

econ144hw3

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```
library(timeSeries)

## Loading required package: timeDate
library(marima)
library(strucchange)

## Loading required package: zoo
##
## Attaching package: 'zoo'
## The following object is masked from 'package:timeSeries':
##
##     time<-
## The following objects are masked from 'package:base':
##
##     as.Date, as.Date.numeric
## Loading required package: sandwich
library(seasonal)

##
## Attaching package: 'seasonal'
## The following objects are masked from 'package:timeSeries':
##
##     outlier, series
library(dynlm)

## Warning: package 'dynlm' was built under R version 3.5.2
library(gdata)

## gdata: read.xls support for 'XLS' (Excel 97-2004) files ENABLED.
##
## gdata: read.xls support for 'XLSX' (Excel 2007+) files ENABLED.
##
## Attaching package: 'gdata'
## The following object is masked from 'package:stats':
##
##     nobs
## The following object is masked from 'package:utils':
##
##     object.size
```

```
## The following object is masked from 'package:base':
##
##     startsWith
require(graphics)
library("readxl")

## Warning: package 'readxl' was built under R version 3.5.2
library('xts')

##
## Attaching package: 'xts'

## The following objects are masked from 'package:gdata':
##
##     first, last
library('forecast');

## Warning: package 'forecast' was built under R version 3.5.2
library('fma')
library('expsmooth')
library('lmtest')
library('tseries')
library('Quandl')
library('fpp');
library('urca')
library(Hmisc)

## Warning: package 'Hmisc' was built under R version 3.5.2
## Loading required package: lattice
## Loading required package: survival
## Loading required package: Formula
## Loading required package: ggplot2
##
## Attaching package: 'Hmisc'

## The following objects are masked from 'package:base':
##
##     format.pval, units
setwd("/Users/Renaissance/Desktop/econ144/econ144hw3")
```

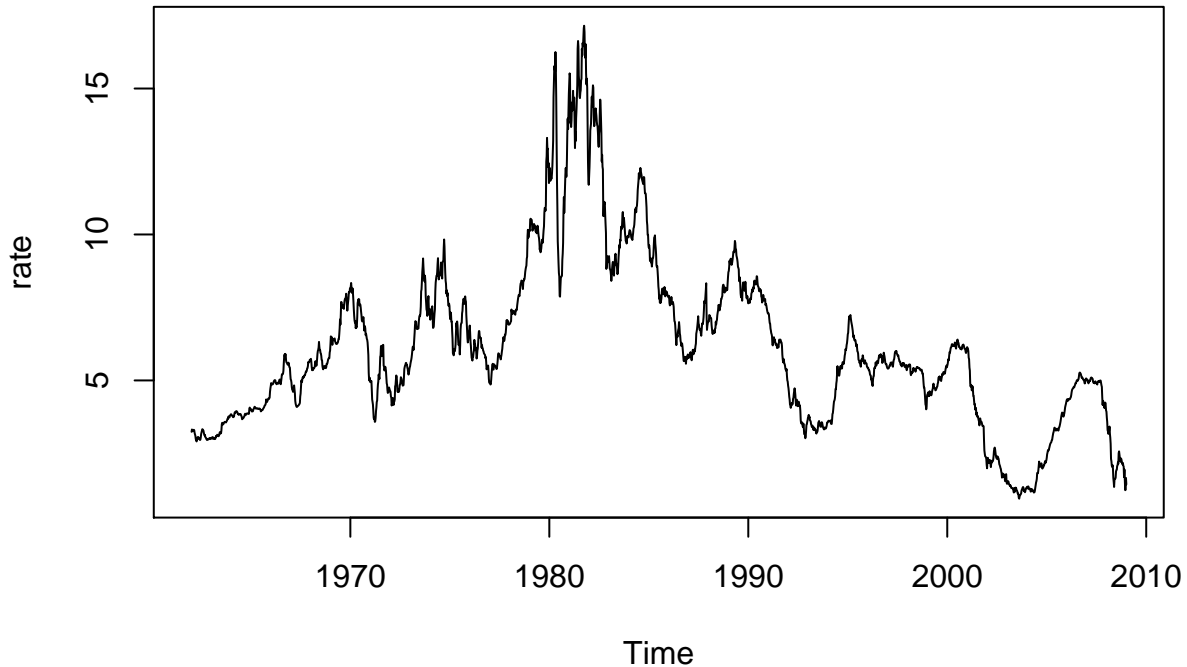
first question

```
hw31 <- read.table("w-gs1yr.txt")
inte <- interpNA(hw31$V4,method="linear")

## Warning in xy.coords(x, y, setLab = FALSE): NAs introduced by coercion
inte2 <- as.numeric(inte)
## a
```

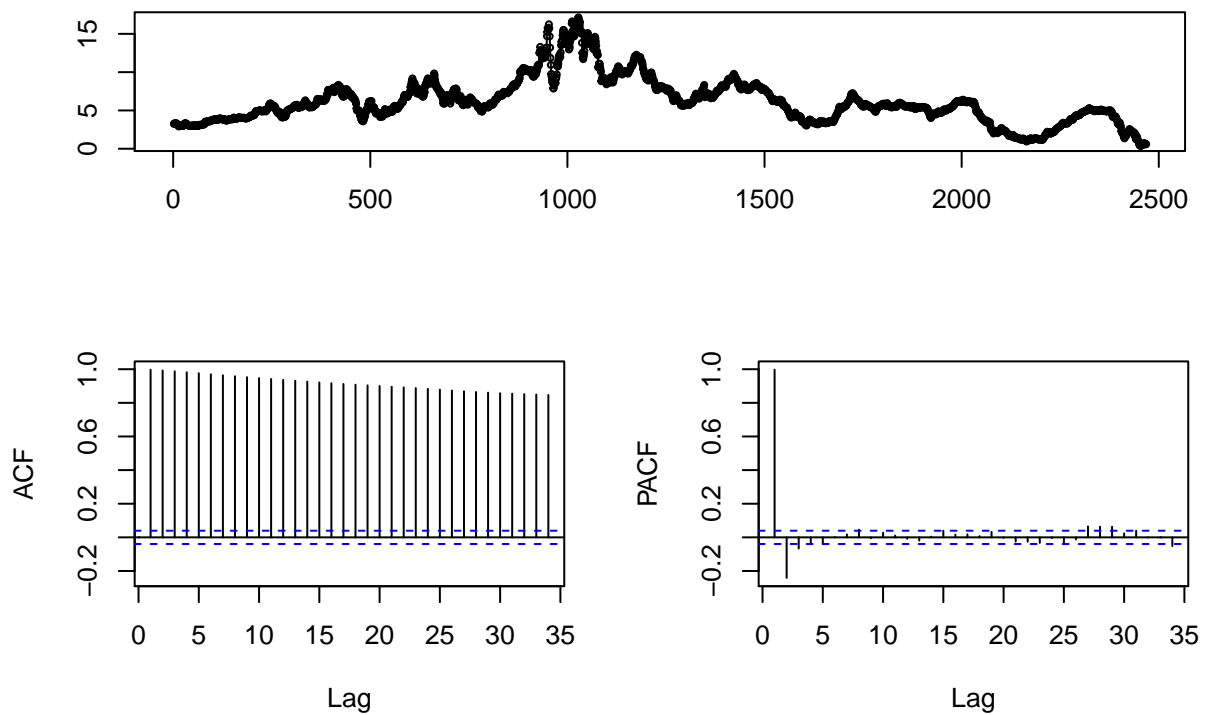
```
t1<-ts(inte,start=1962,2009,freq = 52)
plot(t1,ylab="rate",main="U.S. weekly Interest Rate time series plot")
```

U.S. weekly Interest Rate time series plot



```
tsdisplay(inte2,main = "U.S. weekly Interest Rate time series plot")
```

U.S. weekly Interest Rate time series plot



```

# ACF of this dataset is gradually decreasing in a very slow trend, in PACF graph there is only first t
## b
# try ar2
t12<-as.numeric(t1[2:2445])
ar2_q1 <- arma(t12,order=c(2,0))
summary(ar2_q1)

```

```

##
## Call:
## arma(x = t12, order = c(2, 0))
##
## Model:
## ARMA(2,0)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.5675255 -0.0590173  0.0008931  0.0625970  1.4208073
##
## Coefficient(s):
##      Estimate Std. Error t value Pr(>|t|)
## ar1      1.342579   0.018991  70.696 <2e-16 ***
## ar2     -0.345195   0.018996 -18.172 <2e-16 ***
## intercept 0.015525   0.008424   1.843  0.0653 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 0.03178, Conditional Sum-of-Squares = 77.57, AIC = -1487.67

```

```

# try ar1 +ma2
ar1ma2_q1 <- arma(t12,order=c(1,2))
summary(ar1ma2_q1)

```

```

##
## Call:
## arma(x = t12, order = c(1, 2))
##
## Model:
## ARMA(1,2)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.56965 -0.05907  0.00131  0.06342  1.42291
##
## Coefficient(s):
##      Estimate Std. Error t value Pr(>|t|)
## ar1      0.996452   0.001797  554.597 < 2e-16 ***
## ma1      0.323103   0.020092  16.081 < 2e-16 ***
## ma2      0.122156   0.018829   6.488 8.72e-11 ***
## intercept 0.020984   0.012216   1.718  0.0859 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:

```

```
## sigma^2 estimated as 0.03191, Conditional Sum-of-Squares = 77.89, AIC = -1475.37
```

```
# ar2 +ma1
```

```
ar2ma1_q1 <- arma(t12,order=c(2,1))
```

```
summary(ar2ma1_q1)
```

```
##
```

```
## Call:
```

```
## arma(x = t12, order = c(2, 1))
```

```
##
```

```
## Model:
```

```
## ARMA(2,1)
```

```
##
```

```
## Residuals:
```

```
##           Min           1Q           Median           3Q           Max
```

```
## -1.5995074 -0.0585295  0.0005104  0.0612440  1.4332955
```

```
##
```

```
## Coefficient(s):
```

```
##           Estimate Std. Error t value Pr(>|t|)
```

```
## ar1           1.536201    0.053559   28.683 < 2e-16 ***
```

```
## ar2           -0.538464    0.053460  -10.072 < 2e-16 ***
```

```
## ma1           -0.223032    0.062927   -3.544 0.000394 ***
```

```
## intercept     0.013537    0.006563    2.063 0.039158 *
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
```

```
## Fit:
```

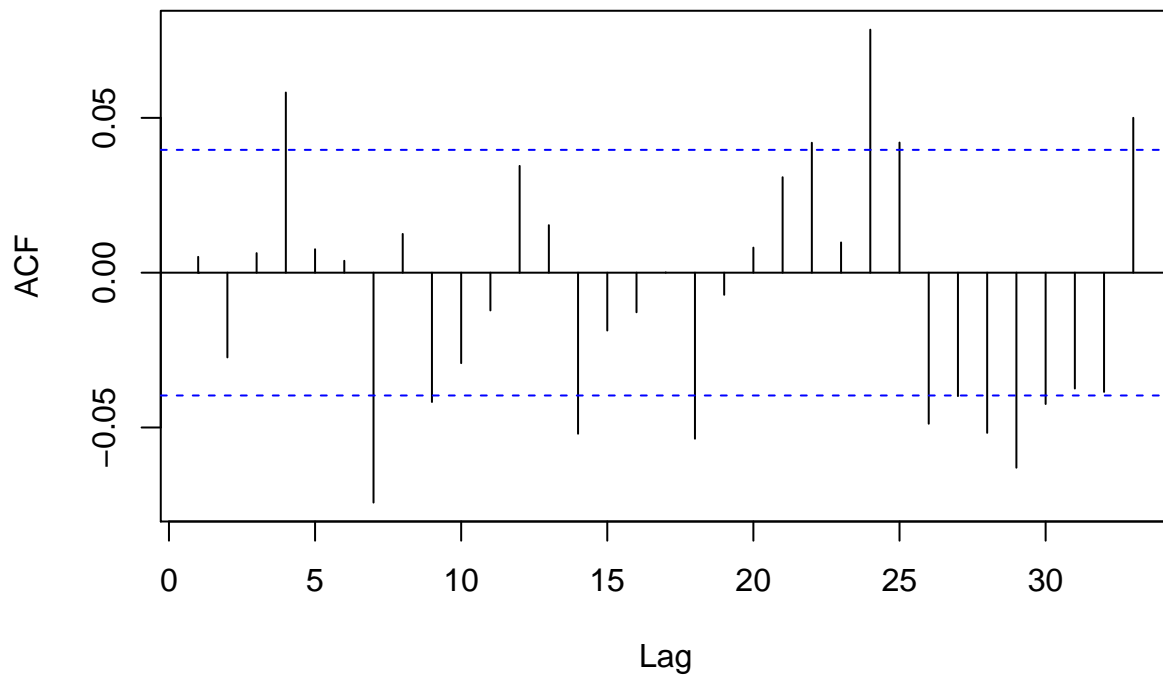
```
## sigma^2 estimated as 0.03163, Conditional Sum-of-Squares = 77.21, AIC = -1496.98
```

```
# I prefer arma(2,1) model. Because as it is shown in the sumamry, all coefficients are of significant.
```

