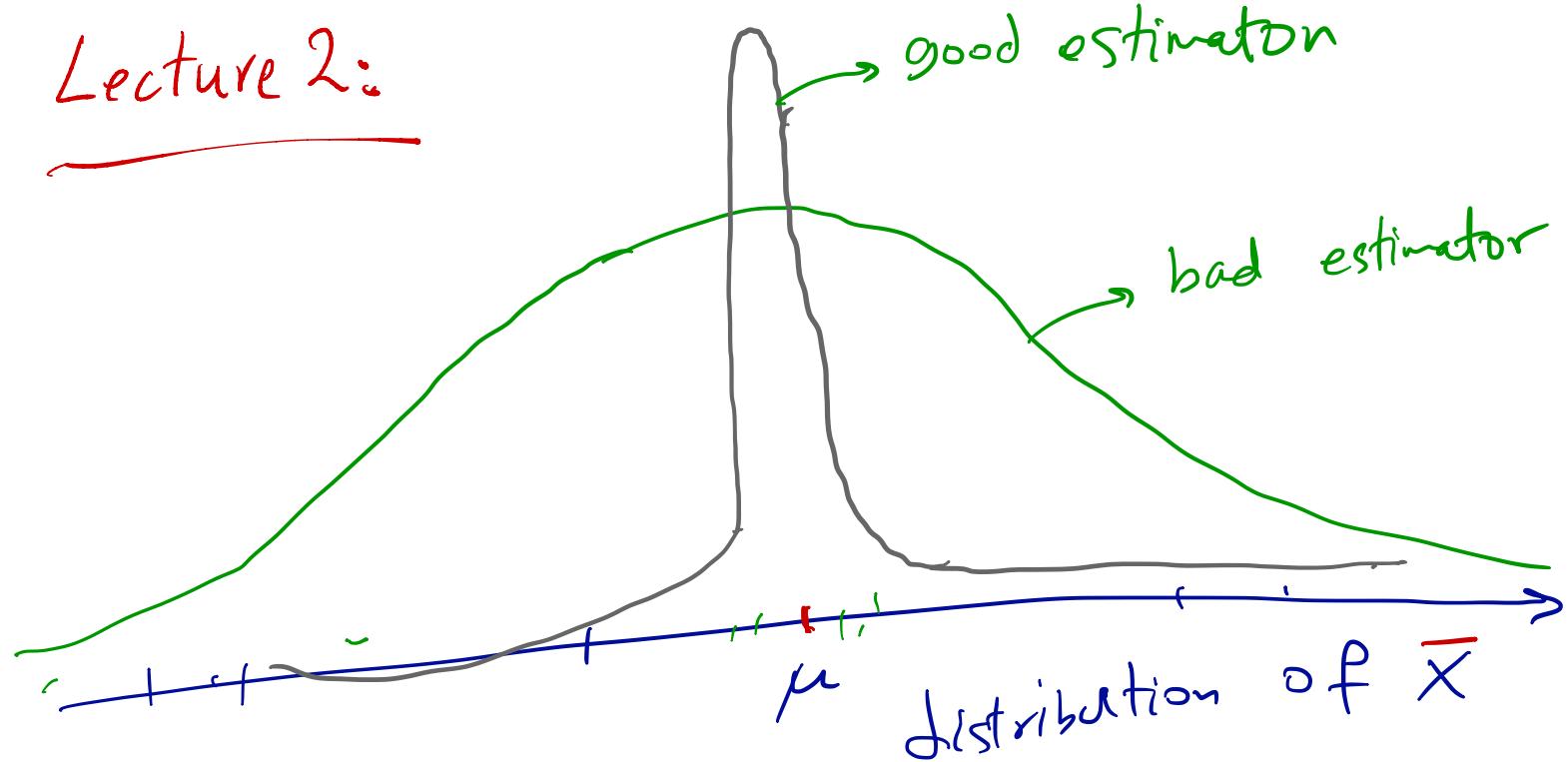


Lecture 2:



In order to see how concentrated \bar{X} is around μ , we need to compute ...

