

ESE 402/542 Recitation 7: Linear Regression

Outline

1. Linear regression in matrix form
2. Polynomial fitting
3. Linear regression and maximum likelihood

Linear Regression in Matrix Form

1. Given (x_i, y_i) pairs, want to fit a line that predicts y from x optimally (least-squares sense)
2. In scalar case, $x_i \in \mathbb{R}$, and assume that the $y_i \in \mathbb{R}$ are generated via $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$, where $\epsilon_i \stackrel{\text{i.i.d.}}{\sim} \text{dist}(0, \sigma^2)$
3. In vector case, $\mathbf{x}_i \in \mathbb{R}^d$. Assume that the $y_i \in \mathbb{R}$ are generated via $y_i = \beta_0 + \beta_1 \mathbf{x}_{i,1} + \dots + \beta_d \mathbf{x}_{i,d} + \epsilon_i$
4. Stacking all n samples into a matrix,

$$\begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} \beta_0 + \beta_1 \mathbf{x}_{1,1} + \dots + \beta_d \mathbf{x}_{1,d} \\ \vdots \\ \beta_0 + \beta_1 \mathbf{x}_{n,1} + \dots + \beta_d \mathbf{x}_{n,d} \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix} \quad (1)$$

$$= \begin{bmatrix} 1 & \mathbf{x}_{1,1} & \dots & \mathbf{x}_{1,d} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & \mathbf{x}_{n,1} & \dots & \mathbf{x}_{n,d} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \vdots \\ \beta_d \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix} \quad (2)$$

Linear Regression in Matrix Form

In matrix form, we can assume our data has been generated

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}, \text{ where } \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}, \mathbf{X} = \begin{bmatrix} 1 & \mathbf{x}_{1,1} & \dots & \mathbf{x}_{1,d} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & \mathbf{x}_{n,1} & \dots & \mathbf{x}_{n,d} \end{bmatrix},$$

$$\boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \vdots \\ \beta_d \end{bmatrix}, \text{ and } \boldsymbol{\epsilon} = \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix}. \text{ Note that } \mathbb{E}[\boldsymbol{\epsilon}] = \vec{0} \text{ and } \mathbf{cov}(\boldsymbol{\epsilon}) = \sigma^2 I_n.$$

We want to minimize RSS and solve $\hat{\boldsymbol{\beta}} = \operatorname{argmin}_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2$.

Taking derivative w.r.t. $\boldsymbol{\beta}$ and setting to 0, we get

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$$

Linear Regression and Maximum Likelihood

Natural question: when is the above linear estimator provably "good"?

1. Above, we assumed $\epsilon \sim \text{dist}(0, \sigma^2 I_n)$
2. Additionally assume that $\epsilon \sim \mathcal{N}(0, \sigma^2 I_n)$. Then, $y_i | \mathbf{x}_i \sim \mathcal{N}(\beta^\top \mathbf{x}_i, \sigma^2)$. Can re-derive least-squares linear regression through MLE
3. Solve $\hat{\beta} = \operatorname{argmax}_{\beta} \Pr(\mathbf{x}_1, y_1, \dots, \mathbf{x}_n, y_n | \beta) = \operatorname{argmax}_{\beta} \prod_{i=1}^n \Pr(y_i | x_i, \beta) = \operatorname{argmax}_{\beta} \sum_{i=1}^n \log \Pr(y_i | x_i, \beta) = \operatorname{argmin}_{\beta} \sum_{i=1}^n (\beta^\top \mathbf{x}_i - y_i)^2 = \operatorname{argmin}_{\beta} \|\mathbf{X}\beta - \mathbf{y}\|^2$ which is the same formulation as we had above

Polynomial Fitting

Special case of linear regression: given the degree d of the polynomial, then we want to fit a model of the form

$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \cdots + \beta_d x^d.$$

Note that this is a special case of linear regression by setting $\mathbf{x}_k = x^k$ for $k = 0, \dots, d$. Given n samples of the form

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \cdots + \beta_d x_i^d + \epsilon_i,$$

we turn this into matrix form

$$\begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & x_1 & \cdots & x_1^d \\ 1 & x_2 & \cdots & x_2^d \\ \vdots & & & \\ 1 & x_n & \cdots & x_n^d \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_d \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}.$$

Polynomial Fitting

We can now solve for $\hat{\beta}_0, \dots, \hat{\beta}_d$ like earlier. The matrix

$$\mathbf{X} = \begin{bmatrix} 1 & x_1 & \cdots & x_1^d \\ 1 & x_2 & \cdots & x_2^d \\ \vdots & & & \\ 1 & x_n & \cdots & x_n^d \end{bmatrix},$$

is also known as the Vandermonde matrix.