

Financial Econometrics Instrumental Variables

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Unobservable Omitted Variable



- Even if we can control all observed variable, selection bias might still exist due to **unobservable omitted variables**.
- That is, the treatment and control group may be very different in some characteristics that we CANNOT observe!
- In other words, CIA (selection on observables) is not valid!
- Thus, we CANNOT eliminate selection bias by including more covariates into regression.

Unobservable Omitted Variable



• Suppose the true model is:

$$Y_i = \delta + \alpha D_i + \beta X_i + \epsilon_i$$

- But now X_i is an unobserved characteristics (e.g. ability, preference, health).
- So we cannot include it into our regression and estimate the following model:

$$Y_i = \delta + \alpha D_i + u_i$$

where $u_i = \beta X_i + \epsilon_i$

Unobservable Omitted Variable



 As mentioned before, failure to include key covaraites will lead to omitted variables bias:

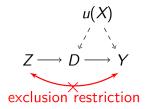
$$\hat{\alpha} \xrightarrow{\mathbf{p}} \alpha + \frac{Cov(u_i, D_i)}{Var(D_i)}$$
$$= \alpha + \beta \frac{Cov(X_i, D_i)}{Var(D_i)}$$

- Remember there is **NO** omitted variable bias if D_i is unrelated to u_i (X_i) .
 - 1 X_i is unrelated to Y_i : $\beta = 0$
 - 2 X_i is unrelated to D_i : $Cov(X_i, D_i) = 0$
- To obtain causal effect (eliminate OVB), we need a variation in D_i that is unrelated to unobserved confounding factor u_i (X_i).



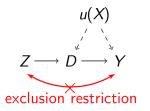
- The Instrumental Variable (IV) is an exogenous sources of variation that drives the treatment D_i but unrelated to other confounding factors u_i (X_i) that affect outcome Y_i .
- Intuitively, IV breaks variation of the treatment D_i into two parts:
 - 1 A part that might be correlated with other confounding factors u_i (X_i).
 - 2 A part that is not.
- We use the variation in D_i that is not correlated with u_i (X_i) to estimate causal effect of the treatment.





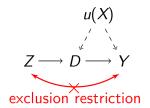
- \bullet Y is an outcome. (e.g. earnings)
- Z is the instrument.
- ullet D is the treatment. (e.g. college degree)
- ullet u(X) is the unobserved confounding factor. (e.g. ability)





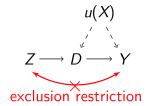
- $\begin{tabular}{ll} \textbf{Unobserved ability} \ u(X) \ \text{might confound with the effect of college} \\ \text{degree} \ D. \end{tabular}$
- We need to find an IV that generate a variation in getting college degree D that is unrelated to ability u(X).
- IV initiates a causal chain: the instrument Z affects D, which in turn affects Y.





- A valid IV needs to satisfy the following conditions:
 - 1 First-stage relationship (Instrument relevance): Z affects D.
 - 2 Exclusion restriction (Instrument exogeneity):
 - * No direct or indirect effect of the instrument Z on the outcome Y NOT through the treatment variable D.
 - ★ The instrument Z affects the outcome Y only through the treatment variable D.





- We can test whether the instrument relevance is satisfied.
- But the instrument exogeneity cannot be tested.
- You have to try to convince your audience that it is satisfied.



Joshua D. Angrist (1990)

"Lifetime Earnings and the Vietnam Era Draft Lottery: Evidence from Social Security Administrative Records", AER

- He wanted to examine the effect of military service on lifetime income.
- We will use Angrist's paper on the effects of military service (D_i) on earnings (Y_i) as an example to go through key concept of IV methods.



- Joining military service is a personal choice.
- Is there any selection bias due to unobservable confounding factors in this example?

• Time preference:

- Less patient people may voluntarily join military service early.
- This myopic thinking may have negative impact on their earnings (e.g. less human capital investment)

• Health condition:

- Better health people can join military service.
- Better health condition also have positive impact on their earnings
- We need a IV for the treatment variable of joining military service!



- Angrist (1990) uses the Vietnam draft lottery (Z_i) as in IV for military service!
- In the 1960s and early 1970s, young American men were draft for military service to serve in Vietnam.
- Concerns about the fairness of the conscription policy lead to the introduction of a draft lottery in 1970.