$$E[(\bar{X} - \mu)^2].$$

$$E[(\bar{X} - \mu)^2] \quad \left(\bar{X} = \frac{x_1 + \dots + x_n}{n} \right)$$

$$E \left[\left(\frac{x_1 + x_2 + \dots + x_n}{n} - \mu \right)^2 \right]$$

$$= E \left[\left(\frac{x_1 + x_2 + \dots + x_n - n\mu}{n} \right)^2 \right]$$

$$= E \left[\left(\frac{(x_1 - \mu) + (x_2 - \mu) + \dots + (x_n - \mu)}{n} \right)^2 \right]$$

$$= \frac{1}{n^2} E \left[\left((x_1 - \mu) + (x_2 - \mu) + \dots + (x_n - \mu) \right)^2 \right]$$

= - -

$$(- - -)^2 = \left(\frac{1}{n} \sum_{i=1}^n (x_i - \mu) \right)^2$$

$$= \frac{1}{n^2} E \left[\left(\sum_{i=1}^n (x_i - \mu) \right)^2 \right]$$

$$= \frac{1}{n^2} E \left[\sum_{i=1}^n (x_i - \mu)^2 + 2 \sum_{\substack{i, j=1 \\ i \neq j}}^n (x_i - \mu)(x_j - \mu) \right]$$

$$= \frac{1}{n^2} \sum_{i=1}^n E[(x_i - \mu)^2] + \underbrace{\frac{2}{n^2} \sum_{\substack{i, j=1 \\ i \neq j}}^n E[(x_i - \mu)(x_j - \mu)]}_{\textcircled{2}}$$

$$(d_1 + d_2 + \dots + d_n)^2 = d_1^2 + d_2^2 + \dots + d_n^2 + 2d_1d_2 + 2d_1d_3 + \dots + 2d_{n-1}d_n$$

②

$$E[(x_i - \mu)(x_j - \mu)] \quad (i \neq j)$$

when $i \neq j$, x_i and x_j are independent. Hence

$$E[(x_i - \mu)(x_j - \mu)]$$

independence $= E[(x_i - \mu)] E[(x_j - \mu)]$

$$= (\underbrace{E[x_i] - \mu}_{=0}) (\underbrace{E[x_j] - \mu}_{=0})$$

$$\Rightarrow ② = 0$$

if two random variables U, V are independent, then $E[U \cdot V] = E[U]E[V]$

$$E[g(U) \cdot h(V)] = E[g(U)] \cdot E[h(V)]$$

$$\begin{aligned}
 ① &= \frac{1}{n^2} \sum_{i=1}^n E[(x_i - \mu)^2] \\
 &= \frac{1}{n^2} (E[(x_1 - \mu)^2] + E[(x_2 - \mu)^2] + \dots + E[(x_n - \mu)^2]) \\
 &\equiv \frac{1}{n^2} \cdot n \cdot E[(x_1 - \mu)^2] \\
 &\quad \uparrow x_i \text{'s are identically distributed (uniform)} \\
 &\equiv \frac{1}{n} E[(x_1 - \mu)^2]
 \end{aligned}$$

$$\equiv \frac{1}{n} \left(\frac{1}{N} (y_1 - \mu)^2 + \frac{1}{N} (y_2 - \mu)^2 + \dots + \frac{1}{N} (y_N - \mu)^2 \right)$$

$$\equiv \frac{1}{n} \cdot \sigma_{\text{population}}^2$$

$$\text{Where } \sigma_{\text{population}}^2 = \frac{1}{N} \left((y_1 - \mu)^2 + \dots + (y_N - \mu)^2 \right)$$

Variance of \bar{X} :

