

**Exercise-1**

We perform best subset, forward stepwise, and backward stepwise selection on a single data set. For each approach, we obtain  $p + 1$  models, containing  $0, 1, 2, \dots, p$  predictors. Explain your answers:

**a) Which of the three models with  $k$  predictors has the smallest training RSS?**

Answer is best using  $p$  predictors so there is no subset model has smallest training RSS as it trains all possible  $2p$  models that couldn't be identified by best subset selection.

**b) Which of the three models with  $k$  predictors has the smallest test RSS?**

As best subset model checks more combinations of predictors so it should result in smallest test RSS, but it is also possible to overfit with any of these three methods. It might be possible that stepwise models perform better on test set compared to best subset. Hence its difficult to decide.

**c) True or False**

- i. True  $K+1$  is obtained by adding one more variable from  $k$  variable model.
- ii. True, same reason we can also obtain by removing 1 variable from  $K+1$  model
- iii. False,
- iv. False
- v. False, may be possible but not all times.

**Exercise-2****a) indicate which of i. through iv. is correct. Justify your answer**

The option III is correct as Lasso is more restrictive model, so it increases the bias, it has the chance of less overfitting to the data and reducing the variance.

**b) Repeat (a) for ridge regression relative to least squares**

The option III is correct as ridge regression is less flexible as it reduces predictors which are not associated with target variable. (Lasso and ridge both are same except shape of constraints)

**c) Repeat (a) for non-linear methods relative to least squares.**

The option II is correct as we make less assumption about function  $f$  which result in reduced bias that is higher than increase in variance hence improve prediction accuracy.

**Exercise-3****a) As we increase  $s$  from 0, the training RSS will**

The option IV is correct as we increase  $s$  the model flexibility increases and reduce restriction on beta which results in steady decrease in the training RSS.

**b) Repeat (a) for test RSS**

The option II is correct as we increase  $s$  model become more flexible as result test RSS decrease initially until increase in variance is less than reduction in bias, but at some point, it starts gradually increase and give the U shape curve.

**c) Repeat (a) for variance**

The option **III** is correct as we increase s, model become more flexible which result on increase in variance.

**d) Repeat (a) for (squared) bias**

The option **IV** is correct as we increase s, the model flexibility increases and bias of model decrease.

**e) Repeat (a) for the irreducible error**

The option **V** is correct as irreducible error does not depend on flexibility of the model and not depend on s value.

**Exercise-4**

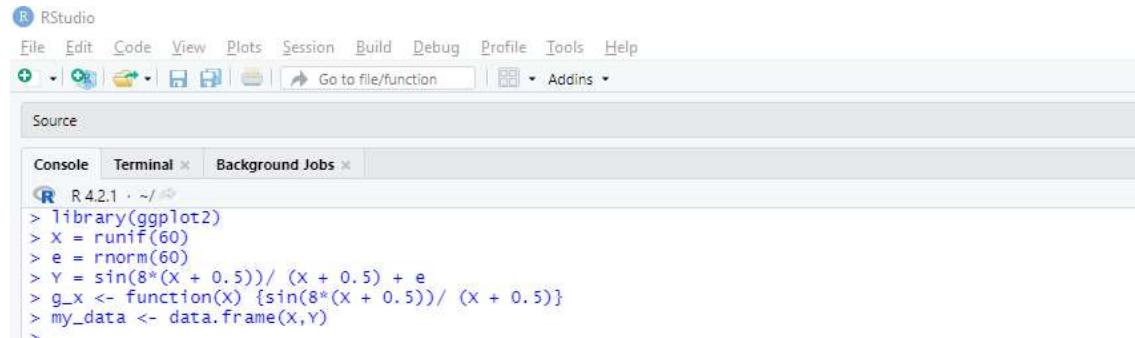
- a) The option **III** is correct as we increase lambda implied decrease flexibility of model, as with increase in lambda we restrict the coefficients which increase in the training RSS.
- b) The option **II** is correct as we increase lambda model become less flexible which at first reduce in the test RSS then start increasing again with overfitting of model and yield U shape curve.
- c) The option **IV** is correct as we increase lambda model become less flexible and variance always decrease as flexibility of model decrease.
- d) The option **III** is correct we increase lambda model become less flexible and more constrained hence results steady increase bias.
- e) The option **V** is correct as irreducible error does not depend on flexibility of the model and not depend on lambda value.

**Exercise-5**

## Chapter-7

### Exercise-2

Provide example sketches of  $\hat{g}$  in each of the following scenarios

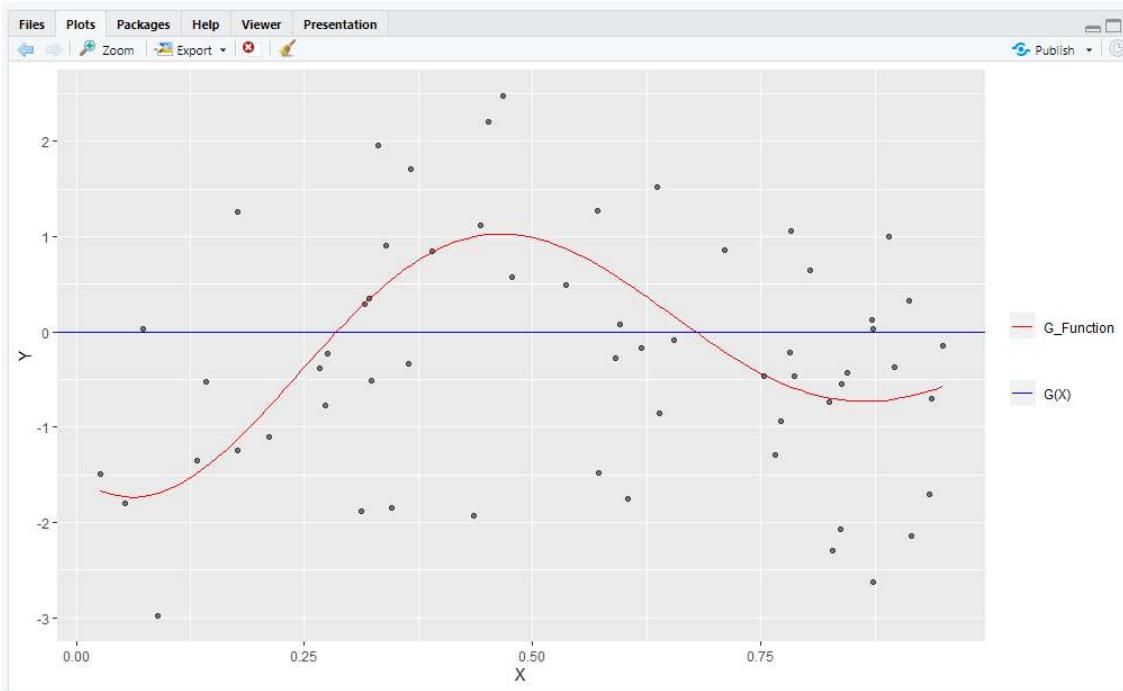


```
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R 4.2.1 · ~/Desktop
> library(ggplot2)
> X = runif(60)
> e = rnorm(60)
> Y = sin(8*(X + 0.5)) / (X + 0.5) + e
> g_x <- function(x) {sin(8*x + 0.5)} / (x + 0.5)
> my_data <- data.frame(X,Y)
```

(a)  $\lambda = \infty, m = 0$

$g = 0$  as large smoothing parameter results in  $g(0) (x) \rightarrow 0$

```
> ggplot(my_data,aes(x=x, y=y))+ geom_point(alpha = 0.5)+ stat_function(fun = g_x, aes(col = "G_Function"))+ geom_hline(aes(yintercept = 0, linetype = "G(X)"),col="blue") + scale_colour_manual(values = "red") + theme(legend.position = "right",legend.title = element_blank())
> |
```



(b)  $\lambda = \infty, m = 1$ .

