Dude, Where's my Alpha? dispatches from a 'quant'

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USF 2013-02-07

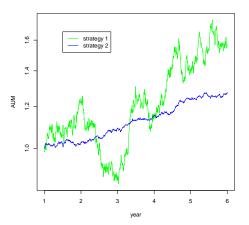
The Billion Dollar Quant Problem

Stipulate the following:

- Generating ideas for trading strategies is easy.
 Generating good ideas is harder.
- Backtesting strategies is an engineering problem.
 Not easy, not insurmountable: lots of possible biases, missing data, some domain knowledge needed.
- If evaluating strategies were easy, a lot of smart people would be rich.

Evaluating strategies must be hard.

A Simple Example: Choose One



Which strategy would you rather invest in? "Slow and Steady" or "Loose Cannon"?

- Let p_t be the 'mark-to-market' (MtM) at time t.
- Compute the *returns*:

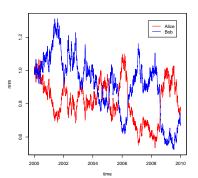
Geometric:
$$l_t =_{\sf df} \log \frac{p_t}{p_{t-1}},$$
 Arithmetic: $r_t =_{\sf df} \frac{p_t}{p_{t-1}} - 1 = e^{l_t} - 1.$

By simple calculus: $l_t \leq r_t$ with equality at zero.

- Sequential Geometric returns are additive by telescoping.
- Contemporaneous Arithmetic returns are additive:
 Arithmetic return of a portfolio is the dollar-weighted average of the components' arithmetic returns. (This includes 'shorting'.)

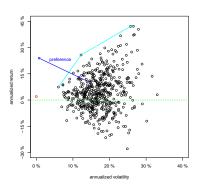
It's a Tough World

- The opposite of a losing strategy might be a losing strategy!
 Suppose Bob sells when Alice buys, and vice versa. Because of compounding, they can both have negative total log return.
- This effect is exacerbated by volatility.
- Factoring in costs, the 'average' strategy is certainly a loser.



the Risk Return Tradeoff

- Let $\mu = \mathsf{E}[l]$, $\sigma^2 = \mathrm{var}(l)$. Throughout, mostly assume l_t are *i.i.d.*, (unrealistic).
- For fixed μ , prefer a smaller σ ; for fixed σ prefer larger μ . The *efficient frontier* is the set of optimal strategies under this preference:



the Sharpe Ratio

- μ, σ are unknown; take sample estimates, $\hat{\mu}$, $\hat{\sigma}$.
- The Sharpe Ratio (SR) is the sample statistic[17]

$$\hat{\zeta} =_{\mathsf{df}} \frac{\hat{\mu}}{\hat{\sigma}},$$

May also include a 'risk-free' rate:

$$\hat{\zeta} = \frac{\hat{\mu} - r_0}{\hat{\sigma}}.$$

- Population analogue: Signal-to-Noise Ratio (SNR) $\zeta =_{\sf df} \mu/\sigma.$
- Connection between SR and t-statistic: $\hat{\zeta} = t/\sqrt{n}$.
- SR is Student's original test statistic. [2] The 'Student Ratio'?

Sharpe Ratio Minutiae I

- No real standard on arithmetic vs. geometric returns. (Although using arithmetic returns looks better ...)
- The units of Sharpe are 'per square root time': $\hat{\mu}$ is '(percent) per time', $\hat{\sigma}$ is '(percent) per root time'.
- Sharpe is often quoted in annualized units, *i.e.*, $yr^{-1/2}$. Avoid ambiguity and *always* include units.
- n.b., $1 \text{yr}^{-1/2} = \frac{1}{2} Q^{-1/2} = \frac{1}{\sqrt{12}} \text{mo.}^{-1/2} = \frac{1}{\sqrt{253}} \text{day}^{-1/2}$
- For equities quant strategies, achieved SR translates as: $1 \text{yr}^{-1/2} \Rightarrow \text{"good"}, 2 \text{yr}^{-1/2} \Rightarrow \text{"great"}, 3 \text{yr}^{-1/2} \Rightarrow \text{"legend"}.$ In HFT, higher SR the norm, but large fixed costs (like r_0).

Cantelli's Inequality gives:

$$\Pr\left\{\text{negative return on period}\right\} \leq \frac{1}{1+\zeta^2}$$

(Often more natural to consider $\zeta^2, \hat{\zeta}^2$.)

