

# Heterogeneous Innovations and Growth under Imperfect Technology Spillovers\*

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## Abstract

We study how frictions in learning others' technology, termed "imperfect technology spillovers," impact firm innovation strategies and the aggregate economy through changes in innovation composition. We develop an endogenous growth model that generates strategic innovation decisions, where multi-product firms improve their products via own-innovation and enter new product markets through creative destruction under learning frictions. In our model, firms with technological advantages intensify own-innovation as learning frictions enable them to protect their markets from competitors, thereby reducing creative destruction of rivals. This pattern gets more pronounced when competitive pressure increases exogenously. Using U.S. administrative firm-level data, we provide regression results supporting the model predictions.

**Keywords:** innovation, technology spillover, endogenous growth, competition

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# 1 Introduction

Firm innovation involves two key steps: first, learning about existing technologies (e.g., [Lucas and Moll, 2014](#)), and then making improvements upon them. Depending on how this process unfolds, innovation can take diverse forms. For example, some innovations build directly on a firm’s existing technologies, while others require learning entirely new ones beyond the firm’s current scope. Recent literature highlights the importance of understanding innovation heterogeneity. It shows that heterogeneous innovations impact firm performance and the aggregate economy in different ways, and firms pursue them for different purposes ([Akcigit and Kerr, 2018](#); [Garcia-Macia et al., 2019](#); [Peters, 2020](#); [Argente et al., 2024](#)).

Despite the richness of the literature on innovation heterogeneity, it often abstracts from the above distinct steps. In particular, limited research explores how learning shapes firm innovation incentives and decisions, and makes these decisions interact and impact the aggregate economy. When learning takes time, it can create strategic incentives for firms to pursue different innovations, influenced by factors such as expected returns, existing knowledge stocks, competitive pressure, and the capacity to overcome learning barriers. This, in turn, can alter the composition of innovations and lead to different implications for the aggregate economy.

In this paper, we investigate two key questions: How do firms use different types of innovation when learning others’ technology takes time? How does this process offer new insights into the aggregate implications of firm innovation, particularly in the context of increasing competition? Theoretically, we develop an endogenous growth model with two types of innovation and learning frictions. This model offers a micro-foundation for understanding firms’ varying innovation incentives and their interactions, as well as how both the level and composition of the two innovation types change when competition intensifies. Empirically, we link U.S. administrative firm-level data to the patent database and document new facts about innovation heterogeneity and compositional changes in response to an exogenous increase in competition. Finally, we calibrate the model and derive the aggregate implications of competition across different economies.

In the model, multi-product firms grow through two types of innovation—“own-innovation” and “creative destruction”—subject to imperfect technology spillovers. Own-innovation improves existing product quality, while creative destruction enables firms to enter new markets by displacing

incumbents.<sup>1</sup> In addition, the two innovations differ along two key dimensions: innovation step size and learning. Individual creative destruction contributes more to product quality improvement, thus to firm and aggregate growth than own-innovation.<sup>2,3</sup> Creative destruction also requires learning others’ technology, which takes time due to imperfect technology spillovers.

The literature widely acknowledges the time-consuming nature of learning others’ technology (e.g., [Lucas and Moll, 2014](#)). Accordingly, we conceptualize innovation in two stages: learning existing technology and building on it. Own-innovation bypasses the learning phase, as the firm already knows the technology, while creative destruction requires it. Thus, entering a market through creative destruction involves learning and improving incumbents’ technology, which demands substantial time and resources.<sup>4</sup> Our model incorporates lagged learning as a form of imperfect technology spillovers, requiring potential rivals to spend one period learning the frontier technology of incumbents. In other words, creative destruction builds on one-period lagged technology.

Imperfect technology spillovers, a novel feature introduced in this class of models, generate unique mechanisms that reshape firm innovation decisions. Spillover frictions create a technology gap between incumbents’ frontier technology and the one-period lagged technology available to rivals. This gap enables incumbents to strategically use own-innovation to defend their markets by improving product quality further and securing a technological advantage—a “*market-protection*” effect.<sup>5,6</sup> Then, the technological advantages of incumbents act as “*technological barriers*” that deter rivals’ entry and stifle creative destruction—a “*technological barrier*” effect. This interaction

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<sup>1</sup>An illustrative real-world example of creative destruction is Apple’s entry into the cell phone industry with the introduction of the iPhone back in 2007 when its major business was computer manufacturing. An example of own-innovation is Apple’s improvement and production of iPhone 16 from iPhone 15.

<sup>2</sup>As highlighted by [Bernard et al. \(2010\)](#) and [Akcigit and Kerr \(2018\)](#), creative destruction plays a pivotal role in driving growth. It is tightly connected to radical innovation that significantly improves existing technologies. We provide empirical evidence supporting this in Section 3.2.

<sup>3</sup>[Garcia-Macia et al. \(2019\)](#) show that the aggregate-level own-innovation contributes more to the aggregate growth than creative destruction, because the latter succeeds less frequently. Our model aligns with these findings at the aggregate level while also distinguishing the contributions of the two types of innovation at the individual level.

<sup>4</sup>Creative destruction involves recruiting new employees to handle new technology, reallocating resources to new projects, training workers, reverse engineering, and preparing production facilities for new products. For example, Apple took three years to enter the cell phone industry, even after leveraging their previously accumulated knowledge from iPod development and production. We show empirical evidence on the difference in learning time in Section 3.2.

<sup>5</sup>In particular, firms with accumulated technological advantage increase own-innovation to defend themselves from competitors, while those without such an advantage reduce it when competition increases exogenously.

<sup>6</sup>Our market-protection effect differs from the escape-competition effect as it focuses on the innovation incentives of market leaders and competitive pressure from the entry margin. This perspective explains why frontier firms like Google and NVIDIA keep innovating intensely, despite their technological lead over competitors. The standard step-by-step innovation model implies these firms should stop innovating due to the assumption of immediate catch-up by rivals.

distinguishes our model from others in the literature on firm innovation and specialization. Competition induces a shift in the composition of firm innovation, driven by the strategic choices of firms and their endogenous interactions. Consequently, the aggregate effect of competition on innovation depends on the relative shifts between the two types of innovation.

The strength of our model lies in capturing the strategic role of own-innovation and its endogenous feedback effects on rivals' creative destruction and market entry. By introducing imperfect technology spillovers into a multi-product firm framework, the model achieves this major theoretical advance.<sup>7</sup> In existing models of multi-product firms growing through product scope expansion, firms cannot protect their markets because rivals can instantly learn and copy frontier technology without any friction (Klette and Kortum, 2004; Akcigit and Kerr, 2018; Peters, 2020).<sup>8</sup> Step-by-step innovation models generate an escape-competition motive, but assume single-product firms that can only do own-innovation (Aghion et al., 2001, 2005; Akcigit et al., 2018). This lacks the feedback effects of incumbents' innovation choices on the innovation of rivals attempting to enter a product market and does not capture the firm-level innovation composition observed in the data.

To this end, our model underscores the important role of innovation composition in understanding the aggregate implications of competition. Unlike earlier frameworks, own-innovation not only improves incumbents' product quality but also suppresses rivals' creative destruction and entry, which contributes to firm and economic growth substantially. Ignoring such shifts in innovation composition obscures the true impact of competition on firms and the aggregate economy.

Next, to validate our model, we construct a unique dataset combining U.S. administrative firm-level data with USPTO patent data from 1976 to 2016. This provides comprehensive information for the entire population of U.S. patenting firms. We use the rise in foreign firm entry into U.S. markets after China's WTO accession in 2001 as an exogenous competition shock, and use the self-citation ratio of patents as a measure of the likelihood that patents are used for own-innovation.

Using the pre-shock period (prior to 2000), we document differences between own-innovation and creative destruction in learning time, quality improvement, and economic outcomes. We find that patents closer to creative destruction have a longer duration between their application year and that of backward-cited patents, higher scientific and market values, and greater contributions

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<sup>7</sup>In this sense, our framework brings together quality-ladder and step-by-step innovation models.

<sup>8</sup>Akcigit and Kerr (2018), for example, is a special case of our model where firms have no technological advantages. In their setup, firms reduce own-innovation when competition increases.

to firm growth. Furthermore, we find regression results consistent with the model predictions on the “*market-protection*” and “*technological barrier*” effects: heightened competition increases own-innovation among firms with technological advantages but decreases creative destruction across all firms; firm entry is lower in industries with higher technological barriers, measured by the TFPR gap as in [Aghion et al. \(2005\)](#).

Lastly, to understand the aggregate implications, we calibrate our model to the U.S. manufacturing sector and conduct two counterfactual exercises by increasing competitive pressure by outside firms exogenously: i) in the U.S. economy, and ii) in economies where creative destruction costs exceed those in the U.S.<sup>9,10</sup> Both exercises yield qualitatively similar results at the firm level: firms increase (decrease) own-innovation for products with a (no) technological advantage, while creative destruction drops across all firms. However, the aggregate implications differ.

Overall innovation—the aggregate-level R&D to sales ratio—declines in the U.S., where firms actively engage in creative destruction with lower associated costs. In contrast, economies with higher creative destruction costs see the opposite result, as the initially low levels of creative destruction leave little room for further decline under increased competition. Despite this, the aggregate growth by domestic firms falls in both economies, even though the latter has seen an increase in overall innovation. This occurs because heightened competition endogenously elevates technological barriers and impedes creative destruction by domestic incumbents and firm entry.

Our paper offers a unified framework that facilitates the comparison of the effects of competition across different countries. Changes in innovation composition, driven by firms’ strategic choices, are an important margin to understand the heterogeneous impact of competition and its aggregate implications across diverse economic landscapes.

**Related Literature** First, our paper is related to an extensive body of research on heterogeneity in innovations. [Aghion et al. \(2004\)](#) and [Akcigit et al. \(2018\)](#) incorporate entry margins and [Atkeson and Burstein \(2010\)](#) explore product and process innovations. However, these models focus on single-product firms that can pursue one innovation type. [Klette and Kortum \(2004\)](#) model multi-product firms but assume a single innovation type. Other studies have expanded the study of multi-product firms, including [Bernard et al. \(2010\)](#) on product switching, [Akcigit and Kerr \(2018\)](#)

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<sup>9</sup>Such costs include heightened frictions related to R&D or labor mobility for creative destruction.

<sup>10</sup>Additional counterfactual analysis of an increasing domestic firm entry is presented in Online Appendix [G](#).

on the two types of innovation by multi-product firms with their distinctive role in driving economic growth, [Peters \(2020\)](#) on the role of creative destruction in mitigating market power accumulated by own-innovation, and [Dhingra \(2013\)](#) and [Argente et al. \(2024\)](#) on the role of cannibalization in firm innovation decisions. In addition, [Garcia-Macia et al. \(2019\)](#) and [Atkeson and Burstein \(2019\)](#) explore the impacts of varied innovation types on growth and policy implications. Our contribution arises from adding learning frictions that create firms' strategic use of own-innovation while creating feedback effects on creative destruction and firm entry into a product market. Our research also provides rich empirical evidence on the properties of heterogeneous innovations and the strategic innovation decisions of multi-product firms that drive shifts in innovation composition.

Second, our paper relates to the growing literature on technology gaps and spillovers. Previous studies have shown that technology gaps between firms shape firm innovation incentives and policy implications ([Aghion et al., 2001, 2005](#); [Dinopoulos and Syropoulos, 2007](#); [Aghion and Griffith, 2008](#); [Acemoglu and Akcigit, 2012](#); [Akcigit et al., 2018](#)); documented the trend of diminishing knowledge diffusion from market leaders to laggards ([Andrews et al., 2016](#); [Bessen et al., 2020](#); [Akcigit and Ates, 2021](#); [Arora et al., 2021](#); [Akcigit and Ates, 2023](#)); and highlighted phenomena consistent with this trend ([Shapiro, 2000](#); [De Ridder, 2024](#); [Olmstead-Rumsey, 2019](#); [Argente et al., 2020](#); [Bessen et al., 2020](#); [Bloom et al., 2020](#); [Aghion et al., 2023](#); [Akcigit and Ates, 2023](#); [Akcigit and Goldschlag, 2023](#)). However, these studies have not identified a definitive mechanism driving this shift. Our contribution is to uncover an underlying endogenous force behind the decline in technology diffusion: firms' strategic responses to increased competition under learning frictions.

Lastly, our paper contributes to the longstanding literature on competition and innovation. The empirical literature reports mixed findings ([Aghion et al., 2005](#); [Bloom et al., 2016](#); [Hombert and Matray, 2018](#); [Shu and Steinwender, 2019](#); [Autor et al., 2020](#)).<sup>11</sup> Some explore this dynamic through Schumpeterian growth models featuring step-by-step innovation ([Aghion et al., 2001, 2004, 2005, 2009](#); [Akcigit et al., 2018](#)), while others employ trapped-factor models where rising competition reduces innovation's opportunity cost ([Bloom et al., 2013, 2021](#); [Medina, 2022](#)).<sup>12</sup> However, their analysis is rooted in several assumptions lacking data support and abstracts from discussing the

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<sup>11</sup>See [Shu and Steinwender \(2019\)](#) for further details.

<sup>12</sup>The former has the "Schumpeterian effect" (laggards) and the "escape-competition effect" (neck-and-neck firms).

composition of different innovations.<sup>13,14</sup> We contribute by providing a rich theoretical framework where multi-product firms strategically leverage technological advantages, employ own-innovation, and endogenously affect rivals’ creative destruction, along with supportive data evidence. Our framework demonstrates how competition changes innovation composition, generating distinct aggregate effects based on relative shifts. This helps reconcile the prior diverging findings and deepens understanding of complex impact of competition on innovation.

The rest of the paper is organized as follows. Section 2 develops the baseline endogenous growth model. Section 3 presents the empirical strategy and findings. Section 4 provides a quantitative analysis of the model. Section 5 concludes.

## 2 Baseline Model

We build a discrete-time infinite horizon endogenous growth model with multi-product firms, two types of innovation, imperfect technological spillovers, and an exogenous source of competitive pressure.<sup>15</sup> The model is distinct in the following three dimensions: we i) introduce a novel friction named “imperfect technology spillovers” by assuming that firms can only learn the incumbent’s technology lagged by one period in the process of creative destruction; ii) generate incumbent firms’ own-innovation decision as an endogenous function of the technology gap—the ratio of the current-period technology  $q_{j,t}$  to the previous-period technology  $q_{j,t-1}$ ,  $\Delta_{j,t} = \frac{q_{j,t}}{q_{j,t-1}}$ , due to the friction; and iii) allow for exogenous shifts of the aggregate creative destruction arrival rate to analyze the effect of increasing competitive pressure on firm innovation and growth. Hereafter, the time subscript is suppressed.<sup>16</sup> The terms product quality and technology are used interchangeably.

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<sup>13</sup>The model assumes single-product firms, a single innovation type, or the immediate imitation by the laggards.

<sup>14</sup>Alternative studies include [Hombert and Matray \(2018\)](#) (product differentiation), [Dhingra \(2013\)](#) (process innovation to avoid cannibalization), and [Helpman \(2023\)](#) (ambiguous impacts on innovation by multi-product firms).

<sup>15</sup>The exogenous competitive pressure emanates from firms outside the economy, which could be foreign firms or domestic incumbent firms in other sectors or states, depending on whether we consider the model economy as an aggregate economy or a specific sector or state.

<sup>16</sup>Superscript  $t$  denotes the forward-period ( $t + 1$ ), and subscript  $-1$  is used for the previous-period ( $t - 1$ ).

