

# Probability

①

The study of probability provides a mathematical framework for such assertions and is essential in every decision making process.

Principle of counting, if an event can happen in  $n_1$  ways and thereafter for each of these events a second event can happen in  $n_2$  ways and for each of these first and second events a third event can happen for  $n_3$  ways and so on, then the number of ways these events can happen is given by the product  $n_1 \cdot n_2 \cdot n_3 \cdots n_m$ .

Permutations: A permutation of a number of objects is their arrangement in some definite order.

Q 1a

The number of permutations of  $\underline{n}$  different things taken over  $\underline{r}$  at time is

$${}^n P_r = \frac{n!}{(n-r)!}$$

Combinations: The number of combinations of  $\underline{n}$  different objects taken  $\underline{r}$  at a time is

$${}^n C_r = \frac{{}^n P_r}{r!} = \frac{n!}{r!(n-r)!}$$

Basic Definitions

Exhaustive events: A sets of events is said to be exhaustive, if it includes all the possible events.

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## Mutually Exclusive events:

If the occurrence of one of the events precludes the occurrence of all other, then such a set of events is said to be mutually exclusive.

## Equally likely Events:

If one of the events cannot be expected to happen in preference to another then such events are said to be equally likely.

Probability: If  $n$  are exhaustive, mutually exclusive and equally likely cases of which  $m$  are favourable to an event, then the probability of the happening of event is  $\underline{\underline{m/n}}$

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Note: ① If  $E$  is any event in a sample space  $S$ , then  $0 \leq P(E) \leq 1$

② to the entire sample space  $S$  there corresponds  $P(S) = 1$

③ If  $A$  &  $B$  are mutually exclusive events then  $P(A \cup B) = P(A) + P(B)$

④ If  $A$  &  $B$  are any events in a sample space  $S$ , then

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

⑤ If  $E$  is an event and  $E^c$  is complement of  $E$  then  $P(E) = 1 - P(E^c)$

⑥ If  $A$  &  $B$  are events in a sample space  $S$  and  $P(A) \neq 0, P(B) \neq 0$  then

$$P(A \cap B) = P(A)P(B|A) = P(B)P(A|B)$$

## Random Variable's

A random variable  $X$  (also called stochastic variable or variate) is a function whose values are real numbers and depend on 'chance'; more precisely, ~~for~~ ~~knowing~~ ~~whether~~ ~~has~~ the following properties

## Discrete Random variables and Discrete Distribution:

A random variable  $X$  [ $X$  is outcome of some experiment] and the corresponding distribution are said to be distribution.

If the probability that  $X$  takes the values  $x_i$  is  $P_i$  then  $P(X=x_i)=P_i, i=1, 2, \dots$

$$\textcircled{i} \quad P(x_i) > 0 \quad (\forall i) \quad P(x_i) = P_i \quad ?$$

$$\textcircled{ii} \quad \sum_i P(x_i) = 1$$

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Let  $x_1, x_2, \dots$  be the values for which  $x$  has positive probability and let  $p_1, p_2, \dots$  be the corresponding probabilities. Then

- i)  $p(x_i) \geq 0$     ii)  $\sum p(x_i) = 1$ , and

$$P(a < x \leq b) = \sum_{a < x_i \leq b} p(x_i)$$

continuous Random Variables, and

continuous Distributions;

A random variable  $X$  and the corresponding distributions are said to be of continuous type, if the distribution function

