

Metropolis-Hastings for Beta distribution

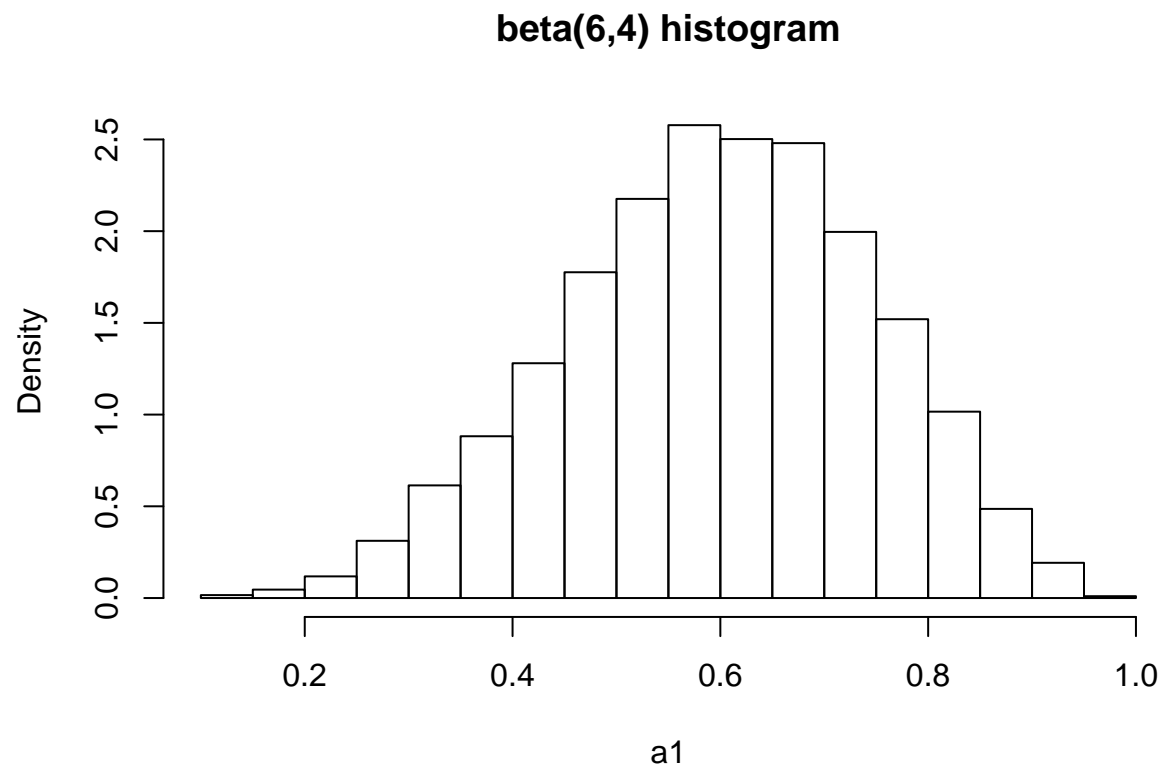
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We write a Metropolis-Hastings algorithm to generate $X \sim \text{beta}(6,4)$. We select a r.v. from $\text{unif}(0,1)$ as a starting value and a $\text{Beta}(c\phi_{old}, c(1 - \phi_{old}))$ jumping distribution where c is a parameter in our function. Later on we can test MH algorithm with different c 's. 10 percent of iterations will be treated as burn-in

```
MHalg<-function(iter,c){  
  
  phi<- c()  
  phi[1]<-runif(1,0,1)#generate a starting value from unif(0,1)  
  
  new_phi<- function(phi) {  
    proposal_phi <- rbeta(1, shape1= c*phi , shape2= c*(1-phi) ) #giving jumping dist.  
    # ratio of accepting proposal_phi  
    r <- (dbeta(proposal_phi, shape1 = 6, shape2 = 4)/  
          dbeta(proposal_phi, shape1= c*phi , shape2=c*(1-phi)))/  
          (dbeta(phi, shape1 = 6, shape2 = 4)/  
           dbeta(phi, shape1= c*proposal_phi, shape2=c*(1-proposal_phi)))  
  
    if (runif(1) <= r)  
      return(proposal_phi)  
    else  
      return(phi)  
  }  
  
  for (i in 1:iter)  
    phi[i+1]= new_phi(phi[i])  
  phi=phi[-(iter/10)] #truncate burn-in  
}
```