$g = c$  because large smoothing parameter results in  $g(1) (x) \rightarrow 0$

RStudio

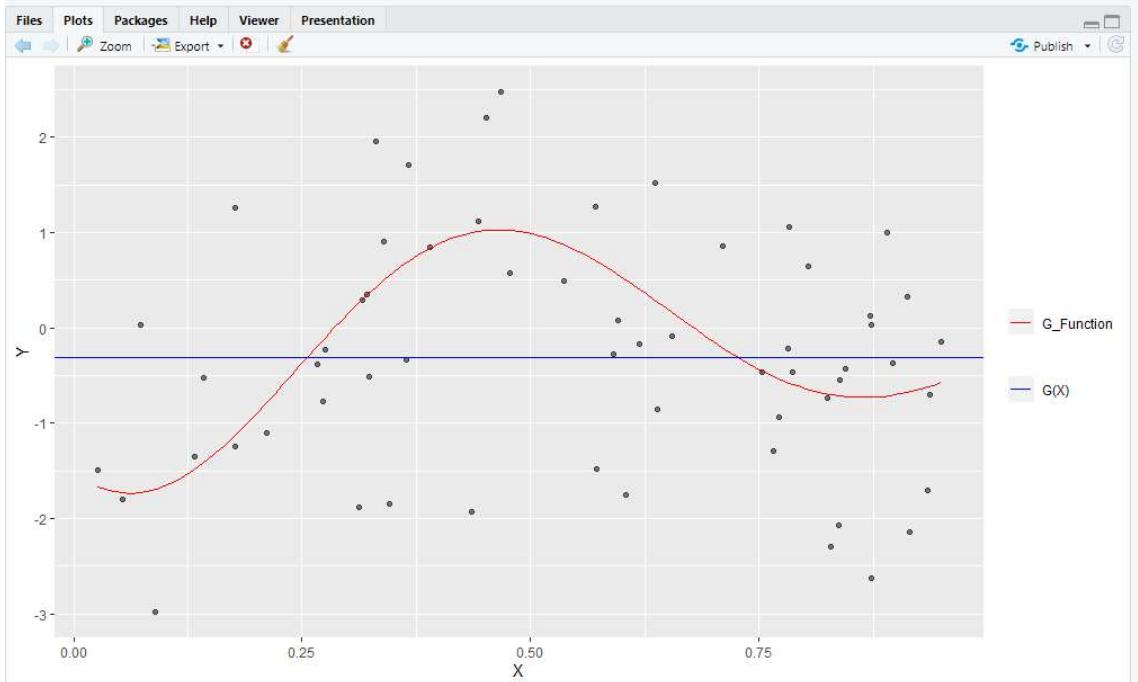
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```
R 4.2.1 - ~/ ~
> ggplot(my_data,aes(x=x, y=Y))+ geom_point(alpha = 0.5)+ stat_function(fun = g_x, aes(col = "G_Function"))+ geom_hline(aes(yintercept = mean(Y), linetype = "G(x)", col="blue") + scale_colour_manual(values = "red") + theme(legend.position = "right",legend.title = element_blank())
> |
```



(c)  $\lambda = \infty, m = 2$ .

$g = cx + d$  because large smoothing parameter results in  $g(2)(x) \rightarrow 0$

RStudio

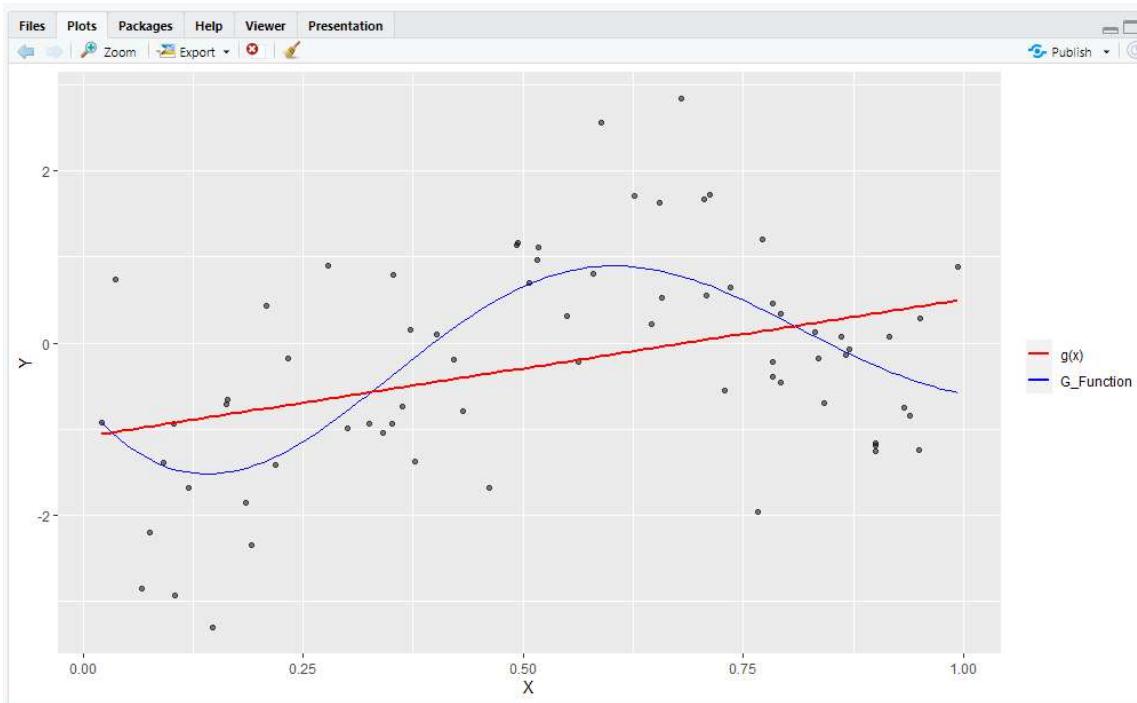
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Console Terminal Background Jobs

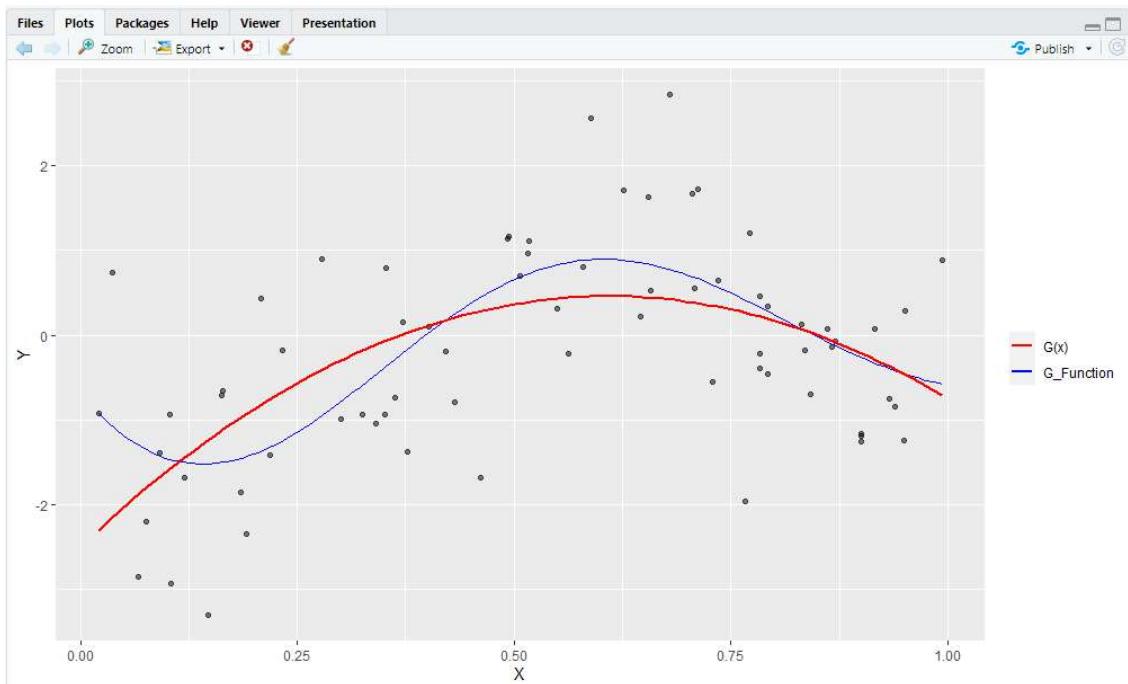
```
R 4.2.1 - ~/ ~
> ggplot(my_data,aes(x=x, y=Y))+ geom_point(alpha = 0.5)+ stat_function(fun = g_x, aes(col = "G_Function"))+ geom_smooth(method = "lm", formula = "y ~ x", se = F,size=0.8, aes(col = "g(x)")) + scale_colour_manual(values = c("red","blue")) + theme(legend.position = "right",legend.title = element_blank())
> |
```



(d)  $\lambda = \infty, m = 3$

$g = cx^2 + dx + e$  because large smoothing parameter results in  $g(3)(x) \rightarrow 0$

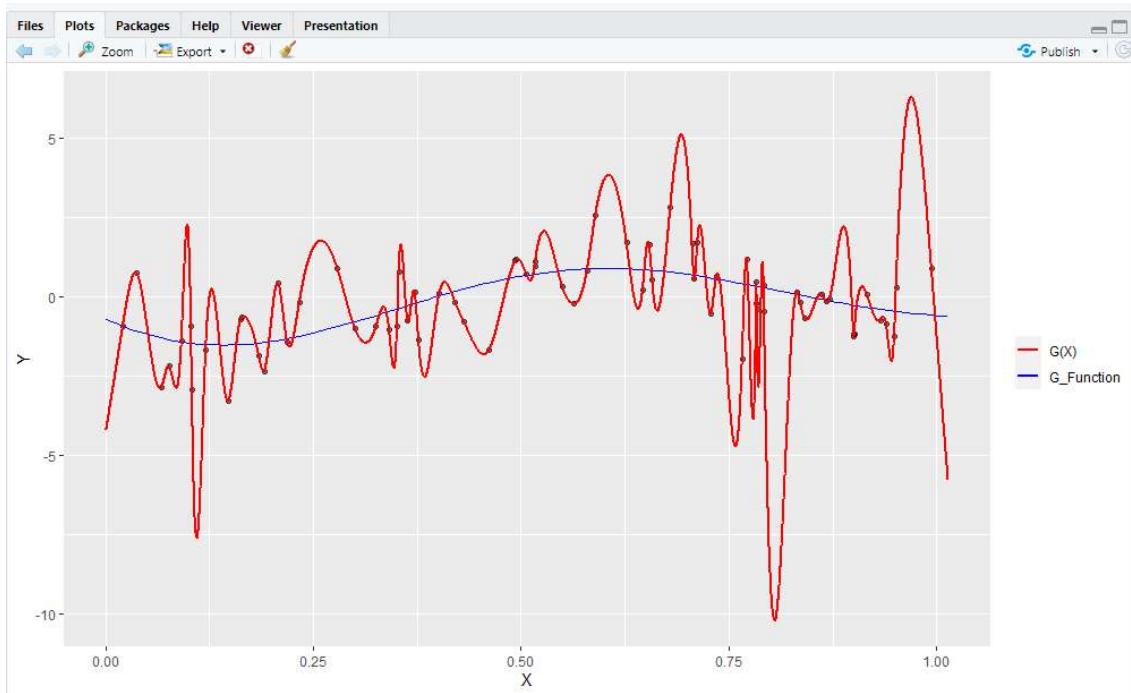
```
R 4.2.1 - ~/r
> ggplot(my_data,aes(x=x, y=Y))+ geom_point(alpha = 0.5)+stat_function(fun = g_x, aes(col = "G_Function"))+ geom_smooth(method = "lm", formula = "y ~ I(x^2)+x", se = F,size=0.8, aes(col = "G(x)")) + scale_colour_manual(values = c("red","blue")) + theme(legend.position = "right",legend.title = element_blank())
>
> |
```



(e)  $\lambda = 0, m = 3$

Here  $g$  is interpolating spline. As an example, it fits low degree polynomials to the small subset of values

