

MULTIPLE LINEAR REGRESSION



Outline: Multiple Linear Regression

Motivation

- Model and Interpretation
- Estimation and Inference

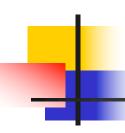
Interaction

Motivation

The response or dependent variable, Y, may depend on several predictors not just one!

 Multiple regression is an attempt to consider the <u>simultaneous</u> influence of several variables on the response

 This may be with the goal of an unbiased estimate of association or for better prediction

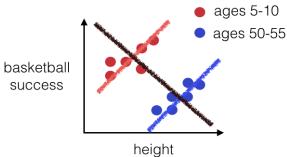


Motivation

- Why not fit multiple separate simple linear regressions?
 - If the goal is to estimate the association between the response and a predictor of interest, a confounder can make the observed association appear
 - stronger than the true association,
 - weaker than the true association, or
 - even the reverse of the true association



- We can adjust for the effects of the confounder by adding a corresponding term to our linear regression
- If the goal is prediction of the response, we may be able to improve prediction by including additional variables in the regression model





Motivation: Cholesterol Example

Data

```
> head(cholesterol)

ID DM age chol BMI TG APOE rs174548 rs4775401 HTN chd

1 1 74 215 26.2 367 4 1 2 1 1

2 1 51 204 24.7 150 4 2 1 1 1

3 0 64 205 24.2 213 4 0 1 1

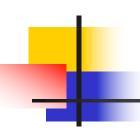
4 0 34 182 23.8 111 2 1 1 1

5 1 52 175 34.1 328 2 0 0 1 0

6 1 39 176 22.7 53 4 0 2 0 0
```

Our goal:

 Investigate the relationship between age (years), BMI (kg/m²) and serum total cholesterol (mg/dl)



Motivation

In general, the multiple regression equation can be written as follows:

$$E[Y | x_1, x_2, ..., x_p] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + ... + \beta_p x_p$$

- We use multiple variables when:
 - The predictor variable is categorical with more than two groups
 - We need polynomials, splines or other functions to model the shape of the relationship(s) accurately
- Estimating association:
 - We want to adjust for confounding by other variables
 - We want to allow the association to differ for different values of other variables (interaction)
- Prediction: we use multiple variables if we think more than one variable will be useful in predicting future outcomes accurately

Model and Interpretation

$$\text{Model:} \quad Y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \ldots + \beta_p x_p + \varepsilon$$
 where we assume
$$\varepsilon \sim N(0, \sigma^2)$$

Extension of simple linear regression!

Systematic component:

$$E[Y|x_1,...,x_p] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + ... + \beta_p x_p$$

Random component:

$$Var[Y | x_1, ..., x_p] = \sigma^2$$

4

Model and Interpretation

For example, let us assume that there are two predictors in the model and so $E[Y|x_1, x_2] = \beta_0 + \beta_1 x_1 + \beta_2 x_2$

Consider two observations with the same value for x_2 , but one observation has x_1 one unit higher, that is,

Obs 1:
$$E[Y|x_1=k+1, x_2=c] = \beta_0 + \beta_1 (k+1) + \beta_2 c$$

Obs 2:
$$E[Y|x_1=k, x_2=c] = \beta_0 + \beta_1(k) + \beta_2 c$$

Thus,
$$E[Y|x_1=k+1, x_2=c] - E[Y|x_1=k, x_2=c] = \beta_1$$

That is, β_1 is the expected mean change in y per unit change in x_1 if x_2 is held constant (adjusted/controlling for x_2)

Similar interpretation applies to β_2



Model and Interpretation

 To facilitate our discussion let's assume we have two predictors with binary values

Model:

$$E[Y | x_1, x_2] = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

| Mean of Y | $X_2 = 0$ | X ₂ =1 | E[Y |
|-------------------|---------------------|-------------------------------|--------------|
| X ₁ =0 | β_0 | $\beta_0 + \beta_2$ | E[\ |
| $X_1 = 1$ | $\beta_0 + \beta_1$ | $\beta_0 + \beta_1 + \beta_2$ | E[\ E[\ |