```
## c
```

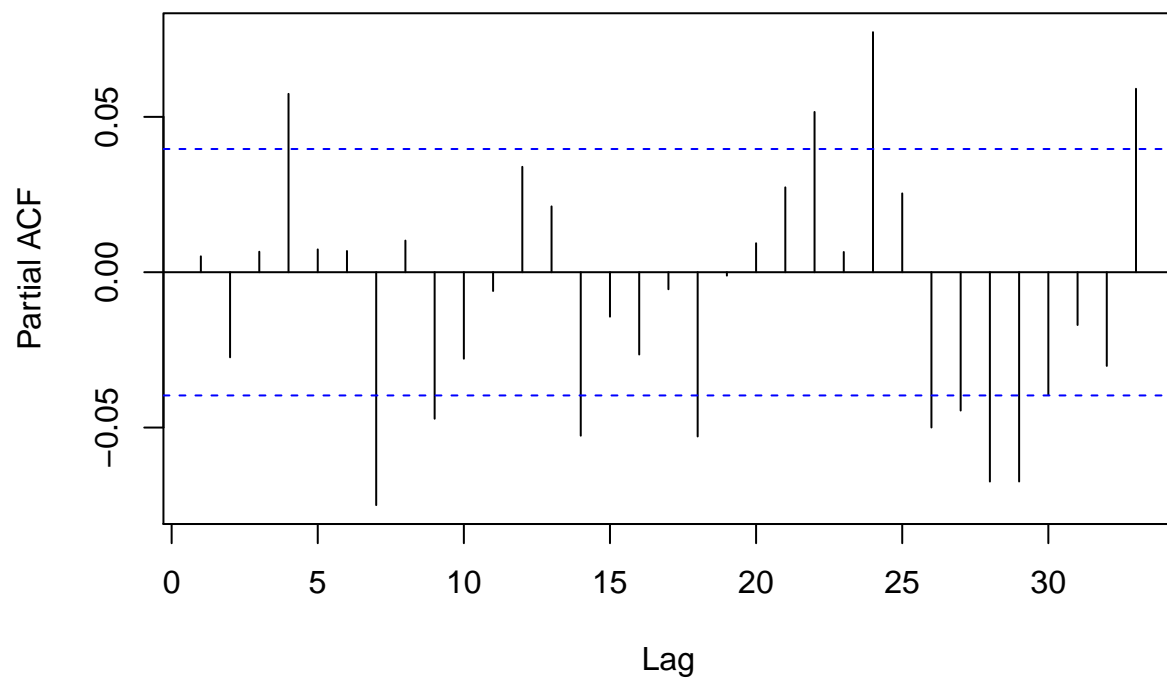
```
Acf(ar2ma1_q1$residuals)
```

Series ar2ma1_q1\$residuals



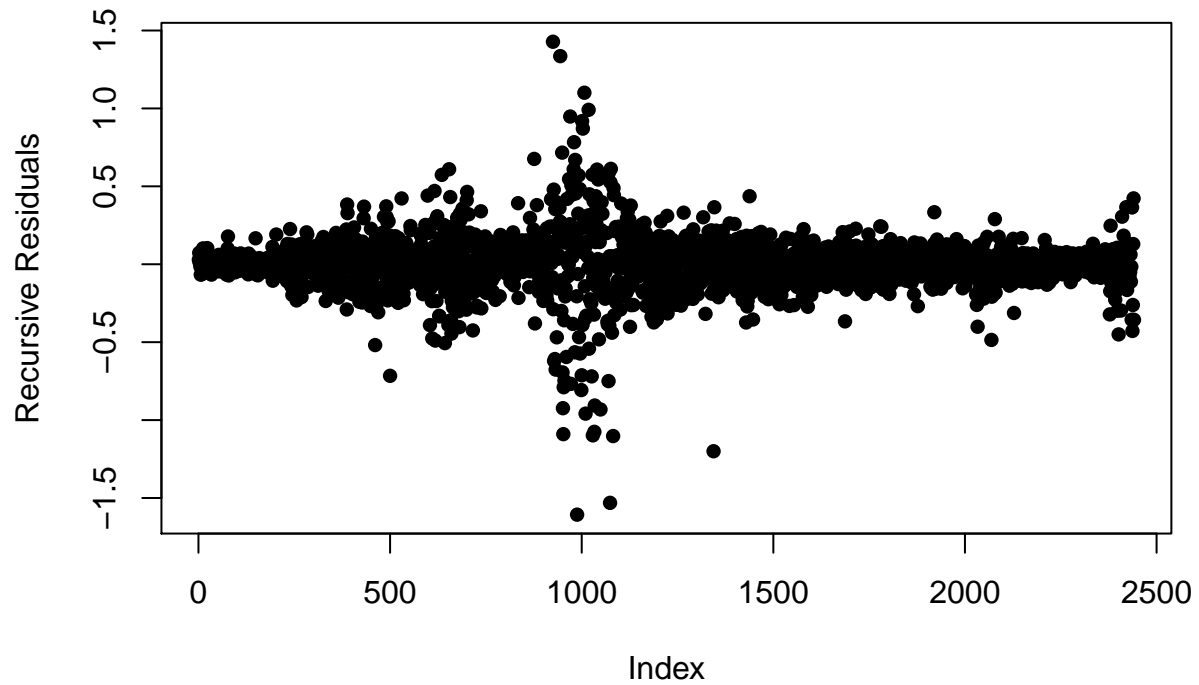
```
Pacf(ar2ma1_q1$residuals)
```

Series ar2ma1_q1\$residuals



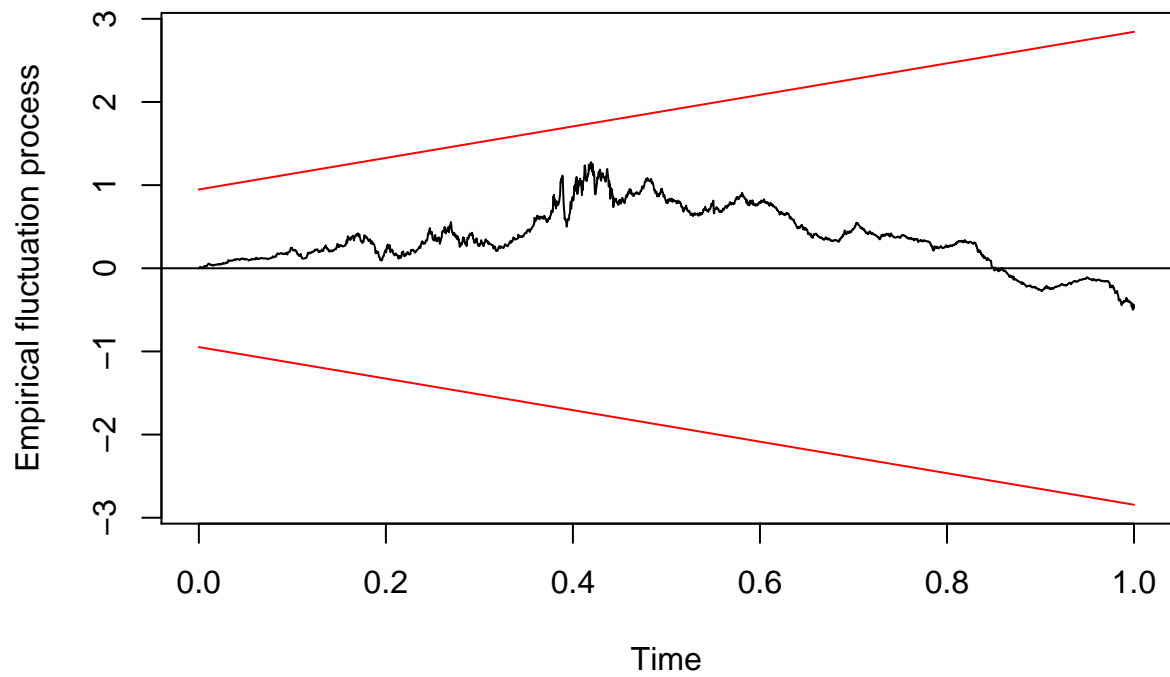
```
# In acf and pacf graph, there are some residuals of significance. Hence arma(2,1) may not be a good fit
## d
```

```
rec_q1=recresid(ar2ma1_q1$residuals~1)
plot(rec_q1, pch=16,ylab="Recursive Residuals")
```



```
# residuals are bounced around 0 randomly. But there are large residuals around index= 1000.
## e
plot(efp(ar2ma1_q1$residuals~1, type = "Rec-CUSUM"))
```

Recursive CUSUM test



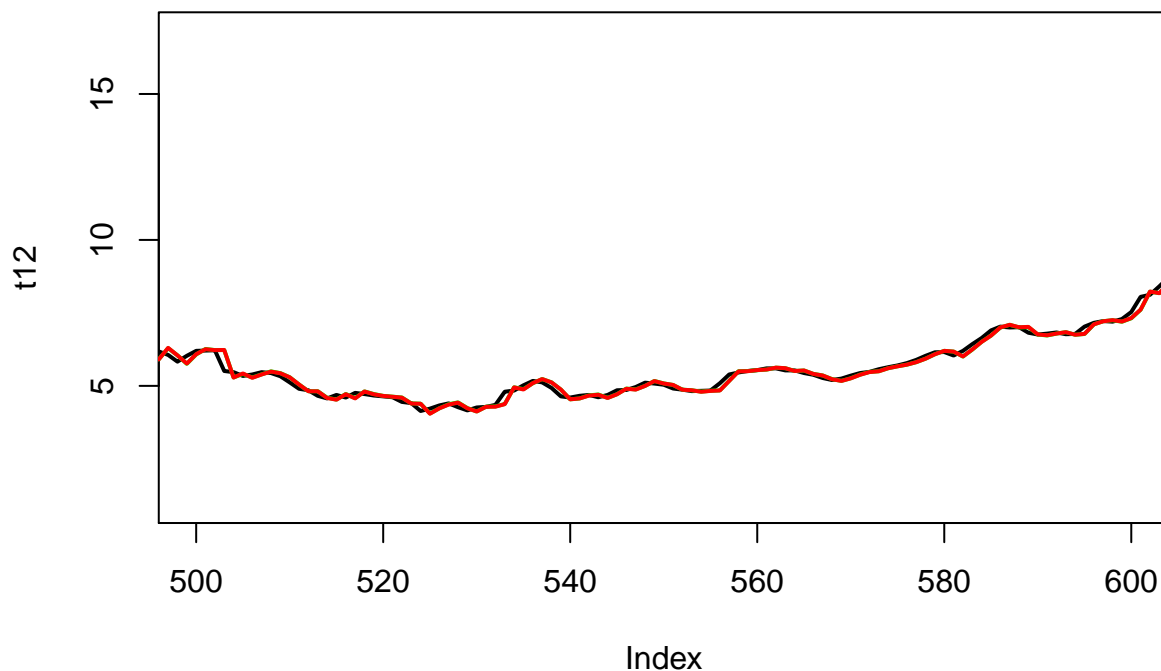
```

# this series passes the recursive cusum test.
## f
auto_p1 <- auto.arima(t12)
summary(auto_p1)

## Series: t12
## ARIMA(1,1,2)
##
## Coefficients:
##          ar1          ma1          ma2
##          0.6265   -0.3053   -0.0517
## s.e.    0.0650    0.0684    0.0301
##
## sigma^2 estimated as 0.03168:  log likelihood=751.7
## AIC=-1495.4   AICc=-1495.38   BIC=-1472.19
##
## Training set error measures:
##              ME      RMSE      MAE      MPE      MAPE
## Training set -0.0004550749 0.1778408 0.1051573 -0.03434224 1.758331
##              MASE      ACF1
## Training set 0.937794 -0.0004524371

# the best fit in r is arima(1,1,2): that is ar order = 1, level of differencing = 1, ma order = 2
plot(t12,type='l',xlim=c(500,600),lwd=2)
lines(ar2ma1_q1$fitted.values,type='l',col='green',lwd=2)
lines(auto_p1$fitted,type='l',col='red',lwd=2)

```

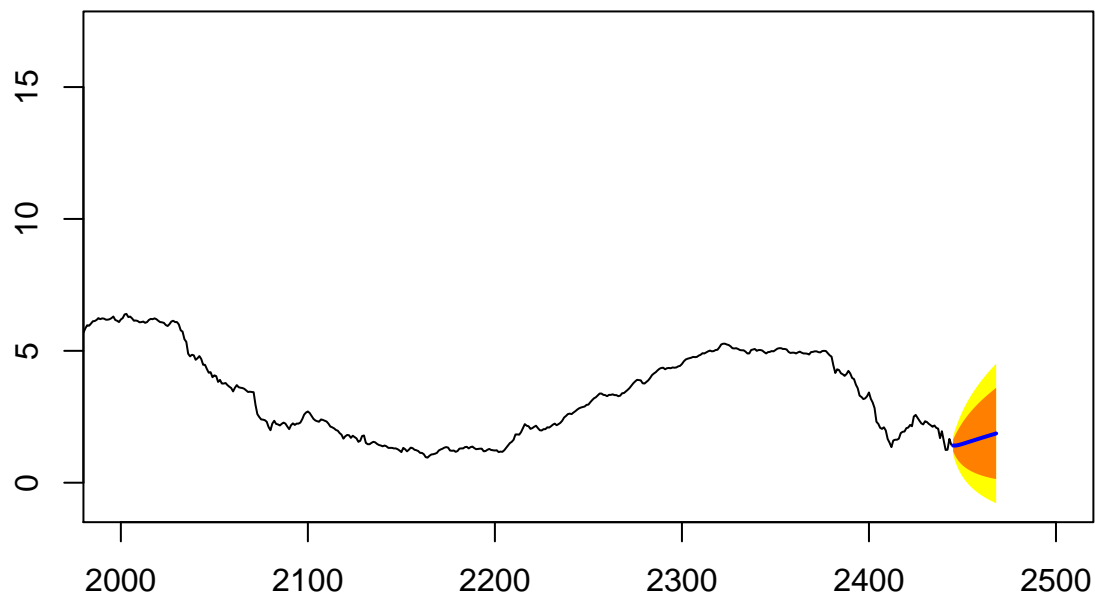


```

# these two simulations coincide each other.
## g
f_q1 = Arima(t12,order=c(2,0,1))
plot(forecast(f_q1,h=24),shadecols="oldstyle",xlim=c(2000,2500))

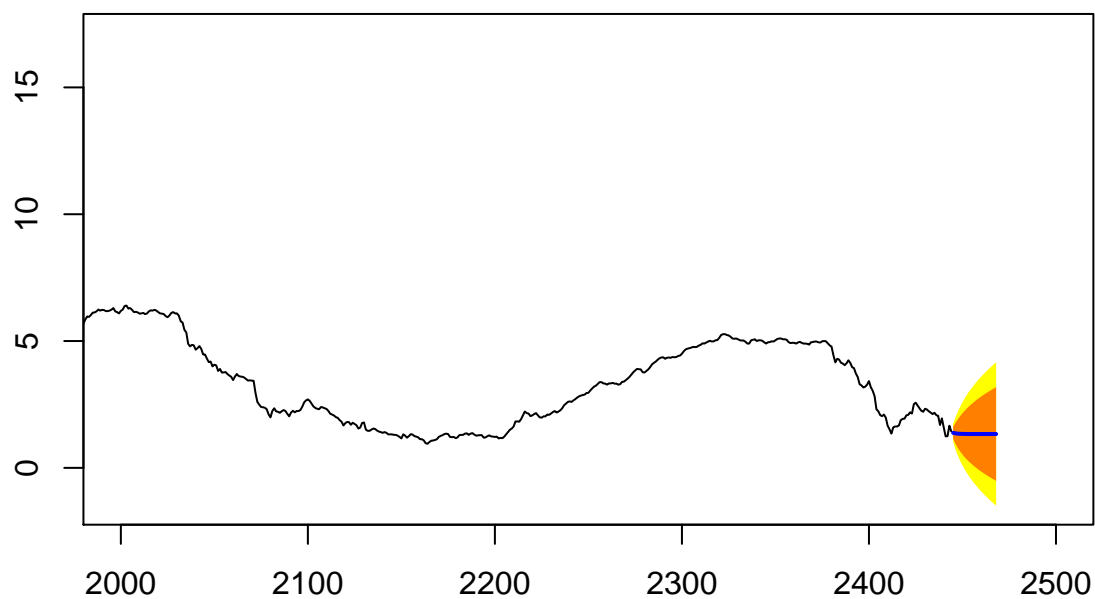
```


Forecasts from ARIMA(2,0,1) with non-zero mean



