

PSTAT 120B HW #3

Reading Outline

- 1)
 - An estimator is a rule, often expressed as a formula, that tells how to calculate the value of an estimate based on the measurements contained in the sample.
 - An experimenter who wants an interval estimate of a parameter must use the sample data to calculate two values, chosen so that the interval formed by the two values includes the target parameter with a specified probability.
- 2)
 - A point estimation produces a guess for the exact value of the target parameter.
 - An interval estimation produces a range of possible values for the target parameter.
 - One way to view this is that a point estimate involves a single numerical value, while an interval estimate is a range of numbers.
- 3)
 - Bias is what quantifies how far an estimator is expected to be from the target parameter.
 - The bias of an estimator is $B(\hat{\theta}) = E(\hat{\theta} - \theta)$, which shows the difference between the estimator's expected value and the true value of the target parameter.
- 4)
 - The mean square error is a function of both its variance and its bias. The MSE quantifies the spread of an estimator about the target parameter.
 - $MSE(\hat{\theta}) = E(\hat{\theta} - \theta)^2 = \frac{\text{Var.}}{n}$
 - $MSE = \text{var}(\hat{\theta}) + (B(\hat{\theta}))^2$

5) - An unbiased estimator for the population mean of a random sample is $\hat{\mu}$ where $\hat{\mu} = \bar{Y}$ if $(Y_1, \dots, Y_n) \stackrel{iid}{\sim} F(y)$

6) - An unbiased estimator for the population variance of random sample is $\hat{\sigma}^2$ where $\hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$ if $(Y_1, \dots, Y_n) \stackrel{iid}{\sim} F(y)$

1. a) $Be^{-\beta y}, y \geq 0$

$$E(Y) = \beta \int_0^{\infty} e^{-\beta y} y dy = \beta \cdot \frac{1!}{\beta^2} = \frac{1}{\beta}$$

$$E(Y^2) = \beta \int_0^{\infty} e^{-\beta y} y^2 dy = \beta \cdot \frac{2!}{\beta^3} = \frac{2!}{\beta^2} = \frac{2}{\beta}$$

$$\text{Var}(Y) = E(Y^2) - E^2(Y) = \frac{2}{\beta^2} - \frac{1}{\beta^2} = \frac{1}{\beta^2}$$

$$\frac{\sqrt{n}(\bar{Y} - \frac{1}{\beta})}{\sqrt{1/\beta^2}} \sim N(0, 1)$$

$$\bar{Y} - \frac{1}{\beta} \sim N(0, \frac{1}{n\beta^2})$$

$$\bar{Y} \sim N(\frac{1}{\beta}, \frac{1}{n\beta^2})$$

b) $f(y) = \frac{1}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} e^{-\frac{y}{2}} y^{\frac{\nu}{2}-1}$

$$E(Y) = \frac{1}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} \cdot \int_0^{\infty} e^{-\frac{y}{2}} y^{\frac{\nu}{2}} dy$$

$$= \frac{1}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} \times \frac{\Gamma(\frac{\nu}{2}+1)}{(\frac{1}{2})^{\frac{\nu}{2}+1}} = \frac{\frac{\nu}{2} \Gamma(\frac{\nu}{2})}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} \times 2^{\frac{\nu}{2}+1} = \boxed{\nu}$$

$$E(Y^2) = \frac{1}{\Gamma(\frac{\nu}{2}) 2^{\nu/2}} \int_0^{\infty} e^{-\frac{y}{2}} y^{\frac{\nu}{2}+1} dy = \frac{(\frac{\nu}{2}+1) \frac{\nu}{2} \Gamma(\frac{\nu}{2}) \times 2^2}{\Gamma(\frac{\nu}{2})} = \boxed{2\nu(\frac{\nu}{2}+1)}$$

$$\text{Var}(Y) = E(Y^2) - E(Y)^2 = 2\nu(\frac{\nu}{2}+1) - \nu^2 = \nu^2 + 2\nu - \nu^2 = \boxed{2\nu}$$

$$\bar{Y} \sim N(\nu, 2\nu/n)$$

c) $E(Y) = \sum_{y=0}^{\infty} \frac{e^{-\lambda} \lambda^y}{y!} y = \lambda \sum_{y=1}^{\infty} \frac{e^{-\lambda} \lambda^{y-1}}{(y-1)!} = \lambda$

