

Unit-Root Nonstatornarity, VAR(p) Model & Cointegration Analysis

[Code ▾](#)

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The file `m_cofi_4rates.txt` contains the monthly rates of the 11th District Cost of Funds Index (COFI), the prime rate of U.S. banks, 1-year and 5-year U.S. Treasury constant maturity rates, and U.S. Treasury 3-month secondary market rates from September 1989 to June 2007. The COFI rates are obtained from the Federal Home Loan Bank of San Francisco, and the other rates are obtained from the Federal Reserve Bank of St. Louis. COFI is a weighted-average interest rate paid by savings institutions headquartered in Arizona, California, and Nevada and is one of the most popular adjustable-rate mortgage (ARM) indices. The prime rate is the interest rate at which banks lend to their most creditworthy customers.

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```
df <- read.table('m_cofi_4rates.txt', header=TRUE)
# this question's code was inspired by https://web.stanford.edu/~xing/statfinbook/_BookFun/ex9.4.5_unitrootCoint.txt
```

(a) Perform the augmented Dickey-Fuller test of the unit-root hypothesis for each of these rates.

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```
print('The augmented Dickey-Fuller test of the unit-root hypothesis for the monthly rates of COFI from September 1989 to June 2007 is:')
```

```
[1] "The augmented Dickey-Fuller test of the unit-root hypothesis for the monthly rates of COFI from September 1989 to June 2007 is:"
```

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```
adf.test(df$cofi)
```

Augmented Dickey-Fuller Test

```
data: df$cofi
Dickey-Fuller = -2.8745, Lag order = 5, p-value = 0.2092
alternative hypothesis: stationary
```

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```
print('The augmented Dickey-Fuller test of the unit-root hypothesis for the 1-year U.S. Treasury constant maturity rates from September 1989 to June 2007 is:')
```

```
[1] "The augmented Dickey-Fuller test of the unit-root hypothesis for the 1-year U.S. Treasury constant maturity rates from September 1989 to June 2007 is:"
```

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```
adf.test(df$X1ycmt)
```

Augmented Dickey-Fuller Test

```
data: df$X1ycmt
Dickey-Fuller = -2.0479, Lag order = 5, p-value = 0.5559
alternative hypothesis: stationary
```

Hide

```
print('The augmented Dickey-Fuller test of the unit-root hypothesis for the 5-year U.S. Treasury constant maturity rates from September 1989 to June 2007 is:')
```

```
[1] "The augmented Dickey-Fuller test of the unit-root hypothesis for the 5-year U.S. Treasury constant maturity rates from September 1989 to June 2007 is:"
```

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```
adf.test(df$X5ycmt)
```

Augmented Dickey-Fuller Test

```
data: df$X5ycmt  
Dickey-Fuller = -2.0488, Lag order = 5, p-value = 0.5555  
alternative hypothesis: stationary
```

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```
print('The augmented Dickey-Fuller test of the unit-root hypothesis for the prime rates of U.S. banks from September 1989 to June 2007 is:')
```

```
[1] "The augmented Dickey-Fuller test of the unit-root hypothesis for the prime rates of U.S. banks from September 1989 to June 2007 is:"
```

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```
adf.test(df$primeRate)
```

Augmented Dickey-Fuller Test

```
data: df$primeRate  
Dickey-Fuller = -2.3052, Lag order = 5, p-value = 0.448  
alternative hypothesis: stationary
```

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```
print('The augmented Dickey-Fuller test of the unit-root hypothesis for the U.S. Treasury 3-month secondary market rates from September 1989 to June 2007 is:')
```

```
[1] "The augmented Dickey-Fuller test of the unit-root hypothesis for the U.S. Treasury 3-month secondary market rates from September 1989 to June 2007 is:"
```

Hide

```
adf.test(df$X3mTbill.2mkt)
```

Augmented Dickey-Fuller Test

```
data: df$X3mTbill.2mkt  
Dickey-Fuller = -2.5454, Lag order = 5, p-value = 0.3472  
alternative hypothesis: stationary
```

(b) Assuming the VAR(2) model for the multivariate time series of these five rates, perform Johansen's test for the number of cointegration vectors.

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```
johansens <- ca.jo(df, type="eigen", K=2, season=NULL)  
summary(johansens)
```

```
#####
# Johansen-Procedure #
#####
```

