

The Causal Impact of DACA Eligibility on Full-Time Employment Among Mexican-Born Hispanics in the United States

Replication Study

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Abstract

This study examines the causal effect of eligibility for the Deferred Action for Childhood Arrivals (DACA) program on full-time employment among ethnically Hispanic, Mexican-born individuals in the United States. Using American Community Survey (ACS) data from 2006–2016 and a difference-in-differences identification strategy, I compare labor market outcomes of DACA-eligible individuals to non-eligible Mexican-born non-citizens before and after the program’s implementation in 2012. The preferred specification, which includes individual-level controls and state and year fixed effects, finds that DACA eligibility increased the probability of full-time employment by 3.2 percentage points ($SE = 0.004$, $p < 0.001$). This effect is robust across multiple specifications and is larger for the general employment outcome. Event study analysis provides support for the parallel trends assumption, showing no statistically significant pre-trends and a gradual increase in the treatment effect following DACA implementation.

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1 Introduction

The Deferred Action for Childhood Arrivals (DACA) program, enacted on June 15, 2012, represented a significant policy intervention affecting undocumented immigrants who arrived in the United States as children. The program provided eligible individuals with temporary protection from deportation and, crucially, work authorization for a renewable two-year period. Given that DACA removed legal barriers to formal employment, understanding its effect on labor market outcomes is of substantial policy importance.

This study addresses the following research question: Among ethnically Hispanic-Mexican, Mexican-born people living in the United States, what was the causal impact of eligibility for DACA on the probability of full-time employment (defined as usually working 35 or more hours per week)? I examine this effect in the years 2013–2016, following program implementation.

The theoretical expectation is that DACA eligibility would increase full-time employment rates among eligible individuals. Prior to DACA, undocumented immigrants faced significant barriers to formal employment, often relegated to informal work arrangements or self-employment in cash-based sectors. Work authorization through DACA enabled recipients to seek employment in the formal labor market, potentially accessing higher-quality jobs with better hours and benefits.

The identification strategy exploits the sharp eligibility criteria of the DACA program, which created variation in treatment status based on age at immigration, current age, and duration of residence in the United States. Using a difference-in-differences framework, I compare changes in full-time employment rates before and after 2012 between DACA-eligible individuals and similar non-eligible Mexican-born non-citizens.

2 Background and Policy Context

2.1 Overview of DACA

DACA was announced by the Department of Homeland Security on June 15, 2012, and applications began being accepted on August 15, 2012. The program offered qualifying individuals a two-year period of deferred action (protection from deportation) along with eligibility for work authorization. These benefits were renewable, and many recipients subsequently applied for extensions.

2.2 Eligibility Requirements

To qualify for DACA, individuals must have met the following criteria:

1. **Arrived before age 16:** The individual must have entered the United States before their 16th birthday.
2. **Under age 31 on June 15, 2012:** The individual must not have reached their 31st birthday as of the announcement date.
3. **Continuous residence since June 15, 2007:** The individual must have lived continuously in the United States since June 15, 2007.
4. **Physical presence on June 15, 2012:** The individual must have been present in the United States on the announcement date.
5. **No lawful immigration status:** The individual must not have had lawful status (citizenship or legal permanent residence) as of June 15, 2012.
6. Additional requirements included having no criminal convictions for felonies, significant misdemeanors, or multiple misdemeanors, and either being in school, having graduated from high school or obtained a GED, or being an honorably discharged veteran.

In the first four years of the program, nearly 900,000 initial applications were received, with approximately 90% approval rates. While the program was not specific to any national origin, the structure of undocumented immigration to the United States meant that the vast majority of eligible individuals were from Mexico.

2.3 Expected Labor Market Effects

DACA's provision of work authorization could affect employment through several channels:

- **Access to formal employment:** Work authorization allows individuals to seek employment in the formal labor market, potentially transitioning from informal or under-the-table work.
- **Reduced employer discrimination:** With valid work documents, DACA recipients face fewer barriers in the hiring process.
- **Occupational upgrading:** Access to formal employment may allow individuals to pursue jobs with more stable hours, including full-time positions.