- From 1970 to 1972 random sequence numbers were assigned to each birth date in cohorts of 19-year-olds.
- Men with lottery numbers below a cutoff were drafted while men with numbers above the cutoff could not be drafted.
- The draft did NOT perfectly determinate military service:
 - Many draft-eligible men were exempted for health and other reasons.
 - Draft-ineligible men volunteered for service.
- Next, let's briefly discuss whether draft eligibility induced by lottery is a good IV or not.



- First-stage relationship (Instrument Relevance): Z_i affects D_i
 - Vietnam veteran status (joining military service) was not completely determined by randomized draft eligibility.
 - ▶ But draft eligibility is highly correlated with Vietnam veteran status.
- Exclusion restriction (Instrument Exogeneity):
 - ▶ The draft eligibility is determined by random numbers.
 - This should not affect one's earnings directly.



Treatment Assignment

$$Z_i = \begin{cases} 1 & \text{if an individual } i \text{ is eligible for a treatment} \\ 0 & \text{if an individual } i \text{ is not eligible for a treatment} \end{cases}$$

- $ightharpoonup Z_i = 1$: those who get draft eligibility (due to lottery results).
- ▶ $Z_i = 0$: those who do not get draft eligibility (due to lottery results).



- Potential Treatments
 - ▶ D_i^z : treatment status given the value of Z.
 - ▶ D_i^1 : treatment status if eligible for a treatment.
 - D_i^0 : treatment status if not eligible for a treatment.
- Observed Treatments

$$D_i = \begin{cases} D_i^1 & \text{if } Z_i = 1\\ D_i^0 & \text{if } Z_i = 0 \end{cases}$$

ullet or in a more compact notation: $D_i=Z_iD_i^1+(1-Z_i)D_i^0$



- **Exclusion Restriction**: the instrument has no direct effect on the outcome, once we fix the value of the treatment.
- Given the exclusion restriction, we know that the potential outcomes for each treatment status only depend on the treatment D_i , not the instrument Z_i .



Potential Outcomes

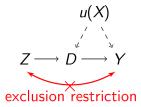
- Y_i^1 : outcome if an individual i get treatment (either $D_i^1 = 1$ or $D_i^0 = 1$).
- ▶ Y_i^0 : outcome if an individual i does not get treatment (either $D_i^1 = 0$ or $D_i^0 = 0$)

Observed Outcomes

$$Y_i = \begin{cases} Y_i^1 & \text{if } D_i^1 = 1 \text{ or } D_i^0 = 1 \\ Y_i^0 & \text{if } D_i^1 = 0 \text{ or } D_i^0 = 0 \end{cases}$$

 \bullet or in a more compact notation: $Y_i = D_i^z Y_i^1 + (1-D_i^z) Y_i^0$





- The IV method characterize a causal chain reaction leading from the instrument Z_i (draft eligibility) to outcome Y_i (earnings).
- Intuitively:

Effect of instrument on outcome

- = (Effect of instrument on treatment)
- × (Effect of treatment on outcome)



• Rearranging:

 $= \frac{\text{Effect of instrument on outcome}}{\text{Effect of instrument on treatment}}$

Formal representation:

$$\alpha_{IV} = \frac{E[Y_i|Z_i = 1] - E[Y_i|Z_i = 0]}{E[D_i|Z_i = 1] - E[D_i|Z_i = 0]}$$



 By using the following identification assumptions, we can prove the causal effect that IV identify is a local average treatment effect (LATE).

Assumption (First-Stage Relationship)

$$E[D_i|Z_i = 1] - E[D_i|Z_i = 0] \neq 0$$

• We need the instrument to have a significant effect on the treatment.



Assumption (Independent Assumption)

$$(Y_i^1, Y_i^0, D_i^1, D_i^0) \perp Z_i$$

- The IV is independent of potential outcomes and potential treatment (i.e. as good as randomly assigned).
- The independence assumption is sufficient for a causal interpretation of the reduced form:

$$E[Y_i|Z_i = 1] - E[Y_i|Z_i = 0] = E[Y_i(D_i^1)|Z_i = 1] - E[Y_i(D_i^0)|Z_i = 0]$$
$$= E[Y_i(D_i^1)] - E[Y_i(D_i^0)]$$

• Independence also means that the first stage captures the causal effect of Z_i on D_i :

$$E[D_i|Z_i = 1] - E[D_i|Z_i = 0] = E[D_i^1|Z_i = 1] - E[D_i^0|Z_i = 0]$$
$$= E[D_i^1 - D_i^0]$$



Assumption (Exclusion Restriction)

$$Y_i^d(Z=1) = Y_i^d(Z=0) = Y_i^d$$

- The potential outcomes for each treatment status only depend on the treatment D_i , not the instrument Z_i .
- Vietnam draft lottery example:
 - An individual's earnings potential as a veteran (join military service) or non-veteran (not join military service) are assumed to be unchanged by draft eligibility status.



