

(due Friday, October 29, by 5:00 p.m. CDT)

No credit will be given without supporting work.

7. Let $\psi > 0$ and let X_1, X_2, \dots, X_n be a random sample from a probability distribution with probability density function

$$f(x; \psi) = \frac{\psi}{2^\psi} \cdot (2-x)^{\psi-1}, \quad 0 < x < 2, \quad \text{zero otherwise.}$$

Recall: $E(X) = \frac{2}{\psi+1}$. A method of moments estimator of ψ is $\tilde{\psi} = \frac{2}{\bar{X}} - 1$.

The maximum likelihood estimator of ψ is

$$\hat{\psi} = \frac{n}{n \cdot \ln 2 - \sum_{i=1}^n \ln(2 - X_i)} = \frac{n}{\sum_{i=1}^n \left(-\ln \left(1 - \frac{X_i}{2} \right) \right)}.$$

$W = -\ln \left(1 - \frac{X}{2} \right)$ has an Exponential distribution with mean $\theta = \frac{1}{\psi}$.

- d) Is the method of moments estimator of ψ , $\tilde{\psi}$, an unbiased estimator of ψ ?

If $\tilde{\psi}$ is not an unbiased estimator of ψ , does $\tilde{\psi}$ underestimate or overestimate ψ (on average)?

Hint: $\tilde{\psi} = g(\bar{X})$. Is $g(x)$ a linear function?

If it is not a linear function, does it curve up or down?

- e) Is $\tilde{\psi}$ a consistent estimator of ψ ? *Justify your answer.*

(NOT enough to say “because it is a method of moments estimator”)

“Hint”: Start with the WLLN and \bar{X} .

- f) Is the maximum likelihood estimator $\hat{\psi}$ an unbiased estimator of ψ ?
If $\hat{\psi}$ is not an unbiased estimator of ψ , construct an unbiased estimator of ψ based on $\hat{\psi}$.

“Hint” 0: If U has a $\text{Gamma}(\alpha_1, \theta)$ distribution, V has a $\text{Gamma}(\alpha_2, \theta)$ distribution, U and V are independent, then $U + V$ has a $\text{Gamma}(\alpha_1 + \alpha_2, \theta)$ distribution.

“Hint” 1: $E(a \odot) = a E(\odot)$. “Hint” 2: $\frac{1}{\heartsuit} = \heartsuit^{-1}$.

“Hint” 3: If T_α has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, then

$$E(T_\alpha^m) = \frac{\theta^m \Gamma(\alpha + m)}{\Gamma(\alpha)} = \frac{\Gamma(\alpha + m)}{\lambda^m \Gamma(\alpha)}, \quad m > -\alpha.$$

- g) Find $\text{MSE}(\hat{\psi}) = (\text{bias}(\hat{\psi}))^2 + \text{Var}(\hat{\psi})$.

“Hint” 1: $\text{bias}(\hat{\psi}) = E(\hat{\psi}) - \psi$. You have $E(\hat{\psi})$ from part (f).

“Hint” 2: $\text{Var}(a \odot) = a^2 \text{Var}(\odot)$. $\text{Var}(\odot) = E(\odot^2) - [E(\odot)]^2$.

“Hint” 3: If T_α has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, then

$$E(T_\alpha^m) = \frac{\theta^m \Gamma(\alpha + m)}{\Gamma(\alpha)} = \frac{\Gamma(\alpha + m)}{\lambda^m \Gamma(\alpha)}, \quad m > -\alpha.$$

- h) Is $\hat{\psi}$ a consistent estimator of ψ ? *Justify your answer.*
(NOT enough to say “because it is the maximum likelihood estimator”)

“Hint”: Start with the WLLN and \bar{W} .

8. Let $\xi > 0$ and let X_1, X_2, \dots, X_n be a random sample from a probability distribution with probability density function

$$f(x; \xi) = \frac{1}{2} \xi^4 x^{11} e^{-\xi x^3}, \quad x > 0, \quad \text{zero elsewhere.}$$

Recall: The maximum likelihood estimator of ξ is $\hat{\xi} = \frac{4n}{\sum_{i=1}^n X_i^3}$.

$W = X^3$ has a Gamma ($\alpha = 4, \theta = \frac{1}{\xi}$) distribution.

- c) Is the maximum likelihood estimator $\hat{\xi}$ an unbiased estimator of ξ ?

If $\hat{\xi}$ is not an unbiased estimator of ξ , construct an unbiased estimator of ξ based on $\hat{\xi}$.

- d) Find $\text{MSE}(\hat{\xi}) = (\text{bias}(\hat{\xi}))^2 + \text{Var}(\hat{\xi})$.

- e) Is $\hat{\xi}$ a consistent estimator of ξ ? *Justify your answer.*

(NOT enough to say “because it is the maximum likelihood estimator”)

- f) Suppose $n = 5$ and $\xi = 1.5$. Find the probability $P\left(\sum_{i=1}^5 X_i^3 > 10\right)$.