

# The Projected Power Method: An Efficient Algorithm for Joint Alignment from Pairwise Differences

Yuxin Chen

Emmanuel Candès



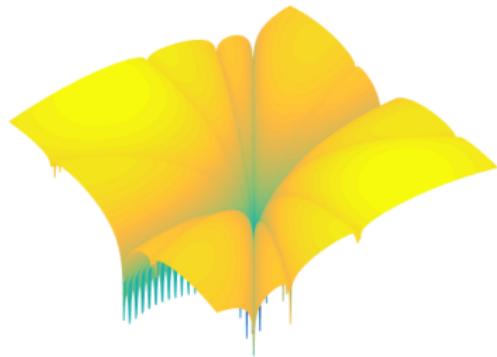
*Department of Statistics, Stanford University, Sep. 2016*

# Nonconvex optimization is everywhere

For instance, maximum likelihood estimation is nonconvex in numerous problems

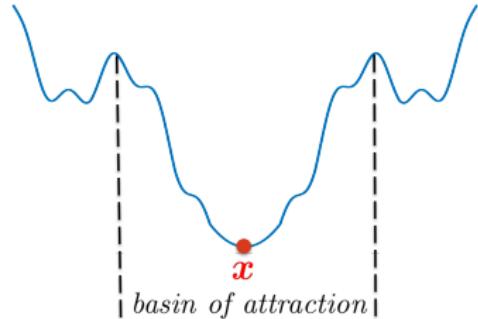
$$\begin{aligned} & \text{maximize}_{\boldsymbol{x}} && \ell(\boldsymbol{x}; \boldsymbol{y}) \\ & \text{subject to} && \boldsymbol{x} \in \mathcal{S} \end{aligned}$$

- matrix completion
- phase retrieval
- dictionary learning
- blind deconvolution
- robust PCA
- ...



# Recent flurry of research in nonconvex procedures

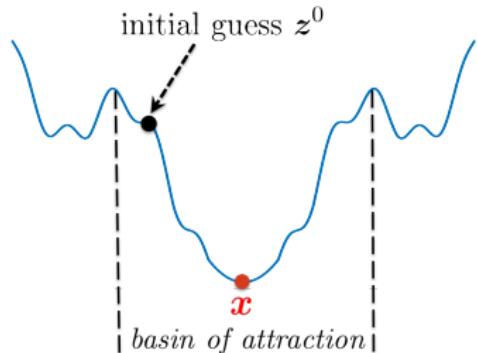
Nice geometry within a neighborhood around  $x$  (basin of attraction)



Keshavan et al'08, Netrapalli et al'13, Candès et al'14, Soltanolkotabi'14, Jain et al'14, Sun et al'14, Chen et al'15, Cai et al'15, Tu et al'15, Sun et al'15, White et al'15, Li et al'16, Yi et al'16, Zhang et al'16, Wang et al'16, ...

# Recent flurry of research in nonconvex procedures

Nice geometry within a neighborhood around  $x$  (basin of attraction)

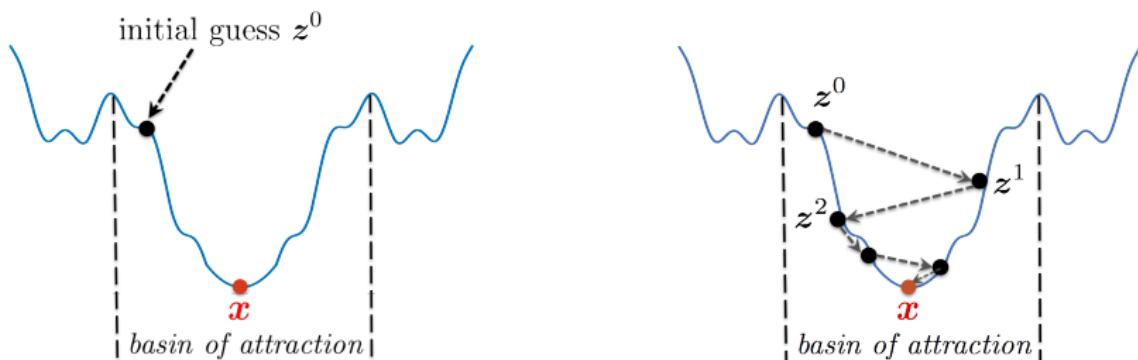


Suggests two-stage paradigms

1. Start from an appropriate initial point

# Recent flurry of research in nonconvex procedures

Nice geometry within a neighborhood around  $x$  (basin of attraction)



Suggests two-stage paradigms

1. Start from an appropriate initial point
2. Proceed via some iterative updates

---

Keshavan et al'08, Netrapalli et al'13, Candès et al'14, Soltanolkotabi'14, Jain et al'14, Sun et al'14, Chen et al'15, Cai et al'15, Tu et al'15, Sun et al'15, White et al'15, Li et al'16, Yi et al'16, Zhang et al'16, Wang et al'16, ...

*This talk: a discrete nonconvex problem*

## Joint alignment from pairwise differences

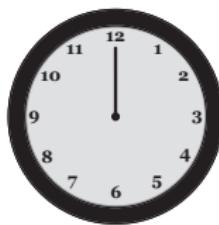
- $n$  unknown variables:  $x_1, \dots, x_n$
- $m$  possible states:  $x_i \in \{1, 2, \dots, m\}$



$$x_1 = 1$$



$$x_2 = 6$$

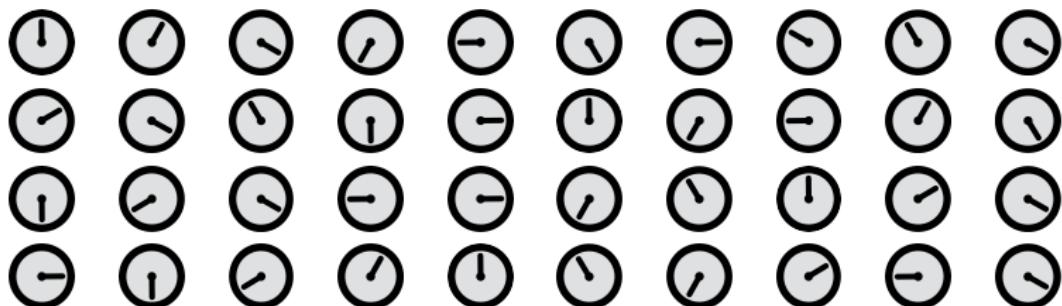


$$x_3 = 12$$

...

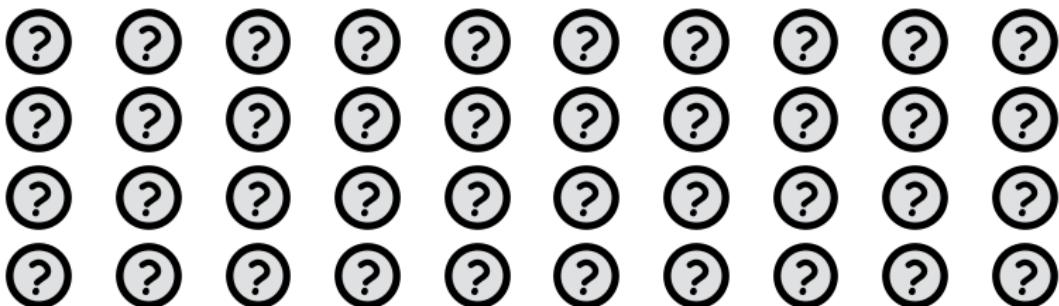
## Joint alignment from pairwise differences

- $n$  unknown variables:  $x_1, \dots, x_n$
- $m$  possible states:  $x_i \in \{1, 2, \dots, m\}$



## Joint alignment from pairwise differences

- $n$  unknown variables:  $x_1, \dots, x_n$
- $m$  possible states:  $x_i \in \{1, 2, \dots, m\}$



## Joint alignment from pairwise differences

- **Measurements:** pairwise differences

$$y_{i,j} \stackrel{\text{ind.}}{=} x_i - x_j + \underbrace{\eta_{i,j}}_{\text{noise}} \pmod{m}, \quad i \neq j$$



$$x_i - x_j \pmod{m}$$

# Joint alignment from pairwise differences

- **Measurements:** pairwise differences

$$y_{i,j} \stackrel{\text{ind.}}{=} x_i - x_j + \underbrace{\eta_{i,j}}_{\text{noise}} \pmod{m}, \quad i \neq j$$

- e.g. random corruption model



$$x_i - x_j \pmod{m}$$

$$y_{i,j} \stackrel{\text{ind.}}{=} \begin{cases} x_i - x_j \pmod{m} & \text{with prob. } \pi_0 \\ \text{Uniform}(m) & \text{else} \end{cases}$$

- $\pi_0$ : non-corruption rate

# Joint alignment from pairwise differences

- **Measurements:** pairwise differences

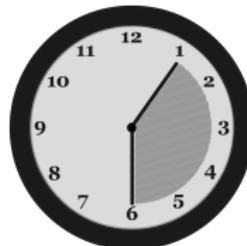
$$y_{i,j} \stackrel{\text{ind.}}{=} x_i - x_j + \underbrace{\eta_{i,j}}_{\text{noise}} \pmod{m}, \quad i \neq j$$

- e.g. random corruption model



$$y_{i,j} \stackrel{\text{ind.}}{=} \begin{cases} x_i - x_j \pmod{m} & \text{with prob. } \pi_0 \\ \text{Uniform}(m) & \text{else} \end{cases}$$

- $\pi_0$ : non-corruption rate



- **Goal:** recover  $\{x_i\}$  (up to global offset)

