

Introduction to Machine Learning (67577)

Exercise 1

Estimation Theory & Mathematical Background

Second Semester, 2023

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1 Submission Instructions

Please make sure to follow the general submission instructions available on the course website. In addition, for the following assignment, submit a single `ex1_ID.tar` file containing:

- An `Answers.pdf` file with the answers for all theoretical and practical questions (include plotted graphs *in* the PDF file).
- The following python files (without any directories): `gaussian_estimators.py`, `fit_gaussian_estimators.py`

The `ex1_ID.tar` file must be submitted in the designated Moodle activity prior to the date specified *in the activity*.

- Late submissions will not be accepted and result in a zero mark.
- Plots included as separate files will be considered as not provided.

2 Theoretical Part

2.1 Mathematical Background

2.1.1 Linear Algebra

Based on Recitation 1

1. Prove that orthogonal matrices are isometric transformations. That is, let $T : V \mapsto W$ be some linear transformation and A the corresponding matrix. Show that if A is an orthogonal matrix then $\forall x \in V \ ||Ax|| = ||x||$.
2. Calculate the SVD of the following matrix A . That is, find the matrices U, Σ, V^\top where U, V are orthogonal matrices and Σ diagonal.

$$A = \begin{bmatrix} 1 & 1 & 0 \\ 1 & -1 & 2 \end{bmatrix}$$

Recall, that to find the SVD of A we can calculate $A^\top A$ to deduce V, Σ and then calculate AA^\top to deduce U . Equivalently, once we deduced V, Σ we can find U using the equality $AV = U\Sigma$.

3. Show that the outer product of two vectors $\mathbf{v} \in \mathbb{R}^n, \mathbf{u} \in \mathbb{R}^m$, which is denoted by $\mathbf{v} \otimes \mathbf{u}$ or $\mathbf{v} \cdot \mathbf{u}^\top$ is a matrix $A \in \mathbb{R}^{n \times m}$ with $\text{rank}(A) = 1$. That is, show that all rows (or columns) in A are linearly dependent.
4. Show that for any orthonormal basis $(\mathbf{u}_1, \dots, \mathbf{u}_n)$ and any arbitrary vector $\mathbf{x} \in \mathbb{R}_n$ such that $\mathbf{x} = \sum_{i=1}^n a_i \cdot \mathbf{u}_i$, it holds that $a_i = \langle \mathbf{x}, \mathbf{u}_i \rangle$ for any $i \in [1, n]$. That is, show that the i 'th coefficient of representing \mathbf{x} in the basis $(\mathbf{u}_1, \dots, \mathbf{u}_n)$, is the inner product between \mathbf{x} and \mathbf{u}_i .

2.1.2 Multivariate Calculus

Based on Recitation 2

5. Let $x \in \mathbb{R}^n$ be a fixed vector and $U \in \mathbb{R}^{n \times n}$ a fixed orthogonal matrix. Calculate the Jacobian of the function $f : \mathbb{R}^n \rightarrow \mathbb{R}^n$:

$$f(\sigma) = U \cdot \text{diag}(\sigma) U^\top x$$

Where $\text{diag}(\sigma)$ is an $n \times n$ matrix where

$$\text{diag}(\sigma)_{ij} = \begin{cases} \sigma_i & i = j \\ 0 & i \neq j \end{cases}$$

6. Use the chain rule to calculate the gradient of $h(\sigma) = \frac{1}{2} \|f(\sigma) - y\|^2$
7. Calculate the Jacobian of the softmax function $S : \mathbb{R}^d \rightarrow [0, 1]^k$

$$S(\mathbf{x})_j = \frac{e^{x_j}}{\sum_{l=1}^k e^{x_l}}$$

8. Let $f : \mathbb{R}^d \rightarrow \mathbb{R}$ be defined as $f(x, y) = x^3 - 5xy - y^5$. Calculate the Hessian of f .

2.1.3 convexity

Based on Recitation 2

9. Prove that the intersection $C := \bigcap_{i \in I} C_i$ for $\{C_i : i \in I\}$ a collection of convex sets is convex.
10. Prove that the vector sum $C_1 + C_2 := \{c_1 + c_2 : c_1 \in C_1, c_2 \in C_2\}$ of two convex sets is convex.
11. Prove that the set $\lambda C := \{\lambda c : c \in C\}$ is convex, for any convex set C , and every scalar λ .

2.2 Estimation Theory

Based on Lecture 1

12. Let $x_1, x_2, \dots \stackrel{iid}{\sim} \mathcal{P}$ be a sample of infinity size drawn from some probability distribution function \mathcal{P} with finite expectation and variance. Show that the sample mean estimator $\hat{\mu}_n = \frac{1}{n} \sum x_i$ calculated over the first n samples is a consistent estimator. Hint: for any given fixed value of $n \in \mathbb{N}$ bound from above the probability of deviating more than ε .
13. Let $\mathbf{x}_1, \dots, \mathbf{x}_m \stackrel{iid}{\sim} \mathcal{N}(\mu, \Sigma)$ be m observations sampled i.i.d from a multivariate Gaussian with expectation of $\mu \in \mathbb{R}^d$ and a covariance matrix $\Sigma \in \mathbb{R}^{d \times d}$. Provide an expression for the log-likelihood function of $\mathcal{N}(\mu, \Sigma)$. Develop the expression as much as you can. Hint: follow the approach used to derive the likelihood function for the univariate case.

3 Practical Part

Before starting the practical part please make sure to have cloned/downloaded the IML.HUJI GitHub repository and set up a working virtual environment. Write the necessary code in the files specified in the questions.