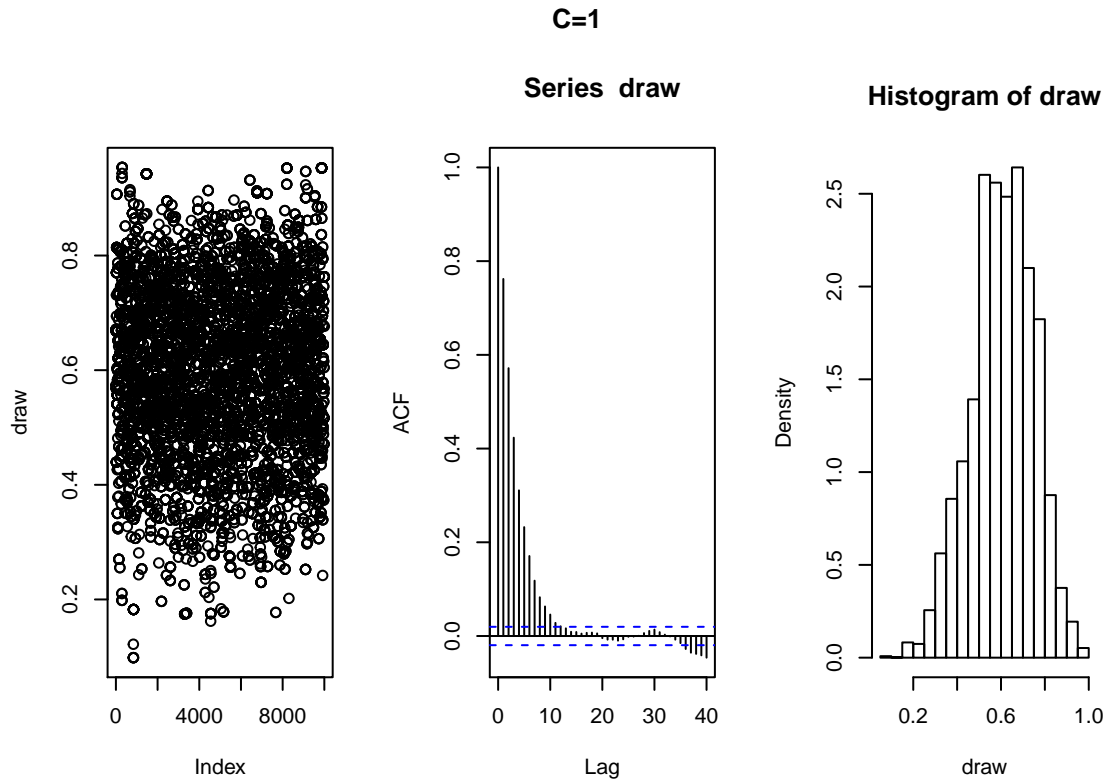
For comparison, generate a histogram with random variables $\sim \text{beta}(6,4)$

```
a1=rbeta(10000,6,4)
hist(a1,freq=FALSE,main="beta(6,4) histogram")
```



We first start a simulation with $c=1$, and 10000 draws. Perform KS test to determine whether our generated random variables are numerically approximately follow $\text{Beta}(6,4)$

```
draw=MHalg(10000,1)
par(mfrow=c(1,3),oma=c(2,2,2,2))
plot(draw); acf(draw); hist(draw,freq=FALSE)
title("C=1", outer=TRUE)
```



```
ks.test(draw,a1)
```

```
## Warning in ks.test(draw, a1): p-value will be approximate in the presence
## of ties
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: draw and a1
## D = 0.0379, p-value = 1.156e-06
## alternative hypothesis: two-sided
```

