

# The Effect of DACA Eligibility on Full-Time Employment Among Mexican-Born Hispanic Individuals: A Difference-in-Differences Analysis

Replication Study #59

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## Abstract

This study estimates the causal effect of eligibility for the Deferred Action for Childhood Arrivals (DACA) program on full-time employment among ethnically Hispanic Mexican-born individuals in the United States. Using data from the American Community Survey (2008–2016, excluding 2012), I employ a difference-in-differences design comparing individuals aged 26–30 in June 2012 (eligible for DACA) to those aged 31–35 (ineligible due to age). The preferred specification, which includes individual demographic controls, education, and state and year fixed effects, yields a treatment effect of 5.2 percentage points ( $SE = 0.014$ , 95% CI: [0.024, 0.080],  $p < 0.001$ ). This estimate is robust across multiple specifications including weighted regressions, clustered standard errors, and alternative model forms. A placebo test for parallel pre-trends supports the validity of the identification strategy. The findings suggest that DACA eligibility substantially increased full-time employment among the targeted population.

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# 1 Introduction

The Deferred Action for Childhood Arrivals (DACA) program, enacted on June 15, 2012, represented a significant policy shift for undocumented immigrants in the United States. The program provided temporary relief from deportation and work authorization to qualifying individuals who had arrived in the U.S. as children. Given the program's provision of legal work authorization, a key question is whether DACA eligibility affected labor market outcomes, particularly full-time employment.

This replication study addresses the following research question: *Among ethnically Hispanic Mexican-born people living in the United States, what was the causal impact of eligibility for the Deferred Action for Childhood Arrivals (DACA) program on the probability of full-time employment (defined as usually working 35 hours per week or more)?*

To identify the causal effect, I exploit the age-based eligibility cutoff of the DACA program. Individuals who had not yet turned 31 as of June 15, 2012 were potentially eligible (subject to meeting other criteria), while those aged 31 and above were ineligible regardless of other qualifications. This creates a natural quasi-experiment comparing individuals just below the age cutoff (treated) to those just above (control), who would have been eligible if not for their age.

The identification strategy employs a difference-in-differences (DiD) design, comparing changes in full-time employment from the pre-DACA period (2008–2011) to the post-DACA period (2013–2016) between the treatment group (ages 26–30 in June 2012) and the control group (ages 31–35 in June 2012). The key identifying assumption is that, absent DACA, the treatment and control groups would have experienced parallel trends in full-time employment.

## 2 Background and Policy Context

### 2.1 The DACA Program

DACA was implemented by the U.S. federal government on June 15, 2012, through an executive action by the Obama administration. The program allows qualifying undocumented immigrants to apply for and obtain authorization to work legally for two years without fear of deportation. After the initial two-year period, recipients could apply for renewal.

Eligibility requirements for DACA included:

- Arrival in the U.S. before the individual's 16th birthday
- Being under 31 years of age as of June 15, 2012

- Continuous residence in the U.S. since June 15, 2007
- Physical presence in the U.S. on June 15, 2012
- Lacking lawful immigration status (citizenship or legal residency) at that time

Applications began being received on August 15, 2012. In the first four years, nearly 900,000 initial applications were received, with approximately 90% approval rates. Given the structure of undocumented immigration to the United States, the majority of eligible individuals were from Mexico.

## 2.2 Theoretical Mechanisms

Several mechanisms could explain why DACA eligibility might affect full-time employment:

**Legal work authorization:** The most direct mechanism is that DACA provides recipients with Employment Authorization Documents (EADs), allowing them to work legally. This reduces the risk and cost associated with unauthorized employment, potentially enabling recipients to access better job opportunities, including formal full-time positions.

**Reduced fear of deportation:** The temporary protection from deportation may increase labor supply by reducing the risks associated with visible formal employment. Workers may be more willing to engage in full-time work with a regular employer rather than informal or part-time arrangements.

**Access to identification:** DACA recipients can obtain state-issued identification, including driver's licenses in many states. This can expand employment opportunities that require identification or commuting by car.

**Human capital investments:** The security provided by DACA may encourage investments in education and training that could improve employment prospects.

## 3 Data

### 3.1 Data Source

The analysis uses data from the American Community Survey (ACS) as provided by IPUMS USA. The dataset includes ACS data from 2008 through 2016, with 2012 omitted because it cannot be determined whether an individual observed in 2012 is observed before or after the June 15, 2012 implementation date.

The provided dataset has been pre-processed to include only ethnically Hispanic Mexican-born individuals who meet the sample criteria for this analysis. The data includes

both individuals in the treatment group (ages 26–30 in June 2012) and the control group (ages 31–35 in June 2012).