3.1 Univariate Gaussian Estimation

Based on lecture 1

Implement the `UnivariateGaussian` class in the `learners.gaussian_estimators.py` file. Follow details specified in class and function documentation.

1. Using `numpy.random.normal` draw 1000 samples $x_1, \dots, x_{1000} \stackrel{iid}{\sim} \mathcal{N}(10, 1)$ and fit a univariate Gaussian. Print the estimated expectation and variance. Output format should be `(expectation, variance)`.
2. Over previously drawn samples, fit a series of models of increasing samples size: 10, 20, ..., 100, 110, ..., 1000. Plot the absolute distance between the estimated- and true value of the expectation, as a function of the sample size. Provide meaningful axis names and title.
3. Compute the PDF of the previously drawn samples using the model fitted in question 1. Plot the empirical PDF function under the fitted model. That is, create a scatter plot with the ordered sample values along the x-axis and their PDFs (using the `UnivariateGaussian.pdf` function) along the y-axis. Provide meaningful axis names and title. What are you expecting to see in the plot?

3.2 Multivariate Gaussian Estimation

Based on Lecture 1

Implement the `Multivariate` class in the `learners.gaussian_estimators.py` file. Follow details specified in class and function documentation.

NOTICE: When implementing the `log_likelihood` function you are required to use the expression developed in the q13 above. That is, the expression for $\ell(\mu, \Sigma | \mathbf{x}_1, \dots, \mathbf{x}_m)$.

4. Using `numpy.random.multivariate_normal` draw 1000 samples $\mathbf{x}_1, \dots, \mathbf{x}_{1000} \stackrel{iid}{\sim} \mathcal{N}(\mu, \Sigma)$

$$\mu = \begin{bmatrix} 0 \\ 0 \\ 4 \\ 0 \end{bmatrix}, \quad \Sigma = \begin{bmatrix} 1 & 0.2 & 0 & 0.5 \\ 0.2 & 2 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0.5 & 0 & 0 & 1 \end{bmatrix}$$

Fit a multivariate Gaussian and print the estimated expectation and covariance matrix. Print each in a separate line.

5. Using the samples drawn in the question above calculate the log-likelihood for models with expectation $\mu = [f_1, 0, f_3, 0]^\top$ and the true covariance matrix defined above, where f_1, f_3 get values returned from `np.linspace(-10, 10, 200)`. Plot a heatmap of f_1 values as rows, f_3 values as columns and the color being the calculated log likelihood. Provide meaningful axis names and title. What are you able to learn from the plot?
6. Of all values tested in question 5, which model (pair of values for feature 1 and 3) achieved the maximum log-likelihood value? Round to 3 decimal places

2.1 Mathematical Background

2.1.1 Linear Algebra

1. Prove that orthogonal matrices are isometric transformations. That is, let $T : V \mapsto W$ be some linear transformation and A the corresponding matrix. Show that if A is an orthogonal matrix then $\forall x \in V \ ||Ax|| = ||x||$.

$$\|Ax\|^2 = (Ax)^t Ax = \overset{\text{transpose}}{x^t} \overset{\text{orthogonal } A}{A^t A} x = \overset{A^t A = I_n}{x^t} x = \|x\|^2$$

2. Calculate the SVD of the following matrix A . That is, find the matrices U, Σ, V^T where U, V are orthogonal matrices and Σ diagonal.

$$A = \begin{bmatrix} 1 & 1 & 0 \\ 1 & -1 & 2 \end{bmatrix}$$

Recall, that to find the SVD of A we can calculate $A^T A$ to deduce V, Σ and then calculate AA^T to deduce U . Equivalently, once we deduced V, Σ we can find U using the equality $AV = U\Sigma$.

בכדי למצוא את ה-SVD של A , נמצא את EVD של $A^t A$ (2)

$$A^t A = \begin{pmatrix} 1 & 1 \\ 1 & -1 \\ 0 & 2 \end{pmatrix} \begin{pmatrix} 1 & 1 & 0 \\ 1 & -1 & 2 \end{pmatrix} = \begin{pmatrix} 2 & 0 & 2 \\ 0 & 2 & -2 \\ 2 & -2 & 4 \end{pmatrix}$$

נמצא את הערכים העצמיים של $A^t A$ ואת הווקטורים העצמיים שלה

$$\det(\lambda I_n - A^t A) = \det \begin{pmatrix} \lambda-2 & 0 & -2 \\ 0 & \lambda-2 & 2 \\ -2 & 2 & \lambda-4 \end{pmatrix} = (\lambda-2) \begin{vmatrix} \lambda-2 & 2 \\ 2 & \lambda-4 \end{vmatrix} - 2 \begin{vmatrix} 0 & -2 \\ \lambda-2 & 2 \end{vmatrix} = (\lambda-2) ((\lambda-2)(\lambda-4) - 4)$$

$$\underbrace{-2(0 - (-2)(\lambda-2))}_{-4(\lambda-2)} = (\lambda-2) \underbrace{(\lambda^2 - 6\lambda)}_{\text{הפולינום העצמי של } A^t A} = (\lambda-2)\lambda(\lambda-6) \Rightarrow \lambda: 0, 6 \text{ הם הערכים העצמיים}$$

