

# Factorial Designs I

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## Introduction

- We have covered the case examining changes in a single factor.
- In practice, we may want to study the effects of multiple factors.
- Classical approach: *study each factor separately while holding others constant*.
  - Inefficient, and can cause the experimenter to miss interactions.
- **Factorial designs** accentuate the factor effects, allow for estimation of inter-dependency of effects (or interactions).
  - They are the first technique in the category of what is called treatment design.

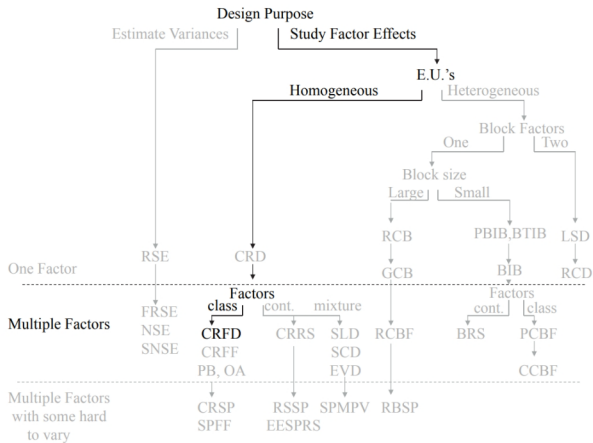


Figure: Source: (1)

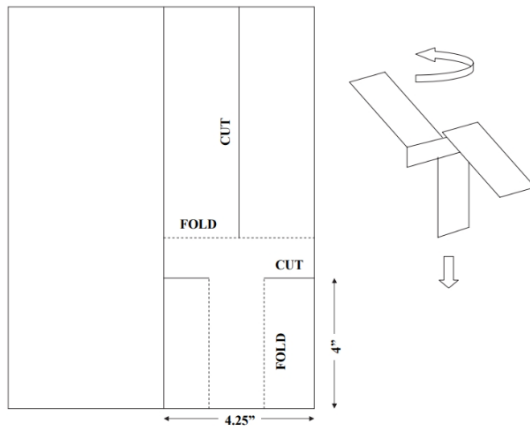
## Factorial Designs

- Factorial Designs examine all possible combinations of factor levels.
- The number of replicates of a specific level of one factor is increased by the product of the number of levels of all other factors in the design.
- The same power or precision can be obtained with fewer replicates.

## Illustrative Example 1 I

- Suppose we would like to create paper helicopters from a fixed sheet of paper.
- We then drop the helicopters from a fixed height and time how long it takes for them to hit the ground.
- We can trim paper off of the wings to reduce weight, but this will also reduce wing surface area.
- **We would like to conduct an experiment to determine if changing the wing length changes the flight time**

# Paper Helicopter



## Illustrative Example 1 II

- Identify the following:
  - Experimental unit
  - Difference between replicates and duplicates for this design
  - Treatment factor
  - Possible lurking variables
  - Why randomize



## Illustrative Example 2 I

- Now, let us assume that we can also vary the body width of the helicopters, thus giving us two factors.
- Assume that we examine four levels for each of the two factors:
  - Wing Length: 4, 4.75, 5.5, 6 (inches)
  - Body Width: 3.25, 3.75, 4, 4.25 (inches)
- How does the *classical* design approach differ from a *factorial* design?

