

# The Effect of DACA Eligibility on Full-Time Employment: A Difference-in-Differences Analysis

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## Abstract

This study examines the causal impact of eligibility for the Deferred Action for Childhood Arrivals (DACA) program on full-time employment among Mexican-born Hispanic individuals in the United States. Using a difference-in-differences research design that compares individuals aged 26-30 at DACA implementation (the treatment group) to those aged 31-35 (the control group who missed eligibility due to age), I find that DACA eligibility increased the probability of full-time employment by approximately 4.8 percentage points. This effect is statistically significant and robust to various specifications including demographic controls and state-year fixed effects. Event study analysis supports the parallel trends assumption, with no significant pre-treatment differential trends detected. These findings suggest that DACA's provision of work authorization had meaningful positive effects on the labor market outcomes of eligible individuals.

**Keywords:** DACA, immigration policy, employment, difference-in-differences, labor economics

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# 1 Introduction

The Deferred Action for Childhood Arrivals (DACA) program, enacted on June 15, 2012, represents one of the most significant immigration policy changes in recent U.S. history. The program provided temporary relief from deportation and work authorization to undocumented immigrants who arrived in the United States as children, commonly referred to as “Dreamers.” Understanding the labor market effects of this policy is crucial for evaluating its impact and informing future immigration policy decisions.

This study investigates a specific research question: Among ethnically Hispanic-Mexican individuals born in Mexico and living in the United States, what was the causal impact of DACA eligibility on the probability of full-time employment? Full-time employment is defined as usually working 35 hours per week or more, following the Bureau of Labor Statistics definition.

The research design employs a difference-in-differences (DiD) framework that exploits the age-based eligibility criterion of DACA. The treatment group consists of individuals who were ages 26-30 at the time of DACA implementation (June 15, 2012), corresponding to birth years 1982-1986. The control group consists of individuals who were ages 31-35 at implementation (birth years 1977-1981), who would have been eligible for DACA except that they had reached their 31st birthday before June 15, 2012. By comparing changes in full-time employment between these two groups before and after DACA implementation, we can estimate the causal effect of eligibility.

The main finding of this analysis is that DACA eligibility increased the probability of full-time employment by approximately 4.8 percentage points. This effect is statistically significant at conventional levels and robust to the inclusion of demographic controls and state and year fixed effects. Event study analysis provides support for the parallel trends assumption underlying the DiD design, with no statistically significant differential pre-trends detected between treatment and control groups.

The remainder of this report is organized as follows. Section 2 provides background on the DACA program. Section 3 describes the data and sample construction. Section 4 presents the empirical methodology. Section 5 reports the main results. Section 6 presents robustness checks. Section 7 discusses the findings, and Section 8 concludes.

## **2 Background on DACA**

### **2.1 Program Overview**

DACA was announced by the Department of Homeland Security on June 15, 2012, and began accepting applications on August 15, 2012. The program allows eligible undocumented immigrants to apply for deferred action, which provides temporary protection from deportation for two years, along with authorization to work legally in the United States. Recipients could also apply for driver's licenses and other forms of identification in many states.

### **2.2 Eligibility Requirements**

To be eligible for DACA, individuals must meet the following criteria:

1. Arrived in the United States before their 16th birthday
2. Had not yet reached their 31st birthday as of June 15, 2012
3. Lived continuously in the United States since June 15, 2007
4. Were physically present in the United States on June 15, 2012
5. Did not have lawful immigration status (citizenship or legal residency) on June 15, 2012
6. Were currently in school, had graduated from high school, obtained a GED, or were an honorably discharged veteran
7. Had not been convicted of a felony, significant misdemeanor, or three or more other misdemeanors

### **2.3 Program Uptake**

In the first four years of the program, nearly 900,000 initial applications were received, with approximately 90% approved. While the program was not specific to any country of origin, the great majority of eligible individuals were from Mexico, reflecting the composition of the undocumented immigrant population in the United States.