$$E[(\bar{X} - \mu)^2] = ① + ②$$

$\stackrel{=}{\approx}$

σ^2 fixed number

$= \frac{1}{n} \sigma_{\text{population}}^2$

So the variance of \bar{X}

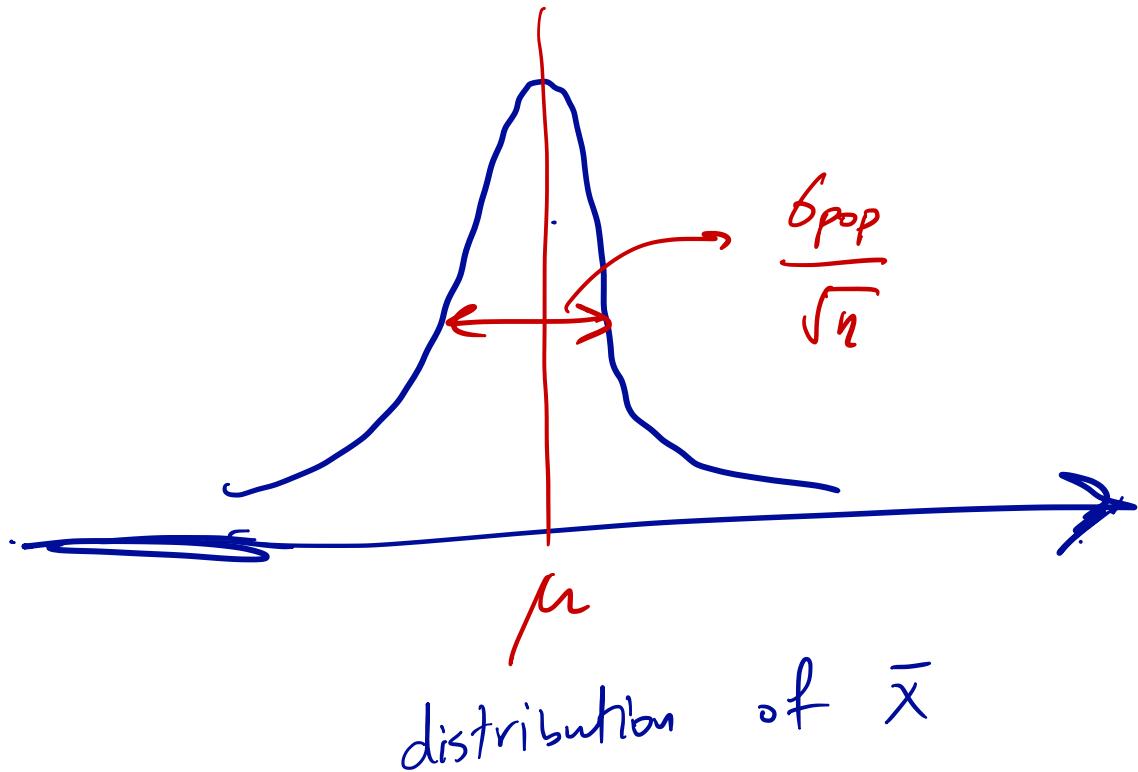
decays wrt the number of data points, n , like $\frac{1}{n}$.

