

Lecture 3 :

$$x_i \stackrel{iid}{\sim} \text{dist}(\mu, \sigma^2)$$

$$\bar{X} = \frac{x_1 + \dots + x_n}{n}$$

CLT :

$$F_{\bar{X}}(x) = \Pr\{\bar{X} \leq x\}$$

$$\Phi_{\mu, \frac{\sigma^2}{n}}(x) : \text{CDF of } N(\mu, \frac{\sigma^2}{n})$$

Then :

$$\left| F_{\bar{X}}(x) - \Phi_{\mu, \frac{\sigma^2}{n}}(x) \right| \leq \frac{C}{n}$$

(for all $x \in \mathbb{R}$)

$\forall a, b$

$$\Pr\{\bar{X} \in [a, b]\} = \Pr\{N(\mu, \frac{\sigma^2}{n}) \in [a, b]\} + \frac{C}{n}$$

Gaussian Distribution:

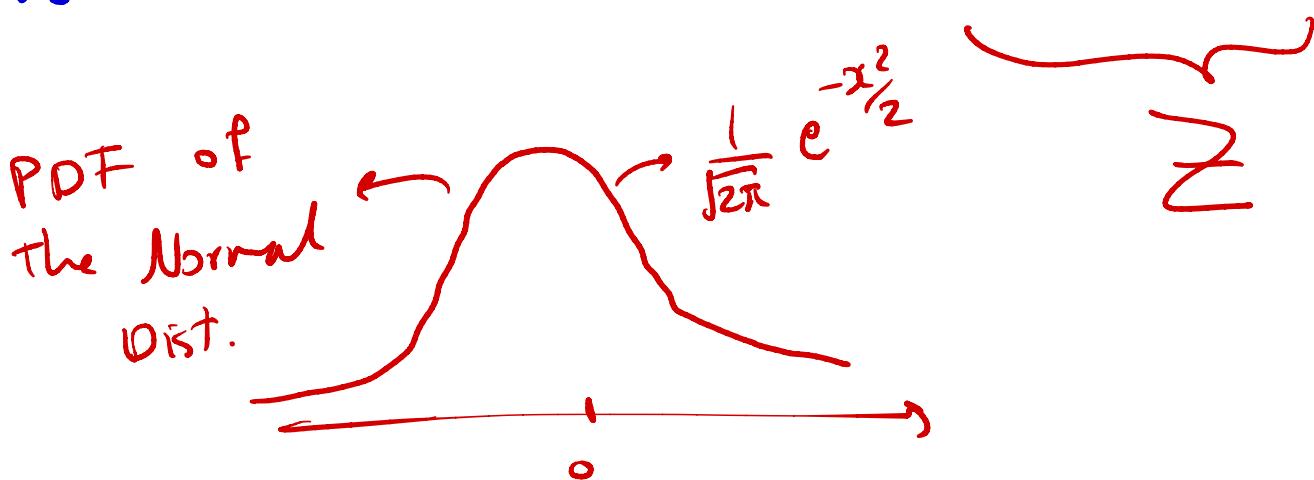
$$X \sim N(\mu, \sigma^2) \rightarrow f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

↳ mean: μ

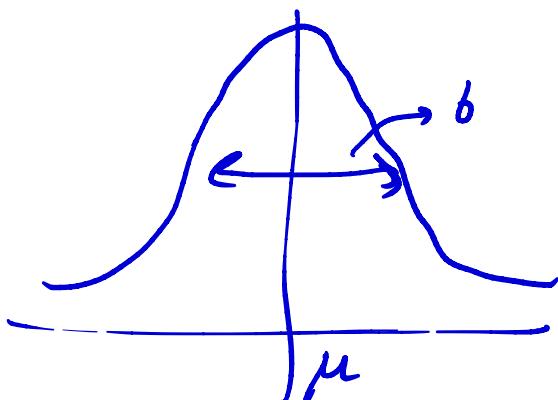
Variance: σ^2

Notation: Note that
we'll be using $N(\mu, \sigma^2)$
and σ^2 is the variance

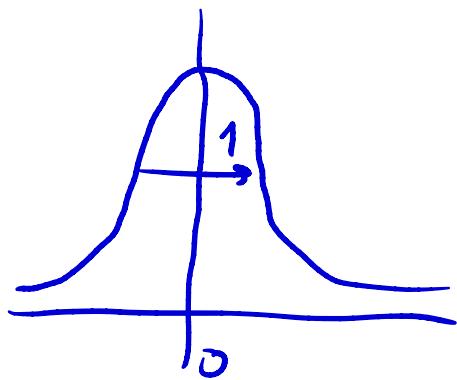
Normal Distribution $\triangleq N(0, 1)$