```
R 4.2.1 · ~/r
> interp_spline <- smooth.spline(x = my_data$x, y = my_data$y, all.knots = T, lambda = 0.0000000000001)
> fitted <- predict(interp_spline, x = seq(min(X) - 0.02, max(X) + 0.02, by = 0.001))
>
> fitted <- data.frame(x = fitted$x, fitted_y = fitted$y)
> ggplot(my_data,aes(x=x, y=y))+ geom_point(alpha = 0.5)+stat_function(fun = g_x, aes(col = "G_Function"))+geom_line(data = fitted, aes(x = x, y = fitted_y, col = "G(X)", size = 0.8) + scale_colour_manual(values = c("red", "blue")) + theme(legend.position = "right", legend.title = element_blank())
>
> |
```



### Exercise-3

Suppose we fit a curve with basic functions  $b_1(X) = X$ ,  $b_2(X) = (X - 1)^2 I(X \geq 1)$ . (Note that  $I(X \geq 1)$  equals 1 for  $X \geq 1$  and 0 otherwise.) We fit the linear regression model  $Y = \beta_0 + \beta_1 b_1(X) + \beta_2 b_2(X) + \epsilon$ , and obtain coefficient estimates  $\hat{\beta}_0 = 1$ ,  $\hat{\beta}_1 = 1$ ,  $\hat{\beta}_2 = -2$ . Sketch the estimated curve between  $X = -2$  and  $X = 2$ .

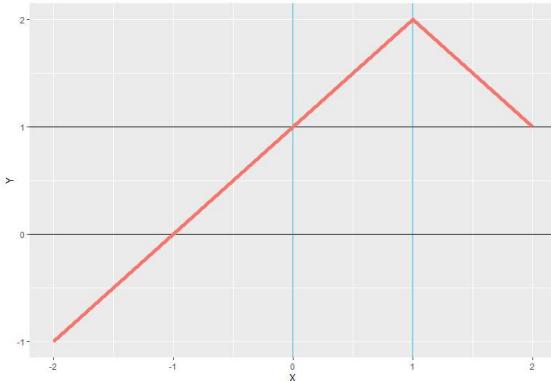
### Answer:

Range of X is -2 :2

Value of Y from given Equation  $Y = \beta_0 + \beta_1 b_1(X) + \beta_2 b_2(X) + \epsilon$

$$Y = 1 + 1b_1(x) - 2b_2(x) \Rightarrow Y = 1 + x - 2(x-1)^2 * I(x \geq 1)$$

```
R 4.2.1 --> 
> x<-seq(-2,2)
> y<-1 + x - 2 * (x-1)^2 * (x >= 1)
> data_frame<-data.frame(x,y)
> ggplot(data_frame,aes(x=x,y=y,col="red"))+geom_vline(xintercept = 0,col="sky blue",size=1)+geom_vline(xintercept = 1,col="sky blue",size=1)+geom_hline(yintercept = 0)+geom_hline(yintercept = 1)+geom_line(size=2)
> |
```



The given curve is:

Linear between  $-2$  and  $1$

$$Y=1+x$$

Quadratic between  $1$  and  $2$

$$Y=1+x-2(x-1)^2.$$

For  $x < 1$ ,  $Y = 1 + x$  so intersects  $= 1$  and slope  $= 1$

For  $x \geq 1$   $Y = 1 + x - 2(x-1)^2$  slope will vary

#### Exercise-4

Suppose we fit a curve with basic functions  $b_1(X) = I(0 \leq X \leq 2) - (X-1)I(1 \leq X \leq 2)$ ,  $b_2(X) = (X-3)I(3 \leq X \leq 4) + I(4 < X \leq 5)$ . We fit the linear regression model  $Y = \beta_0 + \beta_1 b_1(X) + \beta_2 b_2(X) + \epsilon$ , and obtain coefficient estimates  $\hat{\beta}_0 = 1$ ,  $\hat{\beta}_1 = 1$ ,  $\hat{\beta}_2 = 3$ . Sketch the estimated curve between  $X = -2$  and  $X = 6$ . Note the intercepts, slopes, and other relevant information.

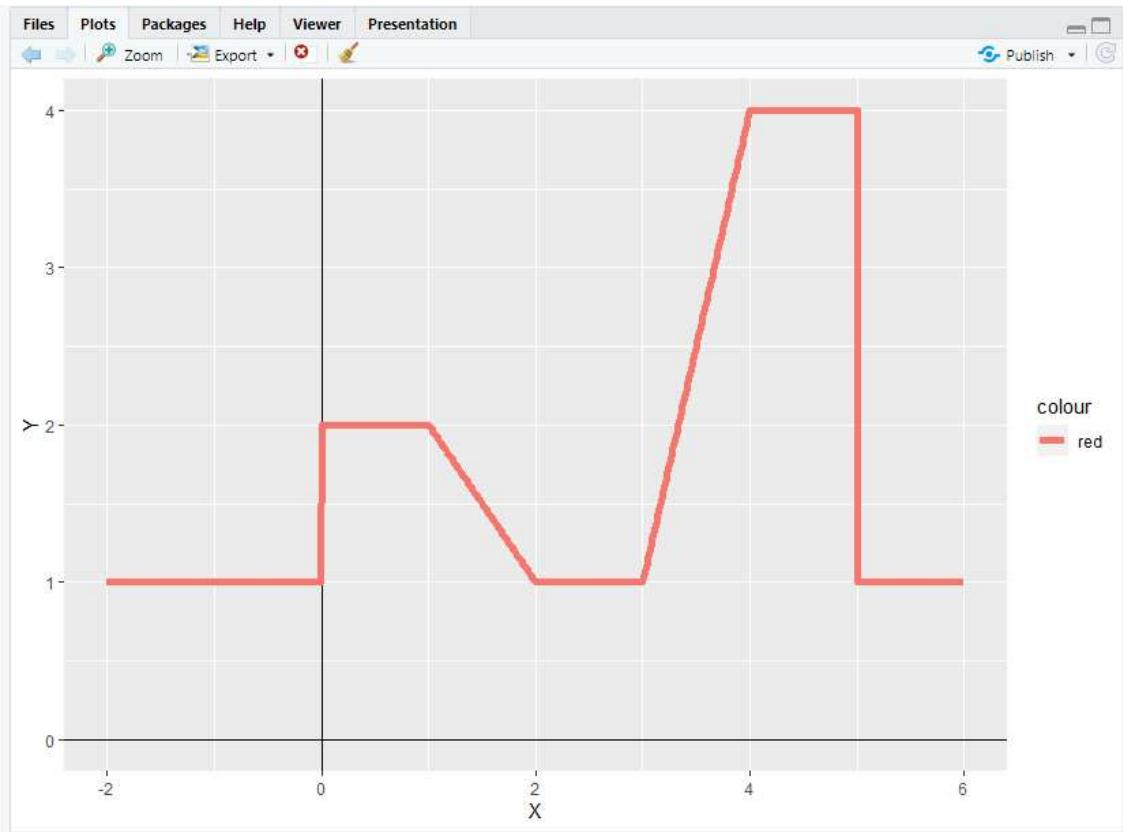
**Answer:**

$$Y = \beta_0 + \beta_1 b_1(X) + \beta_2 b_2(X) + \epsilon$$

$$Y = 1 + 1 * I(0 \leq X \leq 2) - (X-1)I(1 \leq X \leq 2) + 3 * ((X-3)I(3 \leq X \leq 4) + I(4 < X \leq 5))$$

$$Y = 1 + 1 * I(0 \leq X \leq 2) - (X-1)I(1 \leq X \leq 2) + 3 * (X-3)I(3 \leq X \leq 4) + 3 * I(4 < X \leq 5)$$

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R 4.2.1 · ~/Desktop
> X=seq(-2,6,0.01)
> Y=1+(X>=0 & X<=2)-(X-1)*(X>=1 & X<=2)*3*(X-3)*(X>=3 & X<=4)+3*(X>4 & X<=5)
> data_frame=data.frame(X,Y)
> ggplot(data_frame,aes(x=X,y=Y,col="red"))+geom_vline(xintercept = 0)+ geom_hline(yintercept = 0)+geom_line(size=2)
```



The curve is constant between  $-2$  and  $0$  with slope=0, Constant between  $0$  and  $1$  with  $y=2$  and slope=0, curve is Linear between  $1$  and  $2$  with  $y=3-x$  and slope=-1, curve is constant from  $2$  to  $3$  with  $y=1$  and slope=0. For  $x$  value greater than  $3$  and less or equal to  $4$  slope is  $y=2x-8$  and slope is 2, for value between  $4$  to  $5$  remain constant with  $y=4$  and slope=0

### Exercise-5

**(a) As  $\lambda \rightarrow \infty$ , will g1 or g2 have the smaller training RSS?**

**Answer:** -

Smoothing spine g2 has higher order polynomial and smaller training RSS due to order of penalty term.

**(b) As  $\lambda \rightarrow \infty$ , will g1 or g2 have the smaller test RSS?**

**Answer:** -

Smoothing spine g1 has smaller test RSS. As Smoothing spine g2 have higher order polynomial and more flexibility so it may overfit the data.

**(c) For  $\lambda = 0$ , will g1 or g2 have the smaller training and test RSS**

**Answer:** -

We have  $\lambda = 0$  implied  $g1=g2$ , so for both g1 and g2 we will have same training and testing RSS

### Exercise-5

$$n=2, p=2 \quad x_{11}=x_{22}, x_{21}=x_{22}$$

$$y_1 + y_2 = 0 \quad x_{11} + x_{21} = 0 \quad x_{12} + x_{22} = 0$$

Q)

$$\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2$$

$$\Rightarrow \beta_0 = 0$$

$$\sum_{i=1}^n \left( y_i - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2$$

from given information

$$(y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_2 - \beta_1 x_{21} - \beta_2 x_{22})^2 + (\beta_1^2 + \beta_2^2)$$

$$\Rightarrow (y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_2 - \beta_1 x_{21} - \beta_2 x_{22})^2 + \lambda (\beta_1^2 + \beta_2^2)$$

(b)

From equation

$$\Rightarrow (y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_2 - \beta_1 x_{21} - \beta_2 x_{22})^2 + \lambda (\beta_1^2 + \beta_2^2)$$

$\Rightarrow$  we have

$$x_{11} + x_{21} = 0 \Rightarrow x_{11} = -x_{21}$$

$$x_{12} + x_{22} = 0 \Rightarrow x_{12} = -x_{22}$$

$$y_1 + y_2 = 0 \Rightarrow y_1 = -y_2$$

$$\Rightarrow (y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (-y_1 + \beta_1 x_{11} + \beta_2 x_{12})^2 + \lambda (\beta_1^2 + \beta_2^2)$$

$$\Rightarrow (y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_1 - \beta_1 x_{11} + \beta_2 x_{12})^2 + \lambda (\beta_1^2 + \beta_2^2)$$

$$\Rightarrow 2 \left[ y_1 - \beta_1 x_{11} - \beta_2 x_{12} \right]^2 + \lambda [\beta_1^2 + \beta_2^2]$$