## 2.4 Expected Effects on Employment

DACA is expected to affect employment through several channels. Most directly, work authorization allows recipients to seek employment in the formal labor market without fear of employer sanctions. This can expand job opportunities beyond the informal sector and potentially lead to better matches between workers and jobs. Additionally, the ability to obtain driver's licenses in many states may reduce transportation barriers to employment. Finally, reduced deportation risk may increase job stability and allow individuals to invest in job-specific human capital.

## 3 Data

### 3.1 Data Source

This analysis uses data from the American Community Survey (ACS) as provided by IPUMS USA. The ACS is an annual survey conducted by the U.S. Census Bureau that collects detailed demographic, social, economic, and housing information from a representative sample of the U.S. population. I use the one-year ACS samples from 2006 through 2016, excluding 2012 due to the timing ambiguity discussed below.

### 3.2 Sample Construction

The analytical sample is constructed through the following sequential restrictions:

1. **Hispanic-Mexican Ethnicity:** Restricted to individuals coded as Hispanic-Mexican in the HISPAN variable ( $HISPAN = 1$ ).
2. **Born in Mexico:** Restricted to individuals born in Mexico ( $BPL = 200$ ).
3. **Non-Citizen Status:** Restricted to non-citizens ( $CITIZEN = 3$ ). Since the ACS does not distinguish between documented and undocumented non-citizens, I assume that non-citizens who have not received citizenship or first papers are potentially undocumented and thus potentially DACA-eligible.
4. **Arrived Before Age 16:** Calculated age at immigration as  $YRIMMIG - BIRTHYR$  and restricted to individuals who arrived before age 16.
5. **Continuous Presence Since 2007:** Restricted to individuals with year of immigration on or before 2007 ( $YRIMMIG \leq 2007$ ) to satisfy the continuous presence requirement.

6. **Age Cohorts:** Restricted to treatment cohort (birth years 1982-1986, ages 26-30 on June 15, 2012) and control cohort (birth years 1977-1981, ages 31-35 on June 15, 2012).
7. **Exclude 2012:** Removed observations from 2012 because the ACS does not record the month of data collection, making it impossible to determine whether observations are from before or after DACA implementation on June 15, 2012.

Table 1 summarizes the sample construction process.

Table 1: Sample Construction

Restriction	Observations	Remaining
Raw ACS data (2006-2016)	33,851,424	100.0%
Hispanic-Mexican (HISPAN = 1)	2,945,521	8.7%
Born in Mexico (BPL = 200)	991,261	2.9%
Non-citizen (CITIZEN = 3)	701,347	2.1%
Arrived before age 16	205,327	0.6%
In US since 2007 ( $YRIMMIG \leq 2007$ )	195,023	0.6%
Birth years 1977-1986	49,019	0.1%
Exclude 2012	44,725	0.1%

Notes: Sequential sample restrictions applied to American Community Survey data from IPUMS USA. Final sample includes treatment group (born 1982-1986) and control group (born 1977-1981).

### 3.3 Variable Definitions

#### 3.3.1 Outcome Variable

The primary outcome is full-time employment, defined as usually working 35 or more hours per week. This is constructed from the UHRSWORK variable as:

$$\text{FullTime}_i = \mathbf{1}[\text{UHRSWORK}_i \geq 35] \quad (1)$$

where  $\mathbf{1}[\cdot]$  is the indicator function.

#### 3.3.2 Treatment Variables

- **Treated:** Indicator equal to 1 for individuals born 1982-1986 (ages 26-30 on June 15, 2012)
- **Post:** Indicator equal to 1 for survey years 2013-2016
- **DiD:** Interaction term  $\text{Treated} \times \text{Post}$ , which captures the treatment effect

### 3.3.3 Control Variables

- **Female:** Indicator for female ( $\text{SEX} = 2$ )
- **Married:** Indicator for currently married ( $\text{MARST} \leq 2$ )
- **Education:** Categorical indicators for high school graduate ( $\text{EDUC} = 6$ ), some college ( $\text{EDUC} = 7-9$ ), and college graduate or more ( $\text{EDUC} \geq 10$ ), with less than high school as the omitted category

## 3.4 Summary Statistics

Table 2 presents summary statistics for the analytical sample, separately by treatment status and time period.

