

348HWw5

Shay Lebovitz

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#5 a)

```
paper <- read.table('/Users/shaylebovitz/R/paper.txt', header = TRUE)
(cor_mat <- cor(paper))
```

```
##          bl          em          sf          bs
## bl 1.0000000 0.9138256 0.9838790 0.9875554
## em 0.9138256 1.0000000 0.9422199 0.8746665
## sf 0.9838790 0.9422199 1.0000000 0.9745114
## bs 0.9875554 0.8746665 0.9745114 1.0000000
```

```
apply(paper, 2, sd)
```

```
##          bl          em          sf          bs
## 2.8814703 0.7164910 1.4628895 0.6930166
```

bl has a high standard deviation, sf is medium, and em and bs have low standard deviations. All of the variables are very highly correlated with each other, the lowest correlation being bs and em with a correlation of 0.8747, which is still very high.

b) I think that the PC analysis should be based on the correlation matrix because there is a significant spread in the variable standard deviations, and they are also measured in different units.

c)

```
paper_pca <- prcomp(paper, scale = TRUE)
paper_pca$rotation
```

```
##          PC1          PC2          PC3          PC4
## bl 0.5061685 0.26110200 -0.56517738 -0.5968196
## em 0.4854922 -0.81904792 -0.19350510 0.2366720
## sf 0.5080684 0.02020866 0.80019598 -0.3180323
## bs 0.4999573 0.51046828 -0.05307262 0.6976017
```

```
summary(paper_pca)
```

```
## Importance of components:
```

```
##          PC1          PC2          PC3          PC4
## Standard deviation    1.9595 0.37457 0.11227 0.0871
## Proportion of Variance 0.9599 0.03508 0.00315 0.0019
## Cumulative Proportion 0.9599 0.99495 0.99810 1.0000
```

Based on this analysis, just using the first principal component explains 96% of the variance. By the second PC, it is already not 'pulling its weight' in that it explains less than 1/4 of the variance. For these reasons, I would just take the first PC in this case.

d)

```
paper_sd <- apply(paper, 2, sd)
10*paper_pca$rotation[,1]/paper_sd
```

```
##          bl          em          sf          bs
## 1.756633 6.775970 3.473047 7.214218
```

Here we see most weight is on em and bs, which have the smallest standard deviations.

```
paper_pca_scores <- predict(paper_pca)
head(paper_pca_scores)
```

```
##          PC1          PC2          PC3          PC4
## [1,] -0.3968987 0.15425573 -0.021786814 -0.009607294
## [2,] -0.5310917 0.16707751 -0.028802579 -0.049526457
## [3,] -0.9084521 0.25320477 0.035755303 -0.104024036
## [4,] -1.6010530 0.17423169 0.012458192 -0.025313914
## [5,] -1.1137187 0.08777323 0.014111553 -0.178803281
## [6,] -0.7434306 0.13672636 -0.004906303 -0.053276700
```

```
cor(paper, paper_pca_scores)
```

```
##          PC1          PC2          PC3          PC4
## bl 0.9918199 0.097801441 -0.063450892 -0.05198256
## em 0.9513052 -0.306792239 -0.021724279 0.02061396
## sf 0.9955425 0.007569592 0.089835776 -0.02770039
## bs 0.9796492 0.191207016 -0.005958315 0.06076061
```

We see that the first principal component is extremely highly correlated with all the variables, and the rest of the PCs are essentially non-correlated with any. I should also note that because the PC1 coefficients are so similar for every variable, it is essentially an average of the variables.

#6 a)

```
emplmnt <- read.table('/Users/shaylebovitz/R/employment.txt', header = TRUE)
emplmnt_pca <- prcomp(~AGR + MAN + CON + SER + FIN + SPS + TC,
                      data = emlmnt, scale = TRUE)
emplmnt_pca$rotation
```

```
##          PC1          PC2          PC3          PC4          PC5          PC6
## AGR 0.5982957 -0.06475831 0.03837215 -0.06426076 0.088358609 0.02948154
## MAN -0.1166513 0.41333890 0.29670088 0.82325922 -0.114998870 -0.02958918
## CON -0.2660980 -0.23118320 0.72664207 -0.29334483 -0.473901175 -0.07079849
## SER -0.4377825 -0.37921662 0.19288538 0.11736096 0.688108273 0.30824232
## FIN -0.1286410 -0.55516778 -0.43075986 0.32287232 -0.510847542 0.32195688
## SPS -0.5172193 0.10225300 -0.37282378 -0.11254728 -0.005356358 -0.61880152
## TC -0.2861972 0.55591317 -0.14398285 -0.31839692 -0.141063088 0.64160581
##          PC7
## AGR -0.7896647
## MAN -0.1888280
## CON -0.1791411
## SER -0.2122653
## FIN -0.1442851
## SPS -0.4329204
## TC -0.2353450
```