## Motivation: multi-image alignment

Jointly align a collection of images/shapes of the same physical object

# Motivation: multi-image alignment

Jointly align a collection of images/shapes of the same physical object

- $x_i$ : angle of rotation associated with each shape



computer vision/graphics

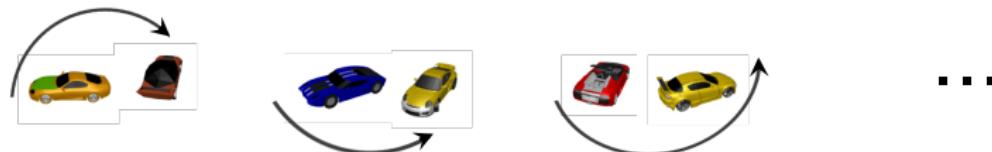
# Motivation: multi-image alignment

**Step 1:** compute pairwise estimates of relative angles of rotations



# Motivation: multi-image alignment

**Step 1:** compute pairwise estimates of relative angles of rotations



**Step 2:** aggregate these pairwise information for joint alignment



## Maximum likelihood estimates (MLE)

$$\begin{aligned} & \text{maximize}_{\{x_i\}} \quad \sum_{i,j} \ell(x_i, x_j; y_{i,j}) \\ & \text{subj. to} \quad x_i \in \{1, \dots, m\}, \quad 1 \leq i \leq n \end{aligned}$$

- Log-likelihood function  $\ell$  may be complicated

## Maximum likelihood estimates (MLE)

$$\begin{aligned} & \text{maximize}_{\{x_i\}} \quad \sum_{i,j} \ell(x_i, x_j; y_{i,j}) \\ & \text{subj. to} \quad x_i \in \{1, \dots, m\}, \quad 1 \leq i \leq n \end{aligned}$$

- Log-likelihood function  $\ell$  may be complicated
- Discrete input space

## Maximum likelihood estimates (MLE)

$$\begin{aligned} & \text{maximize}_{\{x_i\}} \quad \sum_{i,j} \ell(x_i, x_j; y_{i,j}) \\ & \text{subj. to} \quad x_i \in \{1, \dots, m\}, \quad 1 \leq i \leq n \end{aligned}$$

- Log-likelihood function  $\ell$  may be complicated
- Discrete input space
- Looks daunting

## Another look in lifted space

Discrete variables  $\rightarrow$  orthogonal vectors in higher-dimensional space

$$x_i = 1 \iff \boldsymbol{x}_i = \boldsymbol{e}_1 = \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \quad \begin{array}{c} \text{blue square} \\ \text{white squares} \end{array}$$

$$x_i = 2 \iff \boldsymbol{x}_i = \boldsymbol{e}_1 = \begin{bmatrix} 0 \\ 1 \\ \vdots \\ 0 \end{bmatrix} \quad \begin{array}{c} \text{white square} \\ \text{blue square} \\ \text{white squares} \end{array}$$

⋮

$$x_i = j \iff \boldsymbol{x}_i = \boldsymbol{e}_j$$

## Another look in lifted space

Pairwise sample  $y_{i,j} \rightarrow$  encode  $\ell(x_i, x_j)$  by  $L_{i,j} \in \mathbb{R}^{m \times m}$

$$[L_{i,j}]_{\alpha,\beta} = \ell(x_i = \alpha, x_j = \beta)$$

## Another look in lifted space

Pairwise sample  $y_{i,j} \rightarrow$  encode  $\ell(x_i, x_j)$  by  $L_{i,j} \in \mathbb{R}^{m \times m}$

$$[L_{i,j}]_{\alpha,\beta} = \ell(x_i = \alpha, x_j = \beta)$$

- e.g. random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{w.p. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases} \Rightarrow \ell(x_i, x_j) = \begin{cases} \log(\pi_0 + \frac{1-\pi_0}{m}), & \text{if } x_i - x_j = y_{i,j} \\ \log(\frac{1-\pi_0}{m}), & \text{else} \end{cases}$$

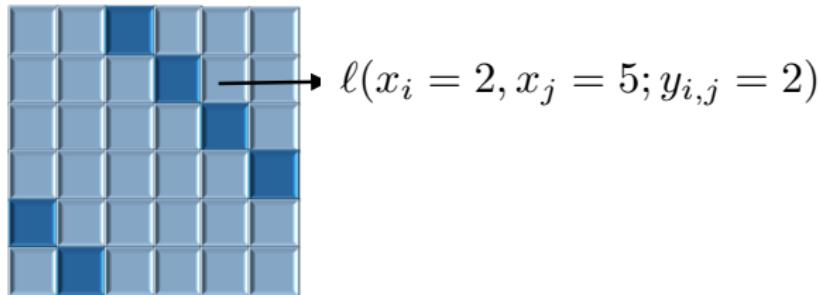
## Another look in lifted space

Pairwise sample  $y_{i,j} \rightarrow$  encode  $\ell(x_i, x_j)$  by  $\mathbf{L}_{i,j} \in \mathbb{R}^{m \times m}$

$$[\mathbf{L}_{i,j}]_{\alpha,\beta} = \ell(x_i = \alpha, x_j = \beta)$$

- e.g. random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{w.p. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases} \Rightarrow \ell(x_i, x_j) = \begin{cases} \log(\pi_0 + \frac{1-\pi_0}{m}), & \text{if } x_i - x_j = y_{i,j} \\ \log(\frac{1-\pi_0}{m}), & \text{else} \end{cases}$$



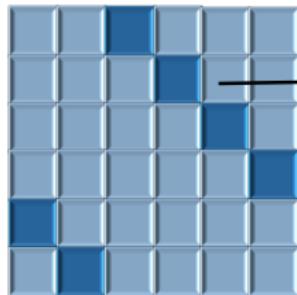
# Another look in lifted space

Pairwise sample  $y_{i,j} \rightarrow$  encode  $\ell(x_i, x_j)$  by  $L_{i,j} \in \mathbb{R}^{m \times m}$

$$[L_{i,j}]_{\alpha,\beta} = \ell(x_i = \alpha, x_j = \beta)$$

- e.g. random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{w.p. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases} \Rightarrow \ell(x_i, x_j) = \begin{cases} \log(\pi_0 + \frac{1-\pi_0}{m}), & \text{if } x_i - x_j = y_{i,j} \\ \log(\frac{1-\pi_0}{m}), & \text{else} \end{cases}$$

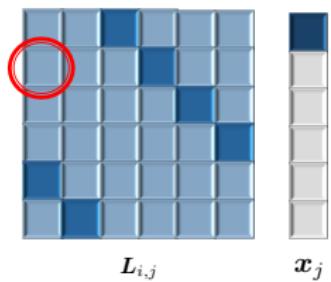


$$\ell(x_i = 2, x_j = 5; y_{i,j} = 2)$$

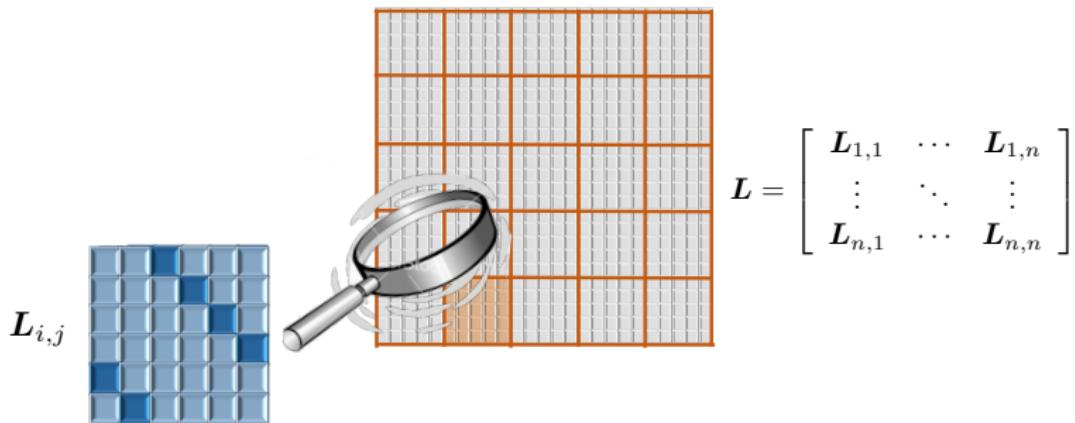
This enables quadratic representation

$$\ell(x_i, x_j) = \mathbf{x}_i^\top L_{i,j} \mathbf{x}_j$$

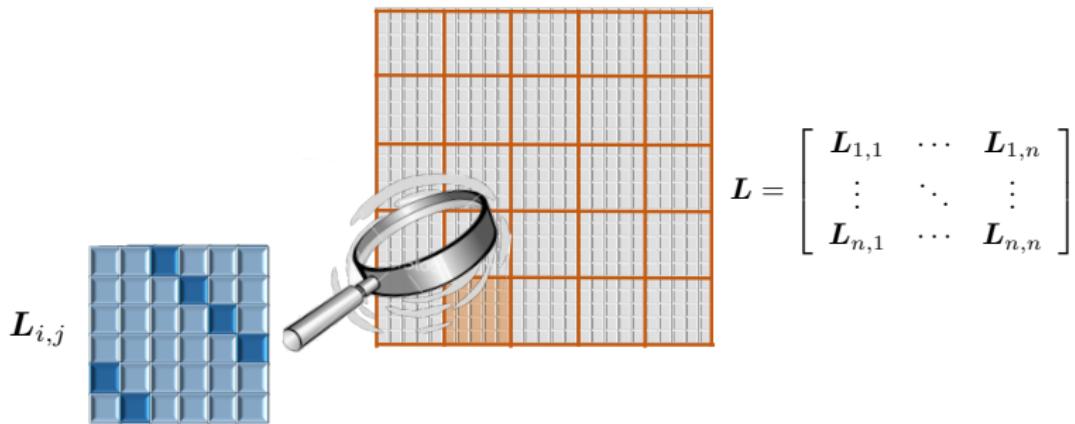
$$\mathbf{x}_i^\top$$



MLE is equivalent to a binary quadratic program



# MLE is equivalent to a binary quadratic program

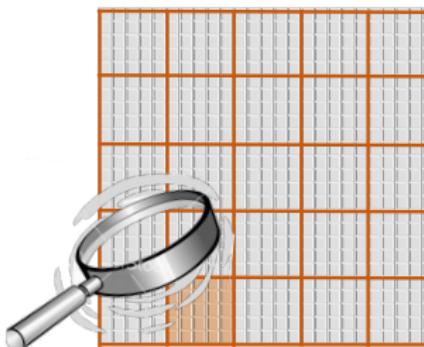


$$\begin{array}{ll}\text{maximize}_{\boldsymbol{x}} & \sum_{i,j} \ell(x_i - x_j; y_{ij}) \\ \text{subj. to} & x_i \in \{1, \dots, m\}\end{array}$$

