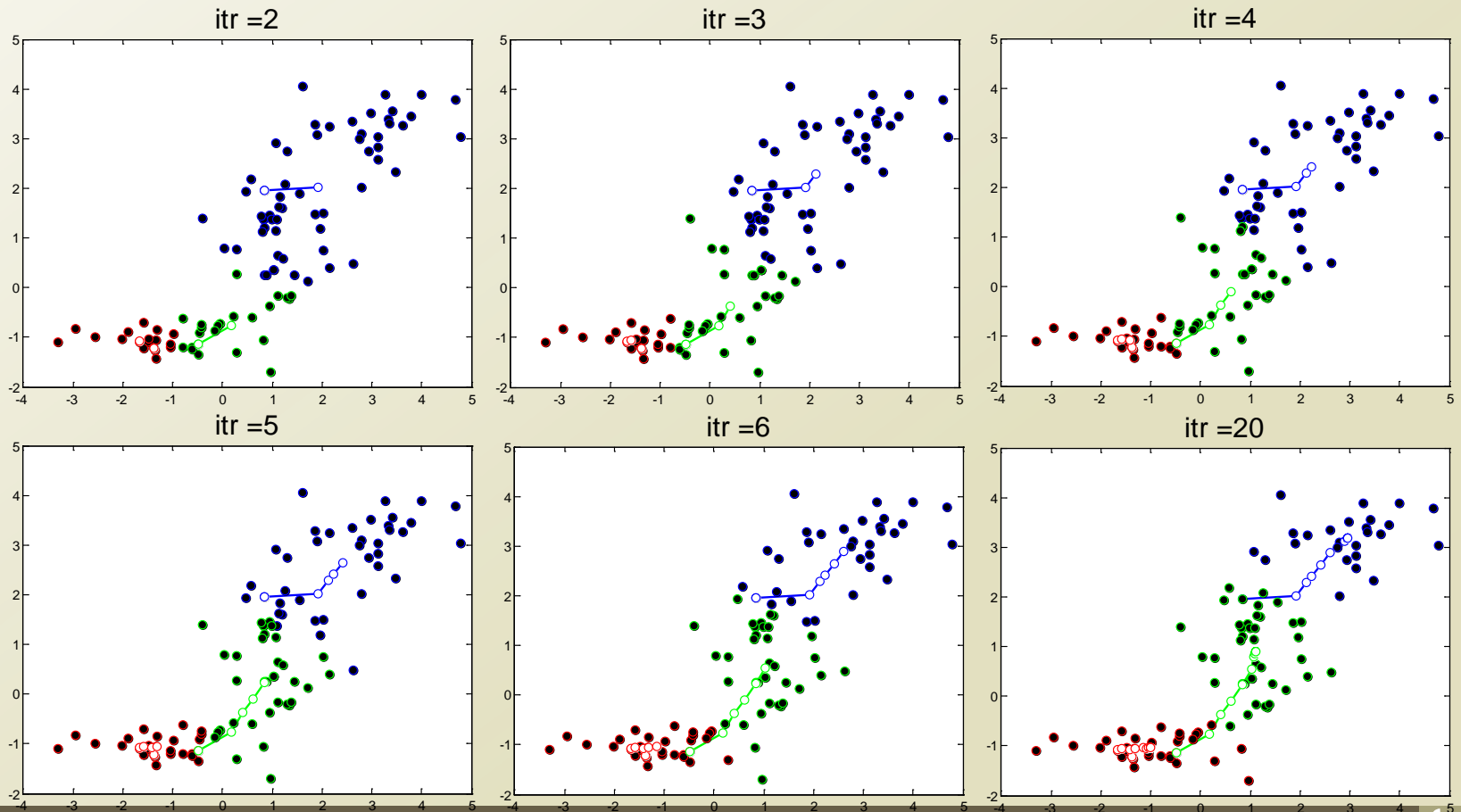
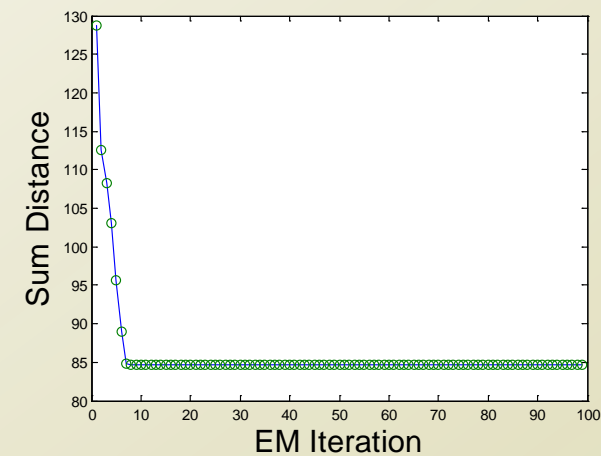
# Expectation and Maximization

- $J = \sum_{n=1}^N \sum_{k=1}^K r_{nk} ||x_n - \mu_k||^2$ 
  - Expectation
    - Expectation of the log-likelihood given the parameters
    - Assign the data points to the nearest centroid
  - Maximization
    - Maximization of the parameters with respect to the log-likelihood
    - Update the centroid positions given the assignments
- $r_{nk}$ 
  - $r_{nk} = \{0, 1\}$
  - Discrete variable
  - Logical choice: the nearest centroid  $\mu_k$  for a data point of  $x_n$
- $\mu_k$ 
  - $$\frac{dJ}{d\mu_k} = \frac{d}{d\mu_k} \sum_{n=1}^N \sum_{k=1}^K r_{nk} ||x_n - \mu_k||^2 = \frac{d}{d\mu_k} \sum_{n=1}^N r_{nk} ||x_n - \mu_k||^2 = \sum_{n=1}^N -2r_{nk}(x_n - \mu_k) = -2(-\sum_{n=1}^N r_{nk}\mu_k + \sum_{n=1}^N r_{nk}x_n) = 0$$
  - $$\mu_k = \frac{\sum_{n=1}^N r_{nk}x_n}{\sum_{n=1}^N r_{nk}}$$