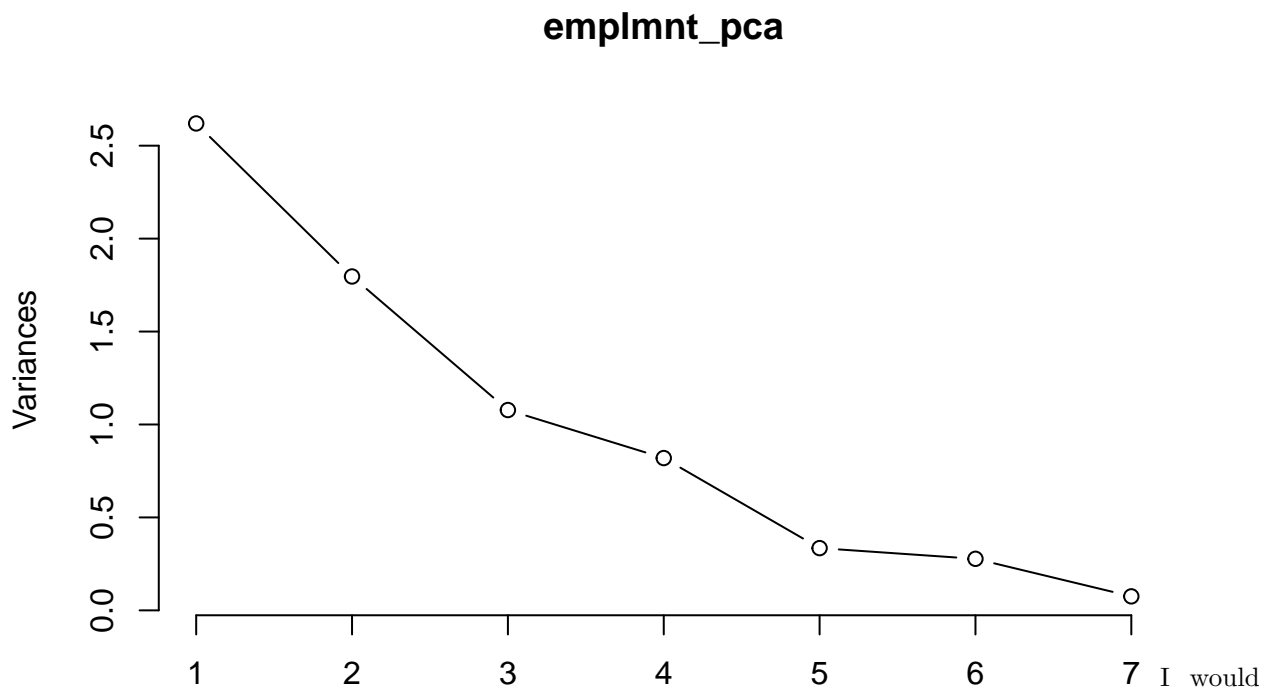
```
emplmnt <- emlmnt[, -8]
```

b)

```
summary(emplmnt_pca)
```

```
## Importance of components:
##          PC1      PC2      PC3      PC4      PC5      PC6      PC7
## Standard deviation  1.6185 1.3402 1.0381 0.9051 0.57866 0.52653 0.27442
## Proportion of Variance 0.3742 0.2566 0.1540 0.1170 0.04784 0.03961 0.01076
## Cumulative Proportion 0.3742 0.6308 0.7848 0.9018 0.94964 0.98924 1.00000
```

```
plot(emplmnt_pca, type = 'l')
```