$$E(Y(Y-1)) = \sum_{y=0}^{\infty} \frac{e^{-\lambda} \lambda^y}{y!} y(y-1) = \lambda^2 \sum_{y=2}^{\infty} \frac{e^{-\lambda} \lambda^{y-2}}{(y-2)!} = \lambda^2$$

$$E(Y^2) - E(Y) = \lambda^2$$

$$\text{Var}(Y) = E(Y^2) - E(Y)^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

$$\frac{\sqrt{n}(\bar{Y} - \lambda)}{\sqrt{\lambda}} \sim N(0, 1)$$

$$\bar{Y} \sim N(\lambda, \frac{\lambda}{n})$$

d) $f_1(y) = N(\mu_1, \sigma_1^2), f_2(y) = N(\mu_2, \sigma_2^2)$

$$c f_1(y) = N(c\mu_1, c^2\sigma_1^2)$$

$$(1-c) f_2(y) = N((1-c)\mu_2, (1-c)\sigma_2^2)$$

$$E(Y) = c\mu_1 + (1-c)\mu_2$$

$$\text{Var}(Y) = c^2\sigma_1^2 + (1-c)^2\sigma_2^2$$

$$\bar{Y} = (c\mu_1 + (1-c)\mu_2) \sim N(0, \frac{c^2\sigma_1^2 + (1-c)^2\sigma_2^2}{n})$$

$$\bar{Y} \sim N(c\mu_1 + (1-c)\mu_2, \frac{c^2\sigma_1^2 + (1-c)^2\sigma_2^2}{n})$$

$$2.) a) \bar{X} = \frac{1}{m} \sum_{i=1}^m X_i$$

$$E[\bar{X}] = \frac{1}{m} \sum_{i=1}^m E[X_i]$$

$$= \frac{1}{m} \sum_{i=1}^m \mu_1$$

$$= \mu_1$$

$$V[\bar{X}] = V\left(\frac{1}{m} \sum_{i=1}^m X_i\right)$$

$$= \frac{1}{m^2} \sum_{i=1}^m V[X_i]$$

$$= \frac{1}{m^2} \times m\sigma^2$$

$$= \sigma^2/m$$

Knowing this,

$$E[\bar{Y}] = \mu_2, V[\bar{Y}] = \frac{\sigma_2^2}{n}$$

$$E[\bar{X} - \bar{Y}] = E[\bar{X}] - E[\bar{Y}]$$

$$= \mu_1 - \mu_2$$

$$b) V[\bar{X} - \bar{Y}] = V[\bar{X}] + V[\bar{Y}] - 2 \text{Cov}[\bar{X}, \bar{Y}]$$

$$= \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n} - 0$$

$$V[\bar{X} - \bar{Y}] = \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}$$

$$c) \sigma_1^2 = 2, \sigma_2^2 = 2.5, m=n$$

$$P(|\bar{X} - \bar{Y} - (\mu_1 - \mu_2)| \leq 1) = .95$$

$$P(-1 \leq \{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)\} \leq 1) = .95$$

$$V[\bar{X} - \bar{Y}] = \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}$$

$$= \frac{2+2.5}{n}$$

$$= 4.5/n$$

$$P\left(\frac{-1}{\sqrt{4.5/n}} \leq Z \leq \frac{1}{\sqrt{4.5/n}}\right) = .95$$

$$2 \times P(Z \leq .47\sqrt{n}) - 1 = .95$$

$$P(Z \leq .47\sqrt{n}) = .975$$

$$.47\sqrt{n} = 1.96$$

$$n = 17.39$$

$$\boxed{\text{At least } 18}$$

$$3.) E(X_i) = \mu_1, E(Y_i) = \mu_2$$

$$V(X_i) = \sigma_1^2, V(Y_i) = \sigma_2^2$$

$$W_i = X_i - Y_i$$

$$\text{for } i = 1, 2, \dots, n$$

$$\text{Let } \mu = E(W_i)$$

$$E(W_i) = E(X_i) - E(Y_i)$$

$$E(W_i) = \mu_1 - \mu_2$$

$$\text{Var}(W_i) = \text{Var}(X_i - Y_i)$$

$$= V(X_i) + V(Y_i) - 2 \text{Cov}(X_i, Y_i)$$

$$\text{Var}(W_i) = \sigma_1^2 + \sigma_2^2 - 0$$

$$U_n = \frac{\sum_{i=1}^n W_i - n\mu}{\sigma/\sqrt{n}} = \frac{\bar{W} - \mu}{\sigma/\sqrt{n}}$$

$$= \frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{\sqrt{(\sigma_1^2 + \sigma_2^2)/n}}$$

converges to std.
normal as $n \rightarrow \infty$.