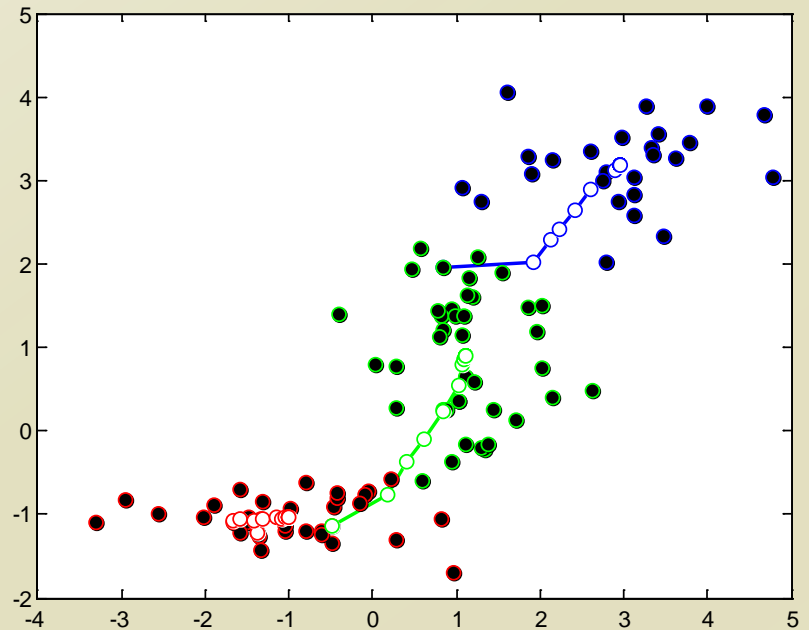
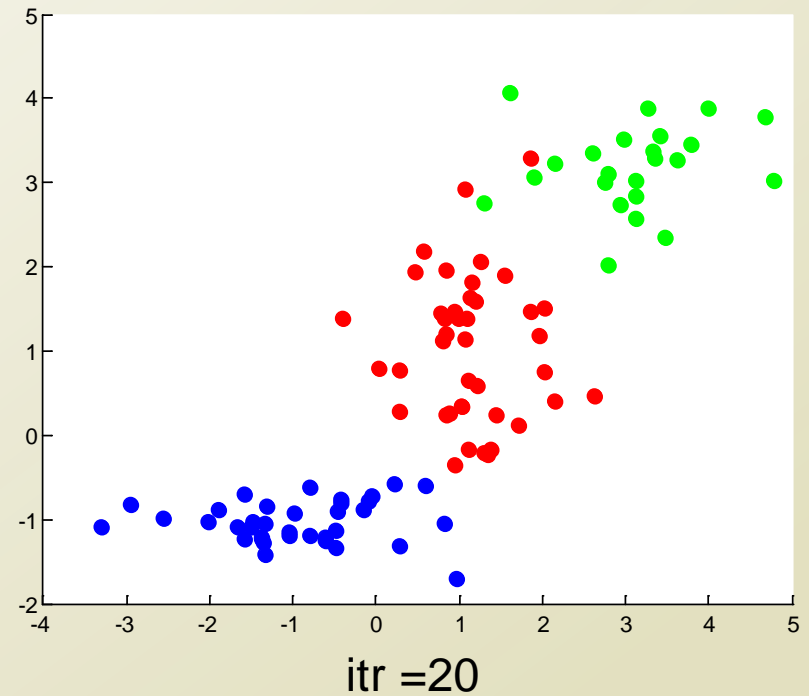
# Progress of K-Means Algorithm

- EM iterations to
  - Optimize the assignments with respect to the sum of distances
  - Optimize the parameters with respect to the sum of distances



# Properties of K-Means Algorithm

- # of clusters is uncertain
- Initial location of centroids
  - Some initial locations might not result in the reasonable results
- Limitation of distance metrics
  - Euclidean distance is very limited knowledge of information
- Hard clustering
  - Hard assignment of data points to clusters
    - $r_{nk} = \{0,1\}$ 
      - This can be the smoothly distributed probability
    - Any alternatives?
    - Soft clustering



# GAUSSIAN MIXTURE MODEL

# Multinomial Distribution

- Binary variable
  - Selecting 0 or 1  $\rightarrow$  binomial distribution
- How about K options?
  - $X=(0,0,1,0,0,0)$  when  $K=6$  and selecting the third option
  - $\sum_k x_k = 1, P(X|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{x_k}$  such that  $\mu_k \geq 0, \sum_k \mu_k = 1$
  - A generalization of binomial distribution  $\rightarrow$  Multinomial distribution
- Given a dataset D with N selections,  $x_1 \dots x_n$ 
  - $P(X|\boldsymbol{\mu}) = \prod_{n=1}^N \prod_{k=1}^K \mu_k^{x_{nk}} = \prod_{k=1}^K \mu_k^{\sum_{n=1}^N x_{nk}} = \prod_{k=1}^K \mu_k^{m_k}$ 
    - When  $m_k = \sum_{n=1}^N x_{nk}$
    - Number of selecting  $k^{\text{th}}$  option out of N selections
  - How to determine the maximum likelihood solution of  $\boldsymbol{\mu}$ ?
    - Maximize  $P(X|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{m_k}$
    - Subject to  $\mu_k \geq 0, \sum_k \mu_k = 1$

# Lagrange Method

Maximize  $P(X|\mu) = \prod_{k=1}^K \mu_k^{m_k}$   
Subject to  $\mu_k \geq 0, \sum_k \mu_k = 1$   
When  $m_k = \sum_{n=1}^N x_{nk}$

- Method of finding a local maximum subject to constraints
  - Maximize  $f(x,y)$
  - Subject to  $g(x,y)=c$
  - Assuming that  $f$  and  $g$  have continuous partial derivatives
  - 1) Lagrange function and multiplier (do you recall this?)
    - $L(x, y, \lambda) = f(x, y) + \lambda(g(x, y) - c)$
    - $L(\mu, m, \lambda) = \sum_{k=1}^K m_k \ln \mu_k + \lambda(\sum_{k=1}^K \mu_k - 1)$ 
      - Using the log likelihood
  - 2) Take the partial first-order derivative of variables, and set it to be zero
    - $\frac{d}{d\mu_k} L(\mu, m, \lambda) = \frac{m_k}{\mu_k} + \lambda = 0 \rightarrow \mu_k = -\frac{m_k}{\lambda}$
  - 3) Utilize the constraint to get the optimized value
    - $\sum_k \mu_k = 1 \rightarrow \sum_k -\frac{m_k}{\lambda} = 1 \rightarrow \sum_k m_k = -\lambda \rightarrow \sum_k \sum_{n=1}^N x_{nk} = -\lambda \rightarrow N = -\lambda$
    - $\mu_k = \frac{m_k}{N}$ : MLE parameter of multinomial distribution

# Multivariate Gaussian Distribution

- Probability density function of the Gaussian distribution

- $N(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{1}{2\sigma^2} (x - \mu)^2)$

- $N(x|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu))$

- $\ln N(x|\mu, \Sigma) = -\frac{1}{2} \ln |\Sigma| - \frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) + C$

- $\ln N(X|\mu, \Sigma) = -\frac{N}{2} \ln |\Sigma| - \frac{1}{2} \sum_{n=1}^N (x_n - \mu)^T \Sigma^{-1} (x_n - \mu) + C$

- $\propto -\frac{N}{2} \ln |\Sigma| - \frac{1}{2} \sum_{n=1}^N \text{Tr}[\Sigma^{-1} (x_n - \mu)(x_n - \mu)^T]$

- $= -\frac{N}{2} \ln |\Sigma| - \frac{1}{2} \text{Tr}[\Sigma^{-1} \sum_{n=1}^N ((x_n - \mu)(x_n - \mu)^T)]$

- $\frac{d}{d\mu} \ln N(X|\mu, \Sigma) = 0 \rightarrow -\frac{1}{2} \times 2 \times -1 \times \Sigma^{-1} \sum_{n=1}^N (x_n - \hat{\mu}) = 0 \rightarrow \hat{\mu} = \frac{\sum_{n=1}^N x_n}{N}$