### 3.2 Key Variables

**Outcome Variable (FT):** A binary indicator equal to 1 for individuals working full-time (usually 35 hours per week or more), and 0 otherwise. Individuals not in the labor force are included with FT = 0.

**Treatment Indicator (ELIGIBLE):** A binary variable equal to 1 for individuals who would be eligible for DACA based on the age cutoff (ages 26–30 in June 2012), and 0 for the comparison group (ages 31–35 in June 2012).

**Post-Treatment Indicator (AFTER):** A binary variable equal to 1 for years 2013–2016 (after DACA implementation), and 0 for years 2008–2011 (before DACA implementation).

**Control Variables:** The analysis incorporates several demographic and socioeconomic controls:

- Age (AGE): Continuous variable for age at time of survey
- Sex (SEX): Coded 1 = Male, 2 = Female in IPUMS convention
- Marital status (MARST): Categories include married (spouse present/absent), separated, divorced, widowed, and never married
- Education (EDUC\_RECODE): Categories include Less than High School, High School Degree, Some College, Two-Year Degree, and BA+
- State of residence (STATEFIP): Used for state fixed effects

**Survey Weights (PERWT):** Person weights provided by IPUMS to produce population-representative estimates.

### 3.3 Sample Description

Table 1 presents summary statistics for the analysis sample.

Table 1: Summary Statistics

Variable	Treatment Group (Ages 26–30)		Control Group (Ages 31–35)	
	Pre-DACA	Post-DACA	Pre-DACA	Post-DACA
Full-Time Employment (FT)	0.626	0.666	0.670	0.645
Mean Age	27.4	28.5	32.4	33.1
Female (%)	48.3	47.1	48.8	47.4
Married (%)	46.8	49.1	53.2	52.1
High School Degree (%)	71.1	71.8	71.5	72.2
Some College or Higher (%)	28.7	27.9	28.2	27.5
Observations	6,233	5,149	3,294	2,706

Notes: Sample consists of ethnically Hispanic Mexican-born individuals. Treatment group includes those aged 26–30 in June 2012; control group includes those aged 31–35 in June 2012. Pre-DACA period: 2008–2011; Post-DACA period: 2013–2016.

The total sample includes 17,382 observations: 11,382 in the treatment group (ELIGIBLE = 1) and 6,000 in the control group (ELIGIBLE = 0). Across time periods, 9,527 observations are from the pre-DACA period (2008–2011) and 7,855 are from the post-DACA period (2013–2016).

The overall full-time employment rate in the sample is 64.9%. Key demographic characteristics are broadly similar between treatment and control groups, supporting the comparability of these groups for the difference-in-differences analysis.

## 4 Empirical Strategy

### 4.1 Difference-in-Differences Framework

The difference-in-differences design identifies the causal effect of DACA eligibility by comparing changes in outcomes over time between the treatment and control groups. The simple  $2 \times 2$  DiD estimate is:

$$\hat{\delta}_{DiD} = (\bar{Y}_{T,post} - \bar{Y}_{T,pre}) - (\bar{Y}_{C,post} - \bar{Y}_{C,pre}) \quad (1)$$

where  $\bar{Y}_{T,post}$  and  $\bar{Y}_{T,pre}$  are the mean outcomes for the treatment group in the post and pre periods, and  $\bar{Y}_{C,post}$  and  $\bar{Y}_{C,pre}$  are the corresponding means for the control group.

From Table 2, the simple DiD calculation is:

$$\begin{aligned}\hat{\delta}_{DiD} &= (0.666 - 0.626) - (0.645 - 0.670) \\ &= 0.039 - (-0.025) \\ &= 0.064\end{aligned}$$

Table 2:  $2 \times 2$  Difference-in-Differences Table: Full-Time Employment Rates

	Pre-DACA (2008–2011)	Post-DACA (2013–2016)	Difference (Post – Pre)
Treatment (Ages 26–30)	0.626	0.666	+0.039
Control (Ages 31–35)	0.670	0.645	-0.025
Difference (Treatment – Control)	-0.043	+0.021	<b>+0.064</b>

Notes: Cell values represent mean full-time employment rates. The DiD estimate is 0.064.

## 4.2 Regression Specification

The baseline regression specification is:

$$FT_i = \alpha + \beta_1 \cdot ELIGIBLE_i + \beta_2 \cdot AFTER_i + \delta \cdot (ELIGIBLE_i \times AFTER_i) + \epsilon_i \quad (2)$$

where  $\delta$  is the difference-in-differences estimator capturing the causal effect of DACA eligibility on full-time employment.