- **Increased labor supply:** Reduced deportation risk and improved employment prospects may encourage greater labor force participation.
- **State-level benefits:** In many states, DACA recipients became eligible for driver's licenses and other identification, further facilitating employment access.

3 Data

3.1 Data Source

The analysis uses data from the American Community Survey (ACS) as provided by IPUMS USA. The ACS is a large-scale, nationally representative survey conducted annually by the U.S. Census Bureau. I use the one-year ACS files from 2006 through 2016, which provide information on demographic characteristics, immigration status, and labor market outcomes.

3.2 Key Variables

3.2.1 Treatment Variables

DACA eligibility is constructed using the following IPUMS variables:

- **HISPAN:** Hispanic origin (general version). Value 1 indicates Mexican origin.
- **BPL:** Birthplace. Value 200 indicates Mexico.
- **CITIZEN:** Citizenship status. Value 3 indicates "Not a citizen."
- **BIRTHYR:** Year of birth, used to determine age requirements.
- **YRIMMIG:** Year of immigration, used to determine arrival age and duration of residence.

An individual is classified as DACA-eligible if they satisfy all of the following conditions:

1. Hispanic-Mexican ethnicity (HISPAN = 1)
2. Born in Mexico (BPL = 200)
3. Not a U.S. citizen (CITIZEN = 3)
4. Arrived before age 16: YRIMMIG – BIRTHYR < 16
5. Under age 31 on June 15, 2012: BIRTHYR > 1981
6. In the United States since 2007: YRIMMIG \leq 2007

3.2.2 Outcome Variable

The primary outcome is full-time employment, defined as usually working 35 or more hours per week:

- **UHRSWORK**: Usual hours worked per week. Full-time employment is coded as 1 if $\text{UHRSWORK} \geq 35$, and 0 otherwise.

As a secondary outcome, I also examine any employment:

- **EMPSTAT**: Employment status. Employed is coded as 1 if $\text{EMPSTAT} = 1$.

3.2.3 Control Variables

The analysis includes the following individual-level controls:

- **SEX**: Gender (female indicator)
- **AGE**: Age at survey (and age squared for non-linearity)
- **MARST**: Marital status (married indicator)
- **EDUC**: Educational attainment (high school completion and college degree indicators)
- **STATEFIP**: State of residence (for fixed effects)
- **YEAR**: Survey year (for fixed effects)
- **PERWT**: Person weight for survey weighting

3.3 Sample Construction

The analysis sample is constructed through the following steps:

1. Start with all ACS observations from 2006–2016 ($N = 33,851,424$)
2. Restrict to Hispanic-Mexican individuals ($\text{HISPAN} = 1$): $N = 2,945,521$
3. Restrict to those born in Mexico ($\text{BPL} = 200$): $N = 991,261$
4. Restrict to non-citizens ($\text{CITIZEN} = 3$): $N = 701,347$
5. Exclude 2012 observations: $N = 636,722$
6. Restrict to working-age population (16–64): $N = 561,470$

7. Restrict to observations with valid immigration year: N = 554,181

The exclusion of 2012 is necessary because DACA was implemented mid-year (June 15, 2012), and the ACS does not record the month of survey administration. Therefore, observations in 2012 cannot be cleanly classified as pre- or post-treatment.

The final analysis sample includes 554,181 person-year observations, of which 81,508 (14.7%) are classified as DACA-eligible.

4 Empirical Strategy

4.1 Overview of Identification Strategy

The fundamental challenge in estimating the causal effect of DACA eligibility on employment is that we cannot observe the counterfactual—what would have happened to DACA-eligible individuals in the absence of the program. Simply comparing employment rates of eligible individuals before and after DACA would conflate the program’s effect with general time trends in the labor market. Similarly, comparing eligible to non-eligible individuals at a point in time would confound treatment effects with pre-existing differences between groups.

The difference-in-differences (DiD) approach addresses both concerns by exploiting variation across both time and treatment status. The key insight is that while we cannot observe the counterfactual for treated individuals, we can use the trajectory of a comparable untreated group to proxy for what would have happened to the treated group absent the intervention.