Test type: maximal eigenvalue statistic (lambda max) , with linear trend

Eigenvalues (lambda):

```
[1] 0.28115617 0.18230111 0.12979406 0.07225424 0.03464075
```

Values of teststatistic and critical values of test:

| | test | 10pct | 5pct | 1pct |
|--------|-------|-------|-------|-------|
| r <= 4 | 7.47 | 6.50 | 8.18 | 11.65 |
| r <= 3 | 15.90 | 12.91 | 14.90 | 19.19 |
| r <= 2 | 29.47 | 18.90 | 21.07 | 25.75 |
| r <= 1 | 42.67 | 24.78 | 27.14 | 32.14 |
| r = 0 | 69.98 | 30.84 | 33.32 | 38.78 |

Eigenvectors, normalised to first column:

(These are the cointegration relations)

| | cofi.l2 | X1ycmt.l2 | X5ycmt.l2 | primeRate.l2 | X3mTbill.2mkt.l2 |
|------------------|-----------|------------|------------|--------------|------------------|
| cofi.l2 | 1.000000 | 1.000000 | 1.000000 | 1.000000 | 1.000000 |
| X1ycmt.l2 | -5.351875 | 2.7399460 | 2.4996108 | 1.384167 | 2.9771100 |
| X5ycmt.l2 | 1.534273 | -0.9231277 | -1.5278353 | -1.699396 | -4.6375957 |
| primeRate.l2 | 1.908798 | 0.3260260 | 0.9516442 | -3.724769 | -1.7744334 |
| X3mTbill.2mkt.l2 | 1.443581 | -3.2981390 | -3.0118510 | 2.822877 | -0.6152458 |

Weights W:

(This is the loading matrix)

| | cofi.l2 | X1ycmt.l2 | X5ycmt.l2 | primeRate.l2 |
|-----------------|-------------|--------------|-------------|--------------|
| cofi.d | -0.02811300 | -0.028599110 | 0.00204150 | -0.005949579 |
| X1ycmt.d | -0.02835492 | 0.032656460 | -0.02063724 | -0.003506453 |
| X5ycmt.d | -0.02717289 | 0.007413592 | 0.03393117 | 0.017832126 |
| primeRate.d | -0.04463161 | 0.011492886 | -0.04974080 | 0.009034110 |
| X3mTbill.2mkt.d | -0.04605080 | 0.116701546 | -0.02176690 | -0.009198674 |

