## Assignment 1: CS 215

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- (a) The situation is equivalent to distributing n books to n people. The total number of ways of doing that is n!. There is only 1 way in which everyone gets his book back. So the probability of it happening is 1/n!
- (b) There is only 1 way of distributing m books to their respective m owners. And for this way there are (n-m)! ways of distributing the left n-m books among left n-m people for a total of  $1 \times (n-m)! = (n-m)!$  ways. So the probability of it happening is (n-m)!/n!
- (c) There are m! way of distributing the m books belonging to the last m people to the first m people. And for each such way there are (n-m)! ways of distributing the left n-m books among left n-m people for a total of  $m! \ge (n-m)! = m!(n-m)!$  ways. So the probability of it happening is m!(n-m)!/n!
- (d) Every book has a probability p of getting unclean which is independent of who picked up which book and independent of whether other books became unclean. So the probability of first m people picking up books that are unclean will be  $\prod_m p = \boxed{p^m}$
- (e) The probability of some particular m people picking up books that are unclean and the rest clean books is  $p^m(1-p)^{n-m}$ . But there are  ${}^nC_m$  different ways of choosing a particular set of m people from a group of n. So the probability of it happening is  ${}^nC_mp^m(1-p)^{n-m}$

	L'age_
2.	For n distinct values sxi3 i=1
	we have
	σ= 1 2:(xj-μ) <sup>2</sup>
	V n-1
	where $\mu$ is mean and $\sigma$ is standard deviation
	consider the inequality
***	consider the inequality $q^2 + q^2 + \dots + q^2 \in  q^i   \forall i = 1, 2, \dots n$
	because
-	
-	
-	
	10-1 15 (xi-tu) 2 8 > 1 1xi-tr1 4 i=1 to u
-	70-1
+	·. (Jn-T) 5 > 1xi-µ1
	· · · · · · · · · · · · · · · · · · ·
1	consider the above proof for not and for nee!
	μ= × and σ=0, therefore inequality 1xi-μ1≤σ√m-1
1	holds.
1	The Thus the inequality is proved for all n >1.
	Now left up consider chebyshev's inequality for
	calculating fraction (estimitated) of 14 values satisfying
	this
	1xi-µ1 > 1n-1 o
	By using chebyshev's inequality,
	19KI < 1
	n n-1
	the same of the sa

Therefore freuction safisfying
1xi-41 50 Th-1 will be greater than
on equal to 1-1
n-1
As when increase 1 >0 and the above fraction >1
n-t
Thus we can say that as n increases chebyshev's
inequality starts starting that 1xi-141 & 5 5 17-1
is true of i=1 to n which is actually true.
as n increases chebyshev's inequality gives a very
good approximination.

3. We will try to prove it by contradiction.  Suppose   \mu - \tau  > \sigma  Case(1): When \tau > \mu:  We have \tau - \mu > \sigma = \tau > \sigma + \mu  As median is a element number will almost divides  the dafa in such a manner that median is a  central element.  Let say n=2k and [yi3] is in ascending order  thand yi are elements of xi then \mu lies  between yk and yk+1  \[ \frac{1}{2} \text{ yk+1} \]  Thus we can say that fraction of element such that  \[ \frac{1}{2} \text{ yk+1} \]  Thus we can say that fraction of element  \[ \frac{1}{2} \text{ yk+1} \]  \[ \frac{1}{2} \text{ yk+1} \text{ yk+1} \text{ yk+2} \]  \[ \frac{1}{2} \text{ yk+1} \text{ yk+2} \text{ yk+1} \text{ yk+2} \]  \[ \frac{1}{2} \text{ yk+1} \text{ yk+2} \text{ yk+2} \text{ yk+2} \text{ yk+2} \text{ yk+2} \]  \[ \frac{1}{2} \text{ yk+2} \]  \[ \frac{1}{2} \text{ yk+2} \]  \[ \frac{1}{2} \text{ yk+2} \]
Suppose Ip-T/>  Case(1): When t>p;  We have t-p>0 > t>0+p.  As median is a elem-number will almost divides  the dafa in such a manner that median is a  central element.  Let say n=2k and [yi3] is in ascending order  tond yi are elements of xi then p lies  befween yk and yk+1  - Yk & p & yk+1  Thus we can say that fraction of element such that  **X** **Y** **
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As median is a elem-number will almost divides  the data in such a manner that median is a  central element.  Let say n=2k and [yi3] is in ascending order  tond yi are elements of xi then u lies  between ye and yet!  Thus we can say that fraction of element such that  KE YESE TO STHU,  we can say that fraction of element  such that to yi > 5+ \mu must be > 1  Such that to yi > 5+ \mu must be > 1  Dut from one sided chebushev's inequality fraction
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Let say n=2k and [yi3] is in ascending order  tond yi are elements of xi then µ lies  before yh and yht  Thus we can say that praction of element such that  Xx yx x x y x x y x x y x x y x x x y x
before ye and yet!  Sefween ye and yet!  Thus we can say that praction of element such that  He yes to yet yet.  As the to > other,  we can say that fraction of element  we can say that fraction of element  such that to yi > other must be > 1  Such that to yi > other must be > 1  Dut from one sided chebusher's inequality fraction
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before ye and yet!  It ye & yet!  Thus we can say that fraction of element such that  Yet yet yet you you to must be > 1  As per to > o+per,  we can say that fraction of element  such that to yi > o+per must be > 1  Put from one sided chebushen's inequality fraction
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Thus we can say that fraction of element such that  HE HEST THE HISTORIAN TO MUST be > 1  NS HEST TO STHE,  WE can say that fraction of element  Such that To History must be > 1  Put from one sided chebushevis inequality fraction
As $\mu > \tau > \tau + \mu$ ,  we can say that fraction of element  such that $\tau > \tau + \mu$ must be $\tau > 1$ Put from one sided chebushevis inequality fraction
such that & yi > 0+ p must be > 1  But from one sided chebushev's inequality fraction
such that & yi >0+4 must be >1  But from one sided chebushev's inequality fraction
such that & yi >0+4 must be >1  But from one sided chebushev's inequality fraction
But from one sided chebushev's inequality fraction
But from one sided chebushev's inequality fraction
But from one sided chebyshev's inequality fraction of element such that yi-&> at must be \$ < 1 -1 1+(1)22
of element such that yi-to the must be so < 1 -1
1+(D <sup>2</sup> 2
We got contradiction
this means 14-51>0 cannot be true when
5>4
case 1: When t< 4;
By similar argument above and using the (K(O) form
of Chebyshev-Cantelli inequality.
14-5/20 connot be true when 5<4.
Teacher's Signature