$\Rightarrow$  equality yields

$$2y_1^2 - 4\beta_1 x_{11} y_1 - 4\beta_2 x_{11} y_1 +$$

$$4\beta_1 \beta_2 x_{11}^2 + 2\beta_1^2 x_{11}^2 + 2\beta_2^2 x_{11}^2 +$$

$$\lambda \beta_1^2 + \lambda \beta_2^2.$$

$\Rightarrow$  maximization with respect to

$\beta_1$  und  $\beta_2$  and set to zero.

gives:  $-4y_1 x_{11} + 4x_{11}^2 \beta_2 +$   
 $4x_{11}^2 \beta_1 + 2\lambda \beta_1 = 0$

$\Rightarrow 4y_1 x_{11} - 4x_{11}^2 \beta_2 =$

$$4x_{11}^2 \beta_1 + 2\lambda \beta_1$$

$$\Rightarrow 2(2y_1 x_{11} - 2x_{11}^2 \beta_2) =$$
  
$$2\beta_1(2x_{11}^2 + \lambda)$$

$$\Rightarrow 2y_1 x_{11} - 2x_{11}^2 \beta_2 = \beta_1(\lambda + 2x_{11}^2)$$

$$\Rightarrow \beta_1 = \frac{(2y_1 x_{11} - 2x_{11}^2 \beta_2)}{(\lambda + 2x_{11}^2)}$$

Similarly

$$B_2 = \frac{(a_2 x_{11} - a_1 x_{11} \beta_1)}{x + a x_{11}}$$

To prove

$\Rightarrow$  Simplify  $B_1$  &  $B_2$  equal

By putting value of  $B_2$  in  $B_1$   
and  $B_2$  into  $B_1$  will

~~get~~ get some result

So can prove

$B_1 = B_2$ . as required.

(c)

Write down two optimized

$$\Rightarrow \min_{B_j} \sum_{j=1}^P (y_j - \sum_{j=1}^P B_j x_{pj})^2$$

$$+ \lambda \sum_{j=1}^P |B_j|$$

$$\Rightarrow \min \left[ (y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_2 - \beta_1 x_{21} - \beta_2 x_{22})^2 \right] + (\lvert \beta_1 \rvert + \lvert \beta_2 \rvert)$$

(d) also coefficient  $\beta_1$  and  $\beta_2$  are not unique.

$$\Rightarrow \sum_{j=1}^n (y_j - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2$$

where  $\sum_{j=1}^p |\beta_j| \leq s$

$$\Rightarrow 2(y_1 - \beta_1 x_{11} - \beta_2 x_{12})^2 + (y_2 - \beta_1 x_{21} - \beta_2 x_{22})^2$$

$$\beta_1 + \beta_2 \leq s$$

$$\Rightarrow 2(y_1 - (\beta_1 + \beta_2)x_{11})^2 \geq 0$$

$$\begin{cases} 3x_{11} + x_{12} = 0 \\ x_{21} + x_{22} = 0 \\ x_{11} = x_{12} \\ x_{21} = x_{22} \end{cases}$$

$\Rightarrow$  RSS has minimum at zero.

$$\text{so } y_1 - (B_1 + B_2)x_1 = 0 \quad (\epsilon_{11} = 0)$$

$$B_1 + B_2 x_1 = y_1$$

$$B_1 + B_2 = y_1/x_1.$$

will be ps of zero.

$\Rightarrow$  we also have to consider lasso constraints.

sum is the intersection of contours of the function

$$[y_1 - (B_1 + B_2)x_1]^2 \text{ with}$$

~~the~~ lasso diamond.

$\Rightarrow$  so as

$B_1, B_2$  vary along the line

$$B_2 = B_1 + y_1/x_1$$

contours will touch lasso diamond at many

- points,

$\Rightarrow$  Thus, Lasso optimization has many solution not just one.

Solution

$$\beta \in \{\beta_1, \beta_2 : \quad$$

$$\beta_1 + \beta_2 = 5, \quad \beta_1 \in [0, 5], \\ \beta_2 \in [0, 5] \}$$

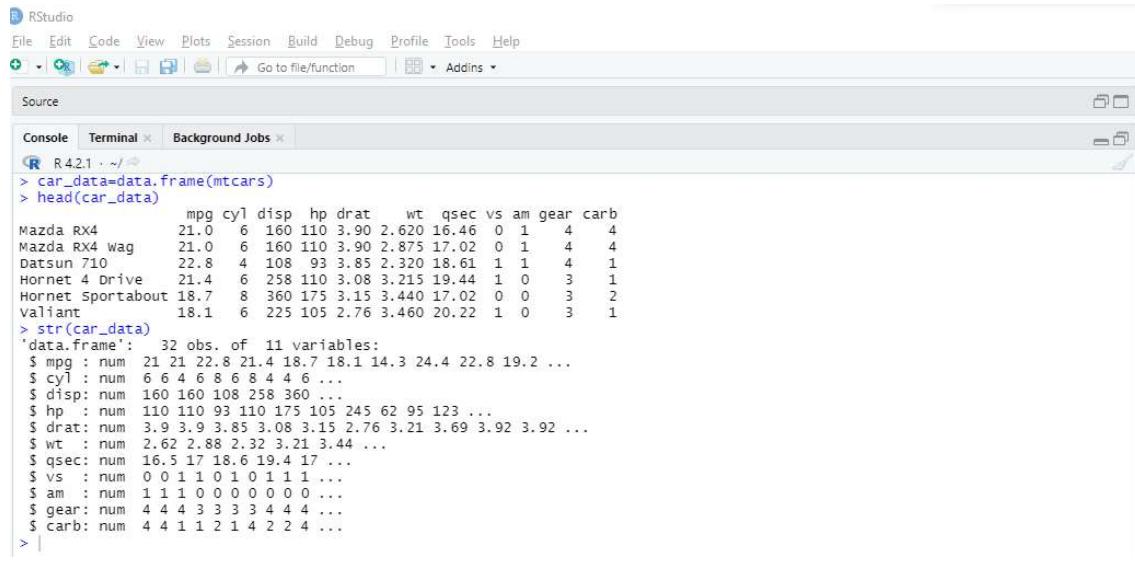
$\cup$

$$\{(\beta_1, \beta_2) : \quad \beta_1 + \beta_2 = -5, \quad \beta_1 \in [-5, 0], \\ \beta_2 \in [-5, 0] \}$$

## Assignment-3

### Problem 2.1

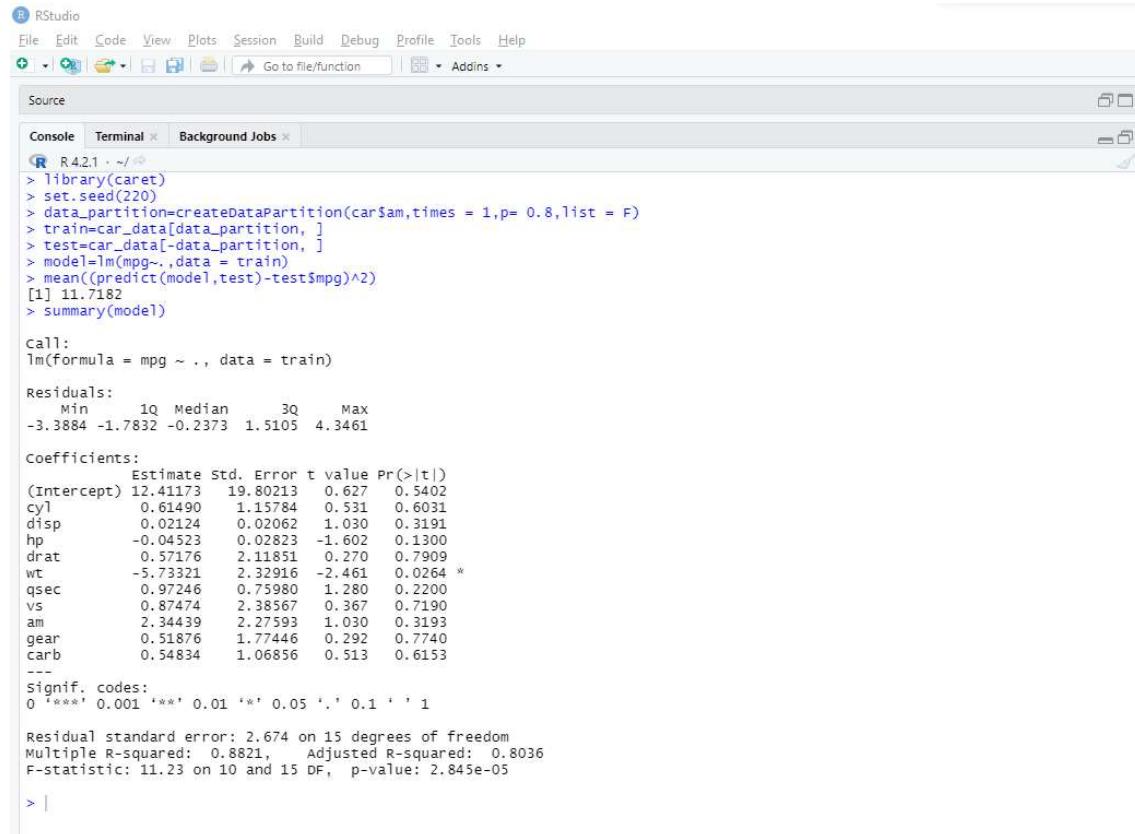
Load the mtcars sample dataset from the built-in datasets (data(mtcars)) into R using a dataframe.