```
plot(forecast(auto_p1,h=24),shadecols="oldstyle",xlim=c(2000,2500))
```

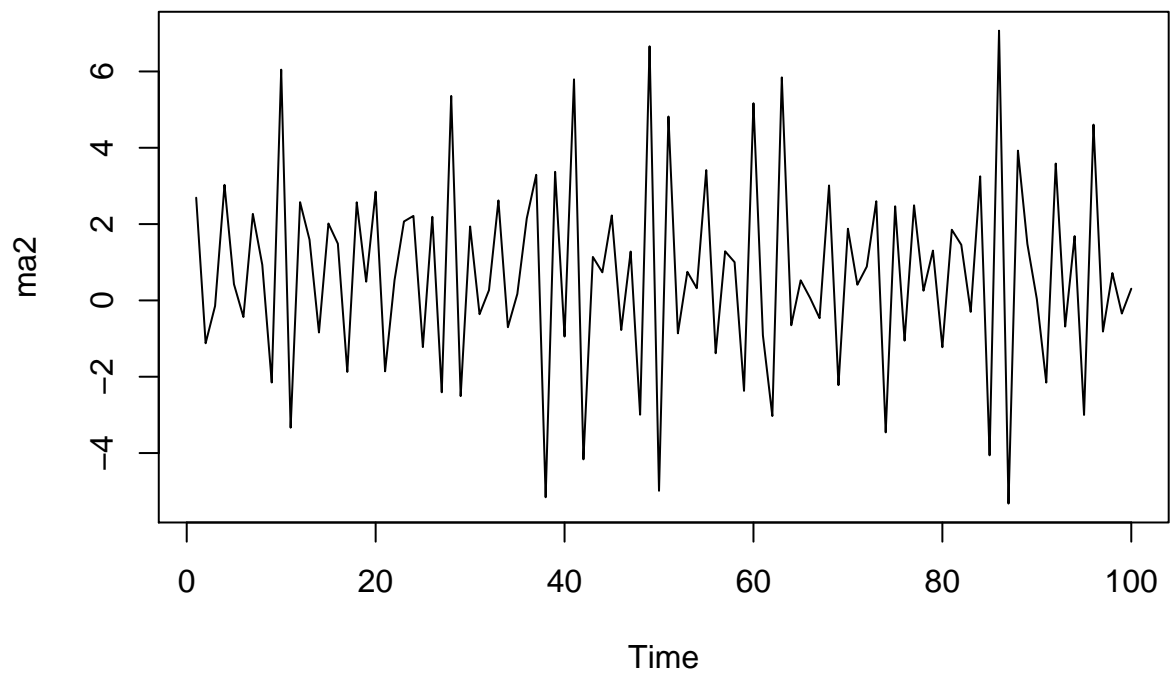
Forecasts from ARIMA(1,1,2)



arima(2,0,1) gives a upward trend approxiamtion, while arima(1,1,2) generated by autoarima gives a re

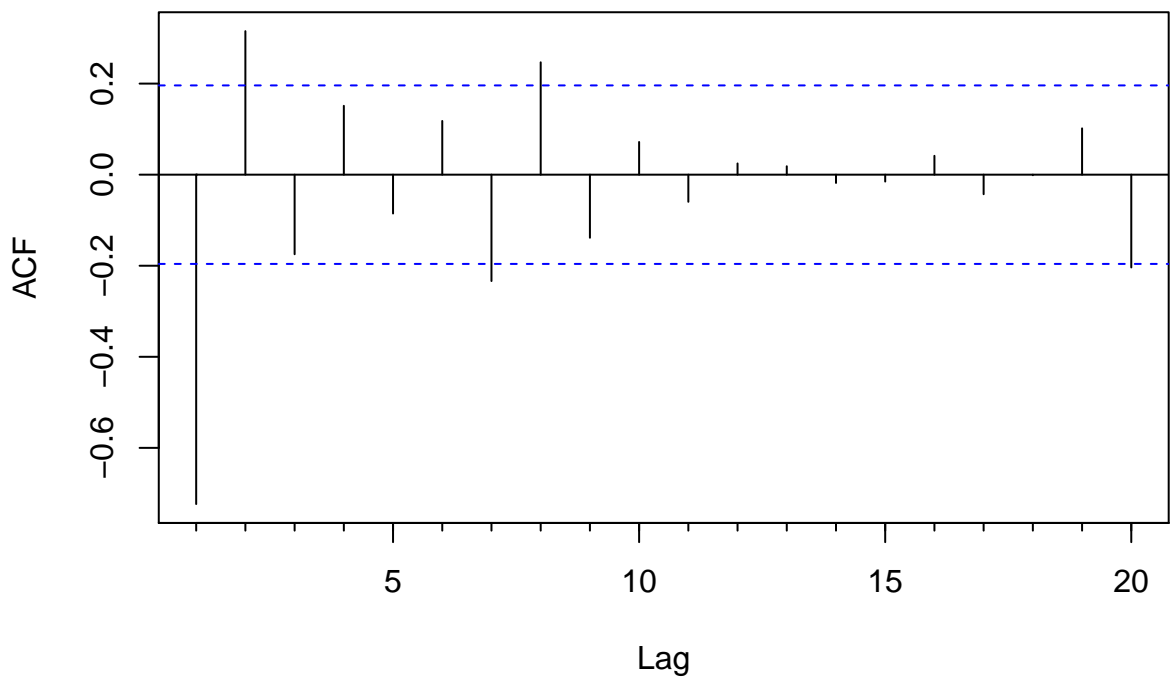
6.5 from book a

```
# generate data set in 6.4
ma2 <- arima.sim(model=list(ma=c(-2,1.35)),n=100)+0.7
plot(ma2)
```



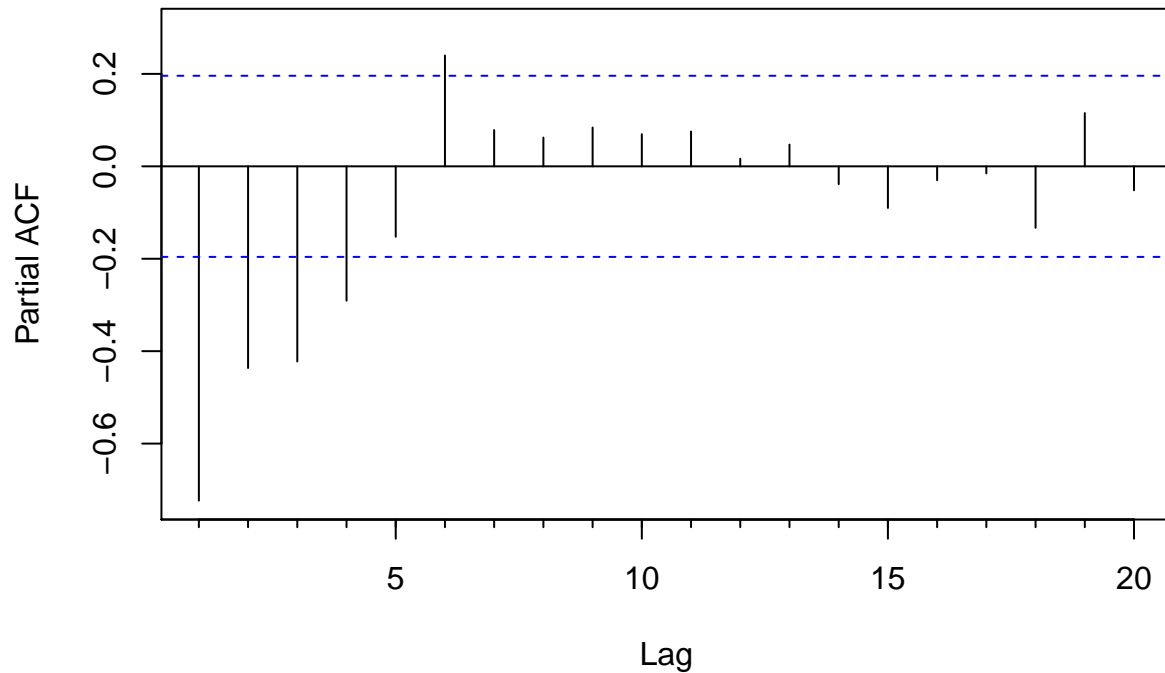
`Acf(ma2)`

Series ma2



`Pacf(ma2)`

Series ma2

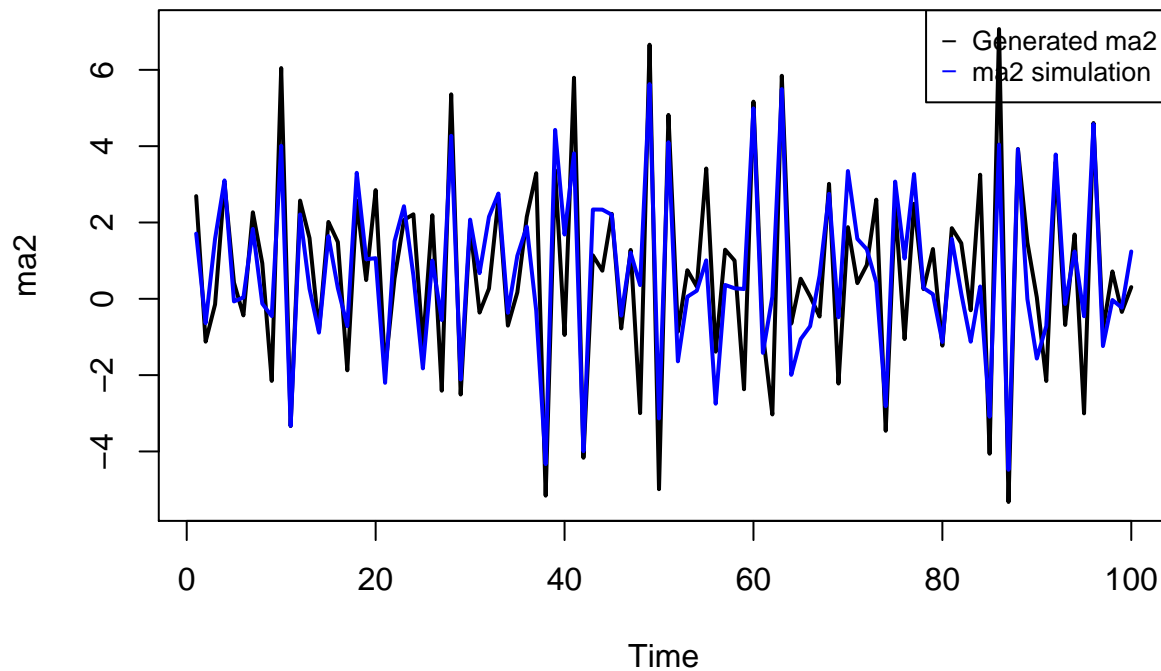


```
## a
ma2_summary=arma(ma2,order=c(0,2))
summary(ma2_summary)

##
## Call:
## arma(x = ma2, order = c(0, 2))
##
## Model:
## ARMA(0,2)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.2617417 -0.8255906  0.0009851  0.9141638  3.8073091
##
## Coefficient(s):
##              Estimate Std. Error t value Pr(>|t|)
## ma1          -1.44618    0.08253  -17.52  <2e-16 ***
## ma2           0.71614    0.07547   9.49   <2e-16 ***
## intercept     0.69900    0.03677  19.01  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 1.892,  Conditional Sum-of-Squares = 183.55,  AIC = 353.56
# We can see that ma1, ma2 and intercept are all of significant, so this is a ma2 process.