E[Y|x₁=1, x₂=0] - E[Y|x₁=0,x₂=0] =
$$\beta_1$$

E[Y|x₁=1, x₂=1] - E[Y|x₁=0,x₂=1] = β_1
E[Y|x₁=0, x₂=1] - E[Y|x₁=0,x₂=0] = β_2

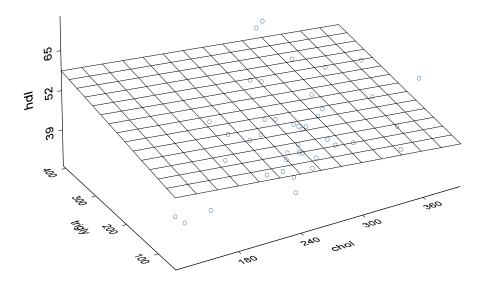
$$E[Y|X_1=1, X_2=1] - E[Y|X_1=1, X_2=0] = \beta_2$$

Estimation

- Least Squares Estimation:
 - As in linear regression, chooses the coefficient estimates that minimize the residual sum of squares

$$D = \sum_{i} (y_i - \hat{y}_i)^2$$

Computation more difficult, but statistical software (R) will do that for you!



Estimation and Inference

- Inference
 - About regression model parameters
 - Hypothesis Testing H_0 : $\beta_j=0$ (j=0,1,2,...,p)

<u>Interpretation:</u> Is there a statistically significant relationship between the response y and x_j after adjusting for all other factors (predictors) in the model?

Test Statistic:
$$\frac{\hat{\beta}_{j} - (null \ hyp)}{se(\hat{\beta}_{j})} \sim t_{n-p-1}$$

Note: The square of the t-statistic gives the F-statistic and the test is known as the **partial F-Test**

Confidence Intervals

$$\hat{\beta}_{j} \pm (critical\ value) \times se(\hat{\beta}_{j})$$



Estimation and Inference

- About the full model
 - Hypotheses

H₀:
$$\beta_1 = \beta_2 = ... = \beta_p = 0$$
 vs. H₁: At least one β_j is not null

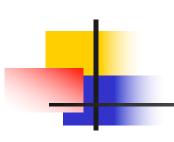
Analysis of variance table

| Source | df | SS | MS | F |
|------------|-------|---|---------------------|---------|
| Regression | р | $SSR = \sum (\hat{y}_i - \overline{y}_i)^2$ | MSR= SSR/p | MSR/MSE |
| Residual | n-p-1 | $SSE = \sum (y_i - \hat{y}_i)^2$ | MSE= SSE/(n-p-1) | |
| Total | n-1 | $SST = \sum (y_i - \overline{y})^2$ | | |



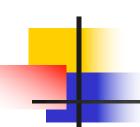
Estimation and Inference

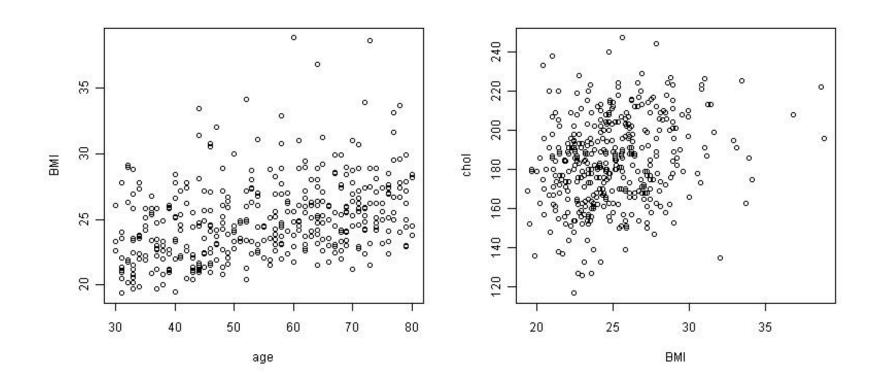
- The F-value is tested against a F-distribution with p, n-p-1 degrees of freedom
 - If we reject the null hypothesis, then the predictors do aid in predicting Y [in this analysis we do not know which ones are important!]
 - Failing to reject the null hypothesis does not mean that none of the covariates are important, since the effect of one or more covariates may be "masked" by others. The hard part is choosing which covariates to include or exclude.
- This is known as the global (multiple) F-test

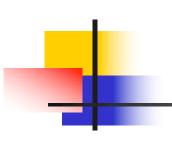


 We have seen that there is a significant relationship between age and cholesterol

Can we better understand variability in cholesterol by incorporating additional covariates?







