

# MF 790 HW 3 - SOLUTIONS

This assignment is due on Thursday, October 7th at 8 AM. The first problem is worth 20 points, while problems 2 and 3 are worth 15 points each.

**1. Itô Integrals Using Midpoint, Right Endpoint.** As seen in class, for a fixed  $t > 0$  if we approximate  $\int_0^t W_u dW_u$  by

$$(0.1) \quad \sum_{i=1}^n W_{t_{i-1}} (W_{t_i} - W_{t_{i-1}})$$

for a given partition  $\Pi$ , then taking  $\|\Pi\| \downarrow 0$  we obtain

$$\int_0^t W_u dW_u = \frac{1}{2} (W_t^2 - t), t \geq 0,$$

which (and this can be shown directly too) shows that  $t \rightarrow W_t^2 - t$  is a martingale.

In this exercise we will see what happens if, instead of evaluating  $W$  at the left side of the interval in (0.1), we evaluate it at the midpoint or the right side of the interval. To make calculations easier we will assume, for each  $t \geq 0$  that the partition is  $0 < t/n < 2t/n < \dots < nt/n = t$  and see what happens when  $n \uparrow \infty$ .

(a) (right hand side) Show

$$\lim_{n \uparrow \infty} \sum_{i=1}^n W_{t_i} (W_{t_i} - W_{t_{i-1}}) = \frac{1}{2} (W_t^2 + t)$$

Is  $t \rightarrow W_t^2 + t$  a martingale?

(b) (midpoint) Write  $W_i := W_{(t_i+t_{i-1})/2}$ . Show

$$\lim_{n \uparrow \infty} \sum_{i=1}^n W_i (W_{t_i} - W_{t_{i-1}}) = \frac{1}{2} W_t^2$$

Is  $t \rightarrow W_t^2$  a martingale?

**Solution**

(a) Note that

$$\begin{aligned} \sum_{i=1}^n W_{t_i} (W_{t_i} - W_{t_{i-1}}) &= \sum_{i=1}^n (W_{t_i} - W_{t_{i-1}})^2 + \sum_{i=1}^n W_{t_{i-1}} (W_{t_i} - W_{t_{i-1}}) \\ &\xrightarrow{n \uparrow \infty} t + \int_0^t W_u dW_u, \\ &= t + \frac{1}{2} (W_t^2 - t) = \frac{1}{2} (W_t^2 + t). \end{aligned}$$

$t \rightarrow W_t^2 + t$  is not a martingale. Indeed, using that  $t \rightarrow W_t^2 - t$  is martingale we find

$$\mathbb{E} [W_t^2 + t | \mathcal{F}_s] = 2t + \mathbb{E} [W_t^2 - t | \mathcal{F}_s] = 2(t - s) + W_s^2 + s > W_s^2 + s,$$

so that in fact, it is a sub-martingale.

(b) Note that

$$\begin{aligned} W_i (W_{t_i} - W_{t_{i-1}}) &= (W_i - W_{t_{i-1}}) (W_{t_i} - W_{t_{i-1}}) + W_{t_{i-1}} (W_{t_i} - W_{t_{i-1}}) \\ &= \frac{1}{2} (W_{t_i} - W_{t_{i-1}})^2 + W_{t_{i-1}} (W_{t_i} - W_{t_{i-1}}) \end{aligned}$$

so that

$$\sum_{i=1}^n W_i (W_{t_i} - W_{t_{i-1}}) = \frac{1}{2} \sum_{i=1}^n (W_{t_i} - W_{t_{i-1}})^2 + \sum_{i=1}^n W_{t_{i-1}} (W_{t_i} - W_{t_{i-1}})$$

The first term goes to  $t/2$  and the second term goes to  $\int_0^t W_u dW_u = \frac{1}{2}(W_t^2 - t)$ . This gives the first result. As for the martingale statement, note that

$$\mathbb{E} [W_t^2 | \mathcal{F}_s] = t + \mathbb{E} [W_t^2 - t | \mathcal{F}_s] = (t - s) + W_s^2,$$

so the process is not a martingale.

