Mathematics for Robotics (ROB-GY 6013 Section A)

- Week 11:
 - Probability and Estimation

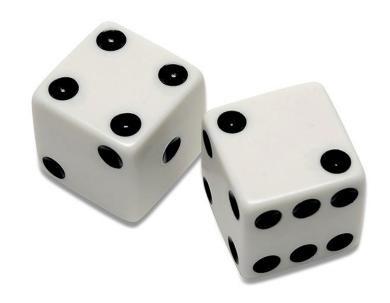
Rolling a Die (Discrete)

• Sample space Ω :

• {"rolling a 1," "rolling a 2," "rolling a 3," "rolling a 4," "rolling a 5," "rolling a 6"}



- P("rolling a 1") = 1/6
- P("rolling an even number") = 3/6
- Expected value: e.g., average points rolled
 - Assign point value to each event
 - Sum of point value × event probability



$$1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6} = \frac{21}{6}$$

Probability Distribution (Continuous)

- Sample space Ω :
 - Real numbers $x \in [-\infty, \infty]$
- Probability of an event
 - Integral of a probability density function f(x)

$$P(x \in [c,d]) = \int_{c}^{d} f(x)dx$$

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$$P(x \in [-\infty,\infty]) = 1 = \int_{-\infty}^{\infty} f(x)dx$$

- Expectation operator: e.g., average function value g(x)
 - Assign g(x) to each $x \in [-\infty, \infty]$
 - Integral of $g(x) \times f(x)$

$$E\{g(x)\} = \int_{-\infty}^{\infty} g(x)f(x)dx$$

Probability Space: The RIGHT way to begin

• Disclaimer: This part is just for fun.



Definition: Probability Space

- (Ω, \mathcal{F}, P) is called a **probability space**.
 - Ω is the sample space. Think of it as the set of all possible outcomes of an experiment.
 - $E \subset \Omega$ is an event
 - \mathcal{F} is the collection of allowed events. It must at least contain \emptyset and Ω . It is closed with respect to set complement, countable unions, and countable intersections. Such sets are called sigma algebras.
 - $P:\mathcal{F} \to [0, 1]$ is a probability measure. It has to satisfy a few basic operations:
 - 1. $P(\emptyset) = 0$ and $P(\Omega) = 1$.
 - 2. For each $E \in \mathcal{F}$, $0 \le P(E) \le 1$
 - 3. If the sets E_1, E_2, \ldots are disjoint (i.e., $E_i \cap E_j = \emptyset$ for $i \neq j$), then $P(\bigcup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$

Continuous (real-valued) random variable X

Probability density function

- Continuous (real-valued) random variable X
 - $X: \Omega \to \mathbb{R}$
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 - $X: \Omega \to \mathbb{R}$
- Probability density function
 - $f: \mathbb{R} \rightarrow [0,\infty)$
 - The value of f at a given point is not the probability
 - You have to integrate f to get the probability

- A function $X: \Omega \to \mathbb{R}$ is a **continuous random variable** with **density** $f: \mathbb{R} \to [0,\infty)$ if:
 - a) it is a random variable, and
 - b) $\forall x \in \mathbb{R}, P(\{\omega \in \Omega \mid X(\omega) \le x\}) = \int_{-\infty}^{x} f(\bar{x}) d\bar{x}.$

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$$P(x_1 \le X \le x_2) = \int_{x_1}^{x_2} f(x) dx$$

Examples: Uniform Random Variable

• Parameters are a, b

$$f(x) = \begin{cases} \frac{1}{b-a} & x \in [a,b] \\ 0 & otherwise \end{cases}$$

Examples:

• Parameters are $\sigma > 0, \mu \in \mathbb{R}$

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

Examples: Gaussian or Normal Random Variable

- Parameters are $\sigma > 0, \mu \in \mathbb{R}$
- $N(\mu, \sigma^2)$

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

Expected Value/Expectation Operator

$$E\{g(x)\} = \int_{-\infty}^{\infty} g(x)f(x)dx$$

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- Mean: $\mu := \mathcal{E}\{X\} := \int_{-\infty}^{\infty} x f(x) dx$
- Variance: $\sigma^2 := \mathcal{E}\{(X-\mu)^2\}\} := \int_{-\infty}^{\infty} (x-\mu)^2 f(x) dx$ (Var. for short)
- Standard Deviation: $\sigma := \sqrt{\sigma^2}$ (Std. Dev. for short)