I would choose the first 4 principal components, for a few reasons. Firstly, using the first four accounts for over 90% of the error, which is should be adequate for any analysis. Second, based on the scree plot, we see that the variance significantly plateaus after PC4, meaning they are not very useful. Choosing 3 PCs would work well too, as PC4 doesn't 'pull its weight', as its variance is less than 1/7.

c)

```
emplmnt_sd <- apply(emplmnt, 2, sd)
100*emplmnt_pca$rotation[,1]/emplmnt_sd
```

```
##          AGR      MAN      CON      SER      FIN      SPS      TC
##  4.861465 -1.233519 -9.736175 -8.483897 -3.226770 -5.923220 -23.204533
```

Here we see that only AGR is positive, the rest are negative.

```
emplmnt_pca_scores <- predict(emplmnt_pca)
head(emplmnt_pca_scores)
```

```
##          PC1      PC2      PC3      PC4      PC5      PC6
## 1 -1.1931943  0.07086769 -0.9785269  0.20434044 -0.001740173 -0.27339172
## 2 -0.8723157  0.23293159 -1.0857976  0.07703399 -0.386479020 -0.23204432
## 3 -0.8556792 -0.45515661 -0.7374718  0.32236092 -0.286008258 -0.11585267
## 4 -0.7877522 -0.80746205  0.3900067  0.71575214 -0.516455050 -0.29606174
## 5  0.7197939  0.07827075  0.2997098 -0.14216014  0.679818185  0.83251243
## 6  0.1263532 -0.70803017 -0.0927552  0.37419949  0.231938444  0.05303738
##          PC7
```

```
## 1 0.002028068
## 2 -0.113200353
## 3 0.019824050
## 4 0.285948904
## 5 -0.358136025
## 6 -0.019239188
```

```
cor(emplmnt, emplmnt_pca_scores)
```

```
##          PC1          PC2          PC3          PC4          PC5          PC6
## AGR  0.9683415 -0.08678933  0.03983426 -0.05816447  0.05112959  0.01552303
## MAN -0.1888001  0.55395834  0.30800618  0.74515825 -0.06654525 -0.01557971
## CON -0.4306797 -0.30983260  0.75432957 -0.26551578 -0.27422765 -0.03727781
## SER -0.7085509 -0.50822754  0.20023496  0.10622716  0.39818073  0.16230007
## FIN -0.2082055 -0.74403794 -0.44717326  0.29224206 -0.29560704  0.16952124
## SPS -0.8371194  0.13703985 -0.38702962 -0.10187015 -0.00309951 -0.32582004
## TC  -0.4632101  0.74503691 -0.14946908 -0.28819123 -0.08162757  0.33782728
##
##          PC7
## AGR -0.21669623
## MAN -0.05181732
## CON -0.04915909
## SER -0.05824890
## FIN -0.03959406
## SPS -0.11880006
## TC  -0.06458231
```

We see that AGR is highly correlated with PC1, and the rest are negatively correlated. PC2 shows strong correlation with TC and strong negative correlation with FIN. PC3 really only shows strong correlation with CON.

#7 a)

```
pollution <- read.table('/Users/shaylebovitz/R/pollution.txt', header = TRUE)
(pollution_sd <- apply(pollution, 2, sd))
```

```
##          CO          NO          NO2          O3          HC
## 1.2337209 1.0873574 3.3709837 5.5658345 0.6917466
```

```
cor(pollution)
```

```
##          CO          NO          NO2          O3          HC
## CO  1.0000000  0.5021525  0.5565838  0.4109288  0.1660323
## NO  0.5021525  1.0000000  0.2968981 -0.1339521  0.2347043
## NO2 0.5565838  0.2968981  1.0000000  0.1666422  0.4477678
## O3  0.4109288 -0.1339521  0.1666422  1.0000000  0.1544506
## HC  0.1660323  0.2347043  0.4477678  0.1544506  1.0000000
```