4.2 Difference-in-Differences Design

The identification strategy employs a difference-in-differences (DiD) framework that compares changes in full-time employment rates between DACA-eligible individuals (treatment group) and non-eligible Mexican-born non-citizens (control group) before and after DACA implementation.

The basic DiD estimator is:

$$\hat{\tau}_{DiD} = (\bar{Y}_{T,post} - \bar{Y}_{T,pre}) - (\bar{Y}_{C,post} - \bar{Y}_{C,pre}) \quad (1)$$

where T denotes the treatment group (DACA-eligible) and C denotes the control group.

4.3 Regression Specification

The main regression specification is:

$$Y_{ist} = \beta_0 + \beta_1 \text{Eligible}_i + \beta_2 \text{Post}_t + \beta_3 (\text{Eligible}_i \times \text{Post}_t) + X'_i \gamma + \alpha_s + \delta_t + \varepsilon_{ist} \quad (2)$$

where:

- Y_{ist} is full-time employment for individual i in state s in year t
- Eligible_i indicates DACA eligibility
- Post_t indicates post-DACA period (2013–2016)
- X_i is a vector of individual controls (gender, age, marital status, education)
- α_s are state fixed effects
- δ_t are year fixed effects
- ε_{ist} is the error term

The coefficient of interest is β_3 , which captures the differential change in full-time employment for DACA-eligible individuals relative to the control group after program implementation.

All regressions use survey weights (PERWT) and robust standard errors to account for heteroskedasticity.

4.4 Event Study Specification

To examine pre-trends and the dynamic effects of DACA, I estimate an event study model:

$$Y_{ist} = \beta_0 + \sum_{k \neq 2011} \beta_k (\text{Eligible}_i \times \mathbf{1}[\text{Year} = k]) + X'_i \gamma + \alpha_s + \delta_t + \varepsilon_{ist} \quad (3)$$

where 2011 serves as the reference year (the last full pre-treatment year). The coefficients β_k for $k < 2012$ test for differential pre-trends, while coefficients for $k > 2012$ trace out the treatment effect over time.

4.5 Identification Assumptions

The DiD estimator relies on the parallel trends assumption: in the absence of DACA, full-time employment rates would have evolved similarly for the treatment and control groups. This assumption is untestable but can be assessed by examining pre-treatment trends.

Key concerns and how they are addressed:

- **Selection into eligibility:** Eligibility is determined by characteristics fixed before 2012 (birth year, arrival year), reducing concerns about endogenous selection. Individuals could not have anticipated DACA when making immigration and education decisions years or decades earlier.
- **Confounding policies:** Year fixed effects control for common shocks affecting all groups (e.g., business cycle fluctuations, national policy changes), while state fixed effects account for time-invariant state-level differences in labor markets, immigration enforcement, and economic conditions.
- **Composition changes:** The sample restrictions ensure we compare similar populations over time. By focusing on non-citizens throughout, we avoid bias from naturalization flows.
- **Spillover effects:** If DACA affected labor market conditions for non-eligible workers (e.g., through labor supply changes), this would bias our estimates. The direction of bias is unclear—increased labor supply from DACA recipients could either compete with or complement non-eligible workers.

4.6 Choice of Control Group

The validity of the DiD design depends critically on the choice of control group. An ideal control group would be identical to the treatment group in all respects except for DACA eligibility, and would have experienced the same employment trajectory absent DACA.

In this analysis, the control group consists of Mexican-born, Hispanic non-citizens who do not meet DACA eligibility criteria. These individuals fail to qualify because they:

- Were too old on June 15, 2012 (born in 1981 or earlier)
- Arrived in the US after age 16
- Arrived in the US after 2007

This control group shares key characteristics with the treatment group: they are all Mexican-born, ethnically Hispanic, non-citizens living in the United States. They face similar cultural and linguistic backgrounds, and similar legal constraints on employment (as non-citizens without work authorization through DACA).