נמצא את הווקטורים העצמיים של $A^t A$ עבור $\lambda = 0, 6$ ואת U ו- Σ (2)

$$\begin{pmatrix} 0 & 0 & -2 \\ 0 & 0 & 2 \\ -2 & 2 & -2 \end{pmatrix} \rightarrow \begin{pmatrix} 1 & -1 & 1 \\ 0 & 0 & 2 \\ 0 & 0 & -2 \end{pmatrix} \rightarrow \begin{pmatrix} 1 & -1 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & -1 \end{pmatrix} \rightarrow \begin{pmatrix} 1 & -1 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 0 \end{pmatrix} \quad \lambda=2 \text{ נמצא}$$

$$\text{ל } \lambda=0 \text{ נמצא } \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \text{ ו- } V_2 = \left\{ \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} t \mid t \in \mathbb{R} \right\} \text{ שם } x_1 = t, x_2 = 0, x_3 = t$$

$$\begin{pmatrix} -2 & 0 & -2 \\ 0 & -2 & 2 \\ -2 & 2 & -4 \end{pmatrix} \longrightarrow \begin{pmatrix} 1 & 0 & 1 \\ 0 & 1 & -1 \\ 1 & -1 & 2 \end{pmatrix} \longrightarrow \begin{pmatrix} 1 & 0 & 1 \\ 0 & 1 & -1 \\ 0 & -1 & 1 \end{pmatrix} \longrightarrow \begin{pmatrix} 1 & 0 & 1 \\ 0 & 1 & -1 \\ 0 & 0 & 0 \end{pmatrix}$$

0 לא \vec{x} $\begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix}$ נבחר $x_2=t, x_1=-t, x_3=t$ שם

$$\begin{pmatrix} 4 & 0 & -2 \\ 0 & 4 & 2 \\ -2 & 2 & 2 \end{pmatrix} \longrightarrow \begin{pmatrix} 2 & 0 & -1 \\ 0 & 2 & 1 \\ -1 & 1 & 1 \end{pmatrix} \longrightarrow \begin{pmatrix} 0 & 2 & 1 \\ 0 & 0 & 0 \\ -1 & 1 & 1 \end{pmatrix} \longrightarrow \begin{pmatrix} 1 & -1 & -1 \\ 0 & 1 & \frac{1}{2} \\ 0 & 0 & 0 \end{pmatrix}$$

נקודות 6:

6 לא $\begin{pmatrix} 1 \\ -1 \\ 2 \end{pmatrix}$ נבחר $x_3=t, x_2=-\frac{1}{2}t, x_1=\frac{1}{2}t$

סה"כ מצאנו $\begin{pmatrix} 1 \\ -1 \\ 2 \end{pmatrix}, \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix}$ לא ה"ע 0, 2, 6 בהתאמה. מכיון ש- $A^t A$ מטריצה סימטרית, לא \vec{x} לא

שנני נבדוק זכר הקטורים הללו מנורמלים ונחזיר בסים אורתונורמליים ב- \mathbb{R}^3 .

$$(v_1, v_2, v_3) = \left(\begin{pmatrix} 1 \\ -1 \\ 2 \end{pmatrix} \cdot \frac{1}{\sqrt{6}}, \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \cdot \frac{1}{\sqrt{2}}, \begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix} \cdot \frac{1}{\sqrt{3}} \right) \Leftarrow \left\| \begin{pmatrix} 1 \\ -1 \\ 2 \end{pmatrix} \right\| = \sqrt{1+1+4} = \sqrt{6}, \quad \left\| \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \right\| = \sqrt{2}, \quad \left\| \begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix} \right\| = \sqrt{3}$$

בסיס אורתונורמלי ב- \mathbb{R}^3

$$A^t A = \begin{pmatrix} 1 & 1 & 1 \\ v_1 & v_2 & v_3 \\ 1 & 1 & 1 \end{pmatrix} \begin{pmatrix} 6 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ v_1 & v_2 & v_3 \\ 1 & 1 & 1 \end{pmatrix}^t$$

שם מתקין V^T

$$V^T = \begin{pmatrix} 1/\sqrt{6} & -1/\sqrt{6} & 2/\sqrt{6} \\ 1/\sqrt{2} & 1/\sqrt{2} & 0 \\ -1/\sqrt{3} & 1/\sqrt{3} & 1/\sqrt{3} \end{pmatrix}$$

כך נס'ק ל- V^T , והערות הס'פול"פ הם הס'פול"פ של ה"ע לא

$$A^t A \text{ שם } 0, \text{ כלומר } \sqrt{6}, \sqrt{2}, \text{ ולכן } \Sigma = \begin{pmatrix} \sqrt{6} & 0 & 0 \\ 0 & \sqrt{2} & 0 \end{pmatrix} \quad (\Sigma \text{ היא קאדן מנורמל כן } A)$$

נמצא AA^t שם AA^t הקטורים הס'פול"פ (השאלה פ-)

$$AA^t = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 2 \\ 1 & -1 & 2 \end{pmatrix} \begin{pmatrix} 1 & 1 \\ 1 & -1 \\ 0 & 2 \end{pmatrix} = \begin{pmatrix} 2 & 0 \\ 0 & 6 \end{pmatrix}$$

נבין ש- (e_1, e_2) בסיס אורתונורמלי של \mathbb{R}^2 (בס'פול"פ לא ה"ע) מכיון ש- $AA^t e_1 = 2e_1, AA^t e_2 = 6e_2$, $e_1 \perp e_2, \|e_1\| = \|e_2\| = 1$. לכן $V = (e_2, e_1) = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$

$$A = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \begin{pmatrix} \sqrt{6} & 0 & 0 \\ 0 & \sqrt{2} & 0 \end{pmatrix} \begin{pmatrix} 1/\sqrt{6} & -1/\sqrt{6} & 2/\sqrt{6} \\ 1/\sqrt{2} & 1/\sqrt{2} & 0 \\ -1/\sqrt{3} & 1/\sqrt{3} & 1/\sqrt{3} \end{pmatrix} : \text{שם}$$

$V \quad \Sigma \quad V^T$

3. Show that the outer product of two vectors $\mathbf{v} \in \mathbb{R}^n, \mathbf{u} \in \mathbb{R}^m$, which is denoted by $\mathbf{v} \otimes \mathbf{u}$ or $\mathbf{v} \cdot \mathbf{u}^\top$ is a matrix $A \in \mathbb{R}^{n \times m}$ with $\text{rank}(A) = 1$. That is, show that all rows (or columns) in A are linearly dependent.