ma2_sim <- Arima(ma2,order=c(0,0,2),include.drift=TRUE)
plot(ma2,lwd=2)
```

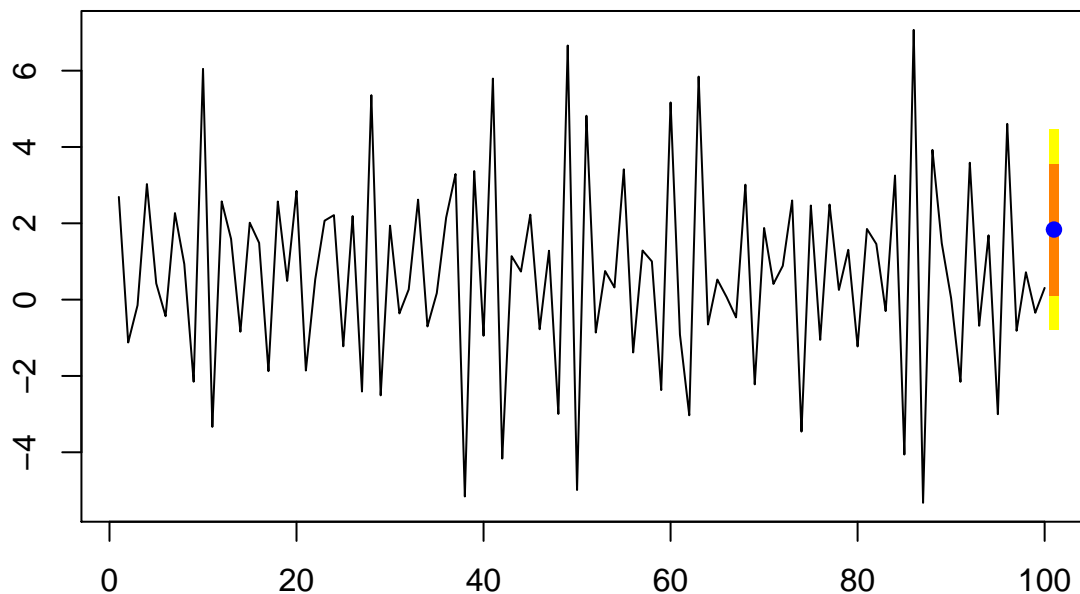
```
lines(ma2_sim$fitted,col="blue",lwd=2)
legend("topright",pch= c("-", "-"),legend=c("Generated ma2", "ma2 simulation"),col=c("black", "blue"),
```



From the graph we can see that simulated ma2 model is very close to the theoretical model, but the ma.

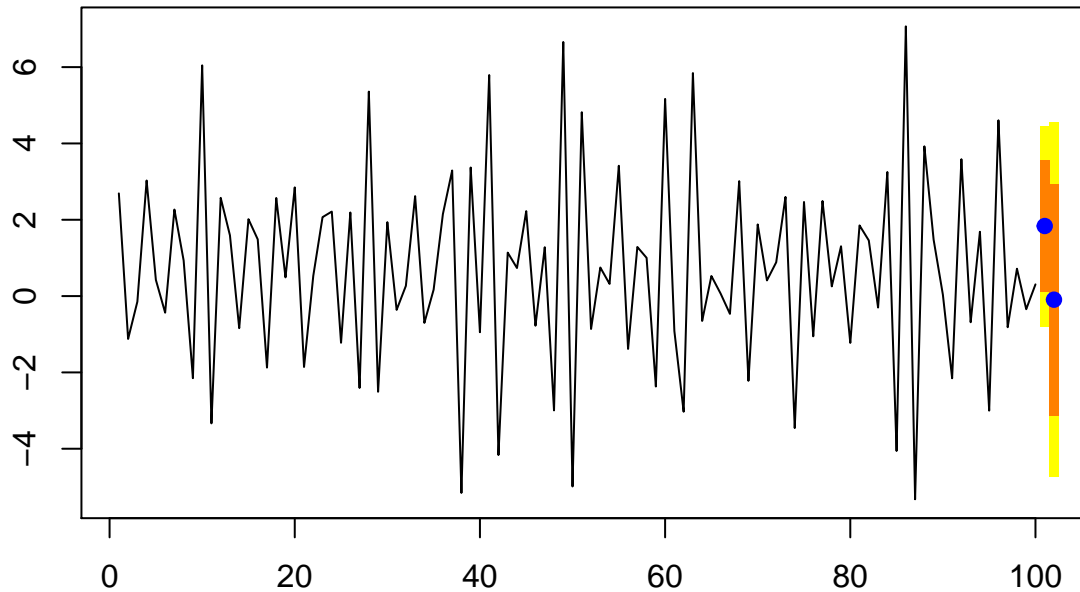
```
## b
# 1-step ahead
plot(forecast(ma2_sim,h=1),shadecols="oldstyle")
```

Forecasts from ARIMA(0,0,2) with drift



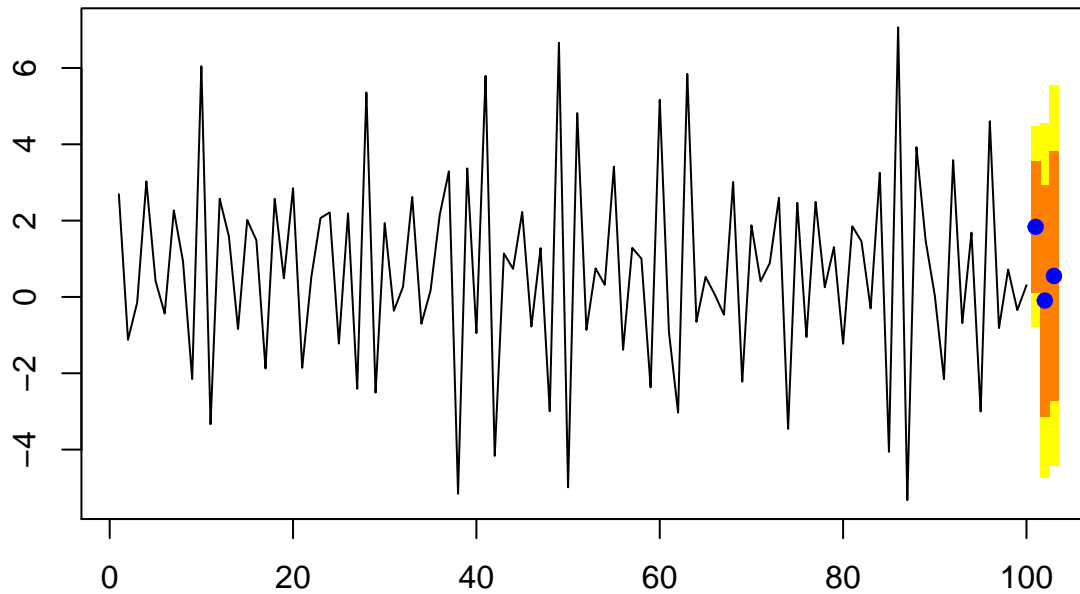
```
# 2-step ahead
plot(forecast(ma2_sim,h=2),shadecols="oldstyle")
```

Forecasts from ARIMA(0,0,2) with drift



```
# 3-step ahead
plot(forecast(ma2_sim,h=3),shadecols="oldstyle")
```

Forecasts from ARIMA(0,0,2) with drift



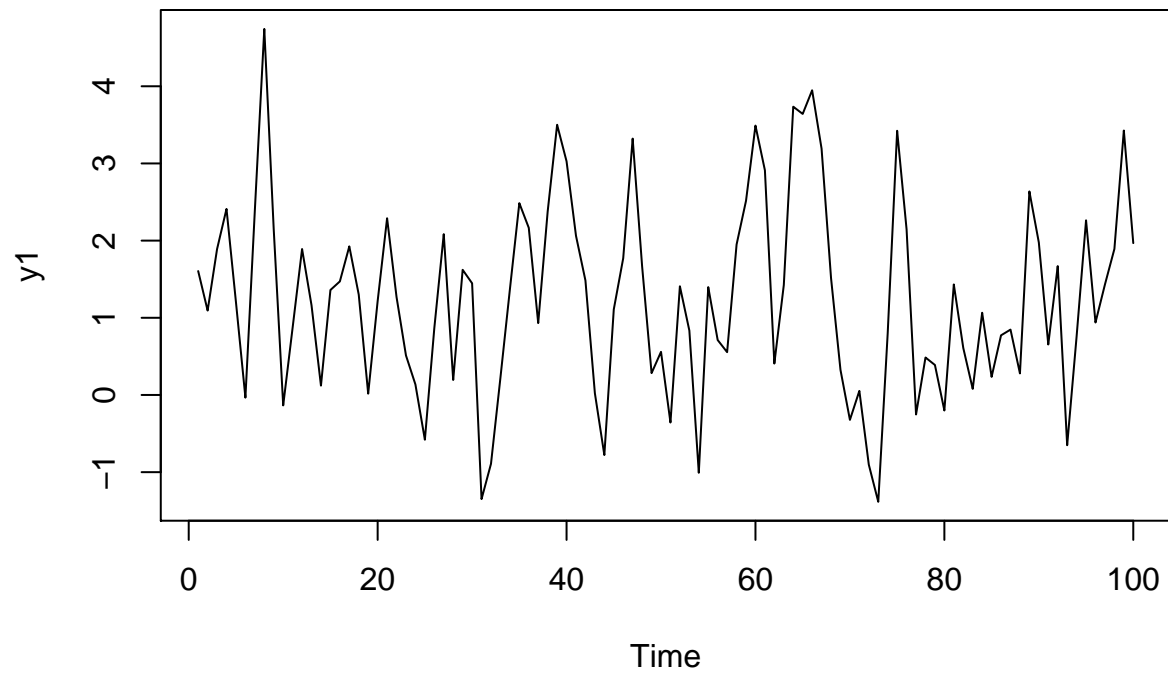
6.6 from book a

```
# first
y1 <- arima.sim(model = list(ma=c(0.8)),n=100)+1.2
# autoregressive:
# 
$$Y_t = 1.2 + 0.8*(Y_{t-1} - 1.2) - 0.8^2*(Y_{t-2} - 1.2) + \dots + (-1)^t * 0.8^{(t-1)} * (Y_1 - 1.2) + e_t$$

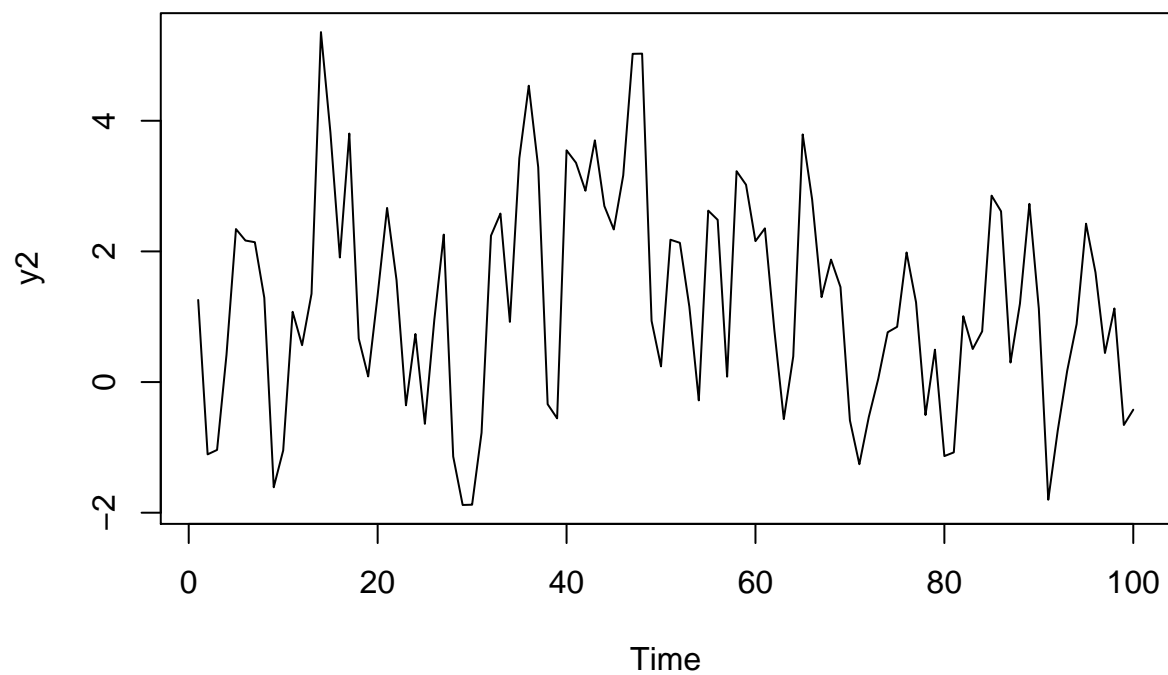
```

```
# second
y2 <- arima.sim(model = list(ma=c(1.25)),n=100)+1.2
# autoregressive:
#  $-1/0.8 * Y_{t+1} = 1/0.8^2 * Y_{t+2} + 1/0.8^3 * Y_{t+3} + \dots + e_t$ 