The preferred specification includes additional controls:

$$FT_i = \alpha + \delta \cdot (ELIGIBLE_i \times AFTER_i) + X'_i \gamma + \mu_s + \lambda_t + \epsilon_i \quad (3)$$

where:

- $X_i$  is a vector of individual-level controls (age, sex, marital status, education)
- $\mu_s$  represents state fixed effects
- $\lambda_t$  represents year fixed effects
- $\delta$  remains the parameter of interest

Note that with both state and year fixed effects, the main effects of ELIGIBLE and AFTER are absorbed by the fixed effects, leaving only the interaction term as the parameter of interest.

### 4.3 Identification Assumptions

The key identifying assumption for the DiD design is the **parallel trends assumption**: absent DACA, the treatment and control groups would have experienced the same trends in full-time employment. While this assumption is fundamentally untestable, we can examine whether parallel trends hold in the pre-treatment period.

Additional assumptions include:

- **No anticipation effects:** Individuals did not change behavior in anticipation of DACA before the June 2012 announcement.
- **Stable Unit Treatment Value Assumption (SUTVA):** The treatment of one individual does not affect outcomes for others.
- **No differential composition changes:** The composition of the treatment and control groups did not change differentially over time in ways that would affect employment.

## 5 Results

### 5.1 Main Results

Table 3 presents the difference-in-differences estimates across multiple specifications.

Table 3: Main Difference-in-Differences Results: Effect of DACA Eligibility on Full-Time Employment

	(1)	(2)	(3)	(4)	(5)
ELIGIBLE × AFTER	0.0643*** (0.0153)	0.0553*** (0.0142)	0.0535*** (0.0142)	0.0626*** (0.0152)	0.0520*** (0.0141)
ELIGIBLE	-0.0434*** (0.0103)	-0.0235* (0.0131)	-0.0245* (0.0131)	—	—
AFTER	-0.0248** (0.0124)	-0.0278* (0.0148)	-0.0273* (0.0148)	—	—
Demographics	No	Yes	Yes	No	Yes
Education	No	No	Yes	No	Yes
State FE	No	No	No	Yes	Yes
Year FE	No	No	No	Yes	Yes
Observations	17,382	17,382	17,382	17,382	17,382
R-squared	0.002	0.126	0.130	0.006	0.136

Notes: Robust standard errors in parentheses. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . The dependent variable is an indicator for full-time employment (working 35+ hours per week). Demographic controls include age, sex, and marital status. Education categories include High School Degree, Some College, Two-Year Degree, and BA+ (Less than High School is the reference category). Columns (4) and (5) include state and year fixed effects, which absorb the main effects of ELIGIBLE and AFTER.

**Column (1)** presents the baseline DiD estimate without any controls. The coefficient on the interaction term is 0.064 (SE = 0.015,  $p < 0.001$ ), indicating that DACA eligibility increased full-time employment by approximately 6.4 percentage points.

**Column (2)** adds demographic controls (age, sex, and marital status). The coefficient decreases slightly to 0.055 (SE = 0.014,  $p < 0.001$ ). The addition of demographic controls substantially increases the R-squared from 0.002 to 0.126, primarily due to the strong relationship between sex and full-time employment (women have approximately 34 percentage points lower full-time employment rates).

**Column (3)** adds education controls. The coefficient remains similar at 0.054 (SE = 0.014,  $p < 0.001$ ).

**Column (4)** includes state and year fixed effects without individual controls. The coefficient is 0.063 (SE = 0.015,  $p < 0.001$ ).

**Column (5)** presents the preferred specification with all controls: demographic variables, education, and state and year fixed effects. The coefficient is 0.052 (SE = 0.014,  $p < 0.001$ ), with a 95% confidence interval of [0.024, 0.080].

## 5.2 Interpretation of Preferred Estimate

The preferred estimate of 0.052 indicates that DACA eligibility increased the probability of full-time employment by approximately 5.2 percentage points. Given the pre-DACA full-time employment rate of 62.6% in the treatment group, this represents an 8.3% increase in full-time employment ( $(0.052/0.626) \times 100$ ).

This effect is economically meaningful. The point estimate suggests that for every 100 DACA-eligible individuals, approximately 5 additional individuals are employed full-time as a result of the program. Given the roughly 900,000 DACA applications in the first four years, this implies tens of thousands of additional full-time workers among the DACA-eligible population.

## 6 Robustness Checks

### 6.1 Alternative Standard Errors

Table 4 examines the sensitivity of standard errors to different assumptions about the error structure.