↔

$$\begin{array}{ll}\text{maximize}_{\boldsymbol{x}} & \boldsymbol{x}^\top \mathbf{L} \boldsymbol{x} \\ \text{subj. to} & \boldsymbol{x} = \begin{bmatrix} \boldsymbol{x}_1 \\ \vdots \\ \boldsymbol{x}_n \end{bmatrix} \\ & \boldsymbol{x}_i \in \{\mathbf{e}_1, \dots, \mathbf{e}_m\}\end{array}$$

# MLE is equivalent to a binary quadratic program


$$\mathbf{L}_{i,j}$$
$$\mathbf{L} = \begin{bmatrix} \mathbf{L}_{1,1} & \cdots & \mathbf{L}_{1,n} \\ \vdots & \ddots & \vdots \\ \mathbf{L}_{n,1} & \cdots & \mathbf{L}_{n,n} \end{bmatrix}$$
$$\text{maximize}_{\mathbf{x}} \quad \mathbf{x}^\top \mathbf{L} \mathbf{x}$$
$$\text{maximize} \quad \sum_{i,j} \ell(x_i - x_j; y_{ij})$$
$$\text{subj. to} \quad x_i \in \{1, \dots, m\}$$
$$\iff$$
$$\text{subj. to} \quad \mathbf{x} = \begin{bmatrix} \mathbf{x}_1 \\ \vdots \\ \mathbf{x}_n \end{bmatrix}$$
$$\mathbf{x}_i \in \{\mathbf{e}_1, \dots, \mathbf{e}_m\}$$

This is essentially nonconvex constrained PCA

# How to solve nonconvex constrained PCA?

## PCA

$$\begin{aligned} \text{maximize}_x \quad & \mathbf{x}^\top \mathbf{L} \mathbf{x} \\ \text{subj. to} \quad & \|\mathbf{x}\| = 1 \end{aligned}$$

### Power method:

for  $t = 1, 2, \dots$

$$\mathbf{z}^{(t)} = \mathbf{L} \mathbf{z}^{(t-1)}$$

$$\mathbf{z}^{(t)} \leftarrow \text{normalize}(\mathbf{z}^{(t)})$$

# How to solve nonconvex constrained PCA?

## PCA

$$\begin{array}{ll}\text{maximize}_{\boldsymbol{x}} & \boldsymbol{x}^\top \mathbf{L} \boldsymbol{x} \\ \text{subj. to} & \|\boldsymbol{x}\| = 1\end{array}$$

## Constrained PCA

$$\begin{array}{ll}\text{maximize}_{\boldsymbol{x}} & \boldsymbol{x}^\top \mathbf{L} \boldsymbol{x} \\ \text{subj. to} & \boldsymbol{x}_i \in \{\boldsymbol{e}_1, \dots, \boldsymbol{e}_m\}\end{array}$$

### Power method:

for  $t = 1, 2, \dots$

$$\boldsymbol{z}^{(t)} = \mathbf{L} \boldsymbol{z}^{(t-1)}$$

$$\boldsymbol{z}^{(t)} \leftarrow \text{normalize}(\boldsymbol{z}^{(t)})$$

### Projected power method:

for  $t = 1, 2, \dots$

$$\boldsymbol{z}^{(t)} = \mathbf{L} \boldsymbol{z}^{(t-1)}$$

$$\boldsymbol{z}^{(t)} \leftarrow \text{Project}_{\Delta^n}(\mu \boldsymbol{z}^{(t)})$$

- $\mu$ : scaling factor

## Projection onto standard simplex

$$\underset{\mathbf{x} \in \{\mathbf{x}_i\}}{\text{maximize}} \quad \mathbf{x}^\top \mathbf{Lx} \quad \text{s.t. } \mathbf{x}_i \in \{\mathbf{e}_1, \dots, \mathbf{e}_m\}$$

$$\mathbf{z}^{(t)} = \mathbf{Lz}^{(t-1)}$$

$$\mathbf{z}^{(t)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{z}^{(t)})$$

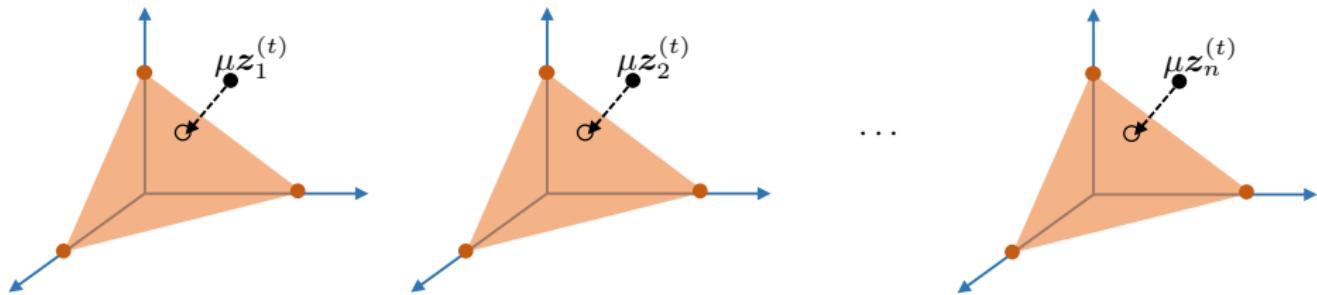
# Projection onto standard simplex

$$\underset{\mathbf{x} = \{\mathbf{x}_i\}}{\text{maximize}} \quad \mathbf{x}^\top \mathbf{Lx} \quad \text{s.t. } \mathbf{x}_i \in \{\mathbf{e}_1, \dots, \mathbf{e}_m\}$$

$$\mathbf{z}^{(t)} = \mathbf{Lz}^{(t-1)}$$

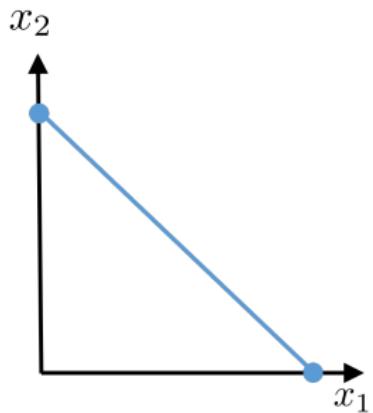
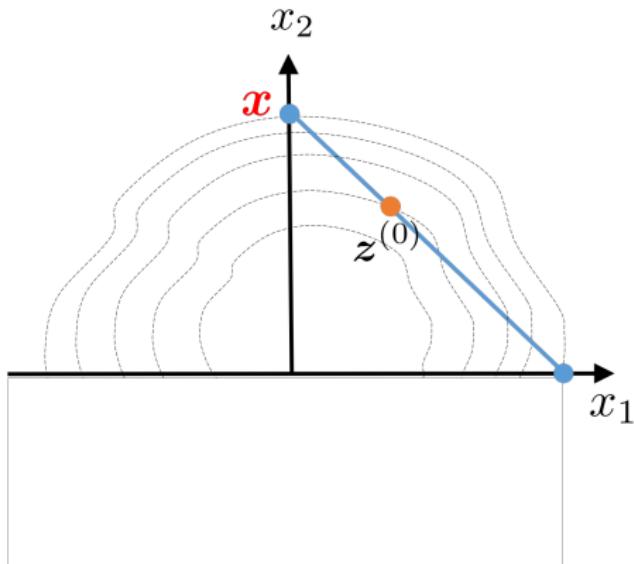
$$\mathbf{z}^{(t)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{z}^{(t)})$$

$\Delta^n$  is convex hull of feasibility set, i.e.  $\left\{ \mathbf{z} = [\mathbf{z}_i]_{1 \leq i \leq n} \mid \forall i : \mathbf{1}^\top \mathbf{z}_i = 1; \mathbf{z}_i \geq \mathbf{0} \right\}$



