

lec. 20

$$P(X_* | x) \propto \int_0^\infty \int_{\mathbb{R}} k(x_* | \theta, \sigma^2) k(\theta, \sigma^2 | x) d\theta d\sigma^2$$

$$= \int_0^\infty \int_{\mathbb{R}} \frac{1}{\sqrt{\sigma^2}} e^{-\frac{1}{2\sigma^2}(x_* - \theta)^2} (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)S^2/2}{\sigma^2}} e^{-\frac{n}{2\sigma^2}(\theta - \bar{x})^2} d\theta d\sigma^2$$

$$= \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{(n-1)S^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{1}{2\sigma^2}((x_* - \theta)^2 + n(\theta - \bar{x})^2)} d\theta d\sigma^2$$

$$= \int_0^\infty (\sigma^2)^{-\frac{n+1}{2}-1} e^{-\frac{(n+1)S^2/2 + x_*^2/2 + n\bar{x}^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{x_* + n\bar{x}}{\sigma^2}\theta - \frac{(n+1)}{2\sigma^2}\theta^2} d\theta d\sigma^2 \rightarrow e^{\rho\theta + \sigma^2}$$

$$= \int_0^{\infty} \sigma^{\frac{-n+1}{2}-1} \sqrt{\frac{\pi}{n+1}} \frac{e^{\left(\frac{x^*+n\bar{x}}{\sigma^2}\right)^2 / 4^2 \frac{n+1}{2\sigma^2}}}{\sigma^2}$$

↓

$$\frac{(x^*+n\bar{x})^2}{2\sigma^2(n+1)}$$

$$\int_0^{\infty} \sqrt{\frac{2\pi}{n+1}} (\sigma^2)^{\frac{-n+1}{2}-1} e^{-\frac{(n-1)s^2/2 + x_*^2/2 + n\bar{x}^2/2 - (x^*+n\bar{x})^2/(2(n+1))}{\sigma^2}} (\sigma^2)^{1/2} d\sigma^2$$

$$\int_0^{\infty} (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{\beta}{\sigma^2}} d\sigma^2 = \Gamma(\alpha) \beta^{-\alpha} = \Gamma\left(\frac{n}{2}\right) \beta^{-\alpha} \propto \beta^{-\alpha}$$

$$= \left(\frac{(n-1)s^2}{2} + \frac{x_*^2}{2} + \frac{n\bar{x}^2}{2} - \frac{(x_*+n\bar{x})^2}{2n+2} \right)^{-\frac{n}{2}} = (ax_*^2 + bx_* + c)^{-n/2}$$

$$a = \frac{1}{2} - \frac{1}{2n+2} = \frac{1}{2} \left(1 - \frac{1}{n+1} \right) = \frac{1}{2} \frac{n}{n+1}$$

$$b = -\frac{2n\bar{x}}{2n+2} = -\frac{n\bar{x}}{n+1}, \quad c = \frac{(n-1)s^2}{2} + \frac{n\bar{x}^2}{2} - \frac{n^2\bar{x}^2}{2n+2} = \frac{1}{2} \left((n-1)s^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1} \right)$$

$$= \left(\frac{1}{a} \right)^{n/2} \left(\frac{1}{a} \right)^{n/2} (ax_*^2 + bx_* + c)^{-n/2} = \left(\frac{1}{a} \right)^{n/2} \left(x_*^2 - \frac{b}{a} x_* + \frac{c}{a} \right)^{-n/2}$$

$$\propto \left(x_*^2 - \frac{b}{a} x_* + \frac{c}{a} \right)^{-n/2} = \left(\left(x_* + \frac{b}{2a} \right)^2 - \frac{b^2}{4a^2} + \frac{c}{a} \right)^{-n/2} \left(\frac{1}{\frac{c}{a} - \frac{b^2}{4a^2}} \right)^{n/2}$$

→

$$\propto \left(1 + \frac{\left(\bar{x} + \frac{b}{2a} \right)^2}{\frac{c}{a} - \frac{b^2}{4a^2}} \right)^{-\frac{(n-1)+1}{2}} = \left(1 + \frac{\frac{1}{n+1} \left(\bar{x} - \frac{b}{2a} \right)^2}{\left(\frac{c}{a} - \frac{b^2}{4a^2} \right) (n-1)} \right)^{-\frac{(n-1)+1}{2}}$$