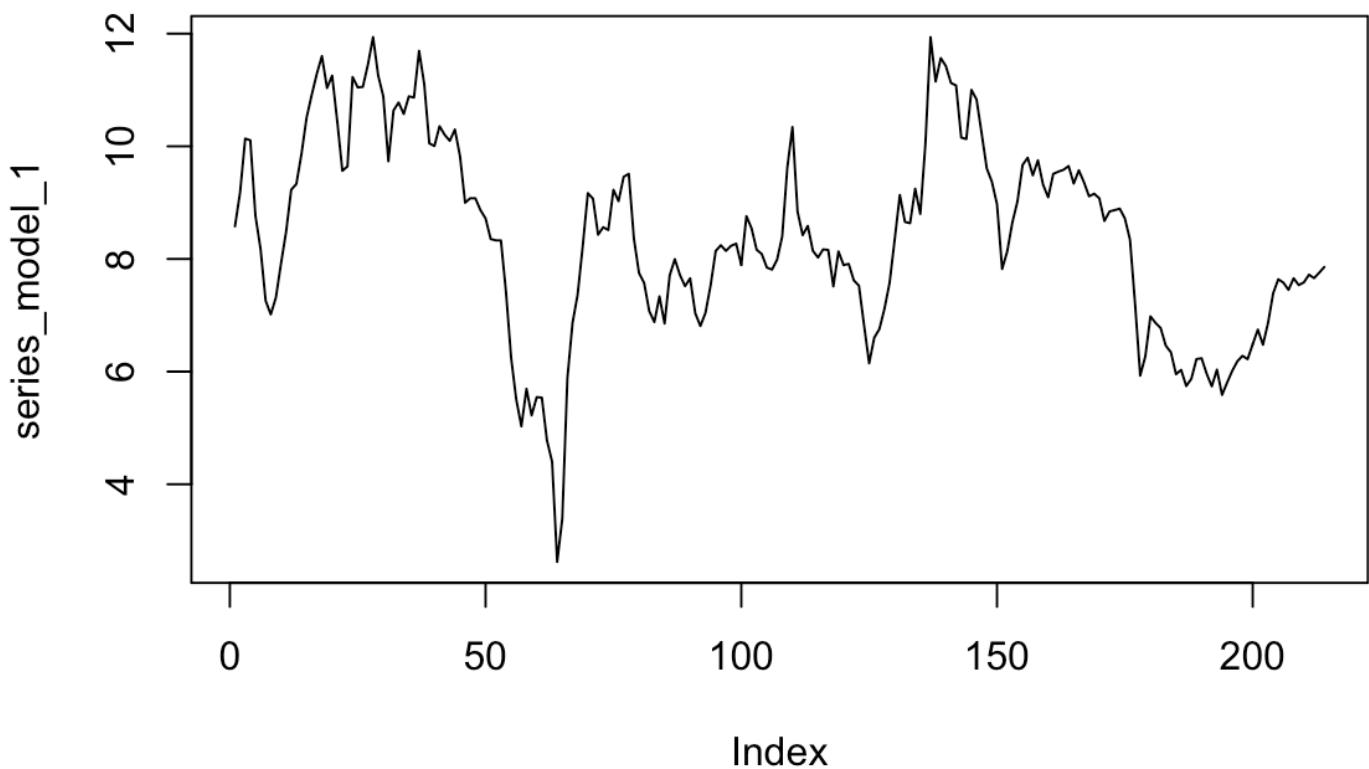
| | X3mTbill.2mkt.l2 |
|-----------------|------------------|
| cofi.d | 2.823119e-05 |
| X1ycmt.d | 8.259722e-03 |
| X5ycmt.d | 9.537910e-03 |
| primeRate.d | 3.817692e-04 |
| X3mTbill.2mkt.d | 4.400697e-03 |

