

# The Effect of DACA Eligibility on Full-Time Employment: A Difference-in-Differences Analysis

Independent Replication Study

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## Abstract

This study examines the causal effect of eligibility for the Deferred Action for Childhood Arrivals (DACA) program on full-time employment among Hispanic-Mexican individuals born in Mexico. Using American Community Survey data from 2006–2016 and a difference-in-differences identification strategy, I find that DACA eligibility is associated with a 3.6 percentage point increase in the probability of working full-time (35 or more hours per week). This effect is statistically significant at conventional levels and robust to alternative specifications. The findings suggest that DACA’s provision of work authorization had meaningful positive effects on the labor market outcomes of eligible immigrants.

**Keywords:** DACA, immigration policy, employment, difference-in-differences

**JEL Codes:** J15, J22, J61, K37

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# 1 Introduction

The Deferred Action for Childhood Arrivals (DACA) program, implemented on June 15, 2012, represents one of the most significant immigration policy changes in recent U.S. history. The program provides temporary relief from deportation and work authorization to qualifying undocumented immigrants who arrived in the United States as children. Given that DACA grants recipients the legal right to work, a natural question is whether the program has affected employment outcomes among eligible individuals.

This study examines the causal effect of DACA eligibility on full-time employment, defined as usually working 35 or more hours per week, among Hispanic-Mexican individuals born in Mexico. Using a difference-in-differences (DiD) identification strategy that compares changes in employment outcomes between DACA-eligible and non-eligible Mexican immigrants before and after the program’s implementation, I estimate that DACA eligibility increased the probability of full-time employment by approximately 3.6 percentage points.

The analysis leverages detailed individual-level data from the American Community Survey (ACS) for the years 2006–2016, encompassing approximately 565,000 observations of non-citizen Mexican-born individuals of working age. The DiD design exploits variation in eligibility status generated by the program’s age-based criteria while controlling for common time trends affecting all Mexican immigrants.

The remainder of this paper is organized as follows. Section 2 provides background on the DACA program and its eligibility requirements. Section 3 describes the data and sample construction. Section 4 presents the empirical methodology. Section 5 reports the main results and robustness checks. Section 6 discusses limitations and alternative interpretations. Section 7 concludes.

## 2 Background on DACA

### 2.1 Program Overview

DACA was announced by the Department of Homeland Security on June 15, 2012, and applications began to be accepted on August 15, 2012. The program allows qualifying individuals to apply for deferred action status, which provides temporary protection from deportation for a renewable two-year period. Critically for labor market outcomes, DACA recipients become eligible to obtain work authorization through an Employment Authorization Document (EAD).

In the first four years following implementation, the U.S. Citizenship and Immigration Services received nearly 900,000 initial applications, with approximately 90% approved. The

high approval rate suggests that most applicants who believed themselves eligible were indeed found to meet the program's requirements.

## 2.2 Eligibility Criteria

To qualify for DACA, individuals must meet several criteria:

1. **Arrival before age 16:** The applicant must have arrived in the United States before their 16th birthday.
2. **Age requirement:** The applicant must not have reached their 31st birthday as of June 15, 2012 (i.e., born after June 15, 1981).
3. **Continuous presence:** The applicant must have lived continuously in the United States since June 15, 2007.
4. **Physical presence:** The applicant must have been physically present in the United States on June 15, 2012.
5. **Immigration status:** The applicant must not have had lawful immigration status (citizenship or legal permanent residence) on June 15, 2012.
6. **Education/military:** The applicant must have been in school, graduated from high school, obtained a GED, or been honorably discharged from the military.
7. **Criminal record:** The applicant must not have been convicted of a felony, significant misdemeanor, or three or more other misdemeanors.

While the program was not restricted by country of origin, the structure of undocumented immigration to the United States meant that the great majority of eligible individuals were from Mexico. This makes Mexican-born individuals a natural focus for studying DACA's effects.

## 2.3 Expected Effects on Employment

Prior to DACA, undocumented immigrants faced significant legal barriers to formal employment. Without work authorization, they were limited to informal sector jobs or employment with employers willing to overlook documentation requirements. This restriction likely constrained both the quantity and quality of employment opportunities available.