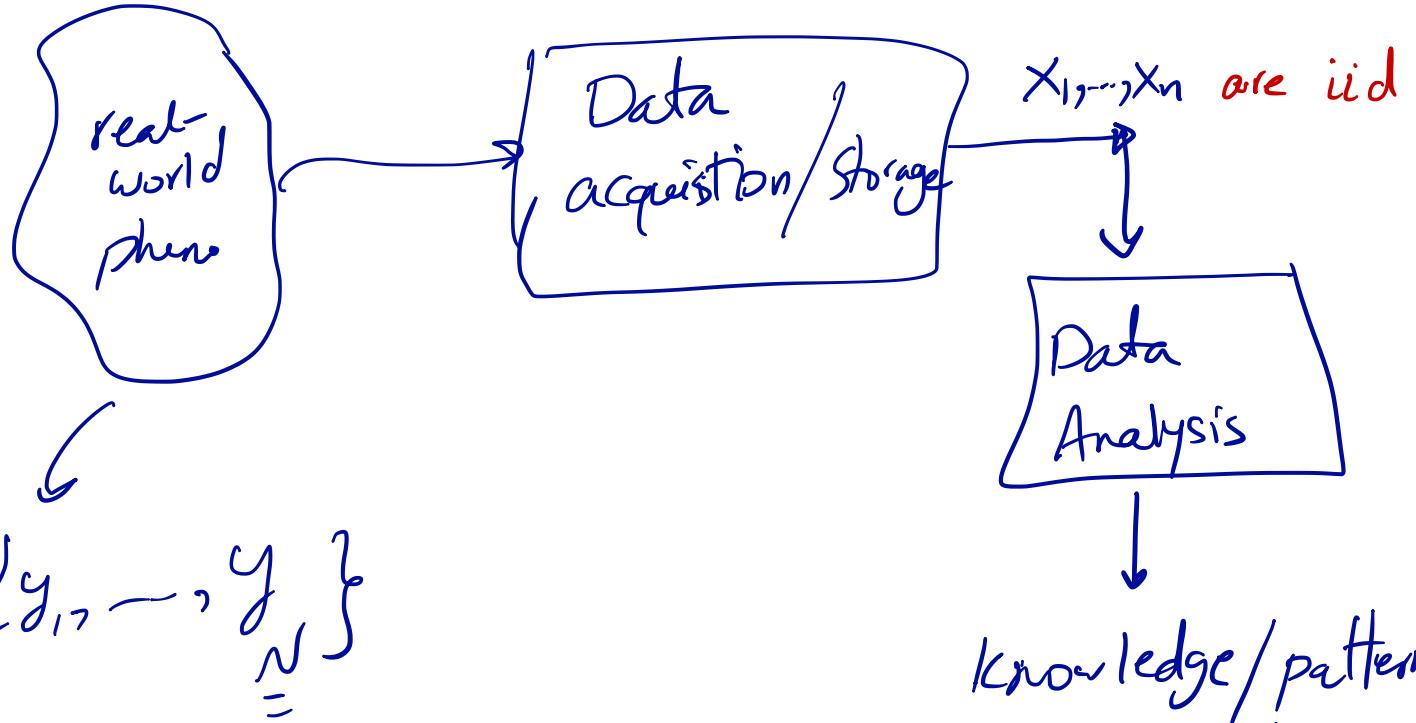
\approx

Standard deviation of \bar{X} :

$$\sqrt{E(\bar{X} - \mu)^2} = \frac{\sigma_{\text{pop}}}{\sqrt{n}}$$

Conclusion:





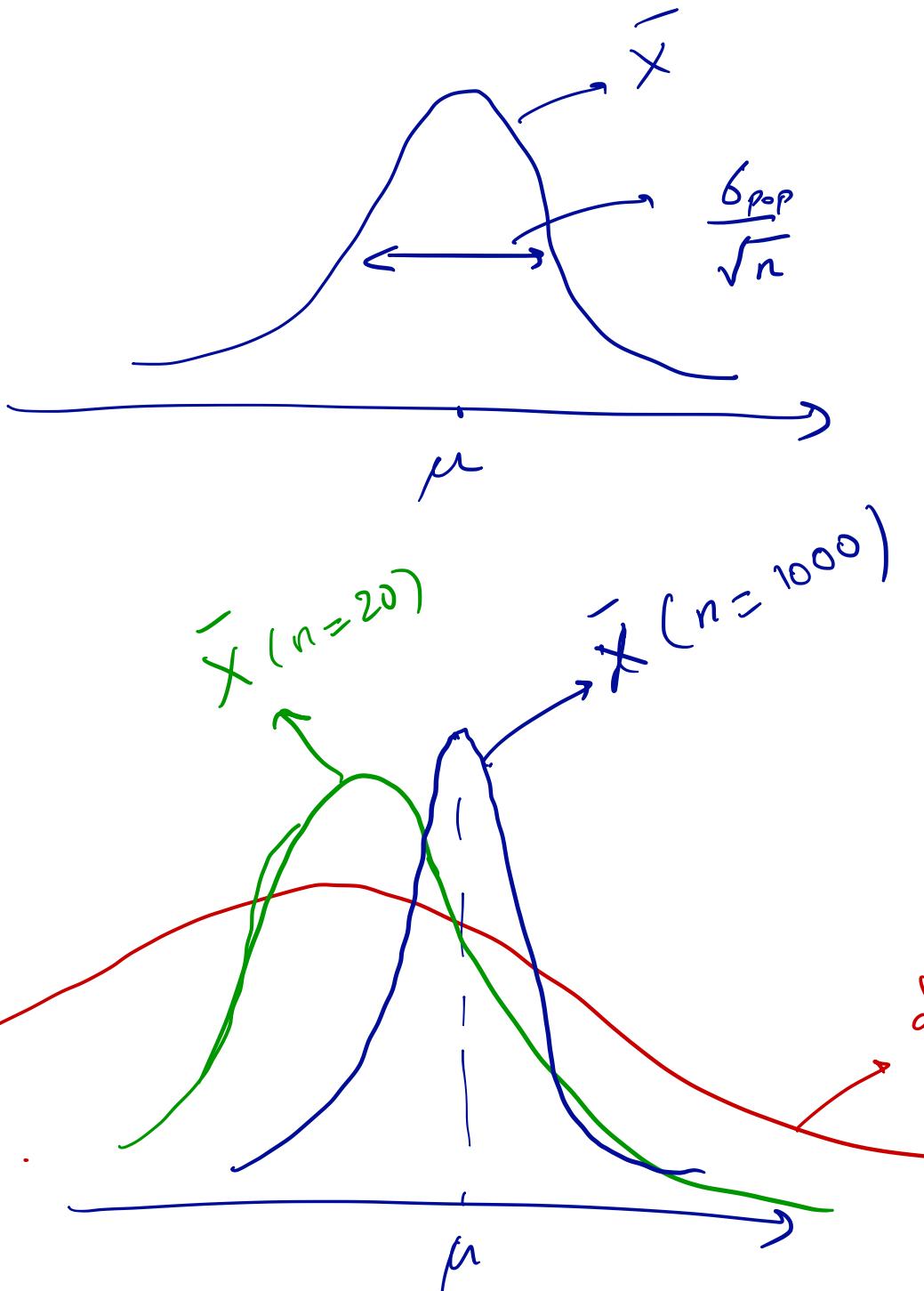
N very large

$$\mu = \frac{\sum_{i=1}^N y_i}{N} \quad , \quad \sigma_{\text{pop}} = \sqrt{\frac{\sum_{i=1}^N (y_i - \mu)^2}{N}}$$

$$x_i = \left\{ \begin{array}{l} y_1 \\ \vdots \\ y_N \end{array} \right. \quad \overline{x} = \frac{x_1 + \dots + x_n}{n}$$

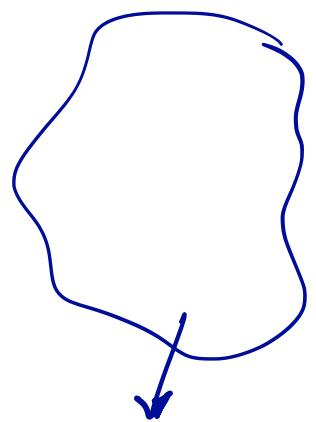
$$E[\bar{x}] = \mu$$

$$\text{Var}(\bar{x}) = \frac{\sigma_{\text{pop}}^2}{n}$$



Another example:

We'd like to find the percentage of people who would vote for a certain candidate in a state.



$$\{y_1, \dots, y_N\}$$

$$y_i = \begin{cases} 1 & \text{if } i \text{ would vote for the candidate} \\ 0 & \text{o.w.} \end{cases}$$

$N = \#$ of people
in the state

$$\mu = \frac{y_1 + \dots + y_N}{N} \rightarrow \% \text{ of people who'd vote for the candidate}$$

Let's summarize what we've learned so far. (and this is how we think of data acquisition/gathering procedure throughout the course).

In order to learn about a physical real-world phenomenon, we will gather sample data.

Random Sample: Statistically,

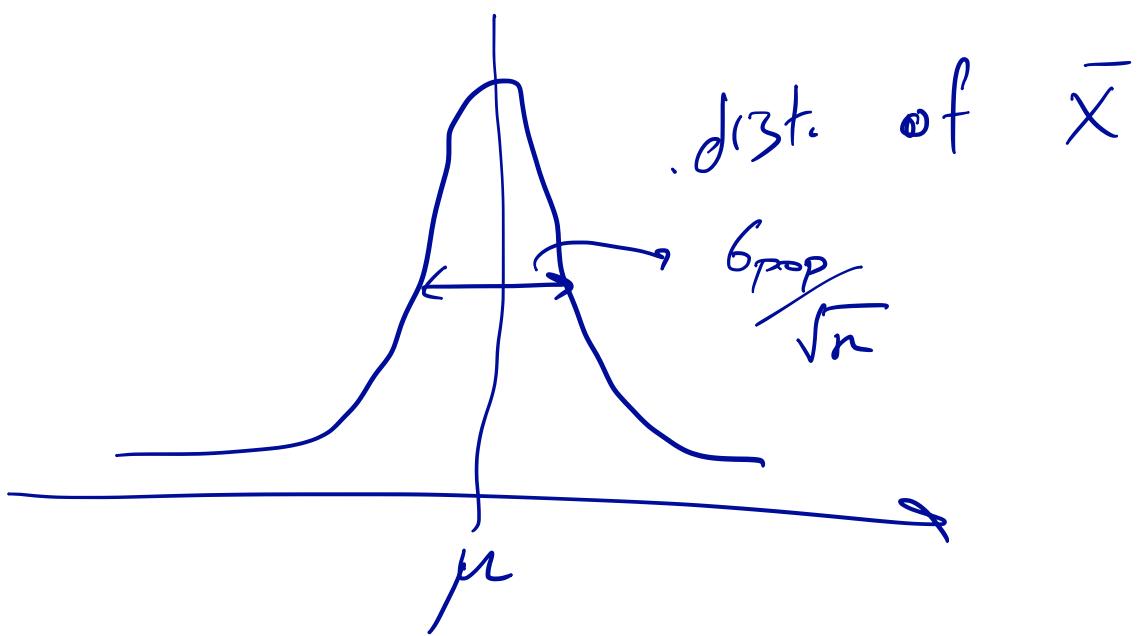
we model the sample data as

x_1, x_2, \dots, x_n where x_i 's are assumed to be independent and identically distributed (i.i.d.) and the samples are generated from the same phenomenon.

