

Mach 17

Iec 1

$$X \sim \text{Ber}(\theta), \theta = P(X=1)$$

There is another way to "parameterize" the Bernoulli.



1-1 function

consider

$$\phi = t(\theta) = \frac{\theta}{1-\theta}, \theta \in [0, \infty) \Rightarrow \theta - \theta\phi = \theta$$

$$\phi = \theta + \theta\phi$$

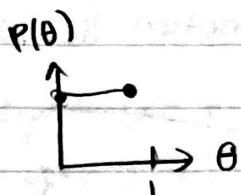
$$\phi = \theta(1+\theta)$$

$$\theta = t^{-1}(\phi) = \frac{\phi}{1+\phi}$$

"odds"

$$\text{Laplace } P(\theta) = U(0,1)$$

$p(t) \stackrel{?}{=} \text{uniform}$



No, it is impossible to have a prior on the support $(0, \infty)$

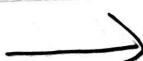
$$\text{If } x \sim \text{Ber}(\phi) = \left(\frac{\phi}{1+\phi}\right)^x \left(\frac{1}{1+\phi}\right)^{1-x} = \frac{\phi^x}{1+\phi}$$

$$P_\phi(x) = P_\theta(t^{-1}(\theta)) \left| \frac{d}{d\theta}[t^{-1}(\theta)] \right| = P_\theta\left(\frac{\theta}{1+\theta}\right) \left| \frac{d}{d\theta}\left[\frac{\theta}{1+\theta}\right] \right|$$

= 1 \cdot \left| \frac{d}{d\theta}\left[\frac{\theta}{1+\theta}\right] \right| \quad \text{quotient rule}

$$= \left| \frac{1+\theta - \theta(1)}{(1+\theta)^2} \right| = \frac{1}{1+\theta^2}$$

Fisher
Shepard
distribution



$p(\theta)$

not uniform

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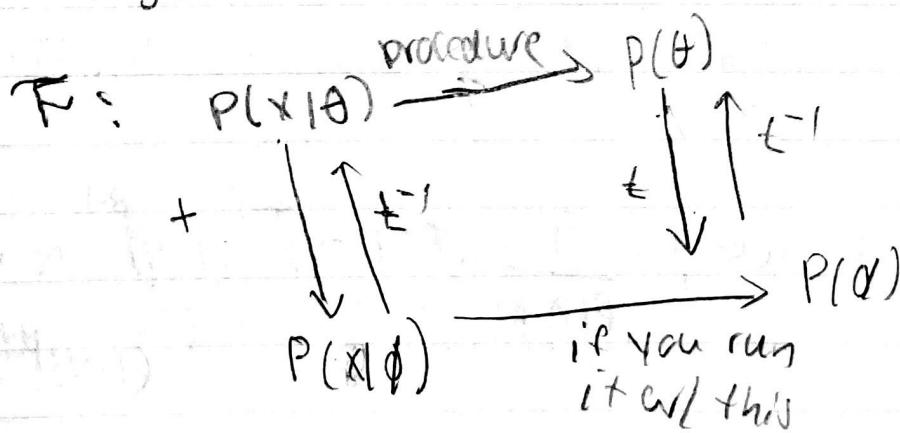
Is this a valid density?

$$\int_0^\infty \frac{1}{(1+\phi)^2} d\phi$$

$$= \left[\frac{\phi}{1+\phi} \right]_0^\infty = 1 - 0 = 1$$

what did we prove? We proved that if you're indifferent on the probability scale then you're *not* indifferent on the odds scale. Fisher used this example to show how stupid Laplace's prior and to further show how stupid Bayesian stats is in general. If you change the parametrization, yes the inference can change.

Can we address this problem in parts? Can we do something? Can this something pick a prior for us? Consider the following. Let θ be the parameter of curly F and $t(\theta) = \phi$, a 1:1 reparametrization. Is there a procedure that can accomplish the following?



It was Harold Jeffreys' idea that solved this puzzle. The prior that is the result of the procedure is then called the "Jeffreys prior". In order to derive the procedure, we need two more tools:

(1) Kernel (2) Fisher Information

$$\text{Kernel } f(x; \theta) \propto k(x; \theta) \Rightarrow \exists c \in \mathbb{R} \quad f(x; \theta) = ck(x; \theta)$$

↑
there exist

This also means k and f are 1-1 because they differ only by c . This is also valid for PMF as well but I'll use the f notation.

$$\int f(x; \theta) dx = 1 \Rightarrow \int ck(x; \theta) dx = 1 \Rightarrow k(x; \theta) dx = \frac{1}{c}$$

