

$$4. f: \mathbb{R}^d \rightarrow \mathbb{R}^p, f(x) = \omega^T x$$

$\omega \in \mathbb{R}^{d \times p}$ $x \in \mathbb{R}^d$ $y \in \mathbb{R}^p$

$$\{(x_i, y_i)\}_{i=1}^n$$

$$\text{Let } X \in \mathbb{R}^{n \times d}, Y \in \mathbb{R}^{n \times p}$$

$$J(\omega) = \sum_{i=1}^n \| \omega^T x_i - y_i \|_2^2$$

Consider i^{th} instance

$$x_i = \langle x_{i1}, \dots, x_{id} \rangle^T \xrightarrow{\text{extend w/bias}}$$

extend w/bias

$$\begin{aligned} & \text{extend } \left(\begin{array}{c} \text{w/bias} \\ \langle 1, x_{i1}, \dots, x_{id} \rangle^T \end{array} \right)^T \\ & \quad \left(\begin{array}{c} 1 \\ w_{01} \dots w_{0p} \\ \vdots \\ w_{d1} \dots w_{dp} \end{array} \right) \end{aligned}$$

$$\omega =$$

$$y_i = \langle y_{i1}, \dots, y_{ip} \rangle^T$$

$$\Rightarrow \omega^T x_i - y_i = \langle z_{i1}, \dots, z_{ip} \rangle^T$$

$$\| \omega^T x_i - y_i \|_2^2 = (z_{i1}^2 + \dots + z_{ip}^2)$$

We'll convert $\tilde{J}(w)$ in terms of X, Y :

$$Y_i = w_{0i} + w_{1i}x_{i1} + \dots + w_{di}x_{id} + \underbrace{\epsilon_i}_{\text{error}}$$

Thus,

$$X = \begin{pmatrix} 1 & x_{11} & \dots & x_{1d} \\ 1 & ; & ; & \\ ; & ; & ; & \\ 1 & x_{n1} & \dots & x_{nd} \end{pmatrix}$$

$$Y = \begin{pmatrix} y_{11} & y_{12} & \dots & y_{1p} \\ ; & ; & ; & \\ y_{np} & \dots & y_{np} \end{pmatrix}$$

Get L₂ norm equivalent:

$$\left(\text{Consider } \|Xw - Y\| \right)$$

$$\text{Now } Xw - Y \in \mathbb{R}^{np}$$

$$= \begin{pmatrix} z_{11} & z_{12} & \dots & z_{1p} \\ ; & ; & ; & \\ z_{n1} & \dots & z_{np} \end{pmatrix}$$

$$\text{where } z_{ij} = (w_{0j} + w_{1j}x_{i1} + \dots + w_{dj}x_{id})$$

Consider $(x\omega - y)(x\omega - y)^T$

The summation for \sum

$$\sum_{i=1}^n \|\omega^T x_i - y_i\|_2^2$$

occurs on the diagonal of

$$(x\omega - y)(x\omega - y)^T$$

$$\Rightarrow J(\omega) = \text{tr} \left[(x\omega - y)(x\omega - y)^T \right]$$

Closed form:

$$= \text{tr}(y y^T) - 2\text{tr}((x\omega)^T y) + \text{tr}((x\omega)^T x\omega)$$

$$J(\omega) = \text{tr}(y^T y) - 2\text{tr}(y^T x\omega) + \text{tr}(\omega^T x^T x\omega)$$

Closed form:

$$\frac{\partial J(\omega)}{\partial \omega} = -2 \frac{\partial}{\partial \omega} (y^T x\omega) + \frac{\partial}{\partial \omega} (\omega^T x^T x\omega)$$

$$0 = -2x^T y^T + 2(x^T x\omega)$$

$$x^T y = x^T x\omega$$

$$\hat{\omega} = (x^T x)^{-1} x^T y$$

4 b)

$$Y \in \mathbb{R}^{m \times p}, W \in \mathbb{R}^{d \times p}, X \in \mathbb{R}^{m \times d}$$