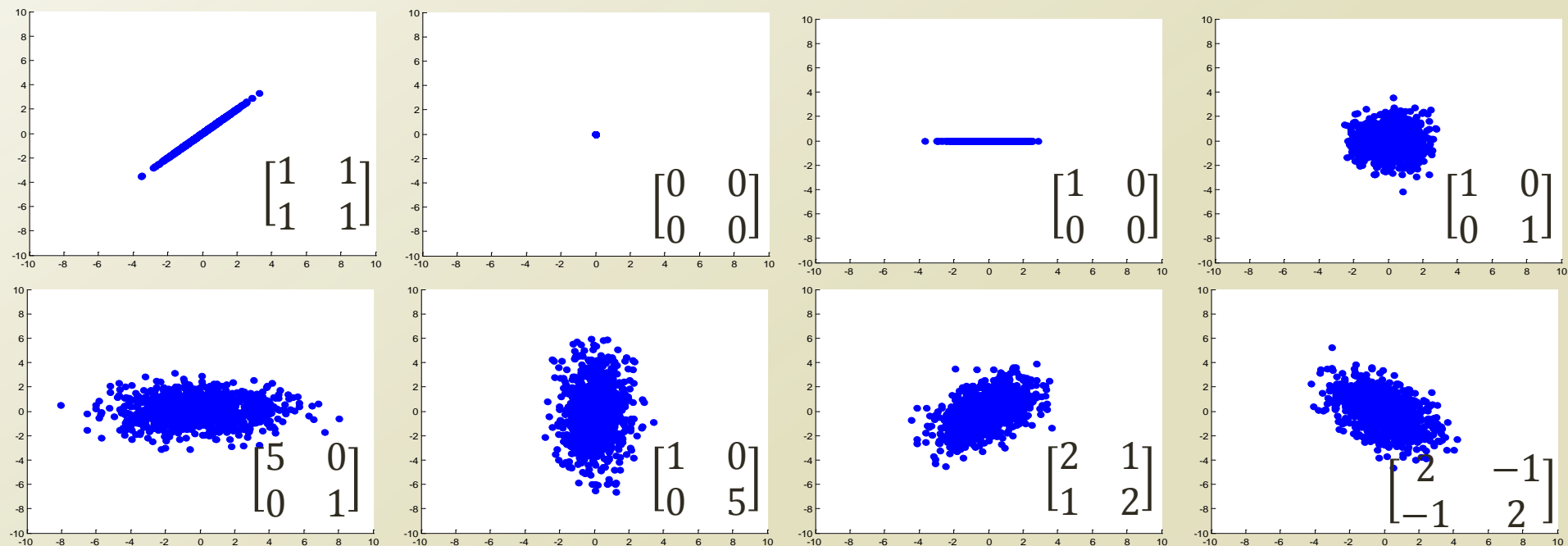
- $\frac{d}{d\Sigma^{-1}} \ln N(X|\mu, \Sigma) = 0 \rightarrow \hat{\Sigma} = \frac{1}{N} \sum_{n=1}^N (x_n - \hat{\mu})(x_n - \hat{\mu})^T$

- Beyond the scope of the course

- Use “trace trick” and 1)  $\frac{d}{dA} \log |A| = A^{-T}$ , 2)  $\frac{d}{dA} \text{Tr}[AB] = \frac{d}{dA} \text{Tr}[BA] = B^T$

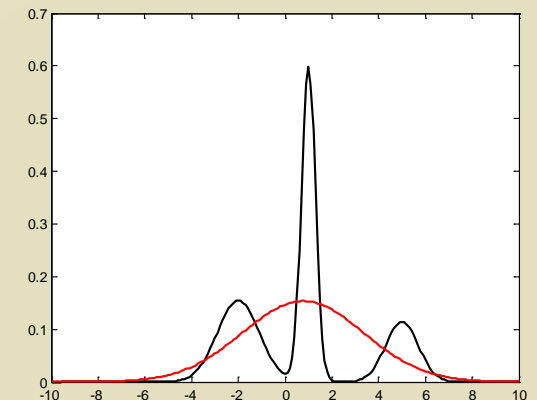
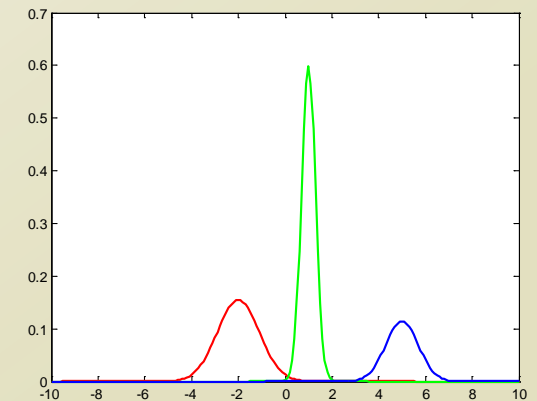
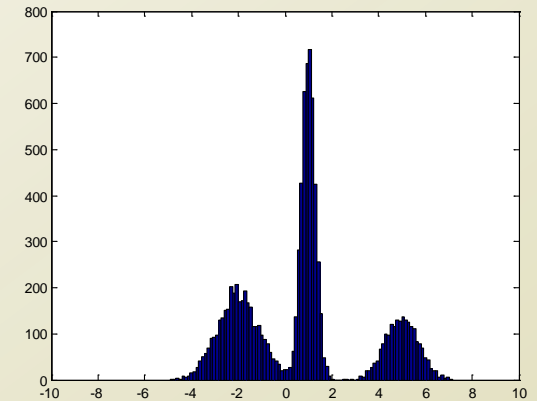
# Samples of Multivariate Gaussian Distribution

- Samples of multivariate Gaussian distributions
  - With various covariance matrixes
  - Covariance matrix should a positive-definite matrix
    - $z^T \Sigma z > 0$  for every non-zero column vector  $z$
    - $\begin{bmatrix} a & b \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} a \\ b \end{bmatrix} = \begin{bmatrix} a & 0 \\ 0 & b \end{bmatrix} \begin{bmatrix} a \\ b \end{bmatrix} = a^2 + b^2 > 0$  when  $a, b$  are non-zero



# Mixture Model

- Imagine that the samples are drawn from three different normal distributions
  - Subpopulation
  - The conventional distributions cannot explain the distribution accurately
  - We need to mix the three normal distribution → Create a new distribution adapted to the samples
  - Mixture distribution
- $P(x) = \sum_{k=1}^K \pi_k N(x|\mu_k, \sigma_k)$ 
  - Mixing coefficients,  $\pi_k$ : A normal distribution is chosen out of K options with probability
    - Works as weighting
    - $\sum_{k=1}^K \pi_k = 1, 0 \leq \pi_k \leq 1$
    - This is a probability (as well as weighting!)
    - Then, which distribution?
    - New variable? Let's say Z!
  - Mixture component,  $N(x|\mu_k, \sigma_k)$ : A distribution for the subpopulation
- $P(x) = \sum_{k=1}^K P(z_k)P(x|z_k)$ 
  - Why this ordering of variables?





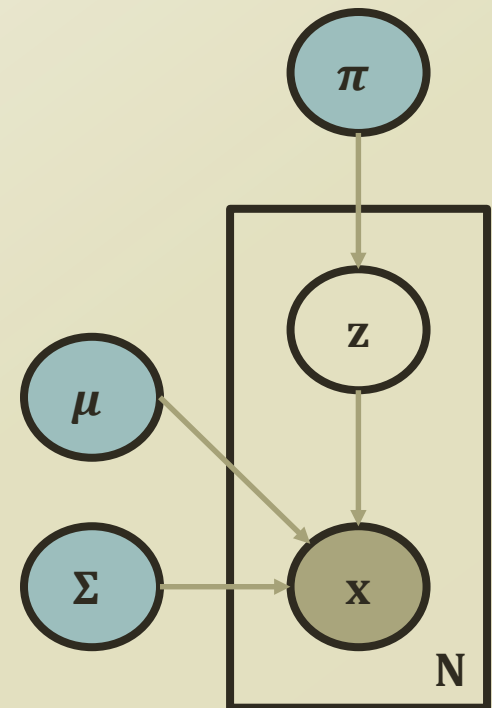
# Gaussian Mixture Model

- Let's assume that the data points are drawn from a mixture distribution of multiple multivariate Gaussian distributions

