

Homework 4

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Problem 13.9: Determine the mean, ACVS and stationarity of a random process

9. A random process $x[n]$ is characterized by

$$x[n] = A(\zeta) \cos [\Omega(\zeta)n + \Theta(\zeta)],$$

where random variables $A(\zeta)$, $\Omega(\zeta)$, and $\Theta(\zeta)$ are mutually independent. Random variables $A(\zeta) \sim U(0, 1)$ and $\Theta(\zeta) \sim U(-\pi, \pi)$ are of continuous type while $\Omega(\zeta)$ is of discrete type taking values 10 and 20 radians with equal probability.

[✓] a) **Determine the mean sequence $m_x[n]$**

We need to compute:

$$E[x[n]] = E[A \cos(\Omega n + \Theta)]$$

Since the three random variables are independent then we can write:

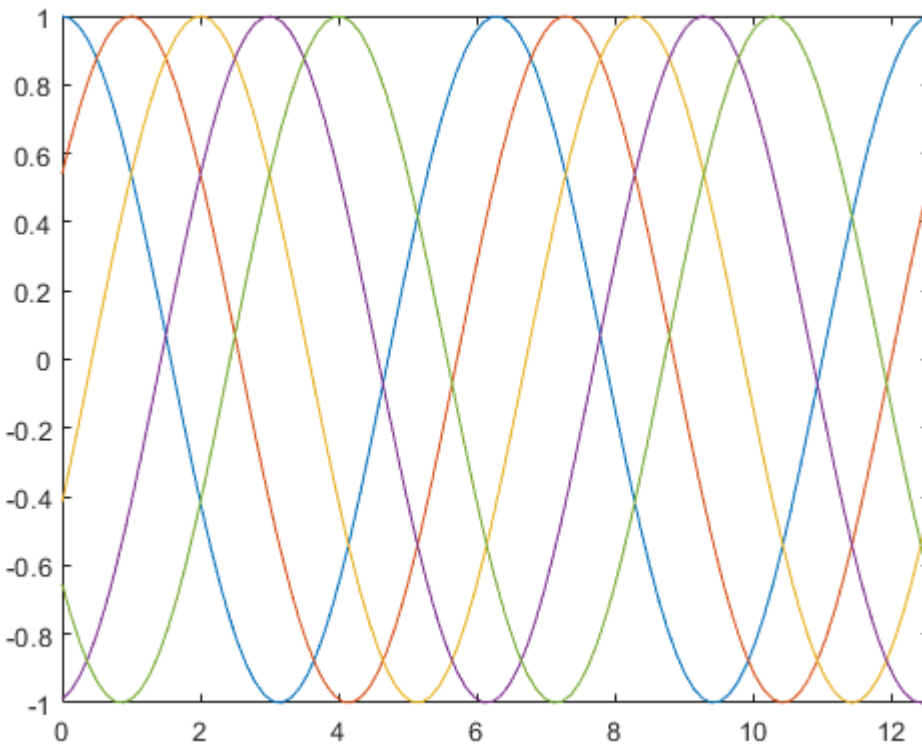
$$E[x[n]] = E[A]E[\cos(\Omega n + \Theta)]$$

We know that $E[A] = \frac{1}{2}$ since $A \sim U(0, 1)$.

Computing the expression $E[\sin(\Theta)]$ requires a bit of an explanation.

Suppose we want to compute $E[\cos(\Omega n + \Theta)]$ where $\Theta \sim U(-\pi, \pi)$. Let us pick one frequency ω (realise one value of Ω). Then let us pick a lot of realisations of Θ . Now if we plot the function $\cos(\omega n + \theta)$ for different values of θ then we will see something like this:

```
n = linspace(0, 4*pi);
plot(n, cos(n), n, cos(n-1), n, cos(n-2), n, cos(n-3), n, cos(n-4));
xlim([0, 4*pi]);
```



If we plot hundreds of cosine functions shifted slightly, we get a large blob of points from -1 to 1. For this reason, the quality $E[\cos(\Omega n + \Theta)]$ will be zero because the mean value is 0. Formally, we can write:

$$E[\cos(\Omega n + \Theta)] = \int_{-\pi}^{\pi} f_{\Theta}(\Theta) \cdot \cos(\Omega n + \Theta) d\Theta = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos(\Omega n + \Theta) d\Theta$$

So we are integrating cosine function over 2π which is zero.