## Illustration

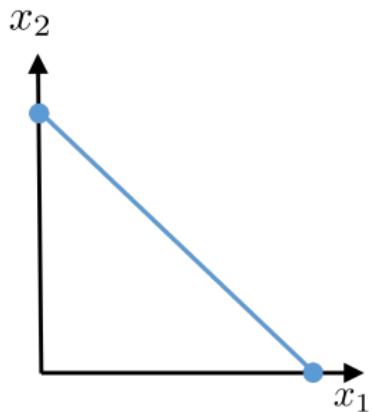
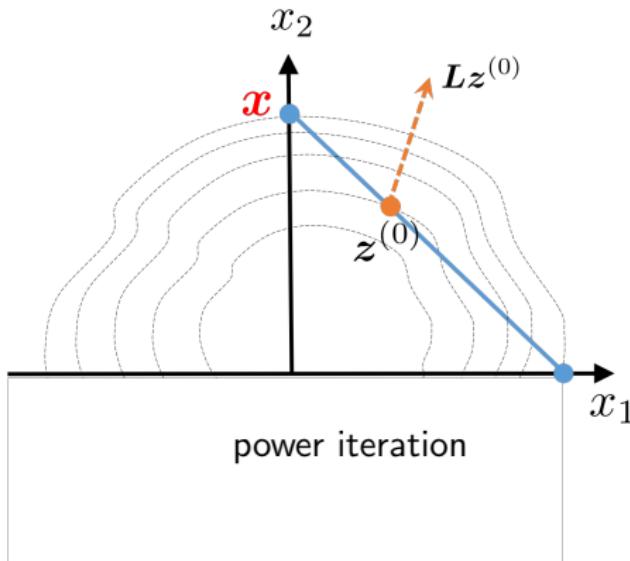
Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$



## Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

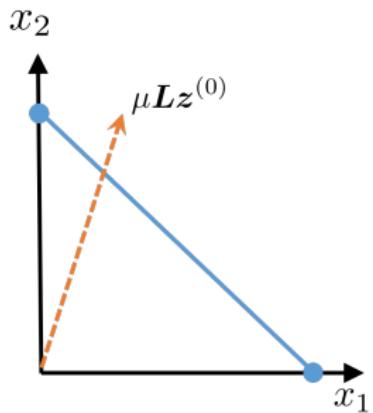
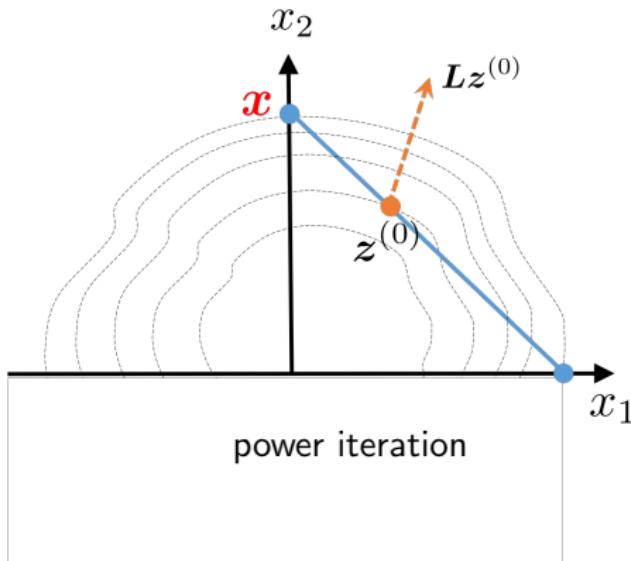
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



## Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

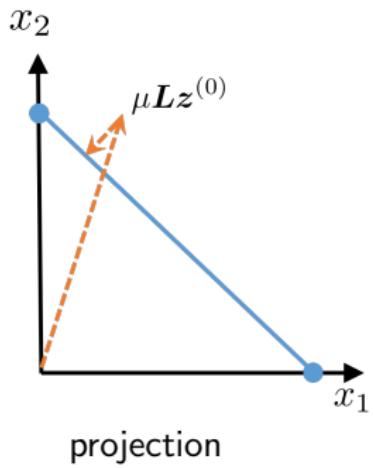
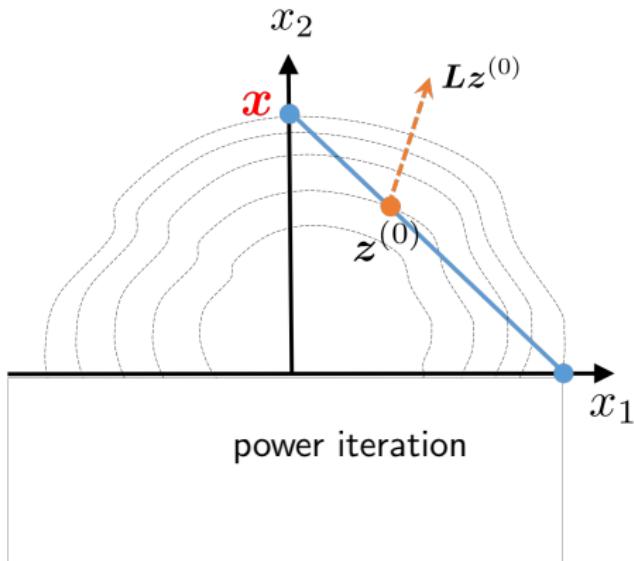
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



# Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

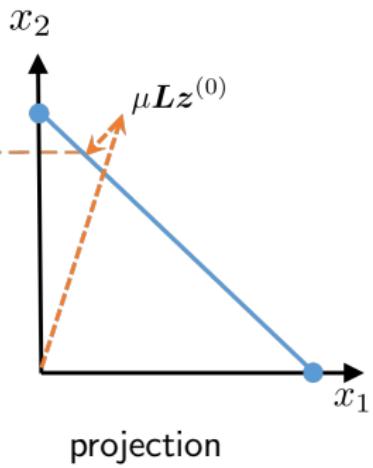
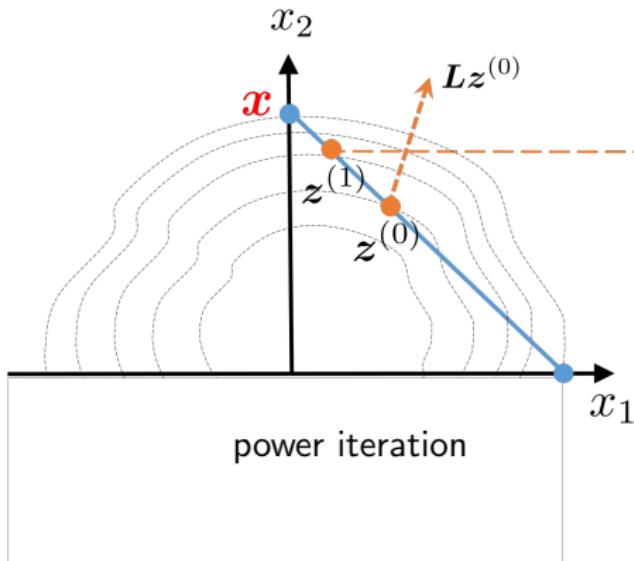
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



## Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

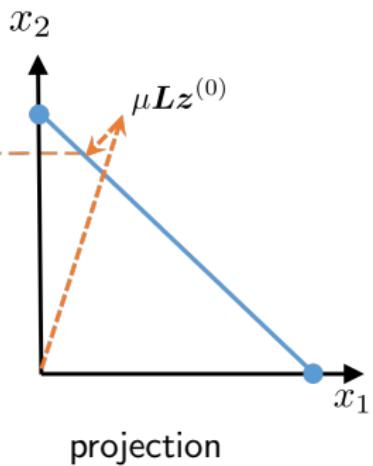
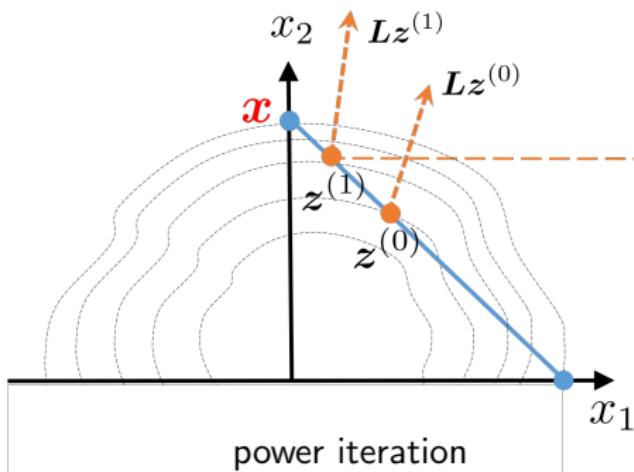
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



## Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

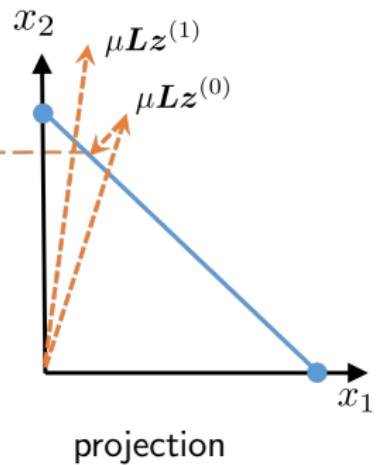
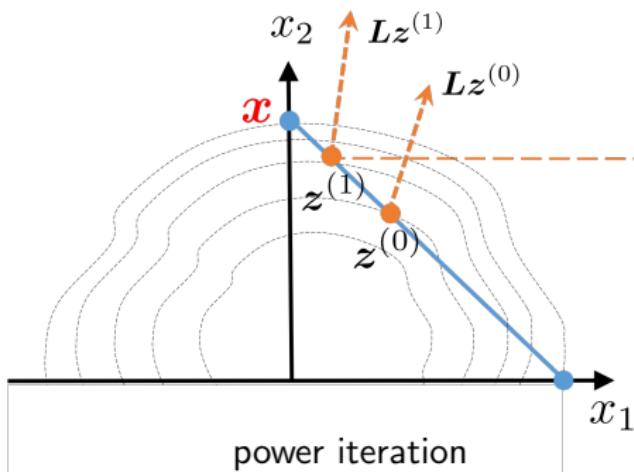
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



## Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

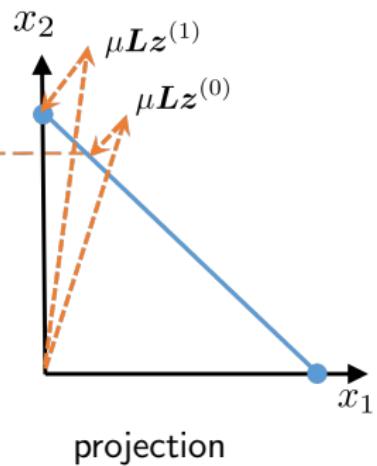
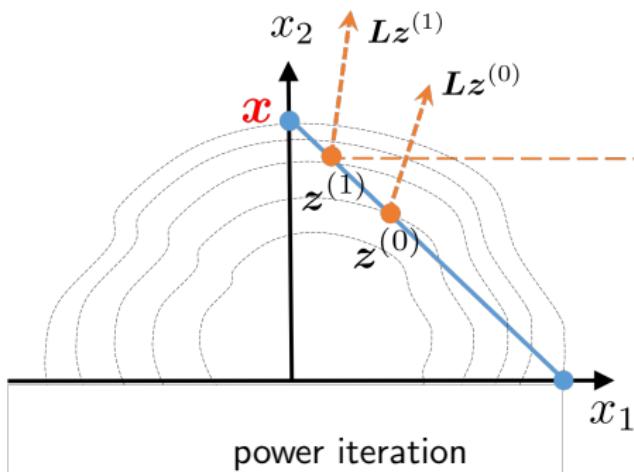
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



# Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

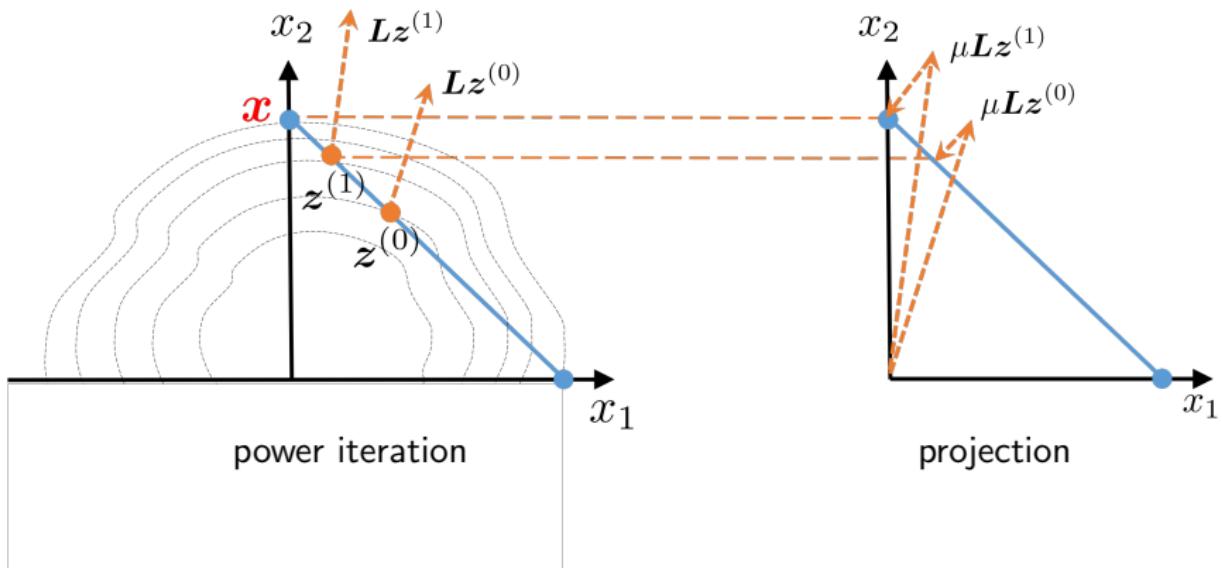
- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



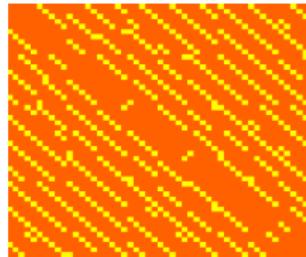
# Illustration

Projected power method:  $\mathbf{z}^{(t+1)} \leftarrow \text{Project}_{\Delta^n} (\mu \mathbf{L} \mathbf{z}^{(t)})$

- $\mathbf{L} \mathbf{z}$  is gradient of  $\frac{1}{2} \mathbf{z}^\top \mathbf{L} \mathbf{z}$



# Initialization?



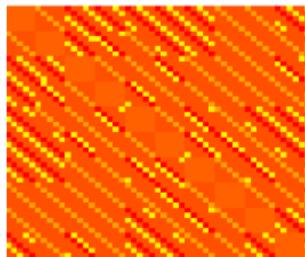
$$\mathbf{L}$$

=



$$\underbrace{\mathbb{E}[\mathbf{L}]}_{\text{approx. low-rank}}$$

+

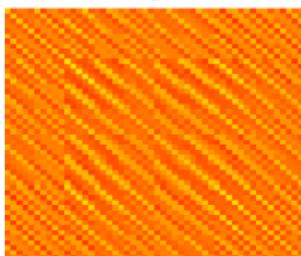


$$\mathbf{L} - \mathbb{E} [\mathbf{L}]$$

# Initialization?



$$\mathbf{L} = \underbrace{\mathbb{E}[\mathbf{L}]}_{\text{approx. low-rank}} + \mathbf{L} - \mathbb{E}[\mathbf{L}]$$

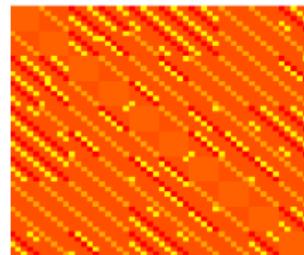
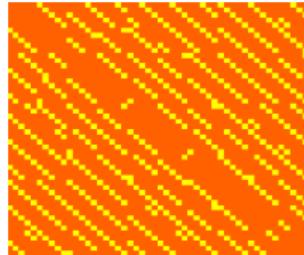


$\hat{\mathbf{L}}$

**Spectral initialization**

1.  $\hat{\mathbf{L}} \leftarrow$  rank- $m$  approximation of  $\mathbf{L}$

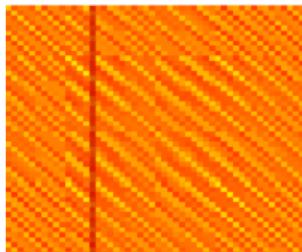
# Initialization?

 $L$  $=$ 

$$\underbrace{\mathbb{E}[L]}_{\text{approx. low-rank}}$$

 $+$ 

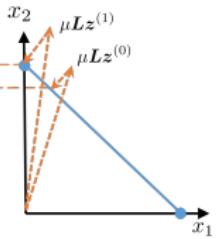
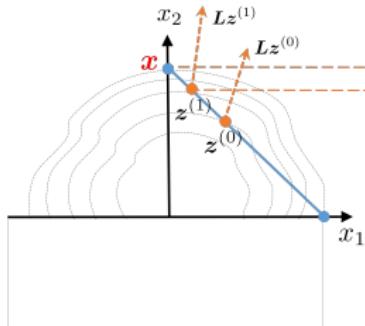
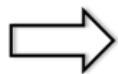
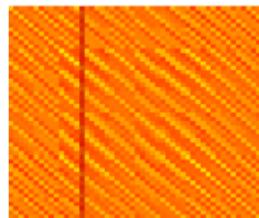
$$L - \mathbb{E}[L]$$

 $\hat{L}$ 

## Spectral initialization

1.  $\hat{L} \leftarrow$  rank- $m$  approximation of  $L$
2.  $z^{(0)} \leftarrow \text{Project}_{\Delta^n}(\mu \hat{z})$ , where  $\hat{z}$  is a random column of  $\hat{L}$

# Summary of projected power method (PPM)



1. Spectral initialization
2. For  $t = 1, 2, \dots$

$$\mathbf{z}^{(t)} \leftarrow \text{Project}_{\Delta^n} \left( \mu \mathbf{L} \mathbf{z}^{(t-1)} \right)$$

# Random corruption model

$$y_{i,j} \stackrel{\text{ind}}{=} \begin{cases} x_i - x_j \bmod m & \checkmark \text{ with prob. } \pi_0 \\ \text{Uniform}(m) & \text{else} \end{cases}$$



# Random corruption model

$$y_{i,j} \stackrel{\text{ind}}{=} \begin{cases} x_i - x_j \bmod m & \checkmark \text{ with prob. } \pi_0 \\ \text{Uniform}(m) & \text{else} \end{cases}$$



**Theorem (Chen-Candès'16)** Fix  $m > 0$  and set  $\mu \gtrsim 1/\sigma_2(\mathbf{L})$ . With high prob., PPM recovers the truth exactly within  $O(\log n)$  iterations if

- signal-to-noise ratio (SNR) not too small:  $\pi_0 > 2\sqrt{\frac{\log n}{mn}}$

## Implications

**Theorem (Chen-Candès'16)** PPM succeeds within  $O(\log n)$  iterations if

$$\text{non-corruption rate } \pi_0 > 2\sqrt{\frac{\log n}{mn}}$$

- PPM succeeds even when most (i.e.  $1 - O(\sqrt{\frac{\log n}{n}})$ ) entries are corrupted

# Implications

**Theorem (Chen-Candès'16)** PPM succeeds within  $O(\log n)$  iterations if

$$\text{non-corruption rate } \pi_0 > 2\sqrt{\frac{\log n}{mn}}$$

- PPM succeeds even when most (i.e.  $1 - O(\sqrt{\frac{\log n}{n}})$ ) entries are corrupted
- Nearly linear time algorithm

