

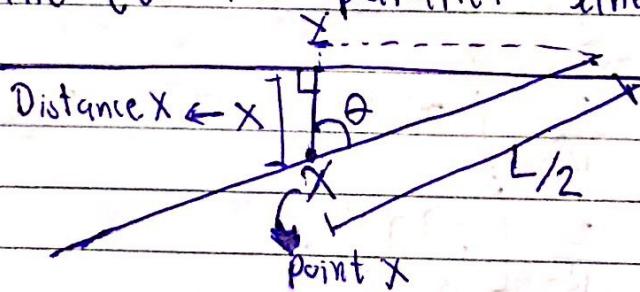
# Tute g Sol<sup>n</sup>.

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## Sol<sup>n</sup> I

This problem is known as "Buffon's needle problem".

- A table has equidistant parallel lines a distant  $D$  apart.
  - Length of a needle =  $L$ ,  $L \leq D$ .
  - Needle is thrown on table. What is the probability that it will intersect with line?
- Here the position of needle in the table can be shown by using 2 parameters.  
 $X$  = Distance between middle point of needle and the nearest line.  
 $\theta$  = Angle between needle and the perpendicular line (to the parallel lines).



- The needle will intersect if the projection of that half portion (of length  $L/2$ ) in that perpendicular line is greater than or equal to the perpendicular distance  $X$ .

$$\therefore \frac{L \cos \theta}{2} = \overrightarrow{XY} \geq X$$

$$\therefore X \leq \frac{L \cos \theta}{2}$$

→  $X$  is uniformly distributed between 0 and  $D/2$

$$\therefore f_X(x) = \begin{cases} \frac{2}{D} & 0 < x < \frac{D}{2} \\ 0 & \text{otherwise.} \end{cases}$$

→  $\theta$  is uniformly distributed between 0 &  $\pi/2$

$$\therefore f_\theta(y) = \begin{cases} 2/\pi & 0 < y < \pi/2 \\ 0 & \text{otherwise.} \end{cases}$$

$$\begin{aligned} \therefore P\left\{X \leq \frac{L \cos \theta}{2}\right\} &= \iint_{\substack{x \leq \frac{L}{2} \cos \theta \\ 0 \leq y \leq \frac{\pi}{2}}} f_X(x) f_\theta(y) d\theta dy \\ &= \int_0^{\pi/2} \int_0^{\frac{L}{2} \cos y} \frac{2}{D} \times \frac{2}{\pi} dx dy \\ &= \frac{4}{D\pi} \int_0^{\pi/2} \frac{L}{2} \cos y dy \\ &= \frac{4}{\pi D} \left[ \frac{L}{2} \sin y \right]_0^{\pi/2} \\ &= \boxed{\frac{2L}{\pi D}} \end{aligned}$$

Soln - 2

First of all Let's prove the convolution function.

- $X$  &  $Y$  are independent random variables.  
 $\rightarrow f_x(x) \rightarrow f_y(y) \rightarrow$  Density fns.  
 $\therefore F_{x+y}(a) = P[(X+Y) \leq a]$

$$= \int \int f_x(x) f_y(y) dx dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_x(x) f_y(y) dx dy$$

$$\therefore F_{x+y}(a) = \int_{-\infty}^{\infty} F_x(a-y) f_y(y) dy$$

$$\therefore f_{x+y}(y) = \frac{d}{da} [F_{x+y}(a)]$$

$$\boxed{f_{x+y}(y) = \int_{-\infty}^{\infty} f_x(a-y) f_y(y) dy}$$

- $x_i, i=1, 2, \dots, n$  are normally distributed independent random variables.

With parameters  $\mu_i$  &  $\sigma_i^2, i=1, 2, \dots, n$ .

- Let's count for 2 R.V.

$$x_1 \rightarrow \sigma_1^2, \mu_1$$

$$x_2 \rightarrow \sigma_2^2, \mu_2$$

$$f_{x_1}(x) = \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}$$

$$f_{x_2}(x) = \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{(x-\mu_2)^2}{2\sigma_2^2}}$$