- Mean
- Variance
- Skewness
- Kurtosis

$$\mu := E\{X\}$$

$$\sigma^2 := E\{(X - \mu)^2\}$$

$$\gamma_1 := \tilde{\mu}_3 = \frac{E\{(X - \mu)^3\}}{\sigma^3}$$

$$\text{Kurt}[X] := \frac{E\{(X - \mu)^4\}}{\sigma^4}$$



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$$E\{(X-\mu)^{2}\} = \int_{-\infty}^{\infty} (X-\mu)^{2} f(x) dx$$

2-norm squared for random variables

$$\gamma_1 := \tilde{\mu}_3 = \frac{E\{(X - \mu)^3\}}{3}$$

• Variance • Skewnesc • Kurtc 11^{12} • Kurtc 11^{12} • Minimum Variance • Within Minimum Variance • Kurtc 11^{12} • Minimum Variance • Kurtc 11^{12} • Minimum Variance • Kurtc 11^{12} • Minimum Variance • Kurtc 11^{12}

$$E\{(X-\mu)^2\} = \int_{-\infty}^{\infty} (X \text{Weighted Least squares})$$

2-norm squared for random variables

Random Vectors

• For example, the joint angles of a 6-DOF robot

$$X = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{bmatrix} \qquad \mu = E\{X\} = \begin{bmatrix} E\{X_1\} \\ E\{X_2\} \\ \vdots \\ E\{X_p\} \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix}$$

Random Vectors are painful

Definition 5.15 Let (Ω, \mathcal{F}, P) be a probability space. A function $X : \Omega \to \mathbb{R}^p$ is called a **random vector** if each component of

$$X = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix} \text{ is a random variable, that is, } \forall \ 1 \leq i \leq p, \ X_i : \Omega \to \mathbb{R} \text{ is a random variable.}$$

Consequently, $\forall x \in \mathbb{R}^p$, the set $\{\omega \in \Omega \mid X(\omega) \leq x\} \in \mathscr{F}$ (i.e., it is an allowed event), where the inequality is understood **pointwise**, that is,

$$\{\omega \in \Omega \mid X(\omega) \leq x\} := \left\{ \omega \in \Omega \mid \begin{bmatrix} X_1(\omega) \\ X_2(\omega) \\ \vdots \\ X_p(\omega) \end{bmatrix} \leq \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_p \end{bmatrix} \right\} := \left\{ \omega \in \Omega \mid \begin{bmatrix} X_1(\omega) \leq x_1 \\ X_2(\omega) \leq x_2 \\ \vdots \\ X_p(\omega) \leq x_p \end{bmatrix} \right\} = \bigcap_{i=1}^p \{\omega \in \Omega \mid X_i(\omega) \leq x_i\}.$$

Random Vectors are painful

Definition 5.16 $X: \Omega \to \mathbb{R}^p$ is a continuous random vector if there exists a density $f_X: \mathbb{R}^p \to [0, \infty)$ such that,

$$\forall x \in \mathbb{R}^P, \ P(\{X \le x\}) = \int_{-\infty}^{x_p} ... \int_{-\infty}^{x_2} \int_{-\infty}^{x_1} f_X(\bar{x}_1, \bar{x}_2 ... \bar{x}_p) d\bar{x}_1 d\bar{x}_2 ... d\bar{x}_p.$$

More generally, for all $A \subset \mathbb{R}^p$ such that the indicator function I_A has bounded variation,

$$P(\{X \in A\}) = \int_{-\infty}^{\infty} ... \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} I_A(\bar{x}_1, \bar{x}_2 ... \bar{x}_p) f_X(\bar{x}_1, \bar{x}_2 ... \bar{x}_p) d\bar{x}_1 d\bar{x}_2 ... d\bar{x}_p.$$

Notation 5.17 The notation $X \sim f$ is read as X is distributed with density f or that X is a random vector with density f.

Definition 5.18 (Moments) Suppose $g: \mathbb{R}^p \to \mathbb{R}^k$

$$\mathcal{E}\{g(X)\} := \int_{\mathbb{R}^p} g(x) f_X(x) dx := \int_{-\infty}^{\infty} ... \int_{-\infty}^{\infty} g(x_1, ..., x_p) f_X(x_1, ..., x_p) dx_1 ... dx_p$$

Covariance matrix and Variance

- Covariance matrix is a matrix
- Variance remains a scalar

$$\Sigma := \operatorname{cov}(X) = \operatorname{cov}(X, X) = E\{(X - \mu)(X - \mu)^T\}$$

$$Var(X) := trace(\Sigma) = \sum_{i=1}^{p} \sum_{i} cov(X, X) = E\{(X - \mu)^{T} (X - \mu)\}$$

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- Note is μ a vector
- Our trick is to get all relevant information from mean and covariance matrix so we can avoid working with the density, which requires a lot of integration

Covariance Matrix is positive semi-definite

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Exercise 5.19 $\Sigma := \text{cov}(X) = \text{cov}(X, X)$ is a positive semi-definite matrix.