## 2. Practice with Itô's formula : Finding Martingales Associated to GBM.

Let  $\mu \in \mathbb{R}$  and  $\sigma > 0$  be constants. Recall that we say  $S$  is a Geometric Brownian Motion (GBM) with drift  $\mu$  and volatility  $\sigma^2$  if

$$S_t = S_0 e^{(\mu - \sigma^2/2)t + \sigma W_t}; \quad t \geq 0,$$

where  $S_0$  is a positive constant and  $W$  is a Brownian Motion. We write  $S \sim \text{GBM}(\mu, \sigma^2)$ .

- Give the Itô process decomposition for  $f(S_t)$  when  $f(x) = x^p, p \in \mathbb{R}$ . Is  $S^p \sim \text{GBM}(\alpha, \theta^2)$  for some  $\alpha, \theta$ ? For a fixed  $\mu, \sigma^2$ , is there a value of  $p$  for which  $S^p$  is a martingale?
- Give the Itô process decomposition for  $f(S_t)$  when  $f(x) = \log(x)$ . Is there a value of  $\mu, \sigma^2$  for which  $\log(S)$  is a Martingale?
- Give the Itô process decomposition for  $f(t, S_t)$  when  $f(t, x) = e^{-\lambda t} x^p$  for  $\lambda, p \in \mathbb{R}$ . Find  $\hat{\lambda}$  so that for any given  $\mu, \sigma$  and  $p$ ,  $t \rightarrow e^{-\hat{\lambda} t} S_t^p$  is a martingale. With this  $\hat{\lambda}$  compute  $\mathbb{E}[S_t^p]$  without using the p.d.f. for  $S_t$ .

**Hint:** For a given Itô process  $X$  with decomposition  $dX_t = \Theta_t dt + \Delta_t dW_t$  if  $\Theta_t(\omega) = 0$  for  $(t, \omega)$  then  $X$  is a Martingale.

### Solution:

- We have  $\dot{f}(x) = (p/x)f(x)$ ,  $\ddot{f}(x) = (p(p-1)/x^2)f(x)$  so that

$$\begin{aligned} dS_t^p &= \frac{p}{S_t} S_t^p dS_t + \frac{1}{2} \frac{p(p-1)}{S_t^2} S_t^p d[S, S]_t, \\ &= p S_t^{p-1} \mu dt + p S_t^{p-1} \sigma dW_t + \frac{1}{2} p(p-1) S_t^{p-2} \sigma^2 dt, \\ &= S_t^p \left( p\mu + \frac{1}{2} p(p-1) \sigma^2 \right) dt + S_t^{p-1} \sigma dW_t. \end{aligned}$$

Therefore,  $S^p \sim \text{GBM}(p(\mu + (1/2)(p-1)\sigma^2), \sigma^2)$ . Furthermore,  $S^p$  will be a martingale when the  $dt$  terms vanish, or equivalently, when

$$p = 1 - \frac{2\mu}{\sigma^2}.$$

(b) Since  $\log(x) = 1/x$ ,  $\ddot{\log}(x) = -1/x^2$ :

$$\begin{aligned} d\log(S)_t &= \frac{dS_t}{S_t} - \frac{1}{2} \frac{d[S, S]_t}{S_t^2}, \\ &= \left( \mu - \frac{1}{2}\sigma^2 \right) dt + \sigma dW_t. \end{aligned}$$

Thus, for  $\mu = (1/2)\sigma^2$  we have that  $\log(S)$  is a martingale.

(c) For the given  $f(t, x)$ , since  $x > 0$

$$\begin{aligned} f_t(t, x) &= -\lambda f(t, x); & f_x(t, x) &= \frac{p}{x} f(t, x); \\ f_{xx}(t, x) &= \frac{p(p-1)}{x^2} f(t, x). \end{aligned}$$

Using the time-dependent Itô's formula gives

$$\begin{aligned} df(t, S_t) &= -\lambda f(t, S_t)dt + \frac{p}{S_t} f(t, S_t)dS_t + \frac{1}{2} \frac{p(p-1)}{S_t^2} d[S, S]_t \\ &= f(t, S_t) \left( \left( -\lambda + \mu p + \frac{1}{2} p(p-1)\sigma^2 \right) dt + \sigma p dW_t \right) \end{aligned}$$