## 2.1 Representative Household

A representative household consists of a measure-one continuum of individuals, each supplying one unit of labor inelastically to the final goods sector and consuming a portion  $C_t$  of the economy's final goods each period. The household has the following lifetime utility:

$$U = \sum_{t=0}^{\infty} \beta^t \log(C_t).$$

## 2.2 Final Goods Producer

The final goods producer produces a final good with labor ( $L$ ) and a continuum of differentiated intermediate goods indexed by  $j \in [0, 1]$  (produced by either domestic firms  $j \in \mathcal{D}$  or outsiders  $j \notin \mathcal{D}$ ). The production function has the constant returns-to-scale technology as follows:

$$Y = \frac{L^\theta}{1-\theta} \left[ \int_0^1 q_j^\theta y_j^{1-\theta} \mathcal{I}_{\{j \in \mathcal{D}\}} dj + \int_0^1 q_j^\theta y_j^{1-\theta} \mathcal{I}_{\{j \notin \mathcal{D}\}} dj \right],$$

where  $y_j$  and  $q_j$  are the quantity and quality of good  $j$ , and  $\mathcal{I}_{\{\cdot\}}$  are indicator functions. The market is competitive with the price normalized to one, and input prices taken as given.

## 2.3 Intermediate Producers

Domestic and outside firms have the mass of  $\mathcal{F}_d$  and  $\mathcal{F}_o$ , respectively, with  $\mathcal{F} = \mathcal{F}_d + \mathcal{F}_o \in (0, 1)$ . They produce and sell differentiated intermediate goods in monopolistically competitive domestic markets. Each firm operates with at least one product, and each product is owned by a single firm. Thus, firm  $f$  can be characterized by the collection of its products  $\mathcal{J}^f = \{j : j \text{ is owned by firm } f\}$ . The intermediate good is produced at a unit marginal cost in terms of final goods.

## 2.4 Innovation by Intermediate Producers

Intermediate producers engage in two types of R&D, own-innovation and creative destruction, by spending expenditures in units of final goods. Firms improve the quality of their own products



through own-innovation, while taking over other markets through creative destruction.<sup>17</sup> The R&D output manifests as improving product quality and is realized at the beginning of the next period.

On top of this, we introduce a novel friction named “imperfect technology spillovers,” under which learning others’ technology takes time in the process of creative destruction.<sup>18</sup> We conceptualize it in the form of lagged learning by assuming creative destruction builds on the past-period technology. Thus, only the owner of a product line can observe the frontier level of technology  $q_{j,t}$  in the market, while outsider firms can only see the lagged level  $q_{j,t-1}$ .<sup>19</sup> Following this, a product line can be sufficiently characterized by its quality  $q_j$  and technology gap between current and previous periods  $\Delta_{j,t} = \frac{q_{j,t}}{q_{j,t-1}}$ .<sup>20</sup> This friction induces incumbents to strategically use own-innovation to build technological barriers and protect their markets from competitors. We name it the “market-protection effect.”

When two firms are neck and neck in a product line, a coin-toss tiebreaker rule applies as in [Acemoglu et al. \(2016\)](#) to make sure each product is produced by only one firm.<sup>21</sup> Creative destruction is undirected and the targeted product is randomly assigned among the entire set of products with equal probability. For now, we assume that firms can only attempt one creative destruction each period, which aids in deriving analytic solutions for firm decision rules and distributions with minimal assumptions. In the quantitative analysis, we allow multiple creative destructions as in [Klette and Kortum \(2004\)](#).

**Own-Innovation** Successful own-innovation improves the current quality  $q_{j,t}$  of the product by  $\lambda > 1$ . The quality of good  $j$  evolves as follows, conditional on the owning firm not displaced by

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<sup>17</sup>Note that the quality in this model is a marginal cost of production-adjusted measure, and can be improved through either technological advancement or cost reduction. In this sense, our concept of innovation encompasses both product and process innovations. Thus, firms with lower technology can compete with firms with advanced technology if they have a sufficiently lower marginal cost of production.

<sup>18</sup>Our empirical analysis in Section 3.2 suggests that creative destruction takes longer to develop than own-innovation. The learning time is measured by the gap between the application years of a patent and the patents it backward cites.

<sup>19</sup>All the aggregate variables and technology gap distribution are publicly observable, and firms make optimal innovation decisions by considering them. In this way, a stationary firm-product distribution is well defined.

<sup>20</sup>This technology gap summarizes the technological advantage incumbents have in their market.

<sup>21</sup>Unused technology is assumed to depreciate sufficiently to make it unprofitable for firms that have done creative destruction to build upon it next period. This approach guarantees the undirected nature of creative destruction and restricts own-innovation to the current owner only.

creative destruction:

$$\{q_{j,t+1}^{\text{in}}\} = \begin{cases} \{\lambda q_{j,t}\} & \text{with probability } z_{j,t} \\ \{q_{j,t}\} & \text{with probability } 1 - z_{j,t}, \end{cases} \quad z_{j,t} = \left( \frac{R_{j,t}^{\text{in}}}{\widehat{\chi} q_{j,t}} \right)^{\frac{1}{\widehat{\psi}}}, \quad \widehat{\chi} > 0, \widehat{\psi} > 1.$$

The success probability of own-innovation,  $z_{j,t}$ , depends on R&D investment  $R_{j,t}^{\text{in}}$ .<sup>22</sup>

**Creative Destruction** Successful creative destruction improves the lagged quality of the obtained product by  $\eta > 1$ . We assume  $\lambda^2 > \eta > \lambda$ .  $\eta > \lambda$  reflects our empirical findings (in Section 3.2) that creative destruction is of higher quality than own-innovation, both in terms of market value (Kogan et al., 2017) and forward citations, and contributes more to firm and aggregate growth.<sup>23,24</sup>  $\lambda^2 > \eta$  is based on the idea that consecutive own-innovation has a significant influence, which ensures that firms can protect their own product lines from rivals through own-innovation.<sup>25</sup> These assumptions are innocuous to our results, introduced primarily to simplify the main exposition and reduce computational burden. In Online Appendix C, we develop a framework where innovation step sizes ( $\lambda$  and  $\eta$ ) are drawn from probability distributions, as in Garcia-Macia et al. (2019), and show the robustness of the main results without specific assumptions about their relative magnitudes. These findings remain valid even when both step sizes are drawn from the same distribution.

Firms invest in creative destruction and can obtain the following product quality if not pre-empted by the successful own-innovation of the incumbent in their target market:

$$\{q_{j,t+1}^{\text{ex}}\} = \begin{cases} \{\eta q_{j,t-1}\} & \text{with probability } x_t \\ \emptyset & \text{with probability } 1 - x_t, \end{cases} \quad x_t = \left( \frac{R_t^{\text{ex}}}{\widetilde{\chi} \bar{q}_t} \right)^{\frac{1}{\widetilde{\psi}}}, \quad \widetilde{\chi} > 0, \widetilde{\psi} > 1.$$

The success probability of creative destruction,  $x_t$ , is determined by R&D investment  $R_t^{\text{ex}}$  and the average quality  $\bar{q}_t$  in the economy.<sup>26</sup> With probability  $1 - x_t$ , the creative destruction fails, implying

<sup>22</sup>Hereafter, we represent the quality of product  $j$  as a point set. This makes it easy to describe the case where a firm fails to acquire any product lines—in such cases, the product quality set is an empty set.

<sup>23</sup>Our empirical analysis in Section 3.2 shows that creative destruction is associated with higher firm employment growth compared to own-innovation, consistent with Akcigit and Kerr (2018). Additionally, we find that creative destruction is positively associated with TFPR growth and the number of products added at the firm level.

<sup>24</sup>The latter implications for firm and aggregate growth are also shown in Akcigit and Kerr (2018).

<sup>25</sup>We discuss this in more detail in our quantitative analysis in Section 4.

<sup>26</sup>The average quality matters for creative destruction as the target product is randomly assigned.

neither product takeover nor quality is obtained. See Online Appendix A.1 for examples illustrating how firms choose own-innovation and creative destruction.

**Product Quality Evolution** Imperfect technology spillovers create a gap between current and lagged product quality levels,  $\Delta_{j,t} = \frac{q_{j,t}}{q_{j,t-1}}$ , which represents the technological advantage incumbent firms hold in their markets. This advantage allows them to defend their product lines through own-innovation. The technology gap can take one of four possible values.

**Lemma 1.** *There are four possible values for a technology gap:  $\Delta^1 = 1$ ,  $\Delta^2 = \lambda$ ,  $\Delta^3 = \eta$ , and  $\Delta^4 = \frac{\eta}{\lambda}$ , with  $\Delta^3$  and  $\Delta^4$  arising exclusively through creative destruction. Proof: See the Appendix.*

Then, incumbents' product quality (conditional on a technology gap) evolves as follows:

$$\left\{ q_{j,t+1} \mid \Delta_{j,t} = \Delta^1 \right\} = \begin{cases} \emptyset & , \text{ with prob. } \bar{x} \\ \{q_{j,t}\} & , \text{ with prob. } (1 - \bar{x})(1 - z_j^1) \\ \{\lambda q_{j,t}\} & , \text{ with prob. } (1 - \bar{x})z_j^1 \end{cases} \quad (1)$$

$$\left\{ q_{j,t+1} \mid \Delta_{j,t} = \Delta^2 \right\} = \begin{cases} \emptyset & , \text{ with prob. } \bar{x}(1 - z_j^2) \\ \{q_{j,t}\} & , \text{ with prob. } (1 - \bar{x})(1 - z_j^2) \\ \{\lambda q_{j,t}\} & , \text{ with prob. } z_j^2 \end{cases} \quad (2)$$

$$\left\{ q_{j,t+1} \mid \Delta_{j,t} = \Delta^3 \right\} = \begin{cases} \emptyset & , \text{ with prob. } \frac{1}{2}\bar{x}(1 - z_j^3) \\ \{q_{j,t}\} & , \text{ with prob. } (1 - \frac{1}{2}\bar{x})(1 - z_j^3) \\ \{\lambda q_{j,t}\} & , \text{ with prob. } z_j^3 \end{cases} \quad (3)$$

$$\left\{ q_{j,t+1} \mid \Delta_{j,t} = \Delta^4 \right\} = \begin{cases} \emptyset & , \text{ with prob. } \bar{x}(1 - \frac{1}{2}z_j^4) \\ \{q_{j,t}\} & , \text{ with prob. } (1 - \bar{x})(1 - z_j^4) \\ \{\lambda q_{j,t}\} & , \text{ with prob. } (1 - \frac{1}{2}\bar{x})z_j^4. \end{cases} \quad (4)$$

Note that  $z_j^\ell$  is the optimal own-innovation of the firm owning product  $j$  with its technology gap  $\Delta^\ell$ ,  $\ell \in \{1, 2, 3, 4\}$ .  $\bar{x}$  is the aggregate creative destruction arrival rate, representing the probability that a product market faces a rival firm with successful creative destruction. The symbol  $\emptyset$  indicates that the firm loses product line  $j$  in the next period, and the term  $\frac{1}{2}$  in the probabilities reflects a coin-toss tiebreaker in neck-and-neck scenarios.

In the case of  $\Delta_{j,t} = 1$ , incumbents lack any technological advantage and lose their product lines if a rival firm arrives with successful creative destruction, irrespective of their success in own-innovation.<sup>27</sup> In contrast, for other cases where  $\Delta^\ell > 1$ , firms can lower the probability of losing their product lines by investing more in own-innovation.<sup>28</sup> Hence, firms are more incentivized to augment their own-innovation efforts for products with technological advantages ( $\Delta^\ell > 1$ ).<sup>29</sup>

For rival firms entering a market, the success probability of product takeover depends not only on the success of creative destruction but also on the technology gap and the own-innovation intensity of the product owner, even after successful creative destruction. Thus, the success probability of product takeover  $x_{\text{takeover}} (\equiv x\bar{x}_{\text{takeover}})$  can be decomposed into i) the success probability of creative destruction  $x$ , and ii) conditional takeover probability  $\bar{x}_{\text{takeover}}$ , which is defined as

$$\bar{x}_{\text{takeover}} = \mu(\Delta^1) + (1 - z^2)\mu(\Delta^2) + \frac{1}{2}(1 - z^3)\mu(\Delta^3) + \left(1 - \frac{1}{2}z^4\right)\mu(\Delta^4), \quad (5)$$

with technology gap distribution  $\{\mu(\Delta^\ell)\}_{\ell=1}^4$  (the mass of products with a gap  $\Delta^\ell$ ).<sup>30</sup>

Note that the higher the overall innovation intensity (both own-innovation and creative destruction), the wider the average technology gap becomes in the economy (the mass of products with  $\Delta^1$  decreases). This makes it difficult for firms to successfully take over product markets.<sup>31</sup> This is referred to as the “technological barrier effect,” where increased own-innovation by incumbents or higher  $\bar{x}$  dampens the creative destruction and growth of firms.<sup>32</sup>

**Potential Startups** There is a fixed mass of potential domestic startups  $\mathcal{E}_d$ . To start a business, they invest in creative destruction R&D and attempt to take over a product line from an incumbent firm.

<sup>27</sup>Rivals with successful creative destruction achieve  $q_{j,t+1}^{\text{rival}} = \eta q_{j,t-1}$ , which is greater than  $\lambda q_{j,t-1}$ . The model in [Akcigit and Kerr \(2018\)](#) falls into this category.

<sup>28</sup>The extent of the reduction in the probability of product loss is contingent on the technology gap.

<sup>29</sup>See Online Appendix [A.2](#) for the evolution of product quality for rival firms entering a market.

<sup>30</sup>This shows that if a firm succeeds in creative destruction for a product line with a technology gap of  $\Delta^1$ , then it takes over that product line with a probability of one. For a product line with  $\Delta^2$ , this probability becomes  $1 - z^2$ ; for  $\Delta^3$ , it is  $\frac{1}{2}(1 - z^3)$ ; and for  $\Delta^4$ , it is  $1 - \frac{1}{2}z^4$ . It is assumed that own-innovation intensity  $z$  depends solely on technology gap  $\Delta^\ell$ . In the next section, we prove this assumption holds true.

<sup>31</sup>Higher own-innovation intensity widens the technology gap. Simultaneously, higher creative destruction intensity increases the aggregate creative destruction arrival rate, thereby incentivizing incumbent firms to engage in own-innovation more endogenously, as discussed earlier.

<sup>32</sup>This technological barrier effect is a novel feature of our model, which is distinct from the well-known Schumpeterian effect. The Schumpeterian effect is that firm innovation incentives decline following an increase in  $\bar{x}$  due to lowered expected future profits conditional on successful innovation and business takeover.

Potential startups choose R&D expenditure  $R_e^{\text{ex}}$  and decide the probability of creative destruction  $x_e = (R_e^{\text{ex}} / (\tilde{\chi}_e \bar{q}))^{\frac{1}{\tilde{\psi}_e}}$ , where  $\tilde{\chi}_e > 0$  and  $\tilde{\psi}_e > 1$ . Let  $V(\{(q_j, \Delta_j)\})$  denote the value of a firm that has a product with quality  $q_j$  and a technology gap of  $\Delta_j$ . Then a potential startup's expected profits from entering through R&D is

$$\Pi^e = \tilde{\beta} \mathbb{E} [V(\{(q'_j, \Delta'_j)\}) | x_e] - \tilde{\chi}_e (x_e)^{\tilde{\psi}_e} \bar{q}, \quad (6)$$

where  $\tilde{\beta}$  is the stochastic discount factor, and the expectation conditioning on  $x_e$  is taken over the distribution of incumbents' product quality  $q_j$  and technology gap  $\Delta_j$  due to the undirected nature of creative destruction.<sup>33</sup> Potential startups choose the probability of creative destruction  $x_e$  to maximize expected profits from entry. Since potential startups are ex-ante homogeneous, they all choose the same level of creative destruction intensity  $x_e^*$ . Hence, the mass of potential domestic startups that succeed in creative destruction and attempt to take over product markets is  $\mathcal{E}_d x_e^*$ .