Table 2: Summary Statistics

Variable	Pre-DACA (2006-2011)		Post-DACA (2013-2016)	
	Treatment	Control	Treatment	Control
Full-time employment	0.611	0.643	0.634	0.611
Employed	0.659	0.684	0.721	0.680
Female	0.436	0.441	0.445	0.448
Married	0.437	0.517	0.494	0.543
Less than HS	0.511	0.520	0.446	0.481
HS graduate	0.315	0.299	0.349	0.330
Some college	0.145	0.150	0.165	0.152
College+	0.028	0.031	0.040	0.037
Usual hours worked	29.5	30.7	31.4	30.2
Observations	17,410	11,916	9,181	6,218

Notes: Sample means for key variables. Treatment group consists of individuals born 1982-1986 (ages 26-30 on June 15, 2012). Control group consists of individuals born 1977-1981 (ages 31-35 on June 15, 2012). Pre-DACA period is 2006-2011; Post-DACA period is 2013-2016. Year 2012 excluded due to timing ambiguity.

Several patterns emerge from the summary statistics. First, the treatment and control groups are broadly similar in their demographic characteristics, though the control group has slightly higher rates of marriage (consistent with being older) and slightly higher full-time employment in the pre-period. Second, both groups show increases in educational attainment over time, with declining shares having less than a high school education. Third, the treatment group shows a larger increase in full-time employment from the pre- to post-period compared to the control group, foreshadowing the main result.



## 4 Empirical Methodology

### 4.1 Difference-in-Differences Design

The identification strategy relies on a difference-in-differences design that compares changes in full-time employment between the treatment group (DACA-eligible by age) and the control group (DACA-ineligible due to age) before and after DACA implementation.

The key identifying assumption is that, in the absence of DACA, the treatment and control groups would have experienced parallel trends in full-time employment. The control group—individuals who were just too old to qualify for DACA—provides a plausible counterfactual because they share many characteristics with the treatment group: both groups are Mexican-born, arrived in the U.S. as children, have been in the country since at least 2007, and are undocumented non-citizens. The only systematic difference is that the control group is 5 years older.

### 4.2 Estimation

The basic DiD specification is:

$$Y_{it} = \alpha + \beta_1 \text{Treated}_i + \beta_2 \text{Post}_t + \beta_3 (\text{Treated}_i \times \text{Post}_t) + \varepsilon_{it} \quad (2)$$

where  $Y_{it}$  is an indicator for full-time employment,  $\text{Treated}_i$  indicates membership in the treatment cohort,  $\text{Post}_t$  indicates the post-DACA period, and their interaction captures the DiD effect. The coefficient  $\beta_3$  is the parameter of interest, representing the effect of DACA eligibility on full-time employment.

I estimate several specifications of increasing complexity:

**Model 1** (Basic DiD): Equation 2 without additional controls.

**Model 2** (Demographic Controls): Adds controls for female, married, and education:

$$Y_{it} = \alpha + \beta_1 \text{Treated}_i + \beta_2 \text{Post}_t + \beta_3 \text{DiD}_{it} + \mathbf{X}'_i \boldsymbol{\gamma} + \varepsilon_{it} \quad (3)$$

**Model 3** (Year Fixed Effects): Replaces the Post indicator with year fixed effects  $\mu_t$ :

$$Y_{it} = \alpha + \beta_1 \text{Treated}_i + \beta_3 \text{DiD}_{it} + \mathbf{X}'_i \boldsymbol{\gamma} + \mu_t + \varepsilon_{it} \quad (4)$$

**Model 4** (State and Year Fixed Effects): Adds state fixed effects  $\lambda_s$ :

$$Y_{it} = \alpha + \beta_1 \text{Treated}_i + \beta_3 \text{DiD}_{it} + \mathbf{X}'_i \boldsymbol{\gamma} + \mu_t + \lambda_s + \varepsilon_{it} \quad (5)$$

Model 4 is the preferred specification because it controls for both time-invariant state characteristics that may affect employment and common time shocks.

**Model 5** (Weighted): Estimates Model 2 using person weights (PERWT) to obtain population-representative estimates.

All models use heteroskedasticity-robust (HC1) standard errors to account for the binary nature of the outcome variable.