• Another view: a year-on-year loss is a " ζ sigma event". Central Limit Theorem: If returns are well behaved, enough gambles are made, annual return will be nearly normal.

$$\begin{split} &\zeta = 1 \mathrm{yr}^{-1/2} \Rightarrow \Pr\left\{ \mathrm{down\ year} \right\} \approx 16\% \\ &\zeta = 2 \mathrm{yr}^{-1/2} \Rightarrow \Pr\left\{ \mathrm{down\ year} \right\} \approx 2.3\% \\ &\zeta = 3 \mathrm{yr}^{-1/2} \Rightarrow \Pr\left\{ \mathrm{down\ year} \right\} \approx 0.13\% \\ &\zeta = 4 \mathrm{yr}^{-1/2} \Rightarrow \Pr\left\{ \mathrm{down\ year} \right\} \approx 0.0032\% \end{split}$$

Diversification and Sharpe Ratio

• Squared SNR of independent strategies is subadditive: For k $\bot\!\!\!\bot$ strategies, with SNRs ζ_1,\ldots,ζ_k , optimal rebalancing combination has

$$\zeta_*^2 = \sum_i \zeta_i^2$$

Diminishing returns in ζ : $2yr^{-1/2} + 1yr^{-1/2} \Rightarrow 2.2yr^{-1/2}$.

For dependent strategies, correlation is relevant.
 For two strategies with correlation ρ:

$$\zeta_*^2 = \frac{\zeta_1^2 + \zeta_2^2 - 2\rho\zeta_1\zeta_2}{1 - \rho^2}$$

• In general, want anti-correlated strategies ('hedges'). Avoid strategies correlated to 'retail alpha'.



Diversification vs Correlation

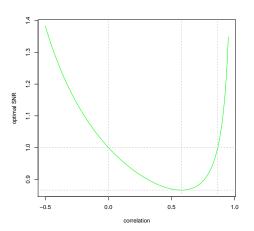


Figure: ζ_* is plotted versus ρ for $\zeta_1=0.5 {\rm yr}^{-1/2}, \zeta_2=0.87 {\rm yr}^{-1/2}$. It is possible to get *no* diversification benefit beyond best strategy.

Inference on Sharpe Ratio I

SR is actually a biased estimator of SNR:[14]

$$\mathsf{E}\left[\hat{\zeta}\right] = \sqrt{\frac{n-1}{2}} \frac{\Gamma\left((n-2)/2\right)}{\Gamma\left((n-1)/2\right)} \zeta = \frac{1}{c_4} \zeta.$$

Though the bias is small ($\leq 0.77\%$ when $n \geq 100$).

Standard error assuming normal returns: [21, 10, 6]

$$s.e. = \sqrt{\frac{1 + \frac{\zeta^2}{2}}{n - 1}} \approx \sqrt{\frac{1 + \frac{\hat{\zeta}^2}{2}}{n - 1}} \approx \sqrt{\frac{1}{n}}.$$

The latter annualizes as you would expect.

Rule of thumb: $\left|\zeta - \hat{\zeta}\right| \leq \frac{2}{\sqrt{n}}$, with probability ≈ 0.95 .

Inference on Sharpe Ratio II

- Fixes for heteroskedasticity, autocorrelation, skew. [20, 10, 15] (Typically these bias $\hat{\zeta}$ by < 10%)
- No easy fix for omitted variable bias i.e., "something changed in the world."
- Perhaps better to generalize to regression:

$$l_t = \alpha + \sum_i \beta_i f_{i,t} + \epsilon_t,$$

with f_i returns from known benchmark strategies.

- Typically want no exposure to f_i because it is available via ETPs, or one believes they are zero mean or cannot be timed.
- Whole literature on regression to perform inference on α . Inference on α/σ is less common.

Power and Sample Size I

A power rule is the relation between sample size, effect size (SNR), and rates of false positives and false negatives.

Figure out if you have enough data or enough 'alpha'.

- Hypothetical Vendor: "I have three years of historical data."
- Hypothetical Strategist: "This strategy has SNR $0.6 \text{yr}^{-1/2}$."
- Hypothetical Investor: "You have one year to prove yourself."

Good approximations of form

$$n \approx \frac{\kappa}{\zeta^2},$$

with κ a function of type I and type II rates, etc. [20, 6]

	one sided	two sided
power = 0.50	2.72	3.86
power = 0.80	6.20	7.87

Table: Value of κ to achieve given power in t-test, $\alpha = 0.05$.