# plot
plot(y1)
```

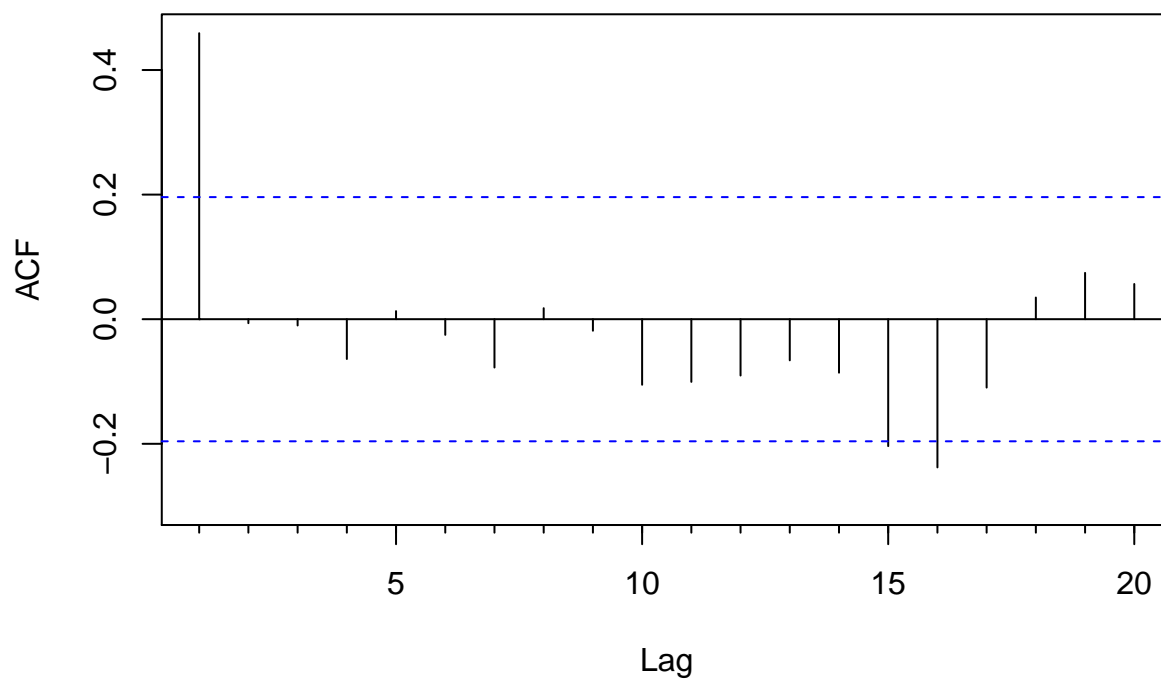


```
plot(y2)
```



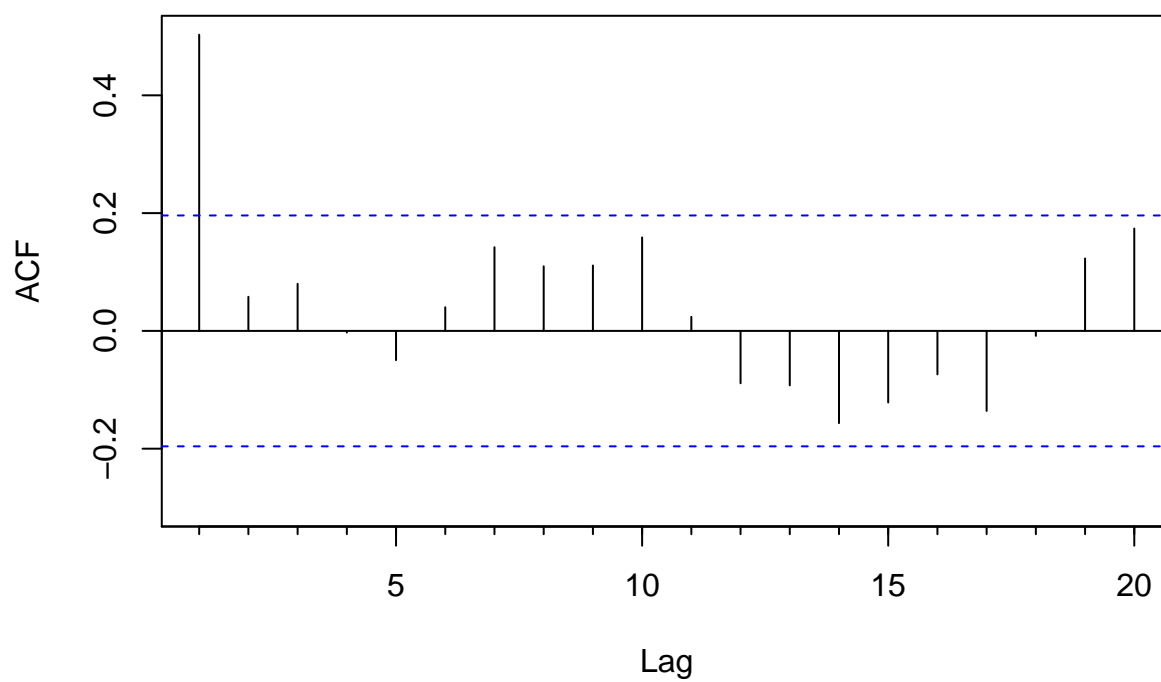
Acf(y1)

Series y1



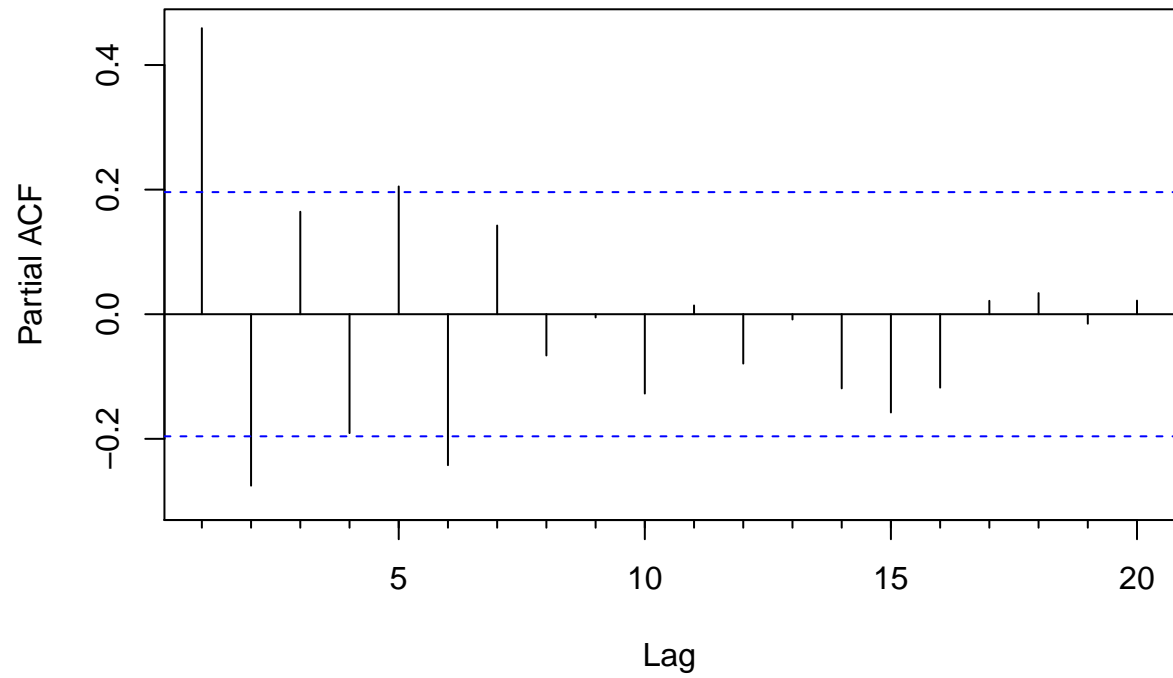
Acf(y2)

Series y2



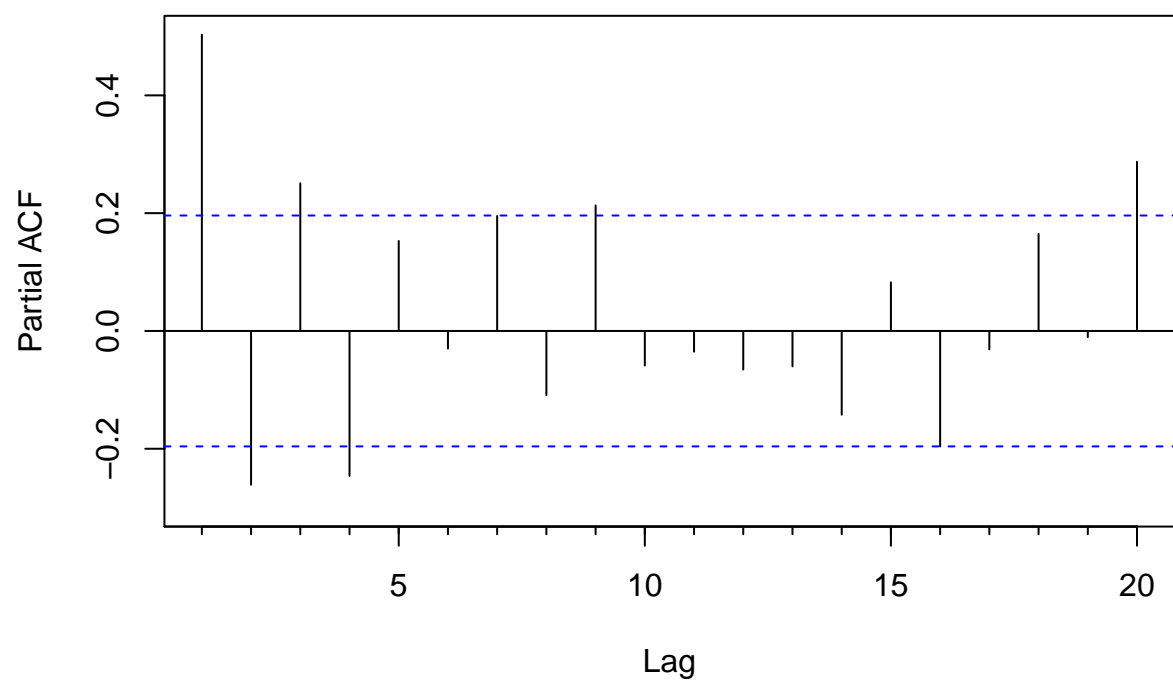
Pacf(y1)

Series y1



Pacf(y2)

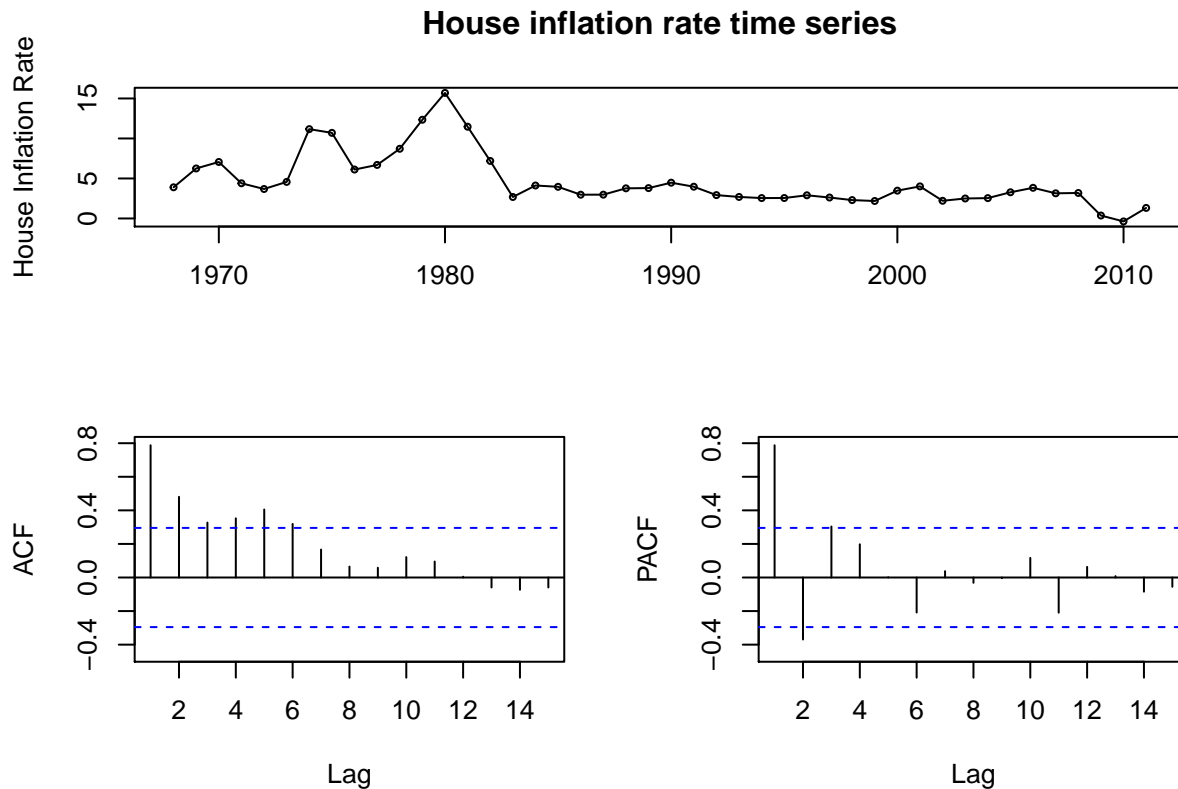
Series y2



I prefer the first model. Because its invertibility, we can convert MA(1) process into its own past,

7.6 from book a

```
hw31 <- read.xls('hw37a.xls',sheet = 3)
# delete na
hw31 <- hw31[-c(1),]
# house inflation
hinf <-hw31$housing.Inflation...
thinft<-ts(hinf,start=1968,2011,freq = 1)
tsdisplay(thinf,ylab="House Inflation Rate",main="House inflation rate time series")
```



```
# from acf and pacf, housing inflation rate can be modeled as ar2
ar21 <- arma(hinf,order=c(2,0)) #Same as MA(1) = AR(0) + MA(1)
summary(ar21)
```

```
##
## Call:
## arma(x = hinf, order = c(2, 0))
##
## Model:
## ARMA(2,0)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.1746 -0.8164 -0.3451  0.4210  6.3067
##
```

```
## Coefficient(s):
##           Estimate Std. Error t value Pr(>|t|)
## ar1         1.0997    0.1382   7.958 1.78e-15 ***
## ar2        -0.3749    0.1422  -2.636  0.00838 **
## intercept    1.2157    0.5125   2.372  0.01769 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 3.493, Conditional Sum-of-Squares = 143.22, AIC = 185.9
```

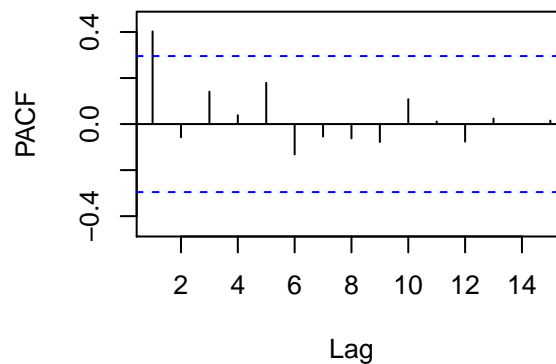
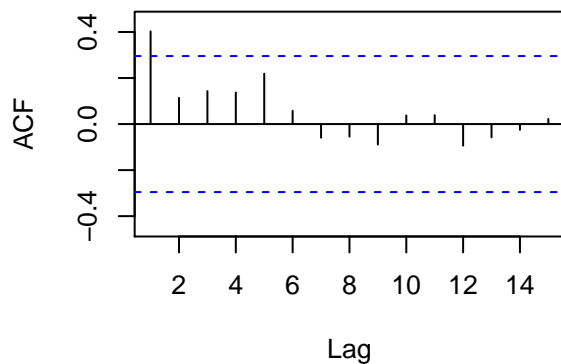
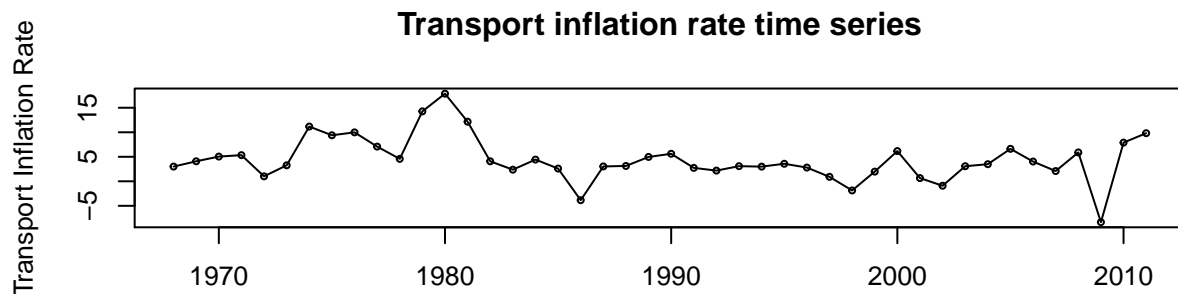
from the t test, all coefficients are of significances. Hence house inflation rate can be modeled as

transportation inflation

```
tinf <-hw31$transportation.Inflation....
```

```
ttinf<-ts(tinf,start=1968,2011,freq = 1)
```

```
tsdisplay(ttinf,ylab="Transport Inflation Rate",main="Transport inflation rate time series")
```



from acf and pacf, transport inflation rate can not be modeled as ar2