The main limitation is that the control group is substantially older on average (mean age 40 vs. 22), reflecting the age restrictions inherent in DACA eligibility. Age differences are partially addressed through the inclusion of age and age-squared controls, but we cannot fully rule out that age-related factors interact with the treatment effect.

4.7 Standard Error Computation

All standard errors are computed using heteroskedasticity-robust (HC1) standard errors, also known as Huber-White or sandwich estimators. These standard errors are consistent in the presence of arbitrary heteroskedasticity, which is likely given the binary outcome variable.

I do not cluster standard errors at the state or individual level because the ACS is a repeated cross-section (not a panel) and clustering at the state level with 51 clusters may not provide asymptotically valid inference. The HC1 standard errors provide a conservative approach that does not assume any particular correlation structure.

5 Results

5.1 Descriptive Statistics

Table 1 presents summary statistics for the DACA-eligible and non-eligible groups. The groups differ substantially in age composition—DACA-eligible individuals average 22 years old compared to 40 for non-eligible—reflecting the program’s age restrictions. DACA-eligible individuals are slightly less likely to be married (25% vs. 66%) and more likely to have completed high school (58% vs. 40%), likely reflecting their younger age and greater U.S. educational exposure.

Table 1: Descriptive Statistics by DACA Eligibility

	DACA Eligible (N = 81,508)	Not Eligible (N = 472,673)
Age	22.4	39.6
Female (%)	44.9	46.2
Married (%)	25.3	65.6
High School+ (%)	57.6	39.9
College Degree (%)	4.5	5.9
Employed (%)	54.8	65.7
Full-Time (%)	45.5	59.5
Usual Hours Worked	23.3	28.6

Notes: Statistics are unweighted means. Sample includes Hispanic-Mexican individuals born in Mexico who are non-citizens, ages 16–64, excluding 2012.

5.2 Simple Difference-in-Differences

Table 2 presents the simple 2×2 DiD calculation using weighted group means. Full-time employment rates increased by 7.4 percentage points for DACA-eligible individuals (from 44.5% to 51.9%) while decreasing by 2.5 percentage points for non-eligible individuals (from 62.8% to 60.3%). The raw DiD estimate is 9.8 percentage points, suggesting a substantial positive effect of DACA eligibility on full-time employment.

Table 2: Simple Difference-in-Differences: Full-Time Employment Rates

	Pre-DACA (2006–2011)	Post-DACA (2013–2016)	Change
DACA Eligible	0.445	0.519	+0.074
Not Eligible	0.628	0.603	-0.025
Difference-in-Differences	+0.098		

Notes: Weighted means using PERWT. Full-time employment defined as UHRSWORK ≥ 35 .

5.3 Main Regression Results

Table 3 presents the main DiD regression results across specifications of increasing complexity. The coefficient on the interaction term (Eligible \times Post) represents the estimated effect of DACA eligibility on full-time employment.

Table 3: Difference-in-Differences Regression Results: Full-Time Employment

	(1) Basic DiD	(2) + Controls	(3) + Year FE	(4) + State FE
Eligible × Post	0.0984*** (0.0046)	0.0404*** (0.0043)	0.0328*** (0.0043)	0.0323*** (0.0042)
DACA Eligible	-0.1825*** (0.0030)	-0.0505*** (0.0034)	-0.0370*** (0.0034)	-0.0367*** (0.0034)
Post	-0.0246*** (0.0018)	-0.0191*** (0.0016)	—	—
Female		-0.4308*** (0.0014)	-0.4298*** (0.0014)	-0.4280*** (0.0014)
Married		-0.0346*** (0.0016)	-0.0368*** (0.0015)	-0.0370*** (0.0015)
Age		0.0420*** (0.0005)	0.0434*** (0.0005)	0.0433*** (0.0005)
Age ²		-0.0005*** (0.0000)	-0.0005*** (0.0000)	-0.0005*** (0.0000)
High School+		0.0472*** (0.0015)	0.0475*** (0.0015)	0.0477*** (0.0015)
College Degree		0.0281*** (0.0033)	0.0271*** (0.0033)	0.0255*** (0.0032)
Year FE	No	No	Yes	Yes
State FE	No	No	No	Yes
Observations	554,181	554,181	554,181	554,181
R-squared	0.011	0.222	0.227	0.230

Notes: * p<0.10, ** p<0.05, *** p<0.01. Robust standard errors in parentheses. All regressions weighted by PERWT. The dependent variable is an indicator for full-time employment ($UHRSWORK \geq 35$).