## Classic Approach

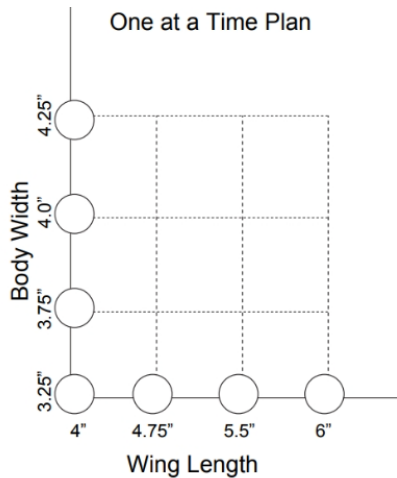


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## Factorial Design

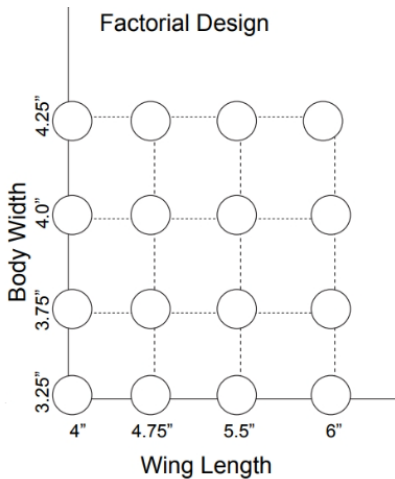


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## Illustrative Example 2 II

- We can see that experiments are ran at all possible combinations in a factorial design.
- By running two replicates at each combination we get 32 total runs (8 replicates at each individual level).
- We would need to make 56 runs if we took a *classic* approach to achieve the same number of replicates.
- Sometimes referred to as **hidden replication**.

## Interpreting Interactions

- If an interaction exists, the effect of one factor on the response will differ depending on the level of the other factor.
- We can examine possible interactions using visualizations using grouped scatterplots.
- Interactions are common in factorial experiments, so it is important to be able to clearly explain the results.

# Scatterplots

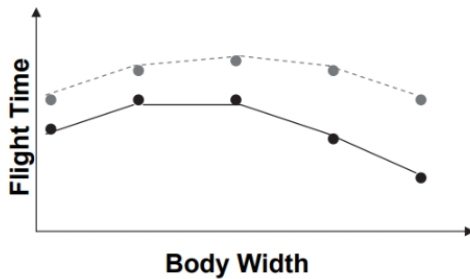


Figure: Source: (1)

## Interaction Terms

- **Interaction terms** occur when one explanatory variable modulates the effect of another on the response variable.
- It does **NOT** refer to a relationship between two variables.
- You just have to remember to multiply the coefficient by the product of all the variables in the term.

## Example 1

- Assuming the bread experiment was conducted again using a factor of two different temperatures (15C and 20C), import the *Bread.csv* data into R.
- Create a plot of the results to see if an interaction may be present.



## Two-Factor Designs in R I

- *There are many ways to create this kind of design in R.*
- We can use the `expand.grid()` function to create a data frame of all possible combinations:
  - ```
D <- expand.grid(Factor.1 =  
  as.factor(c(level.1,...,level.i)),  
  Factor.2 = as.factor(c(level.1,...,level.j)) )
```
- To replicate you can use the `rbind()` function:
  - ```
D <- rbind(D,D)
```

## Two-Factor Designs in R II

- Randomization:
  - `set.seed(2030) # Reproducible`
  - `D <- D[order(sample(1:nrow(D))), ] # Randomization`
  - `write.csv(D, file="Design.csv") # save optional`

## Example 2

- Assume that we examine four levels for each of the two factors from the helicopter experiment:
  - Wing Length: 4, 4.75, 5.5, 6 (inches)
  - Body Width: 3.25, 3.75, 4, 4.25 (inches)
- Use R to come up with a randomized two-factor design with two replications for each combination.

# Analysis

## Mathematical Model (Two Factors)

- The mathematical model for a completely randomized two-factor factorial design can be written as:

$$Y_{ijk} = \mu_{ij} + \epsilon_{ijk}. \quad (1)$$

- $i$  represents the level of the first factor.
- $j$  represents the level of the second factor.
- $k$  represents the replicate number.
- This model is called a *cell means model* and  $\mu_{ij}$  represents the expected response in the  $ij^{th}$  cell.

## Alternative Mathematical Model

- The effects model:

$$Y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \epsilon_{ijk}. \quad (2)$$

- $\alpha_i$  and  $\beta_j$  are the main effects:
  - $\alpha_i$  represents the difference between the marginal average of all experiments at the  $i^{th}$  level of the first factor and the overall average.
  - $\beta_j$  represents the difference between the marginal average at the  $j^{th}$  level of the second factor and the overall average.
- The interaction effects  $\alpha\beta_{ij}$  represents the difference in the cell mean  $\mu_{ij}$  and  $\mu + \alpha_i + \beta_j$ .