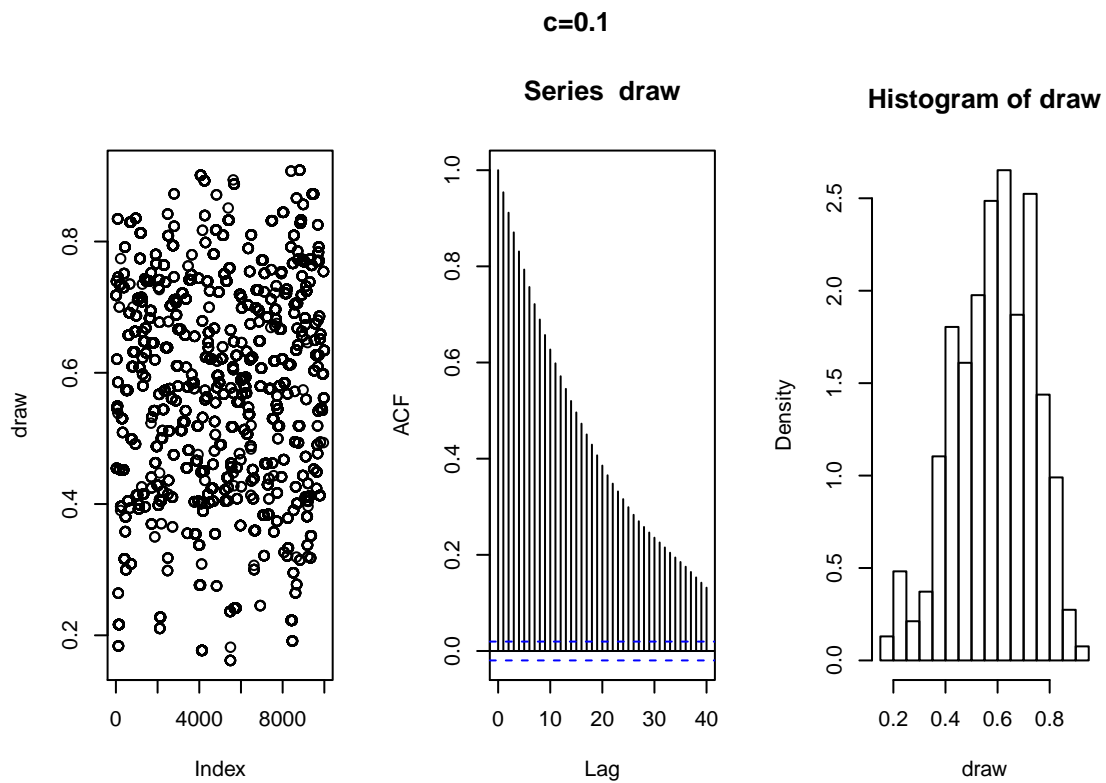
Now we simulate with $c=0.1, 2.5, 10$.

```
draw=MHalg(10000,0.1)
par(mfrow=c(1,3),oma=c(2,2,2,2))
plot(draw); acf(draw); hist(draw,freq=FALSE)
ks.test(draw,a1)
```

```
## Warning in ks.test(draw, a1): p-value will be approximate in the presence
## of ties
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: draw and a1
## D = 0.063, p-value < 2.2e-16
## alternative hypothesis: two-sided
```

```
title("c=0.1", outer=TRUE)
```