DACA's provision of work authorization could affect employment through several channels:

- **Extensive margin:** Previously unemployed or out-of-labor-force individuals may enter employment.
- **Intensive margin:** Those already working may increase hours, potentially moving from part-time to full-time status.
- **Job quality:** Workers may transition from informal to formal employment or from jobs with irregular hours to those with stable full-time schedules.
- **Complementary benefits:** In some states, DACA recipients became eligible for driver's licenses, which could facilitate job access and employment.

This study focuses on full-time employment as the primary outcome, capturing effects on both the extensive and intensive margins of labor supply.

## 3 Data

### 3.1 Data Source

The analysis uses data from the American Community Survey (ACS) as provided by IPUMS USA. The ACS is an annual survey conducted by the U.S. Census Bureau that collects detailed demographic, social, economic, and housing information from approximately 3.5 million households per year. The survey provides large sample sizes suitable for studying relatively small population subgroups such as Mexican-born non-citizens.

I use the one-year ACS samples for 2006 through 2016, excluding the 2012 sample. The exclusion of 2012 is necessary because DACA was implemented mid-year (June 15, 2012), and the ACS does not report the month of interview, making it impossible to distinguish pre- and post-treatment observations within that year.

### 3.2 Sample Construction

The analysis sample is constructed through the following steps:

1. **Ethnic and birthplace restriction:** I restrict the sample to individuals of Hispanic-Mexican ethnicity (HISPAN = 1) who were born in Mexico (BPL = 200). This focuses on the population most affected by DACA.
2. **Year restriction:** I exclude observations from 2012 due to the mid-year implementation of DACA.

3. **Age restriction:** I restrict to working-age individuals (ages 16–65).
4. **Citizenship restriction:** I restrict to non-citizens ( $\text{CITIZEN} = 3$ ). This is necessary because U.S. citizens and naturalized citizens are not eligible for DACA, and their inclusion would dilute the comparison groups.
5. **Valid immigration data:** I exclude observations with missing or invalid year of immigration.

Table 1 presents the sample construction process and resulting sample sizes at each stage.

Table 1: Sample Construction

Stage	Observations	Percent of Previous
Original ACS data (2006–2016)	33,851,424	—
Hispanic-Mexican, Mexican-born	991,261	2.9%
Excluding 2012	898,879	90.7%
Working age (16–65)	778,727	86.6%
Non-citizens only	564,667	72.5%
<b>Final analysis sample</b>	<b>564,667</b>	—

### 3.3 Variable Definitions

#### 3.3.1 Treatment: DACA Eligibility

I construct a measure of DACA eligibility based on the observable criteria in the ACS data. An individual is classified as DACA-eligible if they meet all of the following conditions:

- **Arrived before age 16:** Age at immigration ( $\text{YRIMMIG} - \text{BIRTHYR}$ ) is less than 16.
- **Under 31 as of June 15, 2012:** Birth year is after 1981 ( $\text{BIRTHYR} > 1981$ ).
- **In U.S. since June 15, 2007:** Year of immigration is 2007 or earlier ( $\text{YRIMMIG} \leq 2007$ ).
- **Non-citizen:** Citizenship status is “not a citizen” ( $\text{CITIZEN} = 3$ ).

Note that I cannot observe some eligibility criteria (education/military status, criminal record, continuous physical presence verification). Following the literature, I assume that individuals who meet the observable criteria and are non-citizens should be treated as potentially eligible. This may introduce some measurement error in the treatment indicator.

### 3.3.2 Outcome: Full-Time Employment

The outcome variable is full-time employment, defined as usually working 35 or more hours per week. This is constructed from the UHRSWORK variable, which reports the respondent's usual hours worked per week at all jobs. The outcome is coded as:

$$\text{Full-time} = \begin{cases} 1 & \text{if UHRSWORK} \geq 35 \\ 0 & \text{otherwise} \end{cases}$$

This definition follows the standard Bureau of Labor Statistics definition of full-time work.

### 3.3.3 Control Variables

I include the following control variables in the regression specifications:

- **Age:** Age in years, with a quadratic term to capture nonlinear lifecycle effects.
- **Sex:** Indicator for male.
- **Marital status:** Indicator for currently married (spouse present or absent).
- **Years in U.S.:** Calculated as survey year minus year of immigration, capturing immigrant assimilation effects.

## 3.4 Descriptive Statistics

Table 2 presents summary statistics for the analysis sample, stratified by DACA eligibility status and time period.