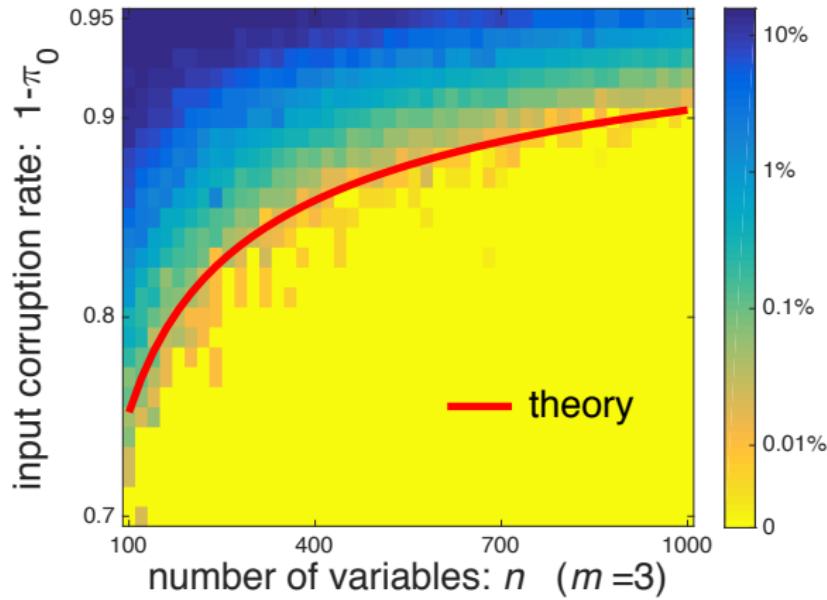
# Implications

**Theorem (Chen-Candès'16)** PPM succeeds within  $O(\log n)$  iterations if

$$\text{non-corruption rate } \pi_0 > 2\sqrt{\frac{\log n}{mn}}$$

- PPM succeeds even when most (i.e.  $1 - O(\sqrt{\frac{\log n}{n}})$ ) entries are corrupted
- Nearly linear time algorithm
- Works for any initialization obeying  $\|\mathbf{z}^{(0)} - \mathbf{x}\| < 0.5\|\mathbf{x}\|$

## Empirical misclassification rate



Misclassification rate when  $n$  and  $\pi_0$  vary    ( $\mu = 10/\sigma_2(\mathbf{L})$ )

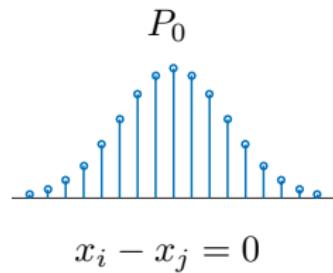
## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

Distributions of  $y_{i,j}$  under different hypotheses

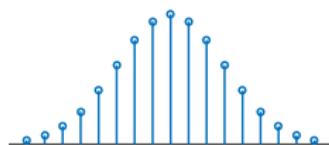


## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

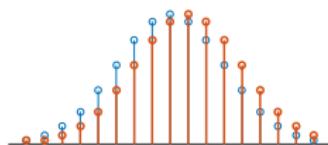
Distributions of  $y_{i,j}$  under different hypotheses

$P_0$



$$x_i - x_j = 0$$

$P_1$

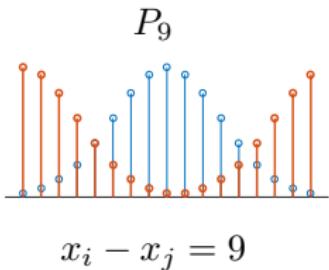
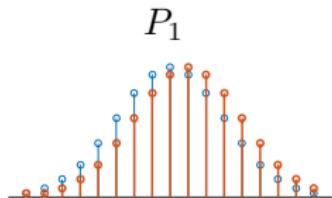
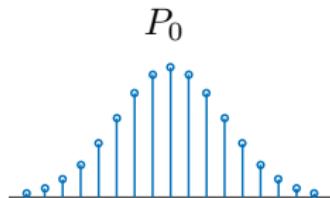


$$x_i - x_j = 1$$

## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

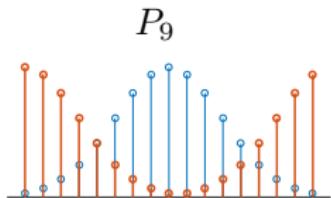
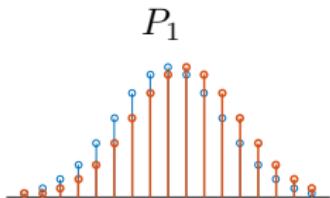
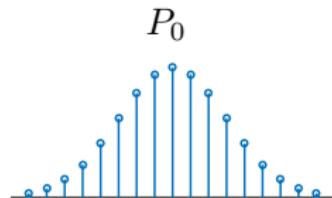
Distributions of  $y_{i,j}$  under different hypotheses



## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

Distributions of  $y_{i,j}$  under different hypotheses



$$\downarrow$$

$$\text{KL}(P_0 \parallel P_1)$$

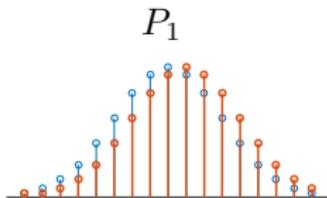
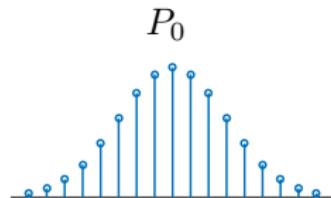
$$\downarrow$$

$$\text{KL}(P_0 \parallel P_9)$$

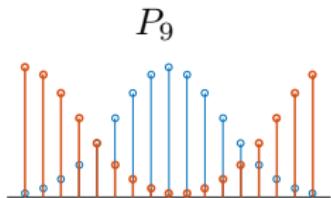
## More general noise models

$$y_{i,j} = x_i - x_j + \eta_{i,j} \bmod m, \quad \text{where } \eta_{i,j} \stackrel{\text{i.i.d.}}{\sim} P_0$$

Distributions of  $y_{i,j}$  under different hypotheses



$$\begin{array}{c} \downarrow \\ \text{KL}(P_0 \parallel P_1) \end{array}$$



$$\begin{array}{c} \downarrow \\ \text{KL}(P_0 \parallel P_9) \end{array}$$

**Theorem (Chen-Candès'16)** Fix  $m > 0$  and set  $\mu \gtrsim 1/\sigma_2(\mathbf{L})$ . Under mild conditions, PPM succeeds within  $O(\log n)$  iterations with high prob., provided that

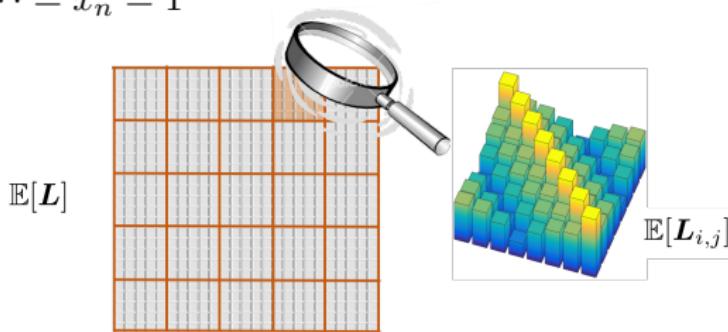
$$\text{KL}_{\min} := \min_{1 \leq l < m} \text{KL}(P_0 \parallel P_l) > \frac{4 \log n}{n}$$

# Interpretation: why $\text{KL}_{\min}$ matters

**Theorem (Chen-Candès'16)** ... PPM succeeds within  $O(\log n)$  iterations if

$$\text{KL}_{\min} := \min_{1 \leq l < m} \text{KL}(P_0 \parallel P_l) > \frac{4 \log n}{n}$$

Suppose  $x_1 = \dots = x_n = 1$

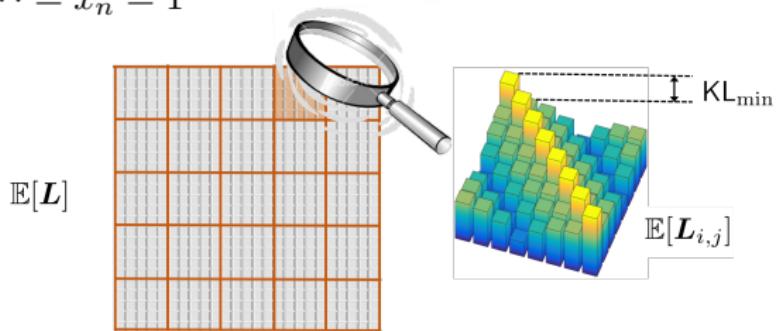


# Interpretation: why $\text{KL}_{\min}$ matters

**Theorem (Chen-Candès'16)** ... PPM succeeds within  $O(\log n)$  iterations if

$$\text{KL}_{\min} := \min_{1 \leq l < m} \text{KL}(P_0 \parallel P_l) > \frac{4 \log n}{n}$$

Suppose  $x_1 = \dots = x_n = 1$

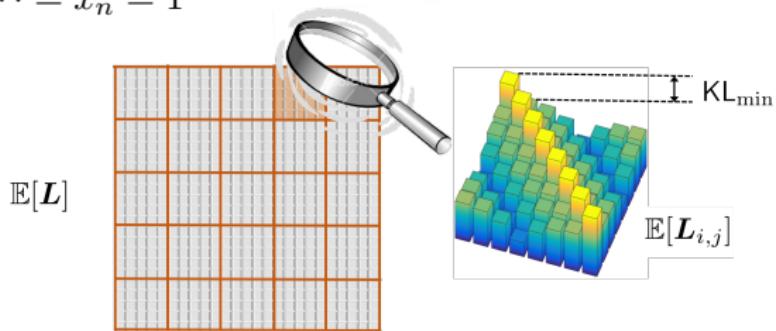


# Interpretation: why $\text{KL}_{\min}$ matters

**Theorem (Chen-Candès'16)** ... PPM succeeds within  $O(\log n)$  iterations if

$$\text{KL}_{\min} := \min_{1 \leq l < m} \text{KL}(P_0 \parallel P_l) > \frac{4 \log n}{n}$$