- It appears that BMI increases with increasing age
- And cholesterol increases with increasing BMI
- What if we want to estimate the association between age and cholesterol while holding BMI constant?
- Multiple regression!



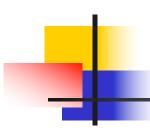
```
> fit2=lm(chol~age+BMI)
> summary(fit2)
Call:
lm(formula = chol \sim age + BMI)
Residuals:
   Min 1Q Median 3Q
                                 Max
-58.994 -15.793 0.571 14.159 62.992
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 137.1612 9.0061 15.230 < 2e-16 ***
        0.2023 0.0795 2.544 0.011327 *
age
           1.4266 0.3822 3.732 0.000217 ***
BMT
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.34 on 397 degrees of freedom
Multiple R-squared: 0.07351, Adjusted R-squared: 0.06884
F-statistic: 15.75 on 2 and 397 DF, p-value: 2.62e-07
```



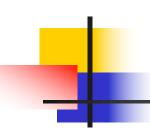
Our estimated regression equation is

$$\hat{y} = 137.16 + 0.20 Age + 1.43 BMI$$

• Question: How do we interpret the age coefficient?



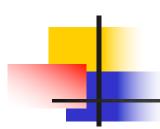
- Our estimated regression equation is $\hat{y} = 137.16 + 0.20 Age + 1.43 BMI$
- Question: How do we interpret the age coefficient?
- Answer: This is the estimated average difference in cholesterol associated with a one year difference in age for two subjects with the same BMI.



Our estimated regression equation is

$$\hat{y} = 137.16 + 0.20 Age + 1.43 BMI$$

- The age coefficient from our simple linear regression model was 0.31.
- Question: Why do the estimates from the two models differ?



Our estimated regression equation is

$$\hat{y} = 137.16 + 0.20 Age + 1.43 BMI$$

- The age coefficient from our simple linear regression model was 0.31.
- Question: Why do the estimates from the two models differ?
- Answer: We are now conditioning on or controlling for BMI so our estimate of the age association is among subjects with the same BMI.



```
Call:
lm(formula = chol \sim age + BMI)
Residuals:
   Min
       10 Median 30
                               Max
-58.994 -15.793 0.571 14.159 62.992
Coefficients:
          Estimate Std. Error t value Pr(>|t|)
(Intercept) 137.1612 9.0061 15.230 < 2e-16 ***
       age
       1.4266 0.3822 3.732 0.000217 ***
BMI
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 21.34 on 397 degrees of freedom
Multiple R-squared: 0.07351, Adjusted R-squared: 0.06884
F-statistic: 15.75 on 2 and 397 DF, p-value: 2.62e-07
```

Cholesterol Example:

Did adding BMI improve our model?

How does the model with age and BMI compare to a model that contains only the mean?



Interaction and Linear Regression

- Statistical interaction (aka effect modification) occurs when the relationship between an outcome variable and one predictor is different depending on the levels of a second predictor
- Interactions are usually investigated because of a priori assumptions/hypotheses on the part of the researchers

 Linear regression models allow for the inclusion of interactions with cross-product terms



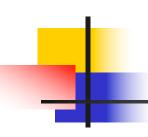
Confounding vs. Interaction/Effect Modification

- Data and scientific understanding help distinguish between confounding and effect modifying variables:
 - Confounder: Associated with predictor and response; Association between response and predictor constant across strata of the new variable
 - Effect modifier/interaction: Association between response and the predictor varies across strata of the new variable