(c) Estimate the cointegration vectors and use them to describe the equilibrium relationship between the five rates.

```

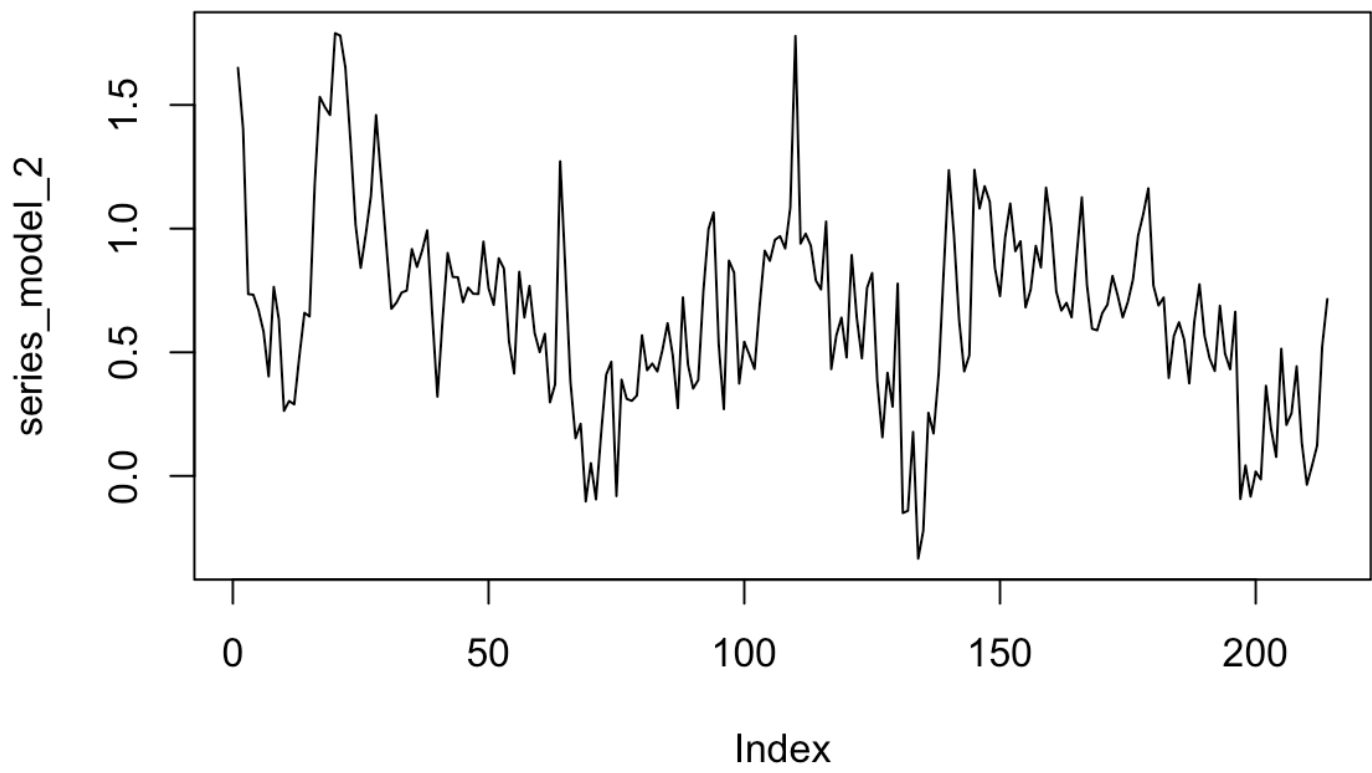
coin_vec_1 <- c(1, -5.351875, 1.534274, 1.908798, 1.443581)
coin_vec_2 <- c(1, 2.739946, -0.9231277, 0.3260260, -3.2981390)
coin_vec_3 <- c(1, 2.4996108, -1.5278353, 0.9516442, -3.0118510)
coin_vec_4 <- c(1, 1.384167, -1.699396, -3.724769, 2.822877)
coin_1 <- df$cofi * 1 - 5.351875 * df$X1ycmt + 1.534274 * df$X5ycmt + 1.908798 * df$primeRate + 1.443581 * df$X3mTbill.2mkt
coin_2 <- df$cofi * 1 + 2.739946 * df$X1ycmt - 0.9231277 * df$X5ycmt + 0.3260260 * df$primeRate - 3.2981390 * df$X3mTbill.2mkt
coin_3 <- df$cofi * 1 + 2.4996108 * df$X1ycmt - 1.5278353 * df$X5ycmt + 0.9516442 * df$primeRate - 3.0118510 * df$X3mTbill.2mkt
coin_4 <- df$cofi * 1 + 1.384167 * df$X1ycmt - 1.699396 * df$X5ycmt - 3.724769 * df$primeRate + 2.822877 * df$X3mTbill.2mkt
series_model_1 <- as.matrix(df) %>% coin_vec_1
series_model_2 <- as.matrix(df) %>% coin_vec_2
series_model_3 <- as.matrix(df) %>% coin_vec_3
series_model_4 <- as.matrix(df) %>% coin_vec_4
plot(series_model_1, type = 'l')