- $P(x) = \sum_{k=1}^K P(z_k)P(x|z_k) = \sum_{k=1}^K \pi_k N(x|\mu_k, \Sigma_k)$
- How to model such mixture?
  - Mixing coefficient, or Selection variable:  $z_k$ 
    - The selection is stochastic which follows the multinomial distribution
    - $z_k \in \{0,1\}, \sum_k z_k = 1, P(z_k = 1) = \pi_k, \sum_{k=1}^K \pi_k = 1, 0 \leq \pi_k \leq 1$
    - $P(Z) = \prod_{k=1}^K \pi_k^{z_k}$
  - Mixture component
    - $P(X|z_k = 1) = N(x|\mu_k, \Sigma_k) \rightarrow P(X|Z) = \prod_{k=1}^K N(x|\mu_k, \Sigma_k)^{z_k}$
- This is the marginalized probability. How about conditional?

$$\begin{aligned} \gamma(z_{nk}) &\equiv p(z_k = 1|x_n) = \frac{P(z_k=1)P(x|z_k = 1)}{\sum_{j=1}^K P(z_j=1)P(x|z_j = 1)} \\ &= \frac{\pi_k N(x|\mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j N(x|\mu_j, \Sigma_j)} \end{aligned}$$

- Log likelihood of the entire dataset is
  - $\ln P(X|\pi, \mu, \Sigma) = \sum_{n=1}^N \ln \{ \sum_{k=1}^K \pi_k N(x_n|\mu_k, \Sigma_k) \}$



# Expectation of GMM

- Similar problem of K-means algorithm
  - Two interacting parameters
  - As before, we apply the expectation and the maximization algorithm
    - Expectation: the assignment between the clusters and the data points
    - Maximization: the update of the parameters
- Expectation step
  - Assign a data point to a nearest cluster  $\rightarrow$  the assignment probability
    - Given the parameters and the data point, calculate the likelihood
  - $$\gamma(z_{nk}) \equiv p(z_k = 1 | x_n) = \frac{P(z_k=1)P(x|z_k = 1)}{\sum_{j=1}^K P(z_j=1)P(x|z_j = 1)} = \frac{\pi_k N(x|\mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j N(x|\mu_j, \Sigma_j)}$$
    - Here,  $x, \pi, \mu, \Sigma$  are given, calculate  $\gamma(z_{nk})$
    - $\gamma(z_{nk})$  are used to calculate  $\pi, \mu, \Sigma$
    - The new  $\gamma(z_{nk})$  motivates the update of the old parameters

# Maximization of GMM

$$N(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)$$

$$\ln N(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = -\frac{1}{2} \ln |\boldsymbol{\Sigma}| - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) + \mathcal{C}$$

$$\ln N(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = -\frac{N}{2} \ln |\boldsymbol{\Sigma}| - \frac{1}{2} \sum_{n=1}^N (\mathbf{x}_n - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x}_n - \boldsymbol{\mu}) + \mathcal{C}$$

$$\frac{d}{d\boldsymbol{\mu}} \ln N(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \rightarrow -\frac{1}{2} \times 2 \times -1 \times \boldsymbol{\Sigma}^{-1} \sum_{n=1}^N (\mathbf{x}_n - \hat{\boldsymbol{\mu}}) = 0 \rightarrow \hat{\boldsymbol{\mu}} = \frac{\sum_{n=1}^N \mathbf{x}_n}{N}$$

$$\frac{d}{d\boldsymbol{\Sigma}^{-1}} \ln N(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0 \rightarrow \hat{\boldsymbol{\Sigma}} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \hat{\boldsymbol{\mu}})(\mathbf{x}_n - \hat{\boldsymbol{\mu}})^T$$

- Maximization step

- Update the parameters given  $\gamma(z_{nk})$

- Parameters to update:  $\pi, \boldsymbol{\mu}, \boldsymbol{\Sigma}$

- $\ln P(\mathbf{X}|\pi, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$

- Typical methods

- Derivative  $\rightarrow$  set the equation to zero when the function is smooth

- Lagrange method when there is a constraint. Which parameter has the constraint?

$$\gamma(z_{nk}) = \frac{\pi_k N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

- $\frac{d}{d\boldsymbol{\mu}_k} \ln P(\mathbf{X}|\pi, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \frac{\pi_k N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \boldsymbol{\Sigma}^{-1}(\mathbf{x}_n - \hat{\boldsymbol{\mu}}_k) = 0$

$$\rightarrow \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \hat{\boldsymbol{\mu}}_k) = 0 \rightarrow \hat{\boldsymbol{\mu}}_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n}{\sum_{n=1}^N \gamma(z_{nk})}$$

- $\frac{d}{d\boldsymbol{\Sigma}_k} \ln P(\mathbf{X}|\pi, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = 0$

$$\rightarrow \boldsymbol{\Sigma}_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \hat{\boldsymbol{\mu}}_k)(\mathbf{x}_n - \hat{\boldsymbol{\mu}}_k)^T}{\sum_{n=1}^N \gamma(z_{nk})}$$

