

CSCI 567 Homework 4

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October 29, 2017

Problem 1. Generative models

• Question 1.1

$$L(\theta; x_1, x_2, \dots, x_n) = \log P(X = x_1, \dots, x_n | \theta) = \log(P(x_1 | \theta)P(x_2 | \theta) \dots P(x_n | \theta)) \quad (1)$$

Thus

$$L(\theta; x_1, x_2, \dots, x_n) = \log\left(\frac{1}{\theta}\right)^n 1[0 < x_1 \leq \theta, \dots, 0 < x_n \leq \theta] = n \log\left(\frac{1}{\theta}\right) 1[0 < x_1 \leq \theta, \dots, 0 < x_n \leq \theta] \quad (2)$$

From $0 < x_1 \leq \theta, 0 < x_2 \leq \theta, \dots, 0 < x_n \leq \theta$, we get: $\theta \geq \max(x_1, x_2, \dots, x_n)$. Thus:

$$\hat{\theta}_{ML} = \operatorname{argmax}_{\theta} \{n \log\left(\frac{1}{\theta}\right)\} = \max(x_1, x_2, \dots, x_n) \quad (3)$$

• Question 1.2

• first part:

$$P(k | x_n, \theta_1, \theta_2, w_1, w_2) = \frac{P(x_n, \theta_1, \theta_2, w_1, w_2 | k) P(k)}{P(x_n, \theta_1, \theta_2, w_1, w_2)} \quad (4)$$

$$P(k | x_n, \theta_1, \theta_2, w_1, w_2) = \frac{w_k \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k]}{\sum_{j=1}^2 w_j \frac{1}{\theta_j} 1[0 < x_n \leq \theta_j]} \quad (5)$$

• second part - assuming $\theta = w_1, w_2, \theta_1, \theta_2$,

$$Q(\theta, \theta^{OLD}) = \sum_{n=1}^N \sum_{k=1}^2 q(z_n = k) \log P(x_n, z_n = k | \theta) \quad (6)$$

$$q(z_n = k) = P(z_n = k | \theta^{OLD}) = P(z_n = k | w_1^{OLD}, w_2^{OLD}, \theta_1^{OLD}, \theta_2^{OLD}) \quad (7)$$

We also have:

$$\log P(x_n, z_n = k | \theta) = \log(P(x_n | z_n = k, \theta_k) P(z_n = k | \theta_k)) = \log P(x_n | z_n = k, \theta_k) + \log P(z_n = k | \theta_k) \quad (8)$$

$$\log P(x_n, z_n = k|\theta) = \log U(x_n|z_n = k, \theta_k) + \log w_k = \log w_k + \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \quad (9)$$

From 6, and using 5, 7 and 9, we get:

$$Q(\theta, \theta^{OLD}) = \sum_{n=1}^N \sum_{k=1}^2 \left(\frac{w_k^{OLD} \frac{1}{\theta_k^{OLD}} 1[0 < x_n \leq \theta_k^{OLD}]}{\sum_{j=1}^2 w_j^{OLD} \frac{1}{\theta_j^{OLD}} 1[0 < x_n \leq \theta_j]} \right) (\log w_k + \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k]) \quad (10)$$

- third part - defining $P_{OLD}(k|x_n) = \frac{w_k^{OLD} \frac{1}{\theta_k^{OLD}} 1[0 < x_n \leq \theta_k^{OLD}]}{\sum_{j=1}^2 w_j^{OLD} \frac{1}{\theta_j^{OLD}} 1[0 < x_n \leq \theta_j]}$, we have:

$$Q(\theta, \theta^{OLD}) = \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) (\log w_k + \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k]) \quad (11)$$

In the M-step, we follow the following update rule: $\theta^{new} \leftarrow \argmax_{\theta} \{Q(\theta, \theta_{OLD})\}$

$$Q(\theta, \theta^{OLD}) = \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) \log w_k + \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \quad (12)$$

which means w_k and θ_k can be optimized separately. Also, term $P_{OLD}(k|x_n)$ is independent of optimization parameters and can be treated as a constant. Thus

- θ_k^* :

$$\begin{aligned} \argmax_{\theta_k} \{ \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \} &= \\ \argmax_{\theta_k} \{ \sum_{n=1}^N \sum_{k=1}^2 \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \} &= \\ \argmax_{\theta_k} \{ \sum_{n=1}^N \sum_{k=1}^2 \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \} &= \\ \argmin_{\theta_k} \{ \sum_{n=1}^N \sum_{k=1}^2 \theta_k 1[0 < x_n \leq \theta_k] \} & \end{aligned}$$

Using $\theta_2^{OLD} \geq \max(x_1, \dots, x_n)$ and $\theta_2^{OLD} \geq \min(x_1, \dots, x_n)$, we get:

$$\argmax_{\theta_1} \{ \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \} = \min(x_1, \dots, x_n)$$

which leads to: $\theta_1^{t+1} \leftarrow \min(x_1, \dots, x_n)$

$$\argmax_{\theta_2} \{ \sum_{n=1}^N \sum_{k=1}^2 P_{OLD}(k|x_n) \log \frac{1}{\theta_k} 1[0 < x_n \leq \theta_k] \} = \max(x_1, \dots, x_n)$$

which leads to: $\theta_2^{t+1} \leftarrow \max(x_1, \dots, x_n)$

- w_k^* : $\argmax \sum_n \sum_k \log(w_k)$

we know that: $0 < w_k \leq 1$, thus: $w_1^{t+1} \leftarrow 1$ and $w_2^{t+1} \leftarrow 1$.

