Time Series Analysis on Airline-Passengers Dataset:

Here we use 12 Years of airline passengers data to build a model that could forecast for the next 1 year the number of Passengers boarding the plane every month.

```
In [2]: import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns

In [3]: df_airline = pd.read_csv("..\\Statistics_and_ML\\dataset\\airline-passengers.csv")

In [4]: df_airline.head()
```

Out[4]:

	Month	Passengers
0	1949-01	112
1	1949-02	118
2	1949-03	132
3	1949-04	129
4	1949-05	121

```
In [5]: df_airline.tail()
```

Out[5]:

	Month	Passengers
139	1960-08	606
140	1960-09	508
141	1960-10	461
142	1960-11	390
143	1960-12	432

We have 12 Years of Airline Passenger data starting from the year Jan 1949 to Dec 1960. The periodicity of the observations is 1 month.

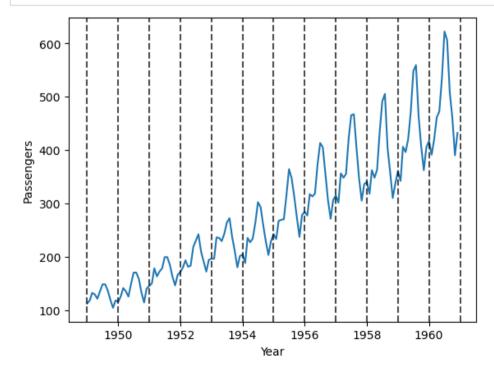
No missing values detected.

Plotting the time series:

```
In [11]: plt.figure()
plt.plot(df_airline)
plt.xlabel('Year')
plt.ylabel('Passengers')

# matplotlib.pyplot.axvline(x=0, ymin=0, ymax=1, **kwargs)
# Add a vertical line across the Axes.
# x position in data coordinates of the vertical line
# https://matplotlib.org/stable/api/_as_gen/matplotlib.pyplot.axvline.html
# Since the index for the time series has been set to DateTimeIndex therefore
# each x-coordinate is of type 'pandas_libs.tslibs.timestamps.Timestamp'.
for year in range(1949, 1962):
    plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

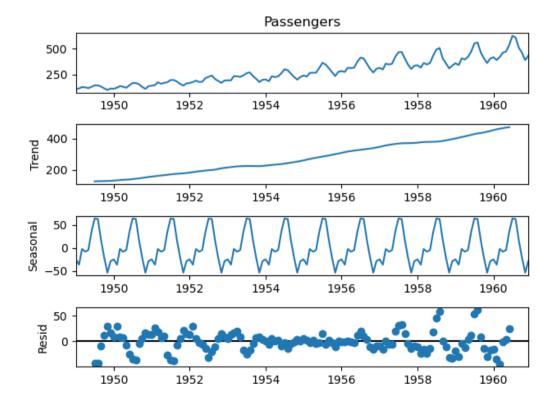
plt.show()
```



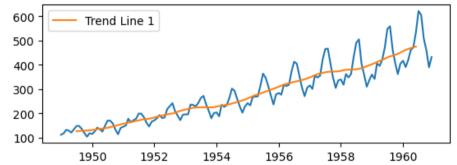
Inference: Data clearly shows an increasing trend and seasonality. Seasonality is easily seen by the occurrence of peaks at seemingly regular intervals of time. Seasonal frequency is 12 periods or 12 months.

Potting the trend, seasonal and residual components separately:

<Figure size 640x480 with 0 Axes>



```
In [14]: plt.figure()
    plt.subplot(2,1,1)
    plt.plot(df_airline['Passengers'])
    plt.plot(ts_decomposed_result.trend, label="Trend Line 1")
    plt.legend()
    plt.show()
```