## 2.5 Exogenous Competitive Pressure and Creative Destruction

As explained before, the aggregate creative destruction arrival rate  $\bar{x}$  is the probability that an incumbent faces a rival firm (either a domestic startup, incumbent, or an outside firm) with successful creative destruction. The aggregate creative destruction arrival rate is equal to the total mass of firms succeeding in creative destruction given the undirected nature of creative destruction and the continuum of unit mass of product lines.<sup>34</sup> Let  $\bar{x}_d$  denote the total mass of domestic firms with successful creative destruction and  $\bar{x}_o$  denote the outside firms' counterpart. The creative destruction arrival rate is defined as

$$\bar{x} = \bar{x}_d + \bar{x}_o.$$

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<sup>33</sup> $\tilde{\beta} = \frac{\beta C}{C'}$  as the household owns all firms.

<sup>34</sup>This follows along with the assumption that each firm can do creative destruction at most one product line each period, which makes the total mass of firms with successful creative destruction equivalent to the total mass of product markets for which an incumbent faces a rival firm. This assumption is extended in our full-fledged version, and this result still holds with additional aggregation across products within successful firms.

Competitive pressure from outside firms is captured by an exogenous increase in  $\bar{x}_o$ , resulting from either increased creative destruction intensity or a larger mass of outside firms.<sup>35</sup>

## 2.6 Equilibrium

**Optimal Production and Employment** The final goods producers choose labor and intermediate goods inputs. Let  $p_j$  denote the price of differentiated product  $j$ , and  $w$  denote the wage in the domestic economy. The inverse demand for intermediate good  $j$  is:

$$p_j = q_j^\theta L^\theta y_j^{-\theta}. \quad (7)$$

Each product is assumed to be supplied by a single firm. We follow [Acemoglu et al. \(2012\)](#) and [Acemoglu et al. \(2018\)](#) and assume that the current and former incumbents engage in the following two-stage price-bidding game for each product line  $j$ : i) each firm pays a fee of  $\varepsilon$  ( $> 0$ ), and ii) those that have paid the fee announce their prices.<sup>36</sup>

Intermediate producers take (7) as given and maximize their operating profits  $\pi(q_j)$  for each product  $j \in \mathcal{J}^f$ .<sup>37</sup> The optimal production and price are derived as follows:

$$y_j = (1 - \theta)^{\frac{1}{\theta}} q_j \quad \text{and} \quad p_j = \frac{1}{1 - \theta}, \quad (8)$$

which simplify the equilibrium profit, wage, and final goods output to the following:

$$\pi(q_j) = \underbrace{\theta(1 - \theta)^{\frac{1-\theta}{\theta}}}_{\equiv \pi} q_j, \quad w = \theta(1 - \theta)^{\frac{1-2\theta}{\theta}} \bar{q}, \quad \text{and} \quad Y = (1 - \theta)^{\frac{1-2\theta}{\theta}} \bar{q}. \quad (9)$$

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<sup>35</sup>[Jo \(2024\)](#) extends our baseline model to a two-country framework and endogenizes the changes in  $\bar{x}_o$  driven by the changes in bilateral tariff rates.

<sup>36</sup>This is to avoid the case where the former market leader, having lost its leadership to the current leader in a market, attempts to produce and sell its product through limit pricing. This ensures that only the firm with the leading-edge technology enters the first stage and announces its price in equilibrium.

<sup>37</sup>Since each intermediate product incurs a unit marginal cost in terms of final goods, the problem is identical for both domestic and outside firms.

**Optimal own-innovation and creative destruction** Let  $\Phi^f \equiv \{(q_j, \Delta_j)\}_{j \in \mathcal{J}^f}$  denote the set of product quality and technology gap for intermediate goods producer  $f$ . The firm value is:

$$V(\Phi^f) = \max_{\substack{x, \\ \{z_j\}_{j \in \mathcal{J}^f}}} \left\{ \sum_{j \in \mathcal{J}^f} \left[ \pi q_j - \hat{\chi} z_j^{\hat{\psi}} q_j \right] - \bar{q} \tilde{\chi} x^{\tilde{\psi}} + \tilde{\beta} \mathbb{E} \left[ V(\Phi^{f'} | \Phi^f) \middle| x, \{z_j\}_{j \in \mathcal{J}^f} \right] \right\}.$$

The first three terms define the current profits (revenue net of production and R&D costs), and the last term is the discounted future expected value. This expectation is computed over various factors, including the success probabilities of own-innovation and creative destruction, the creative destruction arrival rate, the outcomes of winning or losing coin tosses, the current-period product quality distribution, and the current-period technology gap distribution.

**Proposition 1.** *The firm value function and optimal innovation choices are derived as:*

$$V(\Phi^f) = \sum_{\ell=1}^4 A_{\ell} \left( \sum_{j \in \mathcal{J}^f | \Delta_j = \Delta^{\ell}} q_j \right) + B \bar{q} \quad (10)$$

$$z^1 = \left[ \tilde{\beta} \left( (1 - \bar{x}) \lambda A_2 - (1 - \bar{x}) A_1 \right) / \left( \hat{\psi} \hat{\chi} \right) \right]^{\frac{1}{\hat{\psi}-1}} \quad (11)$$

$$z^2 = \left[ \tilde{\beta} \left( \lambda A_2 - (1 - \bar{x}) A_1 \right) / \left( \hat{\psi} \hat{\chi} \right) \right]^{\frac{1}{\hat{\psi}-1}} \quad (12)$$

$$z^3 = \left[ \tilde{\beta} \left( \lambda A_2 - (1 - \bar{x}/2) A_1 \right) / \left( \hat{\psi} \hat{\chi} \right) \right]^{\frac{1}{\hat{\psi}-1}} \quad (13)$$

$$z^4 = \left[ \tilde{\beta} \left( \lambda (1 - \bar{x}/2) A_2 - (1 - \bar{x}) A_1 \right) / \left( \hat{\psi} \hat{\chi} \right) \right]^{\frac{1}{\hat{\psi}-1}} \quad (14)$$

$$x = \left[ \tilde{\beta} A_{\text{takeover}} / \left( \tilde{\psi} \tilde{\chi} \right) \right]^{\frac{1}{\tilde{\psi}-1}}, \quad (15)$$

where

$$A_1 = \pi - \hat{\chi} (z^1)^{\hat{\psi}} + \tilde{\beta} [A_1 (1 - \bar{x}) (1 - z^1) + \lambda A_2 (1 - \bar{x}) z^1] \quad (16)$$

$$A_2 = \pi - \hat{\chi} (z^2)^{\hat{\psi}} + \tilde{\beta} [A_1 (1 - \bar{x}) (1 - z^2) + \lambda A_2 z^2] \quad (17)$$

$$A_3 = \pi - \hat{\chi} (z^3)^{\hat{\psi}} + \tilde{\beta} [A_1 (1 - \bar{x}/2) (1 - z^3) + \lambda A_2 z^3] \quad (18)$$

$$A_4 = \pi - \hat{\chi} (z^4)^{\hat{\psi}} + \tilde{\beta} [A_1 (1 - \bar{x}) (1 - z^4) + \lambda A_2 (1 - \bar{x}/2) z^4] \quad (19)$$

$$B = \left( x \tilde{\beta} A_{\text{takeover}} - \tilde{\chi} x^{\tilde{\psi}} \right) / \left( 1 - \tilde{\beta} (1 + g) \right), \quad (20)$$

$g$  is the growth rate of the average product quality, and  $A_{\text{takeover}}$  is the ex-ante value of a successful takeover of a product line as follows:

$$A_{\text{takeover}} \equiv \frac{1 - z^3}{2} A_1 \mu(\Delta^3) + \left(1 - \frac{z^4}{2}\right) A_2 \lambda \mu(\Delta^4) + A_3 \eta \mu(\Delta^1) + (1 - z^2) A_4 \frac{\eta}{\lambda} \mu(\Delta^2). \quad (21)$$

*Proof:* See the [Appendix](#).

Note that  $A_\ell$  is the sum of discounted expected profits from owning a product with a technology gap of  $\Delta^\ell$ , normalized by the current-period product quality. The first two terms in (16) through (19) denote the normalized instantaneous profits, net of the optimal R&D spending. The terms inside the brackets are the normalized future value from own-innovation.  $B$  is the sum of the discounted expected profits from owning an additional product through creative destruction, normalized by the average product quality.<sup>38</sup>

For the optimal own-innovation (11)-(14), the first term in the brackets (after  $\tilde{\beta}$ ) in the numerator represents the future value from successful own-innovation with the quality increased by  $\lambda$ , and the second term is the counterpart from no successful own-innovation. Thus, holding  $\bar{x}$  fixed, the net future value of successful own-innovation depends on the firm's technology gap, pinning down its optimal choice. Consequently, own-innovation becomes an endogenous function of the technology gap, which is a unique feature of this model due to imperfect technology spillovers.

**Corollary 1.** *In an equilibrium where  $\{z^\ell\}_{\ell=1}^4$  are well defined, the probabilities of own-innovation satisfy  $z^2 > z^3 > z^4 > z^1$ . *Proof:* See the [Appendix](#).*

Corollary 1 shows that own-innovation increases with the technology gap, which helps firms protect their markets. However, beyond a certain point, a wider technology gap can discourage further investment in own-innovation. This occurs because firms are less likely to lose their product line even without doing additional own-innovation.

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<sup>38</sup>To understand this variable clearly, we can rewrite (20) as  $B\bar{q} = x\tilde{\beta}A_{\text{takeover}}\bar{q} - \tilde{\chi}x^{\tilde{\psi}}\bar{q} + \tilde{\beta}B(1+g)\bar{q}$ . After investing  $\tilde{\chi}x^{\tilde{\psi}}\bar{q}$  in creative destruction in the current period, the firm receives the discounted expected profit  $\tilde{\beta}A_{\text{takeover}}\bar{q}$  in the next period if the creative destruction succeeds with probability  $x$ . Then, the firm plans to invest in creative destruction next period and receive an expected profit of  $B\bar{q}'$  in the following period, where  $\bar{q}' = (1+g)\bar{q}$ . Thus, (20) illustrates that  $B$  is the annuity value of an infinite stream of constant payoffs  $x\tilde{\beta}A_{\text{takeover}} - \tilde{\chi}x^{\tilde{\psi}}$ , evaluated at a constant discount rate of  $\tilde{\beta}(1+g)$ , which is the growth rate-adjusted stochastic discount factor.



Furthermore, the optimal own-innovation also depends on creative destruction arrival rate  $\bar{x}$  as shown in Corollary 2, which we label the “market-protection effect.”<sup>39,40</sup>

**Corollary 2** (Market-Protection Effect). *With  $\tilde{\psi} \in (1, 2]$ , the market-protection effect is maximized and is positive for product lines with a technology gap of  $\Delta^2$ , whereas it is minimized and is negative for product lines with  $\Delta^1$ . The market-protection effect is positive for the  $\Delta^3$  case, whereas its sign is ambiguous for the  $\Delta^4$  case. Thus,*

$$\left. \frac{\partial z^2}{\partial \bar{x}} \right|_{A_1, A_2} > \left. \frac{\partial z^3}{\partial \bar{x}} \right|_{A_1, A_2} > 0, \quad \left. \frac{\partial z^3}{\partial \bar{x}} \right|_{A_1, A_2} > \left. \frac{\partial z^4}{\partial \bar{x}} \right|_{A_1, A_2} \leq 0, \quad \text{and} \quad 0 > \left. \frac{\partial z^1}{\partial \bar{x}} \right|_{A_1, A_2}.$$

*Proof:* See the [Appendix](#).

In the  $\Delta^1$  case, own-innovation fails to effectively protect the firm’s product, as shown in (1). Consequently,  $z^1$  decreases as the rate of creative destruction  $\bar{x}$  increases. In contrast, the  $\Delta^2$  case has the strongest market-protection effect, exerting the highest impact on reducing the probability of losing a product as in (2). In the  $\Delta^3$  case, increasing  $z^3$  lowers the probability of product loss, though less than in the  $\Delta^2$  case. This results in a positive but diminished market-protection effect. The effect in the  $\Delta^4$  case remains ambiguous, where higher  $z^4$  leads to a smaller decrease in the probability of losing the product line. This suggests firms that have innovated intensively previously (and thus larger technology gaps) are more likely to intensify own-innovation in response to increased competition (higher  $\bar{x}$ ) than those with less recent innovation. This highlights another crucial and unique aspect of our model: firms strategically employ own-innovation to defend against competitors, leveraging imperfect technology spillovers.

As a result, optimal creative destruction depends on own-innovation, the technology gap distribution among incumbents, and the expected value of products ( $\{A_\ell\}_{\ell=1}^4$ ). Equations (15) and (21) show that higher overall own-innovation and creative destruction intensities reduce firms’ incentive for creative destruction in partial equilibrium, with  $\{A_\ell\}_{\ell=1}^4$  held constant. This is because increased overall innovation shifts the technology gap distribution, which raises the average technology gap,

<sup>39</sup>Note that as  $A_1$  and  $A_2$  also depend on  $\bar{x}$ , it is difficult to analytically determine the signs of the partial derivatives of  $\{z^\ell\}_{\ell=1}^4$  with respect to  $\bar{x}$ . However, by holding the values of  $A_1$  and  $A_2$  fixed, we can explicitly ascertain these signs as in Corollary 2.

<sup>40</sup>The term  $A_2$  in (11)-(14) reflects the well-known Schumpeterian effect—the lower the expected future profits from keeping the product line through own-innovation, the lower the incentive to invest in own-innovation.

and hampers firms' market takeover (the “technological barrier effect”). Furthermore, keeping the probabilities of own-innovation and the technology gap distribution constant, a decrease in the expected product values reduces creative destruction (the “Schumpeterian effect”).<sup>41</sup> Our simple three-period model in Online Appendix B formally proves these predictions.

Similarly, the optimal creative destruction by potential startups  $x_e$  is derived as follows:

$$x_e = \left[ \tilde{\beta} \left( A_{\text{takeover}} + \bar{x}_{\text{takeover}} B(1 + g) \right) / \left( \tilde{\psi}_e \tilde{\chi}_e \right) \right]^{\frac{1}{\tilde{\psi}_e - 1}}, \quad (22)$$

with proof provided in the Appendix.