Definition 1.4 For two vectors $\mathbf{v} \in \mathbb{R}^n$, $\mathbf{u} \in \mathbb{R}^m$, the *outer product* of \mathbf{v} and \mathbf{u} , which is denoted by $\mathbf{v} \otimes \mathbf{u}$ or $\mathbf{v}\mathbf{u}^\top$, is an $n \times m$ matrix with entries:

\uparrow \downarrow \downarrow
 $\mathbb{N} \times \mathbb{N} \Rightarrow \mathbb{N} \times \mathbb{N}$
 $\mathbb{N} \times \mathbb{N}$

$$[\mathbf{v} \otimes \mathbf{u}]_{ij} = v_i \cdot u_j, \quad \mathbf{v} \otimes \mathbf{u} = \begin{bmatrix} v_1 u_1 & v_1 u_2 & \cdots & v_1 u_m \\ \vdots & \vdots & \ddots & \vdots \\ v_n u_1 & v_n u_2 & \cdots & v_n u_m \end{bmatrix} = \mathbf{A}$$

פ.ש. A של $n \times n$ מטריצה, $A = \begin{pmatrix} v_1 \cdot u^t \\ v_2 \cdot u^t \\ \vdots \\ v_n \cdot u^t \end{pmatrix}$ v_i וקטורים, u^t וקטור, $v_i \cdot u^t = 0$ $\forall i$ $\Rightarrow A = 0$ $\Rightarrow \text{rank}(A) = 0$ $\Rightarrow A$ אינו הפיך.

4. Show that for any orthonormal basis $(\mathbf{u}_1, \dots, \mathbf{u}_n)$ and any arbitrary vector $\mathbf{x} \in \mathbb{R}_n$ such that $\mathbf{x} = \sum_{i=1}^n a_i \cdot \mathbf{u}_i$, it holds that $a_i = \langle \mathbf{x}, \mathbf{u}_i \rangle$ for any $i \in [1, n]$. That is, show that the i 'th coefficient of representing \mathbf{x} in the basis $(\mathbf{u}_1, \dots, \mathbf{u}_n)$, is the inner product between \mathbf{x} and \mathbf{u}_i .

$$\langle x, u_i \rangle = \langle \sum_{j=1}^n a_j u_j, u_i \rangle = \sum_{j=1}^n a_j \langle u_j, u_i \rangle = a_i$$

הבס"ס
אחיותי ואני
סמור הזנאי
המנקה הבתים שזהו כי הוא צב"ס ואפ"כ נ"ו המנקה הבתים שזהו כי וי מנקה

Definition 2.4 Let $\mathbf{f} : \mathbb{R}^d \rightarrow \mathbb{R}^m$ where $\mathbf{f}(\mathbf{x}) = (f_1(\mathbf{x}), \dots, f_m(\mathbf{x}))^\top$. The **Jacobian** of f is the $m \times d$ matrix of all partial derivatives:

$$J_{\mathbf{x}}(\mathbf{f}) := \begin{bmatrix} \frac{\partial f_1(\mathbf{x})}{\partial x_1} & \cdots & \frac{\partial f_1(\mathbf{x})}{\partial x_d} \\ \vdots & & \vdots \\ \frac{\partial f_m(\mathbf{x})}{\partial x_1} & \cdots & \frac{\partial f_m(\mathbf{x})}{\partial x_d} \end{bmatrix}$$

2.1.2 Multivariate Calculus

Based on Recitation 2

5. Let $x \in \mathbb{R}^n$ be a fixed vector and $U \in \mathbb{R}^{n \times n}$ a fixed orthogonal matrix. Calculate the Jacobian of the function $f: \mathbb{R}^n \rightarrow \mathbb{R}^n$:

$$f(\sigma) = U \cdot \text{diag}(\sigma) U^T x$$

$\begin{matrix} n \times n & n \times n & n \times n & n \times n \end{matrix}$

Where $\text{diag}(\sigma)$ is an $n \times n$ matrix where

$$\text{diag}(\sigma)_{ij} = \begin{cases} \sigma_i & i = j \\ 0 & i \neq j \end{cases}$$

$$U = \begin{pmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{pmatrix} \quad \text{no } (5)$$

$$\begin{aligned} f(\sigma) &= U \text{diag}(\sigma) U^T x = U \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_n \end{pmatrix} \begin{pmatrix} u_1^T x \\ \vdots \\ u_n^T x \end{pmatrix} = U \begin{pmatrix} \sigma_1 u_1^T x \\ \vdots \\ \sigma_n u_n^T x \end{pmatrix} = \sum_{i=1}^n \underbrace{(\sigma_i u_i^T x)}_{1 \times n \cdot n \times n} u_i = \\ &= \sum_{i=1}^n u_i (u_i^T \sigma_i x) = \sum_{i=1}^n \sigma_i u_i u_i^T x \end{aligned}$$

$$[J_\sigma(f)]_{ij} = \frac{\partial f_j(\sigma)}{\partial \sigma_i} = [u_i u_i^T x]_j \quad \Rightarrow \quad J_\sigma(f) = \begin{pmatrix} | & & | \\ u_1 & \dots & u_n \\ | & & | \end{pmatrix} \begin{pmatrix} u_1^T x & 0 \\ \vdots & \ddots \\ 0 & u_n^T x \end{pmatrix} = U \text{diag}(U^T x)$$