Confounding vs. Interaction/Effect Modification

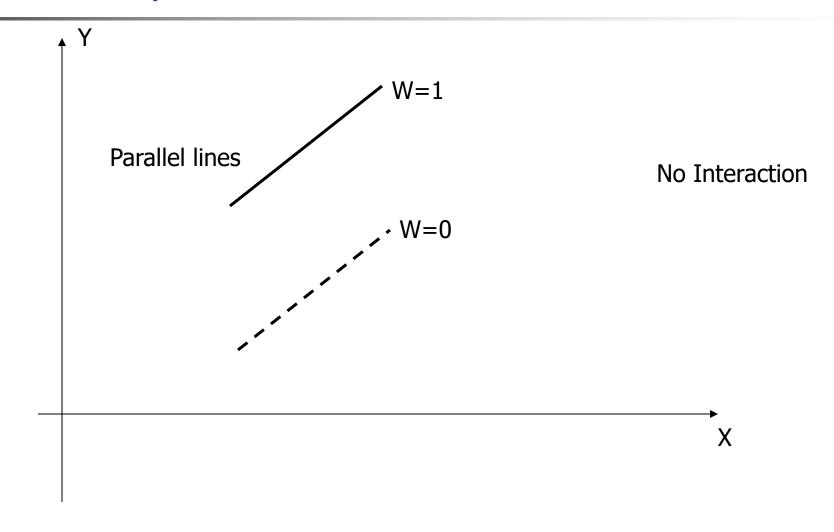
- Confounding: Estimates of association from unadjusted analysis are markedly different from estimates of association from adjusted analysis
 - Association within each stratum is similar, but different from the "crude" association in the combined data (ignoring the strata)
 - In linear regression, these symptoms are diagnostic of confounding
- Effect modification would show differences between adjusted analysis and unadjusted analysis, but would also show different associations in the different strata



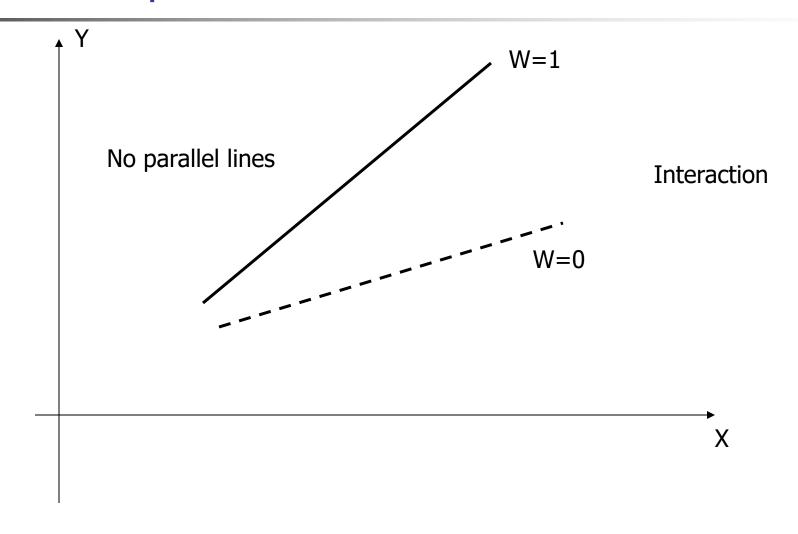
Effect Modification /Interaction

- Even if present, effect modification may not always be of interest in summarizing the effect of a predictor.
- For example, pleconaril, an antiviral drug, reduced the mean duration of symptoms in subjects with a common cold due to rhinoviruses but had no effect in subjects whose cold was due to some other agent.
- In the case of the pleconaril, effect modification was important in checking that the drug did actually work by inhibiting rhinovirus. However, in clinical use of the drug, it would typically not be possible to determine the infectious agent (the tests are expensive and take longer than just recovering from the cold), and so the average effectiveness of the drug across all colds would be a more important quantity.