RStudio interface showing the console tab with the command > car\_data=data.frame(mtcars) and its output. The output shows the first few rows of the mtcars dataset and its structure as a data frame with 32 observations and 11 variables.

```
R 4.2.1 . ~/ 
> car_data=data.frame(mtcars)
> head(car_data)
#> #>   mpg cyl disp  hp drat    wt  qsec vs am gear carb
#> Mazda RX4     21.0   6 160 110 3.90 2.620 16.46  0  1   4   4
#> Mazda RX4 Wag 21.0   6 160 110 3.90 2.875 17.02  0  1   4   4
#> Datsun 710    22.8   4 108  93 3.85 2.320 18.61  1  1   4   1
#> Hornet 4 Drive 21.4   6 258 110 3.08 3.215 19.44  1  0   3   1
#> Hornet Sportabout 18.7   8 360 175 3.15 3.440 17.02  0  0   3   2
#> Valiant      18.1   6 225 105 2.76 3.460 20.22  1  0   3   1
> str(car_data)
'data.frame': 32 obs. of 11 variables:
 $ mpg : num  21 21 22.8 21.4 18.7 ...
 $ cyl  : num  6 6 4 6 8 6 8 4 4 6 ...
 $ disp : num  160 160 108 258 360 ...
 $ hp   : num  110 110 93 110 175 105 245 62 95 123 ...
 $ drat : num  3.9 3.9 3.85 3.08 3.15 2.76 3.21 3.69 3.92 3.92 ...
 $ wt   : num  2.62 2.88 2.32 3.21 3.44 ...
 $ qsec : num  16.5 17 18.6 19.4 17 ...
 $ vs   : num  0 0 1 0 1 0 1 1 1 ...
 $ am   : num  1 1 1 0 0 0 0 0 0 0 ...
 $ gear : num  4 4 4 3 3 3 3 4 4 4 ...
 $ carb : num  4 4 1 1 2 1 4 2 2 4 ...
> |
```

Perform a basic 80/20 test-train split on the data and fit a linear model with mpg as the target response, and all other variables as predictors/features



RStudio interface showing the console tab with the commands for creating a test-train split and fitting a linear model. The output includes the summary of the linear model, showing coefficients, residuals, and statistical results.

```
R 4.2.1 . ~/ 
> library(caret)
> set.seed(220)
> data_partition=createDataPartition(car_data$mpg, times = 1, p= 0.8, list = F)
> train=car_data[data_partition, ]
> test=car_data[-data_partition, ]
> model=lm(mpg~., data = train)
> mean((predict(model,test)-test$mpg)^2)
[1] 11.7182
> summary(model)

Call:
lm(formula = mpg ~ ., data = train)

Residuals:
    Min      1Q  Median      3Q     Max
-3.3884 -1.7832 -0.2373  1.5105  4.3461

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 12.41173  19.80213  0.627  0.5402
cyl         0.61490  1.15784  0.531  0.6031
disp        0.02124  0.02062  1.030  0.3191
hp          -0.04523  0.02823 -1.602  0.1300
drat        0.57176  2.11851  0.270  0.7909
wt          -5.73321  2.32916 -2.461  0.0264 *
qsec        0.97246  0.75980  1.280  0.2200
vs          0.87474  2.38567  0.367  0.7190
am          2.34439  2.27593  1.030  0.3193
gear        0.51876  1.77446  0.292  0.7740
carb        0.54834  1.06856  0.513  0.6153
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.674 on 15 degrees of freedom
Multiple R-squared:  0.8821, Adjusted R-squared:  0.8036
F-statistic: 11.23 on 10 and 15 DF,  p-value: 2.845e-05
> |
```

What features are selected as relevant based on resulting t-statistics? What are the associated coefficient values for relevant features?

```
R 4.2.1 . ~/r
> coef(model)
(Intercept) cyl disp hp drat wt qsec vs am gear
12.41173489 0.61489995 0.02124426 -0.04523187 0.57176446 -5.73320752 0.97245746 0.87473702 2.34438911 0.51875794
> |
```

Only feature `wt` seems to be relevant feature.

Perform a ridge regression using the `glmnet` package from CRAN, specifying a vector of 100 values of  $\lambda$  for tuning. Use cross-validation (via `cv.glmnet`) to determine the minimum value for  $\lambda$  - what do you obtain?

```
R 4.2.1 . ~/r
> library(glmnet)
> x1=model.matrix(mpg~.,train)[,-1]
> y1=train$mpg
> lambdacv=10^seq(5,-5,by=-1)
> ridgecv=cv.glmnet(x1,y1,alpha=0,lambda = lambdacv)
```

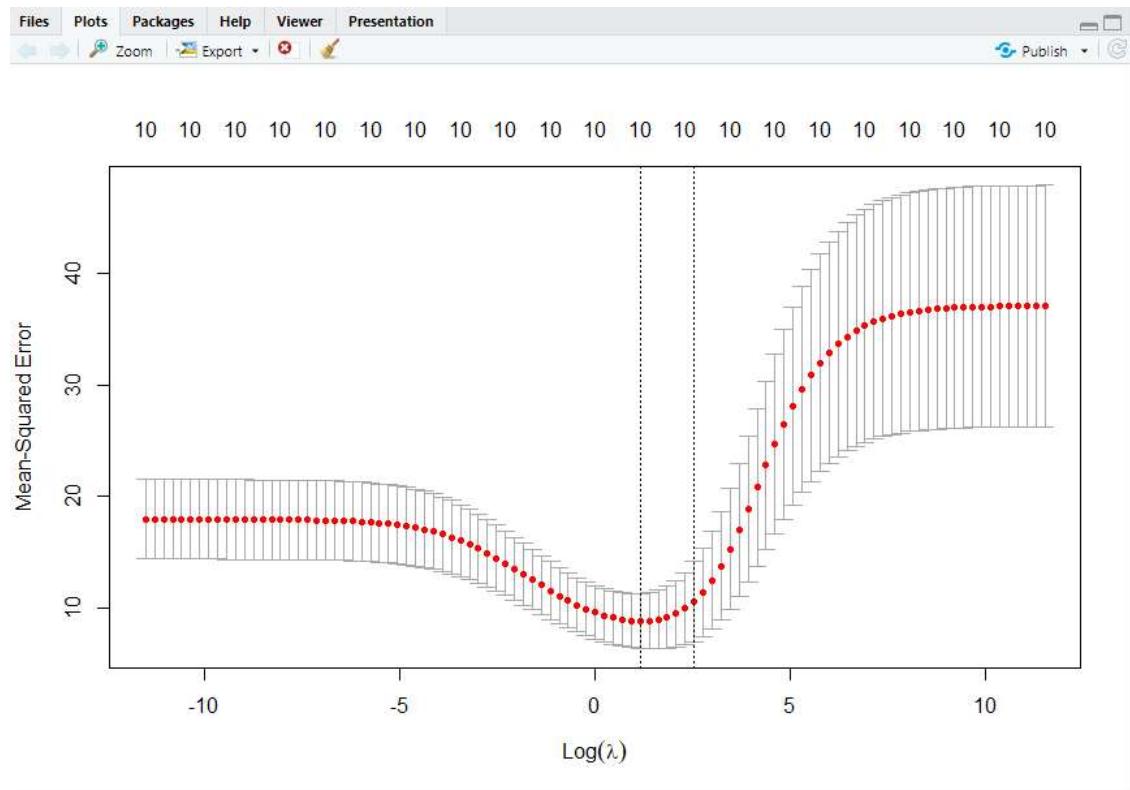
  

```
> plot(ridgecv)
> |
```

```
R 4.2.1 . ~/r
> ridgecv
Call: cv.glmnet(x = x1, y = y1, lambda = lambdacv, alpha = 0)
Measure: Mean-Squared Error

Lambda Index Measure SE Nonzero
min 3.162 46 8.782 2.450 10
1se 12.589 40 10.527 3.584 10
> |
```

Plot training MSE as a function of  $\lambda$



Minimum value for Lambda and regression model using glmnet

```
RStudio
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Source
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R 4.2.1 · ~/d
> min_lambda=ridgecv$lambda.min
> min_lambda
[1] 3.162278
> fit=glmnet(x1,y1,alpha=0,lambda = min_lambda)
> summary(fit)
    Length Class      Mode
a0          1   -none- numeric
beta       10 dgMatrix s4
df          1   -none- numeric
dim         2   -none- numeric
lambda     1   -none- numeric
dev.ratio  1   -none- numeric
nulldev    1   -none- numeric
npasses    1   -none- numeric
jerr        1   -none- numeric
offset     1   -none- logical
call        5   -none- call
nobs       1   -none- numeric
> coef(ridgecv,s="lambda.min")
11 x 1 sparse Matrix of class "dgCMatrix"
           s1
(Intercept) 22.765889374
cyl        -0.313166738
disp       -0.006137202
hp        -0.012575125
drat        0.829486711
wt        -1.494756384
qsec        0.147770130
vs         0.810816246
am         1.370276352
gear        0.583594189
carb       -0.478560452
> |
```

What is out-of-sample test set performance using predict? and how do the coefficients differ versus the regular linear model?

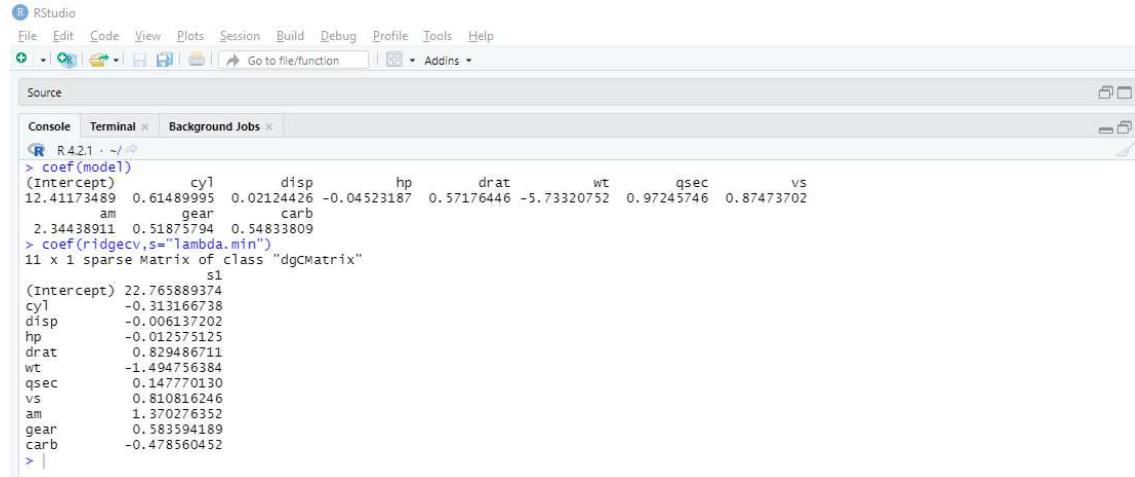


RStudio interface showing the console tab with the following R code and output:

```
R 4.2.1 · ~/r
> rx=model.matrix(mpg~.,test)[,-1]
> predict_model=predict(fit,s=min_lambda,newx = rx,type="response")
> mean((predict_model-test$mpg)^2)
[1] 3.741804
> |
```

So, from the result of ridge regression, it shows MSE on test data reduce from 11.7182 to 3.741804

Has ridge regression performed shrinkage, variable selection, or both?