↓

$$C = \frac{1}{\int k(x; \theta) dx}$$

Supp[X]

$$Y \sim \text{Beta}(\alpha, \beta) = \frac{1}{B(\alpha, \beta)} y^{\alpha-1} (1-y)^{\beta-1}$$

$$(1-y)^{\beta-1} = k(y; \alpha, \beta)$$

$F: \text{Bin}(n, \theta)$, n fixed, $P(\theta) = \text{Beta}(\alpha, \beta)$

$$P(\theta|x) = \frac{P(x|\theta) P(\theta)}{P(x)} \propto P(x|\theta) P(\theta) =$$

$$\binom{n}{x} \theta^x (1-\theta)^{n-x} \frac{1}{\Gamma(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

$$\propto \theta^x (1-\theta)^{n-x} \theta^{\alpha-1} (1-\theta)^{\beta-1} = \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1} \text{Beta}(nx, n-x+\beta)$$

Ex. of kernel

$$Y \sim N(\theta, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y-\theta)^2} \propto e^{-\frac{1}{2\sigma^2}(y-\theta)^2}$$

$$= e^{-\frac{1}{2\sigma^2}(y^2 - 2y\theta + \theta^2)}$$

$$= e^{-\frac{y}{2\sigma^2}} e^{\frac{y\theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}} \propto e^{-\frac{y\theta}{2\sigma^2}} e^{\frac{\theta^2}{2\sigma^2}}$$

$$= K(y; \theta)$$

$$C = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}}$$

Fisher Information $X < x_1, \dots, x_n >$

$$\text{Recall } \mathcal{L}(\theta; x) = P(x; \theta)$$

↓

$$l(\theta; x) := \ln(\mathcal{L}(\theta; x))$$

$$S(\theta; x) = \frac{d}{d\theta} [l(\theta; x)] \rightarrow$$

$S(\theta; x) = 0$ solve for θ is equal to

$\hat{\theta}_{MLE}$ is derived by:

$$I(\theta) = \text{Var}_x [S(\theta; x)] = E_x - \ell''(\theta; x)$$

This is Fisher Information.

An example Fisher Information calculation: $X \sim \text{Bern}(\theta)$

$$\ell(\theta; x) = \theta^x (1-\theta)^{1-x} \Rightarrow \ell'(\theta; x) = x \ln(\theta) + (1-x) \ln(1-\theta)$$

$$\ell''(\theta; x) = \frac{x}{\theta} - \frac{1-x}{1-\theta} = \ell''(\theta; x) = -\frac{x}{\theta^2} - \frac{1-x}{(1-\theta)^2}$$

$$I(\theta) = E_x [-\ell''] = E_x \left[\frac{x}{\theta^2} + \frac{1-x}{(1-\theta)^2} \right] =$$

$$= \frac{1}{\theta^2} E[X] + \frac{1}{(1-\theta)^2} (1 - E[X])$$

$$= \frac{1}{\theta^2} \cancel{\theta} + \frac{1}{(1-\theta)^2} \cancel{(1-\theta)} = \frac{1}{\theta} + \frac{1}{1-\theta} = \frac{1}{\theta(1-\theta)}$$

Thm: the Jeffreys prior $P_J(\theta) \propto \sqrt{I(\theta)}$



Let's see this works first and then provide a proof for the thm later.

$$F: \text{Bin}(n, \theta) = L(\theta; x) = \binom{n}{x} \theta^x (1-\theta)^{n-x} = l(\theta; x) =$$

$$= \ln \binom{n}{x} + x \ln \theta + (n-x) \ln (1-\theta)$$

$$\Rightarrow l'(\theta; x) = \frac{x}{\theta} - \frac{n-x}{1-\theta} = l''(\theta; x) = -\frac{x}{\theta^2} - \frac{n-x}{(1-\theta)^2}$$

$$I(\theta) = E[-l] = E \left[\frac{x}{\theta^2} + \frac{n-x}{(1-\theta)^2} \right] = \frac{1}{\theta^2} E[x] + \frac{1}{(1-\theta)^2} (n - E[x])$$

$$= \frac{1}{\theta^2} + n\theta + \frac{1}{(1-\theta)^2} (n/n\theta)$$

$$= n \left(\frac{1}{\theta} + \frac{1}{1-\theta} \right)$$

$$= \frac{n}{\theta(1-\theta)}$$

$$P_J(\theta) \propto \sqrt{\frac{n}{\theta(1-\theta)}} \propto \sqrt{\frac{1}{\theta(1-\theta)}} = \theta^{-1/2} (1-\theta)^{-1/2} \underset{\alpha=1}{\overset{\beta=1}{\sim}} \text{Beta}(\frac{1}{2}, \frac{1}{2})$$

The Jeffrey's prior is $\text{Beta}(\frac{1}{2}, \frac{1}{2})$! It's amazing that it came out conjugate who knows what could've happened?

Jeffrey's procedure

$$P(x|\theta) \xrightarrow{\text{Jeffrey's procedure}} P_J(\theta) = \text{Beta}(1/2, 1/2)$$

#

$$\downarrow \quad \uparrow t \quad \downarrow \quad \uparrow t^{-1}$$

Jeffrey's procedure

$$P(x|\phi) \xrightarrow{\text{Jeffrey's procedure}} P_J(\phi) = ?$$

We will verify this using $\phi = t|\theta = \frac{\theta}{1-\theta}$, the "odds",

But just b/c it works once, doesn't mean we've
proven the thm in the need to prove it...

midterm 1 ↑