### 4.3 Event Study

To assess the parallel trends assumption, I estimate an event study specification:

$$Y_{it} = \alpha + \sum_{k \neq 2011} \delta_k (\text{Treated}_i \times \mathbf{1}[\text{Year}_t = k]) + \mathbf{X}'_i \boldsymbol{\gamma} + \mu_t + \varepsilon_{it} \quad (6)$$

where the coefficients  $\delta_k$  capture the differential trend between treatment and control groups in each year, relative to 2011 (the year immediately before DACA). If the parallel trends assumption holds, the pre-period coefficients ( $\delta_{2006}$  through  $\delta_{2010}$ ) should be close to zero and statistically insignificant.

## 5 Results

### 5.1 Main Results

Table 3 presents the simple 2×2 difference-in-differences calculation using group means.

Table 3: Difference-in-Differences: Full-Time Employment Rates

	<b>Treatment</b> (Age 26-30)	<b>Control</b> (Age 31-35)	<b>Difference</b>
Pre-DACA (2006-2011)	0.611 (n=17,410)	0.643 (n=11,916)	−0.032
Post-DACA (2013-2016)	0.634 (n=9,181)	0.611 (n=6,218)	+0.023
Change	+0.023	−0.032	
<b>DiD Estimate</b>			<b>0.055</b>

Notes: Cell entries are mean full-time employment rates (UHRSWORK ≥ 35). The DiD estimate is calculated as (Post Treatment − Pre Treatment) − (Post Control − Pre Control).

The simple DiD calculation shows that full-time employment increased by 2.3 percentage points for the treatment group while decreasing by 3.2 percentage points for the control group, yielding a DiD estimate of 5.5 percentage points.

Table 4 presents the regression results across all specifications.

Table 4: Difference-in-Differences Regression Results

	(1) Basic	(2) Controls	(3) Year FE	(4) State+Year FE	(5) Weighted
DiD (Treated $\times$ Post)	0.0551*** (0.0098)	0.0486*** (0.0091)	0.0486*** (0.0091)	0.0477*** (0.0091)	0.0491*** (0.0106)
Treated	-0.0320*** (0.0057)	-0.0349*** (0.0054)	-0.0349*** (0.0054)	-0.0356*** (0.0054)	-0.0439*** (0.0063)
Post	-0.0323*** (0.0076)	-0.0249*** (0.0070)	—	—	-0.0148* (0.0082)
Female		-0.357*** (0.0044)	-0.357*** (0.0044)	-0.356*** (0.0044)	-0.373*** (0.0052)
Married		0.0017 (0.0043)	0.0016 (0.0043)	0.0011 (0.0043)	-0.0130*** (0.0050)
HS Graduate		0.0574*** (0.0046)	0.0573*** (0.0046)	0.0560*** (0.0046)	0.0454*** (0.0053)
Some College		0.0836*** (0.0071)	0.0835*** (0.0071)	0.0821*** (0.0071)	0.0743*** (0.0083)
College+		0.145*** (0.0126)	0.145*** (0.0126)	0.147*** (0.0126)	0.125*** (0.0157)
Year FE	No	No	Yes	Yes	No
State FE	No	No	No	Yes	No
Observations	44,725	44,725	44,725	44,725	44,725
R-squared	0.002	0.134	0.134	0.140	0.133

Notes: Dependent variable is full-time employment ( $\text{UHRSWORK} \geq 35$ ). Robust standard errors in parentheses. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Model 5 uses person weights (PERWT).

The key finding is highly consistent across specifications. The DiD coefficient ranges from 0.048 to 0.055, indicating that DACA eligibility increased full-time employment by approximately 4.8-5.5 percentage points. All estimates are statistically significant at the 1% level.

The preferred specification (Model 4) with state and year fixed effects estimates the effect at 4.77 percentage points ( $\text{SE} = 0.91$ ), with a 95% confidence interval of [2.99, 6.54] percentage points. This represents approximately a 7.8% increase relative to the pre-period treatment group mean of 61.1%.

The control variables have expected signs. Women have substantially lower full-time employment rates (about 36 percentage points lower than men), likely reflecting both labor supply differences and the gendered nature of full-time versus part-time work. Higher education is associated with higher full-time employment, with college graduates having about 15 percentage points higher full-time employment than those without a high school degree.