$F(x) = P(X \leq x)$  of  $X$  can be represented as

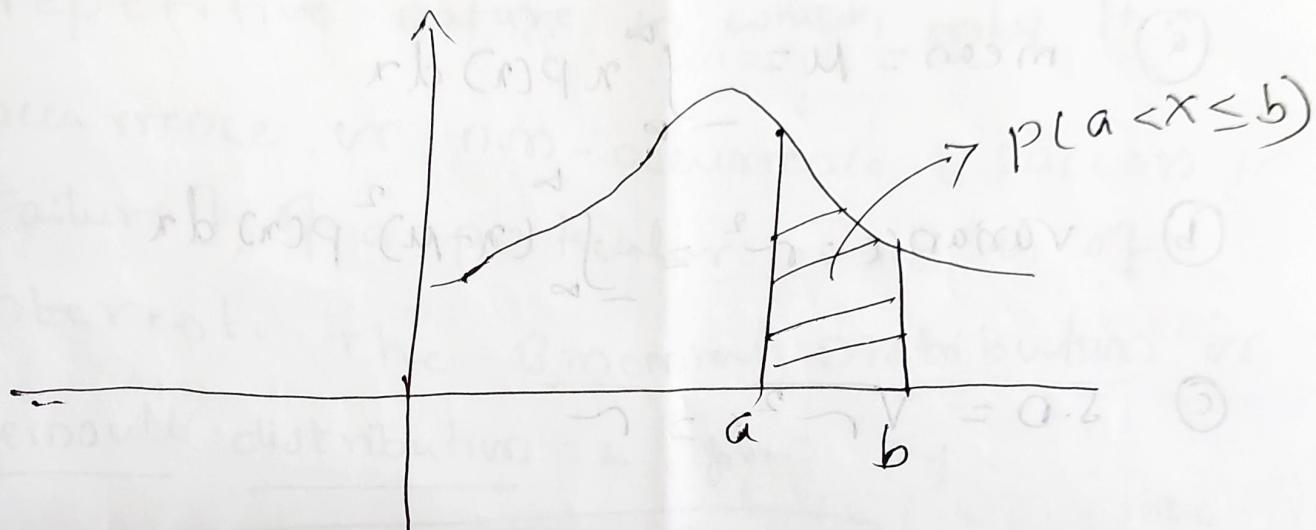
$$F(x) = P(X \leq x) = \int_{-\infty}^x f(u) du$$

The function  $f(x)$  is called probability density function

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And  $P(a < X \leq b) = \int_a^b f(x) dx$

$$= F(b) - F(a)$$



## Mean and Variance of Distribution:

### ① Discrete Distribution:

ⓐ Mean =  $M = \sum x_i P(x_i)$

ⓑ Variance =  $\sigma^2 = \sum x_i^2 P(x_i) - M^2$

or  $= \sum (x_i - M)^2 P(x_i)$

(i)

Standard Deviation =  $\sqrt{\sigma^2} = \sigma$

(ii) Continuous Distribution:

(a) mean =  $\mu = \int_{-\infty}^{\infty} x p(x) dx$

(b) variance =  $\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 p(x) dx$

(c) S.D =  $\sqrt{\sigma^2} = \sigma$

Notes for the frequency Distribution

i)  $N = \sum f_i$

ii) mean =  $\frac{\sum f_i x_i}{\sum f_i}$

(iii)  $Q_1 + Q_3 = M = \text{mean}$  (iv)

$M - (Q_1 + Q_3) = \sigma = \text{standard deviation}$  (v)

(vi)  $Q_1 - (M - Q_3) =$  (vii)

## Bernoulli Binomial Distribution:

It is concerned with trials of a repetitive nature in which only the occurrence or non-occurrence [success or failure], of a particular event is of interest. The Binomial Distribution or Bernoulli distribution is given by

$$P(X=x) = n C_x p^x q^{n-x}$$

where 'p' is the probability of success in one trial and 'q' is failure. If 'n' independent trials constitute the experiment and this experiment be repeated N times then the Binomial frequency Distribution is

$$N \cdot n C_x p^x q^{n-x}$$

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mean of Binomial Distribution

$$\boxed{M = np}$$

or

$$\boxed{\bar{x} = np}$$

variance of Binomial Distribution

$$\boxed{\sigma^2 = npq}$$

$$\boxed{S.D = \sqrt{\sigma^2}}$$

Poisson Distribution:

It is a distribution related to the probabilities of events which have a large number of independent opportunities for occurrence.