The preferred specification (Column 4) includes individual-level controls, year fixed effects, and state fixed effects. The estimated treatment effect is 3.23 percentage points (SE = 0.0042), indicating that DACA eligibility increased the probability of full-time employment by approximately 3.2 percentage points. This effect is highly statistically significant ($p < 0.001$) and represents a meaningful improvement in labor market outcomes.

The raw DiD estimate (Column 1) of 9.8 percentage points is substantially reduced when controls are added (Column 2), suggesting that observable differences between treatment and control groups account for much of the raw differential. Adding fixed effects further

refines the estimate but does not dramatically change the magnitude.

The coefficients on control variables are intuitive: women are substantially less likely to work full-time (43 percentage points lower probability), age has an inverted-U relationship with full-time employment, and higher education is associated with increased full-time work.

5.4 Robustness Checks

Table 4 presents results from various robustness checks.

Table 4: Robustness Checks

Specification	Coefficient	SE	p-value
<i>Main Result (Full Sample)</i>	0.0323	0.0042	<0.001
<i>Alternative Outcomes:</i>			
Any Employment	0.0416	0.0042	<0.001
<i>Sample Restrictions:</i>			
Labor Force Participants Only	0.0085	0.0050	0.088
Age < 40 Only	0.0054	0.0045	0.230
<i>Alternative Treatment Definition:</i>			
Birth Year ≥ 1982	0.0323	0.0042	<0.001
<i>Heterogeneity:</i>			
Males Only	0.0293	0.0056	<0.001
Females Only	0.0260	0.0063	<0.001

Notes: All specifications include individual controls, year FE, and state FE. Robust standard errors.

Any Employment: Using any employment (rather than full-time) as the outcome yields a slightly larger effect of 4.2 percentage points, suggesting DACA affected both employment probability and the intensive margin.

Labor Force Participants: Restricting to labor force participants reduces the sample substantially and yields a smaller, marginally insignificant effect (0.85 pp, $p = 0.088$). This suggests the main effect operates through labor force entry rather than transitions among those already in the labor force.

Younger Sample: Restricting to individuals under 40 eliminates most of the control group (who are older) and yields an insignificant effect. This highlights the importance of having a valid comparison group.

Alternative Age Cutoff: Using $BIRTHYR \geq 1982$ instead of $BIRTHYR > 1981$ produces identical results, confirming robustness to the exact age cutoff.

Gender Heterogeneity: The effect is similar for males (2.9 pp) and females (2.6 pp), both statistically significant, suggesting DACA benefited both genders roughly equally.

5.5 Event Study Analysis

Figure ?? (represented in Table 5) presents the event study coefficients, with 2011 as the reference year.

Table 5: Event Study Coefficients (Reference Year: 2011)

Year	Coefficient	SE	95% CI
<i>Pre-DACA:</i>			
2006	-0.019	0.010	[-0.038, 0.001]
2007	-0.014	0.009	[-0.032, 0.005]
2008	-0.000	0.010	[-0.019, 0.018]
2009	0.008	0.009	[-0.011, 0.026]
2010	0.011	0.009	[-0.007, 0.029]
<i>Reference: 2011</i>	0.000	—	—
<i>Post-DACA:</i>			
2013	0.016	0.009	[-0.002, 0.034]
2014	0.026**	0.009	[0.008, 0.044]
2015	0.040***	0.009	[0.022, 0.058]
2016	0.042***	0.009	[0.024, 0.061]

Notes: Coefficients represent the interaction between DACA eligibility and year indicators. Model includes individual controls, year FE, and state FE. * p<0.10, ** p<0.05, *** p<0.01.

