

Non-parametric methods: comparing groups

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STAT 417

OUTLINE

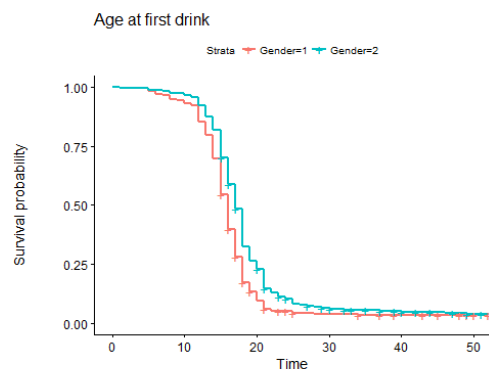
Getting started (2 groups)

Log-rank (2 groups)

Software output (2 groups)

> 2 groups

Comparing survival experiences of independent populations



Which group appears to have the larger median survival time?

1. 1 = males
2. 2 = females

Briefly comment on the survival experiences by gender.

		Number at risk					
Strata	Gender=1	461	435	55	13	6	2
	Gender=2	539	525	138	28	15	5
		0	10	20	30	40	50
		Time					

Comparing survival experiences of independent populations

The previous figure reveals an observed difference in the estimated survival curves among males and females based on the samples of ages. Why can't we stop here?

Inference procedures to compare survival experiences

To compare survival curves over the range of time t among several independent populations, we can conduct formal tests.:

- 1.
- 2.

Age at first drink, by-hand example

Males	43+	15	19	14	18+	16	14
Females	18	15	17	16	40+	24+	16

R Code

```
drinksu <- data.frame(time = c(43, 15, 19, 14, 18, 16, 14,
                             18, 15, 17, 16, 40, 24, 16),
                      censor = c(0, 1, 1, 1, 0, 1, 1,
                                1, 1, 1, 1, 0, 0, 1),
                      gender = c("m", "m", "m", "m", "m", "m", "m",
                                "f", "f", "f", "f", "f", "f", "f"))

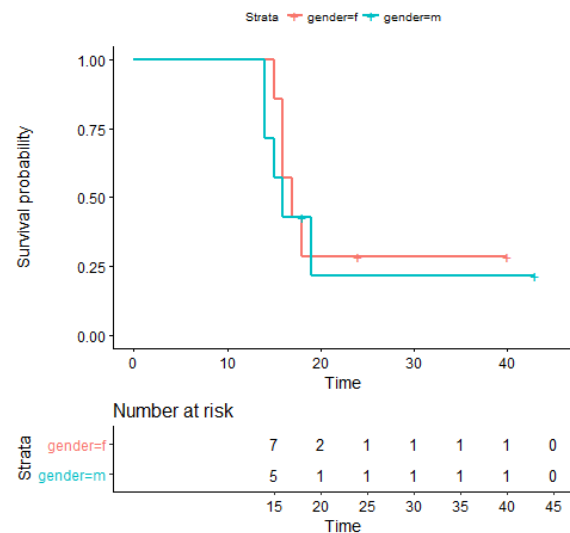
KM_obj <- survfit(Surv(time, censor) ~ gender, data = drinksu)

library(survminer)
ggsurvplot(KM_obj, data = drinksu,
            risk.table = TRUE,
            title = "Age at first drink (subset)")
```

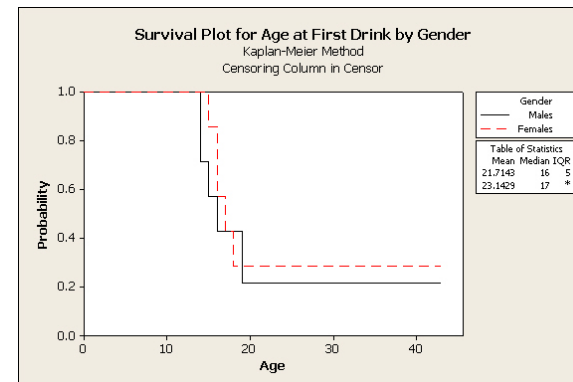
R Code

Comparing survival experiences in R

Age at first drink (subset)



Comparing survival experiences in Minitab



Getting started (2 groups) ○○○○○●○	Log-rank (2 groups) ○○○○○○○○○○	Software output (2 groups) ○○○○○○○	> 2 groups ○○○○○○○○○○
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Log-rank test

The log-rank test can be used to compare survival experiences for **two** populations. The general null and alternative hypotheses are:

H_0 :

H_a :

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Getting started (2 groups) ○○○○○●○	Log-rank (2 groups) ○○○○○○○○○○	Software output (2 groups) ○○○○○○○	> 2 groups ○○○○○○○○○○
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H_0 and H_a for age at first drink