↳ x_i 's are random observations

- A statistic is any quantity whose value can be computed from sample data. Prior to obtaining data, there is uncertainty as to what value any particular statistic will result. Therefore, a statistic is typically a random variable.

e.g. Sample mean \bar{X} is a statistic.



In order to understand more about the distribution of \bar{X} , we need to study the powerful framework of Central limit theorem.

The Central Limit Theorem:

Take $X_1, \dots, X_n \stackrel{iid}{\sim} \text{dist}(\mu, \sigma^2)$

Then (informally):

$\bar{X} \sim \text{gaussian} + \text{small error}$

we know that :

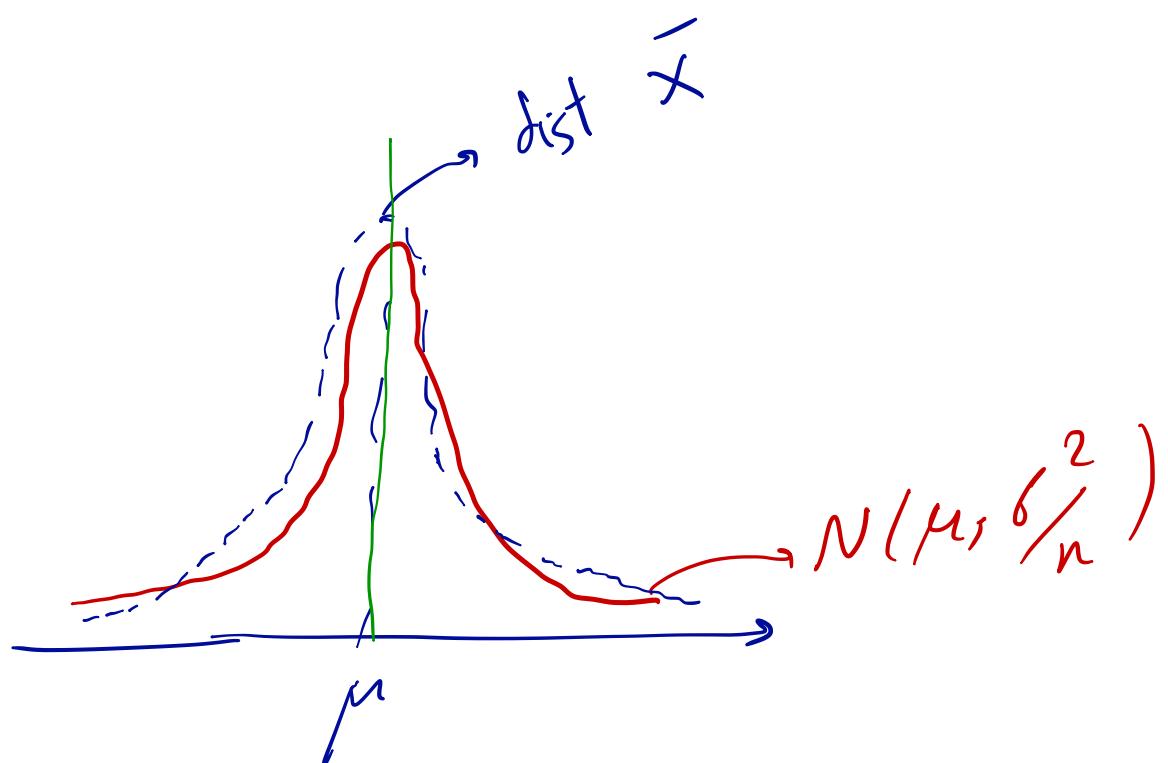
$$\begin{cases} E[\bar{X}] = \mu \\ \text{Var}(\bar{X}) = \frac{\sigma^2}{n} \end{cases}$$

$$\bar{X} = \text{gaussian} + \text{small error}$$

\downarrow

$\frac{c}{n}$

$$= N\left(\mu, \frac{\sigma^2}{n}\right) + \text{small error}$$



Formal Statement (CLT):