b)

```
(pollution_pca_cov <- prcomp(pollution, scale = F))
```

```
## Standard deviations (1, ..., p=5):
## [1] 5.6410664 3.3862185 1.1984182 0.7293890 0.5183157
##
## Rotation (n x k) = (5 x 5):
##          PC1          PC2          PC3          PC4          PC5
## CO  -0.10343698  0.18207274  0.62534826 -0.578141442  0.48046054
## NO   0.01778827  0.12836692  0.74801509  0.500285034 -0.41640584
```

```
## NO2 -0.16191213  0.95490887 -0.22085443 -0.004444182 -0.11461711
## O3  -0.98090512 -0.17661949 -0.01542742  0.054814906 -0.05820662
## HC  -0.02437144  0.08559241 -0.01995722  0.642217201  0.76107736
```

```
(pollution_pca_cor <- prcomp(pollution, scale = T))
```

```
## Standard deviations (1, ..., p=5):
## [1] 1.4875002 1.0644756 0.9522246 0.7415103 0.4445961
##
## Rotation (n x k) = (5 x 5):
##          PC1          PC2          PC3          PC4          PC5
## CO  -0.5621560  0.10573277 -0.4645354 -0.07701329 -0.6716227
## NO  -0.4106208 -0.57912175 -0.3241413  0.45949883  0.4240304
## NO2 -0.5394732 -0.03035633  0.2155503 -0.71908494  0.3801343
## O3  -0.2689306  0.80581727 -0.1087780  0.34033332  0.3881693
## HC  -0.3898924 -0.05635245  0.7879370  0.38732397 -0.2719260
```

O3 has the highest standard deviation, and in the unscaled PC1, carries practically all the weight of the principal component. Using the scaled version, we see that the coefficients of PC1 are much more even across variables, with O3 now being the smallest. Based on the correlation matrix, O3 is the least correlated variable.

c)

```
pollution_pca_scaled <- prcomp(~I(CO/35) + I(NO/25) + I(NO2/5) + I(O3/7.5) +
                                I(HC/25), data = pollution, scale = FALSE)
pollution_pca_scaled$rotation
```

```
##          PC1          PC2          PC3          PC4          PC5
## I(CO/35) -0.026627863  0.01495777 -0.436121552  0.478471590  0.76153145
## I(NO/25) -0.001081216  0.02398402 -0.897456221 -0.176704632 -0.40344938
## I(NO2/5) -0.501023048  0.86458853  0.031632813  0.006518366 -0.02048055
## I(O3/7.5) -0.864945991 -0.50149247 -0.002999093 -0.006768663 -0.01785851
## I(HC/25) -0.011581302  0.01389266 -0.057971489 -0.860088430  0.50651759
```

```
summary(pollution_pca_scaled)
```

```
## Importance of components:
##          PC1          PC2          PC3          PC4          PC5
## Standard deviation  0.7742 0.6378 0.04430 0.02637 0.01829
## Proportion of Variance 0.5939 0.4031 0.00194 0.00069 0.00033
## Cumulative Proportion 0.5939 0.9970 0.99898 0.99967 1.00000
```

I would say that the first 2 PCs are needed, as the first one only covers roughly 60% which is probably not adequate. The first two cover >99% of the variance, which is certainly adequate. NO2 and O3 are the only variables with significant PC1 and PC2 coefficients.

#8 a)

```
qbs <- read.table('/Users/shaylebovitz/R/QBs.txt', header = TRUE)
apply(qbs, 2, sd)
```

```
##      Comp      TD      Int      YPA      Rate
## 3.9040116 1.4322015 1.0264446 0.7053741 12.9562610
```

Because the standard deviations vary so much, I will perform the analysis based on the correlation matrix

```
qbs_pca <- prcomp(~Comp + TD + Int + YPA, data = qbs, scale = TRUE)
qbs_pca$rotation
```

```
##          PC1          PC2          PC3          PC4
```

```
## Comp -0.4782130  0.43465731  0.7630190  0.01368644
## TD   -0.5533956 -0.05905053 -0.3268963  0.76380962
## Int   0.4441221  0.82079809 -0.1946390  0.30192998
## YPA  -0.5175144  0.36589145 -0.5225480 -0.57030328
```

```
summary(qbs_pca)
```

```
## Importance of components:
```

```
##           PC1      PC2      PC3      PC4
## Standard deviation    1.6890 0.7796 0.63236 0.37390
## Proportion of Variance 0.7131 0.1519 0.09997 0.03495
## Cumulative Proportion 0.7131 0.8651 0.96505 1.00000
```

b) From the principal components analysis, we see that about 71% of the total variation can be explained by PC1. That, along with the fact that PC2 only explains about 15% of the variation and thus doesn't carry its weight, leads me to believe that just one PC will be adequate for the analysis.