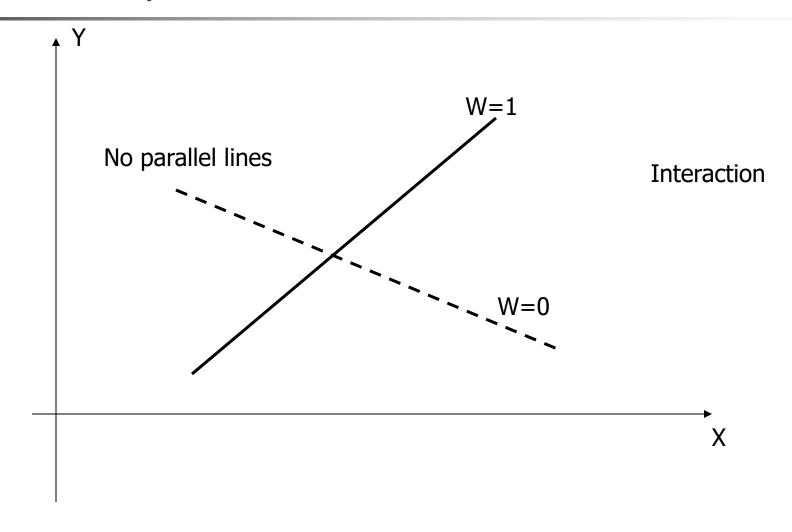




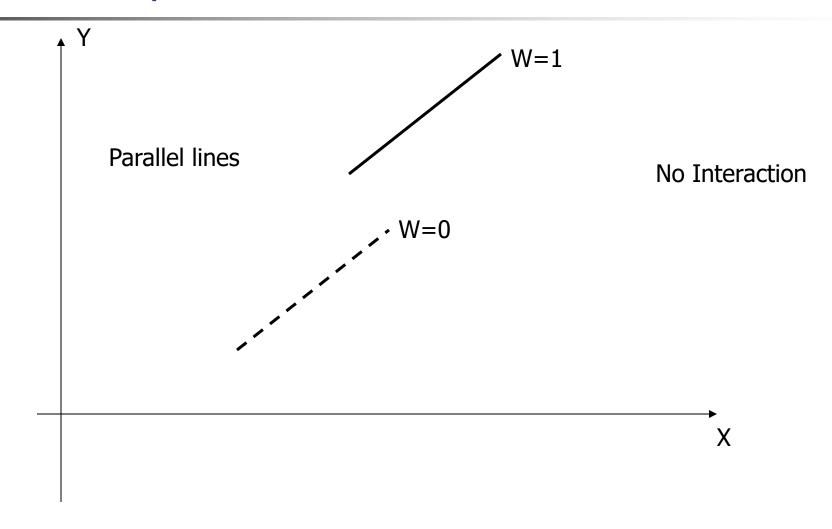


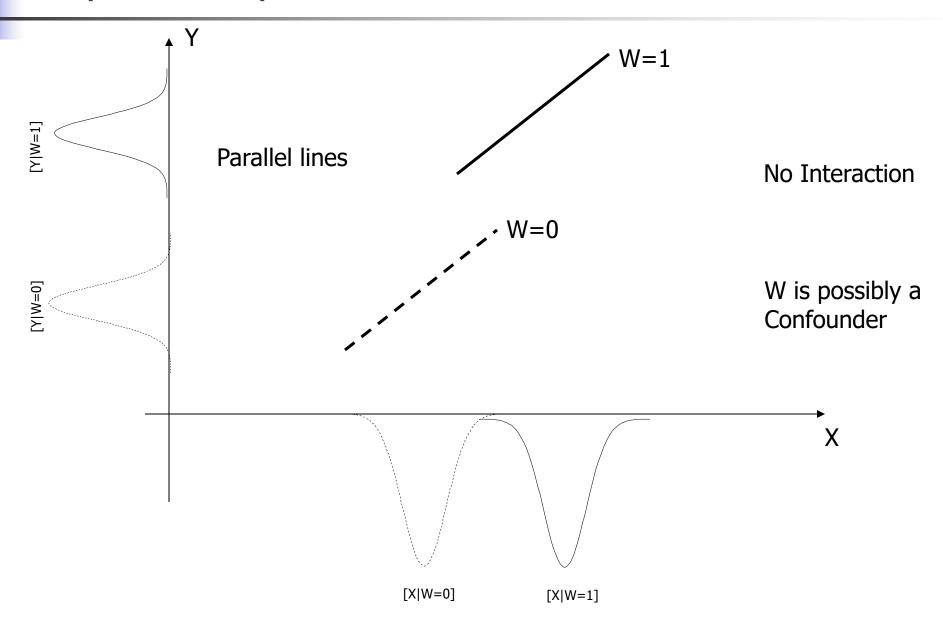


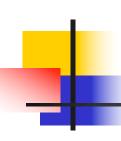












Model and Interpretation: interaction

Assume that there are two predictors in the model

$$E[Y|X_1, X_2] = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1 X_2$$

Consider two observations with the same value, c, for x_2 , but one observation has x_1 one unit higher