$$N(\mu, \sigma^2) = \mu + \sigma Z$$



first $\times \sigma$
then add μ

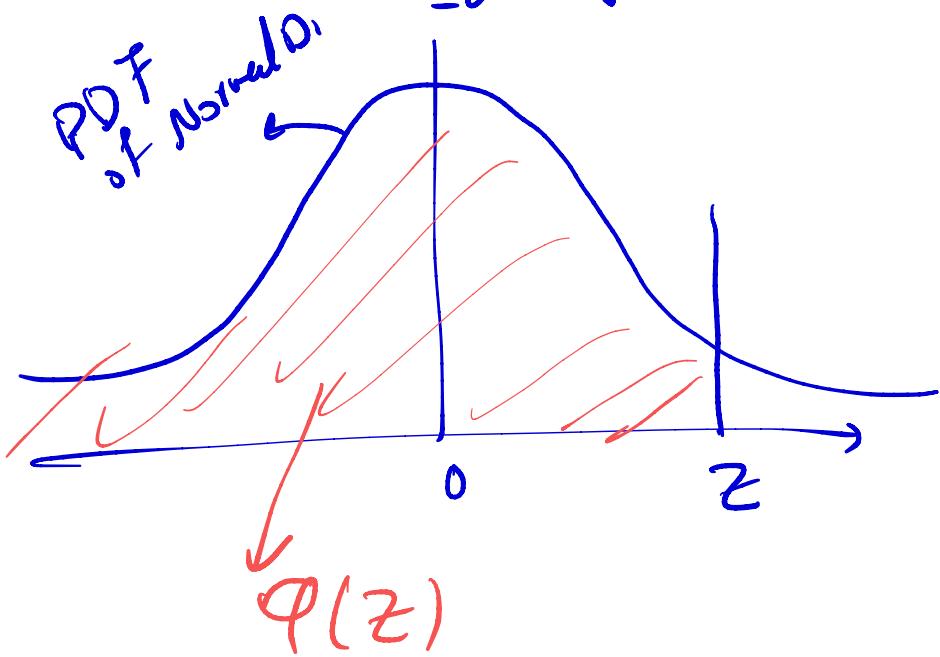


A few definitions:

The Gaussian Cumulative Distribution

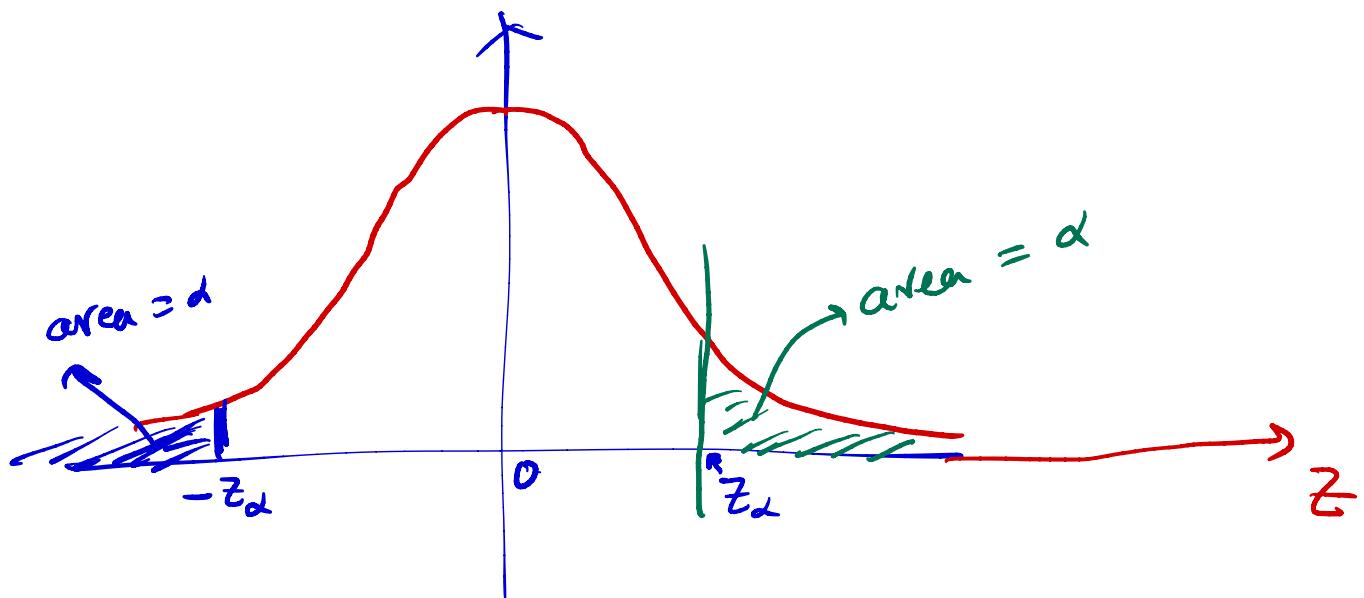
Function

$$\Phi(z) = \Pr\{Z \leq z\}$$
$$= \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$



Z_α : Given a value $\alpha \in [0,1]$ we define Z_α as the point on the z -axis such the area under the

normal PDF after z_α is equal to α .



Important Fact:

Sum of independent Gaussians is a Gaussian.

$$\begin{cases} X_1 \sim N(\mu_1, \sigma_1^2) \\ X_2 \sim N(\mu_2, \sigma_2^2) \end{cases} \quad X_1 \perp\!\!\!\perp X_2$$

$$\text{Then } \begin{cases} X_1 + X_2 = N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2) \\ \text{Var}(X_1 + X_2) = \text{Var}(X_1) + \text{Var}(X_2) \\ = \sigma_1^2 + \sigma_2^2 \end{cases}$$

Two r.v.'s are called independent if

$$\Pr\{X_1 \in A, X_2 \in B\} = \Pr\{X_1 \in A\} \cdot \Pr\{X_2 \in B\}$$

A, B

Q: If X_1, \dots, X_n are iid and $N(\mu, \sigma^2)$

then what is the distribution of

$$\frac{\underline{X_1 + \dots + X_n}}{n} ?$$

A: $X_i \sim N(\mu, \sigma^2)$ then:

$$\{ X_1 + \dots + X_n = N(\mu \cdot n, n\sigma^2) \}$$

$$\begin{aligned} \text{Var}(X_1 + \dots + X_n) &= \text{Var}(X_1) + \dots + \text{Var}(X_n) \\ &= \sigma^2 + \dots + \sigma^2 \\ &= n\sigma^2 \end{aligned}$$

$$\bar{X} = \frac{1}{n} (x_1 + \dots + x_n)$$

$$\rightarrow E[\bar{X}] = \frac{1}{n} E[x_1 + \dots + x_n] = \underline{\mu}$$

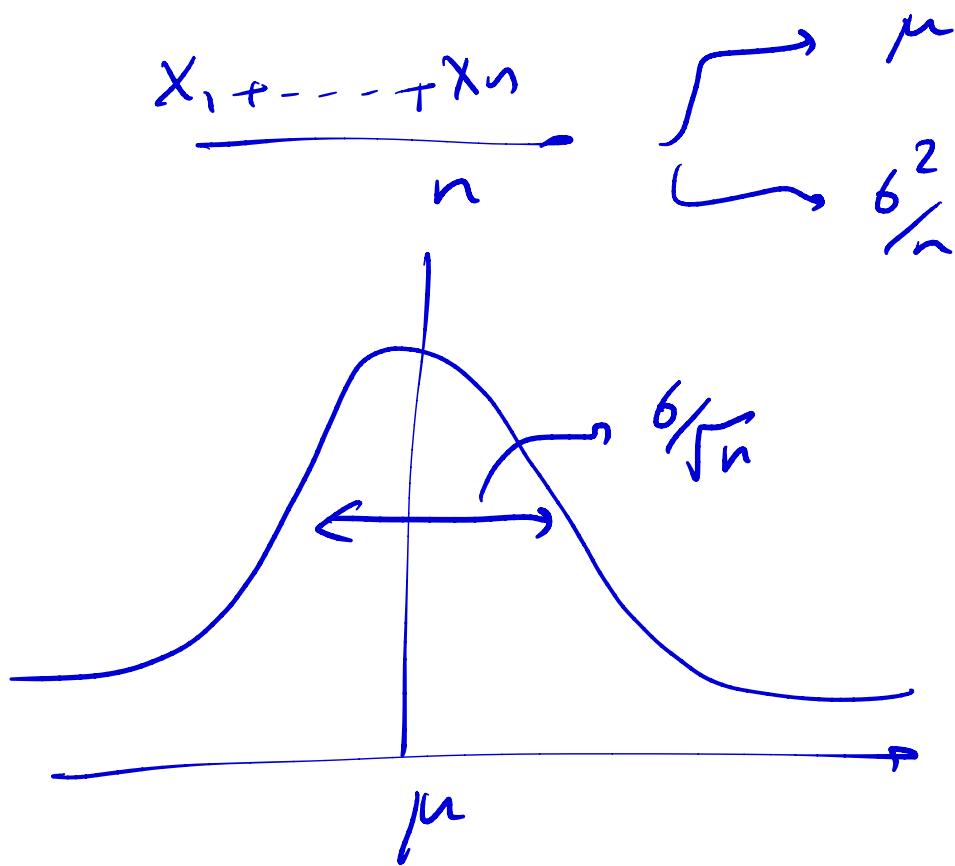
$$\left\{ \begin{array}{l} \text{Var}(\bar{X}) = \frac{1}{n^2} \text{Var}(x_1 + \dots + x_n) \end{array} \right.$$

