



Bitcoin Futures—What use are they?

Shaen Corbet^{a,*}, Brian Lucey^b, Maurice Peat^c, Samuel Vigne^d

^a DCU Business School, Dublin City University, Dublin 9, Ireland

^b Trinity Business School, Trinity College Dublin, Dublin 2, Ireland

^c University of Sydney Business School, Sydney, New South Wales, Australia

^d Queen's Management School, Queen's University Belfast, BT9 5EE, Northern Ireland, United Kingdom

HIGHLIGHTS

- This article investigates the effect of the introduction of Bitcoin futures.
- The introduction of Bitcoin futures has increased the spot volatility of Bitcoin.
- Bitcoin futures are not an effective hedging tool.
- Price discovery is driven by uninformed investors in the spot market.
- Bitcoin futures did not affect the nature of Bitcoin as a speculative asset rather than a currency.

ARTICLE INFO

Article history:

Received 16 June 2018

Received in revised form 11 July 2018

Accepted 21 July 2018

Available online 7 August 2018

Keywords:

Bitcoin
Cryptocurrencies
Futures markets
Volatility
Speculative assets
Currencies

ABSTRACT

Early analysis of Bitcoin concluded that it did not meet the economic conditions to be classified as a currency. Since this conclusion, interest in Bitcoin has increased substantially. We investigate whether the introduction of futures trading in Bitcoin is able to resolve the issues that stopped Bitcoin from being considered a currency. Our analysis shows that spot volatility has increased following the appearance of futures contracts, that futures contracts are not an effective hedging instrument, and that price discovery is driven by uninformed investors in the spot market. We therefore argue that the conclusion that Bitcoin is a speculative asset rather than a currency is not altered by the introduction of futures trading.

© 2018 Elsevier B.V. All rights reserved.

1. Introduction

An early analysis of Bitcoin by Yermack (2015) concluded that it was not a currency but rather a speculative asset. The author argued that Bitcoin failed to satisfy the functions of money: a medium of exchange, unit of account, and store of value. The idea that Bitcoin has no intrinsic value, is supported by others such as Cheah and Fry (2015), but an open discussion remains as to the economic value of Bitcoin (Demir et al., 2018). More details on the advances of Bitcoin literature can be found in Lucey et al. (2018). A recent innovation in the Bitcoin trading environment is the introduction of futures contracts by the Chicago Mercantile Exchange (CME) and the Chicago Board Options Exchange (CBOE) in December 2017. The high volatility of Bitcoin prices was identified by Yermack as a feature which lead to Bitcoin not being a useful unit of account. We examine the relationship between futures and spot prices, finding that by contrast to the norm, cash leads the

futures. This we surmise is related to the very high volatility of bitcoin.

In December 2017 trading in futures contracts on Bitcoins commenced on the Chicago Mercantile Exchange and the Chicago Board Options Exchange. On the 1st of December, both exchanges announced a Bitcoin futures contract. The CBOE contract commenced trading on the 10th of December, with each contract being for one Bitcoin. Three aspects of the introduction of futures on the spot market will be explored. Firstly, the impact of futures trading on spot volatility is examined. Secondly, the hedging effectiveness of futures contracts is evaluated. Finally, the flow of information between spot and futures markets is documented.

2. Data

Both the CME and the CBOE future contracts are cash settled in US Dollars. Table 1 displays stylised facts of these two contracts using data sampled at a one-minute frequency from CBOE futures contract sourced from Thomson Reuters Tick History, and Bitcoin price data from Thomson Reuters Eikon. We explore the impact of

* Corresponding author.

E-mail address: shaen.corbet@dcu.ie (S. Corbet).

Table 1
Stylised facts based on CBOE and CME Bitcoin Futures.

Variable	CBOE Futures	CME Futures
Product Code	XBT	BTC
First Traded	10th of December 2017	18th of December 2017
Contract unit	1 Bitcoin	5 Bitcoins
Minimum Price Fluctuation	10.00 points USD/XBT (equal to \$10.00 per contract).	\$5.00 per bitcoin = \$25.00 per contract.
Position Limits	A person: (i) may not own or control more than 5,000 contracts net long or net short in all XBT futures contract expirations combined and (ii) may not own or control more than 1,000 contracts net long or net short in the expiring XBT futures contract, commencing at the start of trading hours 5 business days prior to the Final Settlement Date of the expiring XBT futures contract.	1,000 contracts with a position accountability level of 5,000 contracts.
Price Limits	XBT futures contracts are not subject to price limits.	7% above and below settlement price, +/-13% previous settlement, +/-20% for prior settlement.
Settlement	The Final Settlement Value of an expiring XBT futures contract shall be the official auction price for Bitcoin in U.S. dollars determined at 4:00 p.m. Eastern Time on the Final Settlement Date by the Gemini Exchange Auction.	Cash settled by reference to Final Settlement Price.