$u_i u_i^T x \text{ se } j\text{-th element}$

6. Use the chain rule to calculate the gradient of $h(\sigma) = \frac{1}{2} \|f(\sigma) - y\|^2$

Definition 2.3 Let $f: \mathbb{R}^d \rightarrow \mathbb{R}$. The gradient of f at \mathbf{x} is the vector of partial derivatives:

$$\nabla f(\mathbf{x}) := \left(\frac{\partial f(\mathbf{x})}{\partial x_1}, \dots, \frac{\partial f(\mathbf{x})}{\partial x_d} \right)^T$$

(6)

$$h(\sigma) = \frac{1}{2} \|f(\sigma) - y\|^2 = \frac{1}{2} \langle f(\sigma) - y, f(\sigma) - y \rangle = \frac{1}{2} (\|f(\sigma)\|^2 + \|y\|^2 - 2\langle f(\sigma), y \rangle) = \frac{1}{2} \|f(\sigma)\|^2 + \frac{1}{2} \|y\|^2 - \langle f(\sigma), y \rangle =$$

$$= \frac{1}{2} \|f(\sigma)\|^2 - f(\sigma)^T y + \frac{1}{2} \|y\|^2$$

$$\nabla h(\sigma) = \frac{\partial h(\sigma)}{\partial \sigma} = \frac{\partial h(\sigma)}{\partial f(\sigma)} \frac{\partial f(\sigma)}{\partial \sigma} = \frac{\partial \left(\frac{1}{2} \|f(\sigma)\|^2 - f(\sigma)^T y + \frac{1}{2} \|y\|^2 \right)}{\partial f(\sigma)} \cdot \frac{\partial f(\sigma)}{\partial \sigma} = (f(\sigma) - y)^T \cdot \frac{\partial f(\sigma)}{\partial \sigma} =$$

$$= \underbrace{(f(\sigma) - y)^T}_{1 \times n} \underbrace{J_{\sigma}(f)}_{n \times h} = \underbrace{(f(\sigma) - y)^T}_{1 \times n} \underbrace{U \text{diag}(U^T X)}_{n \times h}$$

$$\text{for } \mathbb{R}^n \text{ and } \mathbb{R}^h \text{ vectors } x, y \quad \frac{\partial x^T y}{\partial x_i} = \frac{\sum_j x_j y_j}{\partial x_i} = y_i \Rightarrow \frac{\partial x^T y}{\partial x} = y, \quad \frac{\partial x^T x}{\partial x_i} = \frac{\sum_j x_j^2}{\partial x_i} = 2x_i \Rightarrow \frac{\partial \|x\|^2}{\partial x} = 2x$$

7. Calculate the Jacobian of the softmax function $S: \mathbb{R}^d \rightarrow [0, 1]^k$

$$S(\mathbf{x})_j = \frac{e^{x_j}}{\sum_{l=1}^k e^{x_l}}$$

$$\left(\frac{f}{g}\right)' = \frac{f'g - fg'}{g^2} \quad \text{חוק ג' (7) : חוק ג'}$$

$$[J_K(s)]_{ij} = \frac{\partial s(x)_i}{\partial x_j} = \frac{a}{\partial x_j} \frac{e^{x_i}}{\sum_{l=1}^K e^{x_l}} = \begin{cases} \frac{e^{x_i} \sum_{l=1}^K e^{x_l} - e^{x_i} \cdot e^{x_j}}{\left(\sum_{l=1}^K e^{x_l}\right)^2} = \frac{e^{x_i}}{\sum_{l=1}^K e^{x_l}} \cdot \frac{\left(\sum_{l=1}^K e^{x_l} - e^{x_j}\right)}{\sum_{l=1}^K e^{x_l}} = s(x)_i \cdot (1 - s(x)_j) & i=j \\ \frac{0 \cdot \sum_{l=1}^K e^{x_l} - e^{x_i} \cdot e^{x_j}}{\left(\sum_{l=1}^K e^{x_l}\right)^2} = -\frac{e^{x_i}}{\sum_{l=1}^K e^{x_l}} \cdot \frac{e^{x_j}}{\sum_{l=1}^K e^{x_l}} = -s(x)_i \cdot s(x)_j = s(x)_i \cdot (-s(x)_j) & i \neq j \end{cases}$$

$$\delta_{ij} = \begin{cases} 1 & i=j \\ 0 & i \neq j \end{cases} \quad \text{אז} \quad [J_K(f)]_{ij} = s(x)_i \cdot (\delta_{ij} - s(x)_j) \quad \text{כאן}$$

$$[\text{diag}(s)]_{ij} = \begin{cases} s(x)_i & i=j \\ 0 & i \neq j \end{cases} \quad \text{אז} \quad [SS^T]_{ij} = s(x)_i \cdot s(x)_j \quad \text{לכן} \quad J_X(f) = \text{diag}(s) - SS^T \quad \text{המטריצה}$$

$$0 - s(x)_i s(x)_j = -s(x)_i s(x)_j, \quad s(x)_i - s(x)_i s(x)_j = s(x)_i (1 - s(x)_j) \quad \text{אז}$$