```



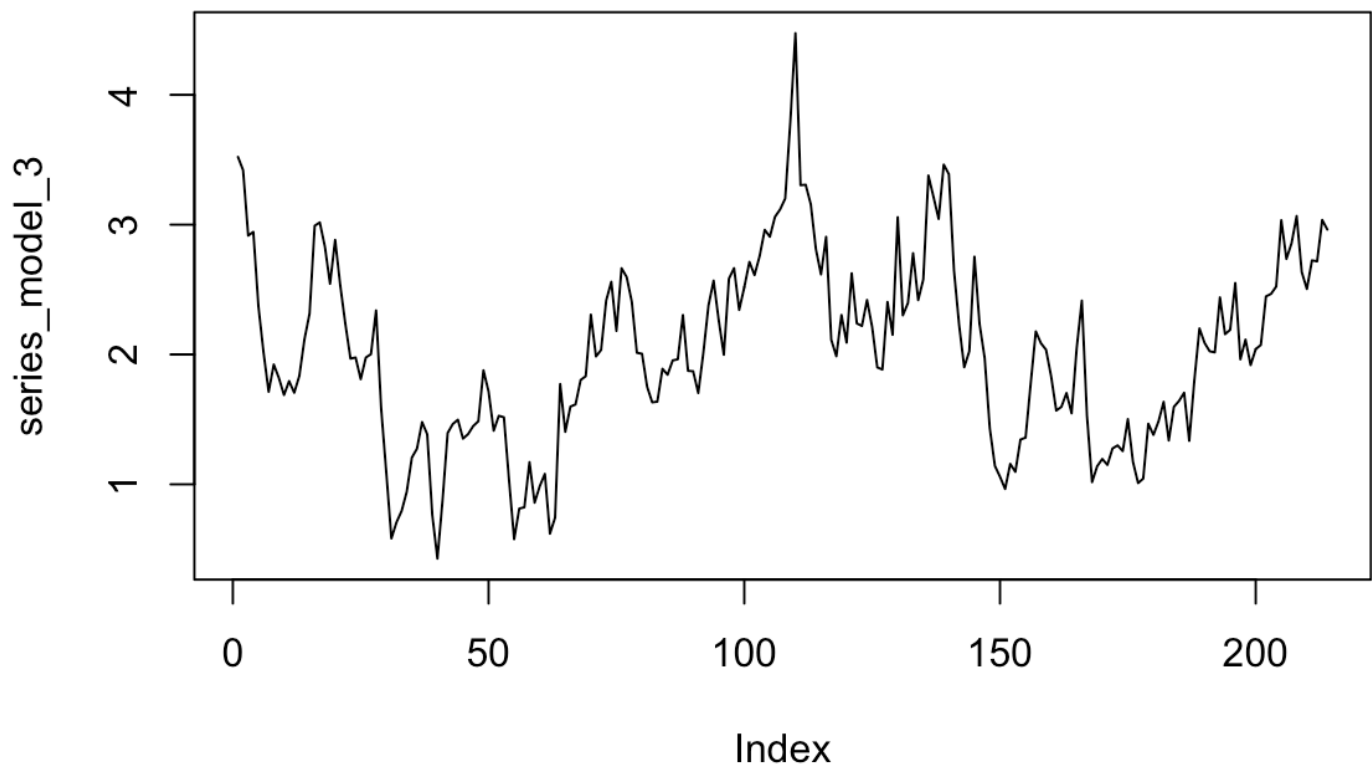
Hide

```
plot(series_model_2, type = 'l')
```