Closed form

$$\hat{W} = (X^T X)^{-1} X^T Y$$

Consider i^{th} column of Y

$$y_i = \begin{pmatrix} y_{1i} \\ \vdots \\ y_{mi} \end{pmatrix}$$

$$\text{Now, } (X(X^T X)^{-1} X^T) \overset{X}{=} H \in \mathbb{R}^{d \times m}$$

and this matrix H is similar for a linear regression problem and a multivariate linear regression problem

For $H y_i \in \mathbb{R}^{d \times 1}$ which corresponds to the i^{th} column of \hat{W} .

Thus, for y_i , $i = 1, \dots, p$, we get the p columns of \hat{W}

$\hat{w}_i = i^{\text{th}}$ column of \hat{W}

$$= \begin{pmatrix} w_{1i} \\ \vdots \\ w_{di} \end{pmatrix} = \begin{pmatrix} (h_{11} y_{1i} + \dots + h_{1m} y_{mi}) \\ \vdots \\ (h_{d1} y_{1i} + \dots + h_{dm} y_{mi}) \end{pmatrix}$$

$$\Rightarrow \hat{\omega} = (\mathbf{x}^T \mathbf{x})^{-1} \mathbf{x}^T \mathbf{y}$$

$$\Leftrightarrow \hat{\omega}_i = (\mathbf{x}^T \mathbf{x})^{-1} \mathbf{x}^T \mathbf{y}_i \quad i=1, \dots, p$$

where $\hat{\omega}_i, y_i$ are i^{th} column of $\hat{\omega}, \mathbf{y}$.

\Rightarrow The multivariate Regression can be broken down into p separate linear regression problems.

This solution isn't appropriate if there exists any sort of relationship b/t the dependant (y) variables.

Ex. Using time, season, weather, temperature predict the number of people visiting a bike shop for the day and the revenue for the day.

In this example, revenue (y_2) may depend on the number of people predicted to visit the shop (y_1), and thus, treating y_1 and y_2 as independent and solving using

$$\hat{\omega} = (\mathbf{x}^T \mathbf{x})^{-1} \mathbf{x}^T \mathbf{y}$$

may not be appropriate.

$$4c) \quad W \in \mathbb{R}^{d \times p}$$

Let W be of low rank R .
 Since every finite-dimensional matrix has
 a rank decomposition

$$W = AB, \quad A \in \mathbb{R}^{d \times R}, \quad B \in \mathbb{R}^{R \times p}, \quad X \in \mathbb{R}^{d \times m}$$

Thus, low rank regression:

$$\min_{A, B} \|Y - X^T AB\|^2, \quad X \in \mathbb{R}^{d \times m}, \quad Y \in \mathbb{R}^{m \times p}$$

$$\equiv \min_{A, B} \|Y - (A^T X)^T B\|^2$$

Thus, low-rank regression can be viewed as standard linear regression but in the $\overline{X} = A^T X \in \mathbb{R}^{R \times m}$ subspace.

Consider i^{th} column of Y

$$= \forall i=1, \dots, p \quad \min_{A, B} \|y_i - (A^T x_i)^T B\|$$

$A^T x_i \in \mathbb{R}^{R \times 1}$ which is a R -dimensional subspace \mathbb{R}^R

let $\bar{X}_i = A^T x_i$

Thus, $y_i^T B = \bar{X}_i^T B R + \epsilon_i^{2 \times p}$

Since this is R adds to linear regression in a R -dimensional subspace of \mathbb{R}^p , this leaves a $(p-R)$ -dimensional of just error terms

$$\epsilon_{i=(p-R), \dots, p} \sim N(0, \sigma^2 I)$$

$\Rightarrow \exists$ $(p-R)$ -dimensional subspace of \mathbb{R}^p containing only noise.