8. Let $f: \mathbb{R}^d \rightarrow \mathbb{R}$ be defined as $f(x, y) = x^3 - 5xy - y^5$. Calculate the Hessian of f .

$$\frac{\partial f}{\partial x} = 3x^2 - 5y$$

$$\frac{\partial f}{\partial y} = -5x - 5y^4$$

$$\frac{\partial f}{\partial x \partial y} = -5$$

$$\frac{\partial^2 f}{\partial x^2} = 6x$$

$$\frac{\partial^2 f}{\partial y^2} = -20y^3$$

$$\frac{\partial^2 f}{\partial y \partial x} = -5$$

אז המטריצה

$$H(f)_{(x,y)} = \begin{pmatrix} \frac{\partial^2 f}{\partial x^2} & \frac{\partial^2 f}{\partial x \partial y} \\ \frac{\partial^2 f}{\partial y \partial x} & \frac{\partial^2 f}{\partial y^2} \end{pmatrix} = \begin{pmatrix} 6x & -5 \\ -5 & -20y^3 \end{pmatrix}$$

2.1.3 convexity

Based on Recitation 2

9. Prove that the intersection $C := \bigcap_{i \in I} C_i$ for $\{C_i : i \in I\}$ a collection of convex sets is convex.
10. Prove that the vector sum $C_1 + C_2 := \{c_1 + c_2 : c_1 \in C_1, c_2 \in C_2\}$ of two convex sets is convex.
11. Prove that the set $\lambda C := \{\lambda c : c \in C\}$ is convex, for any convex set C , and every scalar λ .

9. יהי $u, v \in \bigcap_{i \in I} C_i$ ונראה ש $\alpha u + (1-\alpha)v \in C$ לכל $\alpha \in [0,1]$.

מכיוון ש $\bigcap_{i \in I} C_i \subseteq C_i$ לכל $i \in I$ ו $u, v \in C_i$ לכל $i \in I$ נקבל ש $\alpha u + (1-\alpha)v \in C_i$ לכל $i \in I$ ולכן $\alpha u + (1-\alpha)v \in \bigcap_{i \in I} C_i = C$.

10. נניח $u, v \in C_1 + C_2$ ונראה ש $\alpha u + (1-\alpha)v \in C_1 + C_2$.
 נכתוב $u = c_1 + c_2$ ו $v = c_3 + c_4$ עבור $c_1, c_3 \in C_1$ ו $c_2, c_4 \in C_2$.
 אז $\alpha u + (1-\alpha)v = \alpha(c_1 + c_2) + (1-\alpha)(c_3 + c_4) = (\alpha c_1 + (1-\alpha)c_3) + (\alpha c_2 + (1-\alpha)c_4)$.
 מכיוון ש C_1 ו C_2 קונקסים, נקבל ש $\alpha c_1 + (1-\alpha)c_3 \in C_1$ ו $\alpha c_2 + (1-\alpha)c_4 \in C_2$.
 לכן $\alpha u + (1-\alpha)v \in C_1 + C_2$.

$$\alpha u + (1-\alpha)v = \alpha(c_1 + c_2) + (1-\alpha)(c_3 + c_4) = (\alpha c_1 + (1-\alpha)c_3) + (\alpha c_2 + (1-\alpha)c_4)$$

נראה ש $\alpha c_1 + (1-\alpha)c_3 \in C_1$ ו $\alpha c_2 + (1-\alpha)c_4 \in C_2$.
 מכיוון ש C_1 קונקסי, נקבל ש $\alpha c_1 + (1-\alpha)c_3 \in C_1$.
 מכיוון ש C_2 קונקסי, נקבל ש $\alpha c_2 + (1-\alpha)c_4 \in C_2$.

11. יהי $u, v \in \lambda C$ ונראה ש $\alpha u + (1-\alpha)v \in \lambda C$.
 נכתוב $u = \lambda c_1$ ו $v = \lambda c_2$ עבור $c_1, c_2 \in C$.
 אז $\alpha u + (1-\alpha)v = \alpha \lambda c_1 + (1-\alpha) \lambda c_2 = \lambda (\alpha c_1 + (1-\alpha)c_2)$.
 מכיוון ש C קונקסי, נקבל ש $\alpha c_1 + (1-\alpha)c_2 \in C$.
 לכן $\alpha u + (1-\alpha)v \in \lambda C$.

$$\alpha u + (1-\alpha)v = \alpha \lambda c_1 + (1-\alpha) \lambda c_2 = \lambda (\alpha c_1 + (1-\alpha)c_2) \Rightarrow \alpha u + (1-\alpha)v \in \lambda C$$

המשפט הראשון המשפט השני

2.2 Estimation Theory

Based on Lecture 1

12. Let $x_1, x_2, \dots \stackrel{iid}{\sim} \mathcal{P}$ be a sample of infinity size drawn from some probability distribution function \mathcal{P} with **finite expectation** and **variance**. Show that the sample mean estimator $\hat{\mu}_n = \frac{1}{n} \sum x_i$ calculated over the first n samples is a **consistent estimator**. Hint: for any given fixed value of $n \in \mathbb{N}$ bound from above the probability of deviating more than ε .
13. Let $\mathbf{x}_1, \dots, \mathbf{x}_m \stackrel{iid}{\sim} \mathcal{N}(\mu, \Sigma)$ be m observations sampled i.i.d from a multivariate Gaussian with expectation of $\mu \in \mathbb{R}^d$ and a covariance matrix $\Sigma \in \mathbb{R}^{d \times d}$. Provide an expression for the log-likelihood function of $\mathcal{N}(\mu, \Sigma)$. Develop the expression as much as you can. Hint: follow the approach used to derive the likelihood function for the univariate case.