Solution: For $v \in \mathbb{R}^p$, we need to show that $v^{\top} \Sigma v \geq 0$, where $\Sigma := \mathcal{E}\{(X - \mu) \cdot (X - \mu)^{\top}\}$.

$$v^{\top} \Sigma v := v^{\top} \mathcal{E}\{(X - \mu) \cdot (X - \mu)^{\top}\} v$$

$$= \mathcal{E}\{v^{\top} (X - \mu) \cdot (X - \mu)^{\top} v\}$$

$$= \mathcal{E}\{((X - \mu)^{\top} v)^{\top} \cdot ((X - \mu)^{\top} v)\}$$

$$= \mathcal{E}\{||(X - \mu)^{\top} v||^{2}\}$$

$$= \int_{\mathbb{R}^{p}} ||(X - \mu)^{\top} v||^{2} f_{X}(x) dx$$

$$\geq 0$$

because the integral of a non-negative function over \mathbb{R}^p is non-negative.

Covariance Matrix is positive semi-definite

• If the covariance matrix is **positive definite**, its inverse is the **information matrix**. The interpretation is that "high variance" means "low information" and vice versa.

Eigenvalues and Shape of Covariance Matrix

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Recall: you can always diagonalize a symmetric matrix with an orthogonal matrix

- What does a diagonal covariance matrix mean?
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- Recall: you can always diagonalize a symmetric matrix with an orthogonal matrix
- Subtle: Independence for random vectors → diagonal covariance matrix.
 Converse not true. Covariance matrix does not capture all information about the density.



Idea

• Least squares as a minimum distance problem

Idea

- Least squares as a minimum distance problem
- Estimation as a weighted least squares problem

Overdetermined equations

Too many equations (find best approximation)

$$\widehat{\alpha} = \underset{\alpha \in \mathbb{R}^m}{\arg \min} \|A\alpha - b\|^2 \iff (A^{\top}SA)\widehat{\alpha} = A^{\top}Sb \iff \widehat{\alpha} = (A^{\top}SA)^{-1}A^{\top}Sb$$

Underdetermined equations

- Too many solutions (too few equations)
 - Find "smallest" solution

$$\widehat{x} := \underset{Ax=b}{\operatorname{arg\,min}} ||x|| = \underset{Ax=b}{\operatorname{arg\,min}} ||x||^2$$

Underdetermined equations

- Too many solutions (too few equations)
 - Find "smallest" solution

$$\widehat{x} := \underset{Ax=b}{\operatorname{arg\,min}} ||x|| = \underset{Ax=b}{\operatorname{arg\,min}} ||x||^2$$

$$\widehat{x} = S^{-1}A^{\top}\beta$$
, $AS^{-1}A^{\top}\beta = b$ or, equivalently, $\widehat{x} = S^{-1}A^{\top} \left(AS^{-1}A^{\top}\right)^{-1}b$

Best Linear Unbiased Estimator

• Goal: How to choose the weight matrix in an overdetermined problem

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- Model: $y = Cx + \varepsilon$,
 - **Measurement** (model output) $y \in \mathbb{R}^m$
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 - **Noise** (output) $\varepsilon \in \mathbb{R}^m$

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holds for all $x \in \mathbb{R}^n$

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$$\hat{x} = Ky$$
 $E\{\hat{x} - x\} = 0$ $Var(\hat{x} - x) = E\{(\hat{x} - x)^T(\hat{x} - x)\}$ holds for all $x \in \mathbb{R}^n$ Minimizes variance

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$$cov(\hat{x} - x) = (C^T Q^{-1} C)^{-1}$$

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state and noise are uncorrelated $E\{xx^T\} = 0$

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holds for all $x \in \mathbb{R}^n$

Minimizes variance

• Find:
$$\hat{K}$$
 $\hat{K} = PC^T (CPC^T + Q)^{-1}$ $cov(\hat{x} - x) = P - PC^T (CPC^T + Q)^{-1}CP$