```
ar22 <- arma(tinf,order=c(2,0))
```

```
summary(ar22)
```

```
##
## Call:
## arma(x = tinf, order = c(2, 0))
##
## Model:
## ARMA(2,0)
##
## Residuals:
```

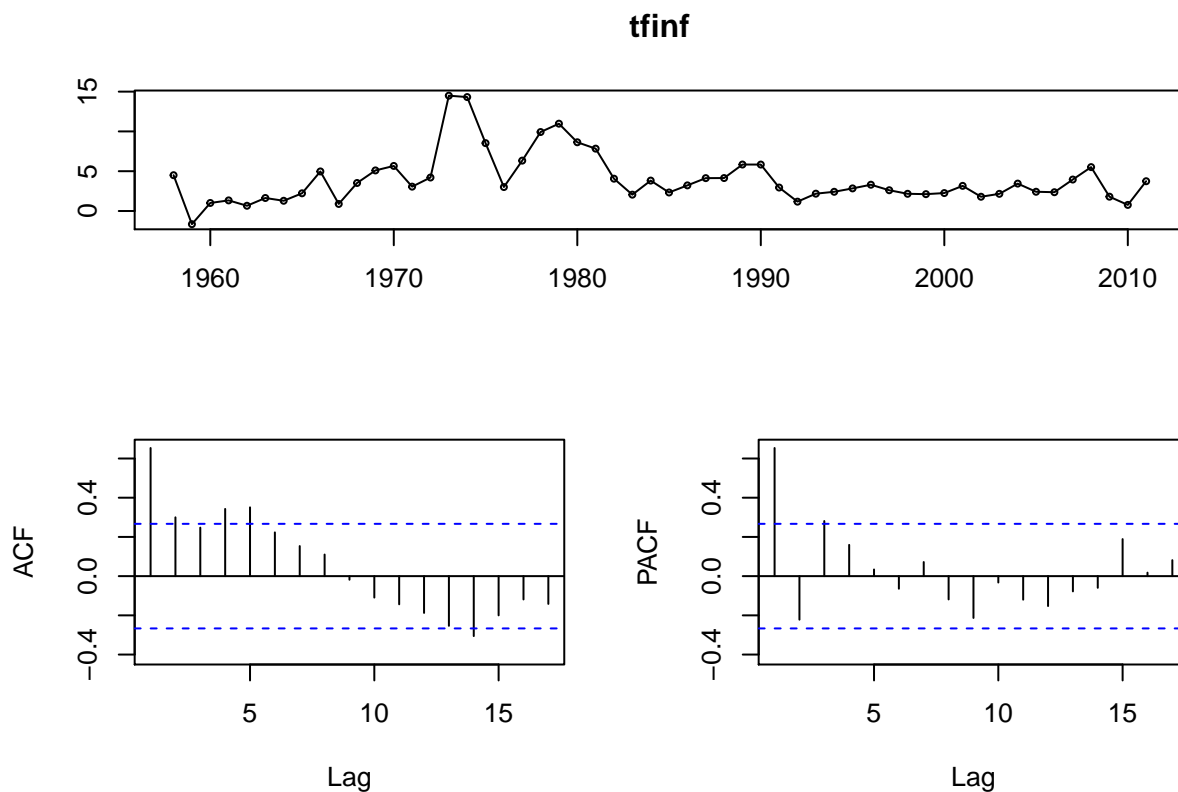
```
##      Min      1Q   Median      3Q      Max
## -13.5817 -1.8746 -0.1395   2.0365   9.8649
##
## Coefficient(s):
##           Estimate Std. Error t value Pr(>|t|)
## ar1           0.4391    0.1522   2.886  0.00391 **
## ar2          -0.0553    0.1532  -0.361  0.71820
## intercept     2.7798    0.9489   2.930  0.00339 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 17.78,  Conditional Sum-of-Squares = 728.81,  AIC = 257.49
```

from the t test, coefficient of ar2 is not of significance. Hence transportation inflation rate can n

7.7 from book a

```
hw32<-read.xls('hw37a.xls',sheet = 4)
hw32<-hw32[-c(1),] # delete na row

# food inflation
finf <- hw32$food.Inflation...
tfinf <- ts(finf,1958,2011,freq=1)
tsdisplay(tfinf)
```



```
ffit <- auto.arima(finf)
summary(ffit) # ar(1)+ma(1) process
```

```

## Series: finf
## ARIMA(1,0,1) with non-zero mean
##
## Coefficients:
##          ar1      ma1      mean
##          0.3709  0.5361  4.0364
## s.e.    0.1677  0.1632  0.7306
##
## sigma^2 estimated as 5.287:  log likelihood=-120.47
## AIC=248.94  AICc=249.75  BIC=256.89
##
## Training set error measures:
##              ME      RMSE      MAE      MPE      MAPE      MASE
## Training set -0.03008777  2.234634  1.526076 -19.88114  60.99807  0.8351221
##              ACF1
## Training set -0.01990706

ffit2 <- arma(finf,order=c(1,0))# try ar1
summary(ffit2) # compare aic and find ffit works better

##
## Call:
## arma(x = finf, order = c(1, 0))
##
## Model:
## ARMA(1,0)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -5.94928 -1.52663  0.02819  0.97395 10.38545
##
## Coefficient(s):
##           Estimate Std. Error t value Pr(>|t|)
## ar1           0.6537     0.1029   6.35 2.16e-10 ***
## intercept     1.3533     0.5205   2.60 0.00932 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 5.824,  Conditional Sum-of-Squares = 302.85,  AIC = 252.39

ffitr = Arima(finf,order=c(1,0,1),include.drift=TRUE) # (1,1) fit
# forecast
p1 <-forecast(ffit,h=1) # 1 step ahead
p2 <-forecast(ffit,h=2) # 2 step ahead
p3 <-forecast(ffit,h=3) # 3 step ahead
# forecast error
e1 <- recresid(ffitr$res~1)
e1

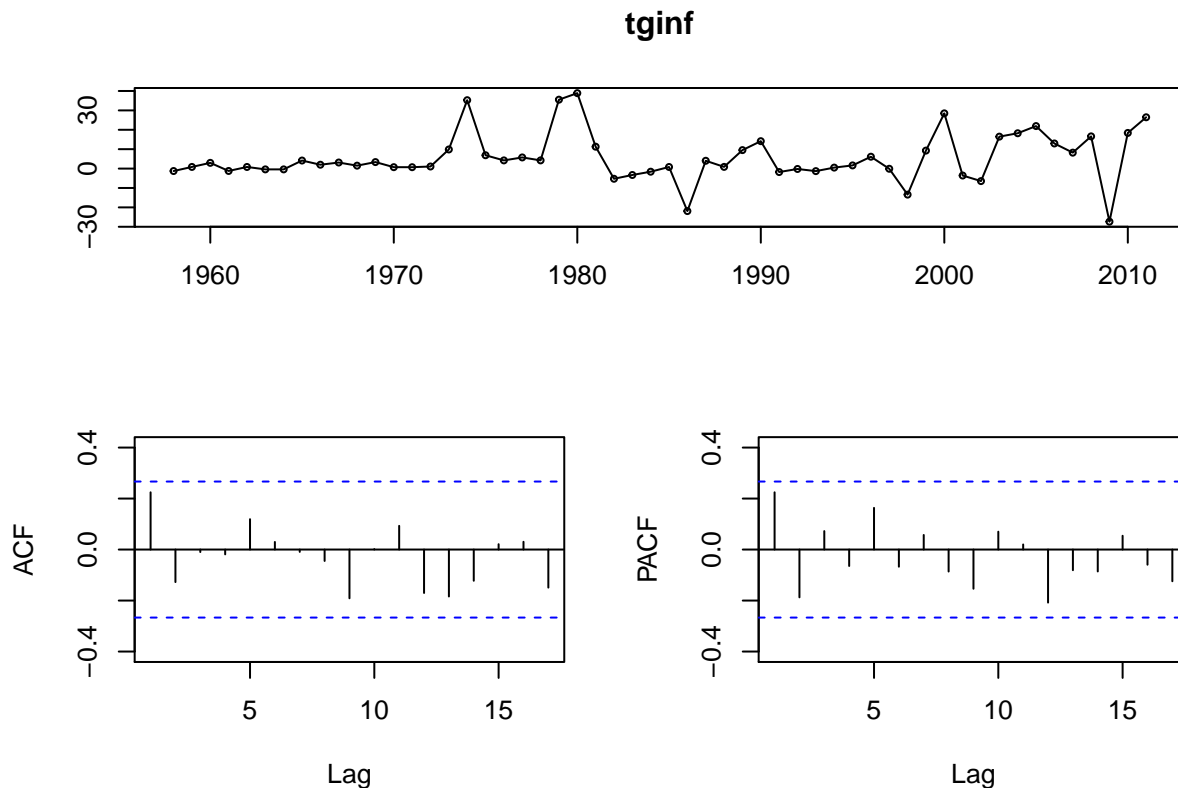
## [1] -3.972323338  3.702457647 -1.144104112  0.502066183  0.773833165
## [6] -0.175201771  1.312931889  2.645990676 -3.236547510  4.055600800
## [11]  0.517795012  2.225096852 -1.574164909  2.541114018  9.769020328
## [16]  1.421478091  0.018057585 -2.597697955  4.143908148  2.706706459
## [21]  3.035665545  0.018102093  1.639172565 -2.708150057 -0.895413848
## [26]  0.671203509 -2.262647667  0.823427494 -0.243383744  0.039118185

```

```
## [31]  1.572105276  0.105939084 -1.954407156 -1.485531651 -0.035306991
## [36] -0.902090064 -0.049714685 -0.177665709 -0.931889140 -0.675130900
## [41] -0.654254898 -0.469670199  0.311457385 -1.769188914  0.262175606
## [46]  0.335786090 -1.170368003 -0.004709248  0.990420844  1.428449276
## [51] -3.065797076 -0.262159508  1.582856512
```

```
# forecast uncertainty
# sigma^2 is 5.353
```

```
# gas inflation
ginf <- hw32$Gas.Inflation....
tginf <-ts(ginf,1958,2011,freq=1)
tsdisplay(tginf) # white noise
```



```
gfit <- auto.arima(ginf)
summary(gfit) # ma1 process
```

```
## Series: ginf
## ARIMA(0,0,1) with non-zero mean
##
## Coefficients:
##      ma1      mean
##      0.3389  5.5646
## s.e.  0.1523  2.1390
##
## sigma^2 estimated as 144.5: log likelihood=-209.94
## AIC=425.87  AICc=426.35  BIC=431.84
##
## Training set error measures:
##              ME      RMSE      MAE      MPE      MAPE      MASE
```

```
## Training set 0.02835135 11.7948 8.357421 190.3038 364.341 0.8604053
##          ACF1
## Training set -0.0401626

gfit2 <- arma(ginf,order=c(1,1)) # try ar1 +ma1 process
summary(gfit2) # compare aic and find gfit works better

##
## Call:
## arma(x = ginf, order = c(1, 1))
##
## Model:
## ARMA(1,1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.9360  -5.5391  -0.9959   6.3187  29.8594
##
## Coefficient(s):
##              Estimate Std. Error t value Pr(>|t|)
## ar1            -0.59545    0.11332   -5.255 1.48e-07 ***
## ma1             0.95597    0.07325   13.051 < 2e-16 ***
## intercept      6.88947    3.60509    1.911  0.056 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Fit:
## sigma^2 estimated as 138, Conditional Sum-of-Squares = 7224.22, AIC = 425.3

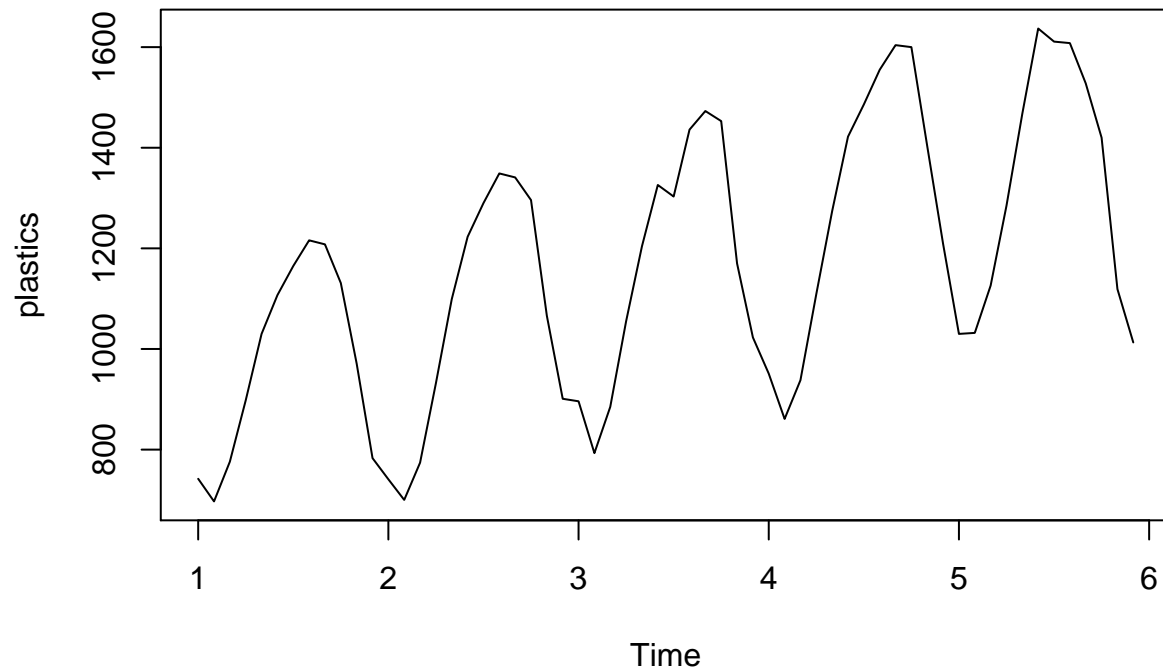
# forecast
p21 <-forecast(ffit,h=1) # 1 step ahead
p22 <-forecast(ffit,h=2) # 2 step ahead
p23 <-forecast(ffit,h=3) # 3 step ahead
# forecast error
e2 <- recresid(gfit$residuals~1)
e2

## [1] 2.71568017 2.30067183 -2.25563173 1.45331976 -1.05337720
## [6] -0.12202155 3.89922379 0.03699689 2.26783568 -0.24467771
## [11] 2.32249722 -1.26208177 -0.01841867 -0.06897297 8.46545398
## [16] 29.69473228 -6.96875651 2.68662591 0.89333956 -0.07131925
## [21] 30.82080503 22.26609895 -3.27066239 -10.99910310 -5.99307033
## [26] -5.65211036 -3.10150329 -26.11722964 8.10582709 -6.56385706
## [31] 7.04748593 6.81108939 -9.05646988 -1.92580823 -5.29580013
## [36] -2.18981805 -2.03234485 2.41192094 -5.26631446 -15.71789008
## [41] 10.67682186 20.57487535 -15.02053551 -5.62080169 14.09784509
## [46] 8.87762588 14.02580085 3.04726326 1.92556374 10.55040191
## [51] -36.12085758 25.61519593 12.44083323