$$= \frac{6^2}{n}$$

↖

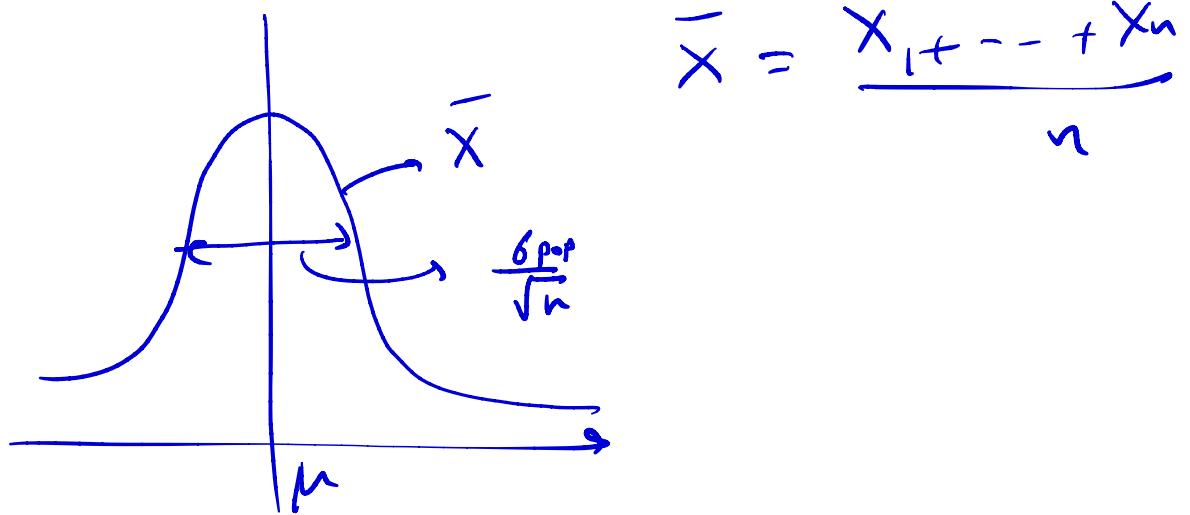
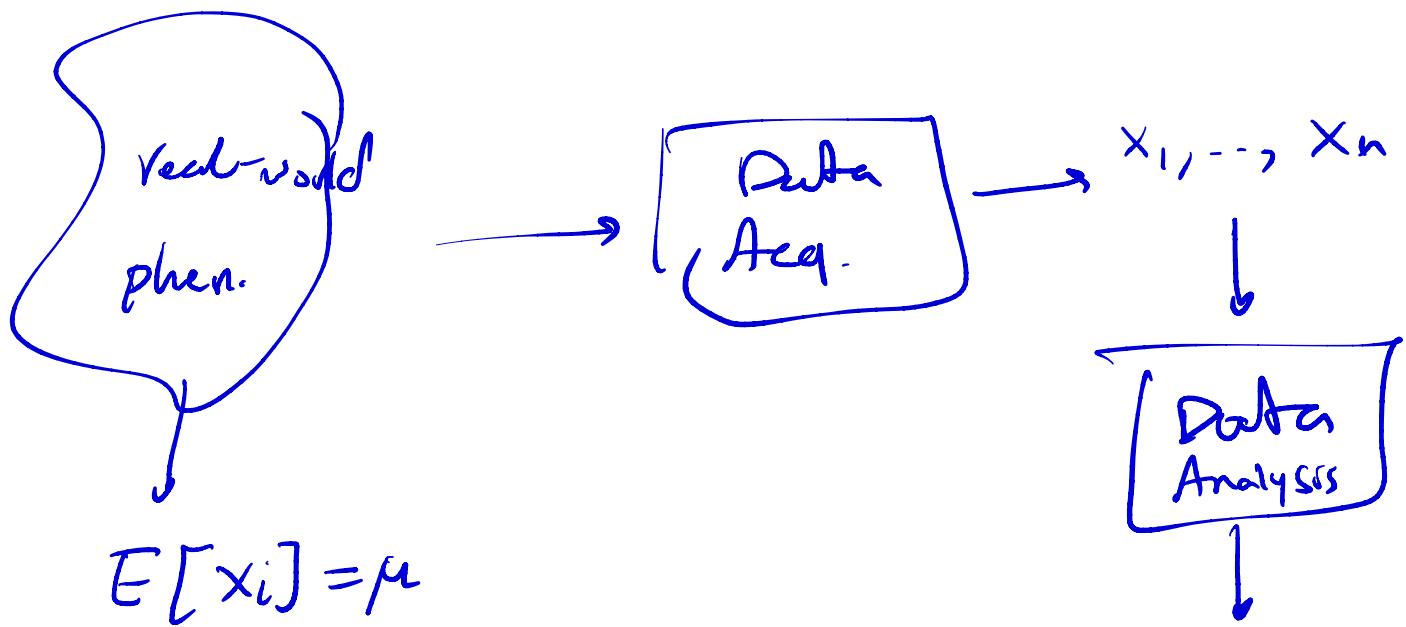
$$\Rightarrow \bar{X} = \mathcal{N}\left(\mu, \frac{6^2}{n}\right)$$