Problem 2. Mixture density models

• Question 2.1

$$P(x) = \sum_{k=1}^K \pi_k P(x|k) \quad (13)$$

$$P(x_a, x_b) = \sum_{k=1}^K \pi_k P(x_a, x_b | k) \quad (14)$$

$$P(x_b | x_a) P(x_a) = \sum_{k=1}^K \pi_k P(x_a, x_b | k) \quad (15)$$

$$P(x_b | x_a) = \frac{1}{P(x_a)} \sum_{k=1}^K \pi_k P(x_a, x_b | k) \quad (16)$$

$$P(x_b | x_a) = \frac{1}{P(x_a)} \sum_{k=1}^K \pi_k P(x_b | x_a, k) P(x_a | k) \quad (17)$$

$$P(x_b | x_a) = \sum_{k=1}^K \left(\frac{\pi_k}{P(x_a)} P(x_a | k) \right) P(x_b | x_a, k) \quad (18)$$

thus:

$$\lambda_k = \frac{\pi_k}{P(x_a)} P(x_a | k) \quad (19)$$

Problem 3. The connection between GMM and K-means

• **Question 3.1**

$$\gamma(z_{nk}) = \frac{\pi_k \exp\left(\frac{-\|x_n - \mu_k\|^2}{2\sigma^2}\right)}{\sum_j \pi_j \exp\left(\frac{-\|x_n - \mu_j\|^2}{2\sigma^2}\right)} \quad (20)$$

when $\sigma \rightarrow 0$: $\exp\left(\frac{-\|x_n - \mu_k\|^2}{2\sigma^2}\right) \rightarrow \frac{-\|x_n - \mu_k\|^2}{2\sigma^2}$, and:

$$\gamma(z_{nk}) \rightarrow \frac{\pi_k \frac{-\|x_n - \mu_k\|^2}{2\sigma^2}}{\sum_j \pi_j \frac{-\|x_n - \mu_j\|^2}{2\sigma^2}} = \frac{\pi_k (-\|x_n - \mu_k\|^2)}{\sum_j \pi_j (-\|x_n - \mu_j\|^2)} \quad (21)$$

the dominant term in the denominator is: $\operatorname{argmax}(-\|x_n - \mu_j\|^2) = \operatorname{argmin}(\|x_n - \mu_j\|^2)$.
Thus:

$$\gamma(z_{nk}) = \begin{cases} \frac{\pi_k (-\|x_n - \mu_k\|^2)}{\pi_k (-\|x_n - \mu_k\|^2)} = 1, & k = \operatorname{argmin}(\|x_n - \mu_j\|^2) = r_{nk} \\ 0, & \text{O.W.} \end{cases}$$

now we only need to show: $\log \pi_k + \log N(x_n | \mu_k, \sigma^2 I) \rightarrow c \|x_n - \mu_k\|^2$ as $\sigma \rightarrow 0$ (c is a constant).

$$\log \pi_k + \log N(x_n | \mu_k, \sigma^2 I) = \log \pi_k + \log\left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{-\|x_n - \mu_k\|^2}{2\sigma^2}\right)\right) \quad (22)$$

$$\log \pi_k + \log N(x_n | \mu_k, \sigma^2 I) = \log \pi_k + \log\left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \frac{\|x_n - \mu_k\|^2}{2\sigma^2} = -\frac{\|x_n - \mu_k\|^2}{2\sigma^2} + \text{const.} \quad (23)$$

Ignoring the constant terms, maximizing the expected complete-data log-likelihood (in GMM) comes down to maximizing the term $-\frac{\|x_n - \mu_k\|^2}{2\sigma^2}$, and since $2\sigma^2$ is the same for all classes, this maximization is equivalent to maximizing the term $-\|x_n - \mu_k\|^2$, which is equivalent to minimizing $\|x_n - \mu_k\|^2$, as in the K-means algorithm.