Now  $Z = X_1 + X_2$  (Using Convolution function)

$$\therefore f_Z(z) = \int_{-\infty}^{\infty} f_{X_2}(z-x) f_{X_1}(x) dx$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{(z-x-\mu_2)^2}{2\sigma_2^2}} \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} dx$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sqrt{\sigma_1^2+\sigma_2^2}} \exp\left[-\frac{(z-(\mu_1+\mu_2))^2}{2(\sigma_1^2+\sigma_2^2)}\right] \frac{1}{\sqrt{2\pi}\sigma_1\sigma_2} \cdot dx$$

$$\exp\left[-\frac{(x - \frac{\sigma_1^2(z-\mu_2)+\sigma_2^2\mu_1}{\sigma_1^2+\sigma_2^2})^2}{2(\frac{\sigma_1\sigma_2}{\sqrt{\sigma_1^2+\sigma_2^2}})^2}\right]$$

$$= \frac{1}{\sqrt{2\pi(\sigma_1^2+\sigma_2^2)}} \exp\left(-\frac{(z-(\mu_1+\mu_2))^2}{2(\sigma_1^2+\sigma_2^2)}\right) * \frac{1}{\sqrt{2\pi}\frac{\sigma_1\sigma_2}{\sqrt{\sigma_1^2+\sigma_2^2}}} \exp\left[-\frac{(x - \frac{\sigma_1^2(z-\mu_2)+\sigma_2^2\mu_1}{\sigma_1^2+\sigma_2^2})^2}{2(\frac{\sigma_1\sigma_2}{\sqrt{\sigma_1^2+\sigma_2^2}})^2}\right] dx$$

$$= \frac{1}{\sqrt{2\pi(\sigma_1^2+\sigma_2^2)}} \exp\left(-\frac{(z-(\mu_1+\mu_2))^2}{2(\sigma_1^2+\sigma_2^2)}\right) \quad \text{①}$$

(Because the term (integral term) will be 1 as the density fn of  $X$  is there in integral).

- So we are getting  $Z$  as a ~~Normal~~ Random variable with parameters  $\mu_1+\mu_2$  &  $\sigma_1^2+\sigma_2^2$

\* To prove for  $n$  normal random variables we will use the induction hypothesis.

•  $P(0)$ :  $X_i, i=1, 2, \dots, n$  are normal random variables with parameters  $\mu_i$  &  $\sigma_i^2$   
 $i=1, 2, \dots, n \Rightarrow \sum_{i=1}^n X_i$  is a normal random variable with parameters  $\sum_{i=1}^n \mu_i$  &  $\sum_{i=1}^n \sigma_i^2$ .

•  $P(0) \Rightarrow$  "Normal random variable itself is a normal random variable"  
 $\Rightarrow P(0)$  is true.

•  $P(1)$  is also true according to result (I).

• Let's assume that  $P(k)$  is true  
 $\therefore \sum_{i=1}^k X_i$  is a normal random variable with parameters  $\sum_{i=1}^k \mu_i$  &  $\sum_{i=1}^k \sigma_i^2$ .

to prove for  $P(k+1)$

$$\sum_{i=1}^{k+1} X_i = \sum_{i=1}^k X_i + X_{k+1}$$

now in result (I) putting  $\sigma_1^2 = \sum_{i=1}^k \sigma_i^2$  &

$$\mu_1 = \sum_{i=1}^k \mu_i$$

$$(\sigma_1^2 = \sigma_{k+1}^2)$$

$$(\mu_1 = \mu_{k+1})$$

$$X_1 = \sum_{i=1}^k X_i \text{ & } X_2 = X_{k+1}$$

So we will get the  $\sum_{i=1}^{k+1} x_i$  also as a normal random variable with parameters  $\sum_{i=1}^{k+1} \mu_i$  &  $\sum_{i=1}^{k+1} \sigma_i^2$ .

$\therefore P(k)$  is true  $\Rightarrow P(k+1)$  is true.

$\Rightarrow P(n)$  is true for all  $n$ .

$\Rightarrow \sum_{i=1}^n x_i$  is a normal R.V. with parameters  $\sum_{i=1}^n \mu_i$  &  $\sum_{i=1}^n \sigma_i^2$ .

### Soln 3

A Basketball team has to play 44 games.

• 26 against team A. Winning probability = 0.4

$X_A$  = number of games won by a team against team A.

$$n_1 = 26 \quad p_1 = 0.4$$

• 18 game against team B.