In signal processing, we like to add random shifts ala $\Theta \sim U(0, 2\pi)$ to avoid that the expected value or the mean value becomes dependent on time.

Alternatively, we can use

Using the rule $\cos(\alpha + \beta) = \cos(\alpha)\cos(\beta) - \sin(\alpha)\sin(\beta)$, we rewrite the expression to:

$$E[x[n]] = E[A(\cos(\Omega n) \cos(\Theta) - \sin(\Omega n) \sin(\Theta))]$$

$$E[x[n]] = E[A \cos(\Omega n) \cos(\Theta) - A \sin(\Omega n) \sin(\Theta)]$$

At this point, we need to employ some expectation rules to decompose the expression.

Let X and Y be two random variables and a and b be two constants. Following expectation identities apply:

1. $E[a] = a$ e.g. $E(42) = 42$
2. $E[aX] = aE[X]$ e.g. if you multiply every value by 2, the expectation doubles
3. $E[a \pm X] = a \pm E[X]$ e.g. if you add 42 to every case, the expectation increases by 42
4. $E[X + Y] = E[X] + E[Y]$
5. If X and Y are independent, then $E[XY] = E[X]E[Y]$
6. $E[a \pm bX] = a \pm bE[X] = a \pm bE[X]$
7. $E[b(a \pm X)] = bE[a \pm X] = b(a \pm E[X])$

Use rule 4:

$$E[x[n]] = E[A \cos(\Omega n) \cos(\Theta)] - E[A \sin(\Omega n) \sin(\Theta)]$$

Use rule 5 multiple times:

$$E[x[n]] = E[A]E[\cos(\Omega n)]E[\cos(\Theta)] - E[A]E[\sin(\Omega n)]E[\sin(\Theta)]$$

But how can we continue from here?

We know that $\Theta \sim U(-\pi, \pi)$ so $E[\Theta] = \frac{\pi + (-\pi)}{2} = \frac{0}{2} = 0$

b) Determine the ACVS $c_X[m, n]$

c) Comment on the stationarity of the random process

[✓] Problem 13.13: MSE objective function

13. Consider the mse objective function (13.56)

$$J(a, b) = E[(Y - aX - b)^2].$$

a) Express the objective function in terms of its parameters

(a) Express $J(a, b)$ in terms of the parameters a , b , and the moments of X and Y .

Use MATLAB to expand the expression inside the expected value:

```
syms a b X Y
expand((Y-a*X - b)^2)
```

$$\text{ans} = X^2 a^2 - 2 X Y a + 2 X a b + Y^2 - 2 Y b + b^2$$

So we have:

$$J(a, b) = E[X^2 a^2 - 2 X Y a + 2 X a b + Y^2 - 2 Y b + b^2]$$

Let X and Y be two random variables and a and b be two constants. Following expectation identities apply:

1. $E[a] = a$ e.g. $E(42) = 42$
2. $E[a X] = a E[X]$ e.g. if you multiply every value by 2, the expectation doubles
3. $E[a \pm X] = a \pm E[X]$ e.g. if you add 42 to every case, the expectation increases by 42
4. $E[X + Y] = E[X] + E[Y]$
5. If X and Y are independent, then $E[XY] = E[X]E[Y]$