(12)

נסו μ בתוחלת של $X \sim \mathcal{P}$ שם $\varepsilon > 0$

$$P(|\hat{\mu}_n - \mu| > \varepsilon) \leq P(|\mu_n - E[\mu_n]| \geq \varepsilon) \leq \frac{\text{var}[\mu_n]}{\varepsilon^2} = \frac{\text{var}[x]}{n \varepsilon^2} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \mu_n \text{ מתכנסת תמידית ל} \mu$$

$E[\frac{1}{n} \sum_{i=1}^n x_i] = \frac{1}{n} \cdot n E[x_i] = E[x]$ $\mu_n = \frac{1}{n} \sum_{i=1}^n x_i$ $\text{var}[\frac{1}{n} \sum_{i=1}^n x_i] = \frac{1}{n^2} \sum_{i=1}^n \text{var}[x_i] = \frac{\text{var}[x]}{n}$

\downarrow \downarrow \downarrow

$X \sim \mathcal{P}$ $X \sim \mathcal{P}$ $X \sim \mathcal{P}$

$$f(X) = \frac{1}{\sqrt{(2\pi)^d |\Sigma|}} \exp \left\{ -\frac{1}{2} (X - \mu)^T \Sigma^{-1} (X - \mu) \right\}$$

$X = \begin{pmatrix} | & & | \\ x_1 & \dots & x_m \\ | & & | \end{pmatrix} \in M_{d \times m}$ (13)

$$L(\mu, \Sigma | x_1, \dots, x_m) = f_{\mu}(x_1, \dots, x_m) = \prod_{i=1}^m f_{\mu}(x_i) = \prod_{i=1}^m \frac{1}{\sqrt{(2\pi)^d |\Sigma|}} \exp \left(-\frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right)$$

$$\downarrow$$

$$e^{x \cdot y} = e^x \cdot e^y$$

$$= (2\pi)^d |\Sigma|^{-\frac{m}{2}} \cdot \exp \left(\sum_{i=1}^m \left(-\frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right) \right) = (2\pi)^d |\Sigma|^{-\frac{m}{2}} \cdot \exp \left(-\frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right)$$

טבת $\log L(\mu, \Sigma | x_1, \dots, x_m)$ את

$$\log \left((2\pi)^d |\Sigma|^{-\frac{m}{2}} \cdot \exp \left(-\frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right) \right) = -\frac{m}{2} \log((2\pi)^d |\Sigma|) - \frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

$\log x \cdot y = \log x + \log y$
 $\log a^r = r \log a$
 $\log(e^x) = x$

$$= -\frac{m}{2} \left(\log((2\pi)^d) + \log(|\Sigma|) \right) - \frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

$\log x \cdot y = \log x + \log y$
 $\log a^r = r \log a$

$$= -\frac{m}{2} \cdot d \log(2\pi) - \frac{m}{2} \log(|\Sigma|) - \frac{1}{2} \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) = -\frac{1}{2} \left(m(d \log(2\pi) + \log(|\Sigma|)) + \sum_{i=1}^m (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right)$$

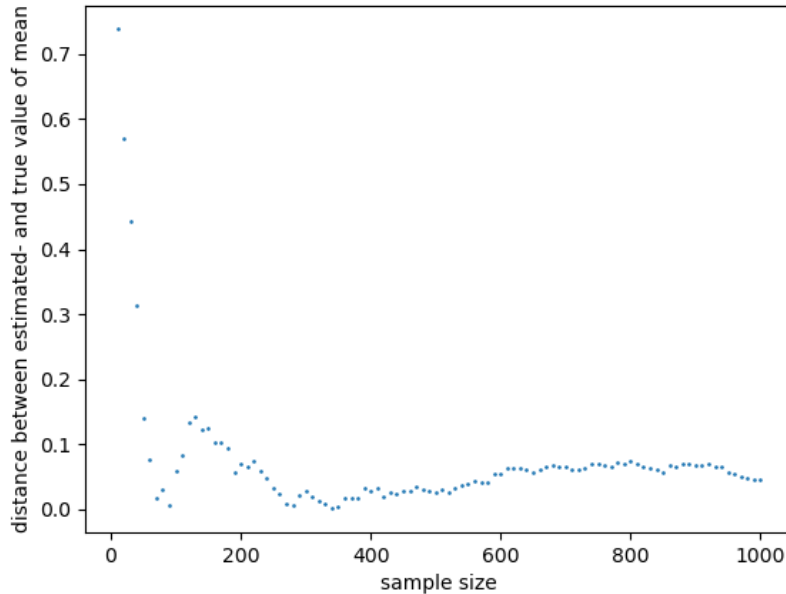
$\log a^r = r \log a$

(9.955, 0.975)

(1)

(2)

Question 2- deviation of estimated expectation from true expectation
as function of sample size for samples from $N(10,1)$ distribution

