

The background of the slide features a complex, abstract pattern of thin red lines and circles. These lines and circles are interconnected, creating a web-like structure that covers the entire slide. The lines vary in thickness and orientation, while the circles are of different sizes and are often centered around specific points where lines intersect.

Introduction to Generalized p -value

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Content

Motivation

Univariate B-F Problem

Multivariate Case

Functional Case (Not covered today)

Remarks

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Motivation

Univariate B-F Problem

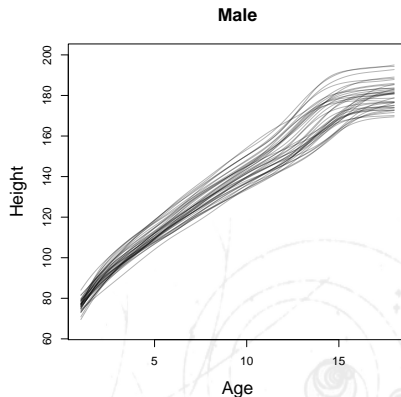
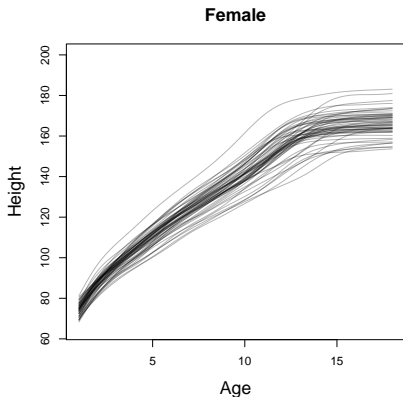
Multivariate Case

Functional Case (Not covered today)

Remarks

Comparing Two Groups of Curves

- Is there any difference between the two sets of curves?



Hypothesis Testing

- You are always “safe” to say, THERE IS, but ...
- How can we measure the confidence of our conclusion?

Hypothesis Testing

- You are always “safe” to say, THERE IS, but ...
- How can we measure the confidence of our conclusion?
- A typical statistical inference problem – Hypothesis Testing
- p -value is a widely used

Challenges

- Not easy to find a suitable test statistic
- Need to study the sampling distribution of the statistics
- Often involves many unknown parameters
- Consequence: Exact solutions only to limited number of problems

Challenges

- Not easy to find a suitable test statistic
- Need to study the sampling distribution of the statistics
- Often involves many unknown parameters
- Consequence: Exact solutions only to limited number of problems
- Reason: Many restrictions on the test statistic
- Generalized p -value relaxes some of the limitations

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Two Sample Mean Test

- The example I mentioned the beginning is similar to a two-sample t test
- We've collected two samples
 - $X_1, X_2, \dots, X_m \stackrel{iid}{\sim} N(\mu_1, \sigma_1^2)$
 - $Y_1, Y_2, \dots, Y_n \stackrel{iid}{\sim} N(\mu_2, \sigma_2^2)$, independent of X_i
- Test problem

$$H_0 : \theta = \mu_1 - \mu_2 = 0 \leftrightarrow H_a : \theta \neq 0$$

Two Sample Mean Test

- If we assume $\sigma_1^2 = \sigma_2^2$, it reduces to a problem we are familiar with

$$T = \frac{\bar{X} - \bar{Y}}{S_p \sqrt{\frac{1}{m} + \frac{1}{n}}}, \quad S_p^2 = \frac{(m-1)S_X^2 + (n-1)S_Y^2}{m+n-2}$$

- Under H_0 , $T \sim t(m+n-2)$, so we reject null hypothesis if

$$|T| > t_{1-\alpha/2}(m+n-2)$$

- Equivalently, we calculate the p -value as

$$p = P(|T(X, Y)| > |T(x, y)|)$$

- Now, what if we drop the assumption of equal variance?