## Effects Model Definitions and Assumptions

- Definitions:

- $\sum_i \alpha_i = 0$
- $\sum_j \beta_j = 0$
- $\sum_i \alpha\beta_{ij} = 0$
- $\sum_j \alpha\beta_{ij} = 0$

- Assumptions:

- Experimental errors are independent (randomization)
- $\epsilon_{ijk} \sim N(0, \sigma^2)$

## Examining Results in R

- To examine the cell means in R:

```
lm1 <- lm(response ~ Factor1:Factor2, data = data)
```

- To examine the effects model in R:

```
lm2 <- lm(response ~ Factor1 + Factor2 +  
Factor1:Factor2, data = data)
```



## Example 3 Preliminaries

- In R, import the *COdata* dataset into your environment.
- This data comes from an experiment that examined the amount of CO emissions released from burning fuel.
- The response variable ( $y$ ) is the CO emissions concentrated in grams/metre<sup>3</sup>.
- There are two factors:
  - Eth: ethanol additions with three levels (0.1, 0.2, 0.3)
  - Ratio: air/fuel ratio with three levels (14, 15, 16)
- There were ( $r = 2$ ) replicate runs for each combination of factor levels.

## Example 3

- Use R to examine the cell means from the CO experiment.
- Use R to examine the effects model from the CO experiment.

## Analysis of Variance Table (ANOVA)

- When there are  $a$  levels for factor A,  $b$  levels for factor B, and  $r$  replications for each group, the ANOVA table for a two-factor design:

Source	df	Sum of Squares (ss)	Mean Squares	F-ratio
A	$a - 1$	$R(\alpha \mu)$	$\frac{ssA}{(a-1)}$	$F = \frac{msA}{msE}$
B	$b - 1$	$R(\beta \alpha, \mu)$	$\frac{ssB}{(b-1)}$	$F = \frac{msB}{msE}$
AB	$(a - 1)(b - 1)$	$R(\alpha\beta \beta, \alpha, \mu)$	$\frac{ssAB}{(a-1)(b-1)}$	$F = \frac{msAB}{msE}$
Error	$ab(r - 1)$	$ssE$	$\frac{ssE}{ab(r-1)}$	

- Where:
  - $R(\alpha|\mu)$ : sum of squares for a model only including A
  - $R(\beta|\alpha, \mu)$ :  $ssE_A - ssE_B$  where  $ssE_B$  is from a reduced model.
  - $R(\alpha\beta|\beta, \alpha, \mu)$ :  $ssE_B - ssE$

## Two Factor Design Analysis R

- *Is there a significant interaction of the explanatory variables?*
- `model <- aov(response ~ Factor1*Factor2, data = data)`
- `summary(model)`

## Example 4

- Use R to construct an ANOVA table for the results of the CO experiment.
- Are there significant differences in the levels of the factors and their interaction?

## `model.tables()`

- The `model.tables()` function breaks down the results from the ANOVA.
- Usage: `model.tables(model, type = "means", se = T)`

## Example 5

- Use the `model.tables()` function to examine *tables of means* from the CO experiment.
- Are any of these results similar to what was found by estimating regression models?

## Main Effects Contrasts I

- To estimate contrasts of the main effects, the `estimable()` function from the *gmodels* package can be used:
  - `c1 <- c(-1/2, 0, 1/2)` # first vs third
  - `c1 <- c(1/2, -1, 1/2)` # orthogonal vector
  - `mod2 <- aov(C0 ~ Eth * Ratio, contrasts =  
list(Eth = cm, Ratio = cm ), data = C0data) #  
restricted model`
  - `c <- rbind("Ethanol 0.3 vs 0.1" = c(0,1,0,0,0,0,0,0,0),  
"Ratio 16 vs 14" = c(0,0,0,1,0,0,0,0,0) ) # specific  
contrasts`
  - `estimable(mod2,c) # hypothesis tests`