Confidence intervals:



So far we've seen that the variance of \bar{X} decays like σ^2/n . The question that we'd like to answer now is "How can we obtain guarantees on how good the

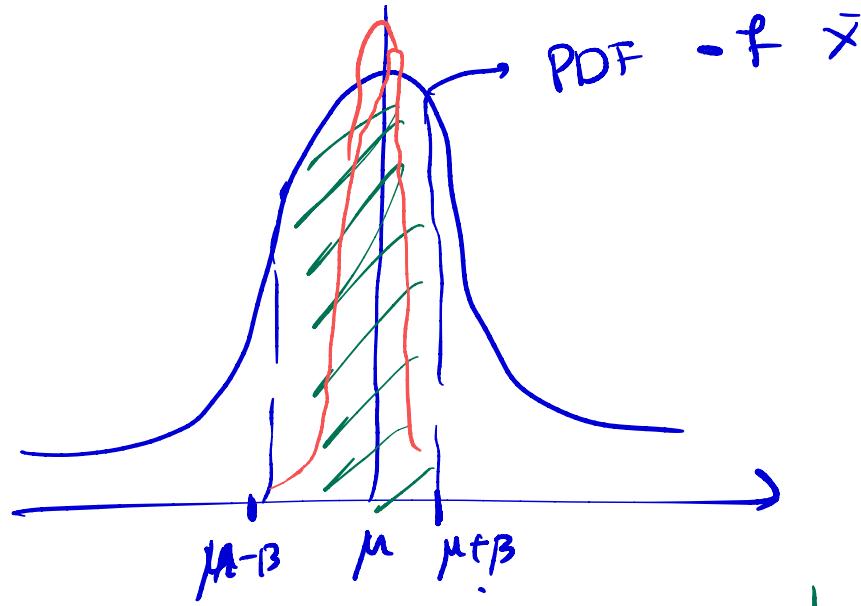
estimator \bar{X} is ?, "