Use rule 4:

$$J(a, b) = E[X^2 a^2 - 2 X Y a + 2 X a b + Y^2 - 2 Y b + b^2]$$

$$J(a, b) = E[X^2 a^2] - E[2 X Y a] + E[2 X a b] + E[Y^2] - E[2 Y b] + E[b^2]$$

Use rule 1 and rule 2:

$$J(a, b) = a^2 E[X^2] - 2 a E[X Y] + 2 a b E[X] + E[Y^2] - 2 b E[Y] + b^2$$

b) Using partial derivatives to determine the values of parameters

(b) Using partial derivatives $\frac{\partial J}{\partial a}$ and $\frac{\partial J}{\partial b}$, determine the values of a and b by solving the equations $\partial J / \partial a = 0$ and $\partial J / \partial b = 0$ that minimize $J(a, b)$ to obtain optimum values given in (13.58) and (13.62).

First, take the partial derivatives:

$$\frac{\partial J(a,b)}{\partial a} = 2a E[X^2] - 2 E[X Y] + 2 b E[X]$$

$$\frac{\partial J(a,b)}{\partial b} = 2 a E[X] - 2 E[Y] + 2b$$

Next, solve the equations:

$$(\text{Eq. 1}) \quad 2a E[X^2] - 2 E[X Y] + 2 b E[X] = 0$$

$$(\text{Eq. 2}) \quad 2 a E[X] - 2 E[Y] + 2b = 0$$

Isolate b in (Eq. 2):

$$2b = -2 a E[X] + 2 E[Y]$$

$$b = -a E[X] + E[Y]$$

$$b = E[Y] - a E[X]$$

This corresponds to (13.58) in the book:

$$(13.58) \quad b_0 = m_y - a m_x$$

Now, plug the expression for b into Eq. 1 in order to find an expression for a :

$$2a E[X^2] - 2 E[X Y] + 2 b E[X] = 0$$

$$2a E[X^2] - 2 E[X Y] + 2 (E[Y] - a E[X]) E[X] = 0$$

$$2a E[X^2] - 2 E[X Y] + 2 E[X]E[Y] - 2 a E[X]E[X] = 0$$

$$2a E[X^2] - 2 a E[X]E[X] - 2 E[X Y] + 2 E[X]E[Y] = 0$$

$$2a (E[X^2] - E[X]E[X]) - 2 E[X Y] + 2 E[X]E[Y] = 0$$

$$2a (E[X^2] - E[X]E[X]) = 2 E[X Y] - 2 E[X]E[Y]$$

$$a (E[X^2] - E[X]E[X]) = E[X Y] - E[X]E[Y]$$

$$a = \frac{E[X Y] - E[X]E[Y]}{E[X^2] - E[X]E[X]} = \frac{E[X Y] - E[X]E[Y]}{E[X^2] - E[X]^2}$$

We have found an expression for a . The numerator looks like it is the covariance:

Covariance The *covariance* of two random variables X and Y is defined by

$$(13.25) \quad c_{xy} \triangleq \text{cov}(X, Y) \triangleq E[(X - m_x)(Y - m_y)] = E(XY) - E(X)E(Y)$$

The denominator looks like it is the variance:

$$(13.11) \quad \text{var}(X) = E[X^2] - E[X]^2 = E[X^2] - m_x^2$$

Therefore, the derived expression is the same as (13.62) in the book.

$$(13.62) \quad a_0 = \frac{c_{xy}}{\sigma_x^2} = \rho_{xy} \frac{\sigma_y}{\sigma_x}$$

Problem 13.22: Computing probabilities

22. Consider two jointly distributed random variables X and Y with pdf

$$f(x, y) = \begin{cases} 8xy, & 0 \leq x \leq 1, 0 \leq y \leq x \\ 0, & \text{otherwise} \end{cases}$$

a) Determine prior and conditional probabilities

(a) Determine $f(x)$, $f(y)$, $f(x|y)$, and $f(y|x)$.

b) Are X and Y independent?

(b) Are X and Y independent?