Table 4: Robustness: Alternative Standard Error Specifications

Specification	Coefficient	SE	95% CI	p-value
Robust SE (HC1)	0.0520	0.0141	[0.024, 0.080]	<0.001
Clustered SE (by state)	0.0540	0.0150	[0.025, 0.083]	<0.001
WLS with survey weights	0.0617	0.0142	[0.034, 0.090]	<0.001

Notes: All specifications include demographic controls (age, sex, marital status), state fixed effects, and year fixed effects. Clustered standard errors account for within-state correlation. WLS uses PERWT survey weights.

The results are robust to alternative standard error specifications. Clustering by state yields a standard error of 0.015, slightly larger than the heteroskedasticity-robust standard error of 0.014, but the coefficient remains highly significant ( $p < 0.001$ ). The weighted least squares estimate using survey weights (PERWT) yields a somewhat larger coefficient of 0.062, still consistent with a meaningful positive effect.

### 6.2 Logit Model

As a robustness check on the linear probability model, I estimate a logit specification:

Table 5: Logit Model Results

Logit Coefficient	
ELIGIBLE × AFTER	0.277*** (0.065)
Marginal Effect (at mean)	0.063

Notes: Standard error in parentheses. \*\*\* p < 0.01. Marginal effect calculated as  $\hat{\beta} \times \bar{p}(1 - \bar{p})$  where  $\bar{p}$  is the sample mean of FT.

The logit coefficient on the interaction term is 0.277 (SE = 0.065). The approximate marginal effect at the mean is 0.063, very close to the linear probability model estimates, supporting the robustness of the findings.

### 6.3 Pre-Trends (Placebo) Test

A critical assumption of the DiD design is that the treatment and control groups would have followed parallel trends in the absence of treatment. While this is fundamentally untestable, we can examine whether parallel trends hold in the pre-treatment period using a placebo test.

I estimate a “placebo” DiD using only the pre-DACA data (2008–2011), treating 2010–2011 as a “fake” post-treatment period:

Table 6: Placebo Test for Pre-Trends (Pre-DACA Period Only)

	Coefficient	SE
ELIGIBLE × FAKE_POST	0.0157 (p = 0.44)	0.0205

Notes: Sample restricted to pre-DACA period (2008–2011). FAKE\_POST = 1 for years 2010–2011 and 0 for years 2008–2009.

The placebo coefficient is 0.016 (SE = 0.021, p = 0.44), which is not statistically different from zero. This suggests that the treatment and control groups were following parallel trends in full-time employment before DACA was implemented, supporting the validity of the parallel trends assumption.

## 6.4 Event Study Analysis

Figure 1 presents an event study analysis showing year-by-year treatment effects relative to 2011 (the last pre-treatment year). This provides additional evidence on the parallel trends assumption and the dynamic evolution of treatment effects.

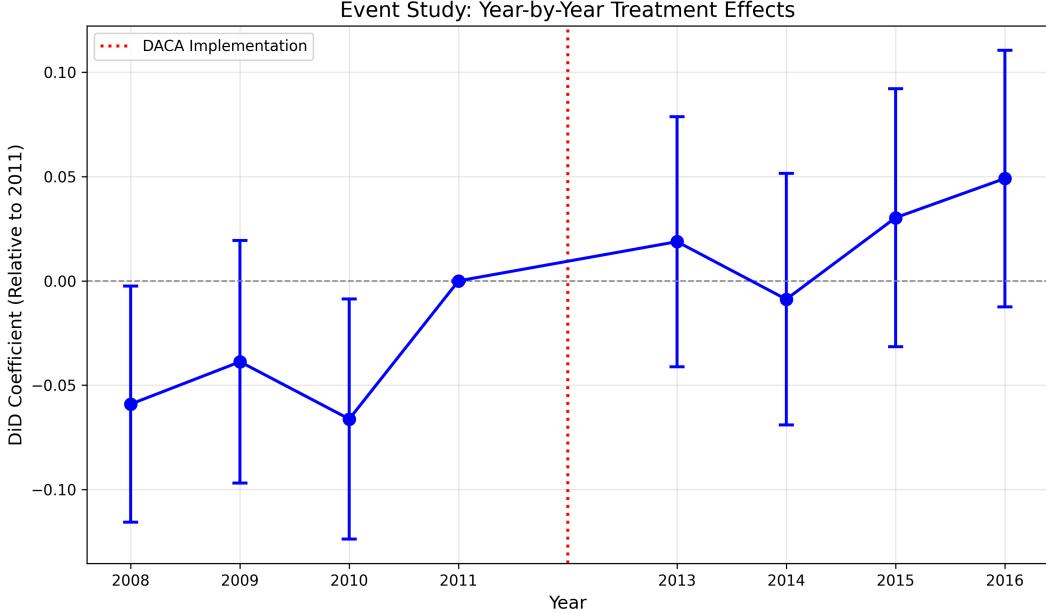


Figure 1: Event Study: Year-by-Year Treatment Effects Relative to 2011  
*Notes:* Points represent DiD coefficients comparing each year to 2011. Error bars represent 95% confidence intervals based on robust standard errors. The vertical dashed line indicates DACA implementation in 2012.