Behrens-Fisher Problem

- This is often referred to as the Behrens-Fisher Problem
- In this case the distribution of T depends on the “nuisance” parameters σ_1^2 and σ_2^2
- As a result, we cannot calculate the p -value $P(|T(X, Y)| > |T(x, y)|)$ exactly
- Some approximate solutions
 - Welch’s approximation t
 - Likelihood ratio test
 - See Das Gupta (2008) and Wikipedia
http://en.wikipedia.org/wiki/Behrens-Fisher_problem

Review of Test Statistic

- In hypothesis testing, we need to construct a test statistic T which satisfies the following properties
 - $T(\cdot)$ is a function only of sample X , so that T doesn't depend on unknown parameters
 - As a result, when plugging observed value of X into $T(\cdot)$, $T(x)$ is a constant (i.e., non-random)
 - Under H_0 , the distribution of T is free of unknown parameters
 - $P(T > t|\theta)$ is nondecreasing in θ
- If satisfied, then the p -value $p = P(T(X) > T(x))$ can be calculated exactly

Generalized Test Variable

- We relax the first requirement, so that T can be a function of the **random sample** X , **observed value** x , and **nuisance parameter** η , written as $T(X; x, \eta)$. We now call $T(X; x, \eta)$ a **generalized test variable**.
- The other three requirements still hold
 - $T(x; x, \eta)$ is a constant, non-random and free of $\xi = (\theta, \eta)$
 - Under H_0 , the distribution of T doesn't rely on ξ either
 - $P(T > t|\theta)$ is nondecreasing in θ for fixed x and η
- If we've found such a $T(X; x, \eta)$, then the **generalized p -value** can be calculated as

$$p = P(T(X; x, \eta) > T(x; x, \eta))$$

which actually doesn't depend on the nuisance parameter η

Solution to B-F Problem by G- p -value

- Consider the following generalized test variable with $\eta = (\sigma_1^2, \sigma_2^2)$

$$T(X, Y; x, y, \eta) = \frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} \cdot \sqrt{\frac{\sigma_1^2 s_1^2}{m S_1^2} + \frac{\sigma_2^2 s_2^2}{m S_2^2}}$$

- We can verify
 - $T(x, y; x, y, \eta) = \bar{x} - \bar{y}$
 - Under $H_0 : \mu_1 = \mu_2$, the distribution of $T(X, Y; x, y, \eta)$ doesn't depend on $\theta = \mu_1 - \mu_2$ or $\eta = (\sigma_1^2, \sigma_2^2)$ (next slide)

Solution to B-F Problem by G-p-value

- We know the following facts

$$\bar{X} - \bar{Y} \sim N\left(\mu_1 - \mu_2, \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}\right)$$

$$\frac{(m-1)S_1^2}{\sigma_1^2} \sim \chi_{m-1}^2, \quad \frac{(m-1)S_2^2}{\sigma_2^2} \sim \chi_{n-1}^2$$

- In the expression of T , given $\mu_1 = \mu_2$,

$$\frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} \sim N(0, 1), \quad \frac{\sigma_1^2 s_1^2}{m S_1^2} \sim \frac{m-1}{m} \frac{s_1^2}{\chi_{m-1}^2}, \quad \frac{\sigma_2^2 s_2^2}{n S_2^2} \sim \frac{n-1}{n} \frac{s_1^2}{\chi_{n-1}^2}$$

Computation

- To compute the generalized p -value, simply simulate random numbers of T to approximate the probability

$$p = P(|T| > |\bar{x} - \bar{y}|)$$

- There is also an efficient method using numerical integration, see Tsui and Weerahandi (1989)

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Multivariate B-F Problems

- Now we move to multivariate case
- $X_1, X_2, \dots, X_m \stackrel{iid}{\sim} N_d(\mu_1, \Sigma_1)$
- $Y_1, Y_2, \dots, Y_n \stackrel{iid}{\sim} N_d(\mu_2, \Sigma_2)$, independent of X_i
- Test problem

$$H_0 : \theta = \mu_1 - \mu_2 = 0 \leftrightarrow H_a : \theta \neq 0$$

Solution by G- p -value

- More tricky and complicated
- The key idea is to generalize the following univariate test variable to multivariate scenario

$$T^2(X, Y; x, y, \eta) = \frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} \left(\frac{\sigma_1^2 s_1^2}{m S_1^2} + \frac{\sigma_2^2 s_2^2}{m S_2^2} \right) \frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}}$$