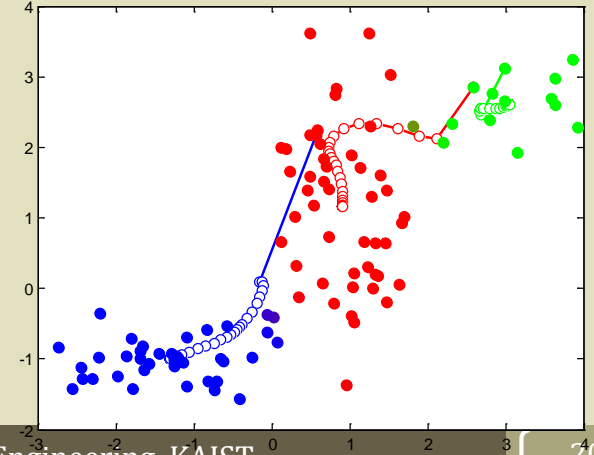
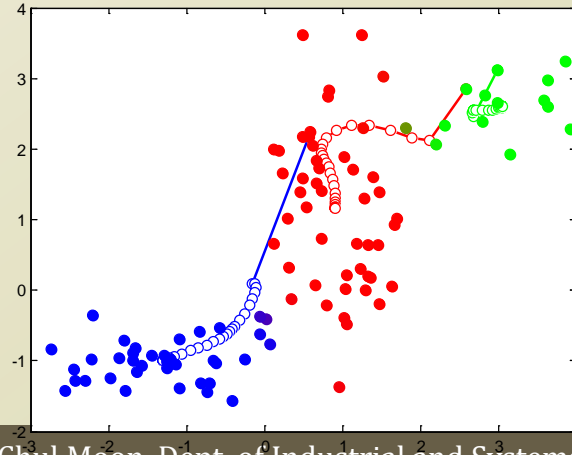
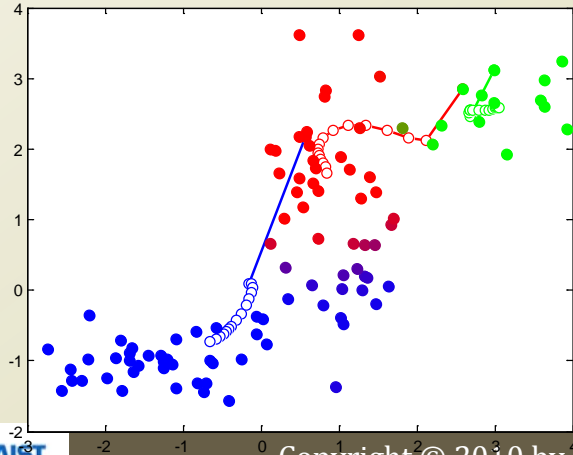
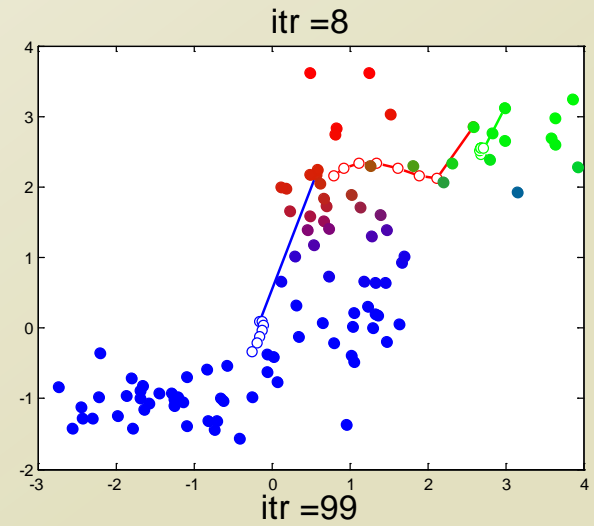
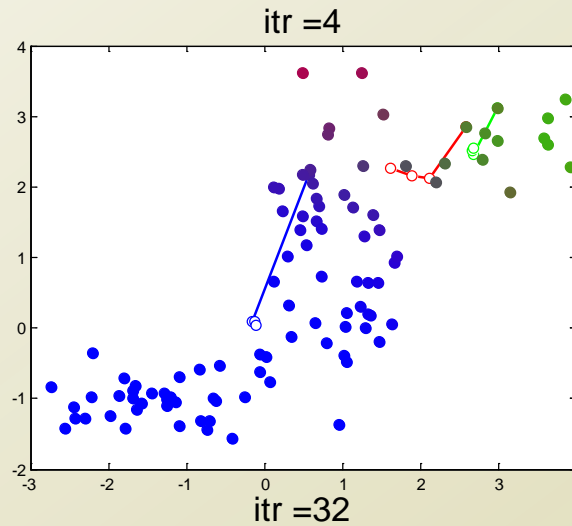
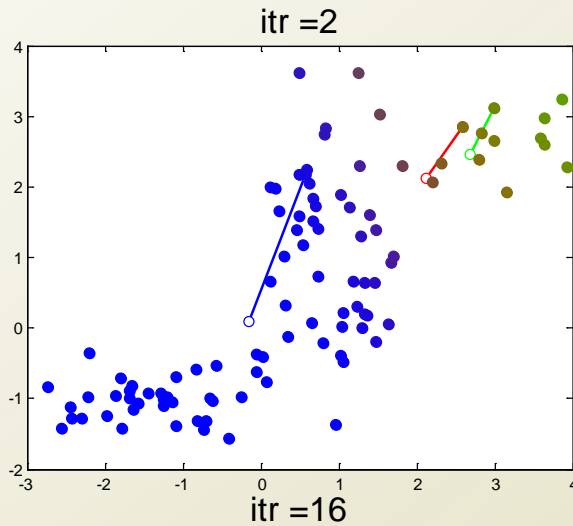
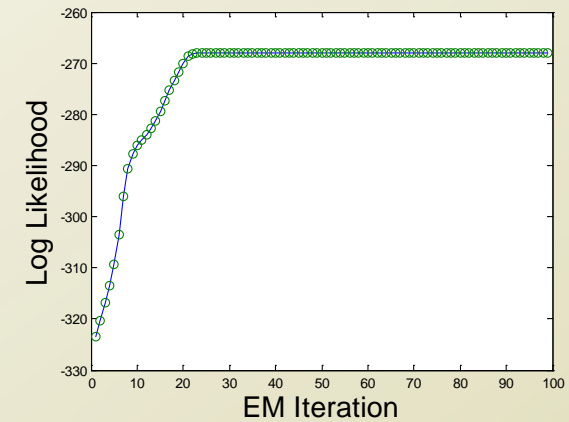
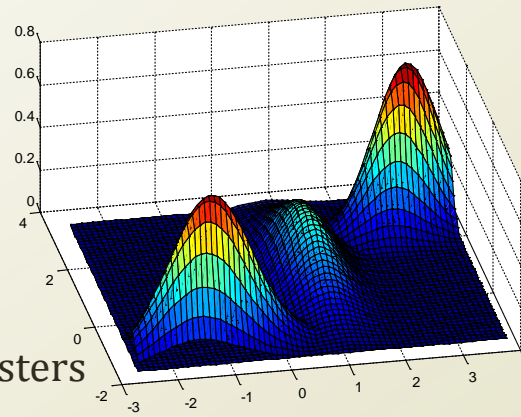
- $\frac{d}{d\pi_k} \ln P(\mathbf{X}|\pi, \boldsymbol{\mu}, \boldsymbol{\Sigma}) + \lambda (\sum_{k=1}^K \pi_k - 1) = 0$

$$\rightarrow \sum_{n=1}^N \frac{N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} + \lambda = 0 \rightarrow \sum_{k=1}^K \left\{ \sum_{n=1}^N \frac{\pi_k N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} + \pi_k \lambda \right\} = 0$$

$$\rightarrow \lambda = -N \rightarrow \pi_k = \frac{\sum_{n=1}^N \gamma(z_{nk})}{N}$$

# Progress of GMM

- Soft clustering
  - Estimated parameters
  - Soft assignment of data points to clusters



# Properties of GMM

- Pros and cons of Gaussian mixture model

- Pros

- More information

- Soft clustering
- Not a simple and discrete assignment
  - Information loss

- More and more information

- Learn the latent distribution
- Distance is not always the answer of the distribution

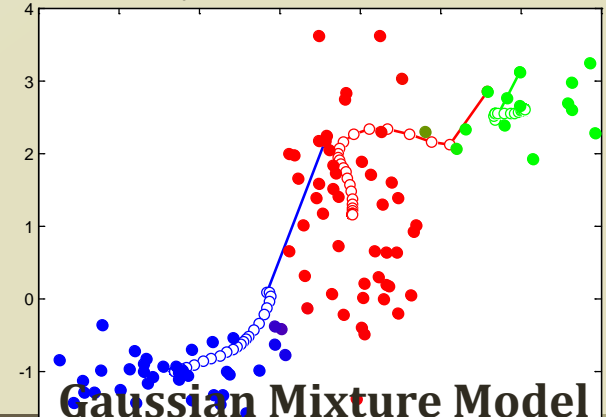
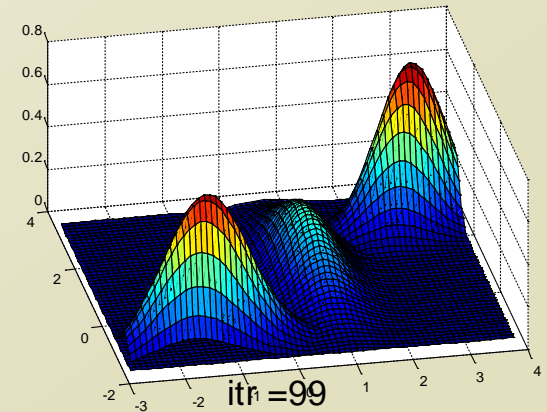
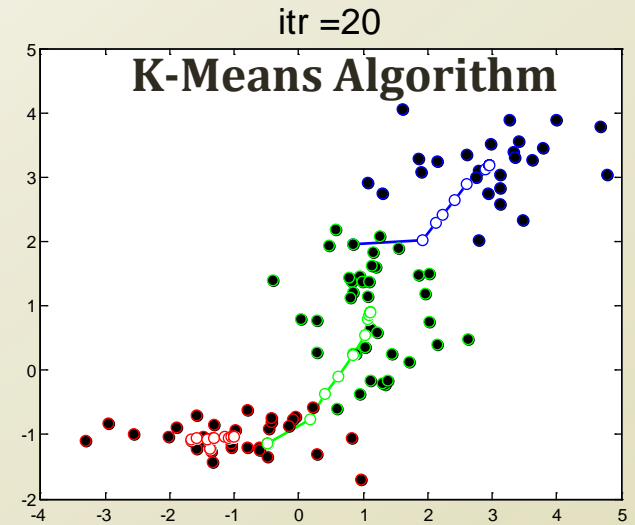
- Cons