# forecast uncertainty
# sigma^2 is 144.5
```

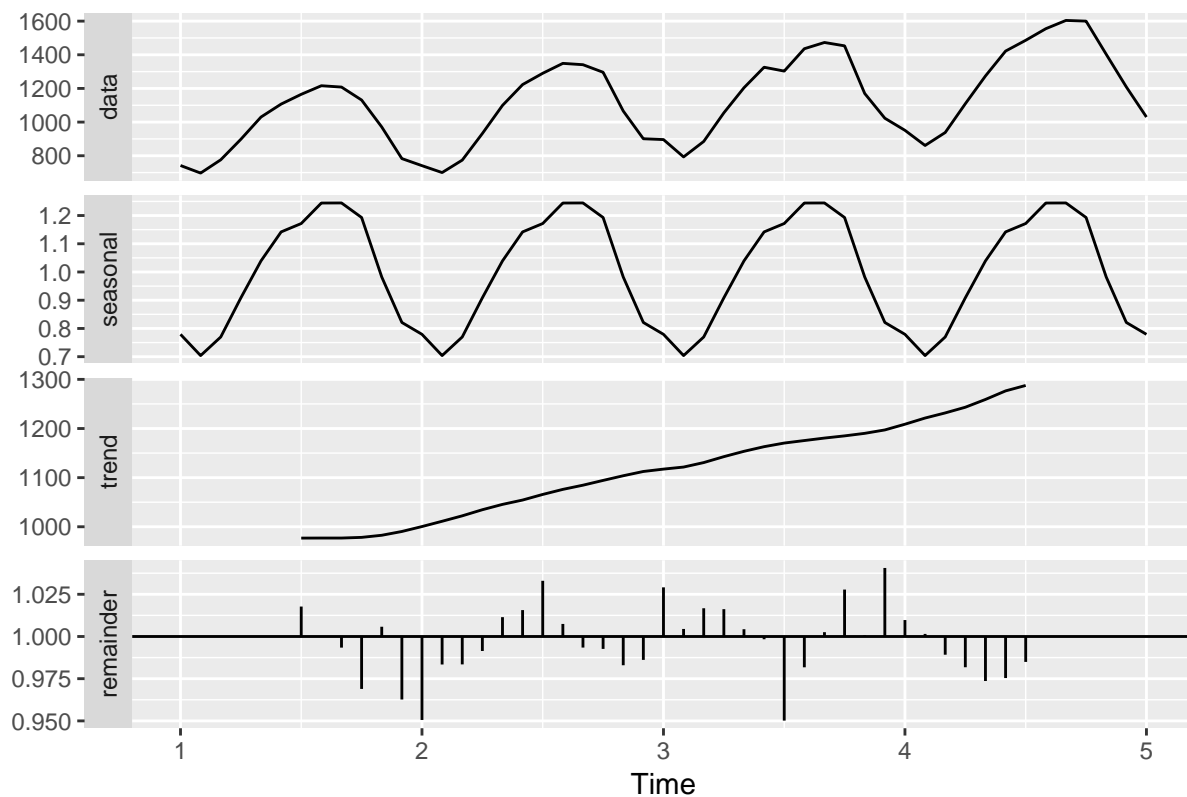
6.2 from book c

```
## a  
plot(plastics)
```



```
# There is a seasonal cycle in between each unit time. A peak appears in the middle of the unit time. T  
## b  
plasticsts<-ts(plastics,1,5,frequency = 12)  
plasticsts %>% decompose(type="multiplicative") %>% autoplot()
```

Decomposition of multiplicative time series



```
splastics<-decompose(plasticts,"multiplicative")
stlplastic <- stl(plasticts, s.window = "periodic")
# trend is in a linear growth
stlplastic # seasonal indices
```

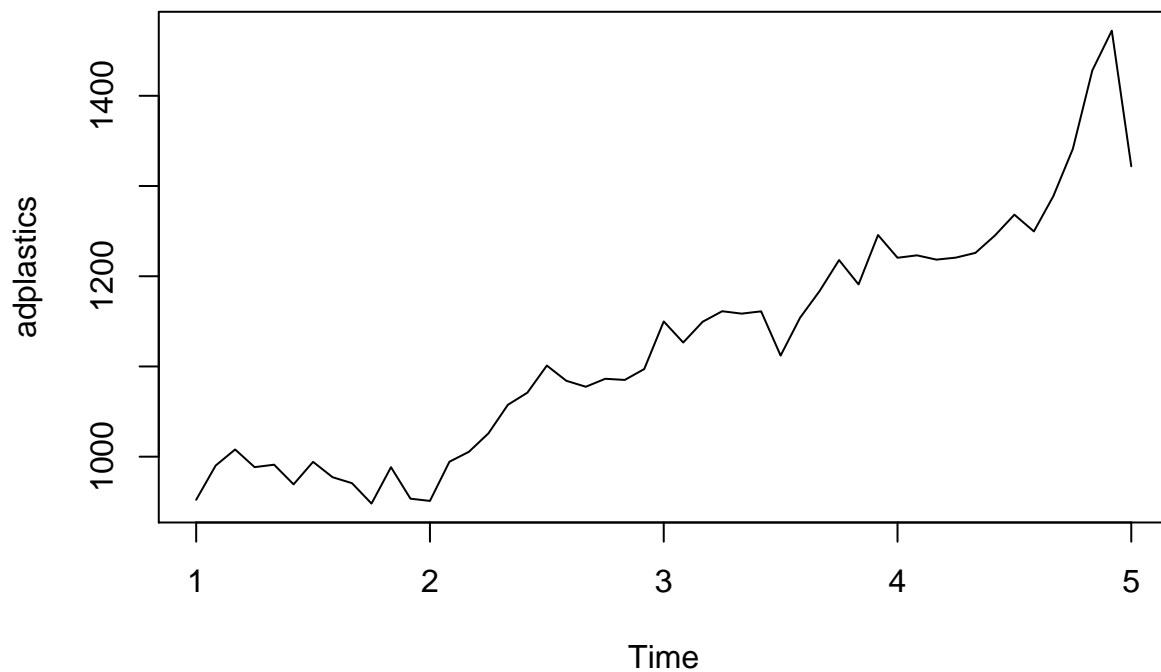
```
## Call:
## stl(x = plasticts, s.window = "periodic")
##
## Components
```

	seasonal	trend	remainder
## Jan 1	-267.401536	1008.6525	0.7489974
## Feb 1	-328.779407	1003.4665	22.3129389
## Mar 1	-254.910287	998.2804	32.6298894
## Apr 1	-107.081843	994.0930	10.9888880
## May 1	38.746597	989.9055	1.3478902
## Jun 1	150.562280	986.7182	-30.2805216
## Jul 1	186.127941	983.5310	-4.6589110
## Aug 1	259.038125	980.8859	-23.9240554
## Sep 1	271.448290	978.2409	-41.6891809
## Oct 1	228.114417	979.6957	-76.8101479
## Nov 1	3.780506	981.1506	-13.9310760
## Dec 1	-179.644999	990.3793	-27.7343497
## Jan 2	-267.401536	999.6081	8.7934080
## Feb 2	-328.779407	1011.8558	16.9235699
## Mar 2	-254.910287	1024.1035	4.8067409
## Apr 2	-107.081843	1035.7099	3.3719315
## May 2	38.746597	1047.3163	12.9371256


```
## Jun 2 150.562280 1056.9442 15.4935119
## Jul 2 186.127941 1066.5721 37.2999206
## Aug 2 259.038125 1075.3960 14.5659196
## Sep 2 271.448290 1084.2198 -14.6680623
## Oct 2 228.114417 1092.8688 -24.9831943
## Nov 2 3.780506 1101.5178 -39.2982873
## Dec 2 -179.644999 1109.1337 -28.4887097
## Jan 3 -267.401536 1116.7496 46.6518993
## Feb 3 -328.779407 1125.2394 -3.4600208
## Mar 3 -254.910287 1133.7292 6.1810680
## Apr 3 -107.081843 1143.4251 18.6567381
## May 3 38.746597 1153.1210 12.1324118
## Jun 3 150.562280 1161.2976 14.1401009
## Jul 3 186.127941 1169.4742 -52.6021877
## Aug 3 259.038125 1175.0520 1.9098819
## Sep 3 271.448290 1180.6297 20.9219705
## Oct 3 228.114417 1186.2888 38.5967548
## Nov 3 3.780506 1191.9479 -25.7284219
## Dec 3 -179.644999 1199.9288 2.7161509
## Jan 4 -267.401536 1207.9098 10.4917551
## Feb 4 -328.779407 1218.7223 -28.9428455
## Mar 4 -254.910287 1229.5347 -36.6244370
## Apr 4 -107.081843 1244.4991 -28.4172294
## May 4 38.746597 1259.4634 -24.2100181
## Jun 4 150.562280 1274.3761 -2.9384245
## Jul 4 186.127941 1289.2889 10.5831916
## Aug 4 259.038125 1304.0947 -8.1327889
## Sep 4 271.448290 1318.9005 13.6512496
## Oct 4 228.114417 1334.1942 37.6913688
## Nov 4 3.780506 1349.4880 49.7315270
## Dec 4 -179.644999 1365.0308 23.6141821
## Jan 5 -267.401536 1380.5737 -83.1721314
```

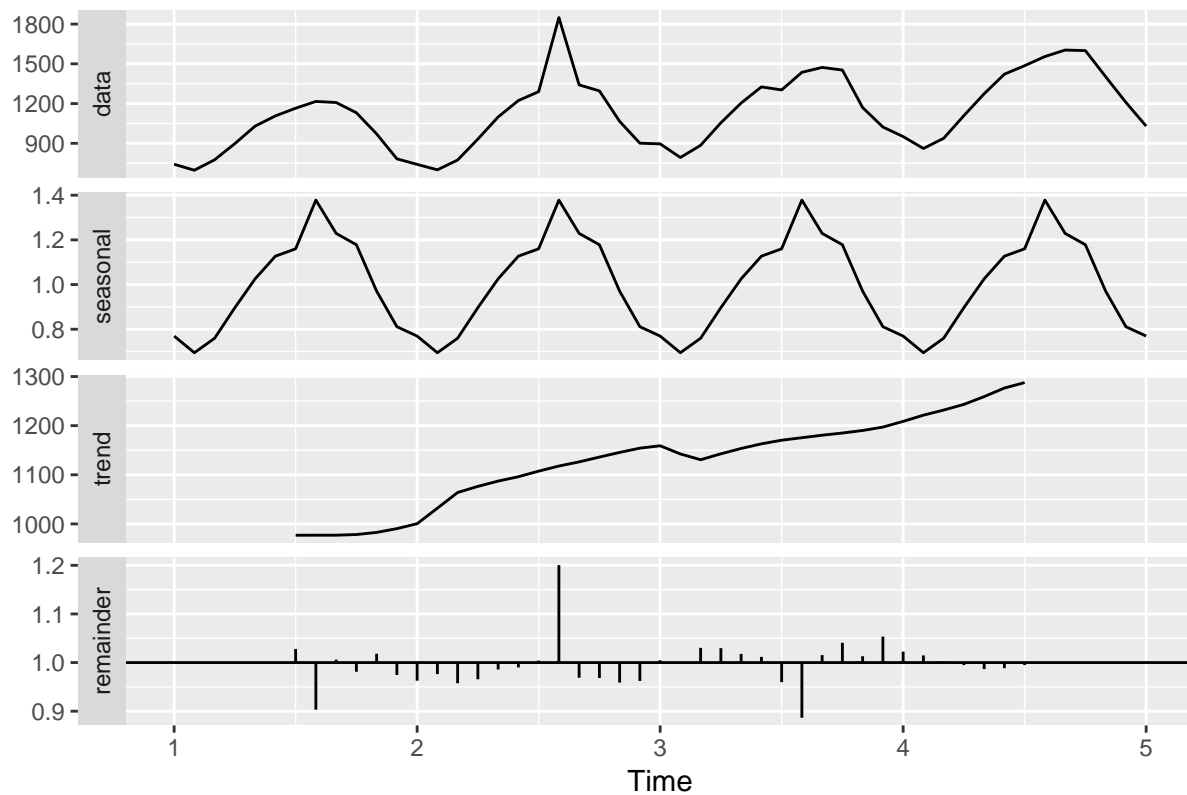
```
## c
# Yes, it supports.

## d
adplastics <- plasticts/splastics$seasonal
plot(adplastics) # adjust plastics data
```

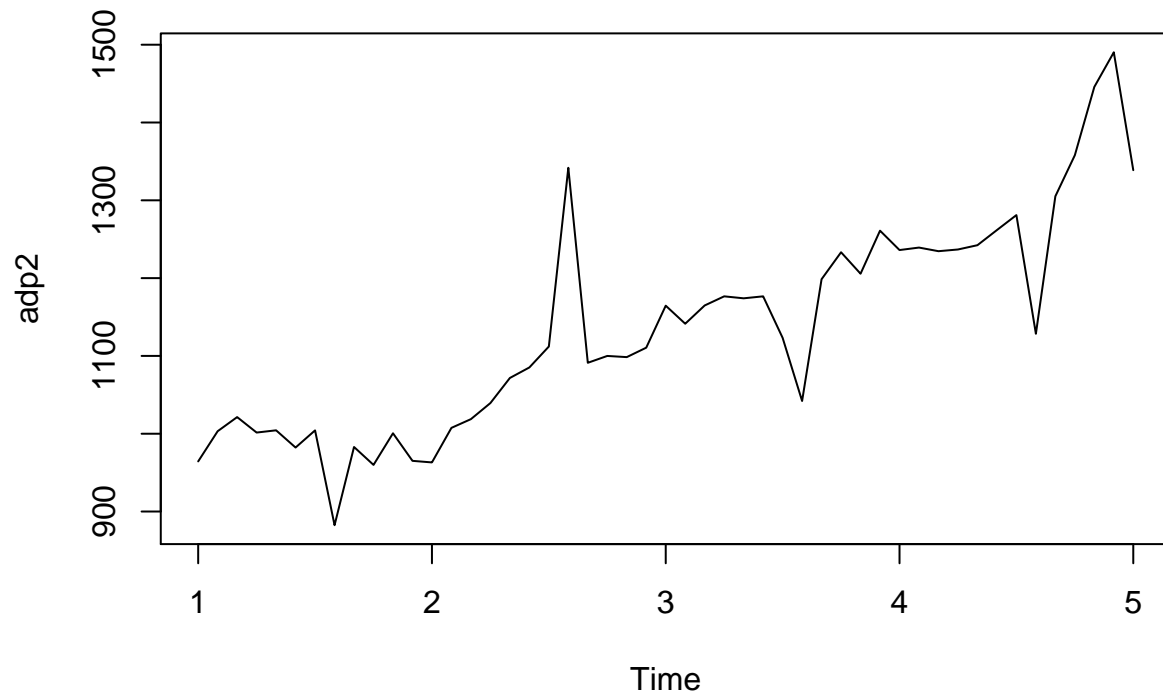