Table 2
Descriptive statistics for bitcoin prices and returns.

Panel A—Full Sample	Price	Return
Mean	9,862.048	4.26E–06
Standard Error	8,579,189	4.33E–06
Median	9,291.53	1.21E–06
Mode	15,000	0.000000
Standard Deviation	3,984.44	0.002009
Sample Variance	15,875,760	4.04E–06
Kurtosis	–0.89573	11.46425
Skewness	0.39184	–0.08776
Range	15,800.5	0.069144
Minimum	3,865.23	–0.03291
Maximum	19,665.73	0.036236
Count	215,696	215,696
Panel B—Pre Futures Introduction	Price	Return
Mean	7,812.788	1.3E–05
Standard Error	10,39188	3.96E–06
Median	6,671.42	1.1E–05
Mode	16,500	0.000000
Standard Deviation	3,559.035	0.001357
Sample Variance	12,666,728	1.84E–06
Kurtosis	0.845098	26.04
Skewness	1.322531	–0.43248
Range	14,152.89	0.053846
Minimum	3,865.23	–0.03166
Maximum	18,018.12	0.022191
Count	117,294	117,294
Panel C—Post Futures Introduction	Price	Returns
Mean	12304.74	–6.1E–06
Standard Error	9,418,187	8.22E–06
Median	11,683.09	0.000000
Mode	15,000	0.000000
Standard Deviation	2,954.4	0.00258
Sample Variance	8,728,479	6.66E–06
Kurtosis	–0.58609	6.020228
Skewness	0.302747	–0.00718
Range	13,741.01	0.069144
Minimum	5,924.72	–0.03291
Maximum	19,665.73	0.036236
Count	98,402	98,402

the introduction of risk management tools on the pricing and risk characteristics of the spot Bitcoin market. From the one-minute transaction prices we calculate the log return for each period, presented in Fig. 1.

The characteristics of Bitcoin data covering the period from the 26st of September 2017 to the 22nd of February 2018 can be found in Table 2. Statistics for the full period and for sub-samples before and after the introduction of futures trading are also presented.

There has been a clear change in the distributional characteristics of Bitcoin returns since the introduction of futures. The mean changed in sign and the standard deviation doubled; this change in volatility is evident in the time series plot of returns.

3. Analysis

The impact of the introduction of futures trading on volatility in the underlying spot market has been investigated for stocks, foreign exchange, interest rates and commodities. The empirical evidence is mixed. Gulen and Mayhew (2000) found a noticeable increase in volatility in the U.S. and Japan, but a negligible effect or decreases in volatility for the remaining 23 markets. A recent study of the introduction of futures on European real estate indices by Lee et al. (2014) found that the volatility of the indices fell after the introduction of the futures contracts.

We apply tests from the process control literature. These are fully described in Ross et al. (2011) and Ross et al. (2015) to which the interested reader is kindly directed. The R statistical package *cpm* from Ross et al. (2015) is used for all estimation.

Two nonparametric statistics are computed, the Mood statistic for changes in volatility (scale) and a Lepage-(type) statistic which tests for changes in location and scale, the results of which are presented in Fig. 2.

Both the Mood and Lepage statistics indicate a significant change in the distribution, driven by the increase in volatility. The date of the change is the 29th of November 2017, two days before the official announcement of the commencement dates for futures trading. As returns for financial assets have often been found to be non independent and identically distributed, the analysis was run on both the raw returns and residuals from a ARMA(1,1)-GARCH (1,1). A significant change in the distribution, associated with the increase in series volatility was detected on the 29th of November 2017 in each case.

We then measure the extent of risk reduction that can be obtained by forming hedge portfolios. It is possible that an appropriately constructed hedge portfolio can be used to manage the volatility of Bitcoin prices. Hedging literature such as Figlewski (1984), Kroner and Sultan (1993), Park and Switzer (1995) and Choudhry (2003), concludes that hedge ratios selected by OLS generally work best when evaluated in sample: we therefore analyse naïve and OLS based hedging strategies. The effectiveness of the hedge can be measured by the percentage reduction volatility that results from holding the hedge portfolio. We also compute hedge effectiveness using semi-variance, which measures the variability of returns below the mean, addressing a shortcoming of the variance and providing a more intuitive measure of risk for hedging focused on downside risk.