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Therefore | \( \mu - \tau \) \( \sigma \) \( \text{full always be true as we showed that | \( \mu - \tau \) \( \text{ is always false.} \)

For \( \tau = \mu \)

\( \text{O} \le \sigma \) \( \text{which is obviously true} \)

\( \text{if \( \text{L} \) \( \text
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4.	P(A) = probability of rickshaw to be rep = 0.01
	P(B) = probability of rickshaw to be blue = 0.99
	Plet c be an event in which the rickshaw is advanty
	red. Observed
- 1	Then, $p(G A) = 0.99$
	P(G B) = 0.02
	Therefore by Bayes theorem,
	P(A C) = P(G A) P(A)
	P(C(A) P(A) + P(C(B) * P(B)
	= 0.99 x0.0
	0.99×0.01+0.99×0.02
	= 0.01 = 1
	= 0.01 = 1 0.01+0.02 3
	Therefore the probability that the rickshow is actually
	a red one when xxz observed it to be red = 1
	3

(a).

$$P(C_1|Z_1) = \frac{1}{3}$$

$$P(C_2|Z_1) = \frac{1}{3}$$

$$P(C_3|Z_1) = \frac{1}{3}$$

(b)  $P(H_3|C_1,Z_1)=\frac{1}{2}$  (if car is present behind first door so rest two doors have equal probability)

 $P(H_3|C_2,Z_1)=1$  (if car is present behind door 2, so the host will always pick door 3)

 $P(H_3|C_3,Z_1)=0$  (if car is present behind door 3, them host will never select it)

(c)

$$P(H_3|C_2, Z_1) = 1$$

$$P(C_2, Z_1) = P(C_2|Z_1) \times P(Z_1) = \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{9}$$

$$P(H_3|Z_1) = \frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot 1 + \frac{1}{3} \cdot 0 = \frac{1}{2}$$

$$P(H_3, Z_1) = P(H_3|Z_1) \times P(Z_1)$$

$$= \frac{1}{2} \cdot \frac{1}{3}$$

$$= \frac{1}{6}$$

$$P(C_2|H_3, Z_1) = \frac{P(H_3|C_2, Z_1)P(C_2, Z_1)}{P(H_3, Z_1)}$$

$$= \frac{1 \times \frac{1}{9}}{\frac{1}{6}}$$

$$= \frac{2}{3}$$

(d)

$$P(C_1|H_3, Z_1) = \frac{P(H_3|C_1, Z_1)P(C_1, Z_1)}{P(H_3, Z_1)}$$
$$= \frac{\frac{1}{2} \times \frac{1}{9}}{\frac{1}{6}}$$
$$= \frac{1}{3}$$