In the age at first drink example, let 1 = males and 2 = females. Then the null and alternative hypotheses are:

H_0 :

H_a :

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Getting started (2 groups) ○○○○○○○	Log-rank (2 groups) ●○○○○○○○○○	Software output (2 groups) ○○○○○○○	> 2 groups ○○○○○○○○○○
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Getting started (2 groups)

Log-rank (2 groups)

Software output (2 groups)

> 2 groups

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Getting started (2 groups) ○○○○○○○	Log-rank (2 groups) ●○○○○○○○○○	Software output (2 groups) ○○○○○○○	> 2 groups ○○○○○○○○○○
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Log-rank test details

The approach is to compare the total number of *observed* events to the total number of *expected* events in Group 1 under the assumption that $S_1(t) = S_2(t)$, where:

- ▶ m
- ▶ n_{1i}
- ▶ n_{2i}
- ▶ $n_i = n_{1i} + n_{2i}$
- ▶ d_{1i} , d_{2i} , and d_i

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Log-rank test details

It is typically assumed that the number of event occurrences that will occur in time interval i , for Group 1 follows a *hypergeometric* probability distribution.

- ▶ Let d_i/n_i be the overall proportion of individuals at time $t_{(i)}$ who experience the event. Then the expected number of event occurrences in Group 1 at time $t_{(i)}$, denoted E_{1i} , is given by:
- ▶ Also, based on the assumption that the number of event occurrences at time $t_{(i)}$ for Group 1 follows a hypergeometric distribution, the variance for the number of event occurrences for Group 1 at time $t_{(i)}$, denoted V_{1i} , is given by:

Age at first drink, by-hand example

Males	43+	15	19	14	18+	16	14
Females	18	15	17	16	40+	24+	16

Write out the ordered event times:

Order	1	2	3	4	5	6	7
Males							
Females							

Age at first drink, by-hand example

i	$t_{(i)}$	Interval	n_i	d_i	d_i/n_i	n_{1i}	d_{1i}	n_{2i}	E_{1i}	V_{1i}
0		[0, 14)								
1		[14, 15)								
2		[15, 16)								
3		[16, 17)								
4		[17, 18)								
5		[18, 19)								
6		[19, 43)								

Example calculations

E_{11} :

V_{11} :

Log-rank test details

We compare the total observed and total expected counts over *all* the complete event times with the statistic:

$$\frac{\text{observed} - \text{expected}}{sd} =$$

- ▶ $\sum_{i=1}^m d_{1i}$ is the sum of the observed event occurrences
- ▶ $\sum_{i=1}^m E_{1i}$ is the sum of the expected event occurrences
- ▶ $\sum_{i=1}^m V_{1i}$ is the variance of the total number of event occurrences over the m complete event times

Log-rank test details

- ▶ This quantity is essentially a z-score, i.e:
- ▶ By the Central Limit Theorem, when each sample size n_i is “reasonably large,” the above term follows approximately a standard normal distribution.

Log-rank test details

- ▶ It is more common to see the square of the statistic reported in statistical software.
- ▶ The square of the above statistic is called the **log-rank test statistic** (for two groups) given by:
- ▶ This test statistic follows a **chi-square distribution** with 1 degree of freedom.

Discussion

If the observed and expected number of events are far apart, then

- ▶ χ_L^2 will be (“big”, “small”),
- ▶ the corresponding p -value will be (“big”, “small”),
- ▶ leading to (evidence / no evidence) to reject H_0 .

Age at first drink, by-hand example

Calculate the value of the log-rank test statistic:

Getting started (2 groups)

Log-rank (2 groups)

Software output (2 groups)

> 2 groups

Minitab output

... Minitab Output

Test Statistics			
Method	Chi-Square	DF	P-Value
Log-Rank	0.131114	1	0.717
Wilcoxon	0.425355	1	0.514

Minitab Output

Conclusion?

Wilcoxon test

- ▶ The log-rank test statistic can be considered a special case of the following statistic:

where the w_i 's are "weights" all equal to _____.

- ▶ The **Wilcoxon test statistic**, denoted X_W^2 , is also a special case of above statistic with weights $w_i = n_i$, i.e:

$$X_W^2 = \frac{[\sum_{i=1}^m n_i (d_{1i} - E_{1i})]^2}{\sum_{i=1}^m n_i^2 V_{1i}}$$

where n_i is the number of subjects at risk just prior to time $t_{(i)}$

- ▶ The Wilcoxon test statistic, X^2 , also follows a χ^2 -distribution with 1 degree of freedom.

Wilcoxon test statistic

Verify that the value of the Wilcoxon test statistics is 0.4253.