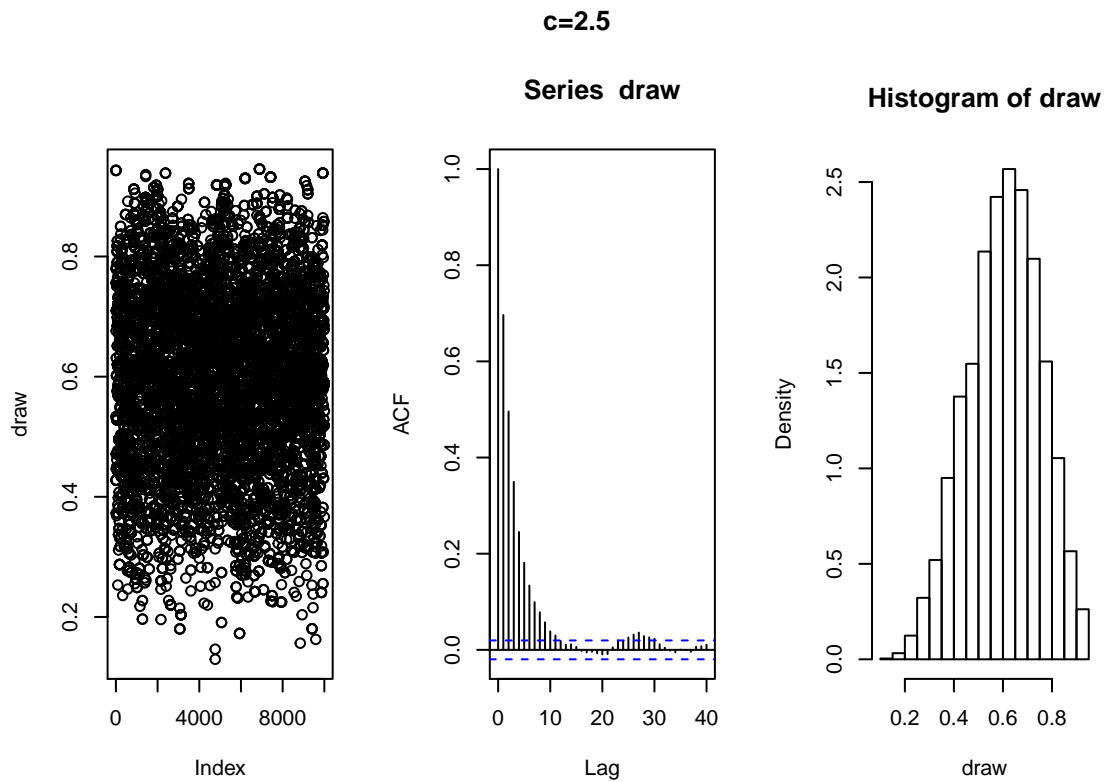


```
draw=MHalg(10000,2.5)
par(mfrow=c(1,3),oma=c(2,2,2,2))
plot(draw); acf(draw); hist(draw,freq=FALSE)
ks.test(draw,a1)
```

```
## Warning in ks.test(draw, a1): p-value will be approximate in the presence
## of ties
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: draw and a1
## D = 0.0217, p-value = 0.01803
## alternative hypothesis: two-sided
```

```
title("c=2.5", outer=TRUE)
```

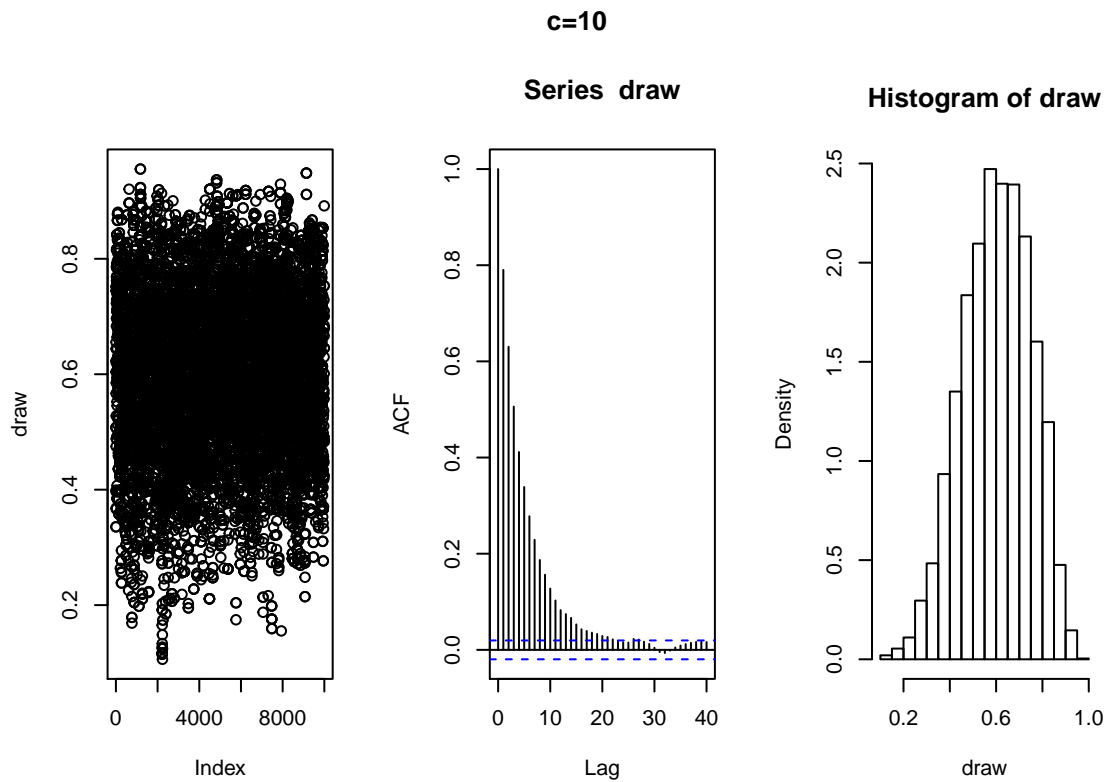


```
draw=MHalg(10000,10)
par(mfrow=c(1,3),oma=c(2,2,2,2))
plot(draw); acf(draw); hist(draw,freq=FALSE)
ks.test(draw,a1)
```

```
## Warning in ks.test(draw, a1): p-value will be approximate in the presence
## of ties
```