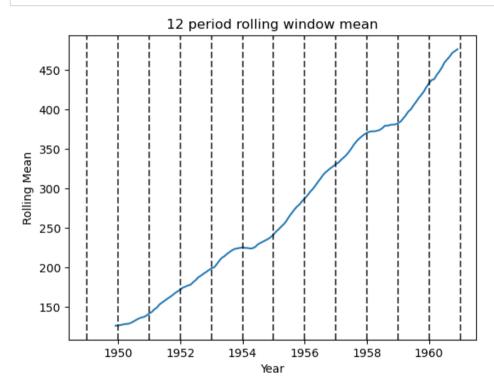
Checking for constant mean and constant variance in the series using Rolling Window:

```
In [15]: rolling window period = 12
         rolling mean = df airline['Passengers'].rolling(rolling window period).mean()
         rolling mean
Out[15]: Month
         1949-01-01
                              NaN
         1949-02-01
                              NaN
         1949-03-01
                              NaN
         1949-04-01
                              NaN
         1949-05-01
                              NaN
                           . . .
         1960-08-01
                       463.333333
         1960-09-01
                       467.083333
         1960-10-01
                       471.583333
         1960-11-01
                       473.916667
         1960-12-01
                       476.166667
         Freq: MS, Name: Passengers, Length: 144, dtype: float64
In [16]: rolling_mean.isna().sum()
Out[16]: 11
```

```
In [17]: plt.figure()
   plt.plot(rolling_mean)
   plt.xlabel('Year')
   plt.ylabel('Rolling Mean')
   plt.title(f'{rolling_window_period} period rolling window mean')

for year in range(1949, 1962):
      plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

plt.show()
```



```
In [18]: rolling_mean.head(30)
Out[18]: Month
         1949-01-01
                              NaN
         1949-02-01
                              NaN
         1949-03-01
                              NaN
         1949-04-01
                              NaN
         1949-05-01
                              NaN
         1949-06-01
                              NaN
         1949-07-01
                              NaN
         1949-08-01
                              NaN
         1949-09-01
                              NaN
         1949-10-01
                              NaN
         1949-11-01
                              NaN
         1949-12-01
                       126.666667
         1950-01-01
                       126.916667
         1950-02-01
                       127.583333
         1950-03-01
                       128.333333
         1950-04-01
                       128.833333
         1950-05-01
                       129.166667
         1950-06-01
                       130.333333
         1950-07-01
                       132.166667
         1950-08-01
                       134.000000
         1950-09-01
                       135.833333
         1950-10-01
                       137.000000
         1950-11-01
                       137.833333
         1950-12-01
                       139.666667
         1951-01-01
                       142.166667
         1951-02-01
                       144.166667
         1951-03-01
                       147.250000
         1951-04-01
                       149.583333
         1951-05-01
                       153.500000
         1951-06-01
                       155.916667
```

Freq: MS, Name: Passengers, dtype: float64

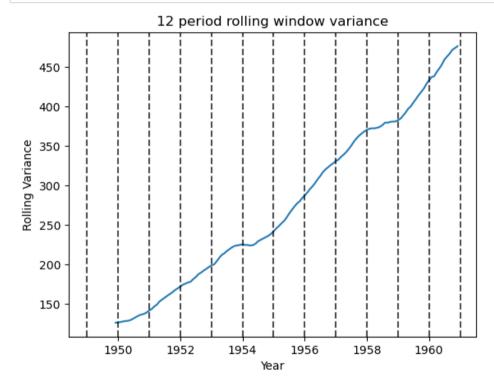
Mean is not constant across time it varies greatly w.r.t. time which again confirms that the time series is not Stationary.

```
In [19]: rolling variance = df airline['Passengers'].rolling(rolling window period).var()
         rolling_variance
Out[19]: Month
         1949-01-01
                               NaN
         1949-02-01
                               NaN
         1949-03-01
                               NaN
         1949-04-01
                               NaN
         1949-05-01
                               NaN
                          ...
         1960-08-01
                       6994.060606
         1960-09-01
                       7160.083333
         1960-10-01
                       6813.174242
         1960-11-01
                       6320.628788
         1960-12-01
                       6043.060606
         Freq: MS, Name: Passengers, Length: 144, dtype: float64
In [20]: rolling variance.isna().sum()
Out[20]: 11
```

```
In [21]: plt.figure()
   plt.plot(rolling_mean)
   plt.xlabel('Year')
   plt.ylabel('Rolling Variance')
   plt.title(f'{rolling_window_period} period rolling window variance')

for year in range(1949, 1962):
      plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

plt.show()
```



Variance is not constant across time it varies greatly w.r.t. time which again confirms that the time series is not Stationary.