**Aggregate Creative Destruction Arrival Rate** With (15) and (22), the aggregate creative destruction arrival rate in this economy is defined as follows:

$$\bar{x} = \underbrace{\mathcal{F}_d x + \mathcal{E}_d x_e}_{\equiv \bar{x}_d} + \underbrace{\mathcal{F}_o x + \mathcal{E}_o}_{\equiv \bar{x}_o}, \quad (23)$$

where  $\mathcal{E}_o$  is the total mass of outside entrants with successful creative destruction, which is exogenously determined.<sup>42,43</sup>

## 2.7 Balanced Growth Path

**Proposition 2.** *The aggregate growth rate  $g$  in a Balanced Growth Path is:*

$$\begin{aligned} g = & \left[ (1 - \bar{x})(1 - z^1) + \Delta^2(1 - \bar{x})z^1 + \Delta^3\bar{x} \right] \mu(\Delta^1) \\ & + \left[ (1 - \bar{x})(1 - z^2) + \Delta^2 z^2 + \Delta^4\bar{x}(1 - z^2) \right] \mu(\Delta^2) \\ & + \left[ 1 - z^3 + \Delta^2 z^3 \right] \mu(\Delta^3) + \left[ (1 - \bar{x})(1 - z^4) + \Delta^2(z^4 + \bar{x}(1 - z^4)) \right] \mu(\Delta^4) - 1, \end{aligned} \quad (24)$$

<sup>41</sup>The direction of the changes in the probabilities of own-innovation and creative destruction in response to changes in the aggregate creative destruction arrival rate  $\bar{x}$  are ambiguous in general equilibrium. They depend on the relative magnitudes and the directions of the market-protection effect, the technological barrier effect, and the Schumpeterian effect. Nonetheless, results from the numerical exercise in Section 4.2 confirm that the partial equilibrium results, given  $\{A_\ell\}_{\ell=1}^4$  and  $B$ , still hold in general equilibrium within plausible parameter ranges. Furthermore,  $\{A_\ell\}_{\ell=1}^4$  and  $B$  decrease with an exogenous increase in  $\bar{x}$ .

<sup>42</sup>Note that an exogenous increase in  $\mathcal{E}_o$  may not increase  $\bar{x}$  by the same amount in equilibrium, as the mass of domestic incumbent firms  $\mathcal{F}_d$  and the probabilities of creative destruction  $x$  and  $x_e$  depend on  $\bar{x}$ . Thus, the level of  $\bar{x}$  is endogenously determined, even when  $\mathcal{E}_o$  changes exogenously.

<sup>43</sup>The outside firms in domestic markets make the same innovation decisions as the domestic firms.

which can be decomposed into the parts attributed to own-innovation and creative destruction by domestic incumbents and startups ( $g_d$ ), as well as outside firms ( $g_o$ ).

*Proof:* See the [Appendix](#).

## 2.8 Firm Distribution

Let  $\mathcal{N} = (n_f, n_f^1, n_f^2, n_f^3, n_f^4)$  denote the technology gap composition of firm  $f$ , where  $n_f$  is the total number of products and  $n_f^\ell$  is the count of products with a technology gap of  $\Delta^\ell$ . Let  $\tilde{\mu}(\mathcal{N})$  denote its distribution. Summing  $\tilde{\mu}(\mathcal{N})$  over all possible  $\mathcal{N}$  gives the total mass of firms  $\mathcal{F}$ .

**Transition of Technology Gap Portfolio** Consider a firm with technology gap composition given by  $\tilde{\mathcal{N}}(n_f, k) \equiv (n_f, n_f - k, k, 0, 0)$ , where  $k \in [0, n_f] \cap \mathbb{Z}$  and  $n_f > 0$ . Ignoring creative destruction, the probability of technology gap composition changing from  $\mathcal{N} = \tilde{\mathcal{N}}(n_f, k)$  to  $\mathcal{N}' = \tilde{\mathcal{N}}(n_f, \tilde{k})$  is

$$\begin{aligned} \tilde{\mathbb{P}}(n_f, \tilde{k} | n_f, k) &= \sum_{\tilde{k}^1 = \max\{0, \tilde{k} - k\}}^{\min\{n_f - k, \tilde{k}\}} \left( \frac{(n_f - k)!}{\tilde{k}^1! (n_f - k - \tilde{k}^1)!} \right) \left( \frac{k!}{(\tilde{k} - \tilde{k}^1)! (k - (\tilde{k} - \tilde{k}^1))!} \right) \\ &\quad \times (1 - \bar{x})^{n_f - (\tilde{k} - \tilde{k}^1)} (1 - z^1)^{n_f - k - \tilde{k}^1} (z^1)^{\tilde{k}^1} (1 - z^2)^{k - (\tilde{k} - \tilde{k}^1)} (z^2)^{\tilde{k} - \tilde{k}^1}, \end{aligned}$$

for  $n_f \geq 1$  and  $0 \leq \tilde{k}, k \leq n_f$ , and zero, otherwise. This follows a binomial process as in [Ates and Saffie \(2021\)](#).

Using the above, we can track general cases transitioning from  $\mathcal{N} = (n_f, n_f^1, n_f^2, n_f^3, n_f^4)$  to  $\mathcal{N}' = (n'_f, n_f^{1'}, n_f^{2'}, n_f^{3'}, n_f^{4'})$  for any  $n'_f \leq n_f + 1$  as products with  $\Delta^3$  or  $\Delta^4$  can only be obtained through creative destruction. Details can be found in Online Appendix [A.4](#).

**Technology Gap Distribution** The aggregate distribution of technology gaps is

$$\mu(\Delta^\ell) = \sum_{n_f=1}^{\bar{n}_f} \sum_{n_f^\ell=0}^{n_f} \sum_{n_f^{-\ell}=0}^{n_f} n_f^\ell \tilde{\mu}(n_f, n_f^1, n_f^2, n_f^3, n_f^4), \text{ for } \ell = 1, 2, 3, 4 \quad (25)$$

where the third summation represents the sum over all possible values for  $n_f^{-\ell}$  other than the focal  $\ell$ . Note  $\sum_{\ell=1}^4 \mu(\Delta^\ell) = 1$  holds in equilibrium.<sup>44</sup>

<sup>44</sup>This is because each product line is occupied by one incumbent and there is a unit mass of products.

**Aggregate Variables and Balanced Growth Path** Given the optimal innovation choices (11), (12), (13), (14), (15), and (22), the aggregate domestic R&D expenses becomes

$$R_d = \hat{\chi} \sum_{\ell=1}^4 \left[ \int_0^1 q_j \mathcal{I}_{\{\Delta_j=\Delta^\ell, j \in \mathcal{D}\}} dj \right] (z^\ell)^{\hat{\psi}} + \mathcal{F}_d \tilde{\chi} \bar{q} x^{\tilde{\psi}} + \mathcal{E}_d \tilde{\chi}_e (x_e)^{\tilde{\psi}_e} \bar{q}, \quad (26)$$

where  $\mathcal{I}_{\{\Delta_j=\Delta^\ell, j \in \mathcal{D}\}}$  is an indicator for product line  $j$  owned by a domestic firm with  $\Delta^\ell$ . With (8), the aggregate demand for final goods by domestic intermediate producers is

$$Y_d = \int_0^1 y_j \mathcal{I}_{\{j \in \mathcal{D}\}} dj = (1 - \theta)^{\frac{1}{\theta}} \int_0^1 q_j \mathcal{I}_{\{j \in \mathcal{D}\}} dj, \quad (27)$$

and the aggregate consumption is determined by

$$C = Y - \int_{j \notin \mathcal{D}} p_j y_j dj - Y_d - R_d, \quad (28)$$

where the second term is the payments to outside intermediate producers.<sup>45</sup> Lastly, the balanced growth path (BGP) equilibrium is characterized by the following:

**Definition 1.** A balanced growth path equilibrium consists of  $y_j^*, p_j^*, w^*, L^*, x^*, \{z^{\ell*}\}_{\ell=1}^4, \bar{x}^*, x_e^*, \mathcal{F}^*, R_d^*, Y^*, C^*, g^*, \tilde{\mu}(\mathcal{N}), \{\mu(\Delta^\ell)\}_{\ell=1}^4$  for  $j \in [0, 1]$  with  $q_j$  such that: (i)  $y_j^*$  and  $p_j^*$  satisfy (8); (ii)  $w^*$  satisfies (9); (iii)  $L^*$  satisfies the labor market clearing condition,  $L = 1$ ; (iv)  $\{z^{\ell*}\}_{\ell=1}^4$  satisfy (11)-(14), and  $x^*$  satisfies (15); (v)  $\bar{x}^*$  satisfies (23); (vi)  $x_e^*$  satisfies (22); (vii)  $Y^*$  satisfies (9); (viii)  $R_d^*$  satisfies (26); (ix)  $C^*$  satisfies (28); (x) the BGP growth rate  $g^*$  satisfies (24); (xi) the distribution of technology gap portfolio composition  $\tilde{\mu}(\mathcal{N})$  and  $\mathcal{F}^*$  satisfy  $\text{inflow}(\mathcal{N}) = \text{outflow}(\mathcal{N})$ ; and (xii) the technology gap distribution  $\{\mu(\Delta^\ell)\}_{\ell=1}^4$  follows (25).

### 3 Empirics

In this section, we empirically examine innovation heterogeneity and validate model predictions on the “market-protection” and “technology-barrier” effects. We identify the causal effect of competition on firm innovation and composition, and analyze the industry-level association between

<sup>45</sup>We assume outside firms use final goods from their economy for production and R&D.

technological barriers and firm entry. We use the rise in foreign firm entry into U.S. markets after China's WTO accession in 2001 as a quasi-experimental increase in competitive pressure.

### 3.1 Data and Measurement

To compile comprehensive data on firm innovation and foreign competition shock, we combine the USPTO PatentsView database, the Longitudinal Business Database (LBD), the Longitudinal Firm Trade Transactions Database (LFTTD), the Census of Manufactures (CMF), the Compustat Fundamental Annual database, the NBER-CES database, and the tariff data in [Feenstra et al. \(2002\)](#).

The LBD tracks the universe of establishments and firms in the U.S. non-farm private sector with at least one paid employee annually from 1976 onward.<sup>46</sup> We aggregate establishment-level data into firm-level using firm identifiers.<sup>47</sup> Firm size is measured by total employment or payroll, and firm age by the age of the oldest establishment of the firm when the firm is first observed in the data. The firm's main industry of operation is based on the six-digit North American Industry Classification System (NAICS) code of the establishment with the highest employment.<sup>48</sup>

The LFTTD tracks all U.S. international trade transactions at the firm level from 1992 onward. It provides information such as the U.S. dollar value of shipments, the origin and destination countries, and a related-party flag indicating whether the U.S. importer and the foreign exporter are related by ownership of at least 6 percent.<sup>49</sup>

The USPTO PatentsView database records all patents ultimately granted by the USPTO from 1976 onward.<sup>50</sup> This database provides comprehensive details for patents, including application and grant dates, technology class, citation, and the name and address of patent assignees. In our analyses, we rely on the citation-adjusted number of utility patents as the main measure of firm innovation.<sup>51</sup> Using the patent-level information, we distinguish domestic innovation from foreign innovation, and assess the extent to which each patent represents own-innovation. The patent application year is

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<sup>46</sup>Details for the LBD and its construction can be found in [Jarmin and Miranda \(2002\)](#).

<sup>47</sup>An establishment corresponds to the physical location where business activity occurs. Establishments that are operated by the same entity, identified through the Economic Census and the Company Organization Survey, are grouped under a common firm identifier.

<sup>48</sup>Time-consistent NAICS codes for LBD establishments are constructed by [Fort and Klimek \(2018\)](#), and the 2012 NAICS codes are used throughout the entire analysis.

<sup>49</sup>[Bernard et al. \(2009\)](#) describe the LFTTD in greater detail.

<sup>50</sup>See <https://patentsview.org/download/data-download-tables>.

<sup>51</sup>See [Cohen \(2010\)](#) for a comprehensive review of the literature on the determination of firm/industry innovative activity and related patent measures.

used for the innovation year.

We link the USPTO patent database to the LBD to track firm patenting over time. Failure to match a patent assignee with its LBD firm counterpart can mismeasure firm innovation changes.<sup>52</sup> Since the USPTO patent data lacks a longitudinally consistent firm identifier, we build our own crosswalk between the two datasets by adopting the internet search-aided algorithm as in [Autor et al. \(2020\)](#).<sup>53</sup> We pool all patents granted up to December 26, 2017, and use patent applications up to 2007 in our main analyses to avoid a right censoring issue arising from patents applied for but not yet granted. Table [G1](#) in the Online Appendix reports summary statistics.

The quinquennial CMF provides detailed information about the U.S. manufacturing establishments and products they produce. It contains product-level details such as product codes and the value of shipment. We use five-digit SIC codes (for the pre-2002 years) or seven-digit NAICS codes (for 2002 onward) to define a product. We obtain the U.S. tariff schedules from [Feenstra et al. \(2002\)](#) to measure the industry-level Trade Policy Uncertainty (TPU) as a proxy for foreign competitive pressure. Lastly, all nominal values are converted to 1997 U.S. dollars, using the industry-level deflator from the NBER CES Manufacturing Industry Database for manufacturing industries and the Consumer Price Index from the BEA for other industries.<sup>54</sup>

For our main analyses, we use LBD and Compustat firms matched to USPTO patents, CMF firms, and industry-level trade data spanning from 1982 to 2007.<sup>55</sup>

**Own-innovation vs. Creative Destruction** We follow [Akcigit and Kerr \(2018\)](#) and use the self-citation ratio, the ratio of self-citations to total citations, as a measure of the likelihood a patent is used for own-innovation vs. creative destruction.<sup>56</sup> A higher (lower) self-citation ratio implies a greater probability that a patent reflects own-innovation (creative destruction).<sup>57</sup> This

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<sup>52</sup>The USPTO assigns patent applications to self-reported firm names, which are frequently misspelled.

<sup>53</sup>This algorithm utilizes the machine-learning capacities of internet search engines. The entire matching methodology is outlined in our accompanying paper [Ding et al. \(2022\)](#). We also apply it when linking Compustat to USPTO.

<sup>54</sup>The NBER CES data are compiled by [Becker et al. \(2013\)](#) (<http://www.nber.org/nberces/>).

<sup>55</sup>Our procedure links patents to the firms initially reported by the USPTO as owners and does not track ownership changes resulting from, for example, M&A activities. We expect our analysis not to be contaminated by firms substituting innovation with acquisitions of other firms, particularly given that U.S. M&A activities began declining around 2000 and did not fully recover by 2007, as shown in [Phillips and Zhdanov \(2023\)](#).

<sup>56</sup>Each granted patent is required to cite all prior patents on which it builds itself. When a cited patent belongs to the owner of the citing patent, these citations are called self-citations.

<sup>57</sup>Thus, 100% self-citation means the patent is used for own-innovation with a 100% probability, and 0% self-citation means the patent is used for creative destruction with a 100% probability.

is because the more an idea is based on the firm’s own knowledge stock (self-citation), the more likely the innovation is used to improve the firm’s existing products (own-innovation). Alternatively, we measure own-innovation by the number of patents with a self-citation ratio above a certain threshold (0% or 10%) and within-firm product sales concentration. In addition, we measure creative destruction by the number of new products added.

**Exogenous Competition Shock** Following [Pierce and Schott \(2016\)](#) and [Handley and Limão \(2017\)](#), we use the removal of trade policy uncertainty (TPU) as a measure of an exogenous competitive pressure shock. Specifically, we use the following industry-level tariff rate gaps between WTO members and non-market economies in the year 1999 as a proxy for the industry-level competitive pressure shock from China occurring in 2001.<sup>58</sup>

$$NTRGap_j = Non\ NTR\ Rate_j - NTR\ Rate_j\ (\text{for industry } j).$$

For multi-industry firms, we use the employment-weighted average of  $NTRGap_j$ .<sup>59</sup>

The removal of TPU encouraged Chinese firms to enter U.S. markets and export their products ([Pierce and Schott, 2016](#)), which captures an exogenous increase in competitive pressure by foreign firms and directly maps into an increase in  $\bar{x}_o$  in our model.