- Long computation time

- Why?
- Falling into local maximum
- Deciding K

- Anyways to mitigate the disadvantage?

- Fast K-means and slow GMM



# Relation between K-Means and GMM

- $N(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}))$
- $P(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}_k|^{1/2}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_k^{-1}(\mathbf{x} - \boldsymbol{\mu}_k))$ 
  - Let's say  $\boldsymbol{\Sigma}_k = \epsilon \mathbf{I}$ 
    - Here,  $\mathbf{I}$  is the identity matrix and  $\epsilon$  is not updated by the EM process
    - $\mathbf{I} = \mathbf{I}^{-1}$

- $= \frac{1}{(2\pi)^{D/2} \epsilon^{1/2}} \exp(-\frac{1}{2\epsilon}(\mathbf{x} - \boldsymbol{\mu}_k)^T(\mathbf{x} - \boldsymbol{\mu}_k))$

- $= \frac{1}{(2\pi)^{D/2} \epsilon^{1/2}} \exp(-\frac{1}{2\epsilon} \|\mathbf{x} - \boldsymbol{\mu}_k\|^2)$

$$J = \sum_{n=1}^N \sum_{k=1}^K r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

## K-Means Algorithm

- $\gamma(z_{nk}) = \frac{\pi_k N(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j N(\mathbf{x}|\boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} = \frac{\pi_k \exp(-\frac{1}{2\epsilon} \|\mathbf{x} - \boldsymbol{\mu}_k\|^2)}{\sum_{j=1}^K \pi_j \exp(-\frac{1}{2\epsilon} \|\mathbf{x} - \boldsymbol{\mu}_j\|^2)}$ 
  - When  $\epsilon \rightarrow 0$ , the term of smallest  $\|\mathbf{x} - \boldsymbol{\mu}_k\|^2$  approaches zero most slowly
  - When all other terms are zero, the term of the smallest  $\|\mathbf{x} - \boldsymbol{\mu}_k\|^2$  has a value
  - Now, it becomes the hard assignment
- Still, GMM with  $\epsilon \mathbf{I}$  is not K-Means. Why?
  - Soft assignment + Covariance matrix learning

# EM ALGORITHM



# Inference with Latent Variables

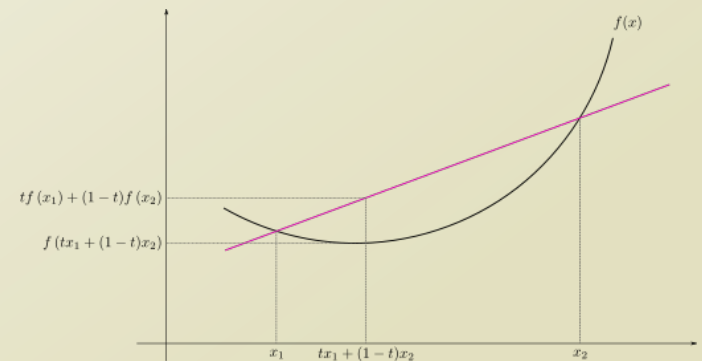
- Difference between classification and clustering
- Let's say
  - $\{X, Z\}$ : complete set of variables
  - $X$ : observed variables
  - $Z$ : hidden (latent) variables
  - $\theta$ : parameters for distributions
  - $P(X|\theta) = \sum_Z P(X, Z|\theta) \rightarrow \ln P(X|\theta) = \ln\{\sum_Z P(X, Z|\theta)\}$ 
    - Any problem here?
    - The locations of summation and log make this complicated
    - Eventually, we want to exchange the locations of the two operators
- What we want to know is
  - The values of  $Z$  and  $\theta$ 
    - Optimizing  $P(X|\theta) = \sum_Z P(X, Z|\theta)$
  - The interacting terms for the optimization



# Probability Decomposition

- $l(\theta) = \ln P(X|\theta) = \ln\{\sum_Z P(X, Z|\theta)\} = \ln\{\sum_Z q(Z) \frac{P(X, Z|\theta)}{q(Z)}\}$ 
  - Use the Jensen's inequality
  - $\ln\{\sum_Z q(Z) \frac{P(X, Z|\theta)}{q(Z)}\} \geq \sum_Z q(Z) \ln \frac{P(X, Z|\theta)}{q(Z)}$
- $= \sum_Z q(Z) \ln P(X, Z|\theta) - q(Z) \ln q(Z)$ 
  - Recall the second term?
  - $H(X) = -\sum_X P(X = x) \log_b P(X = x)$
- $= E_{q(Z)} \ln P(X, Z|\theta) + H(q)$ 
  - $Q(\theta, q) = E_{q(Z)} \ln P(X, Z|\theta) + H(q)$
  - This hold for any distribution of  $q$
  - This is only the lower bound of  $l(\theta)$ 
    - Need to make it tight!
    - How to?

## Jensen's Inequality



When  $\varphi(x)$  is concave

$$\varphi\left(\frac{\sum a_i x_i}{\sum a_j}\right) \geq \frac{\sum a_i \varphi(x_i)}{\sum a_j}$$