Suppose  $x_1 = \dots = x_n = 1$



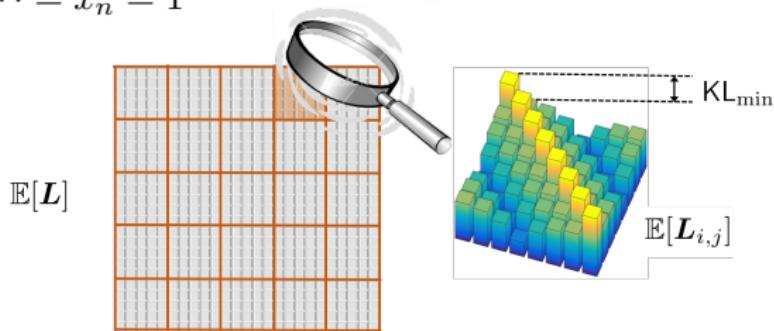
- Peaks of  $E[\mathbf{L}]$  reveal ground truth  $E[\mathbf{L}_{i,j}]$

# Interpretation: why $\text{KL}_{\min}$ matters

**Theorem (Chen-Candès'16)** ... PPM succeeds within  $O(\log n)$  iterations if

$$\text{KL}_{\min} := \min_{1 \leq l < m} \text{KL}(P_0 \parallel P_l) > \frac{4 \log n}{n}$$

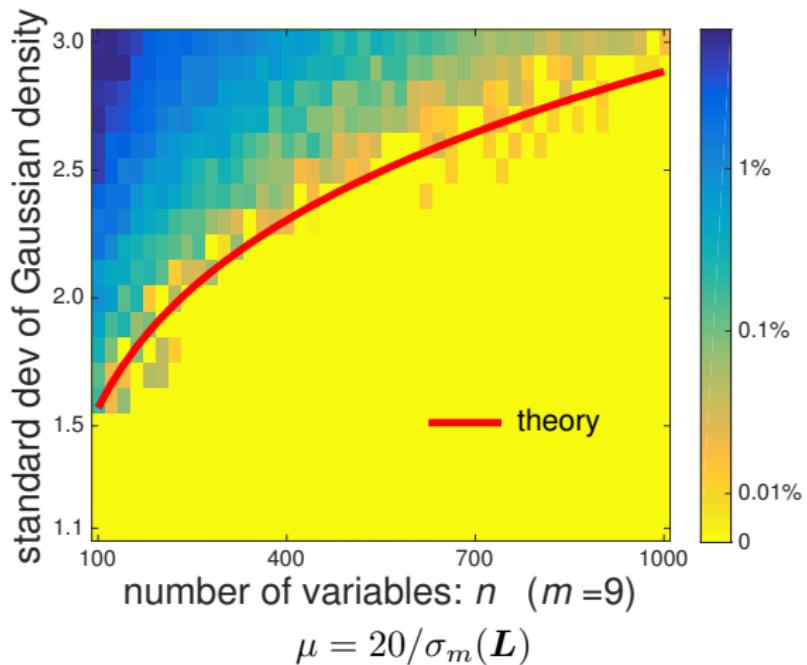
Suppose  $x_1 = \dots = x_n = 1$



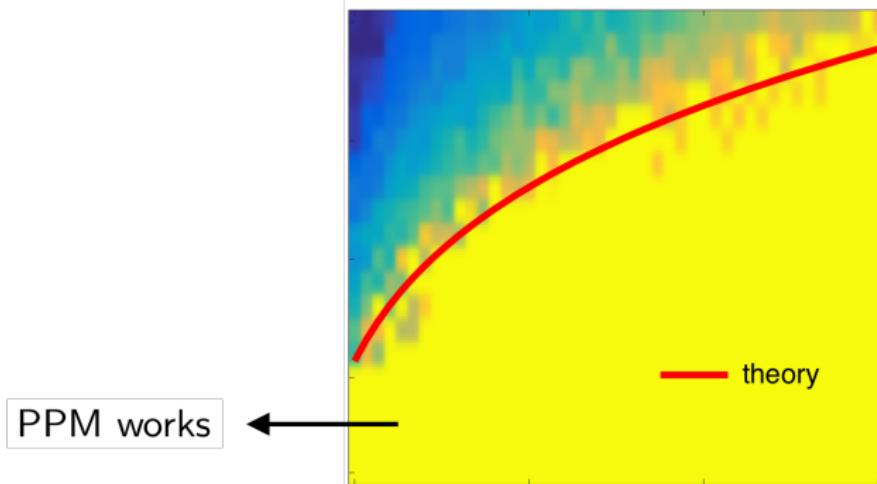
- Peaks of  $\mathbb{E}[\mathbf{L}]$  reveal ground truth  $\mathbb{E}[\mathbf{L}_{i,j}]$
- $\mathbf{L} \approx \mathbb{E}[\mathbf{L}]$  if  $\text{KL}_{\min}$  is sufficiently large

# Empirical misclassification rate

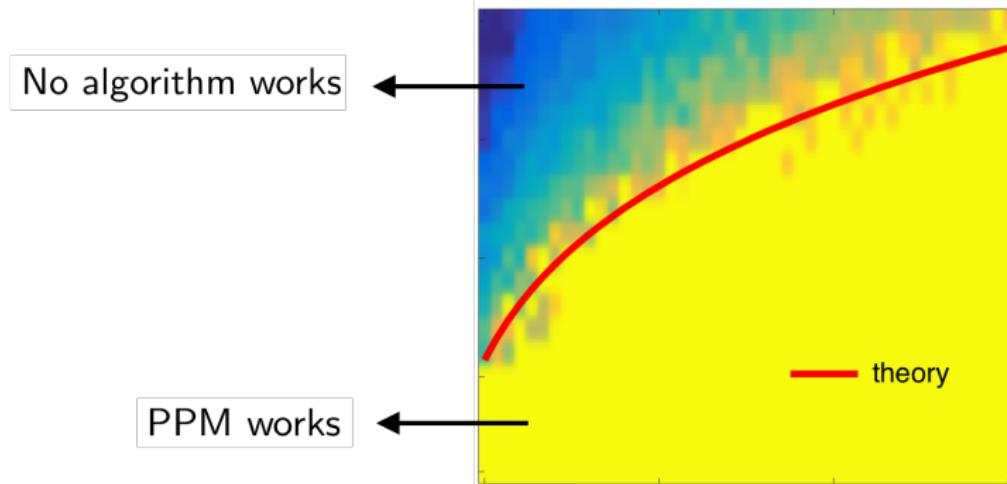
Modified Gaussian noise model:  $\mathbb{P}\{\eta_{i,j} = z\} \propto \exp\left(-\frac{z^2}{2\sigma^2}\right), \quad |z| \leq \frac{m-1}{2}$



# PPM is information-theoretically optimal



# PPM is information-theoretically optimal



**Theorem (Chen-Candès'16)** Fix  $m > 0$ . No method achieves exact recovery if

$$\text{KL}_{\min} < \frac{4 \log n}{n}$$

## Large- $m$ case: random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases}$$

**Theorem (Chen-Candès'16)** Suppose  $\log n \lesssim m \lesssim \text{poly}(n)$ . PPM succeeds if

$$\pi_0 \gtrsim \frac{1}{\sqrt{n}}$$

## Large- $m$ case: random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases}$$

**Theorem (Chen-Candès'16)** Suppose  $\log n \lesssim m \lesssim \text{poly}(n)$ . PPM succeeds if

$$\pi_0 \gtrsim \frac{1}{\sqrt{n}}$$

- Spiky model: when  $m \gg n$ , model converges to



$$x_i \in [0, 1), \quad y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(0, 1), & \text{else} \end{cases}$$

## Large- $m$ case: random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases}$$

**Theorem (Chen-Candès'16)** Suppose  $\log n \lesssim m \lesssim \text{poly}(n)$ . PPM succeeds if

$$\pi_0 \gtrsim \frac{1}{\sqrt{n}}$$

- Spiky model: when  $m \gg n$ , model converges to



$$x_i \in [0, 1), \quad y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(0, 1), & \text{else} \end{cases}$$

- Succeeds even if a dominant fraction  $1 - O(1/\sqrt{n})$  of inputs are corrupted

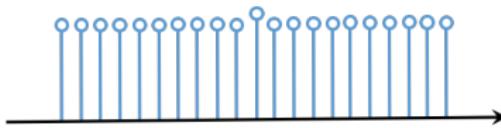
## Large- $m$ case: random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases}$$

**Theorem (Chen-Candès'16)** Suppose  $\log n \lesssim m \lesssim \text{poly}(n)$ . PPM succeeds if

$$\pi_0 \gtrsim \frac{1}{\sqrt{n}}$$

- “Smooth” noise model if  $m \lesssim \sqrt{n}$



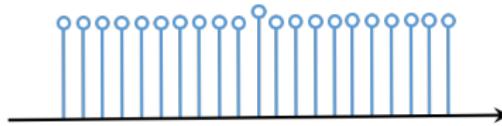
## Large- $m$ case: random corruption model

$$y_{i,j} = \begin{cases} x_i - x_j, & \text{with prob. } \pi_0 \\ \text{Unif}(m), & \text{else} \end{cases}$$

**Theorem (Chen-Candès'16)** Suppose  $\log n \lesssim m \lesssim \text{poly}(n)$ . PPM succeeds if

$$\pi_0 \gtrsim \frac{1}{\sqrt{n}}$$

- “Smooth” noise model if  $m \lesssim \sqrt{n}$



- Recovers each  $x_i \in [0, 1]$  up to a resolution of  $\frac{1}{m} \asymp \frac{1}{\sqrt{n}}$

## Joint shape alignment: Chair dataset from ShapeNet<sup>1</sup>



20 representative shapes (out of 50)

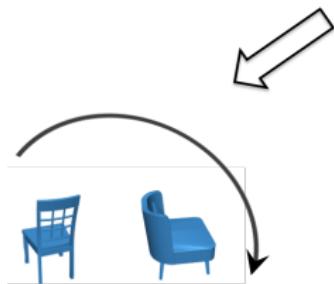
---

<sup>1</sup>We add extra noise to each point of the shapes to make it more challenging.