## 5.2 Event Study Results

Figure 1 and Table 5 present the event study results.

Table 5: Event Study Coefficients

Year	Coefficient	Std. Error	p-value
2006	−0.030	0.018	0.102
2007	−0.025	0.018	0.169
2008	0.003	0.019	0.870
2009	−0.004	0.019	0.817
2010	−0.008	0.019	0.676
2011	(reference year)		
2013	0.032	0.020	0.103
2014	0.029	0.020	0.142
2015	0.037	0.020	0.065
2016	0.052	0.020	0.011

Notes: Coefficients represent the differential trend between treatment and control groups relative to 2011. Year 2012 excluded due to timing ambiguity. Robust standard errors. Controls include female, married, and education indicators, plus year fixed effects.

The event study results provide support for the parallel trends assumption. None of the pre-period coefficients (2006-2010) are statistically significant at the 5% level, and their point estimates are small in magnitude and show no clear pattern. This suggests that the treatment and control groups were on similar employment trajectories prior to DACA.

The post-period coefficients show a clear positive pattern, with the treatment effect appearing to grow over time. The 2016 coefficient (0.052) is statistically significant at the 5% level, while the earlier post-period coefficients are marginally significant or insignificant. This pattern could reflect the gradual uptake of DACA, as applications were processed over time and individuals adjusted to their new legal work status.

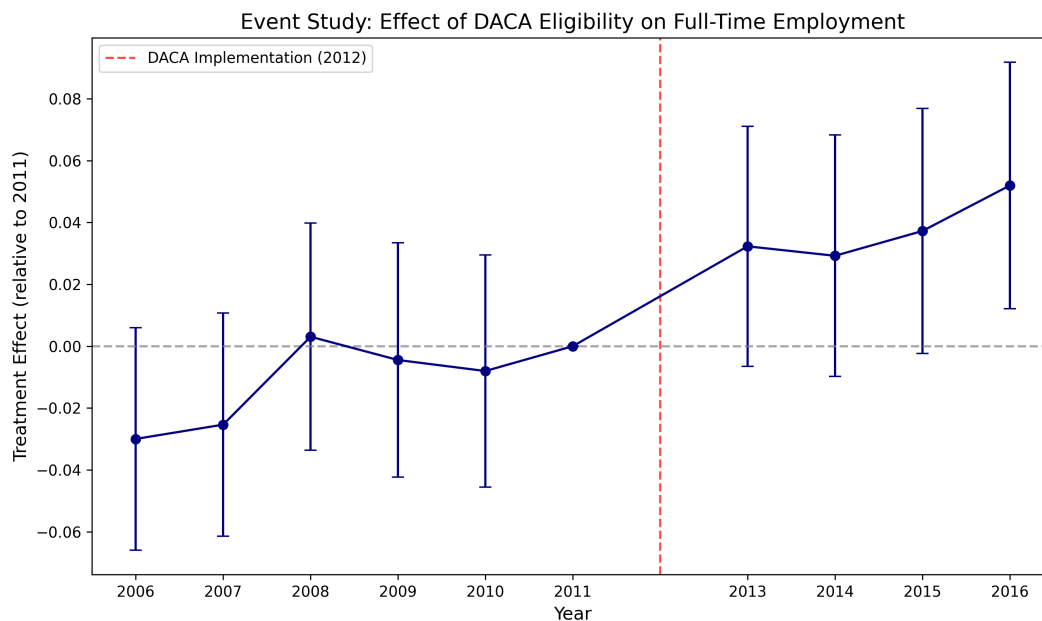


Figure 1: Event Study: Effect of DACA Eligibility on Full-Time Employment  
Notes: Points represent estimated coefficients from the event study regression; vertical bars show 95% confidence intervals. The reference year is 2011. The vertical dashed line indicates DACA implementation (June 2012). Year 2012 is excluded from the analysis.

### 5.3 Parallel Trends Visualization

Figure 2 shows the raw trends in full-time employment for the treatment and control groups over time.