- (e)  $P(C_2|H_3, Z_1) > P(C_1|H_3, Z_1)$
- So, Switching will be beneficial

(f)

$$\forall_i P(H_3|C_i, Z_1) = \frac{1}{2}$$

$$\forall_i P(C_i, Z_1) = \frac{1}{3}$$

$$P(H_3|Z_1) = \frac{1}{2}$$

$$P(H_3, Z_1) = P(H_3|Z_1) \times P(Z_1) = \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{6}$$

$$P(C_2|H_3, Z_1) = \frac{P(H_3|C_2, Z_1)P(C_2, Z_1)}{P(H_3, Z_1)}$$

$$= \frac{\frac{1}{2} \times \frac{1}{9}}{\frac{1}{6}}$$

$$= \frac{1}{3}$$

$$P(C_1|H_3, Z_1) = \frac{P(H_3|C_1, Z_1)P(C_1, Z_1)}{P(H_3, Z_1)}$$

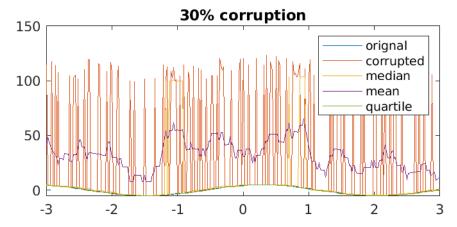
$$= \frac{\frac{1}{2} \times \frac{1}{9}}{\frac{1}{6}}$$

$$= \frac{1}{3}$$

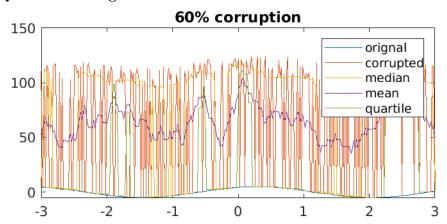
 $P(C_2|H_3,Z_1)=P(C_1|H_3,Z_1)$ . So switching won't create any difference.

The script for this question is under the file named q6.m. Please note that the relative mean squared error values listed below are averaged from multiple runs of q6.m but the plots are representative of only one run. The relative mean squared error values in the case of 30% and 60% corruption are given below respectively:

Moving Median Filtering 42.10 Moving Mean Filtering 103.81 Moving Quartile Filtering 0.01



Moving Median Filtering 744.02 Moving Average Filtering 376.55 Moving Quartile Filtering 76.17



In both the cases **moving quartile filtering** gives the least relative mean squared error. This can be explained as follows:

Lets choose a random x point from the set (-3:0.02:3). Let the set to be considered for filtering be  $S = \{x_{-8}, x_{-7} \dots x_0 \dots x_7, x_8\}$  where  $x_0 = x$  and  $x_i > x_j$  for i > j. On an average if the probability of corruption is p we can expect p|S| values to be corrupted. All the corrupted values are necessarily greater than the uncorrupted ones. And as we consider p values, here, to be 0.3 or 0.6 which is sufficiently greater than 0.75 we can expect the 25th percentile to be in S and hence close to x.

The average will be in between the the corrupted and uncorrupted (hence not close to x) and as the median is just the 50th percentile, chance of it being in S is lower in the case of p = 0.3 and much lower when p = 0.6.

The script of this question is under the file named q7.m

 $\mu = Old\ Mean$ 

 $\mu_n = New\ Mean$ 

 $\sigma = Old \; Standard \; Deviation$ 

 $\sigma_n = New \ Standard \ Deviation$ 

 $X = Old \; Median$ 

 $X_n = New\ Median$ 

 $A_i = Elemnts \ of \ array$ 

 $b = New\ element$ 

 $n = Size \ of \ old \ array$ 

1.

$$\sum_{i} A_{i} = \mu \times n$$

$$\mu_{n} = \frac{\sum_{i} A_{i} + b}{n+1}$$

$$\mu_{n} = \frac{(\mu \times n) + b}{n+1}$$

#### 2. • n is odd.

if b is less than the element left to median then the new median will lie betwee old median and number left to it.

if b lies betwee number left and right to old median then new median will lie between old median and added number.

if b is greater than the number right to meadian the new median will lie between old median and number right to it.

#### • n is even

lets say the old median lies between i and j.

if 
$$b < i, X_n = i$$

else if 
$$b > j, X_n = j$$

else 
$$X_n = b$$

3.

$$\sigma_n^2 = \frac{\sum_i^n (A_i - \mu)^2 + (b - \mu)^2}{n}$$

$$n \cdot \sigma_n^2 = \sum_i^n (A_i - \mu)^2 + (b - \mu)^2$$

$$Now,$$

$$\sum_i^n A_i^2 = (n - 1)\sigma^2 + 2n\mu^2 + -nx^2$$

 $Using\ above\ result$ 

$$n \cdot \sigma_n^2 = (n-1)\sigma^2 + n\mu^2 + n\mu_n^2 - 2n\mu\mu_n + (b-\mu_n)^2$$
$$= (n-1)\sigma^2 + b^2 + n\mu^2 + (n+1)\mu_n^2 - 2n\mu\mu_n - 2b\mu_n$$

 $Using\ result\ of\ first\ qsn$ 

$$n \cdot \sigma_n^2 = (n-1)\sigma^2 + b^2 + n\mu^2 - (n+1)\mu_n^2$$

4. We will increase the frequency of the bin which will contain the new element by 1.