c)

```
qb_pca_scores <- predict(qbs_pca)
qb_pca_scores[,1]
```

```
##           1           2           3           4           5           6
## -3.65831935 -0.38348573 -2.38308495  0.35446489  0.11556998  1.79272556
##           7           8           9          10          11          12
##  0.71186252 -0.27533342 -0.68063062  0.12046717  0.21251653  2.09483279
##          13          14          15          16          17          18
## -2.58825500 -1.17325054  0.03052064  1.44379074 -0.21377860  0.63054126
##          19          20          21          22          23          24
##  2.44837145  0.63175117 -0.55726141 -1.91096316  1.69559675 -0.41841303
##          25          26          27          28          29          30
##  0.55075158  1.43170346 -4.18523072  1.32115054 -2.55659500  2.20824872
##          31          32          33          34          35          36
##  2.27635849 -0.17420932  0.62522057  1.22807276  1.08248146  1.02623722
##          37
## -2.87442540
```

```
cor(qbs$Rate, qb_pca_scores[,1])
```

```
## [1] -0.9963963
```

We see that there is a very strong negative correlation between the first principal component and the quarterback rate. This is surprising as Rate has the highest standard deviation of all the variables, so I'm not exactly sure on how to analyze this. #9) a)

```
properties <- read.table('/Users/shaylebovitz/R/properties.txt', header = T)
cor(properties)
```

```
##           BATH      LOT      SIZE      GAR      ROOM      BED
## BATH  1.0000000  0.4129558  0.7285916  0.22402204  0.5103104  0.4264014
## LOT   0.4129558  1.0000000  0.5715520  0.20466375  0.3921244  0.1516093
## SIZE  0.7285916  0.5715520  1.0000000  0.35888351  0.6788606  0.5743353
## GAR   0.2240220  0.2046638  0.3588835  1.00000000  0.5893871  0.5412988
## ROOM  0.5103104  0.3921244  0.6788606  0.58938707  1.0000000  0.8703883
## BED   0.4264014  0.1516093  0.5743353  0.54129880  0.8703883  1.0000000
## AGE  -0.1007485 -0.3527514 -0.1390869 -0.02016883  0.1242663  0.3135114
## Y     0.7097771  0.6476364  0.7077656  0.46146792  0.5284436  0.2815200
##           AGE      Y
```

```
## BATH -0.10074847  0.7097771
## LOT  -0.35275139  0.6476364
## SIZE -0.13908686  0.7077656
## GAR  -0.02016883  0.4614679
## ROOM  0.12426629  0.5284436
## BED   0.31351144  0.2815200
## AGE   1.00000000 -0.3974034
## Y     -0.39740338  1.0000000
```

We see high correlation between bathroom number and house size, and between room number and house size. This intuitively makes a lot of sense, bigger houses tend to have more bedrooms and bathrooms. All the other variables are somewhat correlated, most ranging from 0.2 to 0.6, besides the ones mentioned previously. For example, garage size and lot size only have a correlation of 0.204, whereas garage size and room number have a correlation of 0.589.

b)

```
properties_reg <- lm(Y ~ BATH + LOT + SIZE + GAR + ROOM + BED + AGE,
                     data = properties)
summary(properties_reg)
```

```
##
## Call:
## lm(formula = Y ~ BATH + LOT + SIZE + GAR + ROOM + BED + AGE,
##     data = properties)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.8792 -1.7515 -0.2857  1.6013  6.1100
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  11.62410    5.52343   2.105  0.0515 .
## BATH         10.85551    3.97518   2.731  0.0148 *
## LOT           0.54253    0.46215   1.174  0.2576
## SIZE          3.92170    4.48393   0.875  0.3947
## GAR           2.94143    1.37361   2.141  0.0480 *
## ROOM          2.39374    1.86928   1.281  0.2186
## BED          -4.78958    2.79845  -1.712  0.1063
## AGE          -0.06972    0.05665  -1.231  0.2362
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.13 on 16 degrees of freedom
## Multiple R-squared:  0.8115, Adjusted R-squared:  0.7291
## F-statistic: 9.843 on 7 and 16 DF,  p-value: 8.944e-05
```