Assumption (Exclusion Restriction)

$$Y_i^d(Z=1) = Y_i^d(Z=0) = Y_i^d$$

- The exclusion restriction would be violated if low lottery numbers may have affected schooling (e.g. to avoid the draft).
- If this was the case the lottery number would be correlated with earnings for at least two cases:
 - 1 through its effect on military service.
 - 2 through its effect on educational attainment.
- The fact that the lottery number is randomly assigned (and therefore satisfies the independence assumption) does NOT ensure that the exclusion restriction is satisfies!

IV and Compliers



- The variation in treatment D_i (veteran status) was not entirely from the draft eligibility Z_i but also from individual choice.
- Thus, $D_i^1 = 1$ or $D_i^1 = 0$?
 - D_i¹ = 1: Those who get draft eligibility choose to join military service.
 D_i¹ = 0: Those who get draft eligibility choose NOT to join.
- Similarly, $D_i^0 = 1$ or $D_i^0 = 0$?

 - $\begin{array}{l} \blacktriangleright \ D_i^0=1 \hbox{: Those who did NOT get draft eligibility choose to join.} \\ \blacktriangleright \ D_i^0=0 \hbox{: Those who did NOT get draft eligibility choose NOT to join.} \end{array}$

IV and Compliers



- We can define four types of individuals based on whether they fellow the draft eligibility results:
 - 1 Compliers: $D_i^1 > D_i^0 \ (D_i^1 = 1, D_i^0 = 0)$
 - ★ David got draft eligibility and joined military service.
 - ★ John did not get draft eligibility and did not join military service.
 - 2 Always Takers: $D_i^1 = D_i^0 = 1$
 - Steve always joined military service no matter the lottery results (whether he got draft eligibility).
 - 3 Never Takers: $D_i^1 = D_i^0 = 0$
 - Trump never joined military service no matter the lottery results (whether he got draft eligibility).
 - 4 Defiers: $D_i^1 < D_i^0 \ (D_i^1 = 0, D_i^0 = 1)$
 - ★ Jimmy got draft eligibility but did NOT join military service.
 - ★ Jonson did NOT get draft eligibility but joined military service.



Assumption (Monotonicity Assumption)

$$D_i^1 > D_i^0$$

- Lastly, we need to make another assumption about the relationship between the instrument and the treatment.
- Monotonicity says that the presence of the instrument never dissuades someone from taking the treatment.
- This is sometimes called **no defiers**!
- In the draft lottery example: draft eligibility should encourage people to join military service.



Theorem (IV Identify LATE)

$$\alpha_{IV} = \frac{E[Y_i|Z_i=1] - E[Y_i|Z_i=0]}{E[D_i|Z_i=1] - E[D_i|Z_i=0]} = E[Y_i^1 - Y_i^0|D_i^1 > D_i^0]$$

- IV estimator represents the causal effect for **compliers**!
- Lottery IV can identify the causal effect of military service on lifetime earnings for those who obey the lottery results (e.g. David and John).



Proof:

$$\alpha_{IV} = \frac{E[Y_i|Z_i = 1] - E[Y_i|Z_i = 0]}{E[D_i|Z_i = 1] - E[D_i|Z_i = 0]}$$

$$= \frac{E[D_i^1Y_i^1 + (1 - D_i^1)Y_i^0|Z_i = 1] - E[D_i^0Y_i^1 + (1 - D_i^0)Y_i^0|Z_i = 0]}{E[D_i^1|Z_i = 1] - E[D_i^0|Z_i = 0]}$$

$$= \frac{E[D_i^1Y_i^1 + (1 - D_i^1)Y_i^0] - E[D_i^0Y_i^1 + (1 - D_i^0)Y_i^0]}{E[D_i^1] - E[D_i^0]}$$

$$= \frac{E[(Y_i^1 - Y_i^0)(D_i^1 - D_i^0)]}{E[D_i^1] - E[D_i^0]}$$

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- Note that since D_i^z is a dummy,
- IV estimates cannot say anything about causal effect for always takers or never takers: $D_i^1 D_i^0 = 0$.
- $D_i^1 D_i^0 = 1$: compliers or $D_i^1 D_i^0 = -1$: defiers?
- Using Monotonicity Assumption, only $D_i^1 D_i^0 = 1$ exists.
- Therefore, $E[(Y_i^1 Y_i^0)(D_i^1 D_i^0)]$ can become the following terms:

$$\begin{split} E[(Y_i^1 - Y_i^0)(D_i^1 - D_i^0)] \\ &= E[(Y_i^1 - Y_i^0) \times (1)|D_i^1 - D_i^0 = 1]Pr(D_i^1 - D_i^0 = 1) \\ &+ E[(Y_i^1 - Y_i^0) \times (-1)|D_i^1 - D_i^0 = -1]Pr(D_i^1 - D_i^0 = -1) \\ &= E[(Y_i^1 - Y_i^0) \times (1)|D_i^1 - D_i^0 = 1]Pr(D_i^1 - D_i^0 = 1) \end{split}$$

• Note: $E[D_i^1] - E[D_i^0] = Pr(D_i^1 - D_i^0 = 1)$





Continue the Proof:

$$\begin{split} \alpha_{IV} &= \frac{E[Y_i|Z_i=1] - E[Y_i|Z_i=0]}{E[D_i|Z_i=1] - E[D_i|Z_i=0]} \\ &= \frac{E[D_i^1Y_i^1 + (1-D_i^1)Y_i^0|Z_i=1] - E[D_i^0Y_i^1 + (1-D_i^0)Y_i^0|Z_i=0]}{E[D_i^1|Z_i=1] - E[D_i^0|Z_i=0]} \\ &= \frac{E[D_i^1Y_i^1 + (1-D_i^1)Y_i^0] - E[D_i^0Y_i^1 + (1-D_i^0)Y_i^0]}{E[D_i^1] - E[D_i^0]} \\ &= \frac{E[(Y_i^1 - Y_i^0)(D_i^1 - D_i^0)]}{E[D_i^1] - E[D_i^0]} \\ &= \frac{E[(Y_i^1 - Y_i^0) \times (1)|D_i^1 - D_i^0 = 1]Pr(D_i^1 - D_i^0 = 1)}{Pr(D_i^1 - D_i^0 = 1)} \\ &= E[(Y_i^1 - Y_i^0)|D_i^1 > D_i^0] = \alpha_{LATE} \end{split}$$



- Never takers and always takers do NOT change their treatment status when the instrument gets switched on and off.
 - So only defiers and compliers contribute to IV estimate.
 - ▶ IV estimate is the sum of those two effects.
- By using monotonicity assumption, we rule out the effect from defiers.
- Therefore, IV estimates the average treatment effect for compliers.



LATE

- $\alpha_{LATE} = E[Y_i^1 Y_i^0 | D_i^1 > D_i^0]$, the Local Average Treatment Effect for compliers.
- ATE for the individuals whose treatment status (join military service) are changed by the instrument (lottery draft).
- LATE (α_{LATE}) is different when using different instruments, Z_i .
- Whether LATE is interesting or not depends on the instrument.



- Without further assumptions (e.g. constant causal effects), LATE is not informative about effects on never-takers or always-takers.
 - → because the instrument does not affect their treatment status.
- In most applications we would be mostly interested in estimating the average treatment effect on the whole population (ATE).

$$= E[Y_i^1 - Y_i^0]$$

This is usually not possible with IV.



Special Case

- If D_i is randomized (e.g. RCT) and everybody is a complier, then $Z_i = D_i$.
- That is, no never taker or always taker.
- One-sided noncompliance, $D_i^0 = 0$, then:

$$\begin{split} E[Y_i^1 - Y_i^0 | D_i^1 > D_i^0] &= E[Y_i^1 - Y_i^0 | D_i^1 = 1] \\ &= E[Y_i^1 - Y_i^0 | Z_i = 1, D_i^1 = 1] \\ &= E[Y_i^1 - Y_i^0 | D_i = 1] \\ &= E[Y_i^1 - Y_i^0] \end{split}$$

• Thus, $\alpha_{LATE} = \alpha_{ATT} = \alpha_{ATE}!$

IV Estimation



• Causal relationship of interest: the effect of military service on earnings

$$Y_i = \delta + \alpha_{IV} D_i + u_i$$

Remember, we just derive:

 $\mbox{Effect of treatment on outcome} = \frac{\mbox{Effect of instrument on outcome}}{\mbox{Effect of instrument on treatment}}$

IV Estimation: Intuition



- We can estimate α_{IV} by running the following two regressions:
- 1 Reduced Form regression: the effect of lottery draft on earnings

$$Y_i = \mu + \alpha_{RF} Z_i + \epsilon_i$$
$$\alpha_{RF} = \frac{Cov(Y_i, Z_i)}{Var(Z_i)}$$