## Main Effects Contrasts II

- We are testing the null hypothesis that there are no differences in the main effects at the specific factor levels.
- **The estimates are only meaningful if there are no significant interactions between the two factors.**

## Example 6

- Adjust the `cm` and `c` matrices from the example code to compare the following fixed-effects contrasts:
  - Ethanol 0.3 vs 0.2
  - Ratio 15 vs 16

## Plotting Interactions in R I

- We can use visualizations to examine possible interactions in R:
- From our example:
  - `with(COdata, (interaction.plot(Eth, Ratio, CO, type = "b", pch = c(18,24,22), leg.bty = "o", main = "Interaction Plot of Ethanol and air/fuel ratio", xlab = "Ethanol",ylab = "CO emissions")))`

## Example 7

- Adjust the example code to visualize how the effects of the Ratio on the CO emissions are modulated by the Ethanol levels.

## Number of Replicates

- Replicates are needed to estimate the variance of the error in order to use the ANOVA  $F$ -test.
- The power of the test will increase as the number of replicates increases (could be expensive).
- Determining the number of replicates needed to achieve a desired power of a test falls into the category of *error control*.
- **The experimenter must select a value of practical difference ( $\Delta$ ).**
  - Usually obtained through the previous/pilot studies.

## Number of Replicates in R (cell means)

- We can use the `Fpower1()` function from the *daewr* package:
  - `library(daewr)`
  - `rmin <- s` # smallest number of replicates considered
  - `rmax <- m` # largest number of replicates considered
  - `alpha <- 0.05`
  - `sigma <- sqrt(sigma_squared)` # from pilot study
  - `nlev <- levels1*levels2` # **total number of factor combinations**
  - `nreps <- c(rmin:rmax)`
  - `Delta <- delta` # practical difference
  - `power <- Fpower1(alpha,nlev,nreps,Delta,sigma)`
  - `power` # results
- Rule of thumb: *The number of replicates that result in a power between 0.80 and 0.90 is usually sufficient for most designs.*

## Number of Replicates in R (marginal means)

- We can use the `Fpower2()` function from the *daewr* package:
  - `library(daewr)`
  - `rmin <- s` # smallest number of replicates considered
  - `rmax <- m` # largest number of replicates considered
  - `alpha <- 0.05`
  - `sigma <- sqrt(sigma_squared)` # from pilot study
  - `nlev <- c(levels1, levels2)` # **numbers of factor levels**
  - `nreps <- c(rmin:rmax)`
  - `Delta <- delta` # practical difference
  - `power <- Fpower2(alpha, nlev, nreps, Delta, sigma)`
  - `power` # results
- **Hidden replication increases the efficiency of the experiments.**

## Example 8

- Use both methods (cell means, marginal means) to determine how many replicates are needed for the CO emissions experiment.
- We can make the following assumptions:
  - fewest: 2
  - most: 8
  - $\hat{\sigma}^2 = 32$
  - $\Delta = 10$
- Comment on your results.



## Unequal Number of Replicates

## Unequal Number of Replicates

- It is uncommon to plan a factorial experiment with an unequal number of replicates per cell.
- In practice, there may be problems with the data collection.
- **As long as the chance of missing observations is not related to the treatment factor levels** the data can still be analysed.
- We need to slightly modify the code we use to perform this analysis.

## Modified Code in R

- Assume that there is now an unequal number of replicates in the data.
- We can use the `contr.sum` argument and the *car* package in R to obtain meaningful results:
  - `library(car)`
  - `model <- lm(response ~ Factor1*Factor2, data = data, contrasts = list(Factor1 = contr.sum, Factor2 = contr.sum))`
  - `Anova(model, type = "III")`
- *Sum contrasts are similar to mean-centering a particular continuous explanatory variable.*
- `type = "III"` indicates a Partial Sum of Squares (used when observations are missing)

## Main Effects Contrasts Unequal Number of Replicates

- We can use estimate the marginal adjusted means for the main effects.
- We can use the *lsmeans* package in R:
  - `library(lsmeans)`
  - `lsmeans(model, "Factor")`
- **Again, the estimates are only meaningful if there are no significant interactions between the two factors.**

## Example 9 from slides

- Assume that we are missing one observation from the CO emissions experiment (see code).
- In R conduct the following analyses:
  - 1 Construct an ANOVA table for the data with an unequal number of replicates.
    - Does the missing observation impact our overall results?
  - 2 Use the `lsmeans()` to examine the marginal adjusted means for both of the main effects factors.
    - Are these results reliable?