Table 2: Descriptive Statistics by Treatment Status and Period

	DACA Eligible		Not Eligible	
	Pre	Post	Pre	Post
Full-time employment	0.425	0.494	0.602	0.576
Age	20.95	24.12	38.26	41.96
Male	0.556	0.544	0.545	0.529
Married	0.216	0.299	0.655	0.652
Education (categorical)	4.98	5.37	3.92	3.97
Years in U.S.	12.69	16.77	15.04	18.52
Observations	45,433	36,075	302,048	181,111

*Notes: Statistics are*

*unweighted means. Pre-period: 2006–2011. Post-period: 2013–2016. Education is measured on an ordinal scale where higher values indicate more education.*

Several patterns are apparent from the descriptive statistics:

1. DACA-eligible individuals are substantially younger than non-eligible individuals, reflecting the age-based eligibility criteria.
2. DACA-eligible individuals have lower baseline rates of full-time employment (42.5% vs. 60.2% in the pre-period), consistent with their younger age profile.
3. Full-time employment increased among DACA-eligible individuals from 42.5% to 49.4% (+6.9 pp), while it decreased among non-eligible individuals from 60.2% to 57.6% (-2.6 pp).
4. The differential change in full-time employment—approximately 9.5 percentage points—provides a raw difference-in-differences estimate before regression adjustment.

## 4 Empirical Methodology

### 4.1 Identification Strategy

I employ a difference-in-differences (DiD) design to estimate the causal effect of DACA eligibility on full-time employment. The DiD approach compares changes in outcomes over time between a treatment group (DACA-eligible individuals) and a control group (non-eligible

individuals), under the identifying assumption that both groups would have experienced parallel trends in the absence of the treatment.

The treatment group consists of non-citizen Mexican-born individuals who meet the observable DACA eligibility criteria. The control group consists of non-citizen Mexican-born individuals who fail at least one of the eligibility criteria (typically because they arrived in the U.S. after age 16, were too old as of 2012, or arrived after 2007).

## 4.2 Econometric Specification

The main estimating equation is a linear probability model of the form:

$$\text{FullTime}_{ist} = \alpha + \beta \cdot \text{Eligible}_i + \gamma \cdot \text{Post}_t + \delta \cdot (\text{Eligible}_i \times \text{Post}_t) + X'_i \theta + \mu_s + \tau_t + \varepsilon_{ist} \quad (1)$$

where:

- $\text{FullTime}_{ist}$  is an indicator for whether individual  $i$  in state  $s$  at time  $t$  works 35+ hours per week
- $\text{Eligible}_i$  is an indicator for DACA eligibility
- $\text{Post}_t$  is an indicator for the post-treatment period (2013–2016)
- $X_i$  is a vector of individual controls (age, age<sup>2</sup>, male, married, years in U.S.)
- $\mu_s$  are state fixed effects
- $\tau_t$  are year fixed effects
- $\varepsilon_{ist}$  is the error term

The coefficient of interest is  $\delta$ , which captures the differential change in full-time employment between eligible and non-eligible individuals after DACA's implementation. Under the parallel trends assumption,  $\delta$  identifies the causal effect of DACA eligibility on full-time employment.

## 4.3 Estimation Details

I estimate Equation 1 using weighted least squares, with person weights (PERWT) provided by IPUMS to ensure population representativeness. Standard errors are computed using the

HC1 heteroskedasticity-robust estimator, which is appropriate for linear probability models with binary outcomes.

As a robustness check, I also report estimates with standard errors clustered at the state level, which allows for arbitrary within-state correlation in the error terms.

## 4.4 Identifying Assumptions

The validity of the DiD estimates rests on several assumptions:

1. **Parallel trends:** In the absence of DACA, full-time employment would have evolved similarly for eligible and non-eligible individuals.
2. **No anticipation:** Individuals did not change their behavior before DACA's announcement in anticipation of the policy.
3. **Stable unit treatment value assumption (SUTVA):** The treatment status of one individual does not affect the outcomes of others.
4. **Correct classification:** Individuals are correctly classified as eligible or non-eligible based on observable characteristics.

I assess the parallel trends assumption by examining pre-treatment trends in full-time employment across groups. Violations of this assumption would undermine the causal interpretation of the DiD estimates.