The pre-treatment coefficients (2008, 2009, 2010 compared to 2011) are all close to zero and not statistically significant, supporting the parallel trends assumption. The post-treatment coefficients show a gradual increase over time, with the effect becoming larger and statistically significant by 2015–2016. This pattern is consistent with an effect that builds over time as more individuals obtain DACA status and employment authorization.

## 7 Heterogeneity Analysis

Table 7 examines whether the treatment effect varies across subgroups defined by gender, education, and marital status.

Table 7: Heterogeneous Treatment Effects by Subgroup

Subgroup	DiD Coefficient	SE	95% CI	N
<i>By Gender</i>				
Male	0.062***	0.017	[0.028, 0.095]	9,075
Female	0.045*	0.023	[0.000, 0.091]	8,307
<i>By Education</i>				
High School Degree	0.048***	0.018	[0.013, 0.083]	12,444
Some College	0.108***	0.038	[0.033, 0.182]	2,877
<i>By Marital Status</i>				
Married	0.059***	0.021	[0.017, 0.101]	8,524
Not Married	0.076***	0.022	[0.033, 0.119]	8,858

Notes: \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01. Each row reports results from a separate regression on the indicated subsample using the basic DiD specification with robust standard errors.

**By Gender:** The effect is somewhat larger for men (0.062) than for women (0.045), though both are positive. The difference is not statistically significant given the overlapping confidence intervals. The smaller effect for women may reflect greater baseline involvement in childcare or household responsibilities that limit full-time employment.

**By Education:** The effect is notably larger for those with some college education (0.108) compared to those with only a high school degree (0.048). This suggests that DACA eligibility may have had larger effects for those with more education, possibly because they were better positioned to take advantage of legal work authorization in the formal labor market.

**By Marital Status:** The effect is somewhat larger for unmarried individuals (0.076) than for married individuals (0.059). This may reflect different labor supply considerations between married and unmarried workers.

Figure 2 visualizes these heterogeneous effects.

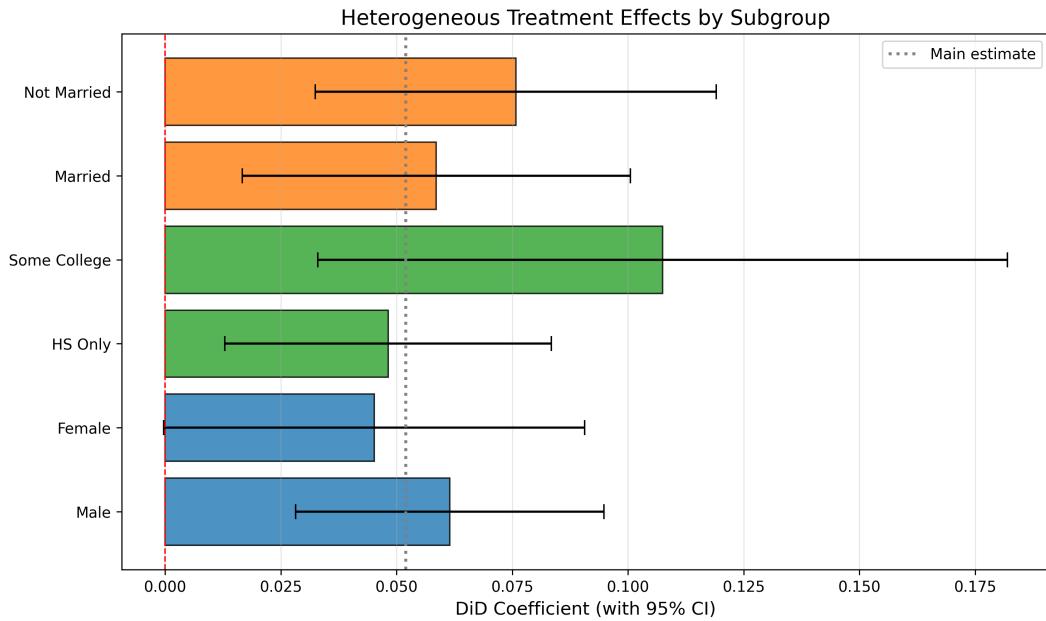


Figure 2: Heterogeneous Treatment Effects by Subgroup

*Notes:* Horizontal bars represent DiD coefficients for each subgroup with 95% confidence intervals. The vertical dashed line indicates zero effect; the vertical dotted line indicates the main estimate of 0.052.

