

Maximum Likelihood Estimation

Session 11

05/04/2021

Methods of Estimation

- So far we have been discussing the requisites of a good estimator. Now we shall briefly outline some (will discuss only one in detail here) of the important methods for obtaining such estimators. Commonly used methods are
 - (I) *Method of Maximum Likelihood Estimation.*
 - (ii) *Method of Minimum Variance.*
 - (iii) *Method of Moments.*
 - (iv) *Method of Least Squares.*
 - (v) *Method of Minimum Chi-square*

Method of Maximum Likelihood Estimation

- From theoretical point of view, the most general method of estimation known is the method of Maximum Likelihood Estimation which was initially formulated by C.F.Gauss but as a general method of estimation was first introduced by Prof.R.A.Fisher.

Likelihood function(Definition)

- Let x_1, x_2, \dots, x_n be a random sample of size n from a population with density function $f(x, \theta)$. Then the likelihood function of the sample values x_1, x_2, \dots, x_n usually denoted by $L = L(\theta)$ is their joint density function given by

$$\begin{aligned} L &= f(x_1, \theta) f(x_2, \theta) \dots f(x_n, \theta) \\ &= \prod_{i=1}^n f(x_i, \theta) \end{aligned}$$

L gives the relative likelihood that the random variables assume a particular set of values x_1, x_2, \dots, x_n

- The principle of maximum likelihood consist in finding an estimator for the unknown parameter $\theta = (\theta_1, \theta_2, \dots, \theta_k)$, say which maximizes the likelihood function $L(\theta)$ for variations in parameter.
- i.e., we wish to find $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_n)$ so that
i.e., $L(\hat{\theta}) = \text{Sup}L(\theta)$ for all $\theta \in \Theta$
- Thus if there exist a function $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_n)$ of sample values which maximizes L for variations in θ , then $\hat{\theta}$ is to be taken as an estimator of θ and $\hat{\theta}$ is usually called M.L.E.

- Thus $\hat{\theta}$ is the solution, if any, of

$$\frac{\partial L}{\partial \theta} = 0 \quad \text{and} \quad \frac{\partial^2 L}{\partial \theta^2} < 0$$

Since $L > 0$, and $\log L$ is a non-decreasing function of L ; L and $\log L$ attain their extreme values (maxima or minima) at the same value of $\hat{\theta}$.

- Thus,

$$\frac{1}{L} \cdot \frac{\partial L}{\partial \theta} = 0 \quad \Rightarrow \quad \frac{\partial \log L}{\partial \theta} = 0,$$

a form which is much more convenient from practical point of view.

- Thus from the practical point of view, we find the following likelihood equation for estimating the parameters.

$$\frac{\partial \log L}{\partial \theta} = 0$$

- If θ is a vector valued parameter, then $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_n)$, is given by the solution of simultaneous equations

$$\frac{\partial}{\partial \theta_i} \log L = \frac{\partial}{\partial \theta_i} \log L(\theta_1, \theta_2, \dots, \theta_n) = 0; i = 1, 2, \dots, n$$

Example 1

- In random sampling from a normal population $N(\mu, \sigma^2)$, find the maximum likelihood estimators for
- (i) μ when σ^2 is known
- (ii) σ^2 , when μ is known
- (iii) the simultaneous estimation for μ and σ^2 .

Solution. $X \sim N(\mu, \sigma^2)$ then

$$\begin{aligned} L &= \prod_{i=1}^n \left[\frac{1}{\sigma\sqrt{2\pi}} \exp \left\{ -\frac{1}{2\sigma^2} (x_i - \mu)^2 \right\} \right] \\ &= \left(\frac{1}{\sigma\sqrt{2\pi}} \right)^n \exp \left\{ -\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2} \right\} \end{aligned}$$

$$\log L = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

Case (i). When σ^2 is known, the likelihood equation for estimating μ is

$$\frac{\partial}{\partial \mu} \log L = 0 \Rightarrow -\frac{1}{2\sigma^2} \sum_{i=1}^n 2(x_i - \mu)(-1) = 0$$

or
$$\sum_{i=1}^n (x_i - \mu) = 0 \Rightarrow \sum_{i=1}^n x_i - n\mu = 0$$

$$\Rightarrow \hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x} \quad \dots(*)$$

Hence M.L.E. for μ is the sample mean \bar{x} .

Case (ii). When μ is known, the likelihood equation for estimating σ^2 is

$$\frac{\partial}{\partial \sigma^2} \log L = 0 \Rightarrow -\frac{n}{2} \times \frac{1}{\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (x_i - \mu)^2 = 0$$

$$\Rightarrow n - \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 = 0, \quad i.e., \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 \quad \dots(**)$$

Case (iii). The likelihood equations for simultaneous estimation of μ and σ^2 are

$$\frac{\partial}{\partial \mu} \log L = 0 \quad \text{and} \quad \frac{\partial}{\partial \sigma^2} \log L = 0, \text{ thus giving}$$

$$\hat{\mu} = \bar{x} \quad \text{[From (*)]}$$

and
$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2 \quad \text{[From (**)]}$$

$$= \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 = s^2, \text{ the sample variance.}$$

- **IMPORTANT :** It may be noted that MLE need not be necessarily unbiased

Remark. Since M.L.E. is the most efficient, we conclude that in sampling from a normal population, the sample mean \bar{x} is the most efficient estimator of the population mean μ .

Example 2

- Find the M.L.E. for the parameter λ of a Poisson distribution on the basis of a random sample of size n .

Solution. The probability function of the Poisson distribution with parameter λ is given by

$$P(X = x) = f(x, \lambda) = \frac{e^{-\lambda} \lambda^x}{x!}; x = 0, 1, 2, \dots$$

Likelihood function of random sample x_1, x_2, \dots, x_n of n observations from this population is

$$L = \prod_{i=1}^n f(x_i, \lambda) = \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{x_1! x_2! \dots x_n!}$$

$$\begin{aligned}\therefore \log L &= -n\lambda + \left(\sum_{i=1}^n x_i \right) \log \lambda - \sum_{i=1}^n \log (x_i !) \\ &= -n\lambda + n\bar{x} \log \lambda - \sum_{i=1}^n \log (x_i !)\end{aligned}$$

The likelihood equation for estimating λ is

$$\frac{\partial}{\partial \lambda} \log L = 0 \quad \Rightarrow \quad -n + \frac{n\bar{x}}{\lambda} = 0 \quad \Rightarrow \quad \lambda = \bar{x}$$

Thus the M.L.E. for λ is the sample mean \bar{x} .