- Gamage et al. (2004) provides a solution

Some Notations

- \bar{X}, \bar{Y} as the sample mean vector
- $W_1 = \sum_{i=1}^m (X_i - \bar{X})(X_i - \bar{X})' = (m-1)S_1$, $W_2 = (n-1)S_2$, the unscaled sample covariance matrices
- It can be derived that

$$\bar{X} - \bar{Y} \sim N_d(\mu_1 - \mu_2, \Sigma_p), \quad \Sigma_p = \frac{1}{m}\Sigma_1 + \frac{1}{n}\Sigma_2$$

$$W_1 \sim W(\Sigma_1, m-1), \quad W_2 \sim W(\Sigma_2, n-1)$$

- $W(\Sigma, n)$ is the Wishart distribution, the generalization of χ^2 distribution to multivariate case

Solution by G-p-value

- Let

$$R_1 = (w_1^{-\frac{1}{2}} \Sigma_1 w_1^{-\frac{1}{2}})^{-\frac{1}{2}} (w_1^{-\frac{1}{2}} W_1 w_1^{-\frac{1}{2}}) (w_1^{-\frac{1}{2}} \Sigma_1 w_1^{-\frac{1}{2}})^{-\frac{1}{2}}$$

R_2 alike, $Z_0 = \Sigma_p^{-\frac{1}{2}} (\bar{X} - \bar{Y})$, $\eta = (\Sigma_1, \Sigma_2)$, then our test variable is

$$T(X, Y; x, y, \eta) = Z_0' \left(\frac{1}{m} w_1^{\frac{1}{2}} R_1^{-1} w_1^{\frac{1}{2}} + \frac{1}{n} w_2^{\frac{1}{2}} R_2^{-1} w_2^{\frac{1}{2}} \right) Z_0$$

1. $T(x, y; x, y, \eta) = (\bar{x} - \bar{y})' (\bar{x} - \bar{y})$
2. Under H_0 ,

$R_1 \sim W(I_d, m - 1)$, $R_2 \sim W(I_d, n - 1)$, $Z_0 \sim N_d(0, I_d)$, so the distribution of T is parameter-free

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To be continued...

I'll cover this topic in the final course presentation

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A Summary of G- p -value

- Provide a p -value in the presence of nuisance parameters
- Relax restrictions on test statistic
- Allow test variable depend on observed values and nuisance parameters, while the final p -value turns out to be parameter-free

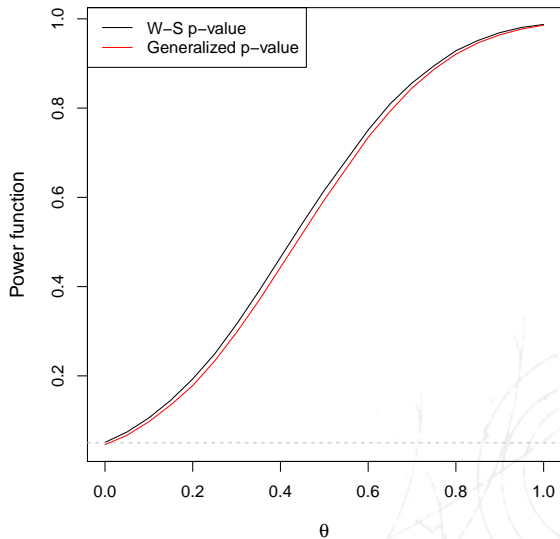
The Good Part

- Provides solutions to the Behrens-Fisher problem, both univariate case and multivariate case
- Also applicable to other problems where traditional method fails

The Not-So-Good Part

- Test variable depends on both random sample and observed value, sometimes hard to explain (in some sense related to Fiducial Inference)
- Tends to give conservative result (slightly less powerful)
- Distribution of test variable usually has no closed form (but easy for sampling)

The Not-So-Good Part



References

- Tsui, K. W., & Weerahandi, S. (1989). Generalized p -values in significance testing of hypotheses in the presence of nuisance parameters. *Journal of the American Statistical Association*, 84(406), 602-607.
- Gamage, J., Mathew, T., & Weerahandi, S. (2004). Generalized p -values and generalized confidence regions for the multivariate Behrens–Fisher problem and MANOVA. *Journal of Multivariate Analysis*, 88(1), 177-189.
- DasGupta, A. (2008). *Asymptotic theory of statistics and probability*. Springer.

Thank you!