(1) idea: $\Pr \{ \bar{X} = \mu \} = ?$

(2) right question $\Pr \{ \mu \in [\bar{X} - \beta, \bar{X} + \beta] \} = ?$

think of β
to be a
small number



$$\Pr \{ \mu \in [\bar{x} - \beta, \bar{x} + \beta] \} = ?$$

Equivalently, we'd like to know
for what value of β we
have:

$$\Pr \{ \mu \in [\bar{x} - \beta, \bar{x} + \beta] \} = \underbrace{1 - \alpha}_{\text{given}}$$

e.g. 0.95
 \sim

As the first step, assume that

$$X_i \stackrel{iid}{\sim} N(\mu, \sigma^2).$$

$$\Pr\{\bar{X} \in [\mu - \beta, \mu + \beta]\} \leftarrow \\ = \Pr\{\mu \in [\bar{X} - \beta, \bar{X} + \beta]\} \leftarrow$$

$$\bar{X} \in [\tilde{\mu} - \beta, \tilde{\mu} + \beta]$$

$$\left\{ \begin{array}{l} \bar{X} \leq \tilde{\mu} + \beta \Leftrightarrow \bar{X} - \beta \leq \tilde{\mu} \\ \bar{X} \geq \tilde{\mu} - \beta \Leftrightarrow \tilde{\mu} \leq \bar{X} + \beta \end{array} \right.$$

$$\left\{ \begin{array}{l} \bar{X} \leq \tilde{\mu} + \beta \Leftrightarrow \bar{X} - \beta \leq \tilde{\mu} \\ \bar{X} \geq \tilde{\mu} - \beta \Leftrightarrow \tilde{\mu} \leq \bar{X} + \beta \end{array} \right.$$

$$\bar{X} \in [\mu - \beta, \mu + \beta] \Leftrightarrow \mu \in [\bar{X} - \beta, \bar{X} + \beta]$$

So we'd like to compute

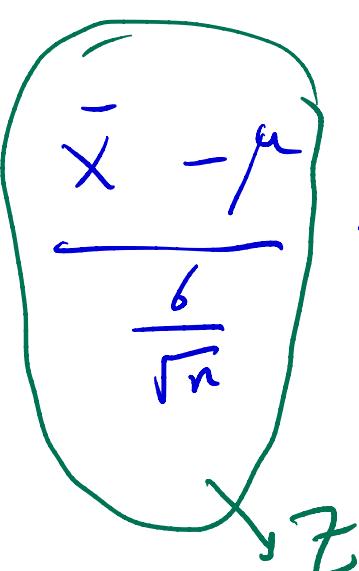
$$\Pr \{ \bar{X} \in [\mu - \beta, \mu + \beta] \} \leftarrow$$

$$x_i \sim N(\mu, \sigma^2) \Rightarrow \bar{X} = N(\mu, \frac{\sigma^2}{n})$$

1

$$\Pr \{ \mu - \beta \leq \bar{X} \leq \mu + \beta \}$$

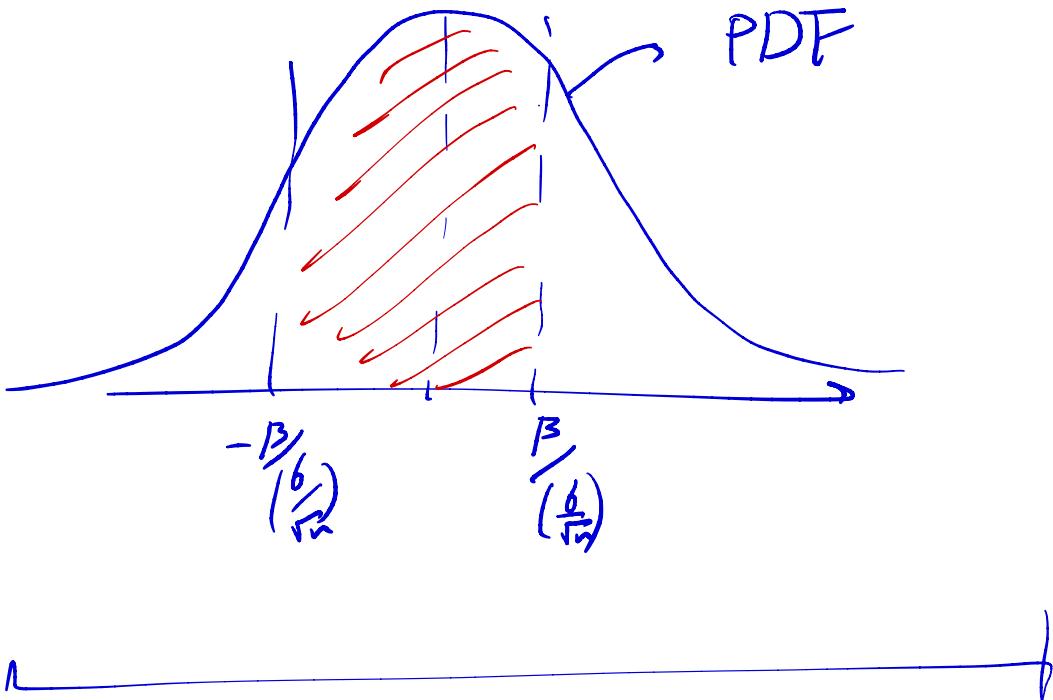
$$\approx \Pr \{ -\beta \leq \bar{X} - \mu \leq \beta \}$$