$\mu \uparrow$
 S

$$\propto T_{\sigma}(\mu, s^2) = T_{n-1}(\bar{X}, \frac{n+1}{n} s^2) \approx N(\bar{X}, s^2)$$

$$= \frac{-b}{2a} = \frac{n\bar{x}}{\frac{n}{n+1}} = \bar{x}$$

$$\frac{c}{a} - \frac{b^2}{4a^2} = \frac{\frac{1}{n+1} \left((n-1)s^2 + n\bar{x}^2 - \frac{n^2 \bar{x}^2}{n+1} \right)}{\frac{n}{n+1}} =$$

$$= \frac{(n-1)(n+1)}{n} s^2 + \underbrace{(n+1)\bar{x}^2 - n\bar{x}^2}_{\downarrow n\bar{x}^2 + \bar{x}^2}$$

$$= \frac{-b^2}{4a^2} = - \left(\frac{b}{2a} \right)^2 = - \left(\frac{b}{2a} \right)^2 = - \bar{x}^2 \Rightarrow \frac{c}{a} - \frac{b^2}{4a^2}$$

$$= \frac{(n-1)(n+1)}{n} s^2$$



$$S_0^2 = \frac{(n-1)(n+1)s^2}{n} = \frac{n+1}{n} s^2$$

under Jeffrey prior...

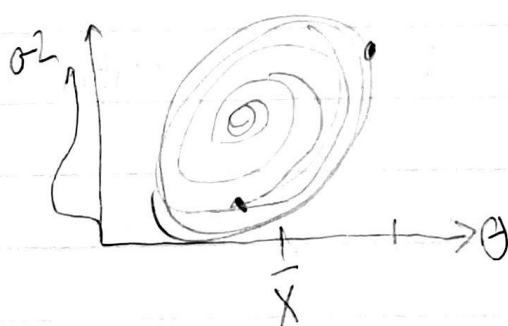
$$P(\theta, \sigma^2 | x) = P(\theta | x, \sigma^2) P(\sigma^2 | x) \quad \text{By def of conditional probability}$$

$$\text{NormInvGamma}(\cdot, \cdot, \cdot)$$

$$N(\bar{x}, \frac{\sigma^2}{n})$$

$$\text{InvGamma}(\frac{n-1}{2}, \frac{(n-1)s^2}{2})$$

You can think of a normal inverse gamma as first sampling from an $\text{InvGamma}(\frac{n-1}{2}, \frac{(n-1)s^2}{2})$ to get a σ^2 value, and then you use that value σ^2 to draw a θ from $N(\bar{x}, \frac{\sigma^2}{n})$ and return the two-dimension point $[\theta]$



$$P(\theta, \sigma^2 | x)$$

This can also be done the other way...

$$P(\theta, \sigma^2 | x) = P(\sigma^2 | x, \theta) P(\theta | x)$$

\downarrow \rightarrow
 $\text{InvGamma}\left(\frac{n}{2}, \frac{n \bar{y}^2}{2}\right)$ $T_{n-1}(\bar{x}, \frac{s^2}{2})$

If we decompose the first way, we draw θ from $N(\bar{x}, \frac{\sigma^2}{n})$ and thus σ^2 must be known.

What if we break this by instead of using the Jeffreys prior, use

$$P(\theta) = N(\mu_0, \tau^2) \text{ and } P(\sigma^2) = \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right)$$

These were the two priors we began with when we started investigating the normal likelihood model.

However, it's important to note we are not allowing $z^2 = \frac{\sigma^2}{n_0}$

what happens?

$$P(\theta, \sigma^2) = P(\theta)P(\sigma^2) \text{ not } P(\theta | \sigma^2)P(\sigma^2)$$

= The two priors are disconnected

Derive the posterior under this two-dim prior

$$P(\theta, \sigma^2 | x) \propto P(x | \theta, \sigma^2) P(\theta, \sigma^2) = P(x | \theta, \sigma^2) P(\theta) P(\sigma^2)$$

$$\propto k(x | \theta, \sigma^2) k(\theta) k(\sigma^2)$$

$$= (\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n(\bar{x} - \theta)^2)} e^{-\frac{1}{2\sigma^2}(\theta - \mu_0)^2} (\theta^2)^{-\frac{n_0+1}{2}} e^{-\frac{n_0+1}{2\sigma_0^2}}$$

$$= (\sigma^2)^{-\frac{n}{2} - \frac{n_0+1}{2} - 1} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n_0\sigma_0^2 + n\bar{x}^2)} e^{\left(\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{\sigma^2}\right)\theta - \left(\frac{n}{2\sigma^2} + \frac{1}{2\sigma_0^2}\right)\theta^2}$$