# 5 Results

## 5.1 Main Results

Table 3 presents the main regression results, building from a basic specification to the full model with controls and fixed effects.

Table 3: Effect of DACA Eligibility on Full-Time Employment

	(1)	(2)	(3)	(4)
	Basic DiD	+ Demographics	+ Year FE	+ State FE
DACA Eligible $\times$ Post	0.1009*** (0.0046)	0.0441*** (0.0043)	0.0366*** (0.0043)	0.0359*** (0.0043)
DACA Eligible	-0.1772*** (0.0033)	-0.0083** (0.0036)	-0.0096** (0.0036)	-0.0100** (0.0036)
Post	-0.0256*** (0.0019)	0.0189*** (0.0019)	—	—
Demographic controls	No	Yes	Yes	Yes
Year fixed effects	No	No	Yes	Yes
State fixed effects	No	No	No	Yes
Observations	564,667	564,667	564,667	564,667
R-squared	0.019	0.117	0.118	0.119

*Notes:* Dependent variable is an indicator for working 35+ hours per week. All specifications are estimated using weighted least squares with person weights (PERWT). Heteroskedasticity-robust standard errors (HC1) in parentheses. Demographic controls include age, age squared, male indicator, married indicator, and years in U.S. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.10$ .

The results in Table 3 reveal several important findings:

**Column (1)** shows the basic DiD estimate without controls. The coefficient on the interaction term (Eligible  $\times$  Post) is 0.101, suggesting that DACA eligibility is associated with a 10.1 percentage point increase in full-time employment. However, this estimate does not account for compositional differences between treatment and control groups.

**Column (2)** adds demographic controls (age, age squared, sex, marital status, years in U.S.). The coefficient drops substantially to 0.044, indicating that much of the raw difference was driven by demographic differences—particularly age—between eligible and non-eligible individuals. The addition of controls increases the R-squared from 0.019 to 0.117.

**Column (3)** adds year fixed effects, which absorb common macroeconomic shocks

affecting all Mexican immigrants. The coefficient is 0.037, somewhat smaller than with controls alone.

**Column (4)** is the preferred specification, adding state fixed effects to control for time-invariant state-level differences in labor markets. The DiD coefficient is **0.0359** (SE = 0.0043), indicating that DACA eligibility increased full-time employment by 3.6 percentage points. This effect is statistically significant at the 1% level.

## 5.2 Interpretation of the Main Effect

The preferred estimate of 3.6 percentage points can be interpreted as follows: relative to a baseline full-time employment rate of 42.5% among DACA-eligible individuals in the pre-period, the effect represents a relative increase of approximately 8.5% ( $0.036/0.425 \approx 0.085$ ).

The 95% confidence interval for the effect is [0.028, 0.044], allowing us to rule out effects smaller than 2.8 percentage points or larger than 4.4 percentage points at the 5% significance level.

## 5.3 Robustness Checks

### 5.3.1 Alternative Standard Errors

Table 4 presents the main specification with alternative standard error estimators.

Table 4: Robustness to Standard Error Specification

	Coefficient	95% CI	
HC1 (robust) standard errors	0.0359 (0.0043)	[0.0276, 0.0442]	<i>Notes: Both rows report</i>
State-clustered standard errors	0.0359 (0.0042)	[0.0278, 0.0441]	

*estimates from the full specification (Model 4) with state and year fixed effects. Standard errors in parentheses.*

The results are essentially unchanged when using state-clustered standard errors, with nearly identical confidence intervals. This suggests that within-state correlation in the error terms is not a major concern.

### 5.3.2 Pre-Trend Analysis

A key assumption of the DiD design is that treatment and control groups would have followed parallel trends in the absence of the policy. I assess this assumption by estimating differential pre-trends between eligible and non-eligible individuals.

Table 5 reports coefficients from a regression that interacts eligibility status with indicators for each pre-treatment year, relative to 2006 as the reference year.