It is the limiting case of Binomial Distribution as  $p \rightarrow 0$  and  $n \rightarrow \infty$ .

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The discrete distribution with the probability function

$$P(X=n) = \frac{m^n e^{-m}}{n!} \text{ or}$$

$$P(X=n) = \frac{m^n e^{-m}}{n!}$$

is called Poisson Distribution.

The mean and S.D. of Poisson Distribution

$$\boxed{\text{Mean} = \mu = n \cdot p = m}$$

$$\boxed{\text{Variance} = \sigma^2 = \mu = m}$$

$$\boxed{\text{S.D.} = \sqrt{n^2} = n}$$

Poisson frequency Distribution

$$= N P C D = N \frac{m^n e^{-m}}{n!}$$

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Note:

- ① prove that the probability of  $x$  success in  $n$ -Bernoulli trials is  ${}^n C_x p^x q^{n-x}$

proof: Let  $x$  be the number of success.

given: In an experiment of  $n$ -Bernoulli

$\Rightarrow (n-x)$  are failures. ~~can be~~  
Hence  $x$  success &  $(n-x)$  failures can be obtained  
in  ${}^n C_x$  ways.

In each of these ways, the probability of  
a success and  $(n-x)$  failures is

$$= P(x \text{ success}) \cdot P((n-x) \text{ failures})$$

$$= \underbrace{P \cdots \cdots P}_{n \text{ times}} \times \boxed{\underbrace{q \quad q \quad \cdots \quad q}_{(n-x) \text{ times}}} \quad V = 0.2$$

$$= p^x q^{n-x}$$

$\therefore$  The probability of  $x$  success in  $n$ -Bernoulli trials is

$$P(X=x) = {}^n C_x p^x q^{n-x}$$

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② prove that mean and S.D. of Bernoulli Distribution (or Binomial Distribution) are  $np$  &  $\sqrt{npq}$

Proof: Let  $\mu$  be the mean of Binomial Distribution

w.k.t

$$\mu = \sum x_i P(x_i)$$

$$= \sum_{i=1}^n x_i {}^n C_{x_i} p^{x_i} q^{n-x_i}$$

since  $x_1=0, x_2=1, x_3=2, \dots$

$$\begin{aligned}\mu &= (0) {}^n C_0 p^0 q^{n-0} + (1) {}^n C_1 p^1 q^{n-1} + (2) {}^n C_2 p^2 q^{n-2} + \dots \\ &\quad + n {}^n C_n p^n q^{n-n} \\ &= np [q^{n-1} + pq^{n-2} + \dots + p^{n-1}]\end{aligned}$$

$$\mu = np [p+q]^{n-1}$$

since  $p+q=1$

$$\therefore \boxed{\mu = np}$$

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w.k.t

$$\text{var} = \sum x_i^2 p(x_i) - \mu^2$$

$$\text{var} = \sum_{i=1}^n x_i^2 p(x_i) - (np)^2$$

$$= \sum_{i=1}^n [x_i(n_i-1) + n_i] p(x_i) - \mu^2$$

$$= \sum_{i=1}^n n_i(n_i-1) p(x_i) + \sum_{i=1}^n x_i p(x_i) - \mu^2$$

$$= \sum_{i=1}^n x_i(n_i-1) \binom{n}{n_i} p^{x_i} q^{n-n_i} + \mu - \mu^2$$

$$= \sum_{i=1}^n x_i(n_i-1) \frac{n!}{x_i!(n-x_i)!} p^{x_i} q^{n-x_i} + \mu - \mu^2$$

$$= \sum_{i=1}^n \frac{n(n-1)(n-2)\dots}{(x_i-2)!(n-x_i)!} p^{x_i} q^{n-x_i} + \mu - \mu^2$$

$$= n(n-1)p^2 \sum_{i=1}^n \frac{(n-2)!}{(x_i-2)!(n-x_i)!} p^{x_i-2} q^{n-x_i} + \mu - \mu^2$$

$$= n(n-1)p^2 (p+q)^{n-2} + np - (np)^2$$

$$= n(n-1)p^2 + np - (np)^2 = n^2 p^2 - np^2 + np - (np)^2$$

$$\therefore \boxed{\text{var} = np(1-p) = npq}$$

$$\boxed{SD = \sqrt{npq}}$$

prove

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$$\textcircled{3} \text{ prove that } p(x) = \frac{m^x e^{-m}}{x!}$$