Problem 4. Naive Bayes

• **Question 4.1**

$$\log P(D|\theta) = \log P(X = x_1, x_2, \dots, x_N; Y = y_1, y_2, \dots, y_N | \pi_c, \mu_{cd}, \sigma_{cd}) \quad (24)$$

since observations are i.i.d, we get:

$$L = \sum_{n=1}^N [\log P(y_n = c | \pi_c) + \sum_{d=1}^D \log P(x_{nd} | y_n = c, \mu_{cd}, \sigma_{cd})] \quad (25)$$

$$L = \sum_{n=1}^N [\log \pi_{y_n} + \sum_{d=1}^D \log(\frac{1}{\sqrt{2\pi\sigma_{nd}^2}} \exp(\frac{-(x_{nd} - \mu_{nd})^2}{2\sigma_{nd}^2})] \quad (26)$$

$$L = \sum_{n=1}^N \log \pi_{y_n} + \sum_{n=1}^N \sum_{d=1}^D \log(\frac{1}{\sqrt{2\pi\sigma_{nd}^2}} \exp(\frac{-(x_{nd} - \mu_{nd})^2}{2\sigma_{nd}^2})) \quad (27)$$

$$L = \sum_{n:y_n=c}^N \log \pi_c + \sum_{c=1}^C \sum_{n:y_n=c,d} \log(\frac{1}{\sqrt{2\pi\sigma_{cd}^2}} \exp(\frac{-(x_{nd} - \mu_{cd})^2}{2\sigma_{cd}^2})) \quad (28)$$

$$L = \sum_{n:y_n=c}^N \log \pi_c - \sum_{c=1}^C \sum_{n:y_n=c,d} [\frac{1}{2} \log(2\pi\sigma_{cd}^2) + \frac{(x_{nd} - \mu_{cd})^2}{2\sigma_{cd}^2}] \quad (29)$$

• **Question 4.2**

$$L(\pi_c) = \sum_{n:y_n=c} \log \pi_c \quad s.t. \sum_{c=1}^C \pi_c = 1 \quad (30)$$

Lagrange multipliers:

$$L(\pi_c, \lambda) = \sum_{n:y_n=c} \log \pi_c - \lambda (\sum_{c=1}^C \pi_c - 1) \quad (31)$$

$$\frac{\partial L}{\partial \pi_c} = \sum_{n:y_n=c} \frac{1}{\pi_c} - \lambda = 0 \rightarrow \pi_c = \frac{1}{\lambda} \sum_{n:y_n=c} 1 \quad (32)$$

$$\frac{\partial L}{\partial \lambda} = 0 \rightarrow \sum_c \pi_c = 1 \quad (33)$$

Using 32 and 33, we get:

$$\sum_c \frac{1}{\lambda} \sum_{n:y_n=c} 1 = 1 \rightarrow \frac{1}{\lambda} \sum_c \sum_{n:y_n=c} 1 = 1 \rightarrow \lambda = \sum_c \sum_{n:y_n=c} 1 \quad (34)$$

now again using 32, we get:

$$\pi_c^* = \frac{\sum_{n:y_n=c} 1}{\sum_c \sum_{n:y_n=c} 1} = \frac{N_c}{N} \quad (35)$$

where N_c is the number of data points labeled as c.

Now,

$$L(\mu_{cd}, \sigma_{cd}) = - \sum_{c=1}^C \sum_{n:y_n=c,d} \left[\frac{1}{2} \log(2\pi\sigma_{cd}^2) + \frac{(x_{nd} - \mu_{cd})^2}{2\sigma_{cd}^2} \right] \quad (36)$$

μ_{cd} and σ_{cd} can be estimated separately for each class c:

$$L_c(\mu_{cd}, \sigma_{cd}) = - \sum_{n:y_n=c,d} \left[\frac{1}{2} \log(2\pi\sigma_{cd}^2) + \frac{(x_{nd} - \mu_{cd})^2}{2\sigma_{cd}^2} \right] \quad (37)$$

$$\frac{\partial L_c}{\partial \mu_{cd}} = \sum_{n:y_n=c,d} \frac{2(x_{nd} - \mu_{cd})}{2\sigma_{cd}^2} = 0 \rightarrow \sum_{n:y_n=c,d} x_{nd} - \mu_{cd} = 0 \quad (38)$$

defining the number of data points with label c as N_c ($N_c = \#(n : y_n = c)$) :

$$\mu_{cd}^* = \frac{\sum_{n:y_n=c,d} x_{nd}}{N_c} \quad (39)$$

Also

$$\frac{\partial L_c}{\partial \sigma_{cd}} = 0 \rightarrow \sum_{n:y_n=c,d} \frac{1}{\sigma_{cd}} \left(\frac{(x_{nd} - \mu_{cd})^2}{\sigma_{cd}^3} \right) = 0 \rightarrow \sum_{n:y_n=c,d} [\sigma_{cd}^2 - (x_{nd} - \mu_{cd})^2] = 0 \quad (40)$$

$$\rightarrow N_c \sigma_{cd}^2 - \sum_{n:y_n=c,d} (x_{nd} - \mu_{cd})^2 = 0 \quad (41)$$

$$\sigma_{cd}^{2*} = \frac{\sum_{n:y_n=c,d} (x_{nd} - \mu_{cd})^2}{N_c} \quad (42)$$