# Name: Shivani Patel

## GSI: Madeline Brandt DISC #: 103



## Math 55, Handout 16.

## PROBABILITY THEORY. 452

1.1. Let S be a sample space of an experiment with a finite or countable number of We assign a probability p(s) to each outcome  $s \in S$  so that two conditions are met:  $0 \le p(s) \le 1$  for each ses

outcomes.

- 1.2. Let |S| = n. The uniform distribution assigns the probability 1/n to each element of S
- 1.3. Selecting an element of S at random means the experiment of selecting an element from a sample space with a uniform distribution
- 1.4. The probability of an event  $E\subseteq S$  is then defined as the sum of the probabilities of the nutcomes in E



Q1. Suppose that a die is biased (loaded) so each odd number is twice as likely to appear as each even number. What is the probability that an odd and an even number appear when we roll this die twice? outcomes: odd (success) or even(fail)

prob of odd=p prob of even=q

$$p=20$$
 \$  $p+9=1$   $p=\frac{2}{3}$  therefore:  $b(1;2,\frac{2}{3})=C(2,1)(\frac{2}{3})^{1}(\frac{1}{5})^{1}=\frac{4}{9}$ 

## UNIONS AND COMPLEMENTS. 465

2.1. The following formulæ continue to hold in the non-uniform case:

$$\begin{array}{rcl} p(\overline{E}) & = & | - \rho(\xi) \\ p(E_1 \cup E_2) & = & \rho(\xi_1) + \rho(\xi_2) - \rho(\xi_1 \cap \xi_2) \end{array}$$

2.2. [Theorem.] Let  $E_1, \ldots, E_k$  be pairwise disjoint events in a sample space S. Then



$$p(\cup_{j=1}^k E_j) = \sum_{\mathbf{j}} \mathsf{P(E_j)}$$

Q2. Prove this Theorem.

PASIC Step: Trne 
$$\rightarrow$$
 n=[ bc  $\rho(\mathcal{E}_{i}) = \rho(\mathcal{E}_{i})$ 

N=2,  $\rho(\mathcal{E}_{i} \cup \mathcal{E}_{2}) = \rho(\mathcal{E}_{i}) + \rho(\mathcal{E}_{2}) - \rho(\mathcal{E}_{i} \cap \mathcal{E}_{2}) = \rho(\mathcal{E}_{i}) + \rho(\mathcal{E}_{2})$ 
 $\rightarrow \mathcal{E}_{1} \triangleq \mathcal{E}_{2} \text{ are disjitht}$ 

Industive of  $\mathcal{E}_{1} = \mathcal{E}_{2} = \mathcal{E}_{2} = \mathcal{E}_{3} = \mathcal{$ 

$$P(V_{j=1}^{k} + | E_{j}) = P((V_{j=1}^{k} E_{j}) | V_{E_{k+1}}) = P(V_{j=1}^{k} E_{j})$$

$$+ P(E_{k+1}) = \sum_{j=1}^{k+1} P(E_{j})$$

2.3. [Boole's Inequality.] Let  $E_1, \ldots, E_k$  be not necessarily disjoint events in a sample space S. Then

$$p(\cup_{j=1}^k E_j) \le \rho(F_1) + \dots + \rho(F_k)$$

Q3. Prove Boole's Inequality.