Consider the Binomial Distribution

$$p(x) = n C_x p^x q^{n-x} \rightarrow \textcircled{1}$$

$$= \frac{n(n-1)(n-2)\dots(n-(x-1))}{x!} p^x q^{n-x}$$

$$p(x) = \frac{n \cdot n^{x-1} (1-\frac{x}{n}) (1-\frac{2}{n}) \dots (1-\frac{(x-1)}{n})}{x!} p^x q^{n-x} \rightarrow \textcircled{2}$$

$$\text{wkt } np = m \quad q^n = (1-p)^n$$

$$\boxed{p = m/n}$$

$$\therefore q^n = \left[1 - \frac{m}{n}\right]^n = \left[\left(1 - \frac{m}{n}\right)^{\frac{n}{m}}\right]^m$$

$$\text{wkt } \lim_{k \rightarrow 0} (1+k)^{1/k} = e$$

$$\therefore \text{as } n \rightarrow \infty \left[1 - \left(\frac{m}{n}\right)\right]^{-\frac{1}{m}} \rightarrow e$$

$$\therefore \text{as } n \rightarrow \infty q^n \rightarrow e^{-m}$$

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from ②

$$\therefore P(x) = \frac{(np)^x (1-p)^{n-x}}{x!} \cdot \frac{\left(1 - \frac{(n-1)}{n}\right)^n}{n^n}$$

as  $n \rightarrow \infty$ 

$$P(x) = \frac{m^x e^{-m}}{x!}$$

④ P.T mean and S.D of Poisson Distribution  
are m.

Proof. For D.R.V mean is given by

$$\text{mean} = \mu = \sum n_i p(x_i)$$

$$= \sum_i n_i \frac{m^{n_i} e^{-m}}{n_i!}$$

$$= m e^{-m} \sum_i \frac{m^{n_i-1}}{(n_i-1)!}$$

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$$= m \bar{e}^m \left[ 1 + \frac{m}{1!} + \frac{m^2}{2!} + \frac{m^3}{3!} + \dots \right]$$

$$= m \bar{e}^m [e^m] = m$$

$\therefore \boxed{\text{mean} = m = M}$

$$\text{variance} = V = \sum x^2 p(x) - \mu^2$$

$$= \sum [(x-1)x + x] p - \mu^2$$

$$= \sum x(x-1)p + \sum np - \mu^2$$

$$= \sum n(n-1) \frac{m^2 \bar{e}^m}{n!} + [np - \mu^2]$$

$$= \sum \frac{m^2 \bar{e}^m}{(n-2)!} + [np - \mu^2]$$

$$= m^2 \bar{e}^m \sum \frac{m^{(n-2)}}{(n-2)!} + [np - \mu^2]$$

$$= m^2 \bar{e}^m [e^m] + m - \mu^2 = m$$

$\therefore \boxed{V = \sigma^2 = m}$

(Q)

## Continuous probability Distribution

If  $x$  belonging to the range of a continuous random variable  $X$ , we assign a real number  $f(x)$  satisfying the conditions

$$\text{i) } f(x) \geq 0 \quad \text{ii) } \int_{-\infty}^{\infty} f(x) dx = 1$$

$f(x)$  is called continuous probability function or probability density function.