## 3.2 Heterogeneity in Firm Innovation

In this section, we present empirical evidence of innovation heterogeneity in terms of i) learning time, ii) quality improvement, and iii) economic outcomes. In this analysis, we focus on the pre-2000 sample, as the rise in competitive pressure after 2000 changed firm innovation decisions and may influence our empirical results.

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<sup>58</sup>Nonmarket economies such as China are by default subject to relatively high tariff rates, known as non-Normal Trade Relations (non-NTR) or column 2 tariffs, when they export to the U.S. On the other hand, the U.S. offers WTO member countries NTR or column 1 tariffs, which are substantially lower than non-NTR tariffs. Although the U.S. granted temporary NTR status to China from 1980, the U.S. Congress voted on a bill to revoke China’s temporary NTR status every year from 1990 to 2001 after the Tiananmen Square protests in 1989. This caused uncertainty about whether the low tariffs would revert to non-NTR rates. Following an agreement on China’s entry into the WTO, the U.S. Congress passed a bill granting China permanent NTR (PNTR), and PNTR was implemented on January 1, 2002. The PNTR has reduced trade policy uncertainty, more for industries with a large prior gap between NTR and non-NTR tariff rates. See [Pierce and Schott \(2016\)](#) for details.

<sup>59</sup>Table G2 in the Online Appendix reports summary statistics of the NTR-related measures.

Table 1: Backward Citation Gap and Self-Citation Ratio

	Citation gaps	Citation gaps	Citation gaps
Self-citation ratio	-2.290*** (0.018)	-2.450*** (0.018)	-2.592*** (0.020)
Observations	728,721	728,721	728,721
Fixed effects	none	<i>ct</i>	<i>it, ct</i>

Notes: Constant terms are omitted for brevity. Robust standard errors are displayed below each coefficient. The mean (standard deviation) of the backward-citation gap and self-citation ratio are 6.81 (3.50), and 0.13 (0.22), respectively. Observations are unweighted. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

**Learning Time** Using U.S. patents and assignee data from 1982 to 1999, we demonstrate that learning others’ technologies takes time in the process of creative destruction, which is both a key assumption and contribution of our model. Although we have limited data to directly estimate learning time, we infer it from the gap between a patent’s application year and the application years of the patents it cites (backward citation gaps).<sup>60</sup>

Suppose that learning and building on other firms’ technologies takes more time. In that case, the gap between the creation year of a technology (patent) and the creation years of the technologies it builds on (i.e., backward-cited patents) will be larger. As a result, patents representing creative destruction should have larger backward citation gaps compared to those representing own-innovation. To test this hypothesis, we estimate the following regression model:

$$CitationGap_{ipct} = \alpha + \beta SelfCite_{ipct} + \delta_{it} + \delta_{ct} + \varepsilon_{ipct}.$$

where  $CitationGap_{ipct}$  represents the average backward citation gap for all patents cited by firm  $i$ ’s patent  $p$  created in year  $t$  and belonging to CPC subsection  $c$ ;  $SelfCite_{ipct}$  is the self-citation ratio of patent  $p$ ;  $\delta_{it}$  is a firm-year fixed effect; and  $\delta_{ct}$  is a CPC technology-year fixed effect.

Table 1 shows a negative relationship between the average backward citation gap and the self-citation ratio, regardless of the fixed effect specifications. This indicates that patents used for creative destruction (with lower self-citation ratios) take longer to develop from existing technologies (larger

<sup>60</sup>The granting of the patent legally affirms the embodied idea as a novel and useful advancement over prior knowledge, as reflected in its citations. Thus, in principle, a citation of Patent X by Patent Y indicates that X is prior knowledge upon which Y builds (Jaffe et al., 1993).



backward citation gaps).<sup>61,62</sup>

**Quality Improvement** We also provide empirical evidence on the differentiated quality improvement between own-innovation and creative destruction ( $\eta > \lambda$ ). Innovation quality is measured in two ways: the number of forward citations as a proxy for scientific value and the stock market response to patent news as a measure of market value (Kogan et al., 2017).<sup>63</sup> Since the market value of innovation is available only for publicly traded firms, we limit this analysis to patenting firms in Compustat from 1982 to 1999. To test if patent quality differs by the two innovation types, we estimate the following regression model:

$$Quality_{ipct} = \alpha + \beta_1 SelfCite_{ipct} + \beta_2 X_{it-1} + \delta_{ct} + \varepsilon_{ipct},$$

where  $Quality_{ipct}$  represents either the log of market value (M-value) or the log of one plus the number of forward citations (S-value) for patent  $p$  by firm  $i$  in year  $t$ , within CPC subsection  $c$ ;  $SelfCite_{ipct}$  is the self-citation ratio for patent  $p$ ;  $X_{it-1}$  is firm  $i$ 's market capitalization in period  $t - 1$  as the baseline measure for firm size; and  $\delta_{ct}$  is a CPC technology-year fixed effect.<sup>64</sup>  $\delta_{ct}$  allows for patent quality comparisons within each market while accounting for varying forward citation trends across technologies and years.<sup>65</sup>

Table 2 presents the results. Note that patent market value is mechanically correlated with market capitalization (0.67 in our sample), as it is the product of its estimated stock return and the firm's market capitalization (Kogan et al., 2017). Also, larger firms tend to produce patents with higher self-citation ratios (correlation = 0.16). After controlling for these factors, we find a negative relationship between market value and closeness to own-innovation, which remains significant with firm fixed effects (column 2-3). Similar results hold for scientific value (columns 4-6). These

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<sup>61</sup>Consistent results are obtained by either using the minimum backward citation gap or limiting the sample to the patents owned by Compustat firms only. The results are available upon request.

<sup>62</sup>We cannot distinguish between learning time and the time needed for successful creative destruction, but this distinction is irrelevant for incumbent firms aiming to protect their markets.

<sup>63</sup>Data sourced from the paper's website (<https://github.com/KPSS2017/Technological-Innovation-Resource-Allocation-and-Growth-Extended-Data>), updated through 2023.

<sup>64</sup>Market capitalization is calculated as the product of the closing market price (PRCC.F) and the number of common shares outstanding (CSHO).

<sup>65</sup>Due to the lack of product market information for patents, we use the main CPC subsection as a proxy for product markets.

Table 2: Patent Quality and Self-Citation Ratio

	M-value	M-value	M-value	S-value	S-value	S-value
Self-citation	0.192*** (0.008)	-0.289*** (0.006)	-0.027*** (0.005)	-0.110*** (0.008)	-0.082*** (0.008)	-0.047*** (0.008)
Market cap <sub>-1</sub>		0.431*** (0.001)	0.289*** (0.003)		-0.025*** (0.001)	-0.043*** (0.005)
Observations	360,750	360,750	360,750	360,750	360,750	360,750
Fixed effects	<i>ct</i>	<i>ct</i>	<i>i, ct</i>	<i>ct</i>	<i>ct</i>	<i>i, ct</i>

*Notes:* The regression sample consists of patenting firms in Compustat from 1982 to 1999. The estimates for firm (*i*), CPC technology-year (*ct*) fixed effects, and the constant are suppressed. Robust standard errors are displayed below each coefficient. Observations are unweighted. The mean (standard deviation) of the market value, scientific value, and self-citation ratio are 2.07 (1.32), 2.88 (1.22), and 0.17 (0.24), respectively. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

findings support the view that creative destructions are of higher quality than own-innovations.<sup>66</sup>

**Economic Outcomes** Lastly, we use either LBD firms linked to USPTO or CMF firms and compare the impacts of the two innovation types on firm performance metrics. We run the following regression for the census years 1982, 1987, 1992, and 1997:

$$\Delta Y_{ijt+5} = \beta_1 Pat_{ijt} + \beta_2 Self_{ijt} + \mathbf{X}_{ijt}\gamma_1 + \delta_{jt+5} + \varepsilon_{ijt+5}.$$

$\Delta Y_{ijt+5}$  represents the DHS growth of firm employment, the number of industries (six-digit NAICS) added, revenue productivity growth, the number of products added, or the growth in within-firm product market concentration between  $t$  and  $t + 5$ ;  $Pat_{ijt}$  is the citation-adjusted number of patents (in log) at  $t$ ; and  $Self_{ijt}$  is the citation-adjusted average self-citation ratio at  $t$  for firm  $i$  in industry  $j$ . Firm and industry controls include firm age, log payroll, the past five years of U.S. patent growth in firm technology fields, innovation intensity, and public firm status. The regression is unweighted, with standard errors clustered at the firm level.

Table 3 shows that firm patenting is positively associated with the growth of employment and productivity, as well as the number of industries or products added. However, this association

<sup>66</sup>The results remain robust with different firm size measures and firm-year fixed effects. The results are available upon request.

Table 3: Real Effect of Innovation on Employment Growth and Industry Added

	$\Delta$ Employment	#Industries added	$\Delta$ TFPR	#Products added	$\Delta$ HHI
#patents	0.036*** (0.010)	0.102*** (0.011)	0.118** (0.055)	0.358** (0.085)	-0.012 (0.023)
Avg. self-citation	-0.256** (0.109)	-0.158** (0.079)	-0.027 (0.053)	-0.274*** (0.102)	0.154** (0.069)
Observations	5,400	5,400	5,700	5,700	5,700
Fixed effects	$jt$	$jt$	$jt$	$jt$	$jt$

*Notes:* The baseline set of controls is included. The estimates for industry ( $j$ ) and year ( $t$ ) fixed effects, and the coefficient associated with the binary indicator are suppressed due to disclosure restrictions. The first two columns are based on the LBD-USPTO firms, and the last three columns are based on the CMF firms. Robust standard errors adjusted for clustering at the firm-level are displayed below each coefficient. Observations are unweighted. Observation counts are rounded due to Census Bureau disclosure avoidance procedures. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

weakens if the new patent has a higher self-citation ratio (or is closer to own-innovation).<sup>67,68</sup> We check the robustness using alternative innovation measures. See Appendix H.2.

### 3.3 Market-Protection Effect

Next, we validate one of the key model predictions: the “market-protection effect.” Following [Pierce and Schott \(2016\)](#), we use a Difference-in-Difference (DD) specification to identify the effect of increasing competitive pressure from Chinese firm entry on the U.S. firm innovation as follows:

$$\begin{aligned}
\Delta y_{ijp} = & \beta_1 Post_p \times NTRGap_{ijp0} \times InnovIntens_{ijp0} + \beta_2 Post_p \times NTRGap_{ijp0} \\
& + \beta_3 Post_p \times InnovIntens_{ijp0} + \beta_4 NTRGap_{ijp0} \times InnovIntens_{ijp0} \\
& + \beta_5 NTRGap_{ijp0} + \beta_6 InnovIntens_{ijp0} + \mathbf{X}_{ijp0}\gamma_1 + \mathbf{X}_{jp0}\gamma_2 + \delta_j + \delta_p + \alpha + \varepsilon_{ijp},
\end{aligned} \tag{29}$$

where  $p \in \{1992 - 1999, 2000 - 2007\}$ ,  $y_{ijp}$  is either i) the total citation-adjusted number of patents (overall innovation), or ii) the citation-weighted average self-citation ratio (closeness to

<sup>67</sup>The mean (and standard deviation) of the citation-adjusted logged number of patents is 1.284 (1.125), and the counterpart for the citation-adjusted average self-citation ratio is 0.050 (0.101). The result, along with this, implies that for average firms, creating one more patent is associated with a 1.32 pp (3.6/2.718) increase in their employment growth as  $\exp(1) \approx 2.718$ . Also, since average firms have the average self-citation ratio of 0.05, an 1% increase in the self-citation ratio is associated with a 0.0128 pp ( $-0.256 \times 0.05 \times 0.01 \times 100$ ) decrease in their employment growth.

<sup>68</sup>Note [Akcigit and Kerr \(2018\)](#) also show that own-innovation contributes less to firm employment growth, which is consistent with our result in the first column of Table 3.

own-innovation) for firm  $i$  in industry  $j$ , and  $\Delta y_{ijp}$  is the DHS (Davis et al., 1996) growth rate of  $y$  between the start-year and end-year for each period  $p$ .<sup>69</sup>

To maximize the sample size, we include firms that applied for at least one patent in the start-year and at least one patent in or before the end-year for each period. We compute the DHS growth rates for the longest available span of years. We also require firms to have at least one patent before the start-year of each period, or to have an age greater than 0, to avoid the impact of firm entry. The sample comprises all patenting LBD firms meeting these three criteria and excludes FIRE (finance, insurance, and real estate) industries.

$Post_p$  is a dummy for the post-treatment period (2000-2007).  $\mathbf{X}_{ijp0}$  and  $\mathbf{X}_{jp0}$  are vectors of firm and industry controls, respectively, measured at the start-year for each period  $p$ .<sup>70</sup>  $\delta_j$  is an industry fixed effect (six-digit NAICS), and  $\delta_p$  is a period fixed effect. The regression is unweighted, and standard errors are clustered by industry (six-digit NAICS). Firms in low TPU industries are the control group, whereas high TPU industry firms are the treatment group. We use the 1992 and 2000 cohorts of firms to gauge firm innovation before and after the policy change in December 2001, minimizing policy-driven changes in firm composition.

$InnovIntens_{ijp0}$  is the lagged five-year average of the ratio of the number of patent applications to total employment for firm  $i$ . This proxies the technological advantage the firm has.<sup>71</sup> It is measured in the start year for each period  $p$  and is normalized by its time-average at the two-digit NAICS level to control for industry effects. The model predicts  $\beta_1 > 0$  when the dependent variable is the changes in self-citation ratio (i.e., the likelihood of patents being used for own-innovation).

Table 4 presents the regression results, which align with our model predictions in several dimensions.<sup>72</sup> First, the first two columns show that the foreign competitive pressure shock has no statistically significant effect on firms' overall innovation, regardless of the firm controls included. According to our model, as competition intensifies, firms adjust their own innovation based on the technological advantages accumulated within their markets. However, firms universally reduce

<sup>69</sup>The long-difference regression specification is standard in settings with a slow-moving process, such as innovation or technological progress (e.g., Acemoglu and Restrepo, 2020). This specification controls for firm fixed effects.

<sup>70</sup>Baseline firm controls include: firm employment, firm age, the past five-year growth of U.S. patents in the CPC technology classes in which the firm operates, and a dummy variable for publicly traded firms. Industry control variables include NTR rates measured at the start of each period.

<sup>71</sup>In addition, we use alternative measures for technological advantages, as discussed in the following section.

<sup>72</sup>To conserve space, Table 4 reports the main coefficient estimates for the triple interaction and the DD-term only. The full results are available upon request.

Table 4: Market-Protection Effect

	$\Delta\text{Patents}$	$\Delta\text{Patents}$	$\Delta\text{Self-cite}$	$\Delta\text{Self-cite}$
NTR gap $\times$ Post	0.238 (0.237)	0.071 (0.283)	-0.075 (0.257)	-0.062 (0.291)
$\times$ Innovation intensity	0.077 (0.231)	-0.054 (0.242)	0.732** (0.299)	0.795*** (0.277)
Observations	6,500	6,500	6,500	6,500
Fixed effects	$j, p$	$j, p$	$j, p$	$j, p$
Controls	no	baseline	no	baseline

*Notes:* The baseline controls include the past five-year U.S. patent growth in firms' own technology fields, log employment, firm age, NTR rate, and a dummy for publicly traded firms. The estimates for industry ( $j$ ) and the period ( $p$ ) fixed effects, along with the coefficient associated with the binary indicator, are suppressed due to disclosure restrictions. The constant is also omitted. Robust standard errors, adjusted for clustering at the level of the firms' major industries, are displayed below each coefficient. Observations are unweighted, and observation counts are rounded due to the Census Bureau disclosure avoidance procedures. For the sake of space, only the main coefficients are presented. Full results are available on request. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

their creative destruction. Considering both changes in own-innovation and creative destruction, the overall effect of competition on firm innovation need not be statistically significant.

