

# Lecture -15

P①

## Recap:

Normal distribution

$$N(\mu, \sigma) \rightarrow N(0, 1)$$

$$X \rightarrow \frac{X - \mu}{\sigma}$$

↓  
we have  
tables for  $\Phi(a)$

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Normal approx. to Binomial

$$\mu = np$$

$$\sigma = \sqrt{np(1-p)}$$

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(continuity correction

$$P(X=i) = \binom{n}{i} p^i (1-p)^{n-i}$$

$$P(i - \frac{1}{2} \leq X \leq i + \frac{1}{2}) \text{ for normal}$$

## Exponential random variables

(2)

Usually used to model  
time: time until next  
earth quake or next war or  
the next accident on S.G. Highway.

let  $\lambda > 0$

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

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Prove that it is a probability  
density function.

$$\int_{-\infty}^{\infty} f(x) dx = 1 = \int_0^{\infty} f(x) dx \quad \text{H.W.}$$



$$E[X] = \frac{1}{\lambda}$$

H.W.

③

$$\text{Var}[X] = \frac{1}{\lambda^2}$$

H.W.

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$$F(a) = P(X \leq a)$$

$$= \int_0^a f(x) dx = 1 - e^{-\lambda a}$$

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e.g. Suppose that the length of a phone call, in minutes, is an exponential random variable with  $\lambda = \frac{1}{10}$ . Suppose

someone arrives at a PCO just before you. What is the probability that you need to wait for more than 10 mins?

$X$  = no. of minutes that ④  
you need to wait

$$P(X > 10) = 1 - P(X \leq 10)$$

$$= 1 - F(10)$$

$$= 1 - (1 - e^{-\frac{1}{10} \cdot 10})$$

$$= e^{-1} = \frac{1}{e}$$

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Memorylessness property (5)

$$P(X > s+t | \cancel{X} > t) = P(X > s)$$

$X$ : battery life

$$t = 50$$

$$s = 50$$

$$\frac{P(X > s+t \cap \underline{X} > t)}{P(X > t)}$$

$$= \frac{P(X > s+t)}{P(X > t)} = \frac{1 - F(s+t)}{1 - F(t)}$$

$$= \frac{e^{-\lambda(s+t)}}{e^{-\lambda t}} = e^{-\lambda s} = P(X > s)$$



# Functions of continuous random variable

⑥

e.g.  $X$  is uniformly distributed over  $(0,1)$ .

$$f(x) = \begin{cases} 1 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise.} \end{cases}$$

$$Y = X^3$$

How is  $Y$  distributed?

$$F_a(x) = \int_0^a f(x) dx = a$$

$$\begin{aligned} P(Y \leq y) &= P(X^3 \leq y) \\ &= P(X \leq y^{1/3}) = F_X(y^{1/3}) \\ P(Y \leq y) &= y^{1/3} = F_Y(y) \end{aligned}$$

$$F_Y(y) = y^{1/3}$$

⑦

$$\begin{aligned} f_Y(y) &= \frac{d}{dy} (y^{1/3}) \\ &= \frac{1}{3} y^{-2/3} \end{aligned}$$

e.g.  $X$  is c.o.v.  $f_X$   
 $P(X < \sqrt{y})$

$$Y = X^2$$

$$P(Y \leq y) = P(X^2 \leq y)$$

$$= P(-\sqrt{y} < X < \sqrt{y})$$

$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

$$f_Y(y) = \frac{1}{2\sqrt{y}} f_X(\sqrt{y}) - \frac{-1}{2\sqrt{y}} f_X(-\sqrt{y})$$

$$= \frac{1}{2\sqrt{y}} (f_X(\sqrt{y}) + f_X(-\sqrt{y}))$$

e.g.  $X : f_X$

⑧

$$Y = |X|$$

$$F_Y(y) = P(Y \leq y)$$

$$= P(|X| \leq y)$$

$$= P(-y \leq X \leq y)$$

$$= F_X(y) - F_X(-y)$$

$$\begin{aligned} f_Y(y) &= f_X(y) - (-1) f_X(-y) \\ &= f_X(y) + f_X(-y) \end{aligned}$$



Theorem:

⑨

$X, f_X(x)$

$g(x)$  is strictly monotonic  
(increasing or decreasing),  
differentiable function of  $x$ .

Then  $Y = g(X)$  has the  
probability density function

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} (g^{-1}(y)) \right| & \text{if } y = g(x) \\ 0 & \text{if } y \neq g(x) \text{ for } \forall x \end{cases}$$

$$f(y)(x)$$

$$g(x) = x^3$$

(10)

$$= f_x(g^{-1}(y)) \left| \frac{d(g^{-1}(y))}{dy} \right|$$

$$= f_x(y^{1/3}) \left| \frac{d(y^{1/3})}{dy} \right|$$

$$= \frac{1}{3 y^{2/3}} f_x(y^{1/3})$$

$$= \frac{1}{3 y^{2/3}}$$