RStudio interface showing the console tab with the following R code and output comparing ridge regression coefficients (s1) with glm model coefficients:

```
R 4.2.1 · ~/r
> coef(model)
(Intercept) cyl disp hp drat wt qsec vs
12.41173489 0.61489995 0.02124426 -0.04523187 0.57176446 -5.73320752 0.97245746 0.87473702
am gear carb
2.34438911 0.51875794 0.54833809
> coef(ridgecv,s="lambda.min")
> coef(ridgecv,s="lambda.min")
11 x 1 sparse Matrix of class "dgCMatrix"
s1
(Intercept) 22.765889374
cyl -0.313166738
disp -0.006137202
hp -0.012575125
drat 0.829486711
wt -1.494756384
qsec 0.147770120
vs 0.810816246
am 1.370276352
gear 0.583594189
carb -0.478560452
> |
```

From above ridge Regression the coefficients and glm model coefficients results, we can see that ridge regression coefficients have shrunk and some of these values are close to zero compared to glm model coefficients. So, answer is yes ridge regression has performed shrinkage.

### **Problem 2.2**

Load the Swiss sample dataset from the built-in datasets (data (Swiss)) into R using a dataframe.

```

R RStudio
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Source
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R 4.2.1 · ~/r/
> swiss_data<-data.frame(swiss)
> head(swiss_data)
  Fertility Agriculture Examination Education Catholic Infant.Mortality
Courteley    80.2        17.0       15     12     9.96          22.2
Delemont     83.1        45.1        6      9    84.84          22.2
Franches-Mnt 92.5        39.7        5      5    93.40          20.2
Moutier      85.8        36.5       12      7   33.77          20.3
Neuveville   76.9        43.5       17     15    5.16          20.6
Porrentruy   76.1        35.3        9      7   90.57          26.6
> str(swiss_data)
'data.frame': 47 obs. of 6 variables:
 $ Fertility : num  80.2 83.1 92.5 85.8 76.9 ...
 $ Agriculture: num  17 45.1 39.7 36.5 43.5 ...
 $ Examination: int  15 6 5 12 17 9 16 14 12 ...
 $ Education   : int  12 9 5 7 15 7 7 13 ...
 $ Catholic    : num  9.96 84.84 93.4 33.77 5.16 ...
 $ Infant.Mortality: num  22.2 22.2 20.2 20.3 20.6 26.6 23.6 24.9 21 24.4 ...
> |

```

Perform a basic 80/20 test-train split on the data (you may use caret, the sample method, or manually) and fit a linear model with Fertility as the target response, and all other variables as predictors/features.

What features are selected as relevant based on resulting t-statistics? What are the associated coefficient values for relevant features?

```

R RStudio
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Source
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R 4.2.1 · ~/r/
> set.seed(140)
> data_split<-createDataPartition(swiss_data$Fertility,p=0.8, list=F)
> train<-swiss_data[data_split,]
> test<-swiss_data[-data_split,]
> model1<-lm(Fertility~.,train)
> summary(model1)

Call:
lm(formula = Fertility ~ ., data = train)

Residuals:
    Min      1Q      Median      3Q      Max 
-15.6942 -5.0406  0.3346  3.7786 16.5156 

Coefficients:
            Estimate Std. Error t value pr(>|t|)    
(Intercept) 65.08317 12.18933  5.340 6.77e-06 ***
Agriculture -0.21914  0.08286 -2.645 0.03242 *  
Examination -0.22739  0.27913 -0.815 0.42112    
Education   -0.94604  0.20220 -4.679 4.74e-05 ***
Catholic    0.11361  0.03963  2.867 0.00717 ** 
Infant.Mortality 1.31036  0.43945  2.982 0.00535 ** 
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7.246 on 33 degrees of freedom
Multiple R-squared:  0.7142, Adjusted R-squared:  0.6708 
F-statistic: 16.49 on 5 and 33 DF,  p-value: 3.73e-16

> model1$coefficients
(Intercept)      Agriculture      Examination      Education      Catholic 
  65.0851684     -0.2191404      -0.2273869     -0.9460447      0.1136121 
Infant.Mortality 1.3103643
> |

```

Based on t-statistics except examination, features like agriculture, Education and infant Mortality are selected as relevant.

```

R RStudio
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Source
Console Terminal Background Jobs

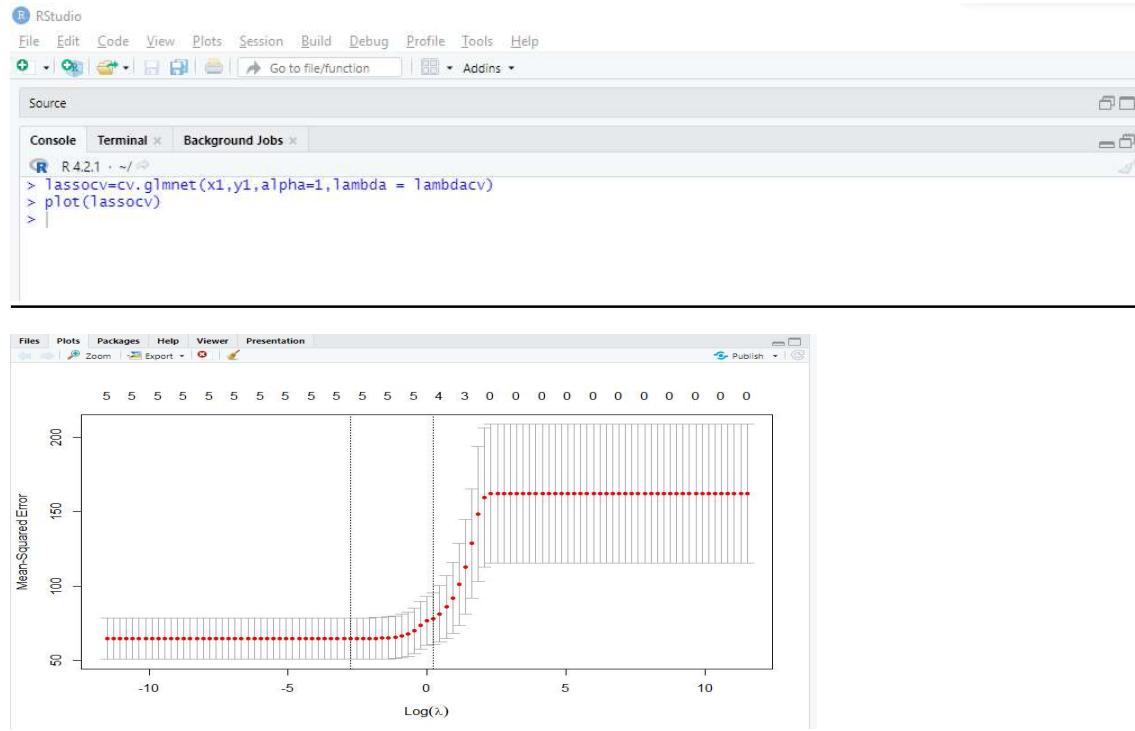
R 4.2.1 · ~/r/
> mean((test$Fertility-predict(model1,test))^2)
[1] 55.80315
> |

```

Perform a lasso regression using the `glmnet` package from CRAN, specifying a vector of 100 values of  $\lambda$  for tuning.

```
R 4.2.1 - ~/ 
> library(glmnet)
Loading required package: Matrix
Loaded glmnet 4.1-4
> library(Matrix)
> x1=model.matrix(Fertility~.,train)[,-1]
> y1=train$Fertility
> lambdacv=10^seq(5,-5,by=-.1)
```

Use cross-validation (via `cv.glmnet`) to determine the minimum value for  $\lambda$  - what do you obtain?  
(Hint: You can use `doMC` in order to speed-up your cross-validation by specifying `parallel=TRUE` in your `glmnet` calls.). Plot training MSE as a function of  $\lambda$  (you may also use  $\log \lambda$ )



```
R 4.2.1 - ~/ 
> min_lambda=lassocv$lambda.min
> min_lambda
[1] 0.06309573
```

Minimum value for Lambda is 0.0630

```
R 4.2.1 . ~/○
> fit=glmnet(x1,y1,alpha = 1,lambda = min_lambda)
> summary(fit)
  Length Class Mode
a0      1   -none- numeric
beta     5   dgCMatrix S4
df       1   -none- numeric
dim      2   -none- numeric
lambda   1   -none- numeric
dev.ratio 1   -none- numeric
nulldev   1   -none- numeric
npasses   1   -none- numeric
jerr      1   -none- numeric
offset    1   -none- logical
call      5   -none- call
nobs      1   -none- numeric
> |
```

What is out-of-sample test set performance (using predict)?

```
R 4.2.1 . ~/○
> xl=model.matrix(Fertility~.,test)[,-1]
> predict_model=predict(fit,s=min_lambda,newx = xl,type = "response")
> mean((predict_model-test$Fertility)^2)
[1] 53.18605
> |
```

How do the coefficients differ versus the regular linear model? Has lasso regression performed shrinkage, variable selection, or both?

```
R 4.2.1 . ~/○
> coef(lassocv)
6 x 1 sparse Matrix of class "dgCMatrix"
  s1
(Intercept) 54.61629460
Agriculture .
Examination -0.14250803
Education    -0.56876487
Catholic     0.04920414
Infant.Mortality 1.10944014
> coef(model1)
(Intercept) Agriculture Examination Education Catholic Infant.Mortality
  65.0851684     -0.2191404    -0.2273869   -0.9460447  0.1136121     1.3103643
> |
```

From the linear model coefficients and lasso regression coefficients, we can say that lasso regression has performed shrinkage.