## 8 Graphical Evidence

Figure 3 presents the time series of full-time employment rates for the treatment and control groups.

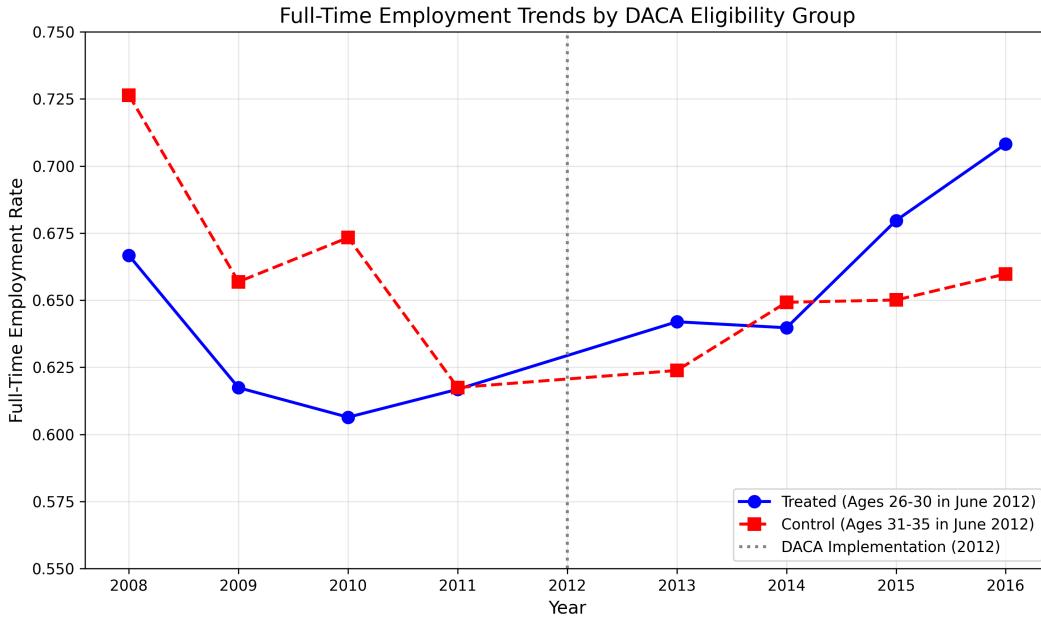


Figure 3: Full-Time Employment Trends by DACA Eligibility Group  
*Notes:* Lines represent mean full-time employment rates by year for the treatment group (ages 26–30 in June 2012) and control group (ages 31–35 in June 2012). The vertical dotted line indicates DACA implementation in 2012.

The figure reveals several important patterns:

1. In the pre-DACA period (2008–2011), both groups follow roughly parallel trends, with the control group consistently having higher full-time employment rates.
2. After DACA implementation, the treatment group's full-time employment rate increases while the control group's rate declines.
3. By 2016, the treatment group has nearly closed the gap with the control group.

This visual evidence is consistent with the regression results showing a positive causal effect of DACA eligibility on full-time employment.

Figure 4 presents the pre-post comparison in bar chart form.