$$\approx \Pr \left\{ \frac{-\beta}{\frac{\sigma}{\sqrt{n}}} \leq \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \leq \frac{\beta}{\frac{\sigma}{\sqrt{n}}} \right\}$$


$$\bar{X} = N(\mu, \frac{\sigma^2}{n}) = \mu + \frac{\sigma}{\sqrt{n}} Z$$

$$\Rightarrow \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} = Z \sim N(0, 1)$$

$$= \Pr \left\{ -\frac{\beta}{\frac{\sigma}{\sqrt{n}}} \leq Z \leq \frac{\beta}{\frac{\sigma}{\sqrt{n}}} \right\}$$

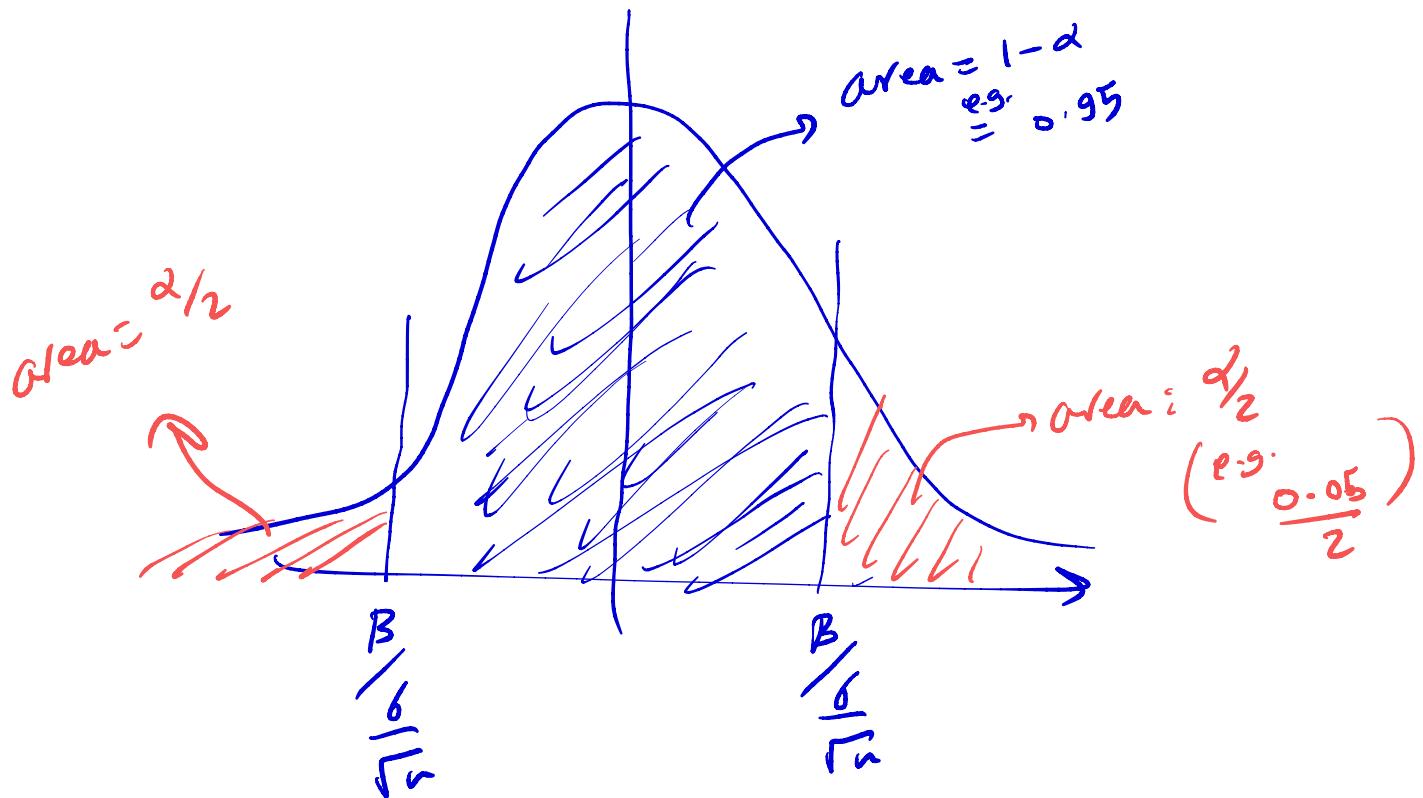


Recall that we wanted to

Solve:

$$\rightarrow \Pr \left\{ \mu \in [\bar{x} - \beta, \bar{x} + \beta] \right\} = \overbrace{1-\alpha}^{0.95}$$

$$= \Pr \left\{ -\frac{\beta}{\frac{\sigma}{\sqrt{n}}} \leq Z \leq \frac{\beta}{\frac{\sigma}{\sqrt{n}}} \right\} = 1-\alpha$$



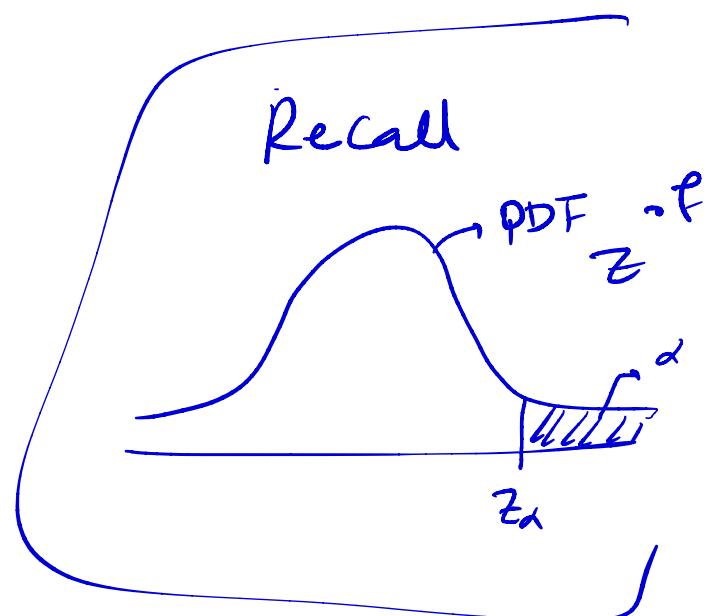
$$\frac{\beta}{\frac{\sigma}{\sqrt{n}}} = z_{\alpha/2} \quad \begin{pmatrix} \text{e.g.} \\ z_{0.025} \\ = 1.96 \end{pmatrix}$$

$$\beta = \frac{6}{\sqrt{n}} \cdot z_{\alpha/2}$$

Equivalently,

$$\Pr \left\{ \mu \in \left[\bar{x} - \frac{6}{\sqrt{n}} z_{\alpha/2}, \bar{x} + \frac{6}{\sqrt{n}} z_{\alpha/2} \right] \right\}$$

$$= 1 - \alpha$$



- So far, all we've obtained was is for the case where $X_i \sim N(\mu, \sigma^2)$.
- The more general setting is when $X_i \stackrel{iid}{\sim} \text{dist}(\mu, \sigma^2)$

In this case, we ask:

for what value of β the following holds:

$$\Pr\{\mu \in [\bar{x} - \beta, \bar{x} + \beta]\} = 1 - \alpha$$

In this case, from the CLT, we know that:

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) + \begin{array}{l} \text{negligible} \\ \text{error} \\ \text{for } n \\ \text{large} \end{array}$$

And in the exact same way as
the Gaussian case derived above,
we obtain that

$$\Pr\left\{\mu \in \left[\bar{X} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}\right]\right\} = 1 - \alpha.$$

(There is always
a CLT error
of $\frac{C}{n}$ which
we neglect.)