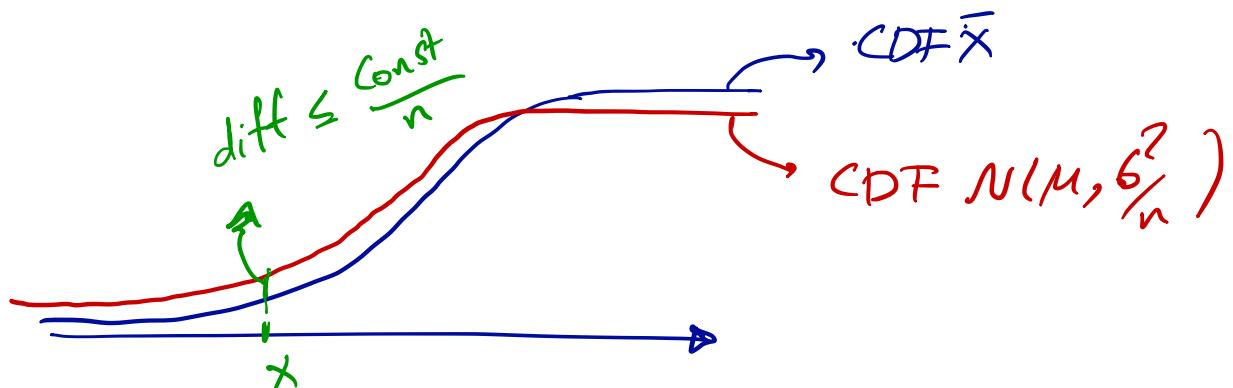
Let $F_{\bar{X}}(x)$ be the CDF of \bar{X} :

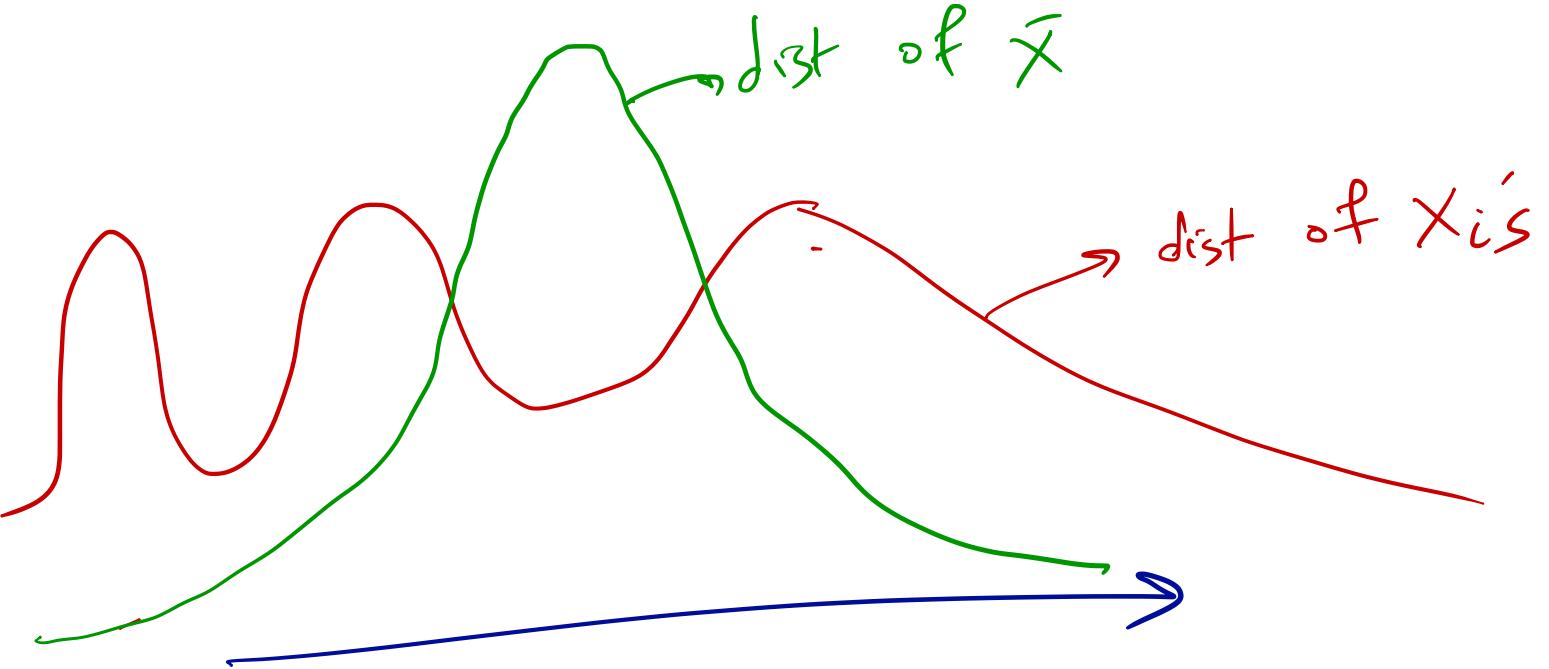
$$F_{\bar{X}}(x) = \Pr \{ \bar{X} \leq x \}$$