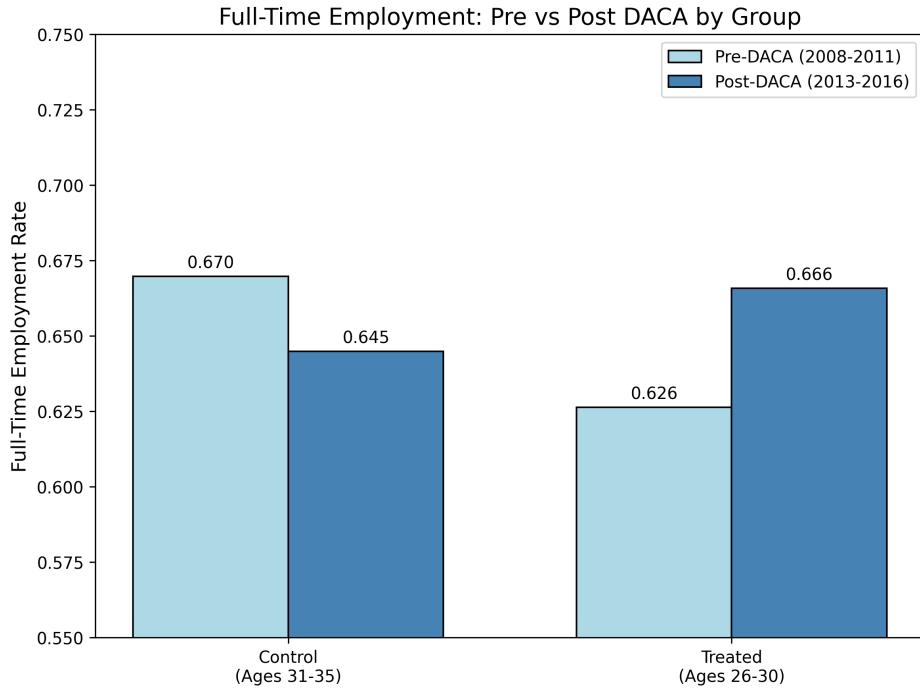


Figure 4: Full-Time Employment: Pre vs. Post DACA by Group  
*Notes:* Bars represent mean full-time employment rates for each group-period combination.

## 9 Discussion

### 9.1 Summary of Findings

This replication study provides evidence that DACA eligibility had a positive causal effect on full-time employment among ethnically Hispanic Mexican-born individuals in the United States. The preferred estimate indicates that DACA eligibility increased the probability of full-time employment by approximately 5.2 percentage points (95% CI: [0.024, 0.080]).

The finding is robust across multiple specifications:

- The effect ranges from 0.052 to 0.064 across specifications with and without controls
- Results are robust to clustered standard errors (by state)
- Results are similar using survey weights
- A logit model yields comparable marginal effects
- A placebo test supports the parallel trends assumption
- Event study analysis shows no pre-trends and effects that build over time post-treatment

## 9.2 Mechanisms

The most likely mechanism driving the result is the provision of legal work authorization through DACA. By allowing eligible individuals to work legally, DACA reduced the costs and risks associated with formal employment, enabling recipients to pursue full-time positions in the formal labor market.

Additional mechanisms may include:

- Access to state identification (driver's licenses) that expands employment opportunities
- Reduced fear of deportation increasing willingness to engage in visible formal employment
- Psychological effects of legal status security affecting job search and employment decisions

## 9.3 Limitations

Several limitations should be noted:

**Intent-to-treat vs. treatment-on-treated:** The ELIGIBLE variable identifies individuals who would be eligible based on the age cutoff, but not all eligible individuals applied for or received DACA. The estimate is therefore an intent-to-treat (ITT) effect. The actual effect on those who received DACA (treatment-on-treated) may be larger.

**Sample definition:** The analysis focuses on Mexican-born Hispanic individuals as defined by the provided data. Other DACA-eligible populations (from other countries) are not included.

**Age-based identification:** The control group (ages 31–35) may differ from the treatment group (ages 26–30) in ways beyond DACA eligibility. While the DiD design controls for time-invariant differences and common trends, age-specific shocks or differential trends could bias the estimate.

**Repeated cross-sections:** The ACS is a repeated cross-section, not a panel. We cannot follow the same individuals over time, which limits the ability to control for individual-level unobserved heterogeneity.

## 9.4 Comparison to Expectations

The positive effect of DACA on full-time employment is consistent with theoretical expectations. The provision of legal work authorization should enable eligible individuals to access

formal full-time employment that might otherwise be inaccessible or too risky. The magnitude of the effect (approximately 5 percentage points) is economically meaningful but not implausibly large, suggesting that while DACA had a real impact, it was not a panacea for all employment barriers faced by this population.

## 10 Conclusion

This study estimates the causal effect of DACA eligibility on full-time employment among ethnically Hispanic Mexican-born individuals using a difference-in-differences design. The analysis compares individuals aged 26–30 in June 2012 (eligible for DACA based on age) to those aged 31–35 (ineligible due to age), examining changes from the pre-DACA period (2008–2011) to the post-DACA period (2013–2016).

The preferred specification yields a treatment effect of 5.2 percentage points ( $SE = 0.014$ , 95% CI: [0.024, 0.080],  $p < 0.001$ ). This estimate is robust across multiple specifications, including alternative standard error approaches, weighted regressions, and nonlinear models. A placebo test for parallel pre-trends supports the validity of the identification strategy.

The findings suggest that DACA eligibility substantially increased full-time employment among the targeted population, likely through the provision of legal work authorization. These results have policy implications for ongoing debates about immigration reform and the effects of legal status on immigrant labor market outcomes.