Two hedging approaches are evaluated. The first is the naïve hedge which is a portfolio with one short futures position for every Bitcoin position, while the second approach is the Ordinary Least Squares (OLS) hedge. A simple OLS regression of the form $r_{spot} = \alpha + \beta_{future}$ is run. The estimated β is used as the hedge ratio. This

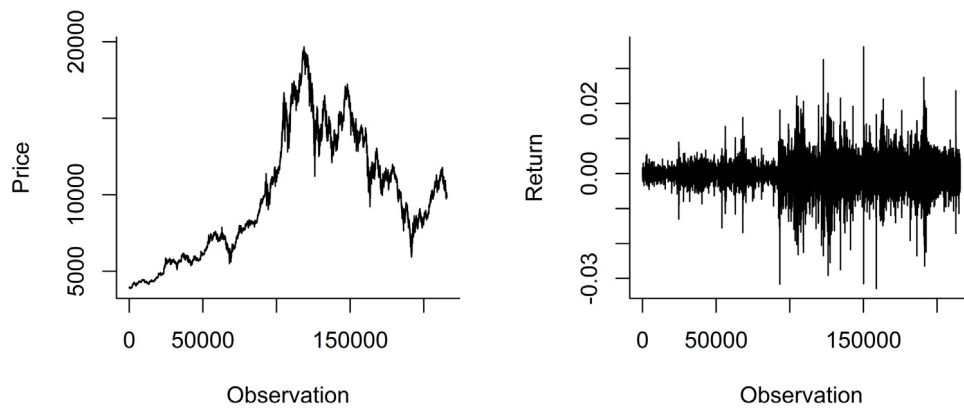


Fig. 1. Price and returns time series over the full sample period.