2 First-Stage regression: the effect of lottery draft on military service

$$D_{i} = \tau + \alpha_{FS}Z_{i} + \eta_{i}$$

$$\alpha_{FS} = \frac{Cov(D_{i}, Z_{i})}{Var(Z_{i})}$$

• The IV estimator is:

$$\widehat{\alpha_{IV}} = \frac{\widehat{\alpha_{RF}}}{\widehat{\alpha_{FS}}} = \frac{\widehat{Cov}(Y_i, Z_i)}{\widehat{Cov}(D_i, Z_i)}$$

IV Estimation: 2-Stage Least Square



- In practice we often estimate IV using Two stage least squares estimation (2SLS).
- If identification assumptions only hold after conditioning on X, covariates are often introduced using TSLS regression.
- It is called 2SLS because you could estimate it as follows:
 - 1 Obtain the first stage fitted values:

$$\widehat{D_i} = \widehat{\tau} + \widehat{\alpha_{FS}} Z_i + X_i' \widehat{\beta}$$

2 Plug the first stage fitted values into the second-stage equation

$$Y_i = X_i' \gamma + \alpha_{2SLS} \widehat{D}_i + u_i^*$$

• The intuition: 2SLS only retains the variation in D_i that is generated by quasi-experimental variation Z_i (and thus hopefully exogenous).

IV Remarks



1 Check IV relevance!

- ▶ Does this IV make sense?
- Do the coefficients have the right magnitude and sign?
- Report the F-statistic in the first stage regression.
 - ★ If F > 10, instruments are strong use 2SLS
 - ★ If F < 10, weak instruments find better IV
- If instruments are weak, then the 2SLS estimator is biased and the t-statistic has a non-normal distribution.

Weak IV Problem



• What are the consequences of weak instruments?

$$\widehat{\alpha_{2SLS}} = \frac{\widehat{Cov}(Y_i, Z_i)}{\widehat{Cov}(D_i, Z_i)}$$

- If $Cov(D_i,Z_i)$ is small (close to zero), a small change in $\widehat{Cov}(Y_i,Z_i)$ due to finite sample perturbation can induce a huge change in $\widehat{\alpha_{2SLS}}$!
- Under this situation, the normal distribution is a poor approximation to the sampling distribution of $\widehat{\alpha_{2SLS}}$.
- If instruments are weak, the usual methods of inference are unreliable.

Weak IV Problem



Weak IV Bias

$$\begin{aligned} plim \ \widehat{\alpha_{OLS}} &= \alpha + \frac{\sigma_{D,u}}{\sigma_D^2} \\ plim \ \widehat{\alpha_{2SLS}} &= \alpha + \frac{\sigma_{\hat{D},u}}{\sigma_{\hat{D}}^2} \end{aligned}$$

Thus,

$$\frac{plim \ \widehat{\alpha_{2SLS}} - \alpha}{plim \ \widehat{\alpha_{OLS}} - \alpha} = \frac{\sigma_{D,u}/\sigma_{\widehat{D},u}}{R_{D,Z}^2} = \frac{\rho_{Z,u}/\rho_{D,u}}{\rho_{D,Z}}$$

• If the instrument is weak, the bias for the 2SLS estimate may be larger than that for the OLS estimate.

IV Remarks



- 2 Check exclusion restriction!
 - ▶ The exclusion restriction cannot be tested directly, but it can be falsified.
 - Placebo test:
 - * Test the reduced form effect of Z_i on Y_i in situations where it is impossible or extremely unlikely that Z_i could affect D_i .
 - * Because Z_i cannot affect D_i , then the exclusion restriction implies that this placebo test should have zero effect.
- 3 If you have many IVs pick your best instrument and report the *just-identified* model.
- 4 Study at the reduced form.
 - lacktriangle Directly estimate the effect of instrument Z on outcome Y
 - If you cannot see the causal relationship of interest in the reduced form it is probably not there.

1 First stage results:

- Having a low lottery number (being eligible for the draft) increases veteran status by about 16 percentage points.
- ▶ Note that the mean of veteran status is about 27 percent.

2 Second stage results:

► Serving in the army lowers earnings by \$2,050 - \$2,741 per year.

3 Placebo test:

- ▶ There is no evidence of an association between draft eligibility (having a low lottery number) and earnings in 1969.
- ▶ Note that 1969 earnings are realized before the 1970 draft lottery.