Power and Sample Size II

rule of thumb for Sharpe power

To test SNR > 0, with 0.05 type I rate, and 50 % power,

$$n pprox rac{2.72}{\zeta^2}.$$
 mnemonic form: $e pprox n \zeta^2$

This rule is sobering:

- Hypothetical Vendor: "I have three years of historical data." Answer: Strategy must have SNR $\geq 0.95 \text{yr}^{-1/2}$.
- Hypothetical Strategist: "This strategy has SNR 0.6yr $^{-1/2}$." Answer: Need 7.6 years of data to backtest. (!)
- Hypothetical Investor: "You have one year to prove yourself." Answer: Strategy must have SNR $> 1.6 \text{yr}^{-1/2}$.

Why do Bad Things Happen to Smart People? I



Figure: "The majority of deployed quant strategies are type I errors." (Image courtesy of Automated Trader magazine.)

Why do Bad Things Happen to Smart People? II

Many explanations are offered:

- Bad backtests: inability to model costs and impact, observer effects, "backtest arbitrage," outright "time-travelling."
- Biased statistical tests. Rejecting the null for uninteresting reasons.
- Claims of "non-stationarity" ("the dog ate my alpha.")

Perhaps most important:

- False positives / Excessive type I rate.
- Sequential model overfit.

Consider the outcomes:

test \ truth	profitable	not profitable	
reject null	win!	type I error	
fail to reject	type II error	business as usual	

Let α , β be the type I and II rates; let c_0 be incidence rate of profitable strategies. Condition on test rejection:

test \ truth	profitable (c_0)	not profitable $(1-c_0)$
reject null	$c_0(1-\beta)$	$(1-c_0)\alpha$
fail to reject	$c_0\beta$	$(1-c_0)(1-\alpha)$

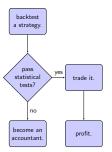
False Discovery Rate (FDR) is ratio of false to all positives:

$$\mathrm{FDR} =_{\mathrm{df}} \frac{(1-c_0)\alpha}{(1-c_0)\alpha + c_0(1-\beta)} \geq 1 - \frac{c_0}{\alpha} \ \mathrm{when} \ \beta \geq c_0.$$

Probably $c_0 \ll 0.05$. But is $c_0 \approx 10^{-5}$? $\approx 10^{-10}$?

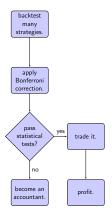
Why does Overfit Happen? I

How it is supposed to happen:



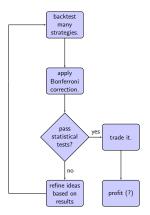
Why does Overfit Happen? II

Or even this:



Why does Overfit Happen? III

How it typically happens:



This is totally broken.

Why does Overfit Happen? IV

Why do quants overfit? A toxic mix:

- Failure is not an option: cannot become an accountant.
- The incidence rate is very low.
- By power rules, you need years of data to confirm alpha.
- There is only one history to work with:
 One collects data at a rate of 1 day per day.
- Often 'refining' ideas while debugging code.
- Sloppy process: often not aware one is overfitting.
- Data snooping tests require good record keeping, cannot deal with sequential overfit. [22, 3, 5]

The Portfolio Problem I

- Suppose you can invest in *p* different assets. (Stocks, ETPs, strategies, *etc.*) How do you choose a portfolio?
- This generalizes the discrete question, "is this asset good or not?" to the continuous "how much shall I invest in it?"
- Consider the arithmetic returns p-vector, x_t . Let μ , Σ be the (vector) mean and (matrix) covariance.
- If \hat{w} is proportion of capital in each asset, return is $\hat{w}^{\top}x_t$. This has mean $\hat{w}^{\top}\mu$ and variance $\hat{w}^{\top} \Sigma \hat{w}$. The SNR is

$$\frac{\hat{\boldsymbol{w}}^{\top}\boldsymbol{\mu}}{\sqrt{\hat{\boldsymbol{w}}^{\top}\Sigma\hat{\boldsymbol{w}}}}$$

In brief, pick a portfolio to maximize SNR.