If  $[a, b]$  is a subinterval of the  $\downarrow$  range of the space  $X$

then

$$P(a \leq x \leq b) = \int_a^b f(x) dx$$

## Cumulative Distribution Function

If  $X$  is a continuous random variable with probability density function  $f(x)$  then

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(x) dx$$

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Note: If  $r$  is any real number

then

$$P(X \geq r) = \int_r^{\infty} f(x) dx$$

$$P(X < r) = 1 - \int_r^{\infty} f(x) dx$$

Mean and Variance:

If  $X$  is continuous random variable with prob. density function  $f(x)$ . Then mean and variance of cont. dist.

$$\text{mean} = \mu = \int_{-\infty}^{\infty} x f(x) dx$$

$$\text{var} = \sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

$$= \int_{-\infty}^{\infty} x^2 f(x) dx - \mu^2$$

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## Exponential Distribution

The continuous probability distribution having the prob. density function  $f(x)$

given by

$$f(x) = \begin{cases} \alpha e^{-\alpha x}, & x > 0 \\ 0, & \text{otherwise,} \end{cases}$$

clearly  $f(x) \geq 0$  and

$$\int_{-\infty}^{\infty} f(x) dx = \int_0^{\infty} \alpha e^{-\alpha x} dx = \alpha \left[ \frac{e^{-\alpha x}}{-\alpha} \right]_0^{\infty}$$

$$= \alpha \left[ -\frac{1}{\alpha} 0 + \frac{1}{\alpha} \right] = 1$$

$$\therefore \int_{-\infty}^{\infty} f(x) dx = 1$$

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## mean and standard Deviation

$$P.T \quad \mu = \frac{1}{\alpha} \quad , \quad S.D = \frac{1}{\alpha}$$

wkt:  $\mu = \int_{-\infty}^{\infty} x f(x) dx$

$$\mu = \alpha \int_0^{\infty} x e^{-\alpha x} dx = \alpha \left[ x \left( \frac{e^{-\alpha x}}{-\alpha} \right) - \frac{e^{-\alpha x}}{(-\alpha)^2} \right]_0^{\infty}$$

$$= \alpha \left\{ [0 - 0] - [0 - \frac{1}{\alpha^2}] \right\} = \alpha \left[ \frac{1}{\alpha^2} \right]$$

$$\boxed{\mu = \frac{1}{\alpha}}$$

$$Var = \sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

$$= \int_0^{\infty} (x - \mu)^2 \cdot \alpha \cdot e^{-\alpha x} dx$$

$$= \alpha \left[ (x - \mu)^2 \frac{e^{-\alpha x}}{(-\alpha)} - 2(x - \mu) \frac{e^{-\alpha x}}{(-\alpha)^2} + 2 \frac{e^{-\alpha x}}{(-\alpha)^3} \right]_0^{\infty}$$

$$= \alpha \left\{ [0 - 0 + 0] - \left[ \frac{\mu^2}{-\alpha} - 2(-\mu) \frac{1}{\alpha^2} + 2 \frac{1}{\alpha^3} \right] \right\}$$

(19) Put  $\mu = \alpha$

$$= \mathcal{E}[\alpha - (\frac{\alpha^2}{\sigma^2} + \frac{\alpha}{\sigma^2} - \frac{2}{\sigma^3})]$$

$$= \alpha \left[ \frac{1}{\sigma^3} - \frac{2}{\sigma^3} + \frac{2}{\sigma^3} \right] = \frac{1}{\sigma^2}$$

$$\therefore \sigma^2 = \frac{1}{\alpha^2} \Rightarrow \boxed{S.D = \sigma = \frac{1}{\alpha}}$$

### Normal Distribution:

This is the most important continuous distribution because in applications many random variables are normal random variables.