However, when examining the effect on own-innovation by substituting the dependent variable with the growth rate of the self-citation ratio, the effect becomes positive and statistically significant, as indicated in the last two columns.<sup>73</sup> This supports the market-protection effect. The estimated coefficient implies a 4.2 percentage points increase in the growth rate of the average self-citation ratio during the period 2000-2007 for a firm with an average lagged innovation intensity (0.18) in an industry with an average NTR gap (0.291). Given that the average value of the seven-year growth rate of the average self-citation ratio between 2000 and 2007 is 28.2 percentage points, this effect corresponds to a 15.0% increase in own-innovation by firms with technological advantages.

The estimated effect is economically important as well. Combined with Table 3, the positive association between firm-level patenting and employment growth decreases by 17.1% (from 1.32pp to 1.10pp) for innovation-intensive firms following the increased competitive pressure.<sup>74</sup> Similar implications can also be inferred for productivity growth and the intensity of adding a new product or entering a new industry.

<sup>73</sup>As firms do not change their overall innovation, the increasing self-citation ratio implies that innovative firms (those above the average innovation intensity) increase their own-innovation while decreasing their creative destruction.

<sup>74</sup>Innovation-intensive firms are those with innovation intensity one standard deviation above the average.

**Validity of the Identification Strategy and Robustness Tests** We confirm the validity of our identification and main results across several dimensions: i) parallel pre-trends assumption, ii) industry-level NTR gap measures, iii) input-output network effects, iv) adjusting sampling bias, v) different standard error clustering, vi) robustness of innovation intensity measures, vii) alternative measures for technological advantage, viii) alternative measures for the two types of innovations, and ix) additional controls (e.g., cumulative patents, firm payroll, industry-level skill and capital intensities, import/export dummies). More details are provided in Online Appendix H.3 and H.4.<sup>75</sup>

### 3.4 Technological Barrier Effect

To test the “technological barrier effect,” we run the following industry-level regression for the four census years during the pre-shock period (1982-1999):

$$FirmEntry_{jt} = \beta TechBarrier_{jt} + \delta_j + \delta_t + \alpha + \varepsilon_{jt}. \quad (30)$$

$FirmEntry_{jt}$  is the firm entry rate, and  $TechBarrier_{jt}$  is the technological barrier in industry  $j$  at year  $t$ . We measure the industry-level technological barrier using the skewness of the firm-level TFPR distribution (normalized by the industry frontier level). A right-skewed distribution indicates more firms near the frontier, implying intensive innovation within that industry. We normalize firm-level TFPR by the industry frontier level and use the top 5th or 10th percentile value.<sup>76</sup>

Table 5 indicates that firm entry is lower in industries with higher technological barriers, consistent with the technological barrier effect predicted in our model.

## 4 Quantitative Analysis

In this section, we calibrate the model to the U.S. manufacturing sector in 1992 and conduct counterfactual exercises to analyze the aggregate effects of increased competition and shifts in firm innovation composition. To better align the model with data, we expand the baseline model

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<sup>75</sup>Tables G11-G13 in the Online Appendix are consistent with our model prediction: higher competitive pressure reduces the number of new products added (creative destruction) for all firms; increases the number of patents with a high level of self-citation (own-innovation) for innovative firms; and increases product market concentration.

<sup>76</sup>Note that this is the inverse of the TFPR gap in Aghion et al. (2005) (i.e., 1-TFPR gap).

Table 5: Technological Barrier Effect

	Firm entry	Firm entry
Technological barriers	-0.012** (0.006)	-0.016** (0.007)
Observation	1,300	1,300
Fixed effects	$j, t$	$j, t$
Tech. barrier thresholds	Top 5%	Top 10%

Note: Industries are NAICS6 defined in the Census of Manufacturers. Technological barriers are measured by either the top 5th or top 10th percentile level of firm-level TFPR (normalized by the frontier level) within each industry. The first column uses the top 5th percentile, and the second column uses the top 10th percentile. Estimates for industry ( $j$ ) and year ( $t$ ) fixed effects as well as the constant are suppressed. Observations are unweighted. Observation counts are rounded due to Census Bureau disclosure avoidance procedures. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table 6: Parameter Estimates

External calibration			Internal calibration		
Param.	Description	Value	Param.	Description	Value
$\beta$	Time discount rate	0.947	$\hat{\chi}$	Scale of own-innovation	0.044
$\hat{\psi}$	Curvature of own-innovation	2.000	$\tilde{\chi}$	Scale of creative destr.	0.405
$\tilde{\psi}$	Curvature of creative destr.	2.000	$\tilde{\chi}^e$	Scale of startup R&D	1.689
$\tilde{\psi}^e$	Curvature of startup R&D	2.000	$\lambda$	Step size, own-innovation	1.040
$\theta$	Quality share, final goods	0.109	$\eta$	Step size, creative destr.	1.075
			$\mathcal{E}_o$	Mass of outside entrants	0.007

in Section 2 by allowing firms to undertake multiple creative destructions simultaneously, which depends on the number of products firms own, as in Klette and Kortum (2004).<sup>77</sup>

## 4.1 Calibration

There are eleven structural parameters in the model, as listed in Table 6. The five parameters in the left column are externally calibrated, while the six parameters in the right column are internally calibrated to match moments associated with firm-level variables and the import penetration ratio in the U.S. manufacturing sector.<sup>78</sup> We use the import penetration ratio because it is an observable

<sup>77</sup>See Online Appendix D and E for further details and computational algorithms used to solve the model.

<sup>78</sup>The average number of products and the number of products added are from the 1992 CMF. The high-growth firm growth rate is sourced from the LBD (Decker et al., 2016). Data on manufacturing imports and exports for the import penetration ratio come from Schott (2008), while data on manufacturing value added and productivity are from the NBER-CES Manufacturing Industry Database. The firm entry rate is taken from the BDS.

Table 7: Target Moments

Moment	Data	Model	Moment	Data	Model
Number of products	2.3	2.3	Avg. productivity growth (%)	1.9	1.9
Number of products added	0.3	0.3	High-growth firm growth (%)	22.5	22.3
Firm entry rate (%)	7.6	7.6	Import penetration rate (%)	15.3	15.3

moment that partially reflects an exogenous change in competitive pressure.

**Externally Calibrated Parameters** The time discount factor ( $\beta$ ) is set to 0.947, which corresponds to an annual interest rate of 5.6%. The curvature parameters of the three R&D cost functions ( $\hat{\psi}$ ,  $\tilde{\psi}$ ,  $\tilde{\psi}^e$ ) are taken from [Acemoglu et al. \(2018\)](#) and [Akcigit and Kerr \(2018\)](#). We set the average profit-to-sales ratio  $\theta (= \int_f \frac{profit_f}{sales_f} df)$  to match the quality share in final goods production (10.9%) reported in [Akcigit and Kerr \(2018\)](#). We normalize the mass of potential domestic startups.

**Internally Calibrated Parameters** The remaining six parameters are internally calibrated to minimize the objective function,  $\min \sum_{i=1}^6 \frac{|\text{model moment}_i - \text{data moment}_i|}{\frac{1}{2}|\text{model moment}_i| + \frac{1}{2}|\text{data moment}_i|}$ , with the six target moments in Table 7. Although the parameters are jointly calibrated, the most relevant moments for each set of parameters can be noted. The R&D scales for own-innovation and creative destruction ( $\hat{\chi}$ ,  $\tilde{\chi}$ ) are set to match the average number of products and the number of products added per firm. The startup R&D scale ( $\tilde{\chi}^e$ ) matches the firm entry rate. We target the average productivity growth rate and the employment growth rate of high-growth firms (90th percentile) to determine the quality multipliers for own-innovation ( $\lambda$ ) and creative destruction ( $\eta$ ). Lastly, the mass of potential outside entrants ( $\varepsilon_o$ ) targets the import penetration ratio in manufacturing sector.

**Model Properties** The calibration results are presented in Table 7, where our model performs well in matching the target moments overall. In particular, it matches well the number of products and products added. Conditional on the number of products, the number of products added reflects both innovation intensity and the duration for learning the frontier technology in data. These variables may vary across products or technologies, and the learning time may not be strictly annual. By assuming a fixed annual learning duration, the calibration adjusts the R&D cost parameters (innovation intensity) to align the data with the annual frequency of the model.



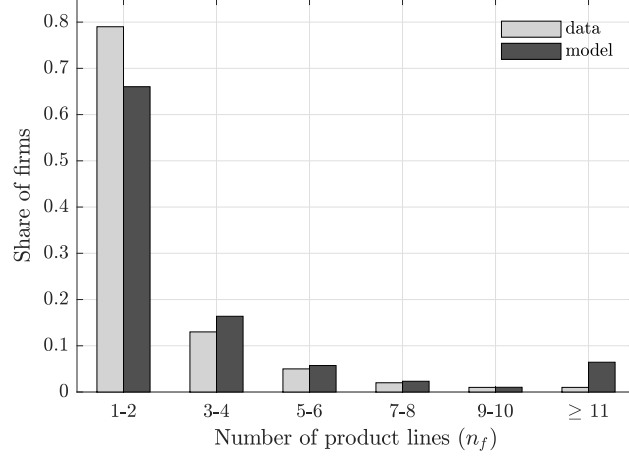


Figure 1: Firm Size Distribution: Theory and Data

Also, note that targeting the growth of high-growth firms helps us pin down the relative size of the two step sizes  $\lambda$  and  $\eta$ , since creative destruction has a greater impact on the right tail of the firm growth distribution. Our model aligns well with these moments, and the estimated parameter values suggest that creative destruction contributes 1.88 ( $\frac{0.075}{0.04}$ ) times more to growth compared to own-innovation. Also, the estimates satisfy the assumption  $\lambda^2 > \eta > \lambda$ , even without imposing any parameter restrictions in the calibration process.

As an external validation of our model, Figure 1 compares the firm size distribution (in terms of the number of products) between the model and the data, which is untargeted. While the model exhibits a thicker right tail, indicating more firms with 11 or more products, it aligns closely with the data. Another important untargeted moment is the aggregate R&D to sales ratio.<sup>79</sup> Our model estimate is 4.6%, which is close to the data estimate of 4.1% in Akcigit and Kerr (2018).<sup>80</sup>

## 4.2 Counterfactual Exercises

Using the calibrated model, we assess the impact of heightened competition on overall firm innovation, the composition of firm innovation, and aggregate growth. To do so, we increase the mass of potential outside entrants  $\mathcal{E}_o$  by 83%, corresponding to the rise in import penetration ratio

<sup>79</sup>The aggregate R&D to sales ratio is defined as the ratio of total R&D expenses (the sum of R&D expenses for own-innovation and creative destruction) of domestic incumbents to their total sales.

<sup>80</sup>Our model incorporates all resources used for product quality improvement and product scope expansion into R&D expenses, some of which might not be fully captured in the data.

in the U.S. manufacturing sector from 1992 to 2007 (from 15.3% to 25.1%).<sup>81,82</sup>

**Increasing Competitive Pressure from Outside Firms** Table 8 shows that an exogenous increase in outside firm entry leads to a rise in the aggregate creative destruction arrival rate  $\bar{x}$  and results in three key effects: i) the expected profits of both types of innovations ( $\{A_\ell\}_{\ell=1}^4$  and  $B$ ) decrease, known as the Schumpeterian effect (Panel B); ii) incumbents intensify own-innovation to protect their existing product lines, especially those with a technology gap of  $\Delta^\ell > 1$ , referred to as the market-protection effect; and iii) the market-protection efforts along with the increase in  $\bar{x}$  raise the average technology gap, making it harder for firms to take over product markets via creative destruction, labeled as the technological barrier effect (Panel C).<sup>83</sup> This effect arises as creative destruction by outside firms and successful own-innovation shift the technology gap distribution. Specifically, the density of  $\Delta^2$ ,  $\Delta^3$ , and  $\Delta^4$  increases, reducing the conditional takeover probability and the ex-ante value of successful product takeover.<sup>84</sup> Consequently, firm incentives for creative destruction and domestic firm entry get reduced, which contributes to the decline observed in  $x$  and  $x_e$  in Table 8. Our novel mechanism comes through ii) and iii).

In the general equilibrium, these effects come into play together and interact. For example, the technological barrier effect in iii) additionally influences the aggregate creative destruction arrival rate  $\bar{x}$ , causing a feedback loop involving i) to iii). The total decline in  $x$  and  $x_e$  in Table 8 results from the combined impact of the Schumpeterian and technological barrier effects.<sup>85</sup>

Finally, Table 9 summarizes how aggregate variables change in response to the increased competitive pressure from outside firms. The aggregate R&D to sales ratio of domestic incumbents drops, indicating that the decrease in creative destruction outweighs the increase in own-innovation. Thus, creative destruction R&D intensity—the ratio of domestic R&D expenses for creative destruction to

<sup>81</sup>Although we use a trade-related moment in our analysis, we do not intend to assess the effect of trade. We borrow the competition aspect embedded in the trade. For a detailed analysis of the effect of trade on innovation, see Jo (2024), which extends our framework to a two-country model to explore the effects of globalization on business dynamism.

<sup>82</sup>In Online Appendix G, we explore an additional counterfactual analysis involving an increase in the creative destruction arrival rate by domestic startups. This comparison allows us to assess the results in light of varying sources of increased competitive pressure.

<sup>83</sup> $\bar{A} \equiv \sum_{\ell=1}^4 A_\ell \mu(\Delta^\ell)$ . Table E1 in Online Appendix F presents details for the changes in  $\{A_\ell\}_{\ell=1}^4$ .

<sup>84</sup> $\bar{x}_{\text{takeover}}$  decreases from 73.2% to 73.0%. The increase in densities  $\mu(\Delta^3)$  and  $\mu(\Delta^4)$  is solely attributed to increased creative destruction by outside firms. The higher density of  $\Delta^2$  reflects both increased own-innovation driven by the market-protection effect and creative destruction by outside firms.

<sup>85</sup>A decomposition reveals that 17.0% and 15.0% of the total change in  $x$  and  $x_e$  are ascribed to the technological barrier effect (due to the shifts in  $\mu(\Delta^\ell)$ , given all else equal).

Table 8: Counterfactual: Increasing Competitive Pressure in the U.S.