$X_B$  = number of games won against team B.

$$n_2 = 18 \quad p_2 = 0.7$$

By the normal approximation of the binomial r.v.,  $X_A$  &  $X_B$  can be seen as a normal random variables.

$$X_A \rightarrow \mu_A = n_1 p_1 = 26 \times 0.4 = 10.4$$

$$\sigma_A^2 = 26 \times 0.4 \times 0.6 = 6.24$$

$$X_B \rightarrow \mu_B = n_2 p_2 = 18 \times 0.7 = 12.6$$

$$\sigma_B^2 = n_2 p_2 (1-p_2) = 18 \times 0.7 \times 0.3 = 3.78$$

$X_A + X_B$  is a normal random variable with parameters  $\mu_A + \mu_B$  &  $\sigma_A^2 + \sigma_B^2$ .

$$= 23$$

$$= 10.02$$

(c)  $P\{ \text{The team wins } 25 \text{ or more games} \}$   
 $= P\{ X_A + X_B \geq 25 \} = P\{ X_A + X_B \geq 24.5 \}$  (continuity correction)

$$= P\left\{ \frac{X_A + X_B - 23}{\sqrt{10.02}} \geq \frac{24.5 - 23}{\sqrt{10.02}} \right\}$$

$$= P\left\{ Z \geq \frac{1.5}{\sqrt{10.02}} \right\}$$

$$= 1 - P\{ Z < 0.4739 \}$$

$$\boxed{T = 0.3178}$$

(b) The team wins more games against team A than it does against team B.  
 $\rightarrow X_A - X_B \rightarrow \text{normal R.V. with parameters}$   
 $\mu_1 - \mu_2 = -2.2$  & variance  
 $= \sigma_1^2 + (-\sigma_2)^2 = \sigma_1^2 + \sigma_2^2 = 10.02.$

 $\therefore P\{ X_A > X_B \} = P\{ X_A - X_B > 0 \}$ 
 $= P\{ X_A - X_B > 0.5 \}$  (continuity correction)
 $= P\left\{ \frac{X_A - X_B - \mu_1}{\sqrt{10.02}} \geq \frac{0.5 + 0.22}{\sqrt{10.02}} \right\}$ 
 $= P\left\{ Z \geq \frac{2.7}{\sqrt{10.02}} \right\}$ 
 $= 1 - P\{ Z < 0.8530 \}$ 

$$\boxed{T = 0.1968}$$

Sol 4  $S(n)$  = Price of a certain security after  $n$  additional weeks.  $n \geq 1$

$\frac{S(n)}{S(n-1)}$ ,  $n \geq 1$ , are independent and identically distributed lognormal random variables.

$$\mu = 0.0165$$

$$\sigma = 0.0730$$

here  $\log\left(\frac{S(n)}{S(n-1)}\right)$  is a normal random

variable.

(a)  $P\{\text{The price of security increases over each of the next two weeks}\}$

$$\left( P\left[ \frac{S(1)}{S(0)} > 1 \right] \right)^2 \text{ Here } \Rightarrow P\left[ \frac{S(1)}{S(0)} > 1 \right] = P\left[ \log\left(\frac{S(1)}{S(0)}\right) > 0 \right]$$

for next  
1 week

$$= P\left[ Z > \frac{0 - 0.0165}{0.0730} \right]$$

$$= P\left[ Z < 0.2260 \right]$$

$$= 0.5894$$

here  $\frac{S(1)}{S(0)} > 1$  denotes that the price has

been increased over one week. for 2 weeks it will be  $(0.5894)^2 = 0.3474$ .  
(as the events are independent)

(b)  $P\{\text{The price is higher than today at the end of 2 weeks}\}$

$$= P\left[ \frac{S(2)}{S(0)} > 1 \right] = P\left[ \frac{S(2)}{S(1)} \frac{S(1)}{S(0)} > 1 \right]$$

$$= P \left\{ \log \left( \frac{s(2)}{s(1)} \right) + \log \left( \frac{s(1)}{s(0)} \right) > 0 \right\}$$

here  $\log \left( \frac{s(2)}{s(1)} \right) + \log \left( \frac{s(1)}{s(0)} \right)$  is a sum of 2

independent normal random variables.

$$\therefore \mu = \mu_1 + \mu_2 = 2 \times 0.0165 = 0.0330$$

$$\therefore \sigma = \sqrt{\sigma_1^2 + \sigma_2^2} = \sqrt{2} \times 0.0730 = 0.1032$$

$$\therefore P \left\{ \log \left( \frac{s(2)}{s(1)} \right) > 1 \right\} = P \left\{ Z > \frac{-0.0330}{0.1032} \right\}$$

$$= P \{ Z < 0.3195 \}$$

$$= 0.6254$$