The normal distribution is defined as the distribution with the density

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$

where  $\sigma > 0$

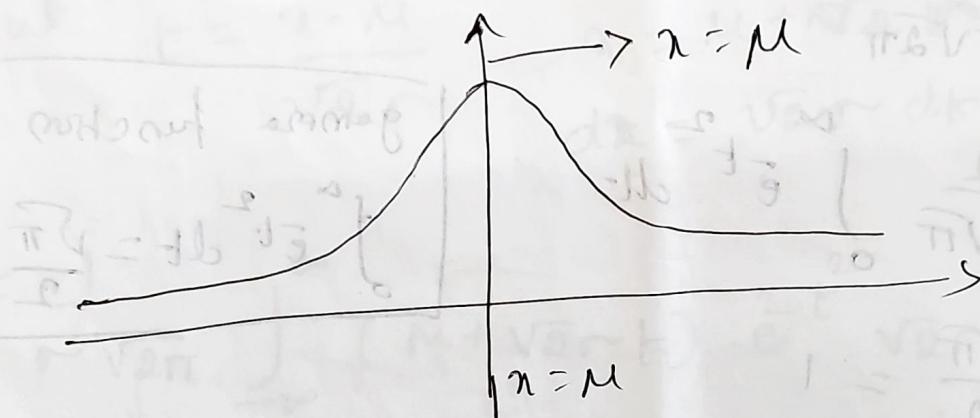
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Note: ①  $\mu$  is mean and  $\sigma$  s.d

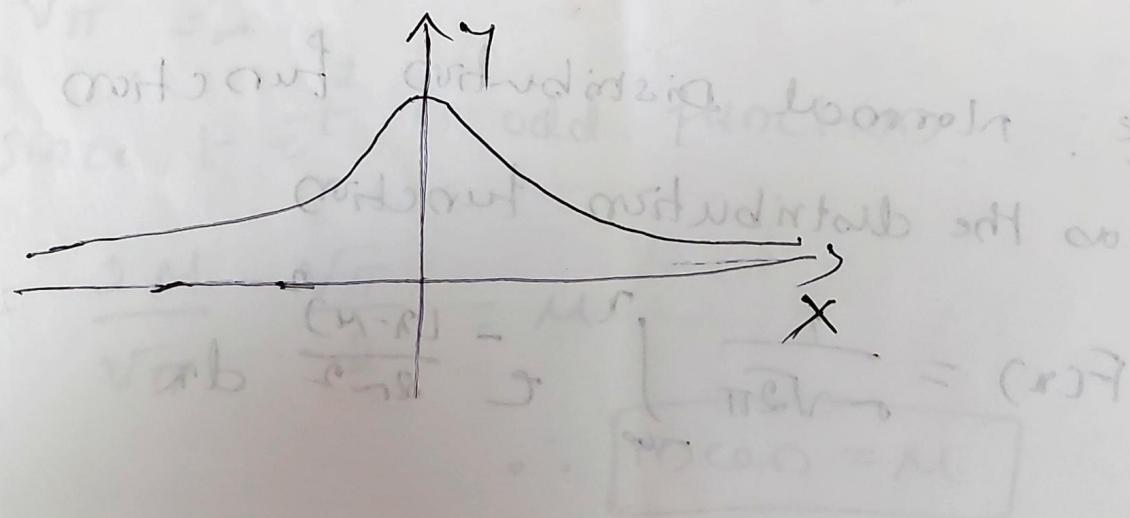
②  $\frac{1}{\sigma\sqrt{2\pi}}$  is a constant that makes the area under the curve equal to 1

$$\text{i.e. } \Rightarrow \int_{-\infty}^{\infty} f(x) dx = 1$$

③  $f(x)$  is symmetric about  $x=\mu$



for  $\mu=0$ ,  $f(x)$  is symmetric about Y axis



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Consider

$$\int_{-\infty}^{\infty} f(x) dx = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

put  $t = \frac{x-\mu}{\sigma\sqrt{2}}$   $x = \mu + \sigma\sqrt{2}t$   
 $dx = \sigma\sqrt{2} dt$

$$= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-t^2} \sigma\sqrt{2} dt$$

$$= \frac{2}{\sqrt{\pi}} \int_0^{\infty} e^{-t^2} dt$$

$$= \frac{2}{\sqrt{\pi}} \cdot \frac{\sqrt{\pi}}{2} = 1$$

gamma function

$$\int_0^{\infty} e^{-t^2} dt = \frac{\sqrt{\pi}}{2}$$

$$\therefore \int_{-\infty}^{\infty} f(x) dx = 1$$

Note: normal distribution function  
has the distribution function

$$F(x) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

# Mean and S.D of Normal Distribution

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W.K.T.