The Portfolio Problem II

- The population parameters are unknown. Given nobservations of returns, compute 'usual' estimates $\hat{\mu}$, Σ .
- Markowitz Portfolio maximizes the SR: [11, 12, 13]

$$\hat{\boldsymbol{w}}_* =_{\mathsf{df}} \operatorname*{argmax} rac{\hat{\boldsymbol{w}}^{ op} \hat{\boldsymbol{\mu}}}{\sqrt{\hat{\boldsymbol{w}}^{ op} \hat{\Sigma} \hat{\boldsymbol{w}}}}.$$

May also incorporate r_0 , but must bound risk:

$$\hat{m{w}}_* =_{\mathsf{df}} \operatorname*{argmax}_{\hat{m{w}}: \hat{m{w}}^{ op} \hat{m{\Sigma}} \hat{m{w}} < s^2} rac{\hat{m{w}}^{ op} \hat{m{\mu}} - r_0}{\sqrt{\hat{m{w}}^{ op} \hat{m{\Sigma}} \hat{m{w}}}}.$$

• The optimal portfolio is $\hat{\boldsymbol{w}}_* \propto \hat{\Sigma}^{-1} \hat{\boldsymbol{\mu}}$. with SR:

$$\hat{\zeta}_* = \sqrt{\hat{\boldsymbol{\mu}}^\top \hat{\Sigma}^{-1} \hat{\boldsymbol{\mu}}} - \frac{r_0}{s}.$$

Drop the r_0 for now, and consider the RV $\hat{\zeta}_*^2 = \hat{\mu}^\top \hat{\Sigma}^{-1} \hat{\mu}$.



• Hotelling T^2 is multivariate generalization of t: [4, 16, 19]

$$T^2 =_{\mathsf{df}} n \hat{\boldsymbol{\mu}}^{\top} \hat{\Sigma}^{-1} \hat{\boldsymbol{\mu}} = n \hat{\zeta}_*^2$$

(Recall that in the univariate case $t = \sqrt{n}\hat{\zeta}$.)

• For normal returns, rescaled T^2 has F-distribution:

$$\frac{(n-p)}{p(n-1)}n\hat{\zeta}_*^2 \sim F\left(n\zeta_*^2, p, n-p\right),\,$$

where ζ_* is SNR of *population* optimal portfolio, $\nu_* = \Sigma^{-1} \mu$. So $\hat{\zeta}_*$ can be used to perform inference on ζ_* .

• n.b. ζ_* is maximal SNR of any portfolio, including $\hat{\boldsymbol{w}}_*$. We expect some loss: $SNR\left(\hat{\boldsymbol{w}}_*\right) \leq SNR\left(\boldsymbol{\nu}_*\right)$. The loss depends on sample size, effect size; not well understood.

- Observe $\hat{\zeta}_*$, perform inference on ζ_* , an upper bound on SNR of $\hat{\boldsymbol{w}}_*$.
- Via F-distribution we have

$$\mathsf{E}\left[\hat{\zeta}_*^2\right] = \frac{\zeta_*^2 + c}{1 - c}, \qquad \text{where } c =_{\mathsf{df}} p/n.$$

Gives unbiased estimator: $\mathsf{E}\left[(1-c)\hat{\zeta}_*^2-c\right]=\zeta_*^2$. (Oops! unbiased estimator of ζ_*^2 can be negative!)

- CI, MLE on ζ_* via F-CDF & PDF. Shrinkage estimators.[9]
- If $\hat{\zeta}_*^2 \leq \frac{c}{1-c}$ then MLE of ζ_* is 0. [18]

rule of thumb, based on F-MLE

"If $\hat{\zeta}_*^2 < c$, don't bother!"

Example: Hotelling on S&P Sector Indices

Index data BASI, INDU, CONG, HLTH, CONS, TELE, UTIL, FINA, TECH from fPortfolio::SPISECTOR, from 2000-01-04 to 2008-10-17. (n=2198 days, p=9 stocks)

- Optimal in-sample Sharpe ratio is $1yr^{-1/2}$.
- MLE for ζ_* is $0.1 {\rm yr}^{-1/2}$. Close to the rule of thumb cutoff: $\hat{\zeta}_*^2=1.05 {\rm yr}^{-1}$ and $c=1.04 {\rm yr}^{-1}$.
- 95% CI for ζ_* is $[0\text{yr}^{-1/2}, 1.3\text{yr}^{-1/2}]$

Conclusion: Markowitz portfolio on S&P sectors not recommended.