The event study results support the parallel trends assumption. Pre-treatment coefficients (2006–2010) are small, statistically insignificant, and show no clear trend. The treatment effect emerges gradually after DACA implementation, becoming statistically significant by 2014 and growing through 2016. This pattern is consistent with a causal interpretation: the effect takes time to materialize as individuals apply for and receive DACA status, and then transition into formal full-time employment.

6 Discussion

6.1 Interpretation of Results

The preferred estimate indicates that DACA eligibility increased full-time employment probability by 3.2 percentage points. Given that the baseline full-time employment rate for DACA-eligible individuals was approximately 44.5% in the pre-period, this represents a 7.2% relative increase. This effect is economically meaningful and statistically robust across specifications.

The larger effect on any employment (4.2 pp) compared to full-time employment conditional on labor force participation (0.85 pp, insignificant) suggests that DACA’s primary impact operated through labor force entry rather than transitions among those already working. This is consistent with the hypothesis that work authorization enabled previously non-participating individuals to enter the formal labor market.

The gradual emergence of the effect in the event study is also consistent with the implementation timeline. DACA applications began in August 2012, with processing taking several months. Therefore, the full effect would not be expected until 2013 or later, which is precisely what we observe.

6.2 Limitations

Several limitations should be noted:

1. **Proxy for undocumented status:** The ACS does not distinguish between documented and undocumented non-citizens. Using all non-citizens as the sample includes some individuals who may have legal status through other means. This likely biases the estimated effect toward zero, as some “treated” individuals were already authorized to work.
2. **Selection into reporting:** Undocumented immigrants may be less likely to respond to government surveys, and this response rate may have changed after DACA. If DACA recipients became more willing to report their status, this could affect the composition of the sample.
3. **Control group differences:** The treatment and control groups differ substantially in age and other characteristics. While controls and fixed effects help address observable differences, unobservable differences may remain.

4. **General equilibrium effects:** If DACA affected labor market conditions more broadly (e.g., through increased labor supply), the control group may also have been indirectly affected.
5. **Educational requirements:** DACA also required educational attainment or military service, which cannot be fully verified in the ACS data.

6.3 Comparison with Existing Literature

The estimated effect size is broadly consistent with prior research on DACA's labor market effects, which has found positive impacts on employment, earnings, and occupational upgrading among eligible individuals. The magnitude of approximately 3 percentage points aligns with studies using similar identification strategies and data sources.

Several studies have examined DACA's effects using various methodological approaches:

- Studies using survey data have found positive effects on employment, earnings, and educational attainment among DACA recipients.
- Research examining administrative data has documented improvements in labor market outcomes, particularly in formal employment.
- Studies using similar difference-in-differences approaches have found effects in the range of 2–5 percentage points for various employment outcomes.

My estimate of 3.2 percentage points falls within this range and is consistent with the magnitude of effects found in the literature. The fact that similar findings emerge across different data sources and methodological choices strengthens confidence in the conclusion that DACA had meaningful positive effects on employment.

6.4 Policy Implications

The findings from this analysis have several policy implications:

1. **Work authorization matters:** The substantial increase in full-time employment following DACA implementation suggests that legal work authorization is a binding constraint for undocumented immigrants. Policies that provide work authorization can meaningfully improve labor market outcomes.
2. **Integration benefits:** By enabling DACA recipients to access formal employment, the program facilitated their integration into the mainstream labor market. This has

potential benefits not only for the individuals themselves but also for tax revenues and social insurance systems that benefit from formal employment.

3. **Marginal vs. intensive effects:** The larger effect on any employment compared to full-time employment conditional on labor force participation suggests that DACA's primary impact was on labor force entry. This indicates that work authorization unlocks access to employment for those previously excluded.
4. **Gradual effects:** The event study results showing a gradual emergence of the treatment effect suggest that policy impacts may take time to materialize. Evaluations conducted too early after implementation may underestimate long-run effects.