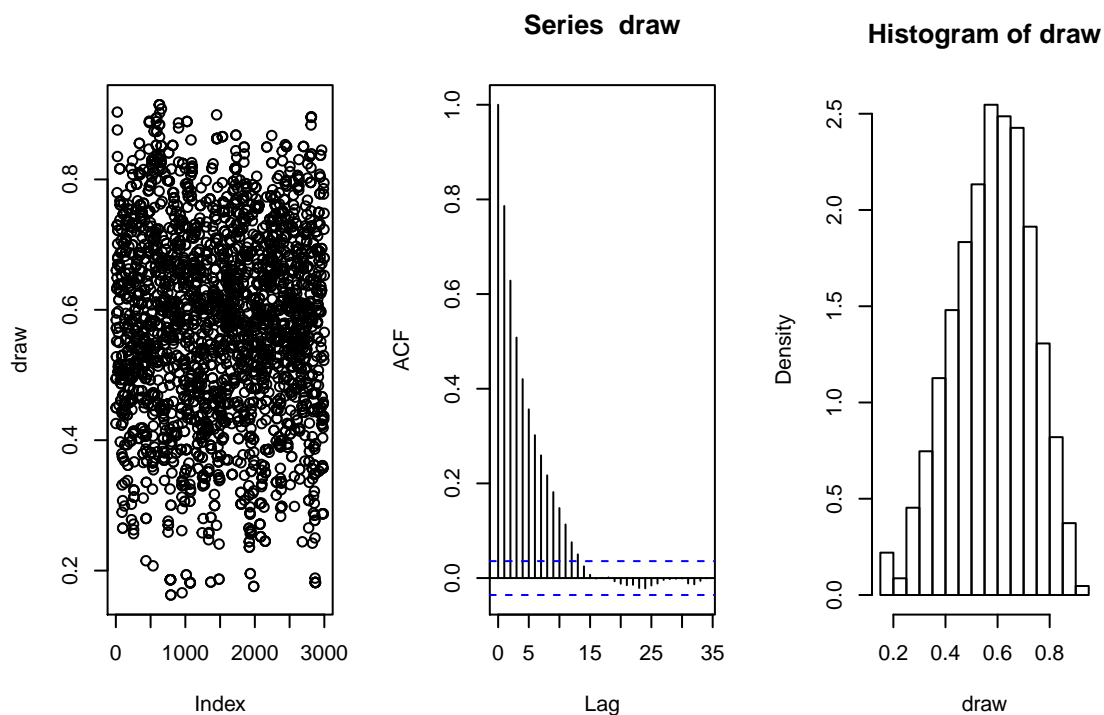
```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: draw and a1
## D = 0.0169, p-value = 0.115
## alternative hypothesis: two-sided
```

```
title("c=10", outer=TRUE)
```



Comparing histograms and Kolmogorov–Smirnov statistic, $c=10$ is the most effective. Now we lower the iteration to 3000. The result is already very close to $\text{beta}(6,4)$

```
draw=MHalg(3000,10)
par(mfrow=c(1,3),oma=c(2,2,2,2))
plot(draw); acf(draw); hist(draw,freq=FALSE)
```



```
a1=rbeta(3000,6,4)
ks.test(draw,a1)
```

```
## Warning in ks.test(draw, a1): p-value will be approximate in the presence
## of ties
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: draw and a1
## D = 0.058, p-value = 8.282e-05
## alternative hypothesis: two-sided
```