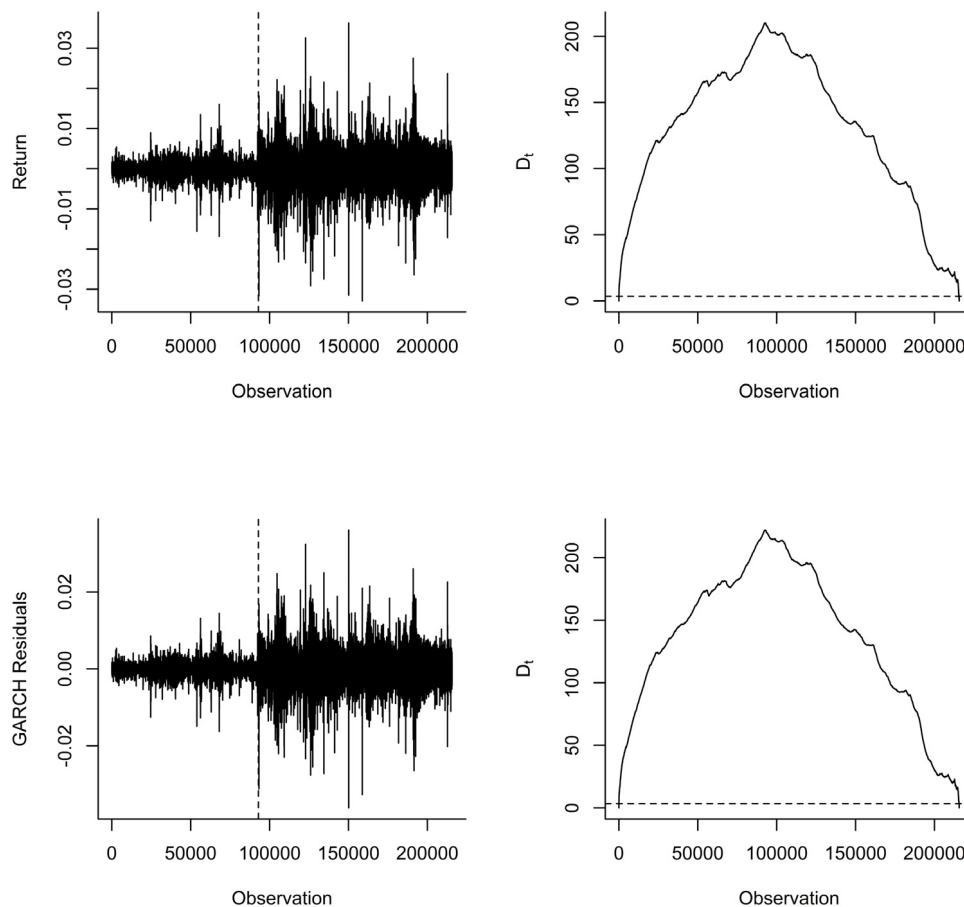


Fig. 2. Change point detection. Note: The above figure presents the Raw Returns Mood Statistics (Top Panel) and GARCH(1,1) Residuals Mood Statistic (Bottom Panel) respectively. These two nonparametric statistics represent the Mood statistic for change in volatility (scale) and a Lepage type statistic which tests for a change in location and scale. The implementation of these statistics for change point detection was executed relying on the *cpm* package in *R* and is used to establish the existence and location of a change point in the Bitcoin price series. Both the Mood and Lepage statistics indicate there is a significant change in the distribution, driven by the increase in volatility. The date of the change is the 29th of November 2017, two days before the official announcement of the commencement dates for futures trading.

approach to hedging is implemented using a rolling regression framework. Here, β is estimated each day then used to compute the hedge portfolio return for the next day. The return series for the hedge is the concatenation of each days hedge portfolio return. Table 3 contains the results of the evaluation of hedge effectiveness.

The first and most striking result is that hedging increases risk, as indicated by the negative sign of the effectiveness and risk reduction results. While the rolling OLS hedge is more effective

than the naïve hedge, as it would be expected, it also increases the pricing risk inherent in physically holding Bitcoin. Using semi-variance in the computation of hedge effectiveness shows an improvement in effectiveness compared to the use of the variance. However, both the hedging strategies are shown to be increasing risk under all evaluation methods.

It is generally accepted that futures contracts lead their respective underlying assets in price discovery: see Bohl et al. (2011), Rosenberg and Traub (2009), Cabrera et al. (2009) and Hauptfleisch

Table 3

Hedge effectiveness.

Naive Hedge	
Risk Reduction	−334.59
Hedge Effectiveness	−3.3459
Hedge Effectiveness (semivariance)	−1.20851
Rolling OLS Hedge	
Risk Reduction	−60.7627
Hedge Effectiveness	−0.60763
Hedge Effectiveness (semivariance)	−0.38919

et al. (2016). These results highlight the importance of market structure and instrument type. The findings of these studies indicate that the centralisation and relative transparency of futures markets contribute to their large role in price discovery. It is also likely that low transaction costs, inbuilt leverage, ease of shorting, and the ability to avoid holding the underlying physical asset make futures contracts an attractive alternative for traders in a wide range of assets. Bohl et al. (2011) argue that the emergence of futures markets generally coincides with the rise of institutional trading. The trades of sophisticated institutional investors contribute to price discovery being focused in futures markets.

There are two standard measures of price discovery commonly employed in the literature: the Hasbrouck (1995) Information Share (IS) and the Gonzalo and Granger (1995) Component Share (CS) measure. Hasbrouck (1995) demonstrates that the contribution of a price series to price discovery (the ‘information share’) can be measured by the proportion of the variance in the common efficient price innovations that is explained by innovations in that price series. Gonzalo and Granger (1995) decompose a cointegrated price series into a permanent component and a temporary component using error correction coefficients. The permanent component is interpreted as the common efficient price, the temporary component reflects deviations from the efficient price caused by trading fractions. We estimate IS and CS, as developed by Hauptfleisch et al. (2016) using the error correction parameters and variance–covariance of the error terms from the Vector Error Correction Model (VECM):

$$\Delta p_{1,t} = \alpha_1(p_{1,t-1} - p_{2,t-1}) + \sum_{i=1}^{200} \gamma_i \Delta p_{1,t-i} + \sum_{j=1}^{200} \delta_j \Delta p_{2,t-j} + \varepsilon_{1,t} \quad (1)$$

$$\Delta p_{2,t} = \alpha_2(p_{1,t-1} - p_{2,t-1}) + \sum_{k=1}^{200} \varphi_k \Delta p_{1,t-k} + \sum_{m=1}^{200} \phi_m \Delta p_{2,t-m} + \varepsilon_{2,t} \quad (2)$$

where $\Delta p_{i,t}$ is the change in the log price ($p_{i,t}$) of the asset traded in market i at time t . The next stage is to obtain the component shares from the normalised orthogonal coefficients to the vector of error correction, or:

$$CS_1 = \gamma_1 = \frac{\alpha_2}{\alpha_2 - \alpha_1}; CS_2 = \gamma_2 = \frac{\alpha_1}{\alpha_1 - \alpha_2} \quad (3)$$