$$\begin{aligned}
 4) \quad \hat{\theta} - \theta &= \hat{\theta} - E(\hat{\theta}) + E(\hat{\theta}) - \theta \\
 &= [\hat{\theta} - E(\hat{\theta})] + [E(\hat{\theta}) - \theta] \\
 &= [\hat{\theta} - E(\hat{\theta})] + B(\hat{\theta})
 \end{aligned}$$

$$\begin{aligned}
 MSE(\hat{\theta}) &= E(\hat{\theta} - \theta)^2 \\
 &= E([\hat{\theta} - E(\hat{\theta})] + B(\hat{\theta}))^2 \\
 &= E([\hat{\theta} - E(\hat{\theta})]^2 + 2[\hat{\theta} - E(\hat{\theta})]B(\hat{\theta}) + [B(\hat{\theta})]^2) \\
 &= V(\hat{\theta}) + 2[E(\hat{\theta}) - E(\hat{\theta})]B(\hat{\theta}) + [B(\hat{\theta})]^2 \\
 &= V(\hat{\theta}) + 2(0) \cdot B(\hat{\theta}) + [B(\hat{\theta})]^2 \\
 &= V(\hat{\theta}) + 0 + [B(\hat{\theta})]^2 \\
 &= \underline{V(\hat{\theta}) + [B(\hat{\theta})]^2}
 \end{aligned}$$

5.) a) Given $\hat{\theta}$ is unbiased,

$$\begin{aligned}
 E(\hat{\theta}) &= \theta \\
 B(\hat{\theta}) &= E(\hat{\theta}) - \theta \\
 &= \theta - \theta \\
 \boxed{B(\hat{\theta}) &= 0}
 \end{aligned}$$

b) $B(\hat{\theta}) = 5$

$$\begin{aligned}
 B(\hat{\theta}) &= E(\hat{\theta}) - \theta \\
 \Rightarrow E(\hat{\theta}) - \theta &= 5 \\
 \boxed{E(\hat{\theta}) &= \theta + 5}
 \end{aligned}$$

$$\begin{aligned}
 6.) \quad a) \quad E(\hat{\theta}) &= a\theta + b \\
 B(\hat{\theta}) &= E(\hat{\theta}) - \theta \\
 &= a\theta + b - \theta \\
 &= (a-1)\theta + b
 \end{aligned}$$

$$\boxed{B(\hat{\theta}) = b + \theta(a-1)}$$

$$\begin{aligned}
 b) \quad E(\hat{\theta}) &= a \cdot \theta + b \\
 E(\hat{\theta}^*) &= E\left(\frac{\hat{\theta} - b}{a}\right) \\
 &= \frac{E(\hat{\theta}) - b}{a} \\
 &= \frac{a\theta + b - b}{a} \\
 &= \theta
 \end{aligned}$$

$$\boxed{\hat{\theta}^* = \frac{\hat{\theta} - b}{a} \text{ is unbiased.}}$$

$$\begin{aligned}
 7.) \quad MSE(\hat{\theta}) &= E[(\hat{\theta} - \theta)^2] \\
 &= V(\hat{\theta}) + [B(\hat{\theta})]^2 \\
 a) \quad B(\hat{\theta}) &= E(\hat{\theta}) - \theta \\
 &= \theta - \theta \\
 &= \underline{0}
 \end{aligned}$$

$$\begin{aligned}
 MSE(\hat{\theta}) &= V(\hat{\theta}) + [B(\hat{\theta})]^2 \\
 &= V(\hat{\theta}) + 0 \\
 &= \boxed{V(\hat{\theta})} \\
 MSE(\hat{\theta}) &= V(\hat{\theta})
 \end{aligned}$$

$$\begin{aligned}
 b) \quad MSE(\hat{\theta}) &= V(\hat{\theta}) + B(\hat{\theta})^2 \\
 &= V(\hat{\theta}) + \text{some value} \\
 MSE(\hat{\theta}) &\geq V(\hat{\theta})
 \end{aligned}$$