We get a fairly high R-squared and adjusted R-squared, showing that the response variables explain the sale price fairly well. However, we see that only BATH and GAR have statistically significant beta values. BATH has a large estimated beta, where as the rest are small, and BED and AGE having negative betas. This seems right for age (an older house should cost less) but not for bed (more bedrooms should cost more).

c)

```
properties_reg2 <- lm(Y ~ BATH + GAR, data = properties)
summary(properties_reg2)
```

```
##
## Call:
## lm(formula = Y ~ BATH + GAR, data = properties)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -5.2211 -3.1169 -0.0322  1.8592 11.3342
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   11.872      4.205   2.824 0.010178 *
## BATH          15.943      3.536   4.509 0.000192 ***
## GAR           3.167      1.408   2.249 0.035372 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.979 on 21 degrees of freedom
## Multiple R-squared:  0.6001, Adjusted R-squared:  0.562
## F-statistic: 15.76 on 2 and 21 DF,  p-value: 6.614e-05

properties_reg3 <- lm(Y ~ BATH + GAR + AGE, data = properties)
summary(properties_reg3)
```

```
##
## Call:
## lm(formula = Y ~ BATH + GAR + AGE, data = properties)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.5769 -2.4129 -0.6249  1.4750  9.1502
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  18.12865    4.34308   4.174 0.000468 ***
## BATH         15.10853    3.11188   4.855 9.6e-05 ***
## GAR          3.17542    1.23329   2.575 0.018086 *
## AGE         -0.14133    0.05202  -2.717 0.013279 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.485 on 20 degrees of freedom
## Multiple R-squared:  0.7079, Adjusted R-squared:  0.6641
## F-statistic: 16.16 on 3 and 20 DF,  p-value: 1.434e-05
```

In both of these regressions, the R-squared and adjusted R-squared values drop, to 0.6001 and 0.562 in `reg2`, and 0.7079 and 0.6641 in `reg3`. We see that `reg3`, which includes `AGE`, does slightly better than `reg2`, showing that age is a valuable predictor. Overall, the drops in R-squared for these models is not too great for the amount of simplification you get from only using 2 or 3 predictors instead of 7.

d)

```
properties_pca <- prcomp(~LOT + SIZE + GAR + ROOM + BED, data = properties, scale = TRUE)
properties_pca$rotation
```

```
##           PC1           PC2           PC3           PC4           PC5
## LOT  0.3096733 -0.7731703 -0.2949459 -0.4222795 -0.202477293
```



```
## SIZE 0.4704060 -0.3395186 0.2899932 0.7611254 0.006113792
## GAR 0.3965764 0.3587489 -0.8130270 0.2250860 -0.048475150
## ROOM 0.5364806 0.1406204 0.1844288 -0.3450071 0.734418524
## BED 0.4875822 0.3721020 0.3659016 -0.2695822 -0.645945167
```

#interpret PC1

For the most part, PC1 is about the average of the five variables, give and take a few percent.

e)

```
PC1 = predict(properties_pca)[,1]
properties_reg4 <- lm(Y ~ BATH + PC1 + AGE, data = properties)
summary(properties_reg4)
```

```
##
## Call:
## lm(formula = Y ~ BATH + PC1 + AGE, data = properties)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -5.8810 -2.3283 -0.2991  2.2205  7.7952
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  27.76255    5.09594   5.448 2.48e-05 ***
## BATH         10.86330    3.76480   2.885 0.00915 **
## PC1           1.37620    0.51665   2.664 0.01492 *
## AGE          -0.15503    0.05184  -2.991 0.00723 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.455 on 20 degrees of freedom
## Multiple R-squared:  0.7129, Adjusted R-squared:  0.6699
## F-statistic: 16.56 on 3 and 20 DF, p-value: 1.209e-05
```

Here, we see that all three of the predictor variables have statistically significant betas. BATH still appears to be the strongest predictor with a beta of 10.86, while PC1 has a smaller beta of 1.38 and AGE has a negative beta of -0.16, which is expected since older houses should cost less. The R-squared and adjusted R-squared are relatively high, higher than the other two alternate models `reg2` and `reg3`, though not as high as when every variable was used as a predictor. This shows us that using PC1 as a predictor does an adequate job for this analysis, as it allows you to reduce the number of predictors but maintain predictive power.