### **Problem 2.3**

Load the Concrete Compressive Strength sample dataset from the UCI Machine Learning Repository (Concrete Data.xls) into R using a dataframe

```
R 4.2.1 - ~/ 
> library(readxl)
> con_data=read_excel("C:/Users/DELL/Desktop/DPA/Assignment-3/Concrete_Data.xls")
> summary(con_data)
Cement (component 1)(kg in a m^3 mixture) Blast Furnace Slag (component 2)(kg in a m^3 mixture)
Min. :102.0 Min. : 0.0
1st Qu.:192.4 1st Qu.: 0.0
Median :272.9 Median : 22.0
Mean :281.2 Mean : 73.9
3rd Qu.:350.0 3rd Qu.:142.9
Max. :540.0 Max. :359.4
Fly Ash (component 3)(kg in a m^3 mixture) water (component 4)(kg in a m^3 mixture)
Min. : 0.00 Min. :121.8
1st Qu.: 0.00 1st Qu.:164.9
Median : 0.00 Median :185.0
Mean : 54.19 Mean :181.6
3rd Qu.:118.27 3rd Qu.:192.0
Max. :200.10 Max. :247.0
Superplasticizer (component 5)(kg in a m^3 mixture) Coarse Aggregate (component 6)(kg in a m^3 mixture)
Min. : 0.000 Min. : 801.0
1st Qu.: 0.000 1st Qu.: 932.0
Median : 6.350 Median : 968.0
Mean : 6.203 Mean : 972.9
3rd Qu.:10.160 3rd Qu.:1029.4
Max. :32.200 Max. :1145.0
Fine Aggregate (component 7)(kg in a m^3 mixture) Age (day) Concrete compressive strength(MPa, megapascals)
Min. :594.0 Min. : 1.00 Min. : 2.332
1st Qu.:731.0 1st Qu.: 7.00 1st Qu.:23.707
Median :779.5 Median :28.00 Median :34.443
Mean :773.6 Mean :45.66 Mean :35.818
3rd Qu.:824.0 3rd Qu.:56.00 3rd Qu.:46.136
Max. :992.6 Max. :365.00 Max. :82.599
> colnames(con_data)=c("cement","Blast_fur","Fly_Ash","water","super_plas","Coarse_agg","fine_agg","age","con_comp_strength")
> con_data=data.frame(con_data)
> summary(con_data)
   cement      Blast_fur      Fly_Ash      water      super_plas      Coarse_agg      fine_agg
Min. :102.0  Min. : 0.0  Min. : 0.00  Min. :121.8  Min. : 0.000  Min. : 801.0  Min. :594.0
1st Qu.:192.4 1st Qu.: 0.0  1st Qu.: 0.00  1st Qu.:164.9  1st Qu.: 0.000  1st Qu.: 932.0  1st Qu.:731.0
Median :272.9  Median : 22.0  Median : 0.00  Median :185.0  Median : 6.350  Median : 968.0  Median :779.5
Mean :281.2  Mean : 73.9  Mean : 54.19  Mean :181.6  Mean : 6.203  Mean : 972.9  Mean :773.6
3rd Qu.:350.0  3rd Qu.:142.9  3rd Qu.:118.27  3rd Qu.:192.0  3rd Qu.:10.160  3rd Qu.:1029.4  3rd Qu.:824.0
Max. :540.0  Max. :359.4  Max. :200.10  Max. :247.0  Max. :32.200  Max. :1145.0  Max. :992.6
   age      con_comp_strength
Min. : 1.00  Min. : 2.332
1st Qu.: 7.00  1st Qu.:23.707
Median :28.00  Median :34.443
Mean :45.66  Mean :35.818
3rd Qu.:56.00  3rd Qu.:46.136
Max. :365.00  Max. :82.599
> |
```

Use the mgcv package to create a generalized additive model (via the gam function) to predict the Concrete Compressive Strength (CCS) as a non-linear function of the input components (C1-C6) - compare the R2 value for a GAM with linear terms as well as smoothed terms

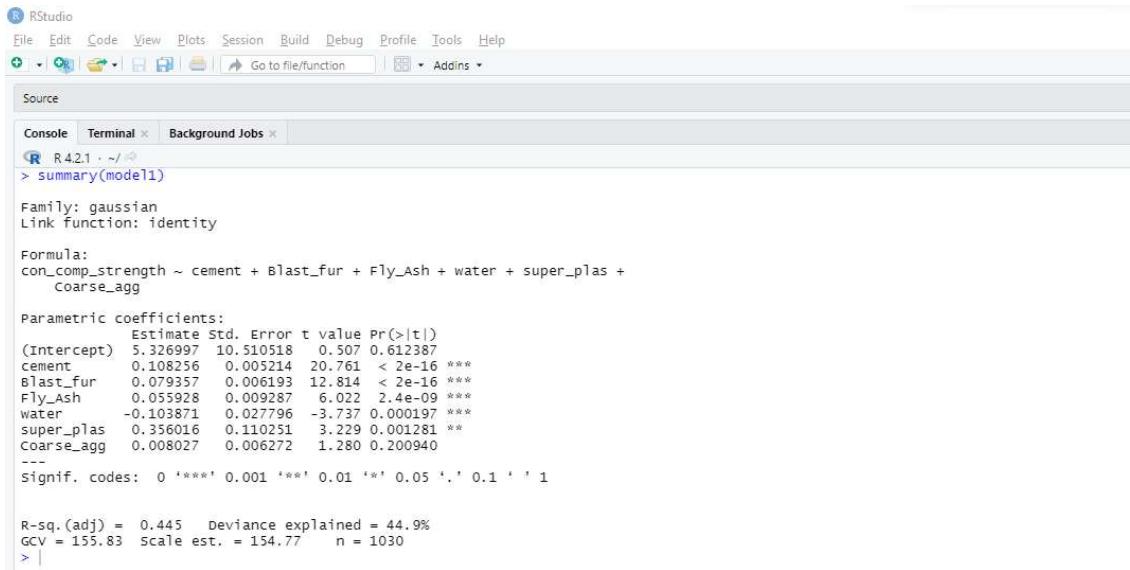
We want only 1<sup>st</sup> six column.

### GAM with Linear Terms

```
R 4.2.1 - ~/ 
> library(mgcv)
Loading required package: nlme
This is mgcv 1.8-40. For overview type 'help("mgcv-package")'.
> model1=gam(con_comp_strength~cement+Blast_fur+Fly_Ash+water+super_plas+Coarse_agg,data = con_data)
> model1
Family: gaussian
Link function: identity

Formula:
con_comp_strength ~ cement + Blast_fur + Fly_Ash + water + super_plas +
    Coarse_agg
Total model degrees of freedom 7

GCV score: 155.8332
> |
```



```

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Source
Console Terminal Background Jobs
R 4.2.1 ~/ ~
> summary(model1)

Family: gaussian
Link function: identity

Formula:
con_comp_strength ~ cement + Blast_fur + Fly_Ash + water + super_plas +
Coarse_agg

Parametric coefficients:
Estimate Std. Error t value Pr(>|t|)
(Intercept) 5.326997 10.510518 0.507 0.612387
cement 0.108256 0.005214 20.761 < 2e-16 ***
Blast_fur 0.079357 0.006193 12.814 < 2e-16 ***
Fly_Ash 0.055928 0.009287 6.022 2.4e-09 ***
water -0.103871 0.027796 -3.737 0.000197 ***
super_plas 0.356016 0.110251 3.229 0.001281 **
Coarse_agg 0.008027 0.006272 1.280 0.200940
---
Signif. codes: 0 '****' 0.001 '***' 0.01 '**' 0.05 '*' 0.1 ' ' 1

R-sq.(adj) = 0.445 Deviance explained = 44.9%
GCV = 155.83 Scale est. = 154.77 n = 1030
>

```

It shows we have statistical effects for cement, Blast furnace but not for coarse aggregate and the adjusted R-squared indicates a significant amount of the variance.

## R<sup>2</sup> Calculation

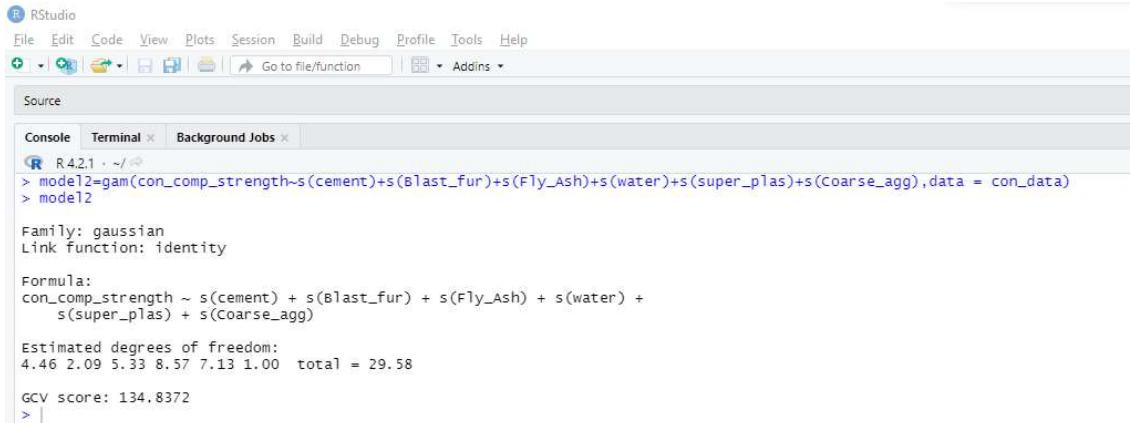


```

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Source
Console Terminal Background Jobs
R 4.2.1 ~/ ~
> model1.sse=sum(fitted(model1)-con_data$con_comp_strength)^2
> model1.sse
[1] 1.560331e-19
> model1.ssr=sum(fitted(model1)-mean(con_data$con_comp_strength))^2
> model1.ssr
[1] 1.539979e-19
> model1.sst=model1.sse+model1.ssr
> model1.sst
[1] 3.10031e-19
> Rsquared_lm=1-(model1.sse/model1.sst)
> Rsquared_lm
[1] 0.4967177
>

```

## GAM with Smoothed term



```

RStudio
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Source
Console Terminal Background Jobs
R 4.2.1 ~/ ~
> model2=gam(con_comp_strength~s(cement)+s(Blast_fur)+s(Fly_Ash)+s(water)+s(super_plas)+s(coarse_agg),data = con_data)
> model2

Family: gaussian
Link function: identity

Formula:
con_comp_strength ~ s(cement) + s(Blast_fur) + s(Fly_Ash) + s(water) +
s(super_plas) + s(coarse_agg)

Estimated degrees of freedom:
4.46 2.09 5.33 8.57 7.13 1.00 total = 29.58

GCV score: 134.8372
>

```

```

RStudio
File Edit Code View Plots Session Build Debug Profile Tools Help
+ - Go to file/function | Addins
Source
Console Terminal Background Jobs
R 4.2.1 - ~/r
> summary(model12)
Family: gaussian
Link function: identity

Formula:
con_comp_strength ~ s(cement) + s(Blast_fur) + s(Fly_Ash) + s(water) +
  s(super_plas) + s(coarse_agg)

Parametric coefficients:
Estimate Std. Error t value Pr(>|t|)
(Intercept) 35.8178    0.3566 100.4 <2e-16 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Approximate significance of smooth terms:
edf Ref.df F p-value
s(cement) 4.464 5.513 69.530 <2e-16 ***
s(Blast_fur) 2.088 2.578 48.091 <2e-16 ***
s(Fly_Ash) 5.332 6.404 1.784 0.101
s(water) 8.567 8.936 13.504 <2e-16 ***
s(super_plas) 7.133 8.143 5.498 1.22e-06 ***
s(coarse_agg) 1.000 1.000 0.018 0.892
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R-sq.(adj) = 0.531 Deviance explained = 54.4%
GCV = 134.84 Scale est. = 130.96 n = 1030
>

```

We can also note that this model accounts for much of the variance in concrete compressive strength, with an adjusted R-squared of .531. Means, it seems like the Cement is associated with concrete compressive strength.