In [22]: from statsmodels.tsa.stattools import adfuller

```
In [23]: def perform_adf_test(series, regression='c'):
    keys = ['adf', 'pvalue', 'usedlag', 'nobs', 'critical values', 'icbest', 'resstore']
    result = adfuller(series, regression=regression)
    for i in range(len(result)):
        print(f"{keys[i]}: {result[i]}")
```

Model used by the ADF(Augmented Dickey Fuller) test is:

```
\Delta y_t = \alpha + \beta t + \gamma y_{t-1} + \delta_1 \Delta y_{t-1} + \dots + \delta_{p-1} \Delta y_{t-p+1} + e_t
```

 H_0 : γ (Co-efficient of the term y_{t-1}) = 0

 $H_1: \gamma < 0$

Let us assume the significance level to be 0.05 i.e. 5%.

```
In [24]: perform_adf_test(df_airline['Passengers'], 'ct')

adf: -2.1007818138446694
pvalue: 0.5456589343124553
usedlag: 13
nobs: 130
critical values: {'1%': -4.030152423759672, '5%': -3.444817634956759, '10%': -3.1471816659080565}
icbest: 993.2814778200581
```

Since the p-value is > 0.05 (our assumed significance level) this implies that we can't reject the null hypothesis H_0 . So the time series is non-stationary.

Approach 1 to find parameters for (p,d,q,P,D,Q) for SARIMA model:

Apply non-seasonal differencing i.e. perform the operation (1-L) * Y_t to get Z_t where L is the Lag operator:

```
In [25]: df1 = df_airline.copy()
    df1["Passengers_diff_1"] = df_airline['Passengers'].diff(1)
In [26]: df1["Passengers_diff_1"].isna().sum()
Out[26]: 1
```

```
In [27]: df1 = df1.dropna()
```

In [28]: df1

Out[28]:

Passengers Passengers_diff_1

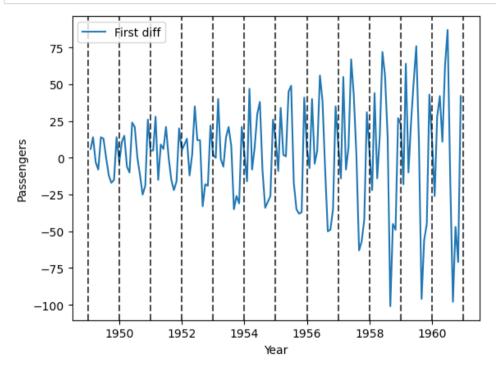
Month		
1949-02-01	118	6.0
1949-03-01	132	14.0
1949-04-01	129	-3.0
1949-05-01	121	-8.0
1949-06-01	135	14.0
1960-08-01	606	-16.0
1960-09-01	508	-98.0
1960-10-01	461	-47.0
1960-11-01	390	-71.0
1960-12-01	432	42.0

143 rows × 2 columns

```
In [29]: plt.figure()
    # plt.plot(df1['Passengers'], label='Actual')
    plt.plot(df1['Passengers_diff_1'], label='First diff')
    plt.xlabel('Year')
    plt.ylabel('Passengers')
    plt.legend()

for year in range(1949, 1962):
        plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

plt.show()
```



```
In [30]: perform_adf_test(df1['Passengers_diff_1'])
```

adf: -2.8292668241700034 pvalue: 0.054213290283824954

usedlag: 12 nobs: 130

critical values: {'1%': -3.4816817173418295, '5%': -2.8840418343195267, '10%': -2.578770059171598}

icbest: 988.5069317854084

Again p-value is > 0.05 which implies that the first difference time series is still non-stationary.