Given the covariance matrix of the reduced form VECM error terms¹ where:

$$M = \begin{pmatrix} m_{11} & 0 \\ m_{12} & m_{22} \end{pmatrix} = \begin{pmatrix} \sigma_1 & 0 \\ \rho\sigma_2 & \sigma_2(1 - \rho^2)^{\frac{1}{2}} \end{pmatrix} \quad (4)$$

we calculate the IS using:

$$IS_1 = \frac{(\gamma_1 m_{11} + \gamma_2 m_{12})^2}{(\gamma_1 m_{11} + \gamma_2 m_{12})^2 + (\gamma_2 m_{22})^2} \quad (5)$$

¹ $\Omega = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$ and its Cholesky factorisation, $\Omega = MM'$.

Table 4

Price discovery results.

Information Share (Hasbrouck)	Lower Bound	Upper Bound	Average
Futures	0.115535	0.183738	0.149637
Bitcoin	0.816261	0.884465	0.850363
Component Share (Gonzalo)	Average		
Futures	0.177028		
Bitcoin	0.822971		
Information Leadership (Yan)	Average		
Futures	0.025636		
Bitcoin	0.827931		
Information Leadership Share (Putnins)	Average		
Futures	0.030034		
Bitcoin	0.969965		

$$IS_2 = \frac{(\gamma_2 m_{22})^2}{(\gamma_1 m_{11} + \gamma_2 m_{12})^2 + (\gamma_2 m_{22})^2} \quad (6)$$

Recent studies show that IS and CS are sensitive to the relative level of noise in each market, they measure a combination of leadership in impounding new information and the relative level of noise in the price series from each market. The measures tend to overstate the price discovery contribution of the less noisy market. An appropriate combination of IS and CS cancels out dependence on noise (Yan and Zivot, 2010; Putnys, 2013). The combined measure is known as the Information Leadership Share (ILS) which is calculated as:

$$ILS_1 = \frac{\left| \frac{IS_1 CS_2}{IS_2 CS_1} \right|}{\left| \frac{IS_1 CS_2}{IS_2 CS_1} \right| + \left| \frac{IS_2 CS_1}{IS_1 CS_2} \right|} \text{ and } ILS_2 = \frac{\left| \frac{IS_2 CS_1}{IS_1 CS_2} \right|}{\left| \frac{IS_1 CS_2}{IS_2 CS_1} \right| + \left| \frac{IS_2 CS_1}{IS_1 CS_2} \right|} \quad (7)$$

We estimate all three price discovery metrics, noting that they measure different aspects of price discovery.

The results in Table 4 show that the spot market leads in price discovery according to all the metrics computed. This result is contrary to what has been found in a range of other asset classes, where futures markets lead. Looking at the Information Leadership Share, 97% of the information affecting Bitcoin prices is reflected in the spot market, while the remaining 3% is reflected in the futures market. The concentration of price discovery in the spot market may be a function of the novelty of the new futures contracts that have been trading for 3 months. It may also be the case that the type of investor attracted to Bitcoin because of its anonymity may not be inclined to begin trading on a registered and regulated futures market, in which personal details have to be recorded before trading is permitted; these investors would, however, in general be classified as uninformed. Because of various restrictions on Bitcoin there is an absence of a large cohort of institutional investors who have positions in physical Bitcoin. The results presented support the argument put forward by Bohl et al. (2011) that the dominance of unsophisticated individual investors in the futures market impedes its contribution to price discovery.

4. Conclusions

A currency has three economic attributes: it is a medium of exchange, a store of value, and a unit of account. Yermack (2015) asserted that Bitcoin was not a currency as it *performs poorly as a unit of account and as a store of value*. The high volatility of Bitcoin prices and the range of prices quoted on various Bitcoin exchanges were seen to damage Bitcoin's usefulness as a unit of account. If the introduction of Bitcoin futures and the ability to trade these would have resulted in a reduction in the variance of Bitcoin prices, or facilitated hedging strategies that could have mitigated pricing risk in the spot market, it is possible that the Bitcoin could have

acted as a unit of account, moving it closer to being a currency. The analysis conducted shows that volatility increased around the announcement of trading in Bitcoin futures. In the period covered by this study hedge portfolios constructed with futures cannot mitigate the risk inherent in the underlying spot market; both hedging strategies considered resulted in an increase in volatility. The price discovery analysis indicated that price discovery is focused on the spot market, which is in keeping with the argument that the traders in the futures market are uninformed noise traders. Together these results support the conclusion of Yermack (2015) that Bitcoin should be seen as a speculative asset rather than a currency.