## Thoughts on Unequal Number of Replicates

- The `aov()` function in R **will not** test the hypotheses in the same way as it does with equal replicates.
- When there is an equal number of replications per cell, the sums of squares computed by the `aov()` function are identical to the type III sums of squares.
- **In general, you can use the methods applied in Example 9 for both cases** (equal and unequal replicates).

## One Replicate per Cell

## One Replicate per Cell I

- When there is enough power, experiments may be ran with one replicate ( $r = 1$ ) per cell.
- With only one replicate per cell, we cannot calculate the ANOVA  $ssE$ 
  - Cannot conduct  $F$ -tests on the main effects and interactions in the traditional way.
- If there is no interaction, we can use an additive model:
$$y_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$$
  - **Dangerous if an interaction actually exists!**



## One Replicate per Cell II

- If both factors are numeric and have equally spaced levels then the sum of squares for the interaction can be partitioned.
  - We can estimate an error term and test the significance of the linear part of the interaction.
- We will order the factors and then examine any significance in a linear interaction.

## One Replicate per Cell Method I

- In R we can order our data, assume a linear interaction, and then test for the significance of the interaction:
  - `Factor1.Order <- as.ordered(data$Factor1)`
  - `Factor2.Order <- as.ordered(data$Factor2)`
  - `Factor1.Lin <-  
 contr.poly(Factor1.Order)[Factor1.Order, ".L"]`
  - `Factor2.Lin <-  
 contr.poly(Factor2.Order)[Factor2.Order, ".L"]`
  - `model <- lm(Response ~ Factor1.Order + Factor2.Order +  
 Factor1.Lin:Factor2.Lin, data = data)`
  - `anova(model)`

## One Replicate per Cell Method II

- In R we can directly test for significance of the interaction using the `Tukey1df()` function:
  - `Tukey1df(data)`
- Note: The response variable needs to be in the first column of the data frame.
  - ```
data <- data %>%  
  relocate(response)
```
- **Much easier to code but slightly different to read than we are used to.**
  - If the `NonAdditivity` term is significant then we may assume that there is an interaction.

## Example 10

- Assuming that we only received one observation from the CO experiment.
- Use Method I and Method II to determine if there is a significant interaction between the Ethanol and the Ratio variables.
- *Note: The data has been adjusted to assume that the group means are our only observation.*

## Example 11

- Import the *virus* dataset into R.
- Take some time to get to know the data.
- Use Method I and Method II to determine if there is a significant interaction between the `Sample` and the `Dilution` variables.
- How would you obtain the coefficient estimates for an additive model if the interactions are not significant?

## Exercise 1

- Using the *Bread.csv* data conduct the following analysis:
  - ① Use visualizations and an appropriate statistical method to determine if an interaction between the two factors exists.
  - ② Examine the main effects using whichever method you prefer.
    - Are these estimates meaningful?
  - ③ How many replicates are needed for this experiment assuming the following information:
    - $\hat{\sigma}^2 = 2.1$  (from pilot)
    - $\Delta = 3$

## Exercise 2

- Remove one observation from the *Bread.csv* data and repeat (1) and (2) from Exercise 1.
- Rework the code from Example 10 to generate a data frame of the cell means and use the appropriate methods to repeat (1) and (2) from Exercise 1.
  - Are the results much different?

## References & Resources

- 1 Lawson, J. (2014). *Design and Analysis of Experiments with R (Vol. 115)*. CRC press.

- model.tables
- gmodels
- Fpower2
- estimable
- Anova
- lsmeans