

# PSTAT 120B - Quiz 1 (Shrawan Shetty)

1.)  $X = -n \log W$   
 $U_i = -\log Y_i$

a) where  $Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Unif}(0,1)$

$$\Rightarrow U_i \sim \text{Exp}(1)$$

$$\Rightarrow E(U_i) = 1$$

$$\Rightarrow \text{Var}(U_i) = 1$$

$$\Rightarrow m_{U_i}(t) = \frac{1}{1-t}$$

b)  $X = -n \log W$   
 $= \sum_{i=1}^n U_i$

$$\begin{aligned} M_X(t) &= E(e^{xt}) \\ &= E(e^{t \cdot \sum_{i=1}^n U_i}) \\ &= \prod_{i=1}^n E(e^{U_i \cdot t}) \\ &= \prod_{i=1}^n m_{U_i}(t) \\ &= \prod_{i=1}^n \frac{1}{1-t} \end{aligned}$$

$$M_X(t) = \frac{1}{(1-t)^n}$$

Gamma Distribution has a similar MGF with the same functional form

c) MGF from previous problem (1b) shows properties of gamma distribution, in which

$$X \sim \text{Gamma}(n, 1)$$

Because of this, we can find the density to be:

$$f_X(x) = \begin{cases} \frac{x^{n-1} e^{-x}}{\Gamma(n)} & x > 0, n \in \mathbb{N}, 1 > 0 \\ 0 & \text{otherwise} \end{cases}$$

Note: A special property of  $\Gamma(n) = (n-1)!$

Because of this, the density can also be written as:

$$f_X(x) = \begin{cases} \frac{x^{n-1} e^{-x}}{(n-1)!} & x > 0, n \in \mathbb{N}, 1 > 0 \\ 0 & \text{otherwise} \end{cases}$$

2.)  $W$  is strictly monotone function  
of  $X = -n \log W$

$$\begin{aligned} a) f(w) &= \frac{d}{dw} \cdot P(X \leq h^{-1}(w)) \\ &= \frac{d}{dw} \cdot F(h^{-1}(w)) \\ &= f(h^{-1}(w)) \left| \frac{d}{dw} h^{-1}(w) \right| \\ &= \frac{e^{n \cdot \log w} \cdot (-n \log w)^{n-1}}{\Gamma(n)} \cdot \left| \frac{-n}{w} \right| \end{aligned}$$

$$f(w) = \begin{cases} \frac{(w^n)n \cdot (-n \log w)^{n-1}}{\Gamma(n) \cdot (w)} & w \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} b) f(w) &= \int f(h^{-1}(w)) \left| \frac{d}{dw} h^{-1}(w) \right| dw \\ &\Rightarrow \int \frac{(-n \log w)^{n-1} (w^n)n}{w \cdot \Gamma(n)} \end{aligned}$$

(The quiz says to not compute  
this integral)

$$3.) -\log w = \bar{U}$$

$$\bar{U} \sim N(\mu, \frac{\sigma^2}{n})$$

$$\text{so } -\log w \sim N(\mu, \frac{\sigma^2}{n})$$

$$-\ln w = \frac{1}{n} \sum_{i=1}^n \ln Y_i$$

since  $w = \left( \prod_{i=1}^n Y_i \right)^{1/n}$   
 For personal reference  $\rightarrow$

$$a) n \cdot -\ln w = -\sum_{i=1}^n \ln Y_i$$

$$n \cdot -\ln w = \text{Gamma}(n, 1)$$

$$-\ln w = \text{Gamma}(n, 1/n)$$

$$E[-\ln w] = \frac{1}{n} \cdot n = 1$$

$$\text{Var}[-\ln w] = \left(\frac{1}{n}\right)^2 \cdot n = 1/n$$

Using CLT  $\left( \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \right)$ :

$$\frac{\bar{U} - \mu}{\sigma/\sqrt{n}} \Rightarrow \frac{-\ln w - E[-\ln w]}{\sigma/\sqrt{n}}$$

$$= \frac{-\ln w - 1}{1/\sqrt{n}} \sim N(0, 1)$$

$$\Rightarrow -\ln w \sim N(1, \frac{1}{n})$$

Now compare

$$N(1, \frac{1}{n}) \text{ and } N(\mu, \frac{\sigma^2}{n})$$

$$\boxed{\mu = 1, \sigma^2 = 1}$$

$$b) \sqrt{n} (-\log w - 1)$$

we already solved:  $-\ln w \sim N(1, \frac{1}{n})$

So, we can now do:

$$(-\log w - 1) \sim N(0, \frac{1}{n})$$

$$\boxed{\sqrt{n}(-\log w - 1) \sim N(0, 1)}$$

$$c) P(w \leq e^{-1}) = P(-\log w - 1 \geq 0) \\ = P(\sqrt{n}(-\log w - 1) \geq 0)$$

We know:

$$\sqrt{n}(-\ln w - 1) \sim N(0, 1) \xrightarrow[n \rightarrow \infty]{\lim_{n \rightarrow \infty} 1/n \rightarrow 0}$$

Thus,

$$\boxed{\Phi(0) \Rightarrow 1/2}$$