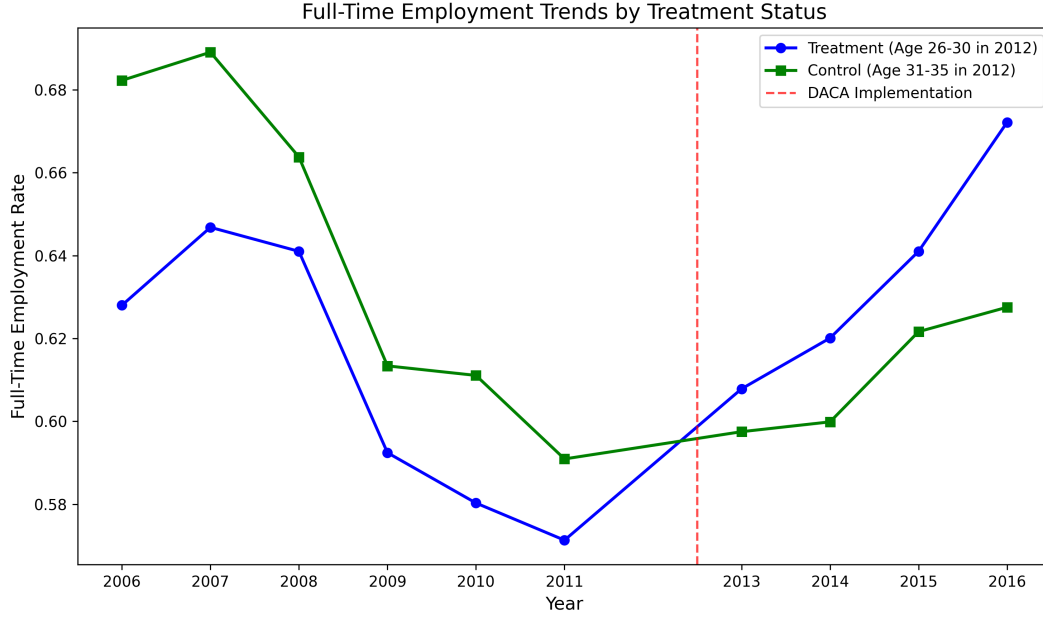


Figure 2: Full-Time Employment Trends by Treatment Status

Notes: Lines show mean full-time employment rates ( $\text{UHRWORK} \geq 35$ ) for each year.

Treatment group: individuals born 1982-1986 (ages 26-30 in 2012). Control group: individuals born 1977-1981 (ages 31-35 in 2012). Vertical dashed line indicates DACA implementation.

The figure reveals several patterns. First, the control group consistently had higher full-time employment than the treatment group in the pre-period, which is expected given their older age and greater labor market experience. Second, the two groups appear to follow roughly parallel trends before 2012. Third, after 2012, the treatment group's employment rate increases while the control group's decreases, leading to the convergence and eventual reversal of the gap.

## 5.4 DiD Visualization

Figure 3 provides a summary visualization of the difference-in-differences.