When  $\varphi(x)$  is convex

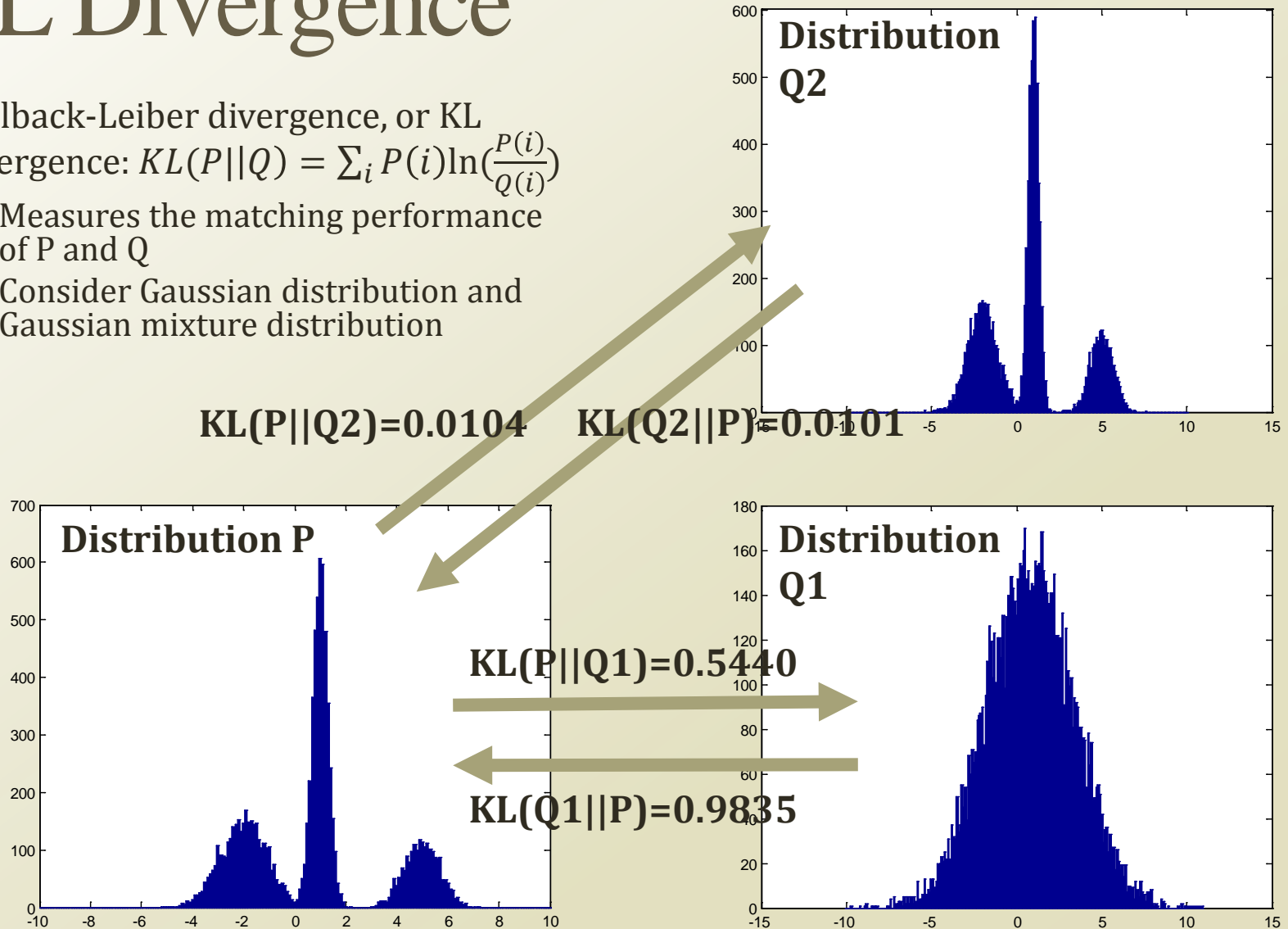
$$\varphi\left(\frac{\sum a_i x_i}{\sum a_j}\right) \leq \frac{\sum a_i \varphi(x_i)}{\sum a_j}$$

# Maximizing the Lower Bound (1)

- $l(\theta) = \ln P(X|\theta) = \ln \left\{ \sum_Z q(Z) \frac{P(X, Z|\theta)}{q(Z)} \right\} \geq \sum_Z q(Z) \ln \frac{P(X, Z|\theta)}{q(Z)} = Q(\theta, q)$ 
  - $Q(\theta, q) = E_{q(Z)} \ln P(X, Z|\theta) + H(q)$
- The other storyline is
  - $l(\theta) \geq \sum_Z q(Z) \ln \frac{P(X, Z|\theta)}{q(Z)} = \sum_Z q(Z) \ln \frac{P(Z|X, \theta) P(X|\theta)}{q(Z)}$
  - $= \sum_Z \{ q(Z) \ln \frac{P(Z|X, \theta)}{q(Z)} + q(Z) \ln P(X|\theta) \} = \ln P(X|\theta) + \sum_Z \{ q(Z) \ln \frac{P(Z|X, \theta)}{q(Z)} \}$
  - $L(\theta, q) = \ln P(X|\theta) - \sum_Z \{ q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)} \}$
- Here, the second term is a very special term
  - $KL(q(Z) || P(Z|X, \theta)) = \sum_Z \{ q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)} \}$
  - Kullback-Leiber divergence, or KL divergence:  $KL(P || Q) = \sum_i P(i) \ln \left( \frac{P(i)}{Q(i)} \right)$
  - Non-symmetric measure of the difference between two probability distributions, or  $KL(P || Q)$
  - Measures the difference
    - $KL(P || Q) \geq 0$
    - When there is no difference between P and Q,  $KL(P || Q) = 0$

# KL Divergence

- Kullback-Leiber divergence, or KL divergence:  $KL(P||Q) = \sum_i P(i) \ln\left(\frac{P(i)}{Q(i)}\right)$ 
  - Measures the matching performance of P and Q
  - Consider Gaussian distribution and Gaussian mixture distribution



# Maximizing the Lower Bound (2)

- $l(\theta) = \ln P(X|\theta) = \ln \left\{ \sum_Z q(Z) \frac{P(X, Z|\theta)}{q(Z)} \right\} \geq \sum_Z q(Z) \ln \frac{P(X, Z|\theta)}{q(Z)} = Q(\theta, q)$ 
  - $Q(\theta, q) = E_{q(Z)} \ln P(X, Z|\theta) + H(q)$
  - $L(\theta, q) = \ln P(X|\theta) - \sum_Z \{q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)}\}$
- Why do we compute  $L(\theta, q)$ ?
  - We do not know how to optimize  $Q(\theta, q)$  without further knowledge of  $q(Z)$
  - The second term of  $L(\theta, q)$  tells how to set  $q(Z)$ 
    - The first term is fixed when  $\theta$  is fixed **at time  $t$**
    - The second term can be minimized to maximize  $L(\theta, q)$ 
      - $KL(q(Z)||P(Z|X, \theta)) = 0 \rightarrow q^t(Z) = P(Z|X, \theta^t)$
  - Now, the lower bound with optimized  $q$  is
    - $Q(\theta, q^t) = E_{q^t(Z)} \ln P(X, Z|\theta^t) + H(q^t)$
- Then, optimizing  $\theta$  to retrieve the tight lower bound is
  - $\theta^{t+1} = \operatorname{argmax}_{\theta} Q(\theta, q^t) = \operatorname{argmax}_{\theta} E_{q^t(Z)} \ln P(X, Z|\theta)$ 
    - $q^t(Z) \rightarrow$  Distribution parameters for latent variable is at time  $t$
    - $\ln P(X, Z|\theta) \rightarrow$  optimized log likelihood parameters is at time  $t+1$