$$\text{mean} = \mu = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

put  $t = \frac{x-\mu}{\sigma}$   $x = \mu + \sigma t$   
 $dx = \sigma dt$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (\mu + \sigma t) e^{-t^2} \sigma dt$$

$$= \frac{\mu}{\sqrt{\pi}} \int_{-\infty}^{\infty} t e^{-t^2} dt + \sigma \sqrt{\frac{2\pi}{\pi}} \int_{-\infty}^{\infty} t^2 e^{-t^2} dt$$

Since  $t e^{-t^2}$  is odd function

$$= \frac{2\mu}{\sqrt{\pi}} \frac{\sqrt{\pi}}{2} = \mu$$

$\therefore \boxed{\text{mean} = \mu}$

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$$\text{var} = V = \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (x - \mu)^2 e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

put  $t = \frac{x - \mu}{\sigma}$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} 2\sigma^2 t^2 e^{t^2} \sqrt{2\pi} dt$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \int_{-\infty}^{\infty} t^2 e^{t^2} dt$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \left[ t \frac{e^{t^2}}{2} \right]_{-\infty}^{\infty} + \frac{2\sigma^2}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{t^2} dt$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \frac{\sqrt{\pi}}{2} = \sigma$$

$$\therefore \boxed{S.D = \sigma}$$

$$\boxed{\mu = 0.0307 \dots}$$

## standard Normal Distribution

(ii)

The distribution function  $F(x)$  of the normal distribution with any parameters is related to the standardized distribution function  $\phi(z)$ .

$$F(x) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{(u-\mu)^2}{2\sigma^2}} du$$

Put  $t = \frac{u-\mu}{\sigma} \Rightarrow dt = \frac{du}{\sigma}$

at ~~u~~  $u=x \Rightarrow t = \frac{x-\mu}{\sigma}$

$$\begin{aligned} \therefore F &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\frac{(x-\mu)}{\sigma}} e^{-t^2/2} dt \\ &= \phi\left[\frac{x-\mu}{\sigma}\right] \end{aligned}$$

put  $\boxed{z = \frac{x-\mu}{\sigma}}$

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$$\therefore \phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-z^2/2} dz$$

$\frac{1}{\sqrt{2\pi}} e^{-z^2/2}$  is called standard normal probability density function.

If  $\mu=0$  &  $\sigma=1$  Then normal probability density function is the standard normal probability density function.

In particular we have

$$\phi(z) = \frac{1}{\sqrt{2\pi}} \int_0^z e^{-z^2/2} dz$$

\*  $z = \frac{x-\mu}{\sigma}$  is called standard normal variable

$$\frac{x-\mu}{\sigma} \sim N(0, 1)$$

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Note:

① The probability that a normal random variable is  $x$  with  $\mu \neq \sigma$ .

i.e.

 $a < x \leq b$  then

$$P(a < x \leq b) = \phi\left(\frac{b-\mu}{\sigma}\right) - \phi\left(\frac{a-\mu}{\sigma}\right)$$