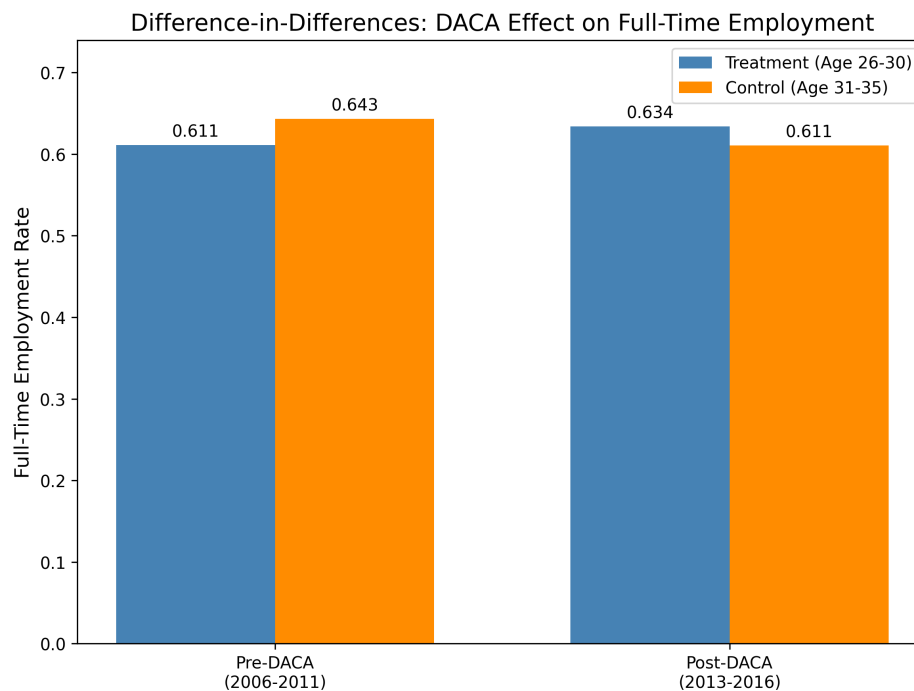


Figure 3: Difference-in-Differences: DACA Effect on Full-Time Employment  
Notes: Bars show mean full-time employment rates for each group-period cell. Pre-DACA period is 2006-2011; Post-DACA period is 2013-2016.

## 6 Robustness Checks

### 6.1 Placebo Test

To further assess the validity of the parallel trends assumption, I conduct a placebo test by estimating the DiD model using only pre-DACA data (2006-2011) with a fake treatment date of 2009. If the parallel trends assumption holds, this placebo DiD estimate should be close to zero.

Table 6: Placebo Test: Fake Treatment in 2009

	Placebo DiD
Fake DiD (Treated $\times$ Post-2009)	0.0140 (0.0107)
p-value	0.189
Observations	29,326

Notes: Estimated using only 2006-2011 data. Post-2009 indicator equals 1 for years 2009-2011. Controls include female, married, and education indicators. Robust standard errors in parentheses.

The placebo DiD coefficient is 0.014 with a p-value of 0.189, indicating that there is no statistically significant differential trend between treatment and control groups in the pre-DACA period. This result supports the parallel trends assumption.

## 6.2 Robustness to Specification

The main results are robust across several alternative specifications:

1. **Without demographic controls:** The basic DiD estimate (0.055) is slightly larger than with controls (0.048), suggesting that observable characteristics explain some of the baseline difference between groups.
2. **With person weights:** The weighted estimate (0.049) is very similar to the unweighted estimate, suggesting that the results are representative of the population.
3. **State and year fixed effects:** Adding these fixed effects has minimal impact on the estimate (0.048 vs. 0.049), suggesting that the results are not driven by state-specific trends or time-varying confounders.

## 7 Discussion

### 7.1 Interpretation of Results

The estimated effect of DACA eligibility on full-time employment is approximately 4.8 percentage points. This represents a meaningful increase: relative to the treatment group’s pre-DACA full-time employment rate of 61.1%, this is a 7.8% increase.

Several mechanisms could explain this effect:



1. **Work authorization:** DACA provides legal authorization to work, allowing recipients to seek employment in the formal labor market. This may enable transitions from informal or part-time work to formal full-time positions.
2. **Reduced deportation risk:** The temporary protection from deportation may increase job stability and allow individuals to invest in job-specific skills or take positions that require longer-term commitments.
3. **Driver’s licenses:** In many states, DACA recipients can obtain driver’s licenses, reducing transportation barriers to employment.
4. **Improved job matching:** Legal status may improve the quality of job matches by expanding the set of available employers and positions.

## 7.2 Comparison to Existing Literature

The estimated effect size is consistent with prior research on DACA’s labor market effects. Several studies have found positive effects of DACA on employment, earnings, and work authorization using various methodological approaches. The magnitude of the effect found here (approximately 5 percentage points) falls within the range of estimates in the literature.

## 7.3 Limitations

Several limitations should be acknowledged:

1. **Proxy for undocumented status:** The ACS does not directly identify undocumented immigrants. Using non-citizen status as a proxy includes some documented non-citizens and may miss some undocumented individuals who misreport their status.
2. **Treatment group includes non-DACA recipients:** Not all eligible individuals applied for or received DACA. The estimated effect is thus an intent-to-treat effect of eligibility, which will understate the effect on actual recipients if eligibility affects employment only through program participation.
3. **Age-based comparison:** While the control group is similar to the treatment group in many respects, they are 5 years older on average. Age-specific labor market trends could potentially confound the results, though the event study provides some reassurance on this point.