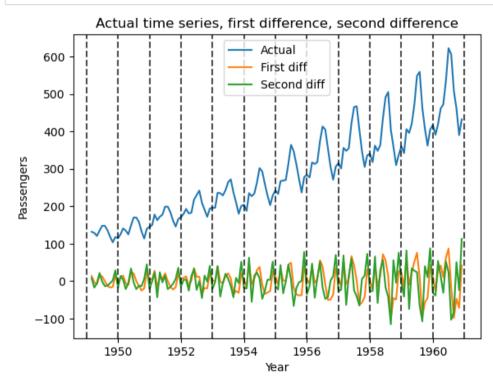
Apply non-seasonal differencing again on Z_t i.e. perform the operation (1-L) * Z_t to get W_t:

```
In [31]: df2 = df1.copy()
          df2["Passengers diff 2"] = df1['Passengers diff 1'].diff(1)
In [32]: df2["Passengers diff 2"].isna().sum()
Out[32]: 1
In [33]: df2 = df2.dropna()
In [34]: df2
Out[34]:
                      Passengers Passengers_diff_1 Passengers_diff_2
               Month
                             132
                                              14.0
                                                                8.0
           1949-03-01
           1949-04-01
                             129
                                              -3.0
                                                               -17.0
           1949-05-01
                             121
                                              -8.0
                                                               -5.0
           1949-06-01
                             135
                                              14.0
                                                               22.0
           1949-07-01
                             148
                                              13.0
                                                               -1.0
           1960-08-01
                             606
                                             -16.0
                                                             -103.0
           1960-09-01
                                             -98.0
                                                               -82.0
                             508
           1960-10-01
                                             -47.0
                                                               51.0
                             461
                                             -71.0
           1960-11-01
                             390
                                                               -24.0
           1960-12-01
                             432
                                              42.0
                                                              113.0
          142 rows × 3 columns
```

```
In [35]: plt.figure()
    plt.plot(df2['Passengers'], label='Actual')
    plt.plot(df2['Passengers_diff_1'], label='First diff')
    plt.plot(df2['Passengers_diff_2'], label='Second diff')
    plt.xlabel('Year')
    plt.ylabel('Passengers')
    plt.title('Actual time series, first difference, second difference')
    plt.legend()

for year in range(1949, 1962):
    plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

plt.show()
```



```
In [36]: perform_adf_test(df2['Passengers_diff_2'])

adf: -16.384231542468488
    pvalue: 2.7328918500143186e-29
    usedlag: 11
    nobs: 130
    critical values: {'1%': -3.4816817173418295, '5%': -2.8840418343195267, '10%': -2.578770059171598}
    icbest: 988.6020417275604

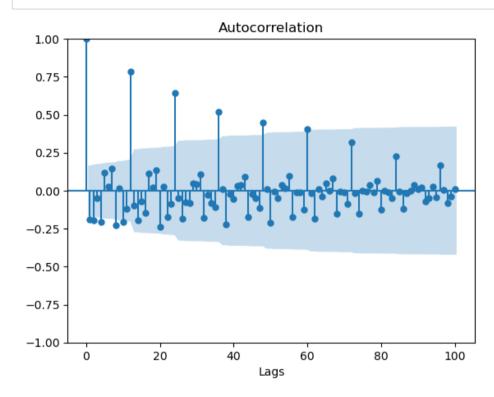
Since the p-value is almost 0 which is < 0.05(our assumed significance level) this implies that we can reject</pre>
```

Analyze the ACF and PACF of the series W_t = (1-L)^2 * Y_t:

the null hypothesis H_0 . So after differencing the time series 2 times the resulting series is now stationary.

In [37]: from statsmodels.graphics.tsaplots import plot_acf, plot_pacf

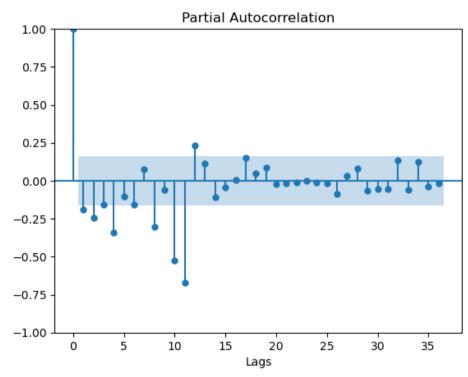
```
In [38]: plot_acf(df2['Passengers_diff_2'], lags = 100)
    plt.xlabel('Lags')
    plt.show()
```



The autocorrelation is decreasing at multiples of seasonal lags. This implies that more number of seasonal moving average terms are required. We could try seasonal lags of 1,2,3,4,5 as values for Q.