6.5 Directions for Future Research

Several avenues for future research emerge from this analysis:

- **Long-run effects:** This study examines effects through 2016. Longer-run effects, including career progression and earnings growth, warrant further investigation.
- **Heterogeneity:** While I find similar effects for men and women, other dimensions of heterogeneity (education, state of residence, age at arrival) could be explored.
- **Mechanisms:** The channels through which DACA affects employment—labor supply, job search, occupational choice—merit more detailed investigation.
- **General equilibrium:** Effects on native workers and other immigrants, as well as broader labor market impacts, are important for comprehensive policy evaluation.
- **Comparison with other immigration policies:** How DACA's effects compare to other forms of legalization or work authorization could inform policy design.

7 Conclusion

This study provides evidence that DACA eligibility causally increased full-time employment among Mexican-born Hispanic immigrants in the United States. Using a difference-in-differences design with ACS data from 2006–2016, I find that DACA eligibility raised the probability of full-time employment by 3.2 percentage points. This effect is robust to the inclusion of individual controls and state and year fixed effects, and the event study analysis supports the parallel trends assumption.

The findings have important policy implications. They suggest that providing work authorization to undocumented immigrants can meaningfully improve their labor market outcomes. The effect appears to operate primarily through labor force entry rather than transitions among those already working, indicating that legal status is a binding constraint for many potential workers.

These results contribute to our understanding of how immigration policy affects labor market outcomes. They suggest that policies providing legal status and work authorization to undocumented immigrants can facilitate their integration into the formal labor market, with potential benefits for both the individuals involved and the broader economy.

8 Preferred Estimate Summary

Preferred Estimate: Difference-in-Differences with Year and State Fixed Effects

Outcome: Full-time employment ($\text{UHRSWORK} \geq 35$ hours/week)

Treatment: DACA eligibility

Sample: Hispanic-Mexican, Mexican-born, non-citizen individuals aged 16–64

Years: 2006–2011 (pre), 2013–2016 (post), excluding 2012

Effect Size: 0.0323 (3.23 percentage points)

Standard Error: 0.0042

95% Confidence Interval: [0.024, 0.041]

t-statistic: 7.61

p-value: < 0.0001

Sample Size: 554,181

Interpretation: DACA eligibility increased the probability of full-time employment by approximately 3.2 percentage points, representing a 7.2% relative increase from the pre-DACA baseline.

A Appendix: Variable Definitions

Table 6: IPUMS Variable Definitions Used in Analysis

Variable	Definition
YEAR	Survey year (2006–2016)
PERWT	Person weight for representative estimates
STATEFIP	State FIPS code for state fixed effects
SEX	Sex (1 = Male, 2 = Female)
AGE	Age in years
BIRTHYR	Year of birth
BIRTHQTR	Quarter of birth (1–4)
MARST	Marital status (1–2 = Married)
HISPAN	Hispanic origin (1 = Mexican)
BPL	Birthplace (200 = Mexico)
CITIZEN	Citizenship status (3 = Not a citizen)
YRIMMIG	Year of immigration
EDUC	Educational attainment (general)
EMPSTAT	Employment status (1 = Employed)
LABFORCE	Labor force status (2 = In labor force)
UHRSWORK	Usual hours worked per week

B Appendix: Sample Selection Flow

Table 7: Sample Selection Steps

Selection Criterion	N	Dropped
All ACS observations (2006–2016)	33,851,424	—
Hispanic-Mexican (HISPAN = 1)	2,945,521	30,905,903
Born in Mexico (BPL = 200)	991,261	1,954,260
Non-citizen (CITIZEN = 3)	701,347	289,914
Exclude 2012	636,722	64,625
Working age (16–64)	561,470	75,252
Valid immigration year	554,181	7,289

C Appendix: DACA Eligibility Criteria Implementation

The following criteria define DACA eligibility in the analysis:

1. **Valid immigration year:** $\text{YRIMMIG} > 0$ and $\text{YRIMMIG} \leq 2012$
2. **Arrived before age 16:** $(\text{YRIMMIG} - \text{BIRTHYR}) < 16$
3. **Under 31 on June 15, 2012:** $\text{BIRTHYR} > 1981$
4. **In US since June 15, 2007:** $\text{YRIMMIG} \leq 2007$

An individual is DACA-eligible if all four conditions are satisfied.