## R<sup>2</sup> calculation

```

RStudio
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Source
Console Terminal Background Jobs
R 4.2.1 - ~/r
> model12.sse=sum(fitted(model12)-con_data$con_comp_strength)^2
> model12.sse
[1] 3.017023e-16
> model12.ssr=sum(fitted(model12)-mean(con_data$con_comp_strength))^2
> model12.ssr
[1] 3.017921e-16
> model12.sst=model12.sse+model12.ssr
> model12.sst
[1] 6.034944e-16
> Rsquared_sm=1-(model12.sse/model12.sst)
> Rsquared_sm
[1] 0.5000744
>

```

## Linear term vs smoothed term

R<sup>2</sup> of the smoothed model is greater than the main R<sup>2</sup> value.

RStudio

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Source

Console Terminal Background Jobs

R 4.2.1 · ~/

```
> anova(model1,model2,test = "chisq")
Analysis of Deviance Table

Model 1: con_comp_strength ~ cement + Blast_fur + Fly_Ash + water + super_plas +
          Coarse_agg
Model 2: con_comp_strength ~ s(cement) + s(Blast_fur) + s(Fly_Ash) + s(water) +
          s(super_plas) + s(Coarse_agg)
Resid. Df Resid. Dev Df Deviance Pr(>chi)
1      1023.00    158334
2      996.43    131019 26.574 < 2.2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
>
```

Results indicate that integrating nonlinear relationships of the covariates improves the model.

**Visualize** the regression using the visreg package, showing the fit as a function of each predictor with confidence intervals - comment on the quality of the fit at extreme values of the predictors.

RStudio

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Source

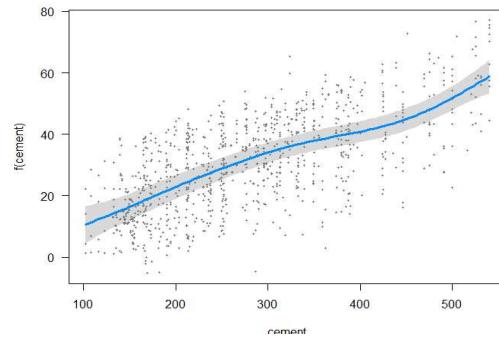
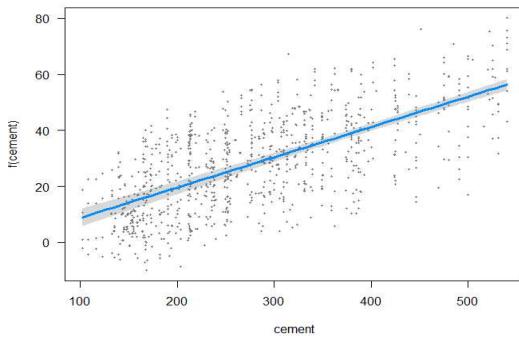
Console Terminal Background Jobs

R 4.2.1 · ~/

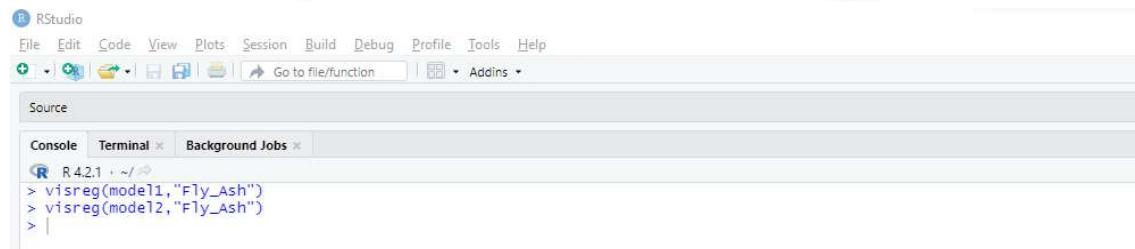
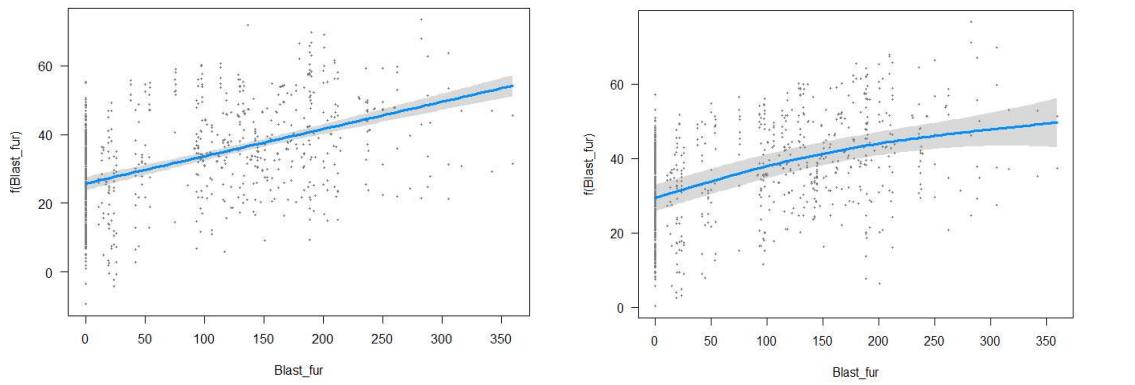
```
> install.packages("visreg")
Installing package into 'C:/Users/DELL/AppData/Local/R/win-library/4.2'
(as "lib" is unspecified)
trying URL 'https://cran.rstudio.com/bin/windows/contrib/4.2/visreg_2.7.0.zip'
Content type 'application/zip' length 299852 bytes (292 KB)
downloaded 292 KB

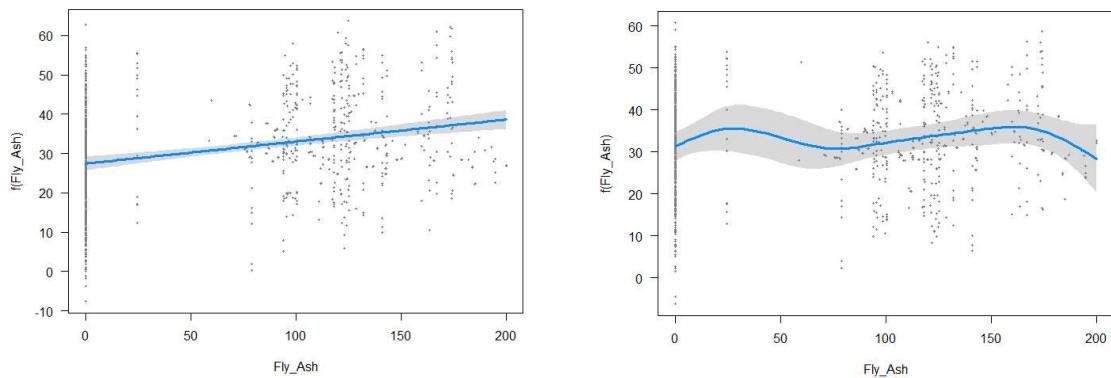
package 'visreg' successfully unpacked and MD5 sums checked

The downloaded binary packages are in
  C:\Users\DELL\AppData\Local\Temp\RtmpgLRhB\downloaded_packages
> library(visreg)
> visreg(model1,'cement')
> visreg(model2,'cement')
> |
```



The above graph displays changes in concrete compressive strengths as a function of cement while all other variables remain fixed. In Graph blue line show expected value, gray colour denote confidence interval and dark gray spots indicate partial residuals .





RStudio

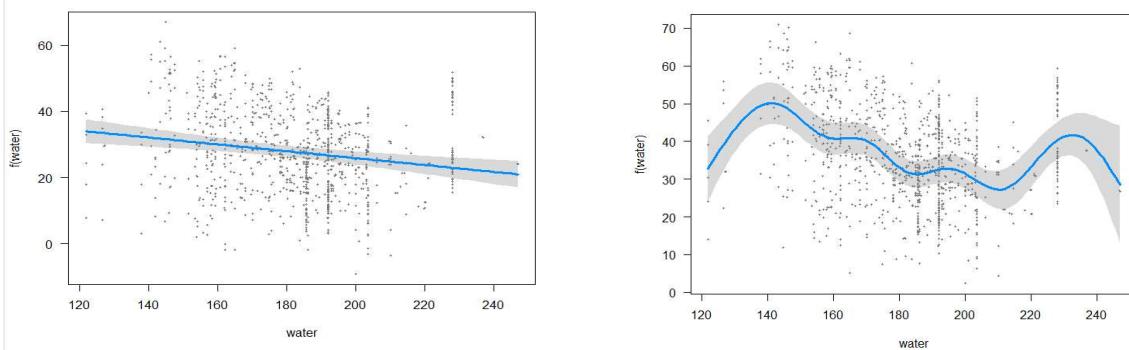
File Edit Code View Plots Session Build Debug Profile Tools Help

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```
R 4.2.1 · ~/ ◊
> visreg(model1,"water")
> visreg(model2,"water")
> |
```



RStudio

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+ - Go to file/function Addins

Source

Console Terminal Background Jobs

```
R 4.2.1 · ~/ ◊
> visreg(model1,"super_plas")
> visreg(model2,"super_plas")
> |
```