After Y_t-1 and Y_t-2 non-seasonal lags the autocorrelation reduces rapidly indicating that the first 2 non-seasonal moving average terms are required i.e. q=2.

```
In [39]: plot_pacf(df2['Passengers_diff_2'], method='ywm', lags=36)
plt.xlabel('Lags')
plt.show()
```



For lags 2,4,8 we see spikes that are significant. So we could try these values for p.

Only at seasonal lag at 12 we see a spike that is significant. So P=1.

Approach 2 to find parameters for (p,d,q,P,D,Q) for SARIMA model:

Applying seasonal differencing on the original time series Y_t to get Z_t i.e. perform the operation (1-L^12) * Y_t where L is the Lag operator:

```
In [43]: df3
```

Out[43]:

	Passengers	Passengers_seasonal_diff
Month		
1950-01-01	115	3.0
1950-02-01	126	8.0
1950-03-01	141	9.0
1950-04-01	135	6.0
1950-05-01	125	4.0
1960-08-01	606	47.0
1960-09-01	508	45.0
1960-10-01	461	54.0
1960-11-01	390	28.0

432

27.0

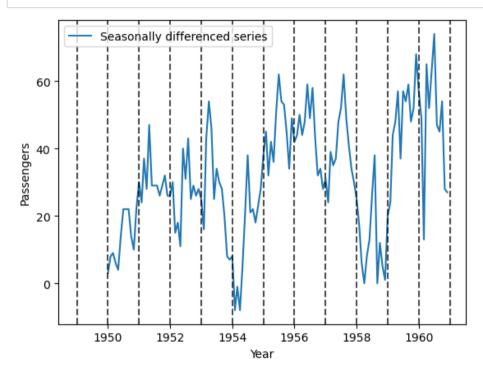
132 rows × 2 columns

1960-12-01

```
In [44]: plt.figure()
    # ptt.plot(df3['Passengers'], label='Actual')
    plt.plot(df3['Passengers_seasonal_diff'], label='Seasonally differenced series')
    plt.xlabel('Year')
    plt.ylabel('Passengers')
    plt.legend()

for year in range(1949, 1962):
        plt.axvline(pd.to_datetime(str(year)), color='k', linestyle='--', alpha=0.7)

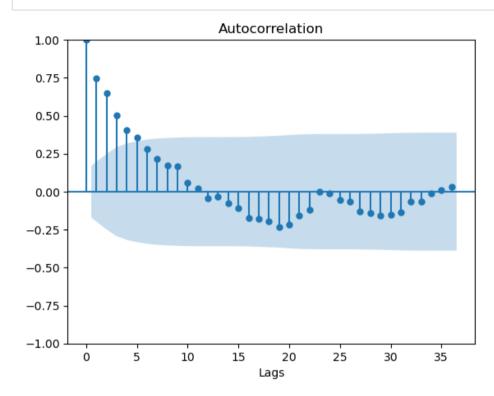
plt.show()
```



```
In [45]: df airline['Passengers'].diff(12).head(30)
Out[45]: Month
         1949-01-01
                        NaN
         1949-02-01
                        NaN
         1949-03-01
                        NaN
         1949-04-01
                        NaN
         1949-05-01
                        NaN
         1949-06-01
                        NaN
         1949-07-01
                        NaN
         1949-08-01
                        NaN
         1949-09-01
                        NaN
         1949-10-01
                        NaN
         1949-11-01
                        NaN
         1949-12-01
                        NaN
         1950-01-01
                        3.0
         1950-02-01
                        8.0
         1950-03-01
                        9.0
         1950-04-01
                        6.0
         1950-05-01
                        4.0
         1950-06-01
                       14.0
         1950-07-01
                       22.0
         1950-08-01
                       22.0
         1950-09-01
                       22.0
         1950-10-01
                       14.0
         1950-11-01
                       10.0
         1950-12-01
                       22.0
         1951-01-01
                       30.0
         1951-02-01
                       24.0
         1951-03-01
                       37.0
         1951-04-01
                       28.0
         1951-05-01
                       47.0
         1951-06-01
                       29.0
         Freq: MS, Name: Passengers, dtype: float64
In [46]: perform_adf_test(df3['Passengers_seasonal_diff'])
         adf: -3.3830207264924805
         pvalue: 0.011551493085514982
         usedlag: 1
         nobs: 130
         critical values: {'1%': -3.4816817173418295, '5%': -2.8840418343195267, '10%': -2.578770059171598}
```