```
## e
p2<-plastics
p2[20]<-plastics[20]+500
p2ts<-ts(p2,1,5,frequency = 12)
p2ts %>% decompose(type="multiplicative") %>% autoplot()
```

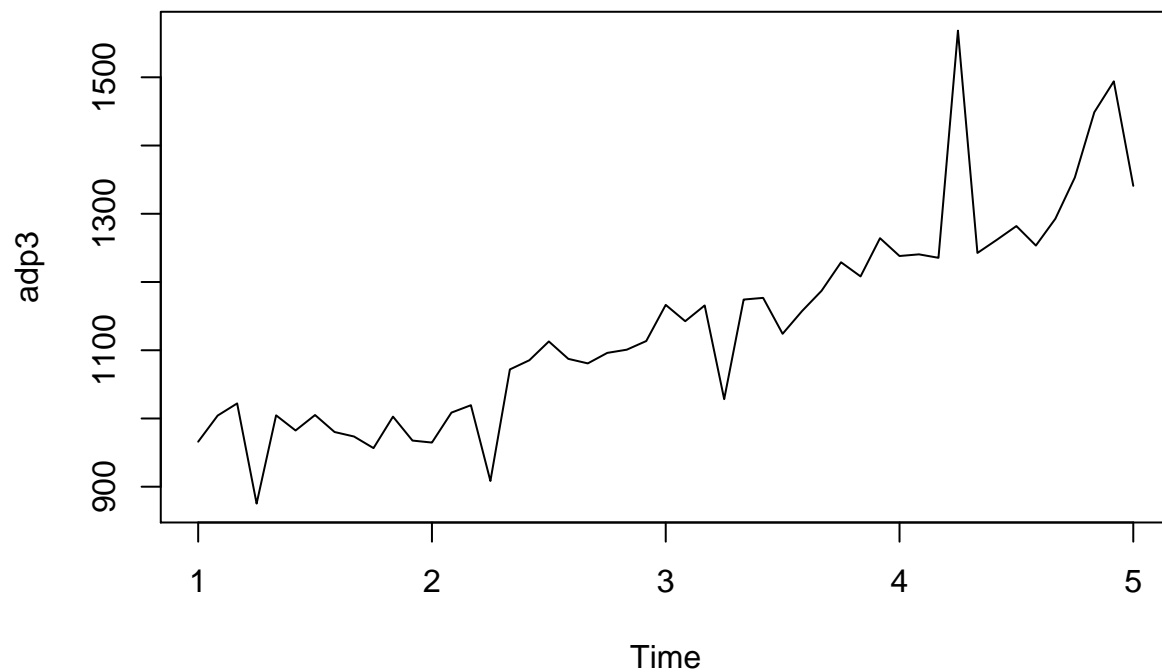
Decomposition of multiplicative time series



```
psd<-decompose(p2ts,type="multiplicative")
adp2<-seasadj(psd)
plot(adp2)
```



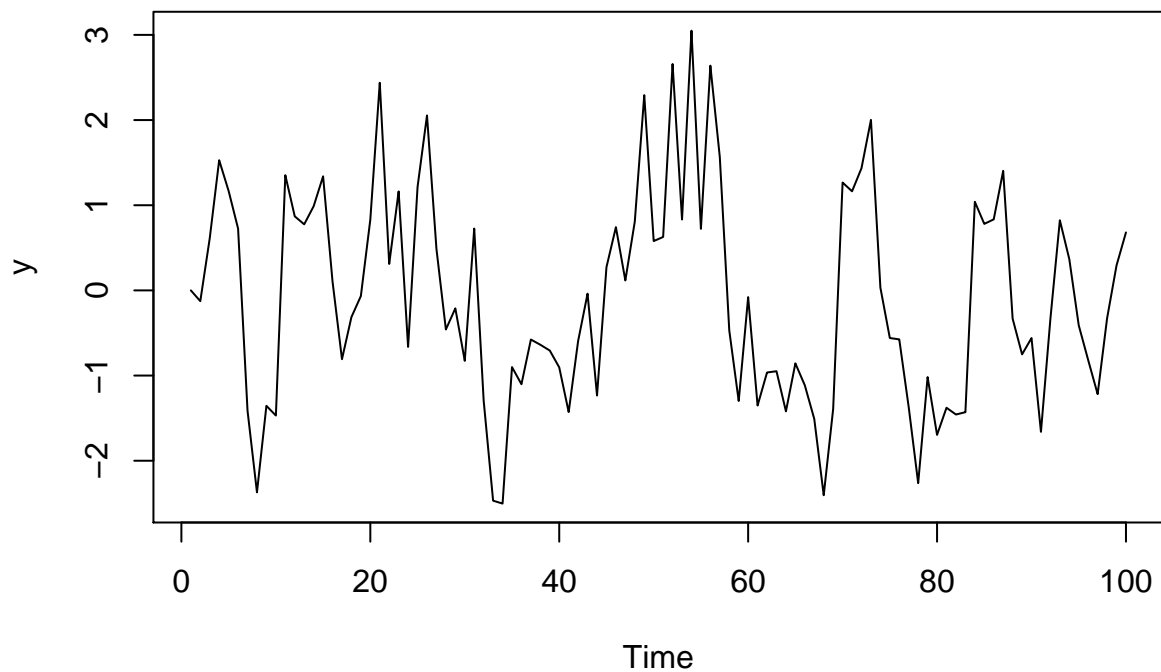
```
# In e, the new outlier drives the seasonal adjusted time series to have a spike at the point we add 500
## f
p3<-plastics
p3[40]<-plastics[40]+500
p3ts<-ts(p3,1,5,frequency = 12)
psd3 <- decompose(p3ts, type='multiplicative')
adp3 <- seasadj(psd3)
plot(adp3)
```



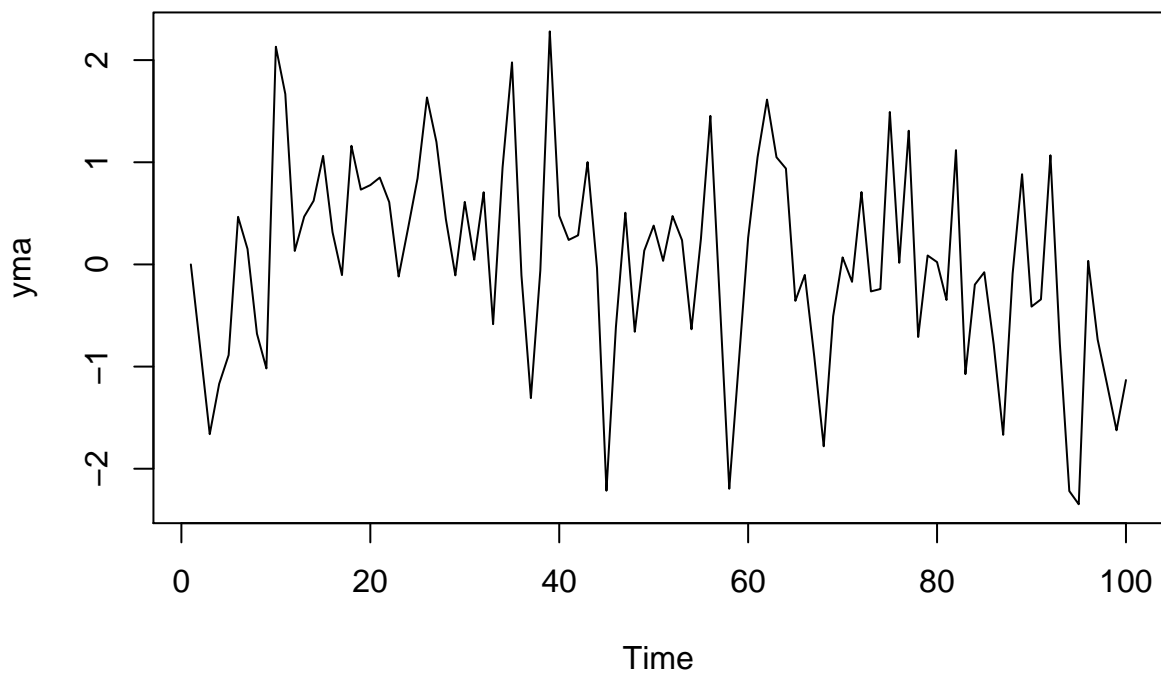
no matter where the outlier exists, it will always generate similar spike but at different location,

8.6 from book c

```
## a
# generate ar(1)
y <- ts(numeric(100))
e <- rnorm(100)
for(i in 2:100)
  y[i] <- 0.6*y[i-1] + e[i]
## b
plot(y)
```



```
# as I change the value of phi1, the graph changes. As phi1 becomes larger, the graph illustrates more pe
## c
# generate ma(1)
yma <- ts(numeric(100))
ema <- rnorm(100)
for (i in 2:100)
  yma[i] <- 0.6*ema[i-1]+ema[i]
## d
plot(yma)
```

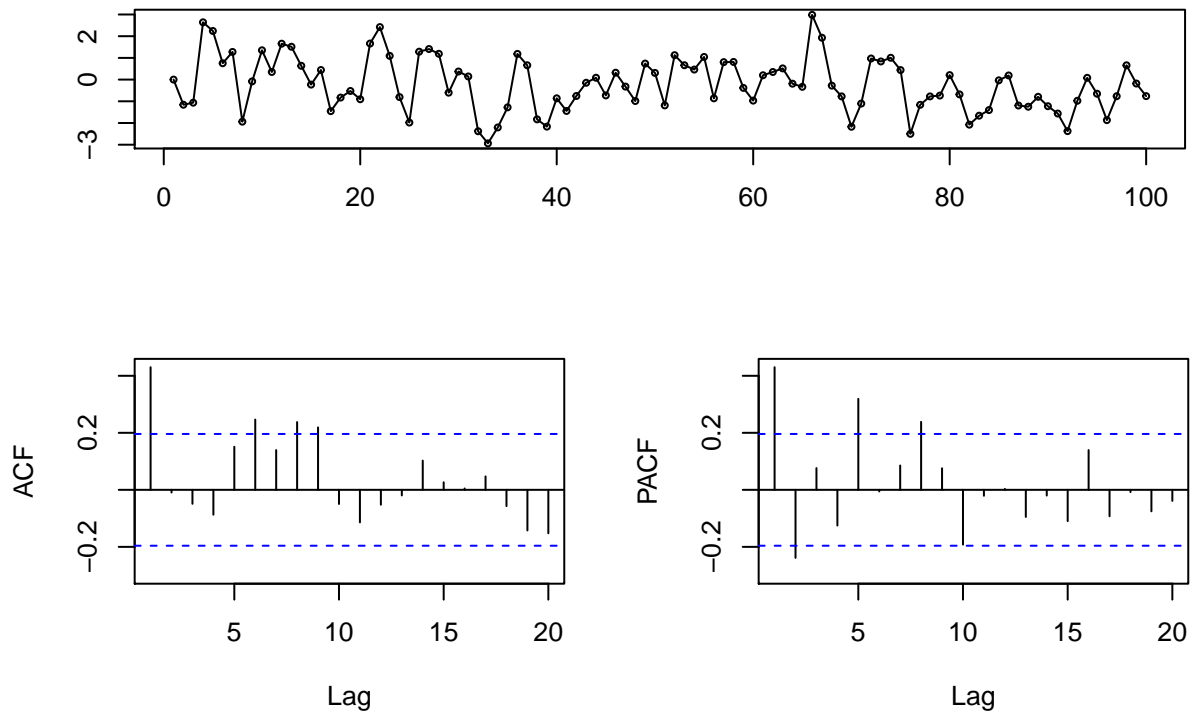


```

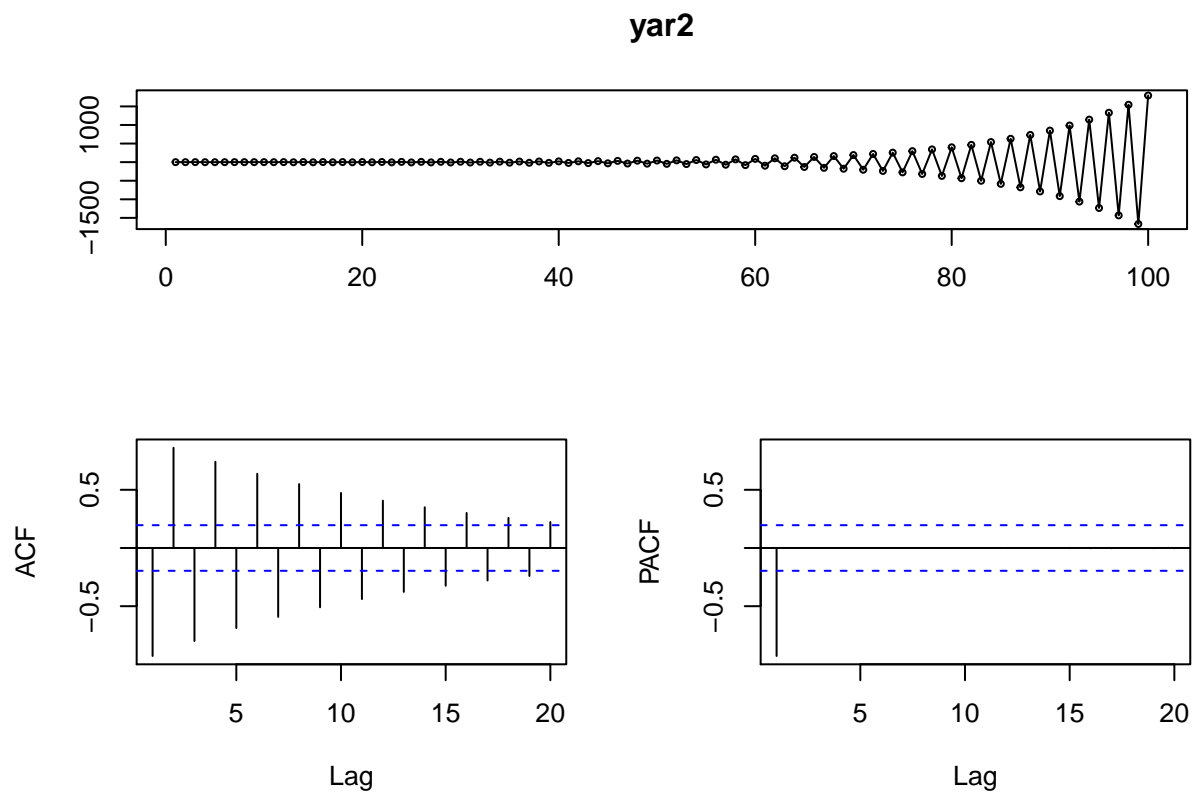
# as I change theta 1 in ma(1) model, the graph changes. As theta 1 becomes smaller, the frequency of f
## e
# generate arma(1,1)
yarma <- ts(numeric(100))
earma <- rnorm(100)
for (i in 2:100)
  yarma[i] <- 0.6*y[i-1]+0.6*earma[i-1]+earma[i]
## f
# generate ar(2)
yar2 <- ts(numeric(100))
ear2 <- rnorm(100)
for(i in 3:100)
  yar2[i] <- -0.8*yar2[i-1] +0.3*yar2[i-2]+ear2[i]
## g
# graph last two and compare
tsdisplay(yarma)

```

yarma



```
tsdisplay(yar2)
```



The oscillation magnitude increases as time increases in ar2 model, which is a non-stationary process.