We thus get a martingale if  $\hat{\lambda} = \mu p + (1/2)p(p-1)\sigma^2$ . Since martingales have constant expectations we know that

$$1 = \mathbb{E} \left[ e^{-\hat{\lambda}t} S_t^p \right] \implies \mathbb{E} [S_t^p] = e^{\hat{\lambda}t} = e^{(p\mu + \frac{1}{2}p(p-1)\sigma^2)t}$$

**3. Derivation of Itô's formula for general  $f''$ .** Do exercise 4.14 on page 198–199 of the class textbook.

**Solution:**

(i)  $Z_j$  is clearly  $\mathcal{F}_{t_{j+1}}$  measurable since  $f''$  is continuous and  $W$  is adapted. Furthermore,

$$\mathbb{E} [Z_j | \mathcal{F}_t] = f''(W_{t_j}) (\mathbb{E} [(W_{t_{j+1}} - W_{t_j})^2] - (t_{j+1} - t_j)) = 0,$$

since  $W$  is adapted and  $W_{t_{j+1}} - W_{t_j}$  is independent of  $\mathcal{F}_{t_j}$  and normally distributed with mean 0 and variance  $t_{j+1} - t_j$ . Using the same properties of  $W$

yields

$$\begin{aligned}
\mathbb{E} [Z_j^2 | \mathcal{F}_t] &= f''(W_{t_j})^2 \mathbb{E} \left[ ((W_{t_{j+1}} - W_{t_j})^2 - (t_{j+1} - t_j))^2 \right], \\
&= f''(W_{t_j})^2 \mathbb{E} \left[ (W_{t_{j+1}} - W_{t_j})^4 - 2(W_{t_{j+1}} - W_{t_j})^2(t_{j+1} - t_j) + (t_{j+1} - t_j)^2 \right], \\
&= f''(W_{t_j})^2 (3(t_{j+1} - t_j)^2 - 2(t_{j+1} - t_j)^2 + (t_{j+1} - t_j)^2), \\
&= 2f''(W_{t_j})^2(t_{j+1} - t_j)^2.
\end{aligned}$$

(ii) Taking out what is known yields

$$\mathbb{E} \left[ \sum_{j=0}^{n-1} Z_j \right] = \sum_{j=0}^{n-1} \mathbb{E} [Z_j] = \sum_{j=0}^{n-1} \mathbb{E} [\mathbb{E} [Z_j | \mathcal{F}_{t_j}]] = 0.$$

(iii) Assuming  $\mathbb{E} \left[ \int_0^T f''(W_t)^2 dt \right] < \infty$  and using part (ii) gives

$$\begin{aligned}
\text{Var} \left[ \sum_{j=0}^{n-1} Z_j \right] &= \mathbb{E} \left[ \left( \sum_{j=0}^{n-1} Z_j \right)^2 \right] \\
&= \sum_{j=0}^{n-1} \mathbb{E} [Z_j^2] + 2 \sum_{j=0}^{n-1} \sum_{k=j+1}^{n-1} \mathbb{E} [Z_j Z_k].
\end{aligned}$$

Now, by part (i) for  $j < k$

$$\begin{aligned}
\mathbb{E} [Z_j^2] &= \mathbb{E} [\mathbb{E} [Z_j^2 | \mathcal{F}_{t_j}]] = 2(t_{j+1} - t_j)^2 \mathbb{E} [f''(W_{t_j})^2]; \\
\mathbb{E} [Z_j Z_k] &= \mathbb{E} [Z_j \mathbb{E} [Z_k | \mathcal{F}_{t_k}]] = 0.
\end{aligned}$$

This yields

$$\begin{aligned}
\text{Var} \left[ \sum_{j=0}^{n-1} Z_j \right] &= 2 \mathbb{E} \left[ \sum_{j=0}^{n-1} f''(W_{t_j})^2 (t_{j+1} - t_j)^2 \right], \\
&\leq 2 \|\Pi\| \mathbb{E} \left[ \sum_{j=0}^{n-1} f''(W_{t_j})^2 (t_{j+1} - t_j) \right].
\end{aligned}$$

As  $\|\Pi\| \downarrow 0$ , the variance goes to 0 since

$$\mathbb{E} \left[ \sum_{j=0}^{n-1} f''(W_{t_j})^2 (t_{j+1} - t_j) \right] \rightarrow \mathbb{E} \left[ \int_0^T f''(W_t)^2 dt \right] < \infty.$$