Name: Example Solutions

STAT 535 — Intro to Bayesian Data Analysis Test 1 — Spring 2022

- 1. Fill in the blanks: Bayesian posterior inference is based on a combination of prior information and data (sample) information. As the sample size increases, the data (likelihood) gets more heavily weighted in the posterior inference.
- 2. Suppose we have two iid observations  $Y_1, Y_2$  that follow a gamma distribution with shape parameter known to be 3 and unknown rate parameter  $\theta > 0$ , i.e., having density

$$f(y|\theta)=\frac{\theta^3}{\Gamma(3)}y^{3-1}e^{-\theta y},\ y>0.$$

The analyst wishes to perform inference about the unknown parameter  $\theta > 0$ .

(a) Write and simplify the likelihood function based on  $Y_1, Y_2$ .

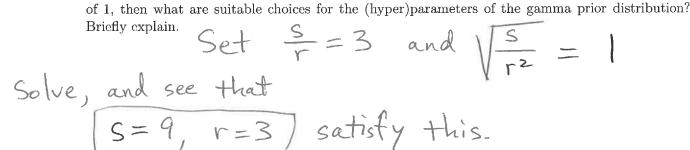
$$L(\theta|y) = \frac{2}{\Gamma(3)} \frac{\theta^{3}}{y_{i}^{2}} e^{-\theta y_{i}^{2}}$$

$$= \left(\frac{\theta^{3}}{\Gamma(3)}\right)^{2} \frac{2}{\Gamma(3)} \frac{2}{\Gamma(3)} e^{-\theta \sum_{i=1}^{2} y_{i}^{2}}$$

$$= \left(\frac{\theta^{3}}{\Gamma(3)}\right)^{2} \frac{2}{\Gamma(3)} e^{-\theta \sum_{i=1}^{2} y_{i}^{2}}$$

(b) Suppose the analyst chooses a gamma prior distribution with shape parameter s and rate parameter r (given on formula sheet). Explain why this choice is sensible, with respect to the support of  $\theta$ .

Since  $\theta > 0$ , the gamma distribution is a good choice of prior since it has support on  $(0, \infty)$ .



(d) Under this likelihood and prior, derive the posterior distribution for  $\theta$  (enough to identify it, including its parameter value(s)), given  $Y_1, Y_2$ .

(c) If the analyst believes, before examining the data, that  $\theta$  has a mean of 3 and a standard deviation

$$P(\theta|y) \propto P(\theta) L(\theta|y)$$

$$C \theta^{9-1} e^{-3\theta} \theta^{6} e^{-\theta \tilde{\Sigma}_{i}} y_{i}$$

$$= \theta^{14} e^{-\theta} (\tilde{\Sigma}_{i} y_{i} + 3)$$
which is gamma (.15),  $\tilde{\Sigma}_{i} y_{i} + 3$ )

(e) Write an expression for the mean of the posterior distribution, and then express it explicitly as a linear combination of the prior mean and the MLE of  $\theta$ , which is  $\hat{\theta}_{ML} = 3/\bar{Y}$ . [This one is slightly challenging (though *not* long and tedious), you may want to save it for the end.]

Posterior mean = 
$$\frac{15}{\overline{\Sigma}y_{i}+3} = \frac{7.5}{\overline{y}+1.5}$$

which can be written as
$$(3)(\frac{1.5}{\overline{y}+1.5}) + (\frac{3}{\overline{y}})(\frac{\overline{y}}{\overline{y}+1.5})$$

prior mean  $\frac{3}{\overline{\Sigma}y_{i}+3} + \frac{3}{\overline{y}}(\frac{\overline{y}}{\overline{y}+1.5})$ 
 $\frac{3}{\overline{\Sigma}y_{i}+3} + \frac{3}{\overline{y}}(\frac{\overline{\Sigma}y_{i}+3}) + \frac{3}{\overline{\Sigma}y_{i}+3} + \frac{3}{\overline{\Sigma}y_{i}+3}$ 

3. Suppose we have iid observations  $Y_1, \ldots, Y_n$  that follow a distribution with pdf:

$$p(y|\theta) = (y-1)\theta^2(1-\theta)^{y-2}, y = 2, 3, 4, \dots$$

where the unknown parameter is  $0 < \theta < 1$ . In this distribution,  $\theta$  represents a success probability.

(a) Suppose you choose as a prior distribution for  $\theta$  a beta(1,3) distribution. Explain briefly why the beta is a reasonable choice as a prior here. Also, briefly explain in words what your specific choice of this beta distribution implies about your prior belief about the success probability.

The beta is a good choice because it has support on (0,1). We believe the success probability is around  $\frac{1}{1+3} = 0.25$ .