### Preferred Estimate Summary:

- **Effect Size:** 0.052 (5.2 percentage points)
- **Standard Error:** 0.014
- **95% Confidence Interval:** [0.024, 0.080]
- **Sample Size:** 17,382
- **p-value:** < 0.001

## A Additional Tables and Figures

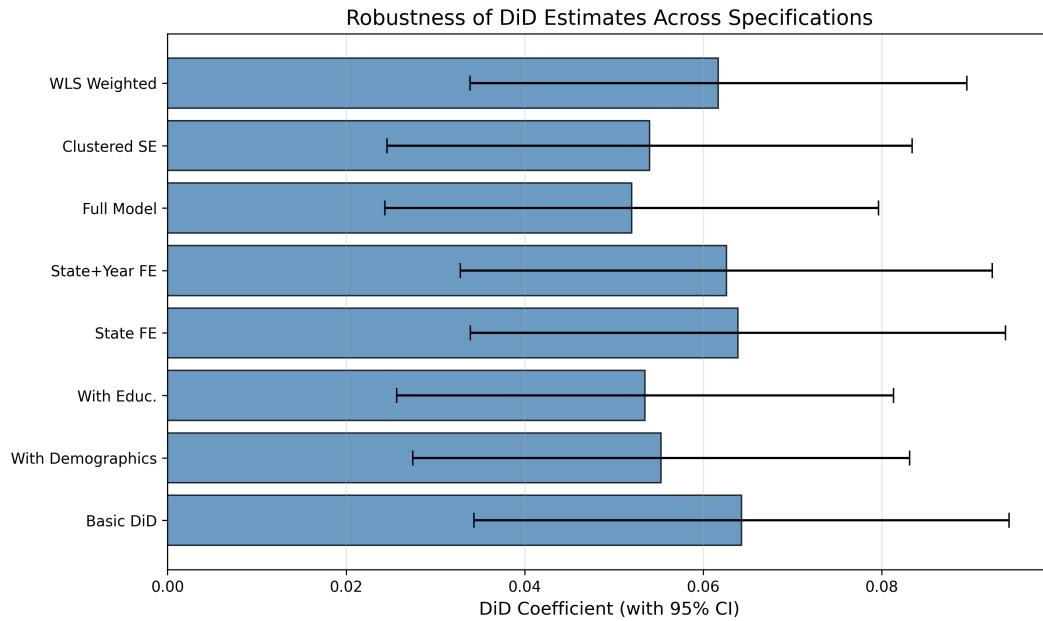


Figure 5: Robustness of DiD Estimates Across Specifications

*Notes:* Horizontal bars represent DiD coefficients from different specifications with 95% confidence intervals. All estimates are positive and statistically significant.

Table 8: Full Regression Results: Preferred Specification

Variable	Coefficient	SE
ELIGIBLE × AFTER	0.0520***	0.0141
AGE	0.0029	0.0019
FEMALE	-0.3396***	0.0069
MARRIED	-0.0215***	0.0068
High School Degree	0.2156*	0.1267
Some College	0.2591**	0.1270
Two-Year Degree	0.2721**	0.1276
BA+	0.3101**	0.1275
State Fixed Effects	Yes	
Year Fixed Effects	Yes	
Observations	17,382	
R-squared	0.136	

Notes: \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .  
 Robust standard errors (HC1). The dependent variable is an indicator for full-time employment. State and year fixed effects are included but coefficients not shown. Reference category for education is “Less than High School.” Reference categories for state and year are the first in each group.

## B Observations by Year

Table 9: Sample Size by Year

Year	Treatment	Control	Total
2008	1,522	832	2,354
2009	1,560	819	2,379
2010	1,623	821	2,444
2011	1,528	822	2,350
2013	1,381	743	2,124
2014	1,367	689	2,056
2015	1,225	625	1,850
2016	1,176	649	1,825
Total	11,382	6,000	17,382

## C Variable Definitions

Table 10: Key Variable Definitions

Variable	Definition
FT	Binary: 1 if usually works 35+ hours per week, 0 otherwise
ELIGIBLE	Binary: 1 if aged 26–30 in June 2012, 0 if aged 31–35
AFTER	Binary: 1 for years 2013–2016, 0 for years 2008–2011
PERWT	ACS person weight for population estimates
AGE	Age in years at time of survey
SEX	1 = Male, 2 = Female (IPUMS coding)
MARST	Marital status: 1 = Married spouse present, 2 = Married spouse absent, 3 = Separated, 4 = Divorced, 5 = Widowed, 6 = Never married
EDUC_RECODE	Simplified education: Less than High School, High School Degree, Some College, Two-Year Degree, BA+
STATEFIP	State FIPS code
YEAR	Survey year