icbest: 919.527129208137

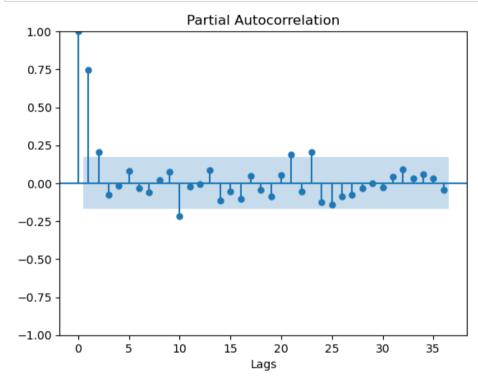
```
In [47]: plot_acf(df3['Passengers_seasonal_diff'], lags=36)
    plt.xlabel('Lags')
    plt.show()
```



For non-seasonal MA terms possible values of q are 1,2,3,4,5 after which the lag terms become insignificant.

For seasonal MA terms, we do not see any significant spike at any seasonal lags that is at lag 12 or lag 24 etc. So Q=0.

```
In [48]: plot_pacf(df3['Passengers_seasonal_diff'], method='ywm', lags=36)
plt.xlabel('Lags')
plt.show()
```



For non-seasonal AR terms, possible values of p are 1, 2 after which the lag terms become insignificant.

For seasonal AR terms, we do not see any significant spike at any seasonal lags that is at lag 12 or lag 24 etc. So P=0.

```
In [49]: perform_adf_test(df3['Passengers_seasonal_diff'])

adf: -3.3830207264924805
pvalue: 0.011551493085514982
usedlag: 1
nobs: 130
critical values: {'1%': -3.4816817173418295, '5%': -2.8840418343195267, '10%': -2.578770059171598}
icbest: 919.527129208137
```

Getting the training and test data for trying out time series models:

```
In [50]: from datetime import timedelta
```

Using the data from 1 Jan 1949 to 31 Dec 1959 as training data.

Using the data from 1 Jan 1960 to 31 Dec 1960 as training data.

```
In [51]: train_end = pd.to_datetime("1959-12-31")
    test_end = pd.to_datetime("1960-12-31")
    train_data = df_airline[:train_end]
    test_data = df_airline[train_end + timedelta(days=1):test_end]
```

In [52]: train_data

Out[52]:

Passengers

Month	
1949-01-01	112
1949-02-01	118
1949-03-01	132
1949-04-01	129
1949-05-01	121
1959-08-01	559
1959-09-01	463
1959-10-01	407
1959-11-01	362
1959-12-01	405

132 rows × 1 columns

```
In [53]: test_data
Out[53]:
```

	Passengers
Month	
1960-01-01	417
1960-02-01	391
1960-03-01	419
1960-04-01	461
1960-05-01	472
1960-06-01	535
1960-07-01	622
1960-08-01	606
1960-09-01	508
1960-10-01	461
1960-11-01	390
1960-12-01	432