Obs 1:
$$E[Y|x_1=k+1, x_2=c] = \beta_0 + \beta_1 (k+1) + \beta_2 c + \beta_3 (k+1)c$$

Obs 2:
$$E[Y|x_1=k, x_2=c] = \beta_0 + \beta_1(k) + \beta_2 c + \beta_3 kc$$

Thus,
$$E[Y|x_1=k+1, x_2=c] - E[Y|x_1=k, x_2=c] = \beta_1 + \beta_3 c$$

That is, the difference in means depends now on the value of x_2 !

Model and Interpretation: interaction

• Model:
$$E[Y|x_1, x_2] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

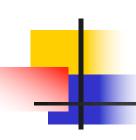
Difference in Means:

$$E[Y|x_1=k+1, x_2=c] - E[Y|x_1=k, x_2=c] = \beta_1 + \beta_3 c$$

The difference in means depends on the value of x_2

- The difference in means is β_1 if c=0.
- The difference in means is $\beta_1 + \beta_3$ if c=1
- The difference in means changes by β_3 for each unit difference in c (that is, in x_2) [that is, β_3 is the difference of differences!]

• H_0 : β_3 =0 tests for interaction



Model and Interpretation: interaction

• Model:
$$E[Y|x_1, x_2] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$$

Another way to look at this

Factor terms involving x₁:

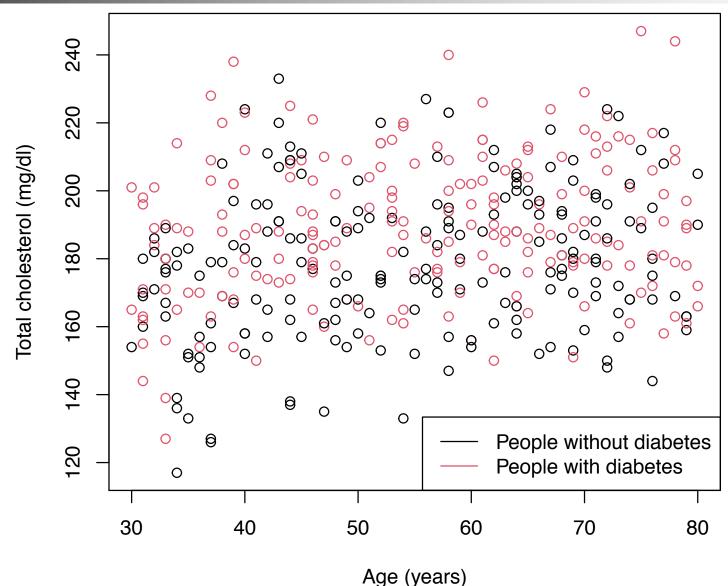
$$E[Y|x_1, x_2] = \beta_0 + (\beta_1 + \beta_3 x_2)x_1 + \beta_2 x_2$$

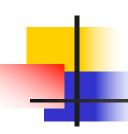
Slope of x_1 changes with x_2 , i.e.

Difference in means for each unit difference in x_1 changes with x_2 (for each one unit difference in x_2 , the difference in means changes by β_3)



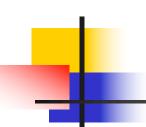
Cholesterol Example: Does diabetes affect the age – cholesterol relationship?

