

Homework 3

Handed out: Monday, September 24, 2018

Due: Wednesday, October 3, 2018 Midnight

Notes:

- We *highly* encourage typed (Latex or Word) homework. Compile as single report containing solutions, derivations, figures, etc.
 - Submit all files including report pdf, report source files (e.g. .tex or .docx files), data, figures produced by computer codes and programs files (e.g. .py or .m files) in a **.zip** folder. Programs should include a Readme file with running instructions.
 - Zipped folder should be turned in on Sakai with the following naming scheme:
HW3_LastName_FirstName.zip
 - Collaboration is allowed however all submitted reports, programs, figures, etc. should be an individual student's write ups. Direct copying could be considered cheating.
 - Software resources for this Homework set can be downloaded from [this link](#) or on Sakai under the Resource folder.
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In this homework, for problems 1 – 7, we will be using the caterpillar regression problem discussed in Chapter 3 of [Bayesian Core](#) by Jean-Michel Marin and Christian Robert¹ using the caterpillar data-set original published by Tomassone in 1993². We suggest reading the entirety of Chapter 3 since this homework involves directly reproducing the results found in the text.

Code Base: For this homework, we will provide the base code for the all the problems found in the links below:

- [\[Matlab\]](#)
- [\[Python 2\]](#)
- [\[R\]](#) from Bayesian Core

The goal is to have an in depth understanding of the algorithms implemented.

Problem 1-5pts

The problem of interest consists of a scalar output y (the logarithmic of the number of caterpillars colonies in a $500m^2$ area) and $k = 10$ explanatory variables x_i ($i = 1, \dots, 10$) which correlate to features such as altitude, landscape slope, vegetation count, etc. The exact meaning does not matter in this context. Using the [caterpillar.mat](#) data, plot the semi-log-y plots of the caterpillar colony (last column of the data file) versus each feature reproducing Figure 3.1 in Bayesian Core.

¹Jean-Michel Marin and Christian Robert. *Bayesian core: a practical approach to computational Bayesian statistics*. Springer Science & Business Media, 2007.

²Richard Tomassone. "Biométrie Modélisation de phénomènes biologiques". In: (1993).

Problem 2-10pts

Reproduce Figure 3.2 in Bayesian Core. For this table you will need calculate four different quantities:

- (a) *Derive* and compute the MLE of the regression coefficients including a bias term. Comment on the dimensionality of \mathbf{X} , $\hat{\beta}$ and \mathbf{y} .
- (b) The standard significance:

$$\hat{\sigma}^2 = \frac{s^2}{n - k - 1}, \quad s^2 = (\mathbf{y} - \mathbf{X}\hat{\beta}) (\mathbf{y} - \mathbf{X}\hat{\beta})^T, \quad (1)$$

which approximates the covariance matrix of $\hat{\beta}$.

- (c) The standard *t-statistic* against null hypothesis:

$$T_i = \frac{\hat{\beta}_i - \beta_i}{\sqrt{\hat{\sigma}^2 \omega_{ii}}} \sim \mathcal{T}(n - k - 1, 0, 1), \quad \omega_{ii} = (\mathbf{X}^T \mathbf{X})^{-1}|_{(i,i)}, \quad (2)$$

where $\hat{\sigma}^2 (\mathbf{X}^T \mathbf{X})^{-1}$ approximates the co-variance matrix of $\hat{\beta}$. For this problem β_i can be taken as 0.

- (d) The p-value:

$$p_i = \Pr(|T_i| > |t_{\text{values}(i)}|), \quad (3)$$

which can be found by adding the left and right values of the CDF of $\mathcal{T}(n - k - 1, 0, 1)$. In the frequentist setting, the *p-value* is commonly used as a probability of the null-hypothesis. However, viewing this idea from a Bayesian perspective what is the logical fallacy frequentists are committing here?