# Joint shape alignment: Chair dataset from ShapeNet<sup>1</sup>



20 representative shapes (out of 50)



pairwise cost  $-\ell_{i,j}(x_i, x_j)$ :

avg nearest-neighbor squared distance

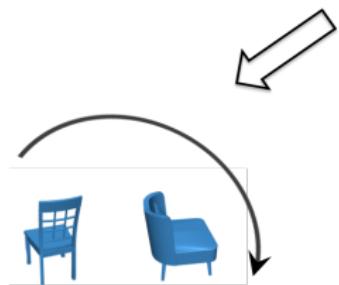
---

<sup>1</sup>We add extra noise to each point of the shapes to make it more challenging.

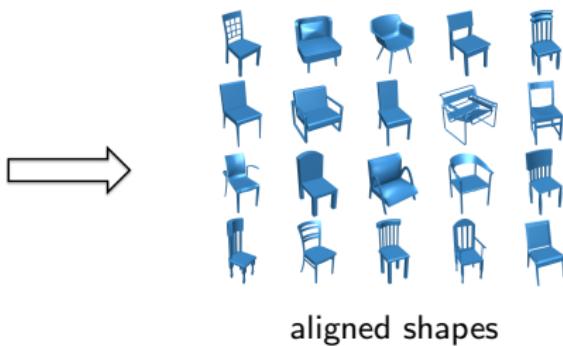
# Joint shape alignment: Chair dataset from ShapeNet<sup>1</sup>



20 representative shapes (out of 50)



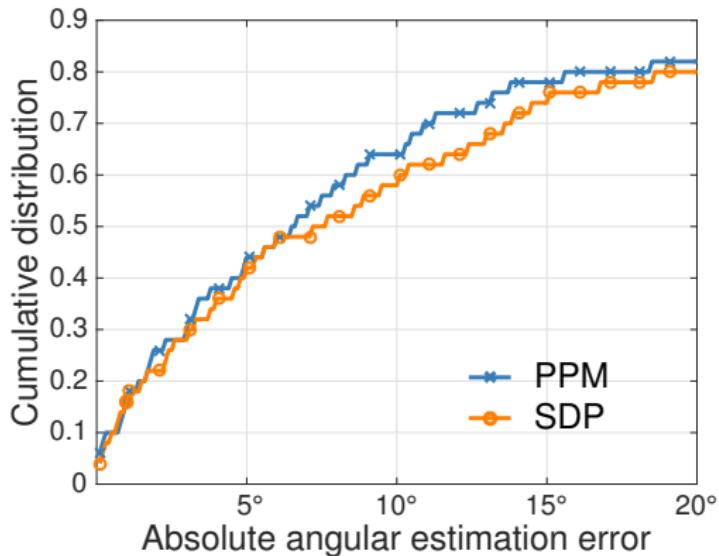
pairwise cost  $-\ell_{i,j}(x_i, x_j)$ :  
avg nearest-neighbor squared distance



aligned shapes

<sup>1</sup>We add extra noise to each point of the shapes to make it more challenging.

## Joint shape alignment: angular estimation errors<sup>2</sup>



	projected power method	semidefinite relaxation
Runtime	2.4 sec	895.6 sec

<sup>2</sup>We add extra noise to each point of the shapes to make it more challenging.

## Joint graph matching: CMU House dataset



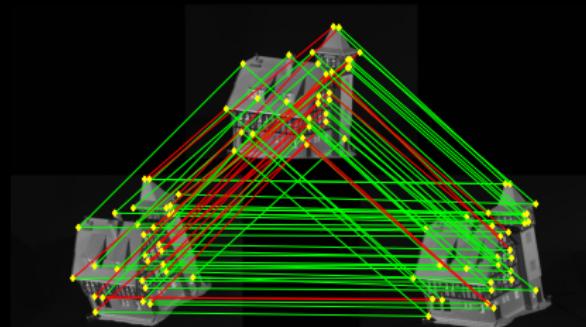
...

111 images of a toy house

# Joint graph matching: CMU House dataset



111 images of a toy house



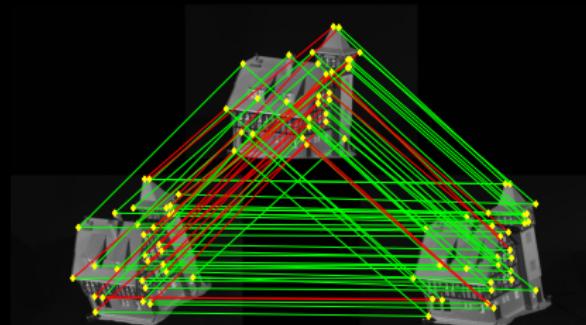
input matches

3 representative images

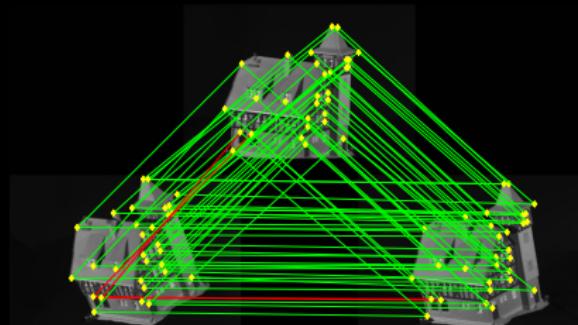
# Joint graph matching: CMU House dataset



111 images of a toy house



input matches



optimized matches

3 representative images

# Dixon imaging in body MRI

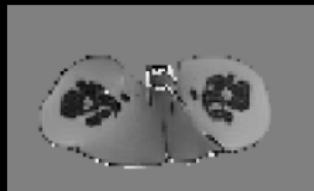
*Zhang et al., Magn. Reson. Med., 2016*

2 phasor candidates for field inhomogeneity at *each voxel*

candidate 1



candidate 2



# Dixon imaging in body MRI

Zhang et al., Magn. Reson. Med., 2016

2 phasor candidates for field inhomogeneity at *each voxel*

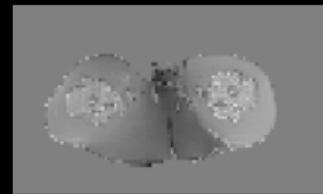
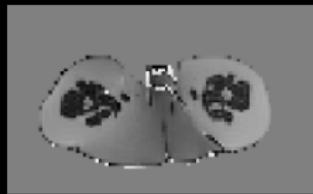
candidate 1



optimize some  
pairwise cost function



candidate 2



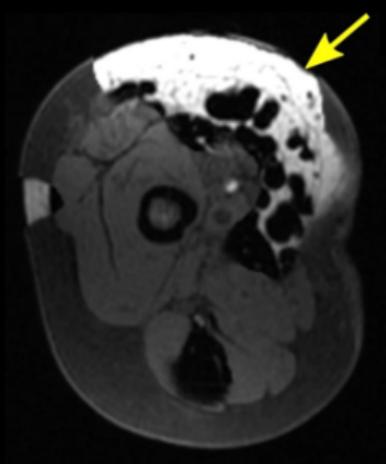
recovery

$$\begin{aligned} & \text{maximize} && \sum \ell(x_i, x_j) \\ & \text{subject to} && x_i \in \{1, 2\} \end{aligned}$$

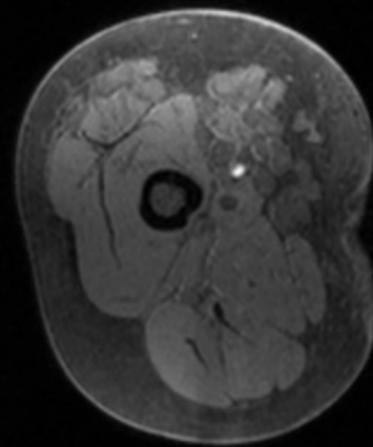
# Dixon imaging in body MRI

Zhang et al., *Magn. Reson. Med.*, 2016

Representative cases of water signal recovery



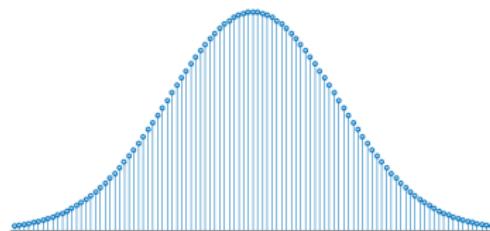
commercial software



projected power method

# Things I have not talked about ...

1. General noise model with large  $m$



2. Incomplete data



## Concluding remarks

A new approach to discrete assignment problems

- Finds MLE in suitable regimes
- Computationally efficient

**Paper:** “The projected power method: an efficient algorithm for joint alignment from pairwise differences”, Y. Chen and E. Candès, 2016