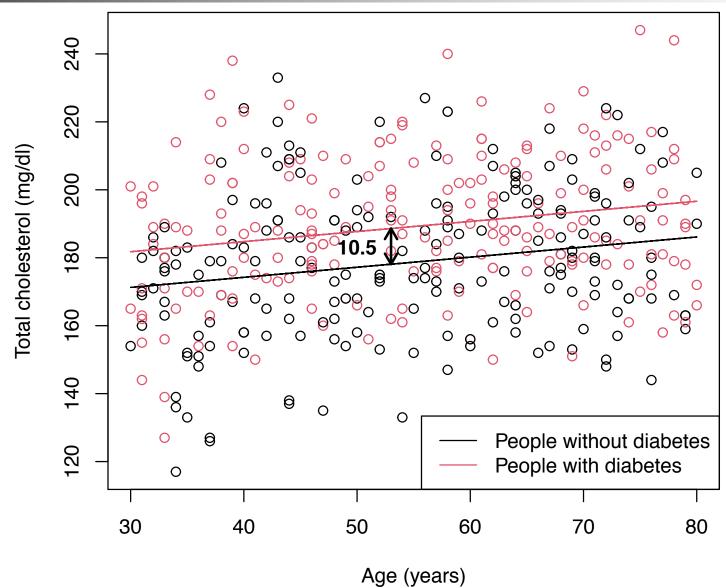


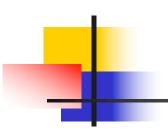


We first fit the model with age and DM terms only (No diabetes: DM=0, With diabetes: DM=1)

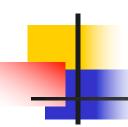
```
> fit3 = lm(chol \sim age+DM)
> summary(fit3)
Call:
lm(formula = chol \sim age + DM)
Residuals:
   Min
            10 Median 30
                                  Max
-55.662 -14.482 -1.411 14.682 57.876
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 162.35445 4.24184 38.275 < 2e-16 ***
        0.29697 0.07313 4.061 5.89e-05 ***
age
            10.50728 2.10794 4.985 9.29e-07 ***
DM
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.06 on 397 degrees of freedom
Multiple R-squared: 0.09748, Adjusted R-squared: 0.09293
F-statistic: 21.44 on 2 and 397 DF, p-value: 1.440e-09
```







- This model indicates that, after controlling for the effect of diabetes, the average cholesterol differs by 0.30 for each additional year of age
- The age effect in this model is very similar to the effect from our simple linear regression (0.31)
- However, this does not mean that the age/cholesterol relationship is the same in people with and without diabetes
- To answer this question we must add the interaction term



Model with age and DM main effects, plus interaction effect

```
> fit4=lm(chol~age*DM)
> summary(fit4)
Call:
lm(formula = chol ~ age * DM)
Residuals:
   Min 10 Median 30 Max
-56.474 -14.377 -1.215 14.764 58.301
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 160.31151 5.86268 27.344 < 2e-16 ***
   age
  14.56271 8.29802 1.755 0.08004 .
DM
age:DM -0.07399 0.14642 -0.505 0.61361
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.08 on 396 degrees of freedom
Multiple R-squared: 0.09806, Adjusted R-squared: 0.09123
F-statistic: 14.35 on 3 and 396 DF, p-value: 6.795e-09
```



```
Call:
lm(formula = chol ~ age * DM)
Residuals:
   Min
            10 Median 30
                                  Max
-56.474 -14.377 -1.215 14.764 58.301
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 160.31151
                       5.86268 27.344 < 2e-16 ***
             0.33460
                       0.10442 3.204 0.00146 **
age
DM
                      8.29802 1.755 0.08004 .
           14.56271
            -0.07399 0.14642 -0.505 0.61361
age:DM
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.08 on 396 degrees of freedom
Multiple R-squared: 0.09806, Adjusted R-squared: 0.09123
F-statistic: 14.35 on 3 and 396 DF, p-value: 6.795e-09
```

Mean cholesterol for people without diabetes (DM = 0) at age 0



```
Call:
lm(formula = chol ~ age * DM)
Residuals:
   Min
        10 Median 30
                                  Max
-56.474 -14.377 -1.215 14.764 58.301
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 160.31151 5.86268 27.344 < 2e-16 ***
             0.33460
                       0.10442 3.204 0.00146 **
            14.56271
                       8.29802 1.755 0.08004
                       0.14642 -0.505 0.61361
            -0.07399
age:DM
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.08 on 396 degrees of freedom
Multiple R-squared: 0.09806, Adjusted R-squared: 0.09123
F-statistic: 14.35 on 3 and 396 DF, p-value: 6.795e-09
```