Description	Variables	Before	After	% Change
Panel A: Changes in Firm Innovation				
creative destruction arrival rate by outside firms (%)	$\bar{x}_o$	3.3	5.5	66.4%
aggregate creative destruction arrival rate (%)	$\bar{x}$	21.5	21.9	1.51%
prob. of own-innovation ( $\Delta^1 = 1$ , %)	$z^1$	16.9	16.8	-0.43%
prob. of own-innovation ( $\Delta^2 = \lambda$ , %)	$z^2$	57.8	57.9	0.19%
prob. of own-innovation ( $\Delta^3 = \eta$ , %)	$z^3$	39.7	39.7	0.13%
prob. of own-innovation (%) ( $\Delta^4 = \frac{\eta}{\lambda}$ , %)	$z^4$	37.3	37.4	0.05%
prob. of creative destruction, incumbents (%)	$x$	16.8	16.5	-1.33%
prob. of creative destruction, potential startups (%)	$x_e$	4.02	3.97	-1.33%
Panel B: Changes in Innovation Values				
Average of own-innovation values	$\bar{A}$	0.167	0.165	-1.04%
Creative destruction value	$B$	0.011	0.011	-2.6%
Panel C: Changes in Technology Gap Distribution				
Technology gap distribution (shares)	$\Delta^1 = 1$	0.541	0.539	-0.4%
	$\Delta^2 = \lambda$	0.314	0.314	0.2%
	$\Delta^3 = \eta$	0.116	0.118	1.1%
	$\Delta^4 = \frac{\eta}{\lambda}$	0.028	0.029	1.4%

Table 9: Changes in Aggregate Moments

Description	Before	After	% Change
Panel A: Changes in the Aggregate Moments			
R&D to sales ratio (%)	4.6	4.5	-1.6%
Creative destruction R&D intensity (%)	63.9	63.1	-1.2%
Average number of products	2.3	2.2	-5.5%
Total mass of domestic firms	0.386	0.361	-6.4%
Panel B: Changes in the Aggregate Growth and Decomposition			
Average productivity growth by domestic firms (%)	1.9	1.7	-11.0%
Growth from domestic own-innovation (%)	1.0	0.9	-11.4%
Growth from domestic creative destruction (%)	0.7	0.6	-13.0%
Growth from domestic startups (%)	0.2	0.2	-1.7%

total R&D expenses—decreases. The average number of products per firm falls, aligning with the empirical findings in [Bernard et al. \(2011\)](#). Furthermore, the total number of domestic firms falls.

In addition, Panel B shows that the average productivity growth of domestic firms ( $g_d$ ) declines. This decrease is attributed to shifts in firm-level innovation intensities and the mass of firms. Keeping the mass of domestic incumbents constant, 12.7% of this decline in growth can be attributed to

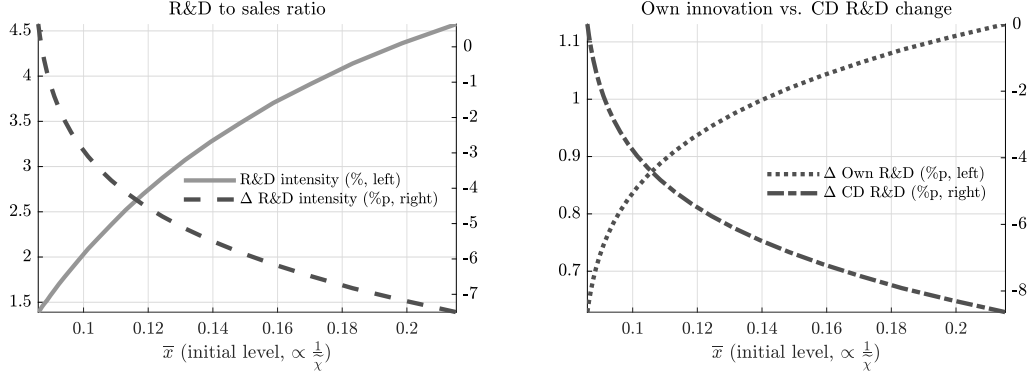


Figure 2: Decomposition of Innovation Change Across different  $\tilde{\chi}$

changes in firm-level creative destruction.<sup>86</sup>

**Comparison: Economy with High Creative Destruction Costs** To compare implications across environments with different innovation structures, we re-calibrate the model to hypothetical economies characterized by lower creativity (less creative destruction due to higher frictions) compared to the U.S. Specifically, we increase the parameter associated with creative destruction costs ( $\tilde{\chi}$ ) up to 80 times higher than the baseline value of 0.405, while keeping other parameters unchanged. We then perform the same counterfactual analysis.

Figure 2 shows the results across different initial levels of  $\bar{x}$  (reflecting different degrees of initial competitive pressure) corresponding to varied values of  $\tilde{\chi}$  (that negatively affects  $\bar{x}$ ). The U.S. economy represents the highest  $\bar{x}$  level in the figures. The left panel shows the initial R&D to sales ratios and their changes following a competitive pressure shock, and the right panel breaks down the latter into the changes in own-innovation and creative destruction (CD).

Own-innovation R&D expenses rise as competitive pressure intensifies, while creative destruction R&D expenses decline. However, the decrease in creative destruction R&D is more pronounced when its cost  $\tilde{\chi}$  is low (high initial  $\bar{x}$ ). While both types of innovation shift similarly across different economies, own-innovation increases more than creative destruction in environments with higher costs of creative destruction (lower initial  $\bar{x}$ ), whereas the opposite holds in environments with lower costs (higher initial  $\bar{x}$ ).<sup>87</sup> Thus, in economies where creative destruction costs are high, aggregate R&D increases in response to competitive pressure, in contrast to the U.S.

Table 10 compares aggregate moments between the U.S. and an economy with high creative

<sup>86</sup>For more detailed breakdowns, refer to Tables E2 and E3 in Online Appendix F. Note that if taking into account

Table 10: Aggregate Moment Comparison: U.S. vs. High Creative Destruction Cost Economy

Moment	Baseline	High CD Costs	After Shock	% Change
R&D to sales ratio (%)	4.58	1.39	1.41	1.0%
Creative destruction R&D intensity (%)	63.9	8.6	7.8	-9.8%
Average number of products	2.3	1.0	1.0	-0.2%
Avg. growth by domestic firms (%)	1.9	1.4	1.3	-9.7%

destruction costs ( $\tilde{\chi} \times 80$ ), as well as the response of the latter economy to a competition shock. The first two columns show that the low creativity economy exhibits lower dynamism than the U.S. with less R&D, fewer products, and lower average productivity growth. The last two columns indicate that both economies respond similarly to increased foreign competition, except for the R&D to sales ratio, where the difference arises from the initially lower level of creative destruction in the low creativity economy. Despite the increased domestic innovation, the growth attributable to domestic innovation drops in this economy. The reduction is associated with decreases in creative destruction by domestic incumbents and startups, coupled with a decline in the mass of domestic incumbents.<sup>88</sup>

**Policy Implications** Our results highlight the importance of examining innovation composition in evaluating the aggregate implications of innovation and informing effective policy design. The two types of innovation contribute differently to the economy, even at the individual level, and interact in ways that shape aggregate outcomes. Increased overall innovation driven by defensive own-innovation may be less beneficial than it appears, as it contributes less to economic growth than creative destruction and restricts firm entry. This underscores the nuanced effects of heightened competition, which, in low-creativity economies, can exacerbate the challenge of insufficient creative destruction.

The differences between the two innovation types extend beyond what is addressed in our paper. Peters (2020), for example, shows that creative destruction reduces misallocation by limiting the accumulation of market power by incumbents. This supports our claim that increases in overall innovation are not always unambiguously beneficial. Given their differing impacts on growth,

the contribution of outside firms, the aggregate growth rate increases.

<sup>87</sup>See Table E4 in Online Appendix F for detailed results when  $\tilde{\chi}$  is 80 times higher than the U.S.

<sup>88</sup>The version with a fixed mass of firms is presented in Table E5 in Online Appendix F. Note that this pattern holds even without the effect of the changes in firm mass.

market power, and resource allocation, policies that fail to account for these distinctions may lead to unintended economic consequences.

Moreover, understanding the composition of innovation can help reconcile disparate findings in the literature. For example, if European economies face higher creative destruction costs due to barriers such as complex approval processes or stringent labor regulations ([Peters, 2020](#); [Aghion et al., 2023](#)), our model predicts that increased foreign competition could raise overall innovation in Europe unlike in the U.S., primarily through shifts in innovation composition. This perspective extends [Aghion et al. \(2005\)](#) by integrating multiple strands of literature and emphasizing compositional changes as a key mechanism in determining the aggregate effects of competition.

In summary, our analysis highlights that innovation composition is pivotal in understanding the effects of competition on economic outcomes. Policies that overlook the distinct contributions of own-innovation and creative destruction risk misjudging their aggregate implications.

## 5 Conclusion

This paper investigates firm innovation incentives in the presence of imperfect technology spillovers and their aggregate implications. We show that learning frictions enable incumbents to strategically use own-innovation to protect their markets, thereby deterring rival entry and stifling creative destruction. Heightened competition amplifies these strategic choices and endogenous interactions, leading to shifts in the composition of innovation. Consequently, the overall impact of competition on innovation depends on the varying magnitudes of these shifts, which can differ across environments with different innovation cost structures. Our paper provides new insights into how firms strategically use innovations under learning frictions and how the strategic innovation decisions of firms impact the composition of innovation and the aggregate outcome. In addition, this framework bridges gaps in the existing literature, reconciles previous findings, and deepens our understanding of the complex relationship between competition and firm innovation.

**Supplementary Materials** The [Online Appendix](#) contains supplementary materials.

## Appendix: Proofs of Propositions

**Proof of Lemma 1 .** Consider the following two cases: 1) no ownership change between  $t - 1$  and  $t$ , and 2) ownership change happens between  $t - 1$  and  $t$ . In scenario 1),  $q_{j,t} = \Delta_{j,t} q_{j,t-1}$  with only  $\Delta_{j,t} \in \{\Delta^1 = 1, \Delta^2 = \lambda\}$  as a result of own-innovation. In scenario 2),  $q_{j,t} = \eta q_{j,t-2}$  holds. Let's consider all possible cases where i)  $\Delta_{j,t} = 1$ , ii)  $\Delta_{j,t} = \lambda$ , iii)  $\Delta_{j,t} = \eta$ , iv)  $\Delta_{j,t} = \frac{\eta}{\lambda}$ , v)  $\Delta_{j,t} = \frac{\eta^n}{\lambda^m}$  with  $n \geq m > 0$ , and vi)  $\Delta_{j,t} = \frac{\lambda^n}{\eta^m}$  with  $n > m > 0$ . These are the only possible values  $\Delta$  can assume, given that product quality can only be adjusted by three step sizes (1,  $\lambda$ , and  $\eta$ ) between two periods without technology regression ( $q_t < q_{t-1}$ ).

- i)  $\Delta_{j,t} = 1$ : For this to be true,  $q_{j,t} = q_{j,t-1}$  should hold. Since  $q_{j,t} = \eta q_{j,t-2}$ , we need  $q_{j,t-1} = \eta q_{j,t-2}$ . This is possible if there was creative destruction between  $t - 2$  and  $t - 1$ , and no own-innovation between  $t - 3$  and  $t - 1$ , leading to  $q_{j,t-2} = q_{j,t-3}$ .
- ii)  $\Delta_{j,t} = \lambda$ : For this to be true,  $\Delta_{j,t-1} = \frac{\eta}{\lambda}$  should hold, as  $\Delta_{j,t} = \frac{q_{j,t}}{q_{j,t-1}} = \frac{\eta q_{j,t-2}}{\Delta_{j,t-1} q_{j,t-2}}$ . This can be possible if there were own-innovation between  $t - 3$  and  $t - 2$ , and creative destruction between  $t - 2$  and  $t - 1$ , but no own-innovation between  $t - 2$  and  $t - 1$ . In this case,  $q_{j,t-2} = \lambda q_{j,t-3}$  and  $q_{j,t-1} = \eta q_{j,t-3}$  holds, and thus  $\Delta_{j,t-1} = \frac{q_{j,t-1}}{q_{j,t-2}} = \frac{\eta q_{j,t-3}}{\lambda q_{j,t-3}} = \frac{\eta}{\lambda}$  follows. So we have shown that both  $\Delta_{j,t} = \lambda$  and  $\Delta_{j,t} = \frac{\eta}{\lambda}$  are possible, and  $\Delta_{j,t} = \frac{\eta}{\lambda}$  can be realized only through creative destruction between  $t - 1$  and  $t$ .
- iii)  $\Delta_{j,t} = \eta$ : For this to be true,  $q_{j,t-1} = q_{j,t-2}$  should hold. This is possible if there was neither ownership change nor own-innovation between  $t - 1$  and  $t - 2$ .
- iv)  $\Delta_{j,t} = \frac{\eta}{\lambda}$ : This follows the illustration in case ii)
- v)  $\Delta_{j,t} = \frac{\eta^n}{\lambda^m}$  with  $n \geq m > 0$ : Suppose this is the case. As  $\Delta_{j,t} \notin \{\Delta^1 = 1, \Delta^2 = \lambda\}$ , there should be an ownership change between  $t - 1$  and  $t$ . Thus  $q_{j,t} = \eta q_{j,t-2}$  holds, implying  $q_{j,t-1} = \frac{\lambda^m}{\eta^{n-1}} q_{j,t-2}$ . Note that  $m \leq n - 1$  is not possible without technology regression. Thus,  $m = n$  (as  $m > n - 1$  and  $n \geq m > 0$ ). If  $\frac{\lambda^m}{\eta^{m-1}} < 1$ , this implies technology regression and can be ruled out. Suppose  $\frac{\lambda^m}{\eta^{m-1}} > 1$ . If  $m = 1$ , we are back to the cases ii) and iv). Suppose  $m > 1$ . As  $\frac{\lambda^m}{\eta^{m-1}} \neq 1$  or  $\lambda$ , there should be an ownership change between  $t - 2$  and  $t - 1$ . Thus,  $q_{j,t-1} = \eta q_{j,t-3}$  holds, implying  $q_{j,t-2} = \frac{\eta^m}{\lambda^m} q_{j,t-3}$ . If  $\Delta_{j,t} = \frac{\eta^n}{\lambda^m}$  is possible,

$q_{j,t-s} = \frac{\eta^m}{\lambda^m} q_{j,t-s-1}$  holds for even numbers  $s$ , and  $\frac{\lambda^m}{\eta^{m-1}} q_{j,t-s-1}$  holds for odd numbers  $s$ . Thus, in this case, either  $q_{j,1} = \frac{\eta^m}{\lambda^m} q_{j,0}$  or  $q_{j,1} = \frac{\lambda^m}{\eta^{m-1}} q_{j,0}$  must hold, which can be ruled out (or we assume this case does not occur). Thus,  $\Delta_{j,t} = \frac{\eta^n}{\lambda^m}$  with  $n \geq m > 0$  is not possible.

- vi)  $\Delta_{j,t} = \frac{\lambda^n}{\eta^m}$  with  $n > m > 0$ : Following the same argument, this case is not possible.

