

Recitation 7

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Panel regression

Motivation

- **Panel data:** We observe multiple individuals for multiple periods of time.

$$Y_{it} = \beta_0 + \beta_1 X_{1,it} + \dots + \beta_k X_{k,it} + u_{it}$$

$i = 1, 2, \dots, N \rightarrow$ individuals, $t = 1, 2, \dots, T \rightarrow$ time periods.

- **Balanced:** There are T datasets for each of the N individuals.
- **Unbalanced:** There are $t \leq T$ datapoints for some of the N individuals.

Advantages

- Panel data allows us to use more datasets.
- Panel data allows us to control for **unobserved heterogeneity** that are
 - 1 different across N entities but always remain same for T periods in a given state (**cross section fixed effect**)
 - 2 different across T time periods but remains the same for all N entities in a particular time period (**time fixed effects**)
 - 3 both of 1) and 2). (**two-way fixed effects**)

Key differences

- Suppose that $T = 2$ and we are interested in the relationship between vehicle related fatality rate (deaths per 10,000 people) and the beer tax. Suppose that we get these result for the two years

$$\hat{Y}_{i1} = 2.01 + 0.15X_{i1}$$

(0.15) (0.20)

$$\hat{Y}_{i2} = 1.86 + 0.44X_{i2}$$

(0.11) (0.20)

- In such case, one might suspect that there is an omitted variable bias that affects these coefficients.
 - Omitted variable specific to the states (Strictness of the relevant law)
 - Time-trends? (Specific to each of years 1 and 2)

Panel regression

How can panel regression do better?

- Let Z_i denote the strictness of state laws on DUI that are unchanging.
- Now write

$$Y_{i1} = \beta_0 + \beta_1 X_{i1} + \beta_2 Z_i + u_{i1}$$

$$Y_{i2} = \beta_0 + \beta_1 X_{i2} + \beta_2 Z_i + u_{i2}$$

- Subtract the second equation from the first to get

$$(Y_{i2} - Y_{i1}) = \beta_1(X_{i2} - X_{i1}) + \beta_2(Z_i - Z_i) + u_{i2} - u_{i1}$$

With Z_i being the same for all periods, the above equation is reduced to

$$(Y_{i2} - Y_{i1}) = \beta_1(X_{i2} - X_{i1}) + (u_{i2} - u_{i1})$$

- The Z_i variable has no role in this equation - because it is now gone.
- If we estimate this particular β_1 , we can obtain much more accurate estimates of the effect of bear tax on fatality rate.

Specific methodologies (cross-sectional FE)

- There are two ways of estimating the data when $T \geq 3$
- Least square dummy variables: Include $N - 1$ individual dummies
- Within estimation: Subtract “demeaned” equation from the original
- Use:

$$Y_{it} = \beta_0 + \beta_1 X_{it} + \beta_2 Z_i + u_{it} \quad (1)$$

where Z_i is the cross section fixed effect.

- Define $\alpha_i = \beta_0 + \beta_1 Z_i$. Then the above equation can be written as

$$Y_{it} = \beta_1 X_{it} + \alpha_i + u_{it} \quad (2)$$

- α_i term can be thought of as an effect of being an entity i , which is **correlated with X_{it}**

LSDV method

- Define a new variable D_{ki} as follows

$$D_{ki} = \begin{cases} 1 & \text{If } i = k \\ 0 & \text{Otherwise} \end{cases}, k \in \{1, 2, \dots, N\}$$

- Since we are going to include β_0 , a common intercept, in our regression we need to remove one of the N (dummy variable trap)
- Then we can write

$$Y_{it} = \beta_0 + \beta_1 X_{it} + \delta_2 D_{2i} + \dots + \delta_N D_{Ni} + u_{it} \quad (\text{LSDV})$$

- This equation gives different intercepts for each i (can you see why?), while keeping the slope on X_{it} constant at β_1
- Control for unobserved cross section fixed effect by allowing the intercept to differ by each i

WE method

- Define \bar{X}_i , \bar{Y}_i as sample mean of X_{it} , Y_{it} for given i over all possible t 's.

$$\bar{X}_i = \frac{1}{T} \sum_{t=1}^T X_{it}$$

Consequently, \bar{Y}_i can be written as

$$\bar{Y}_i = \frac{1}{T} \sum_{t=1}^T Y_{it} = \frac{1}{T} \sum_{t=1}^T (\beta_1 X_{it} + \alpha_i + u_{it}) = \beta_1 \bar{X}_i + \alpha_i + \bar{u}_i$$

- Subtract Y_{it} by \bar{Y}_i to get

$$Y_{it} - \bar{Y}_i = \beta_1 (X_{it} - \bar{X}_i) + (u_{it} - \bar{u}_i) \implies \tilde{Y}_{it} = \beta_1 \tilde{X}_{it} + \tilde{u}_{it}$$

- This process gets rid of α_i . Then, apply OLS estimation on this equation to get the within estimator

TWFE methods

- We have a DGP

$$Y_{it} = \beta_1 X_{it} + \alpha_i + \lambda_t + u_{it}$$

- LSDV: With an overall constant β_0 , we can put $N - 1$ individual and $T - 1$ time dummies
- WE: Demeaning should be done in the following method

$$Y_{it} - \bar{Y}_i - \bar{Y}_t + \bar{Y}$$

This (and only this) would allow us to get rid of both the α_i individual fixed effect and the λ_t time effects

Least square assumptions for panels

- P1** : $E[u_{it}|X_{i1}, \dots, X_{iT}, \alpha_i] = 0$. It means that the conditional mean of the u_{it} term does not depend on any of the X_{it} values for entity i , whether in the future or in the past.
- P2** : $(X_{i1}, \dots, X_{iT}, u_{i1}, \dots, u_{iT})$ is IID across $i = 1, \dots, n$. **This does not rule out the correlation between u_{it}, u_{ij} within entity i for different j and t ,** allowing serial correlation within the same entity
- P3** : (X_{it}, u_{it}) have nonzero finite fourth moments (outliers are very unlikely) so that the panel estimators have a distribution
- P4** : There is no perfect multicollinearity

→ Because of P2, we need to use **clustered standard error** at a cross-sectional level.