i.e. at  $x=a$ ,  $z_1 = \frac{a-\mu}{\sigma}$

$$x=b \quad z_2 = \frac{b-\mu}{\sigma}$$

$$P(a < x \leq b) = F(b) - F(a) = \phi(z_2) - \phi(z_1)$$

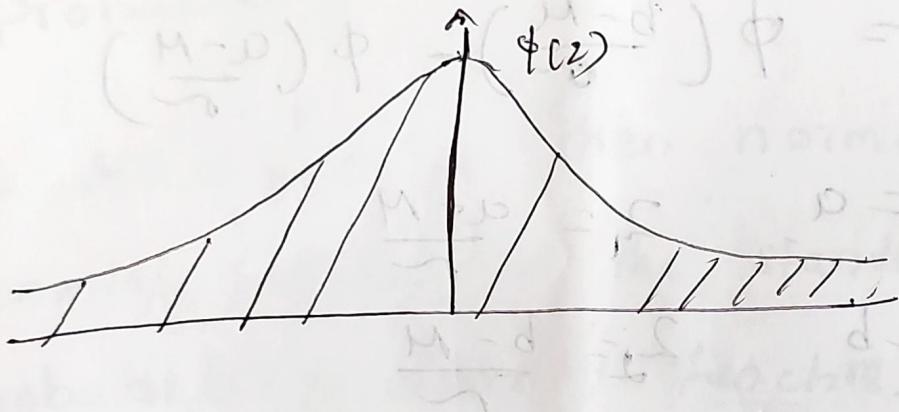
$$\Rightarrow P(z_1 < z < z_2) = \phi(z_2) - \phi(z_1)$$

(15)

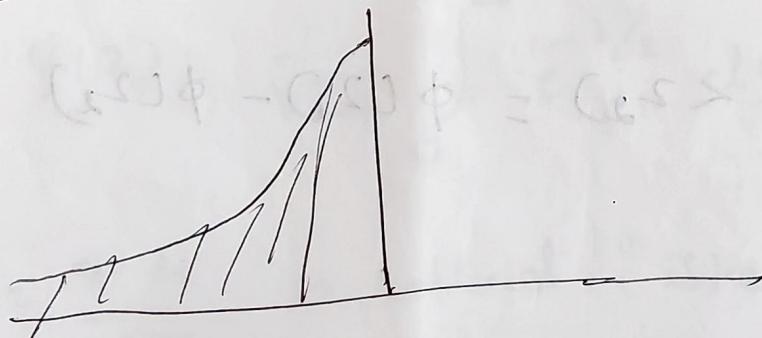
Note: Let  $\phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt$  ①

then

$$\textcircled{1} \quad P(-\infty \leq z \leq \infty) = \int_{-\infty}^{\infty} \phi(z) dz = 1$$



$$\textcircled{2} \quad \int_{-\infty}^0 \phi(z) dz = P(-\infty \leq z \leq 0) = \gamma_2$$

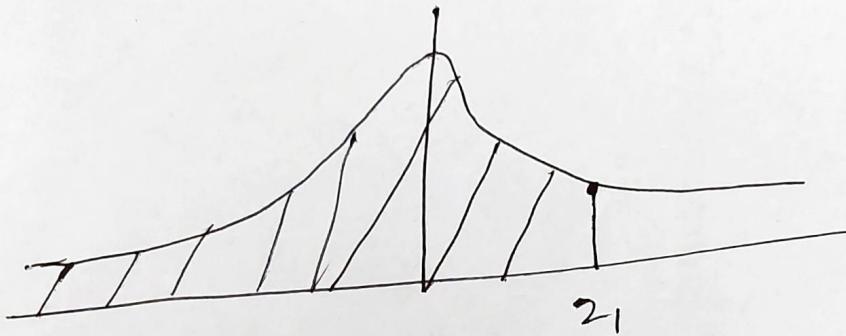


$$\textcircled{3} \quad \int_0^\infty \phi(z) dz = P(0 \leq z \leq \infty) = \gamma_2$$

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$$P(-\infty < Z < z_1) = P(-\infty < Z \leq 0) + P(0 < Z < z_1)$$

$$P(Z < z_1) = 0.5 + \phi(z_1)$$



(ii)  $P(Z > z_2) = P(Z \geq 0) - P(0 \leq Z \leq z_2)$   
 $= 0.5 - \phi(z_2)$