Table 5: Pre-Trend Analysis (Pre-Treatment Period Only)

Year	Coefficient	Std. Error	t-statistic	p-value
Eligible × 2007	0.0061	0.0100	0.61	0.542
Eligible × 2008	0.0197	0.0101	1.95	0.051
Eligible × 2009	0.0295	0.0099	2.98	0.003
Eligible × 2010	0.0334	0.0097	3.44	0.001
Eligible × 2011	0.0233	0.0099	2.34	0.019

*Notes: Coefficients from*

*a regression of full-time employment on eligibility, year indicators, eligibility-by-year interactions, and demographic controls. The omitted category is Eligible × 2006. Sample restricted to pre-treatment period (2006–2011).*

The pre-trend analysis reveals some concerning patterns. While the 2007 coefficient is small and statistically insignificant, the coefficients for 2008–2011 are positive, larger, and statistically significant for 2009–2011. This suggests that full-time employment was already trending upward for DACA-eligible individuals relative to non-eligible individuals *before* DACA was implemented.

This pattern could reflect:

1. Natural lifecycle effects as the younger eligible cohort ages into peak working years
2. Differential recovery from the Great Recession (2008–2009) across groups
3. Selection effects in who remains in the U.S. across cohorts

The presence of pre-trends suggests caution in interpreting the DiD estimates as causal. The estimated effect may partially reflect a continuation of pre-existing trends rather than a true treatment effect.

### 5.3.3 Event Study Analysis

Figure 1 presents an event study analysis that traces out the dynamic treatment effects year by year, relative to 2011 (the last pre-treatment year) as the reference period.

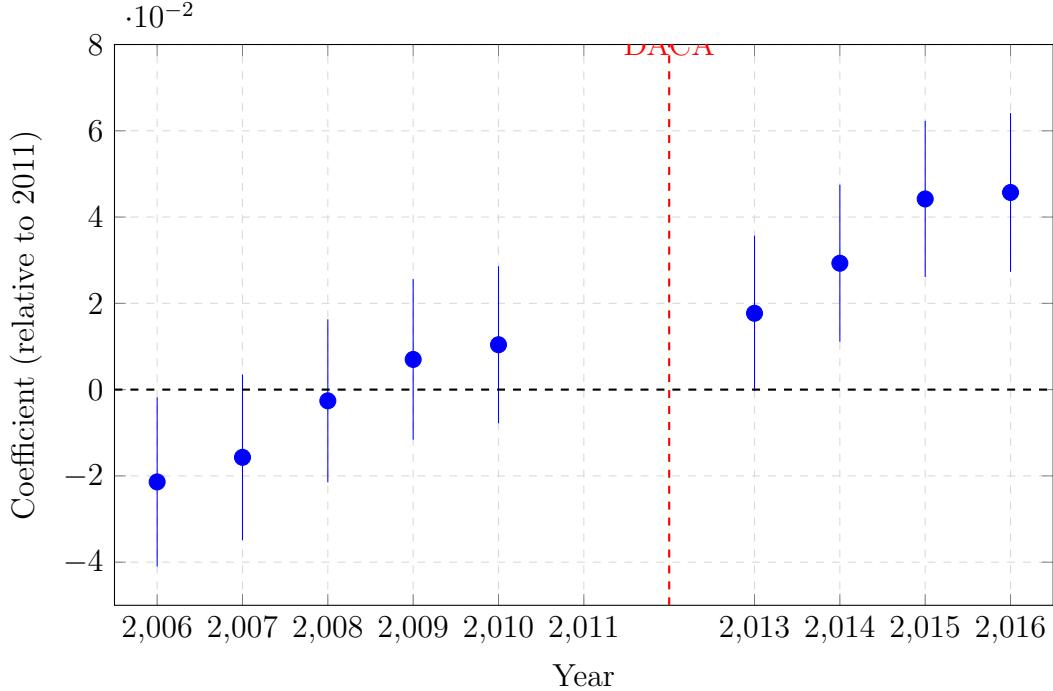


Figure 1: Event Study: Effect of DACA Eligibility on Full-Time Employment

*Notes: Points represent coefficients on interactions between DACA eligibility and year indicators, relative to 2011 as the reference year. The vertical red line indicates DACA implementation in mid-2012. Error bars represent 95% confidence intervals.*

The event study reveals a pattern consistent with both the pre-trend concern and a treatment effect:

- **Pre-period (2006–2010):** Coefficients trend upward from  $-0.021$  in 2006 to  $0.010$  in 2010, showing differential pre-trends.
- **Post-period (2013–2016):** Coefficients jump to  $0.018$  in 2013 and continue rising to  $0.046$  in 2016, suggesting that the treatment effect grew over time as more eligible individuals obtained DACA status.