Problem 3-20pts

Reproduce Tables 3.1 and 3.2 in Bayesian Core. For this problem you will need to compute the following marginal statistics of the posterior:

$$\mathbb{E}_{\pi}(\sigma^2 | \mathbf{y}, \mathbf{X}), \quad \mathbb{E}_{\pi}(\beta_i | \mathbf{y}, \mathbf{X}), \quad \text{Var}_{\pi}(\beta_i | \mathbf{y}, \mathbf{X}), \quad (4)$$

for $i = 0, 1, \dots, 10$ where β_0 is the bias term. For this we will use the conjugate priors:

$$p(\beta | \sigma^2, \mathbf{X}) \sim \mathcal{N}(\tilde{\beta}, \sigma^2 \mathbf{M}^{-1}), \quad (5)$$

$$p(\sigma^2 | \mathbf{X}) \sim \mathcal{IG}(a, b), \quad (6)$$

in which $\mathbf{M} = \frac{\mathbf{I}}{c} \in \mathbb{R}^{(k+1) \times (k+1)}$ where c is a constant. Consider the likelihood of:

$$p(\mathbf{y} | \beta, \sigma^2, \mathbf{X}) = \mathcal{N}(\mathbf{X}\beta, \sigma^2), \quad (7)$$

which is the same you used in Problem 2. What are the PDFs for the conditional and marginal posteriors (Can use the reference for help)? Use the hyper-parameter values of $\tilde{\beta}_i = 0$, $k = 10$, $a = 2.1$ and $b = 2.0$ for computing the table values. c will be governed by which table you are reproducing.

Problem 4-20pts

Reproduce Tables 3.3 and 3.4 in Bayesian Core using Zellner's informative G-prior:

$$p(\boldsymbol{\beta}|\sigma^2, \mathbf{X}) \sim \mathcal{N}\left(\tilde{\boldsymbol{\beta}}, c\sigma^2(\mathbf{X}^T \mathbf{X})^{-1}\right), \quad (8)$$

$$p(\sigma^2|\mathbf{X}) \propto \sigma^{-2} \quad (\text{Improper Jeffery's Prior}), \quad (9)$$

where we will again take $\tilde{\beta}_i = 0$ and consider the cases of $c = 100$ and $c = 1000$. *Derive* and compute the following posterior statistics for $\beta_i, i = 1, \dots, 10$:

$$\mathbb{E}_\pi(\boldsymbol{\beta}|\mathbf{X}, \mathbf{y}) = \frac{1}{c+1} \left(\tilde{\boldsymbol{\beta}} + c\hat{\boldsymbol{\beta}} \right) \quad (10)$$

$$\text{Var}_\pi(\boldsymbol{\beta}|\mathbf{y}, \mathbf{X}) = \frac{c}{c+1} \frac{\left(s^2 + \frac{(\tilde{\boldsymbol{\beta}} + \hat{\boldsymbol{\beta}})^T \mathbf{X}^T \mathbf{X} (\tilde{\boldsymbol{\beta}} + \hat{\boldsymbol{\beta}})}{c+1} \right)}{n-2} (\mathbf{X}^T \mathbf{X})^{-1}. \quad (11)$$

What form does the marginal posterior $p(\boldsymbol{\beta}|\mathbf{X}, \mathbf{y})$ take?

Additionally compute the \log_{10} of the Bayes factors for the exclusion of each feature independently:

$$B_{10}^\pi = \frac{f(\mathbf{y}|\mathbf{X})}{f(\mathbf{y}|\mathbf{X}_0, H_0)} = \frac{(c_0 + 1)^{(k+1-q)/2}}{(c + 1)^{(k+1)/2}} \times \left[\frac{\mathbf{y}^T \mathbf{y} - \frac{c_0}{c_0+1} \mathbf{y}^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{y} + \frac{1}{c_0+1} \tilde{\boldsymbol{\beta}}_0^T \mathbf{X}_0^T \mathbf{X}_0 \tilde{\boldsymbol{\beta}}_0 - \frac{2}{c_0+1} \mathbf{y}^T \mathbf{X}_0 \tilde{\boldsymbol{\beta}}_0}{\mathbf{y}^T \mathbf{y} - \frac{c}{c+1} \mathbf{y}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} + \frac{1}{c+1} \tilde{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \tilde{\boldsymbol{\beta}} - \frac{2}{c+1} \mathbf{y}^T \mathbf{X} \tilde{\boldsymbol{\beta}}} \right]^{n/2}, \quad (12)$$

where f denotes a density function and H_0 denotes the null-hypothesis for a specific model variable (You need not to derive this yourself).

Interpret what the Bayes factor is telling you, and *discuss* if you can pull any conclusions regarding certain features. What is the potential pitfall of using an improper prior in this setting?

Problem 5-5pts

Reproduce Table 3.5 by computing the 90% high posterior density (HPD) intervals for $\beta_i, i = 1, \dots, 10$ (Bayesian Core wrongly indicates 95% HPD). Use again Zellner's information G-prior with $c = 100$.

Problem 6-20pts

Reproduce Table 3.6 by now using the uninformed Jeffery's prior:

$$p(\boldsymbol{\beta}, \sigma^2|\mathbf{X}) \propto \sigma^{-2}, \quad (13)$$

$$p(c) = c^{-1} \mathbf{I}_{N^*}(c), \quad (14)$$

where we have assigned the hyper-parameter c with a simple prior. Again using $\tilde{\beta}_i = 0$, *verify* the marginal takes the following form:

$$f(\mathbf{y}|\mathbf{X}) = \sum_{c=1}^{\infty} f(\mathbf{y}|\mathbf{X}, c) c^{-1} \propto \sum_{c=1}^{\infty} c^{-1} (c+1)^{-(k+1)/2} \left[\mathbf{y}^T \mathbf{y} - \frac{c}{c+1} \mathbf{y}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \right]^{-n/2}. \quad (15)$$

What is the advantage of using the uniformed prior?

Problem 7-20pts

Reproduce Table 3.7–3.11 by computing the most likely models using both Zellner's G-prior (with $c = 100$) and Zellner's uninformative G-prior. This is completed through Gibbs sampling for variable selection, a Markov chain Monte Carlo method, the follows the algorithm below:

Algorithm 1 Gibbs Sampler for Variable Selection

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1: Initialization : $\leftarrow \boldsymbol{\gamma}^0 = (\gamma_1^0, \dots, \gamma_p^0)$ 
2: for  $t=1$ ;  $t \leq T$ ;  $t++$  do
3:   Given  $(\gamma_1^{t-1}, \dots, \gamma_p^{t-1})$ 
4:   1.  $\gamma_1^t$  according to  $\pi_1(\gamma_1 | \gamma_2^{t-1}, \dots, \gamma_p^{t-1})$ 
5:   2.  $\gamma_2^t$  according to  $\pi_2(\gamma_2 | \gamma_1^t, \gamma_3^{t-1}, \dots, \gamma_p^{t-1})$ 
6:   3.  $\gamma_3^t$  according to  $\pi_3(\gamma_3 | \gamma_1^t, \gamma_2^t, \gamma_4^{t-1}, \dots, \gamma_p^{t-1})$ 
7:   ...
8:   p.  $\gamma_p^t$  according to  $\pi_p(\gamma_p | \gamma_1^t, \dots, \gamma_{p-1}^t)$ 
9: end for
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where $\gamma_i = \{0, 1\}$ is a binary indicator that variable x_i is included in the model (I.e. $\boldsymbol{\gamma} = \{1, 1, \dots, 1\}$ indicates a *full* model where all model variables are used). What exactly is this algorithm doing? How do we determine if a model parameter is to be turned on or off? Explain the Gibbs sampler and its implementation in your own words.