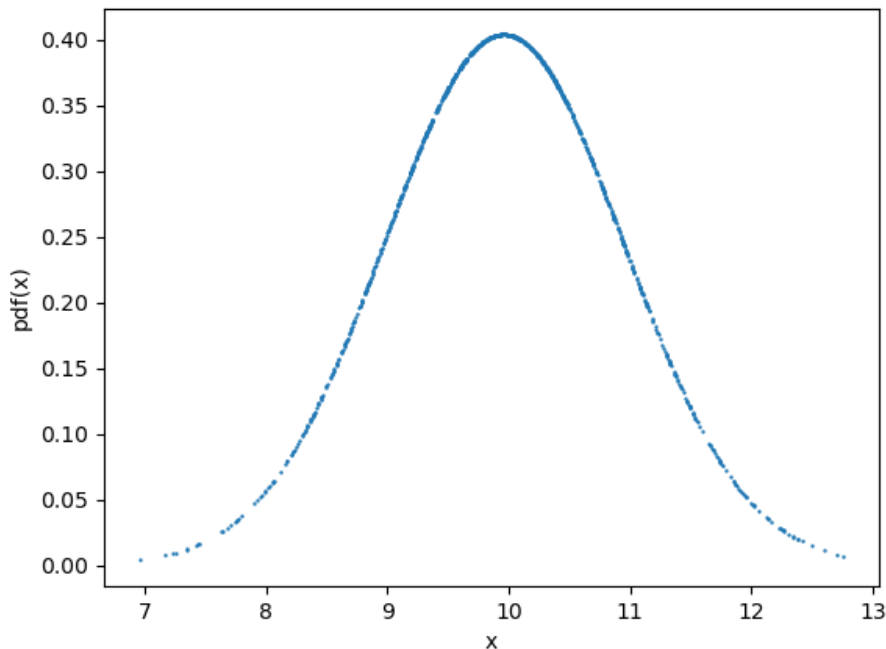


what are you expecting to see in the plot?

הגרלנו את וקטור ה- X מתוך התפלגות $N(10,1)$ ונצפה לראות צורת פעמון באופן דומה לפונקציות צפיפות נורמליות, עם מרכז סביב 10 (התוחלת) וצפיפות של נקודות בעיקר סביב התוחלת. זאת מכיוון שכזכור מקורס הסתברות שלקחנו בקירוב 68% מהדגימות בהתפלגות נורמלית נמצאות במרחק סטיית תקן אחת לכל היותר מהתוחלת, ובערך 95% במרחק שתי סטיות תקן מהתוחלת לכל היותר. מכיוון שהשונות קטנה (1) נצפה לראות נקודות בעיקר בתחום $[9,11]$.

(3)

Question 3- PDF values of samples from $N(10,1)$ distribution



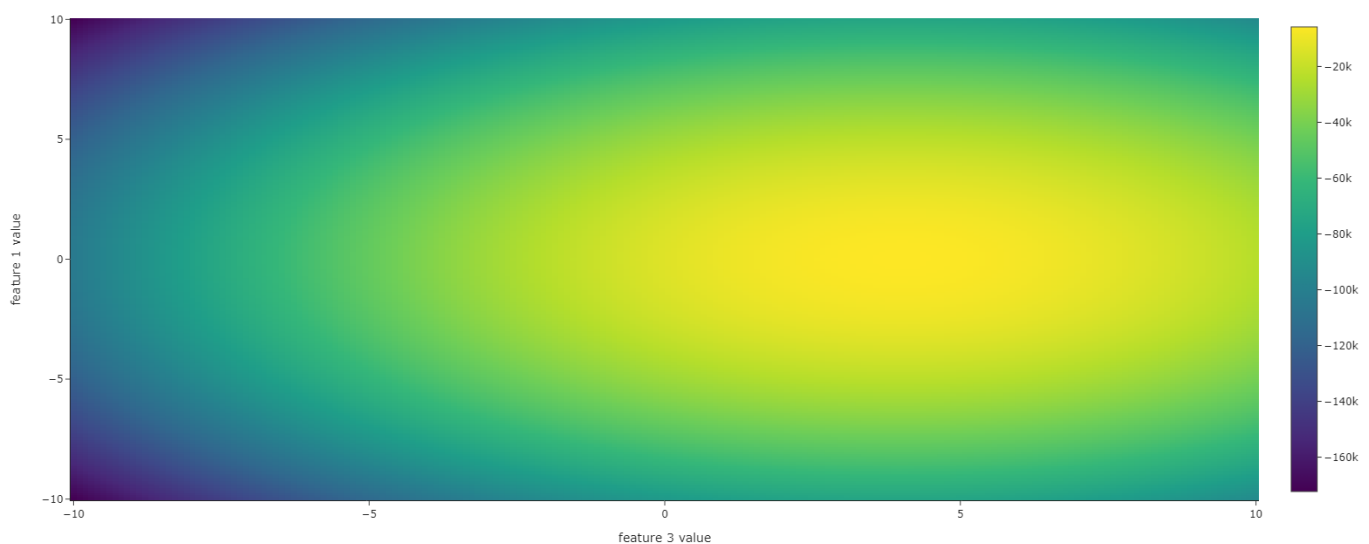
mean =
cov =

```
[-0.023 -0.043  3.993 -0.02 ]  
[[ 0.917  0.166 -0.03  0.463]  
 [ 0.166  1.974 -0.006  0.046]  
 [-0.03  -0.006  0.98  -0.02 ]  
 [ 0.463  0.046 -0.02  0.973]]
```

(2)

(5)

Question 5- Log likelihood of Multivariate Gaussian as function of mean's features 1,3



What are you able to learn from the plot?

נבחין שהאזור הצהוב, בו ה log-likelihood גבוה יותר, הוא האזור שבו ה features קרובים יותר לערך האמת שלהם בתוחלת ($f_1 = 0, f_3 = 4$). כלומר ניתן ללמוד מכך ששערוך לפי עיקרון ה למידה של מקסום פונקציית ה log-likelihood היא בחירה טובה (נזכור שהגענו לחפש מקסימום לפונקציית ה- log-likelihood מתוך מטרה לחפש את ה MLE).

(6)

```
features 1,3 that maximize the log-likelihood are:  
f1= -0.05 ,f3= 3.97
```