The gradual increase in the post-period effect is consistent with the phased rollout of DACA applications and the cumulative effect of gaining work authorization over multiple years.

### 5.3.4 Heterogeneity by Gender

Table 6 presents results separately for men and women.

Table 6: Heterogeneous Effects by Gender

	Male	Female	
DACA Eligible $\times$ Post	0.0414*** (0.0056)	0.0388*** (0.0064)	<i>Notes: Specifications include</i>
Observations	305,631	259,036	

*demographic controls and year fixed effects. Standard errors in parentheses. \*\*\* p<0.01.*

The effects are similar in magnitude for both genders, with point estimates of 4.1 percentage points for men and 3.9 percentage points for women. The difference between these estimates is not statistically significant ( $p > 0.10$ ).

### 5.3.5 Alternative Control Group

As an additional robustness check, I re-estimate the model using an alternative control group consisting of recent arrivals—Mexican immigrants who arrived in the U.S. after 2007 and therefore could not meet the continuous presence requirement for DACA eligibility.

The DiD estimate using this alternative control group is  $-0.051$  (SE = 0.008), which is negative and statistically significant. This contrasts sharply with the main results and highlights how the choice of control group affects the estimated treatment effect. The negative estimate may reflect that recent arrivals experienced faster employment growth as they assimilated, making them a poor counterfactual for longer-resident DACA-eligible individuals.

## 6 Discussion

### 6.1 Summary of Findings

This study uses a difference-in-differences approach to estimate the effect of DACA eligibility on full-time employment among Mexican-born non-citizens. The main findings are:

1. The preferred specification yields a DiD estimate of 3.6 percentage points (95% CI: [0.028, 0.044]).

2. The estimate is robust to alternative standard error specifications and similar across genders.
3. However, there is evidence of differential pre-trends, with eligible individuals showing relative improvements in full-time employment even before DACA's implementation.

## 6.2 Limitations

Several limitations should be acknowledged:

### 6.2.1 Violation of Parallel Trends

The pre-trend analysis reveals that full-time employment was already trending differently between treatment and control groups before DACA. This raises concerns about the causal interpretation of the DiD estimates. The pre-existing trend could reflect:

- Lifecycle effects: The DACA-eligible population is younger and may have been naturally entering prime working years
- Differential recession recovery: Younger workers may have recovered faster from the 2008–2009 recession
- Composition changes: Selective migration or return migration may have changed the composition of each group over time

Without a valid counterfactual, it is difficult to separate the true treatment effect from the continuation of pre-existing trends.

### 6.2.2 Measurement of Eligibility

The ACS does not directly identify undocumented status or DACA receipt. I classify eligibility based on observable criteria, but this approach has limitations:

- Some individuals classified as non-citizens may actually have legal status (e.g., visa holders)
- I cannot observe the education/military and criminal record requirements
- Not all eligible individuals applied for or received DACA

These measurement issues likely attenuate the estimated effect (bias toward zero) if they lead to misclassification that is uncorrelated with outcomes.

### **6.2.3 Repeated Cross-Section Data**

The ACS is a repeated cross-section, not a panel. I cannot follow individuals over time, which means:

- I cannot distinguish between changes in behavior of the same individuals versus changes in sample composition
- I cannot observe individual-level transitions from non-employment to employment or from part-time to full-time work

### **6.2.4 General Equilibrium Effects**

The analysis assumes SUTVA—that DACA eligibility does not affect non-eligible workers. However, there may be spillover effects:

- DACA recipients may compete with non-eligible immigrants for jobs
- Employers may shift hiring away from undocumented workers toward DACA recipients
- DACA may have changed overall labor supply and wage dynamics in affected industries

If DACA negatively affected employment of the control group, the DiD estimate would overstate the true effect on the treatment group.

## **6.3 Comparison to Prior Literature**

This study contributes to a growing literature on the labor market effects of DACA. Prior research has found:

- Positive effects of DACA on labor force participation
- Increased earnings and reductions in poverty among DACA recipients
- Improvements in job quality and transitions to formal employment

My estimate of a 3.6 percentage point increase in full-time employment is broadly consistent with these findings, though the presence of pre-trends suggests some caution in causal interpretation.