(b) Write (and simplify as much as possible) the likelihood function  $L(\theta|y_1,\ldots,y_n)$ .

$$L(\Theta|\mathcal{Y}) = \prod_{i=1}^{n} p(y_i|\Theta) = \prod_{i=1}^{n} [(y_i-1)\Theta^2(1-\Theta)^{y_i-2}]$$

$$= \left[\prod_{i=1}^{n} (y_i-1)\right] \Theta^{2n}(1-\Theta)^{\sum_{i=1}^{n} y_i-2n}$$

(c) Based on your prior distribution and the likelihood here, derive the posterior distribution for  $\theta$ (enough to identify it specifically, including its parameter value(s)).

$$P(\theta|y) \propto P(\theta) L(\theta|y)$$

$$\propto \theta^{|-1|} (1-\theta)^{3-1} \theta^{2n} (1-\theta)^{2y} - 2n$$

$$= \theta^{2n+1-1} (1-\theta)^{2y} - 2n+3-1 \quad \text{which is}$$
beta  $(2n+1) \sum y_i - 2n+3$ 

$$= (d) \text{ If we observe sample values of } 9, 4, 7, 3, 5, \text{ then name the posterior distribution for } \theta, \text{ specifying actual numerical parameter values.}$$

$$n = 5, \quad 5y_i = 28$$

n=5, Sy; =28 actual numerical parameter values.

(e) Based on what you know about the form of the posterior distribution here, give a Bayesian point estimate (a number) for  $\theta$  using your posterior.

$$\hat{\Theta}_{B} = \frac{11}{11+21} = \frac{11}{32} = 0.34375$$

(f) Note that  $E(Y) = 2/\theta$  in this model. What is a Bayesian point estimate of E(Y), given this data set?

$$\widehat{E(Y)} = \frac{2}{\frac{1}{32}} = \frac{64}{11} = 5.82$$

4. My wife has recently gotten into the game Wordle. Mrs. Hitchcock would like to estimate the probability that she is able to solve a randomly selected Wordle puzzle. Inspired by her devoted and, in general, awesome husband, she decides to adopt a Bayesian approach to inference. She specifies a beta prior for the unknown probability of success, and then she tries 10 puzzles (assume the outcomes are independent) and observes how many she is able to solve. A graph of the prior, (scaled) likelihood, and posterior is shown below.

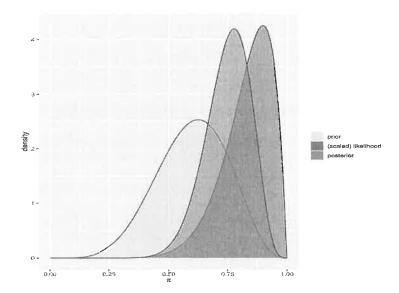


Figure 1: Prior, scaled likelihood, posterior (not necessarily in that order).

- (a) Based on the plot, what is the prior that Mrs. Hitchcock has chosen? [It is one of these listed in the answer choices.]
- (A) beta(2,2) (B) beta(6,4) (C) beta(4,6) (D) beta(1,9)
- (b) Based on the plot, what was the observed number of successes for Mrs. Hitchcock in the 10 attempts?

(A) y = 9 (B) y = 7 (C) y = 6 (D) y = 5

(c) Using your answers to parts (a) and (b), what is a point estimate for her probability of solving a random Wordle puzzle? Indicate how you got your answer.

Posterior is beta 
$$(6+9,4+1) = beta (15,5)$$
  
Posterior mean is  $\frac{15}{15+5} = 0.75$ 

(d) In a sentence or two, discuss how her prior beliefs have been updated into the posterior information. She originally believed she had a bout a 60% chance of success. After seeing the data, she estimates she has a 75% chance of success.

- 5. Suppose for a certain company, 40% of all bolts are made in Boston and the rest are made in Chicago. We also know that 4% of bolts made in Boston are defective and 2% of bolts made in Chicago are defective.
  - (a) What is the probability that a randomly selected bolt is defective? Show work.

Let 
$$A = \text{defective}$$
,  $B = \text{Boston}$   
 $P(A) = P(B)P(A|B) + P(B^c)P(A|B^c)$   
 $= (.40)(.04) + (.60)(.02)$   
 $= [.028]$ 

(b) If a randomly selected bolt is found to be defective, then use Bayes' Rule to find the probability that it was made in Boston. Show work.

$$P(B|A) = P(A|B)P(B) P(A|B)P(B) + P(A|B^{c})P(B^{c}) = \frac{(.04)(.40)}{.028}$$

$$= 0.5714$$

- 6. A random sample of 24 reaction times (in seconds) of world-class sprinters from the 1996 Olympics was taken. Some R output giving the results of the analysis is given in Appendix 1 at the end of the exam. It was assumed that the reaction times were distributed  $N(\mu, \sigma^2)$ , and it was of interest to make inference about the mean reaction time of Olympic sprinters.
  - (a) What is the prior distribution for  $\sigma^2$  (specifying actual numerical parameter values)?

(b) What is the posterior distribution for  $\sigma^2$  given y (specifying actual numerical parameter values)? [Requires a bit of calculator work, but the needed numbers are all in the output.]