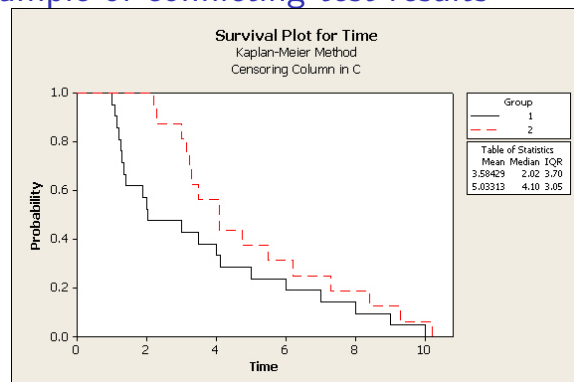
Comparison of log-rank and Wilcoxon tests

1.

2.

3.

Example of conflicting test results



Minitab Output

Method	Chi-Square	DF	P-Value
Log-Rank	2.13966	1	0.144
Wilcoxon	4.27216	1	0.039

Minitab Output

Log-rank test results in R

R Code

```
survdif(Surv(time, censor) ~ gender, data = drinksub)
```

R Code

R Output

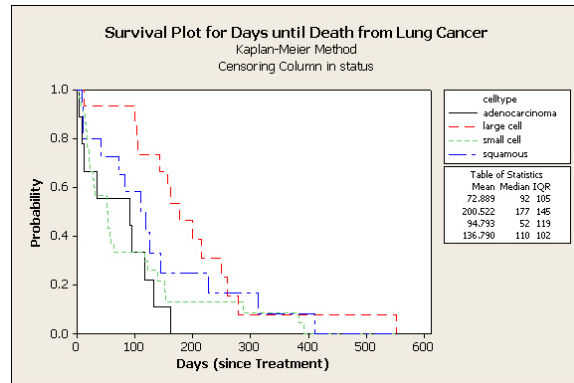
Call:
survdif(formula = Surv(time, censor) ~ gender, data = drinksub)

	N	Observed	Expected	(O-E) ² /E	(O-E) ² /V
gender=f	7	5	5.54	0.0523	0.131
gender=m	7	5	4.46	0.0649	0.131

Chisq= 0.1 on 1 degrees of freedom, p= 0.717

R Output

Lung cancer example (VALCSG)



Lung cancer example (VALCSG)

... Minitab Output

Test Statistics

Method	Chi-Square	DF	P-Value
Log-Rank	9.6408	3	0.022
Wilcoxon	11.5123	3	0.009

Minitab Output

H_0 :

H_a :

Lung cancer example (VALCSG)

► State conclusion based on log-rank test at $\alpha = 0.01$.

► State conclusion based on Wilcoxon test at $\alpha = 0.01$.

Lung cancer example (VALCSG) in R

R Code

```
# focus on treatment group 1 only
veteran1 <- veteran[veteran$trt == 1, ]

KM_obj <- survfit(Surv(time, status) ~ celltype,
                  data = veteran1)

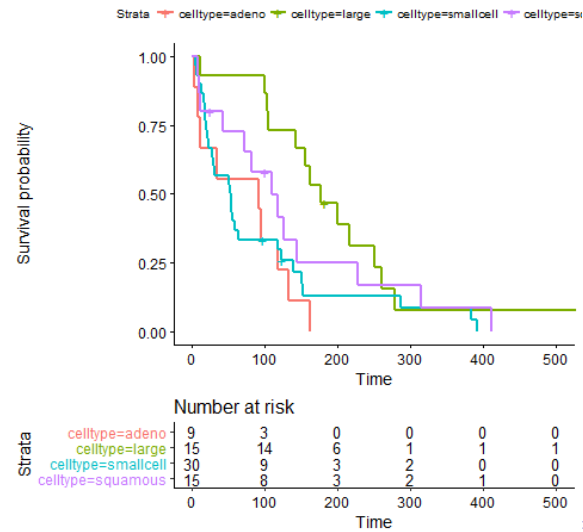
ggsurvplot(KM_obj, data = veteran1,
            risk.table = TRUE,
            title = "VALCSG study, trt=1")

survdif(Surv(time, status) ~ celltype, data = veteran1)
```

R Code

Lung cancer example (VALCSG) in R

VALCSG study, trt=1



Lung cancer example (VALCSG) in R

R Output

Call:

```
survdif(formula = Surv(time, status) ~ celltype, data = veteran1)
```

	N	Observed	Expected	(O-E) ² /E	(O-E) ² /V
celltype=adeno	9	9	5.17	2.832	3.194
celltype=large	15	14	23.17	3.626	6.217
celltype=smallcell	30	28	21.14	2.225	3.503
celltype=squamous	15	13	14.52	0.159	0.211

Chisq= 9.6 on 3 degrees of freedom, p= 0.0219

R Output