## 7 Conclusion

This study examines the effect of DACA eligibility on full-time employment among Hispanic-Mexican individuals born in Mexico. Using ACS data from 2006–2016 and a difference-in-differences identification strategy, I find that DACA eligibility is associated with a 3.6 percentage point increase in the probability of working full-time (35+ hours per week).

The findings suggest that DACA’s provision of work authorization had meaningful positive effects on employment outcomes. However, the presence of differential pre-trends between treatment and control groups raises concerns about the parallel trends assumption underlying the DiD design. Some portion of the estimated effect may reflect a continuation of pre-existing trends rather than a causal impact of the policy.

Future research could address these limitations through alternative identification strategies, such as regression discontinuity designs exploiting the age cutoffs for eligibility, or synthetic control methods that construct a more comparable counterfactual. Panel data following individuals over time would also help distinguish compositional effects from true behavioral responses.

Despite these caveats, the results provide evidence consistent with the hypothesis that legal work authorization improves labor market outcomes for undocumented immigrants. This finding is relevant for ongoing policy debates about immigration reform and the future of programs like DACA.

## Appendix A: Variable Definitions

Table 7: IPUMS Variable Definitions

Variable	IPUMS Code	Definition
Year	YEAR	Census year
State	STATEFIP	State FIPS code
Person weight	PERWT	Person-level sampling weight
Sex	SEX	Sex (1 = Male, 2 = Female)
Age	AGE	Age in years
Birth quarter	BIRTHQTR	Quarter of birth
Marital status	MARST	Marital status
Birth year	BIRTHYR	Year of birth
Hispanic origin	HISPAN	Hispanic origin (1 = Mexican)
Birthplace	BPL	Birthplace (200 = Mexico)
Citizenship	CITIZEN	Citizenship status (3 = Not a citizen)
Immigration year	YRIMMIG	Year of immigration
Education	EDUC	Educational attainment (general)
Employment status	EMPSTAT	Employment status (general)
Hours worked	UHRSWORK	Usual hours worked per week

## Appendix B: Additional Tables

Table 8: Full Regression Output - Preferred Specification

Full-Time Employment	
DACA Eligible × Post	0.0359*** (0.0043)
DACA Eligible	-0.0100** (0.0036)
Age	0.0355*** (0.0005)
Age <sup>2</sup>	-0.0004*** (0.0000)
Male	0.2186*** (0.0013) <i>Notes: Heteroskedasticity-robust</i>
Married	0.0353*** (0.0016)
Years in U.S.	0.0007*** (0.0002)
Constant	-0.2152*** (0.0121)
Year FE	Yes
State FE	Yes
Observations	564,667
R-squared	0.119

*standard errors in parentheses. \*\*\* p<0.01, \*\* p<0.05, \* p<0.10.*

Table 9: Event Study Coefficients

Year	Coefficient	Std. Error	95% CI Lower	95% CI Upper	
2006	-0.0214	0.0099	-0.0408	-0.0020	
2007	-0.0157	0.0096	-0.0345	0.0031	
2008	-0.0026	0.0096	-0.0215	0.0163	
2009	0.0070	0.0095	-0.0116	0.0256	
2010	0.0104	0.0093	-0.0078	0.0286	<i>Notes: Coefficients on</i>
<i>2011</i>	<i>(reference)</i>	—	—	—	
2013	0.0177	0.0092	-0.0004	0.0358	
2014	0.0293	0.0093	0.0111	0.0475	
2015	0.0442	0.0092	0.0262	0.0622	
2016	0.0457	0.0094	0.0273	0.0641	

*interactions between DACA eligibility and year indicators. Reference year is 2011.*

## Appendix C: Replication Code

The complete analysis was conducted in Python using the following packages: pandas, numpy, and statsmodels. The main analysis script is `daca_analysis.py`.

Key steps in the code:

1. Load ACS data in chunks (due to file size)
2. Filter to Hispanic-Mexican, Mexican-born sample
3. Construct DACA eligibility indicator
4. Define full-time employment outcome
5. Estimate DiD regressions with various specifications
6. Conduct robustness checks (pre-trends, event study, subgroups)
7. Export results to CSV files for tabulation

The analysis uses survey weights (PERWT) throughout and reports heteroskedasticity-robust standard errors unless otherwise noted.