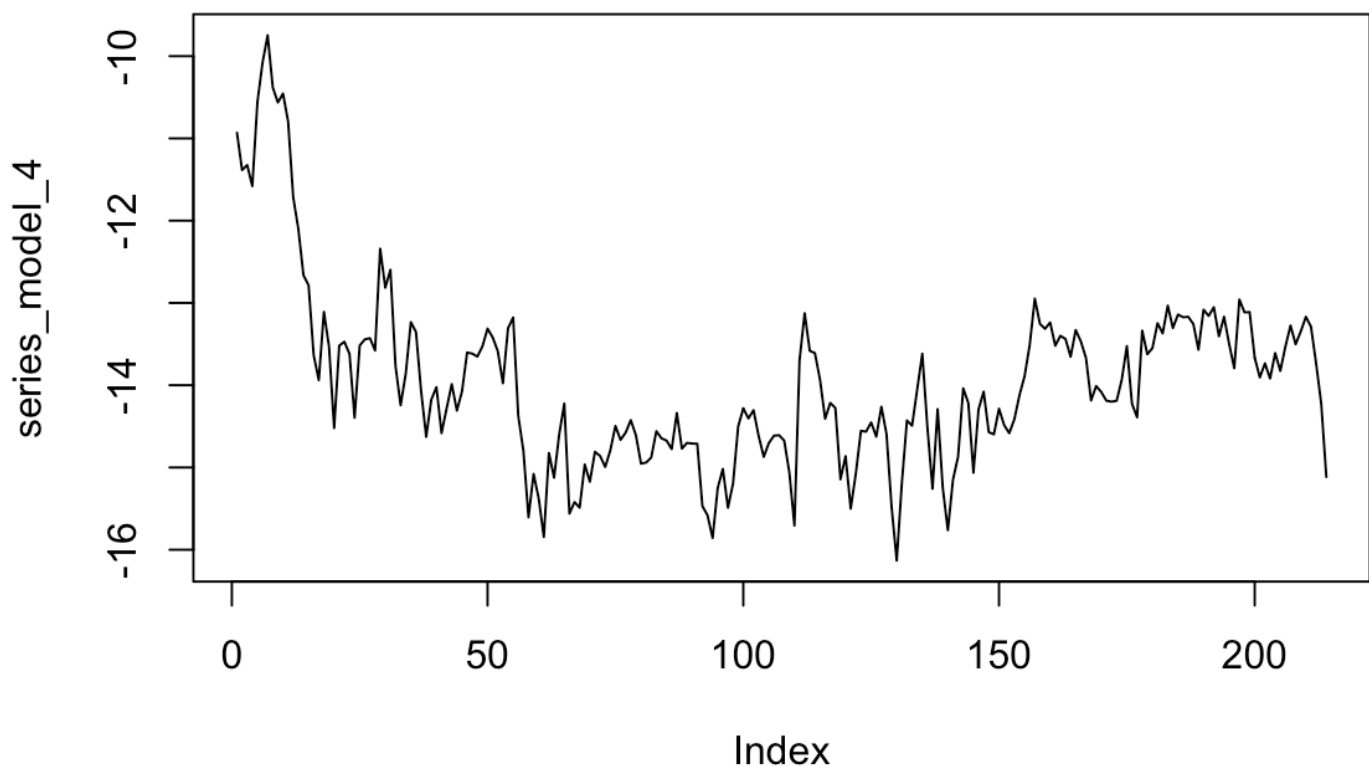
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```
plot(series_model_3, type = 'l')
```



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```
plot(series_model_4, type = 'l')
```

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```
print('Given our estimates of the cointegration vectors and their respective plots, we see that even though the individual components are not stationary and have variances that diverge to infinity, they share common trends that result in  $\beta_i^{\text{top}} y_t$  having long-run equilibrium for  $i=1, \dots, 4$ .')
```

```
[1] "Given our estimates of the cointegration vectors and their respective plots, we see that even though the individual components are not stationary and have variances that diverge to infinity, they share common trends that result in  $\beta_i^{\text{top}} y_t$  having long-run equilibrium for  $i=1, \dots, 4$ ."
```

(d) Regress COFI on the four other rates. Discuss the economic meaning of this regression relationship and whether the regression is spurious.

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```
fit <- lm(df$cofi ~ df$X1ycmt + df$X5ycmt + df$primeRate + df$X3mTbill.2mkt)
summary(fit)
```

```
Call:
lm(formula = df$cofi ~ df$X1ycmt + df$X5ycmt + df$primeRate +
    df$X3mTbill.2mkt)

Residuals:
    Min       1Q   Median       3Q      Max
-0.71351 -0.21455 -0.03518  0.16148  1.26509

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   -0.874633    0.338637  -2.583   0.0105 *
df$X1ycmt     -2.897190    0.116331 -24.905  <2e-16 ***
df$X5ycmt      1.236494    0.046559  26.558  <2e-16 ***
df$primeRate  -0.003372    0.087235  -0.039   0.9692
df$X3mTbill.2mkt 2.859960    0.140077  20.417  <2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3328 on 209 degrees of freedom
Multiple R-squared:  0.9607,    Adjusted R-squared:  0.96
F-statistic: 1279 on 4 and 209 DF,  p-value: < 2.2e-16
```

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```
print('Given the high R^2 we can conclude that the model explains the variability of
the response data around its mean fairly well. The COFI is a regional average of inte
rest expenses incurred by financial institutions, which in turn is used as a base for
calculating variable rate loans. It makes sense that these rates would vary according
to one another so the regression is unlikely to be spurious. Futhermore, from part (c
) since \beta_{ji} \neq 0, the linear regression of y_{tj} on the other components of
y_t is not spurious even though y_t is unit root stationary.')
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