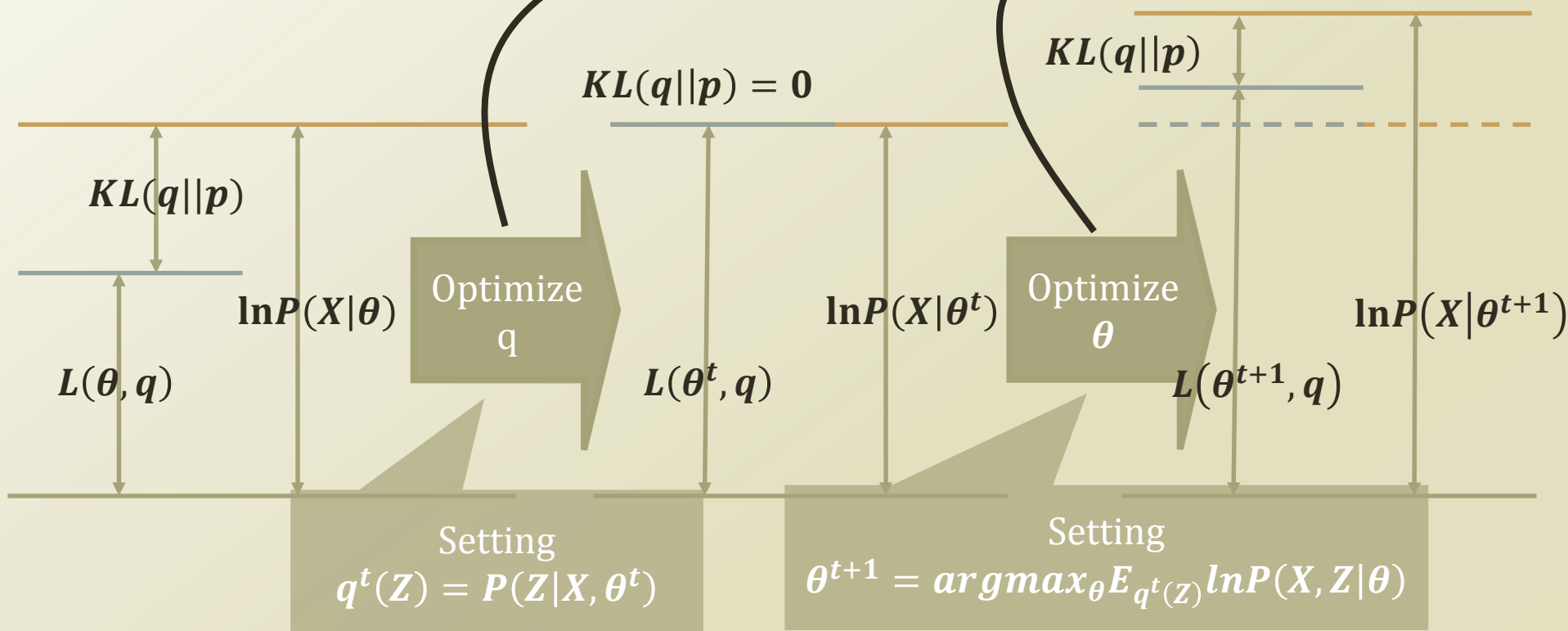
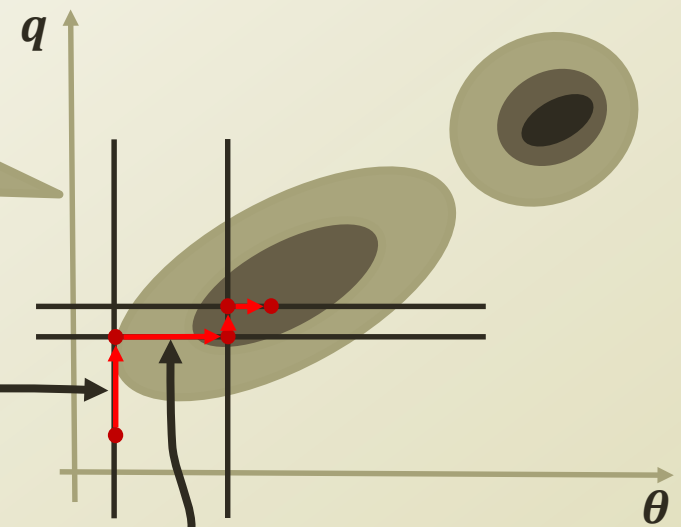
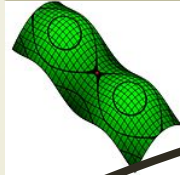
**Tells how to setup Z  
by setting  $q^t(Z) = P(Z|X, \theta^t)$**

**Relax the KL divergence by  
updating  $\theta^t$  to  $\theta^{t+1}$**

# Graphical Interpretation of Lower Bound Maximization

- $l(\theta) = \ln P(X|\theta) \geq L(\theta, q)$   
 $= \ln P(X|\theta) - \sum_Z \left\{ q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)} \right\}$
- $\ln P(X|\theta) = L(\theta, q) + \sum_Z \left\{ q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)} \right\}$   
 $= L(\theta, q) + KL(q||p)$

Fall into a local maxima or ???



# EM Algorithm

$$l(\theta) = \ln P(X|\theta) = \ln \left\{ \sum_Z q(Z) \frac{P(X, Z|\theta)}{q(Z)} \right\} \geq \sum_Z q(Z) \ln \frac{P(X, Z|\theta)}{q(Z)} = Q(\theta, q)$$
$$Q(\theta, q) = E_{q(Z)} \ln P(X, Z|\theta) + H(q)$$
$$L(\theta, q) = \ln P(X|\theta) - \sum_Z \{q(Z) \ln \frac{q(Z)}{P(Z|X, \theta)}\}$$

- EM algorithm
  - Finds the maximum likelihood solutions for models with latent variables
  - $P(X|\theta) = \sum_Z P(X, Z|\theta) \rightarrow \ln P(X|\theta) = \ln \{\sum_Z P(X, Z|\theta)\}$
- EM algorithm
  - Initialize  $\theta^0$  to an arbitrary point
  - Loop until the likelihood converges
    - Expectation step
      - $q^{t+1}(z) = \operatorname{argmax}_q Q(\theta^t, q) = \operatorname{argmax}_q L(\theta^t, q) = \operatorname{argmin}_q KL(q || P(Z|X, \theta^t))$
      - $\rightarrow q^t(z) = P(Z|X, \theta) \rightarrow$  Assign Z by  $P(Z|X, \theta)$
    - Maximization step
      - $\theta^{t+1} = \operatorname{argmax}_\theta Q(\theta, q^{t+1}) = \operatorname{argmax}_\theta L(\theta, q^{t+1})$
      - $\rightarrow$  fixed Z means that there is no unobserved variables
      - $\rightarrow$  Same optimization of ordinary MLE

# Rethinking GMM Learning Process

- GMM, K-Means
  - We used EM algorithm to find the assignment of latent variables and the related distribution parameters
- EM algorithm
  - Initialize  $\theta^0$  to an arbitrary point
  - Loop until the likelihood converges
    - Expectation step
      - Assign Z by  $P(Z|X, \theta)$
      - $\gamma(z_{nk}) \equiv p(z_k = 1|x_n) = \frac{P(z_k=1)P(x|z_k = 1)}{\sum_{j=1}^K P(z_j=1)P(x|z_j = 1)} = \frac{\pi_k N(x|\mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j N(x|\mu_j, \Sigma_j)}$
    - Maximization step
      - Same optimization of ordinary MLE
      - $\frac{d}{d\mu_k} \ln P(X|\pi, \mu, \Sigma) = 0, \frac{d}{d\Sigma_k} \ln P(X|\pi, \mu, \Sigma) = 0, \frac{d}{d\pi_k} \ln P(X|\pi, \mu, \Sigma) + \lambda(\sum_{k=1}^K \pi_k - 1) = 0$
      - $\widehat{\mu}_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) x_n}{\sum_{n=1}^N \gamma(z_{nk})}, \widehat{\Sigma}_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) (x_n - \widehat{\mu}_k)(x_n - \widehat{\mu}_k)^T}{\sum_{n=1}^N \gamma(z_{nk})}, \pi_k = \frac{\sum_{n=1}^N \gamma(z_{nk})}{N}$



# Further Readings

- Bishop Chapter 2 and 9
- Murphy Chapter 11