$$\begin{array}{l} & \text{BASIC step: } n=2 \text{ , } \rho\left(E_{1} \cup E_{2}\right) = \rho(E_{1} \cup E_{2}) = \rho(E_{1}) + \rho(E_{2}) - \rho(E_{1} \cap E_{2}) \leq \rho(E_{1}) + \rho(E_{2}) \\ & \text{Inductive step: } n=k \text{ is } \text{true, } \rho(\bigcup_{j=1}^{k} E_{j}) \leq \rho(E_{1}) + \ldots + \rho(E_{k}) \\ & \text{for } n=k+1 \text{ , } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho((\bigcup_{j=1}^{k} E_{j}) \vee E_{k+1}) \leq \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho(E_{k+1}) \leq \rho\left(E_{1}\right) + \ldots + \rho\left(E_{k+1}\right) \\ & \text{for } n=k+1 \text{ , } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(E_{k+1}\right) \leq \rho\left(E_{1}\right) + \ldots + \rho\left(E_{k+1}\right) \\ & \text{for } n=k+1 \text{ , } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(E_{k+1}\right) \leq \rho\left(E_{1}\right) + \ldots + \rho\left(E_{k+1}\right) \\ & \text{for } n=k+1 \text{ , } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(E_{1}\right) + \dots + \rho\left(E_{k+1}\right) \\ & \text{for } n=k+1 \text{ . } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(E_{1}\right) + \dots + \rho\left(E_{k+1}\right) \\ & \text{for } n=k+1 \text{ . } \rho\left(\bigcup_{j=1}^{k} E_{j}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(E_{1}\right) + \dots + \rho\left(E_{k+1}\right) = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(\bigcup_{j=1}^{k} E_{j}\right) + \rho\left(\bigcup_{j=1}^{k} E_{j}\right) \vee E_{k+1} = \rho\left(\bigcup_{$$

#### CONDITIONAL PROBABILITY.

- 3.1. Let E and F be events with p(F) > 0. The **conditional probability** of E given F is defined by  $p\left(E[F] = \frac{p(E \cap F)}{P(F)}\right)$
- 3.2. If two events E and F are **independent**, then p(E|F) =
- Q3. Are the events E, that a family with three children has children of both sexes, and F, that this family has at most one girl, independent? Assume that the child gender at birth is unambiguous and that the resulting eight ways a family can have three children are equally likely.

B=BOY, G=GIrl  
POSSIBLE OUTCOMED: BBB, BBG, BGB, GBB, GBB, BGG, GGB, GGG, GGB, GGG  
E=BBG, BGB, GBB, BGG, GBB, GGB 
$$\longrightarrow p(E) = \frac{6}{8} = \frac{3}{4}$$
  
F=BBB, BBG, BGB, GBB  $\longrightarrow p(F) = \frac{3}{8} = \frac{3}{8}$   
P(E) × P(F) =  $\frac{3}{4} * \frac{1}{2} = \frac{3}{8} = P(E \cap F)$  SO E & F are independent



- 3.2. Events  $E_1, \ldots, E_k$  are pairwise independent if  $p(\xi_i \cap \xi_j) = p(\xi_i) p(\xi_j)$  for all pairs of integers i and j with  $|\neq i < j \leq n$
- 3.3. Events  $E_1, \ldots, E_k$  are mutually independent if  $\rho(\xi_{i1} \cap \ldots \cap \xi_{ik}) = \rho(\xi_{i1}) \cdots \rho(\xi_{ik})$

#### BERNOULLI TRIALS AND THE BINOMIAL DISTRIBUTION.

Q4. A coin biased so that the probability of heads is 3/4 is tossed 10 times. What is the probability that exactly five heads come up assuming that the flips are mutually independent?

$$P = C(10,5)(3/4)^{5}(1-3/4)^{\frac{1}{5}} = 0.058$$

$$= C(10,5)(3/4)^{5}(1/4)^{\frac{1}{5}} = 0.058$$

4.1. [Theorem.] The probability of exactly k successes in n independent Bernoulli trials, with probability of success p and probability of failure q = 1 - p, is equal to  $C(k, n) p^k q^{n-k}$ 

#### RANDOM VARIABLES.

- 5.1. A random variable is a function from the sample space of an experiment to the set of real numbers
- 5.2. [NB.] A random variable is neither random nor a variable
- 5.3. The distribution of a random variable X on a sample space S is the set of pairs (r, p(x=r)) for all  $r \in S$  where p(x=r) is the probability that x takes the value r
- Q5. What is the distribution of the total number of heads when a fair coin is flipped four times?

$$P(X=0) = C(4,0)(1/2)^{0}(1/2)^{4} = 1/10$$

$$P(X=1) = C(4,1)(1/2)^{1}(1/2)^{3} = 1/4$$

$$P(X=2) = C(4,2)(1/2)^{2}(1/2)^{2} = 3/8$$

$$P(X=3) = C(4,3)(1/2)^{3}(1/2)^{1} = 1/16$$

$$P(X=4) = C(4,4)(1/2)^{4}(1/2)^{0} = 1/16$$