Approach 3 using grid search to find parameters for (p,d,q,P,D,Q) for SARIMA model:

```
In [54]: import itertools
In [55]: p = q = d = range(0,2)
In [56]: # itertools.product(p,d,q)
         # Returns the cartesian product of input iterables. Equivalent to nested for-loops.
         pdq = list(itertools.product(p,d,q))
         print("pdq combinations:")
         print(pdq)
         pdq combinations:
         [(0, 0, 0), (0, 0, 1), (0, 1, 0), (0, 1, 1), (1, 0, 0), (1, 0, 1), (1, 1, 0), (1, 1, 1)]
```

```
(0, 0, 0) \times (0, 0, 0, 12) - AIC: 1853.867132204165
(0, 0, 0) \times (0, 0, 1, 12) - AIC: 1575.2927100520965
(0, 0, 0) \times (0, 1, 0, 12) - AIC: 1183.7768361260032
(0, 0, 0) \times (0, 1, 1, 12) - AIC: 1039.3684612710686
(0, 0, 0) \times (1, 0, 0, 12) - AIC: 1018.1128824449078
(0, 0, 0) \times (1, 0, 1, 12) - AIC: 1008.0082399524545
(0, 0, 0) \times (1, 1, 0, 12) - AIC: 999.189649490497
(0, 0, 0) \times (1, 1, 1, 12) - AIC: 930.7565628459528
(0, 0, 1) \times (0, 0, 0, 12) - AIC: 1670.9731705571419
(0, 0, 1)x(0, 0, 1, 12) - AIC: 1409.6308436189208
(0, 0, 1) \times (0, 1, 0, 12) - AIC: 1071.9868359596408
(0, 0, 1) \times (0, 1, 1, 12) - AIC: 954.9732785958571
(0, 0, 1)x(1, 0, 0, 12) - AIC: 962.7908430127964
(0, 0, 1)x(1, 0, 1, 12)) - AIC: 941.9746021501135
(0, 0, 1) \times (1, 1, 0, 12) - AIC: 947.4386036577736
(0, 0, 1)x(1, 1, 1, 12)) - AIC: 876.8976472270055
(0, 1, 0)x(0, 0, 0, 12)) - AIC: 1267.4912483529856
(0, 1, 0)x(0, 0, 1, 12)) - AIC: 1066.5001391578778
(0, 1, 0) \times (0, 1, 0, 12) - AIC: 898.2577377990535
(0, 1, 0) \times (0, 1, 1, 12) - AIC: 814.9661546286915
(0, 1, 0)x(1, 0, 0, 12)) - AIC: 901.900047525677
(0, 1, 0) \times (1, 0, 1, 12) - AIC: 882.2776805554837
(0, 1, 0)x(1, 1, 0, 12)) - AIC: 821.6405586698072
(0, 1, 0) \times (1, 1, 1, 12) - AIC: 815.1388167150715
(0, 1, 1)x(0, 0, 0, 12)) - AIC: 1245.6319844870782
(0, 1, 1)x(0, 0, 1, 12)) - AIC: 1053.9248192171872
(0, 1, 1)x(0, 1, 0, 12)) - AIC: 887.2936971142087
(0, 1, 1) \times (0, 1, 1, 12) - AIC: 803.4111310810478
(0, 1, 1)x(1, 0, 0, 12)) - AIC: 892.8874405001393
(0, 1, 1)x(1, 0, 1, 12)) - AIC: 863.2617743136642
(0, 1, 1)x(1, 1, 0, 12)) - AIC: 817.7933920131462
(0, 1, 1)x(1, 1, 1, 12)) - AIC: 802.8580194673696
(1, 0, 0) \times (0, 0, 0, 12) - AIC: 1278.2458443888524
(1, 0, 0) \times (0, 0, 1, 12) - AIC: 1077.0858909333813
(1, 0, 0) \times (0, 1, 0, 12) - AIC: 905.752128898712
(1, 0, 0) \times (0, 1, 1, 12) - AIC: 822.7598554087126
(1, 0, 0) \times (1, 0, 0, 12) - AIC: 893.5742879078161
(1, 0, 0)x(1, 0, 1, 12)) - AIC: 880.2330151275008
(1, 0, 0) \times (1, 1, 0, 12) - AIC: 822.5432823654064
(1, 0, 0) \times (1, 1, 1, 12) - AIC: 831.3170042691104
(1, 0, 1) \times (0, 0, 0, 12) - AIC: 1256.5507739583904