Approximation of Strategy Overfit I

Caricature of quant work:

- Construct strategy with free parameters θ ;
- Backtest strategy for $\theta_1, \theta_2, \dots, \theta_m$.
- Pick θ_i that maximizes SR of backtest, θ_* .
- Profit! (or not)

Q: How to estimate the SNR of θ_* ?

Approximation of Strategy Overfit II

Toy Example: Moving Average Crossover:

- θ is vector of 2 window lengths; Long the instrument exactly when one moving average exceeds the other.
- Brute-force backtest for allowable window lengths.

Approximation of Strategy Overfit III

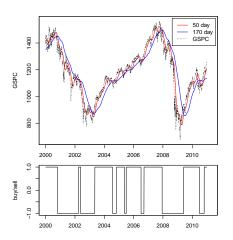


Figure: Two-window MAC on GSPC, with buy/sell signal at bottom.

Approximation of Strategy Overfit IV

Q: How to estimate the SNR of θ_* ?

A(?): Make PCA-like linear approximation of returns vectors:

$$\{oldsymbol{x}_1,\ldots,oldsymbol{x}_m\}pprox\mathcal{K}\subset\mathcal{L}=_{\sf df}\{oldsymbol{\mathsf{Y}}oldsymbol{\hat{w}}\midoldsymbol{\hat{w}}\in\mathbb{R}^p\}$$

Use $\zeta\left(\theta_{*}\right) \approx \max_{\mathcal{L}} \zeta = \zeta_{*}$, then make inference on ζ_{*} . Observe maximal strategy SR

$$\hat{\zeta}_* = \hat{\zeta}\left(\theta_*\right) =_{\mathsf{df}} \max_{1 \le i \le m} \hat{\zeta}\left(\theta_i\right),$$

use it to approximate maximal SR over \mathcal{L} .

You have to estimate p, by Monte Carlo under null, PCA, or SWAG method.

Overfit of Simple 2-Window MAC

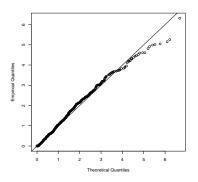


Figure: 2048 Monte Carlo sims of model selection in 2-window MAC over 2500 days, under the null (no population drift or autocorrelation). In-sample $\hat{\zeta}_*$ values transformed to F-statistics with p=2.025.

Use on GSPC adjusted returns (2000-01-03 to 2009-12-31, 2515 days), MLE of ζ_* is $0.65 \text{yr}^{-1/2}$; 95% CI is $(0 \text{yr}^{-1/2}, 1.3 \text{yr}^{-1/2})$.

Conclusions

- I hope I did not scare you.
- Statistics can useful to avoid quant disasters.
- It is better to be lucky than good.
- We are looking for interns!

Bibliography I



Mark Britten-Jones.

The sampling error in estimates of mean-variance efficient portfolio weights.



William Sealy Gosset.

The probable error of a mean.

Biometrika, 6(1):1-25, March 1908.

Originally published under the pseudonym "Student".

The Journal of Finance, 54(2):655-671, 1999.



P. R. Hansen.

A test for superior predictive ability.

Journal of Business and Economic Statistics, 23(4), 2005.



Harold Hotelling

The generalization of Student's ratio.

Annals of Mathematical Statistics, 2(3):360-378, 1931.



Po-Hsuan Hsu, Yu-Chin Hsu, and Chung-Ming Kuan.

Testing the predictive ability of technical analysis using a new stepwise test without data snooping bias. 2008.



N. L. Johnson and B. L. Welch.

Applications of the non-central t-distribution.

Biometrika, 31(3-4):362-389, March 1940.



Raymond Kan and Daniel R. Smith.

The distribution of the sample minimum-variance frontier.

Management Science, 54(7):1364-1380, 2008.

Bibliography II



Raymond Kan and GuoFu Zhou.

Tests of mean-variance spanning.

Annals of Economics and Finance, 13(1), 2012.



T. Kubokawa, CP Robert, and AK Saleh.

Estimation of noncentrality parameters. Canadian Journal of Statistics, 21(1):45–57, 1993.



Andrew W. Lo.

The Statistics of Sharpe Ratios.

Financial Analysts Journal, 58(4), July/August 2002.



H. Markowitz.

Portfolio selection*.

The journal of finance, 7(1):77–91, 2012.



H.M. Markowitz.

The early history of portfolio theory: 1600-1960.

Financial Analysts Journal, pages 5-16, 1999.



H.M. Markowitz.

Foundations of portfolio theory.

The Journal of Finance, 46(2):469-477, 2012.