References

- Bohl, M.T., Salm, C.A., Schuppli, M., 2011. Price discovery and investor structure in stock index futures. *J. Futures Markets* 31 (3), 282–306.
- Cabrera, J., Wang, T., Yang, J., 2009. Do futures lead price discovery in electronic foreign exchange markets? *J. Futures Markets* 29 (2), 137–156.
- Cheah, E.-T., Fry, J., 2015. Speculative bubbles in Bitcoin markets? An empirical investigation into the fundamental value of Bitcoin. *Econ. Lett.* 130, 32–36.
- Choudhry, T., 2003. Short-run deviations and optimal hedge ratio: evidence from stock futures. *J. Multinational Financ. Manag.* 13 (2), 171–192.
- Demir, E., Gozgor, G., Lau, C.K.M., Vigne, S.A., 2018. Does economic policy uncertainty predict the Bitcoin returns? An empirical investigation. *Financ. Res. Lett.* Available online 31 January 2018 In Press, Corrected Proof.
- Figlewski, S., 1984. Hedging performance and basis risk in stock index futures. *J. Financ.* 39 (3), 657–669.
- Gonzalo, J., Granger, C., 1995. Estimation of common long-memory components in cointegrated systems. *J. Business Econ. Stat.* 13 (1), 27–35.
- Gulen, H., Mayhew, S., 2000. Stock index futures trading and volatility in international equity markets. *J. Futures Markets: Futures Opt. Other Deriv. Prod.* 20 (7), 661–685.
- Hasbrouck, J., 1995. One security, many markets: Determining the contributions to price discovery. *J. Financ.* 50 (4), 1175–1199.
- Hauptfleisch, M., Putniņš, T.J., Lucey, B., 2016. Who sets the price of gold? London or New York. *J. Futures Markets* 36 (6), 564–586.
- Kroner, K.F., Sultan, J., 1993. Time-varying distributions and dynamic hedging with foreign currency futures. *J. Financ. Quant. Anal.* 28 (4), 535–551.
- Lee, C.L., Stevenson, S., Lee, M.-L., 2014. Futures trading, spot price volatility and market efficiency: evidence from European real estate securities futures. *J. Real Estate Finance Econ.* 48 (2), 299–322.
- Lucey, B.M., Vigne, S.A., Ballester, L., Barbopoulos, L., Brzezczynski, J., Carchano, O., Dimic, N., Fernandez, V., Gogolin, F., González-Urteaga, A., Goodell, J.W., Helbing, P., Ichev, R., Kearney, F., Laing, E., Larkin, C.J., Lindblad, A., Lončarski, I., Ly, K.C., Marinč, M., McGee, R.J., McGroarty, F., Neville, C., O'Hagan-Luff, M., Piljak, V., Sevic, A., Sheng, X., Stafylas, D., Urquhart, A., Versteeg, R., Vu, A.N., Wolfe, S., Yarovaya, L., Zaghini, A., 2018. Future directions in international financial integration research – A crowdsourced perspective. *Internat. Rev. Financ. Anal.* 55, 35–49.
- Park, T.H., Switzer, L.N., 1995. Bivariate GARCH estimation of the optimal hedge ratios for stock index futures: a note. *J. Futures Markets* 15 (1), 61–67.
- Putniņš, T.J., 2013. What do price discovery metrics really measure? *J. Empir. Finance* 23, 68–83.
- Rosenberg, J.V., Traub, L.G., 2009. Price discovery in the foreign currency futures and spot market. *J. Deriv.* 17 (2), 7–25.
- Ross, G.J., Tasoulis, D.K., Adams, N.M., 2011. Nonparametric monitoring of data streams for changes in location and scale. *Technometrics* 53 (4), 379–389.
- Ross, G.J., et al., 2015. Parametric and nonparametric sequential change detection in R: The cpm package. *J. Stat. Softw.* 66 (3), 1–20.
- Yan, B., Zivot, E., 2010. A structural analysis of price discovery measures. *J. Financ. Markets* 13 (1), 1–19.
- Yermack, D., 2015. Is Bitcoin a real currency? An economic appraisal. In: *Handbook of Digital Currency*. Elsevier, pp. 31–43.