Therefore  $\Delta_{j,t}$  can assume only the four values of  $\{1, \lambda, \eta, \frac{\eta}{\lambda}\}$ .  $\square$

**Proof of Proposition 1.** Using the conjectured value function, we can decompose the expected value into two parts with the linearity of expectation: the expected value of existing product lines  $\mathbb{E} \left[ \sum_{\ell=1}^2 A_\ell \sum_{j \in \mathcal{J}^f | (\Delta'_j | \Delta_j) = \Delta^\ell} \Delta^\ell q_j \right]$  and the expected value for the new product line added through creative destruction  $\mathbb{E} \left[ \sum_{\ell=1}^4 A_\ell I_{\{\eta/\Delta_j = \Delta^\ell\}} \frac{\eta}{\Delta_j} q_j \right]$ . As the realization of own-innovation outcomes and the creative destruction are independent of the realization of creative destruction, the expected value of a new product line becomes:

$$\begin{aligned} \mathbb{E} \left[ \sum_{\ell=1}^4 A_\ell I_{\left\{ \frac{\eta}{\Delta_j} = \Delta^\ell \right\}} \frac{\eta}{\Delta_j} q_j \right] &= \sum_{I^x=0}^1 x^{I^x} (1-x)^{1-I^x} \mathbb{E}_{q_j, \Delta_j} \left[ \sum_{\ell=1}^4 A_\ell I_{\left\{ \frac{\eta}{\Delta_j} = \Delta^\ell \right\}} I^x \frac{\eta}{\Delta_j} q_j \right] \\ &= x \left[ \frac{1-z^3}{2} A_1 \mu(\Delta^3) + \left( 1 - \frac{z^4}{2} \right) A_2 \lambda \mu(\Delta^4) + A_3 \eta \mu(\Delta^1) + (1-z^2) A_4 \frac{\eta}{\lambda} \mu(\Delta^2) \right] \bar{q}. \end{aligned}$$

The terms in the bracket arise from the random property of creative destruction. The assigned product can have a technology gap of  $\Delta^\ell$  with a probability of  $\mu(\Delta^\ell)$ , and the probability of taking over this product line depends on its technology gap. Integrating over all possible qualities  $q_j$  over the entire set of available products gives us  $\bar{q}$ .<sup>89</sup>

The expected value of existing product lines can further be broken down into the four cases of  $\Delta$  and integrated as  $\sum_{\ell=1}^4 \mathbb{E} \left[ \sum_{\ell=1}^2 A_\ell \sum_{j \in \mathcal{J}^f | (\Delta'_j | \Delta_j) = \Delta^\ell} \Delta^\ell q_j \right]$ . To simplify the derivation, we reorder product quality  $q_j$  by its technology gap  $\Delta_j$  and categorize it into the following four groups:  $q_1^f = \{q_{j_1}, q_{j_2}, \dots, q_{j_{n_f^1}}\}$ ;  $q_2^f = \{q_{j_{n_f^1+1}}, \dots, q_{j_{n_f^1+n_f^2}}\}$ ;  $q_3^f = \{q_{j_{n_f^1+n_f^2+1}}, \dots, q_{j_{n_f^1+n_f^2+n_f^3}}\}$ ; and  $q_4^f = \{q_{j_{n_f^1+n_f^2+n_f^3+1}}, \dots, q_{j_{n_f^1+n_f^2+n_f^3+n_f^4}}\}$ ,  $q^f = \cup_{\ell=1}^4 q_\ell^f$ .

If  $\Delta = \Delta^1$  ( $\tilde{\ell} = 1$ ), the expected value can be rephrased as  $\sum_{i=1}^{n_f^1} [A_1(1-\bar{x})(1-z_i^1) + \lambda A_2(1-\bar{x})z_i^1] q_{j_i}$ ; if  $\Delta = \Delta^2$  ( $\tilde{\ell} = 2$ ), it becomes  $\sum_{i=n_f^1+1}^{n_f^1+n_f^2} [A_1(1-\bar{x})(1-z_i^2) + \lambda A_2 z_i^2] q_{j_i}$ ; if  $\Delta = \Delta^3$

<sup>89</sup>Note that individual firms only have information about the distribution of technology gaps  $\{\mu(\Delta^\ell)\}_{\ell=1}^4$  and the average quality level  $\bar{q}$ . That is, for an individual firm, a technology gap and product quality are independent considerations.



( $\tilde{\ell} = 3$ ), it is  $\sum_{i=n_f^1+n_f^2+1}^{n_f-n_f^4} [A_1 (1 - \frac{1}{2}\bar{x}) (1 - z_i^3) + \lambda A_2 z_i^3] q_{j_i}$  ; and if  $\Delta = \Delta^4$  ( $\tilde{\ell} = 4$ ), it is  $\sum_{i=n_f-n_f^4}^{n_f} [A_1(1 - \bar{x})(1 - z_i^4) + \lambda A_2(1 - \frac{1}{2}\bar{x}) z_i^4] q_{j_i}$  .

The  $B\bar{q}$  portion of the conjectured value function in  $\mathbb{E} \left[ V(\Phi^{f'} | \Phi^f) \mid \{z_j\}_{j \in \mathcal{J}^f}, x \right]$  can be expressed as  $\mathbb{E} B\bar{q}' = B(1 + g)\bar{q}$ , where  $g$  denotes the growth rate of product quality in a balanced growth path (BGP) equilibrium. Plugging this into the conjectured value function, we can rephrase the original value function as:

$$\sum_{i=1}^{n_f^1} A_1 q_{j_i} + \sum_{i=n_f^1+1}^{n_f^1+n_f^2} A_2 q_{j_i} + \sum_{i=n_f^1+n_f^2+1}^{n_f-n_f^4} A_3 q_{j_i} + \sum_{i=n_f-n_f^4+1}^{n_f} A_4 q_{j_i} + B\bar{q} =$$

$$\max_{\substack{x \in [0, \bar{x}], \\ \{z_i \in [0, \bar{z}]\}_{i=1}^{n_f}}} \left\{ \begin{aligned} & \sum_{i=1}^{n_f} \left[ \pi q_{j_i} - \hat{\chi} z_i^{\hat{\psi}} q_{j_i} \right] - \bar{q} \tilde{\chi} x^{\tilde{\psi}} \\ & + \tilde{\beta} \sum_{i=1}^{n_f^1} \left[ A_1 (1 - \bar{x}) (1 - z_i^1) + \lambda A_2 (1 - \bar{x}) z_i^1 \right] q_{j_i} \\ & + \tilde{\beta} \sum_{i=n_f^1+1}^{n_f^1+n_f^2} \left[ A_1 (1 - \bar{x}) (1 - z_i^2) + \lambda A_2 z_i^2 \right] q_{j_i} \\ & + \tilde{\beta} \sum_{i=n_f^1+n_f^2+1}^{n_f-n_f^4} \left[ A_1 (1 - \frac{1}{2}\bar{x}) (1 - z_i^3) + \lambda A_2 z_i^3 \right] q_{j_i} \\ & + \tilde{\beta} \sum_{i=n_f-n_f^4+1}^{n_f} \left[ A_1 (1 - \bar{x}) (1 - z_i^4) + \lambda A_2 (1 - \frac{1}{2}\bar{x}) z_i^4 \right] q_{j_i} \\ & + \tilde{\beta} x \left[ \frac{1}{2} (1 - z^3) A_1 \mu(\Delta^3) + (1 - \frac{1}{2} z^4) A_2 \lambda \mu(\Delta^4) \right. \\ & \quad \left. + A_3 \eta \mu(\Delta^1) + (1 - z^2) A_4 \frac{\eta}{\lambda} \mu(\Delta^2) \right] \bar{q} \\ & + \tilde{\beta} B(1 + g) \bar{q} \end{aligned} \right\}.$$

By taking the first-order conditions with respect to each innovation intensity, we get the optimal innovation decision rules, which depend solely on technology gaps. Substituting these optimal innovation intensities into the value function, equating the left-hand side (LHS) to the right-hand side (RHS), and collecting terms, we obtain the five coefficients of the conjectured value function.  $\square$

**Proof of Corollary 1.** Define  $\tilde{z}^\ell = \frac{\hat{\psi} \hat{\chi}}{\beta} (z^\ell)^{(\hat{\psi}-1)}$ . Then  $z^\ell > z^{\ell'} \Leftrightarrow \tilde{z}^\ell > \tilde{z}^{\ell'}$  for all  $\ell, \ell' \in [1, 4] \cap \mathbb{Z}$  under the condition  $\hat{\psi} > 1$ . Given  $\tilde{z}^2 - \tilde{z}^3 = \frac{1}{2}\bar{x}A_1 > 0$ ,  $\tilde{z}^2 - \tilde{z}^1 = \bar{x}\lambda A_2 > 0$ ,  $\tilde{z}^2 - \tilde{z}^4 = \frac{1}{2}\bar{x}\lambda A_2 > 0$ , and  $\tilde{z}^4 - \tilde{z}^1 = \frac{1}{2}\bar{x}\lambda A_2 > 0$ , we can obtain the following relationships:  $z^2 > z^3$ ,  $z^2 > z^1$ ,  $z^2 > z^4$ ,

and  $z^4 > z^1$ . Given  $\tilde{z}^1 = (1 - \bar{x})[\lambda A_2 - A_1] > 0$  in equilibrium,  $\lambda A_2 - A_1 > 0$  holds, and  $\tilde{z}^3 > \tilde{z}^4 \Leftrightarrow z^3 > z^4$  is derived. Thus, the order of  $\{z^\ell\}_{\ell=1}^4$  in equilibrium is  $z^2 > z^3 > z^4 > z^1$ .  $\square$

**Proof of Corollary 2.** The partial derivatives of  $\{z^\ell\}_{\ell=1}^4$  with respect to  $\bar{x}$  are (after removing the common terms)  $\frac{\partial z^1}{\partial \bar{x}} \Big|_{A_1, A_2} : -(z^1)^{2-\hat{\psi}}[\lambda A_2 - A_1] < 0$ ;  $\frac{\partial z^2}{\partial \bar{x}} \Big|_{A_1, A_2} : (z^2)^{2-\hat{\psi}} A_1 > 0$ ;  $\frac{\partial z^3}{\partial \bar{x}} \Big|_{A_1, A_2} : (z^3)^{2-\hat{\psi}} \frac{1}{2} A_1 > 0$ ; and  $\frac{\partial z^4}{\partial \bar{x}} \Big|_{A_1, A_2} : -(z^4)^{2-\hat{\psi}} [\frac{1}{2} \lambda A_2 - A_1] \geq 0$ , with  $A_1$  and  $A_2$  fixed. As  $\lambda A_2 - A_1 > 0$ , it follows that  $\frac{\partial z^1}{\partial \bar{x}} \Big|_{A_1, A_2} < 0$ . Similarly,  $\frac{\partial z^2}{\partial \bar{x}} \Big|_{A_1, A_2} > \frac{\partial z^3}{\partial \bar{x}} \Big|_{A_1, A_2}$  holds with  $z^2 > z^3$ , and  $\frac{\partial z^3}{\partial \bar{x}} \Big|_{A_1, A_2} > \frac{\partial z^4}{\partial \bar{x}} \Big|_{A_1, A_2}$  holds with  $z^3 > z^4$  and  $\lambda A_2 - A_1 > 0$ . However, the sign for  $\frac{1}{2} \lambda A_2 - A_1$  remains ambiguous.  $\square$

**Proof of Potential Startups' Problem.** With the value function defined for incumbents, we have  $\mathbb{E}V(\{(q'_j, \Delta'_j)\}) = x_e \left[ \frac{1}{2}(1-z^3)A_1\mu(\Delta^3) + (1-\frac{1}{2}z^4)A_2\lambda\mu(\Delta^4) + A_3\eta\mu(\Delta^1) + (1-z^2)A_4\frac{\eta}{\lambda}\mu(\Delta^2) \right] \bar{q} + x_e \left[ \frac{1}{2}(1-z^3)\mu(\Delta^3) + (1-\frac{1}{2}z^4)\mu(\Delta^4) + \mu(\Delta^1) + (1-z^2)\mu(\Delta^2) \right] B(1+g)\bar{q}$ , from which the optimal creative destruction choice for potential startups can be derived.  $\square$

**Proof of Proposition 2.** In this model, the output growth rate is the same as the product quality growth rate. For product  $j$  with quality  $q_j$  and a technology gap of  $\Delta_j = \Delta^\ell$ , we can derive the following law of motion of  $q_j$ :

$\Delta^1 : \begin{aligned} q'_j = \Delta^1 q_j & \text{ prob. } (1 - \bar{x})(1 - z^1) \\ q'_j = \Delta^2 q_j & \text{ prob. } (1 - \bar{x})z^1 \\ q'_j = \Delta^3 q_j & \text{ prob. } \bar{x} \\ q'_j = \Delta^4 q_j & \text{ prob. } 0 \end{aligned}$	$\Delta^2 : \begin{aligned} q'_j = \Delta^1 q_j & \text{ prob. } (1 - \bar{x})(1 - z^2) \\ q'_j = \Delta^2 q_j & \text{ prob. } z^2 \\ q'_j = \Delta^3 q_j & \text{ prob. } 0 \\ q'_j = \Delta^4 q_j & \text{ prob. } \bar{x}(1 - z^2) \end{aligned}$
$\Delta^3 : \begin{aligned} q'_j = \Delta^1 q_j & \text{ prob. } 1 - z^3 \\ q'_j = \Delta^2 q_j & \text{ prob. } z^3 \\ q'_j = \Delta^3 q_j & \text{ prob. } 0 \\ q'_j = \Delta^4 q_j & \text{ prob. } 0 \end{aligned}$	$\Delta^4 : \begin{aligned} q'_j = \Delta^1 q_j & \text{ prob. } (1 - \bar{x})(1 - z^4) \\ q'_j = \Delta^2 q_j & \text{ prob. } z^4 + \bar{x}(1 - z^4) \\ q'_j = \Delta^3 q_j & \text{ prob. } 0 \\ q'_j = \Delta^4 q_j & \text{ prob. } 0 \end{aligned}$

Following this, we can compute the expected growth rate of  $q_j$  ( $\mathbb{E}[q'_j | q_j] / q_j - 1$ ) and the aggregate growth rate in (24) by taking the expectation across all product lines.

Using the share of products owned by domestic incumbents ( $s_d = \mathcal{F}_d / \mathcal{F}$ ), the definition of  $\bar{x}$ ,

and the evolution of product quality, the growth rate can be decomposed as follows:

$$\begin{aligned}
g = & \underbrace{(\Delta^2 - 1) s_d \left[ (1 - \bar{x}) z^1 \mu(\Delta^1) + z^2 \mu(\Delta^2) + z^3 \mu(\Delta^3) + \left(1 - \frac{1}{2} \bar{x}\right) z^4 \mu(\Delta^4) \right]}_{\text{own-innovation by domestic incumbent firms}} \\
& + \underbrace{(\Delta^2 - 1) (1 - s_d) \left[ (1 - \bar{x}) z^1 \mu(\Delta^1) + z^2 \mu(\Delta^2) + z^3 \mu(\Delta^3) + \left(1 - \frac{1}{2} \bar{x}\right) z^4 \mu(\Delta^4) \right]}_{\text{own-innovation by foreign firms}} \\
& + \underbrace{(\overline{\Delta^{ex}} - 1) \mathcal{F}_d x \mu(\overline{\Delta^{ex}})}_{\text{creative destr. by domestic incumbents}} + \underbrace{(\overline{\Delta^{ex}} - 1) \mathcal{E}_d x_e \mu(\overline{\Delta^{ex}})}_{\text{creative destr. by domestic startups}} + \underbrace{(\overline{\Delta^{ex}} - 1) \bar{x}_o \mu(\overline{\Delta^{ex}})}_{\text{creative destr. by foreign firms}},
\end{aligned}$$

where  $\overline{\Delta^{ex}} \equiv \frac{\Delta^3 \mu(\Delta^1) + \Delta^4 (1 - z^2) \mu(\Delta^2) + \frac{1}{2} (1 - z^3) \mu(\Delta^3) + \Delta^2 \left(1 - \frac{1}{2} z^4\right) \mu(\Delta^4)}{\mu(\Delta^1) + (1 - z^2) \mu(\Delta^2) + \frac{1}{2} (1 - z^3) \mu(\Delta^3) + \left(1 - \frac{1}{2} z^4\right) \mu(\Delta^4)}$  is an increase in the average product quality due to creative destruction and successful business takeover, and  $\mu(\overline{\Delta^{ex}}) \equiv \mu(\Delta^1) + (1 - z^2) \mu(\Delta^2) + \frac{1}{2} (1 - z^3) \mu(\Delta^3) + \left(1 - \frac{1}{2} z^4\right) \mu(\Delta^4)$  is the share of product lines affected by creative destruction.  $\square$

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