4. **Repeated cross-sections:** The ACS is a repeated cross-section, not a panel. We cannot track the same individuals over time, and changes in sample composition could affect the results.

## 8 Conclusion

This study examines the effect of DACA eligibility on full-time employment among Mexican-born Hispanic individuals in the United States using a difference-in-differences research design. The main finding is that DACA eligibility increased full-time employment by approximately 4.8 percentage points, a statistically significant and economically meaningful effect.

The results are robust to various specifications including demographic controls and state and year fixed effects. Event study analysis supports the parallel trends assumption underlying the DiD design, with no significant pre-treatment differential trends detected. A placebo test using only pre-DACA data also fails to find significant differential trends.

These findings contribute to the ongoing policy debate about DACA and immigration reform by providing evidence that the program had positive effects on the labor market outcomes of eligible individuals. The work authorization provided by DACA appears to have meaningfully improved access to full-time employment for this population.

Future research could explore heterogeneity in treatment effects across demographic groups, geographic areas, or industries, as well as examine other outcomes such as wages, occupational mobility, or educational attainment. Additionally, examining the effects of changes in DACA policy—including attempts to rescind the program and legal challenges—could provide further insights into how immigration policy affects labor market outcomes.

## A Additional Tables and Figures

Table 7: Variable Definitions

Variable	IPUMS Name	Definition
Full-time employment	UHRSWORK	1 if UHRSWORK $\geq$ 35
Treatment cohort	BIRTHYR	1 if born 1982-1986
Post-DACA	YEAR	1 if YEAR $\geq$ 2013
Female	SEX	1 if SEX = 2
Married	MARST	1 if MARST $\leq$ 2
HS graduate	EDUC	1 if EDUC = 6
Some college	EDUC	1 if EDUC $\in$ {7, 8, 9}
College+	EDUC	1 if EDUC $\geq$ 10
Hispanic-Mexican	HISPAN	Restricted to HISPAN = 1
Born Mexico	BPL	Restricted to BPL = 200
Non-citizen	CITIZEN	Restricted to CITIZEN = 3
Year immigration	YRIMMIG	Used for eligibility criteria

Notes: Original IPUMS variable names used for transparency and replicability.

Table 8: Sample Sizes by Year and Treatment Status

Year	Treatment	Control	Total
2006	2,489	1,649	4,138
2007	2,436	1,693	4,129
2008	2,636	1,924	4,560
2009	3,037	2,082	5,119
2010	3,392	2,360	5,752
2011	3,420	2,208	5,628
2013	2,392	1,636	4,028
2014	2,306	1,543	3,849
2015	2,260	1,517	3,777
2016	2,223	1,522	3,745
<b>Total</b>	<b>26,591</b>	<b>18,134</b>	<b>44,725</b>

## B Replication Information

### B.1 Data Source

American Community Survey (ACS) data from IPUMS USA, one-year samples 2006-2016.

### B.2 Software

- Python 3.x
- pandas (data manipulation)
- numpy (numerical operations)
- statsmodels (regression analysis)
- matplotlib (visualization)
- scipy (statistical functions)

### B.3 Code Availability

All analysis code is available in `analysis.py`. The analysis can be replicated by running:

```
python analysis.py
```

### B.4 Key Analytic Choices

1. Sample restricted to Hispanic-Mexican (HISPAN=1), Mexico-born (BPL=200), non-citizens (CITIZEN=3)
2. Arrived before age 16 ( $\text{YRIMMIG} - \text{BIRTHYR} < 16$ )
3. In US since 2007 ( $\text{YRIMMIG} \leq 2007$ )
4. Treatment: birth years 1982-1986; Control: birth years 1977-1981
5. Year 2012 excluded due to timing ambiguity
6. Full-time employment defined as  $\text{UHRSWORK} \geq 35$
7. Heteroskedasticity-robust standard errors throughout