Difference in mean cholesterol between people with and without diabetes at age 0



```
Call:
lm(formula = chol ~ age * DM)
Residuals:
    Min
            10 Median 30
                                  Max
-56.474 -14.377 -1.215 14.764 58.301
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 160.31151
                        5.86268 27.344 < 2e-16 ***
             0.33460
aσe
                       0.10442 3.204 0.00146 **
DM
            14.56271
                        8.29802 1.755 0.08004 .
                        0.14642 -0.505 0.61361
age:DM
            -0.07399
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.08 on 396 degrees of freedom
Multiple R-squared: 0.09806, Adjusted R-squared: 0.09123
F-statistic: 14.35 on 3 and 396 DF, p-value: 6.795e-09
```

Difference in mean cholesterol associated with each one year change in age for people without diabetes



```
Call:
lm(formula = chol ~ age * DM)
Residuals:
   Min
            10 Median 30
                                  Max
-56.474 -14.377 -1.215 14.764 58.301
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 160.31151 5.86268 27.344 < 2e-16 ***
             0.33460 0.10442 3.204 0.00146 **
age
DM
                        8.29802 1.755 0.08004 .
            14.56271
            -0.07399
age:DM
                        0.14642 - 0.505
                                        0.61361
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.08 on 396 degrees of freedom
Multiple R-squared: 0.09806, Adjusted R-squared: 0.09123
F-statistic: 14.35 on 3 and 396 DF, p-value: 6.795e-09
```

Difference in change in mean cholesterol associated with each one year change in age comparing people with and without diabetes



Interpretation?

Estimated model:

```
160.3 + 0.33 Age + 14.56 Diabetes - 0.07 Age x Diabetes
```

```
Subject 1: Age = a+1, diabetes = b

Subject 2: Age = a, diabetes = b

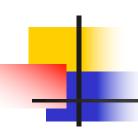
Difference in the estimated cholesterol:

[160.3 + 0.33(a+1) + 14.56(b) - 0.07 (a+1)(b)] - [160.3 + 0.33(a) + 14.56 (b) - 0.07 (a)(b)] = 0.33-0.07b
```

 Diabetes exerts a small (not statistically significant) effect on the age/cholesterol relationship

In people without diabetes: 160.3+0.33 Age

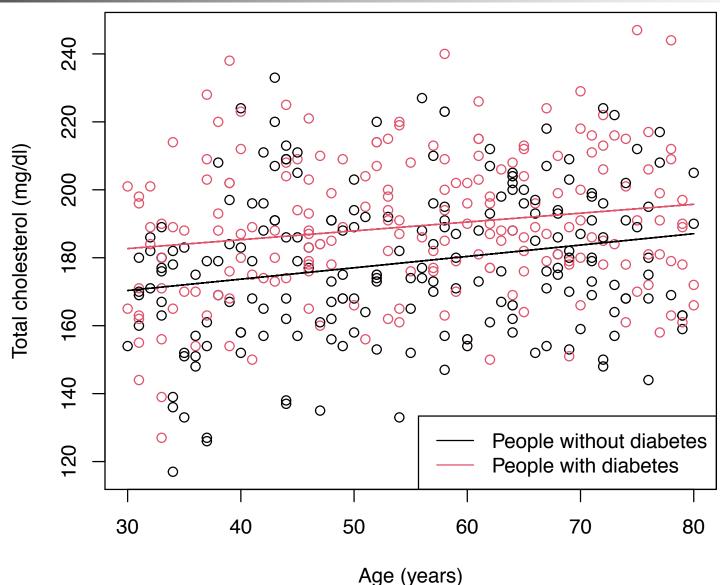
In people with diabetes: 174.9+0.26 Age

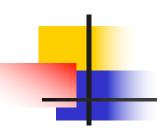


 We can also test the significance of interaction terms using an F-test

Adding the interaction term did not significantly improve model fit







Summary

We have considered:

- Simple linear regression
 - Interpretation
 - Estimation
 - Model checking
- Multiple linear regression
 - Confounding
 - Interpretation
 - Estimation
 - Interaction

Exercise

- Work on Exercise 7-8
 - Try each exercise on your own
 - Make note of any questions or difficulties you have
 - At 1:15PT we will meet as a group to go over the solutions and discuss your questions