and Let $\phi_{\mu, \sigma}(x)$ be the CDF of $N(\mu, \frac{\sigma^2}{n})$. Then :

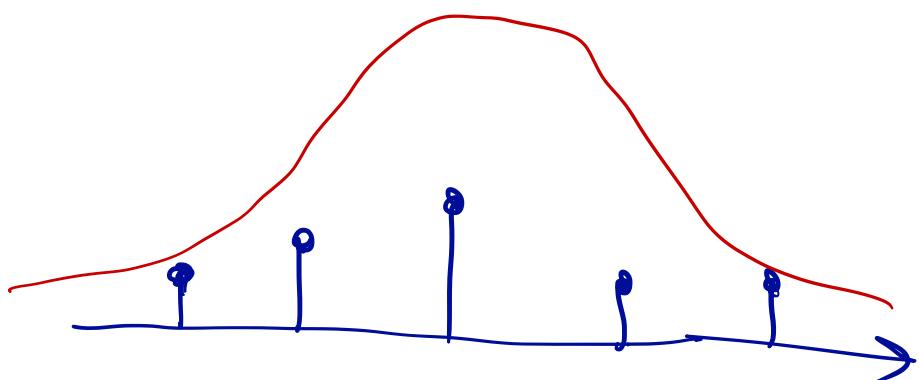
depends on
the dis. X_i 's

$$| F_{\bar{X}}(x) - \phi_{\mu, \sigma}(x) | \leq \frac{\text{Constant}}{n}$$





The theorem does not hold for
Pdfs. (e.g. the dist of X_i could be discrete)



Gaussian Distribution:

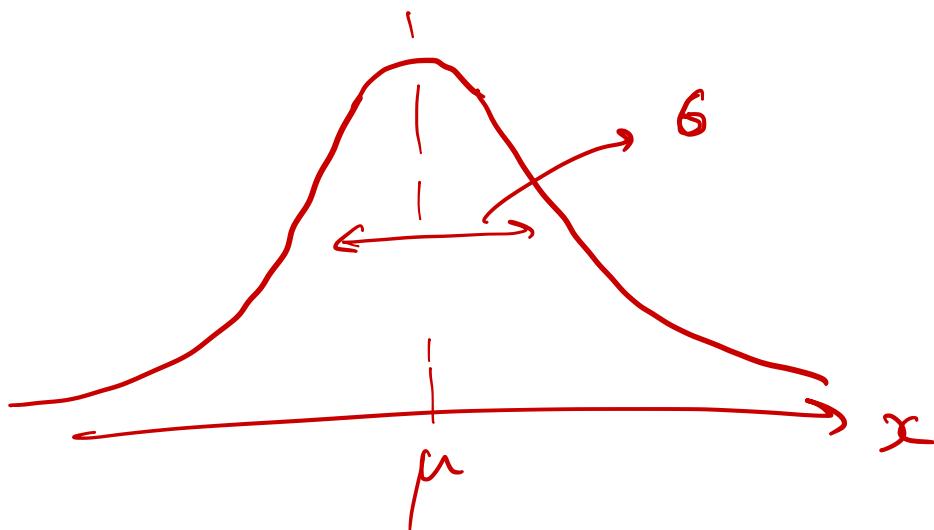
$$N(\mu, \sigma^2)$$

↓ mean ↓ variance

Pdf:

$$\frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

:=



The standard gaussian or the normal distribution:

$$Z = N(0, 1)$$