Robert E. Miller and Adam K. Gehr.

Sample size bias and Sharpe's performance measure: A note.

Journal of Financial and Quantitative Analysis, 13(05):943-946, 1978.

Bibliography III



J. D. Opdyke.

Comparing sharpe ratios: So where are the p-values? Journal of Asset Management, 8(5), 2007.



Pranah Kumar Sen

Union-intersection principle and constrained statistical inference.

Journal of Statistical Planning and Inference, 137(11):3741–3752, 2007.

Special Issue: In Celebration of the Centennial of The Birth of Samarendra Nath Roy (1906-1964).



William F. Sharpe.

Mutual fund performance.

Journal of Business, 39:119, 1965.



M. C. Spruill.

Computation of the maximum likelihood estimate of a noncentrality parameter.

Journal of Multivariate Analysis, 18(2):216-224, 1986.



N.H. Timm.

Applied multivariate analysis: methods and case studies.

Springer Texts in Statistics. Physica-Verlag, 2002.



Gerald van Belle.

Statistical Rules of Thumb (Wiley Series in Probability and Statistics).

Wiley-Interscience, 1 edition, March 2002.



Christian Walck.

Hand-book on STATISTICAL DISTRIBUTIONS for experimentalists.

December 1996.

Bibliography IV



Halbert White.

A reality check for data snooping. *Econometrica*, 68:1097–1127, 2000.

Skew and Sharpe I

distribution	param	skew	ex. kurtosis	type I
Gaussian		0	0	0.05
Student's t	df = 10	0	1	0.047
SP500		-1	23	0.064
symmetric SP500		0	22	0.058
Tukey h	h = 0.1	0	5.5	0.06
Tukey h	h = 0.24	0	1.3e+03	0.047
Tukey h	h = 0.4		Inf	0.17
Lambert W x Gaussian	delta = -0.2	-1.2	5.7	0.054
Lambert W x Gaussian	delta = -0.4	-2.7	18	0.086
${\sf Lambert}\ {\sf W}\times {\sf Gaussian}$	delta = -1.2	-30	5.2e + 03	0.26

Table: Empirical type I rates of the test for $\zeta=1.0$ via distribution of the Sharpe ratio are given for various distributions of returns. The empirical rates are based on 2048 simulations of three years of daily returns, with a nominal rate of $\alpha=0.05$. Skew appears to have a much more adverse effect than kurtosis alone.

Wait, I Want Returns

Idea: lever up until volatility "too high." Set this by a drawdown cap. Sharpe Ratio limits drawdowns.

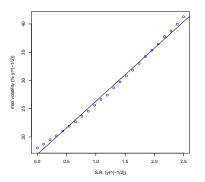


Figure: Maximum volatility (percent per root year) so that probability of a 33 % drawdown over a year is 2.5% or less. Assumes independence, homoskedasticity, mild kurtosis.

- Can generalize Hotelling statistic to get 'Spanning Tests' [1, 8]
- Let T_{p+q}^2, T_p^2 be Hotelling stats on full set of p+q assets and subset of p assets. Do the q marginal assets 'add any value'.
- Let

$$\Delta T^2 = (n-p-1)\frac{T_{p+q}^2 - T_p^2}{n-1 + T_p^2} = (n-p-1)\frac{\hat{\zeta}_{*,p+q}^2 - \hat{\zeta}_{*,p}^2}{1 - (1/n) + \hat{\zeta}_{*,p}^2}.$$

• ΔT^2 also takes a (non-central) Hotelling distribution:

$$\frac{n - (p+q)}{q(n-p-1)} \Delta T^2 \sim F(q, n - (p+q), \delta),$$

$$\delta = (n-p-1) \frac{\zeta_{*,p+q}^2 - \zeta_{*,p}^2}{1 - (1/n) + \zeta_{*,p}^2}.$$

• Same inference on δ can be applied (MLE, CI).

Delta Hotelling on S&P Sector Indices

Delta Hotelling: what do BASI, INDU, CONG, HLTH, CONS, TELE add to UTIL, FINA, TECH?

- MLE for $(\zeta_{*,p+q}^2 \zeta_{*,p}^2)$ is 0yr^{-1} .
- 95% CI for $\left(\zeta_{*,p+q}^2 \zeta_{*,p}^2\right)$ is $\left[0 \text{yr}^{-1}, 0.26 \text{yr}^{-1}\right]$.