Inv. Gamma 
$$(\alpha + \frac{\pi}{2} - 0.5, \beta + \frac{1}{2}(\Sigma y_i^2 - n y^2))$$
  
Inv. Gamma  $(66 + 12 - 0.5, 0.026 + \frac{1}{2}(.685671 - 24(.167958^2))$   
 $\Rightarrow$  Inv. Gamma  $(77.5, 0.0303)$ 

(c) Note that the prior for  $\mu$  was:

$$\mu | \sigma^2 \sim N(\delta, \sigma^2/s_0)$$

Refer to the R code and the form of the posterior for  $\mu$ . Comment on exactly how our prior beliefs about  $\mu$  have been altered by observing the sample data, making reference to the values for the prior mean and posterior mean for  $\mu$ .

Our prior belief was that  $\mu$  was around S=0.18, After seeing the data, our posterior estimate for  $\mu$  has lowered to 0.1684 seconds.

- (d) What do the two given point estimates for  $\mu$  indicate about the symmetry/skewness of its posterior distribution? (Choose the best answer.)
- (A) The posterior for  $\mu$  is skewed, and a 95% (equal-tailed) quantile-based credible interval for  $\mu$  would be the same as a 95% HPD credible interval for  $\mu$ .

(B) The posterior for  $\mu$  is symmetric, and a 95% (equal-tailed) quantile-based credible interval for  $\mu$  would be the same as a 95% HPD credible interval for  $\mu$ .

- (C) The posterior for  $\mu$  is skewed, and a 95% (equal-tailed) quantile-based credible interval for  $\mu$  would be different than a 95% HPD credible interval for  $\mu$ .
- note the (D) The posterior for  $\mu$  is symmetric, and a 95% (equal-tailed) quantile-based credible interval for posterior  $\mu$  would be different than a 95% HPD credible interval for  $\mu$ .

mean and median are the same  $\Rightarrow$  Symmetric posterior.

(e) In a sentence, briefly but carefully interpret (in the context of the variable in the study) the given 95% credible interval for  $\mu$ .

With posterior probability 0.95, the mean reaction time for Olympic sprinters is between 0.1606 and 0.1763 seconds.

## Appendix 1

```
> react <- c(.187,.152,.137,.175,.172,.165,.184,.185,.147,.189,.172,.156,.168,.140,.214,.163,.202,.173,.175,.154,.160,.169,.148,
> y<- react
> ybar <- mean(y); n <- length(y); my.s <- sd(y)
> # prior parameters
> my.alpha <- 66; my.beta <- 0.026
> # prior parameters
> my.delta <- 0.18; s0 <- 1
> ybar
[1] 0.1679583
> n
[1] 24
> sum(y^2)
[1] 0.685671
> my.s
[1] 0.0193716
> library(pscl) # loading pscl package
> library(TeachingDemos) # loading TeachingDemos package
> ### Point estimates for sigma^2:
> p.mean.sig.sq <- (my.beta + 0.5*(sum(y^2) - n*(ybar^2)) ) / (my.alpha + n/2 - 0.5 - 1)
> p.median.sig.sq <- qigamma(0.50, my.alpha + n/2 ~ 0.5, my.beta + 0.5*( sum(y^2) - n*(ybar^2) ) )
> print(paste("posterior.mean for sigma^2=", round(p.mean.sig.sq,4),
        "posterior.median for sigma^2=", round(p.median.sig.sq,4) ))
[1] "posterior.mean for sigma^2= 0.0004 posterior.median for sigma^2= 0.0004"
> ### Marginal Interval estimate for sigma^2:
> hpd.95.sig.sq <- hpd(qigamma, alpha=my.alpha + n/2 - 0.5, beta=my.beta + 0.5*( sum(y^2) - n*(ybar^2) ) )
> hpd.95.sig.sq
[1] 0.0003111014 0.0004875030
> #### AN APPROACH FOR INFERENCE ABOUT mu:
> # Randomly sample many values for the posterior of sigma^2:
> sig.sq.values <- rigamma(n=1000000,alpha=my.alpha + n/2 - 0.5, beta=my.beta + 0.5*( sum(y^2) - n*(ybar^2) ) )
> # Randomly sample many values from the posterior of mu, GIVEN the sampled values of sigma^2 above:
> mu.values <- rnorm(n=1000000,mean=((sum(y)+my.delta*s0)/(n+s0)), sd=sqrt(sig.sq.values/(n+s0)))
> # Point estimates for mu:
> print(paste("posterior.mean for mu=", round(mean(mu.values),4),
        "posterior.median for mu=", round(median(mu.values),4) ))
[1] "posterior.mean for mu= 0.1684 posterior.median for mu= 0.1684"
> # 95% HPD interval estimate for mu:
> round(emp.hpd(mu.values),4)
[1] 0.1606 0.1763
```