(1, 0, 1) \times (0, 0, 1, 12) - AIC: 1064.6671829896356
(1, 0, 1)x(0, 1, 0, 12)) - AIC: 895.4819722663491
(1, 0, 1) \times (0, 1, 1, 12) - AIC: 812.9279426161573
(1, 0, 1)x(1, 0, 0, 12)) - AIC: 889.6927487827884
(1, 0, 1) \times (1, 0, 1, 12) - AIC: 867.7584454866541
(1, 0, 1)x(1, 1, 0, 12)) - AIC: 819.4447145085644
(1, 0, 1)x(1, 1, 1, 12)) - AIC: 819.4243737233215
(1, 1, 0) \times (0, 0, 0, 12) - AIC: 1257.9634107014663
(1, 1, 0)x(0, 0, 1, 12)) - AIC: 1063.0736082639758
(1, 1, 0) \times (0, 1, 0, 12) - AIC: 893.0941782749939
```

```
(1, 1, 0)x(0, 1, 1, 12)) - AIC: 810.3950674867854
(1, 1, 0)x(1, 0, 0, 12)) - AIC: 884.6501491722378
(1, 1, 0)x(1, 0, 1, 12)) - AIC: 869.0250368828597
(1, 1, 0)x(1, 1, 0, 12)) - AIC: 810.1697343725327
(1, 1, 0)x(1, 1, 1, 12)) - AIC: 810.0538746711336
(1, 1, 1)x(0, 0, 0, 12)) - AIC: 1240.0096159726365
(1, 1, 1)x(0, 0, 1, 12)) - AIC: 1047.047612871338
(1, 1, 1)x(0, 1, 0, 12)) - AIC: 887.5809544179544
(1, 1, 1)x(0, 1, 1, 12)) - AIC: 803.7012277205246
(1, 1, 1)x(1, 0, 0, 12)) - AIC: 886.2868951863057
(1, 1, 1)x(1, 1, 0, 12)) - AIC: 864.4184468797822
(1, 1, 1)x(1, 1, 0, 12)) - AIC: 811.2171076081977
(1, 1, 1)x(1, 1, 1, 12)) - AIC: 803.3627295936262
```

```
In [61]: # Sorting the dictionary according to keys.
aic_dict = dict(sorted(aic_dict.items(), key=lambda item: item[1]))
aic_dict
```

```
Out[61]: {((0, 1, 1), (1, 1, 1, 12)): 802.8580194673696,
          ((1, 1, 1), (1, 1, 1, 12)): 803.3627295936262,
          ((0, 1, 1), (0, 1, 1, 12)): 803.4111310810478,
          ((1, 1, 1), (0, 1, 1, 12)): 803.7012277205246,
          ((1, 1, 0), (1, 1, 1, 12)): 810.0538746711336,
          ((1, 1, 0), (1, 1, 0, 12)): 810.1697343725327,
          ((1, 1, 0), (0, 1, 1, 12)): 810.3950674867854,
          ((1, 1, 1), (1, 1, 0, 12)): 811.2171076081977,
          ((1, 0, 1), (0, 1, 1, 12)): 812.9279426161573,
          ((0, 1, 0), (0, 1, 1, 12)): 814.9661546286915,
          ((0, 1, 0), (1, 1, 1, 12)): 815.1388167150715,
          ((0, 1, 1), (1, 1, 0, 12)): 817.7933920131462,
          ((1, 0, 1), (1, 1, 1, 12)): 819.4243737233215,
          ((1, 0, 1), (1, 1, 0, 12)): 819.4447145085644,
          ((0, 1, 0), (1, 1, 0, 12)): 821.6405586698072,
          ((1, 0, 0), (1, 1, 0, 12)): 822.5432823654064,
          ((1, 0, 0), (0, 1, 1, 12)): 822.7598554087126,
          ((1, 0, 0), (1, 1, 1, 12)): 831.3170042691104,
          ((0, 1, 1), (1, 0, 1, 12)): 863.2617743136642,
          ((1, 1, 1), (1, 0, 1, 12)): 864.4184468797822,
          ((1, 0, 1), (1, 0, 1, 12)): 867.7584454866541,
          ((1, 1, 0), (1, 0, 1, 12)): 869.0250368828597,
          ((0, 0, 1), (1, 1, 1, 12)): 876.8976472270055,
          ((1, 0, 0), (1, 0, 1, 12)): 880.2330151275008,
          ((0, 1, 0), (1, 0, 1, 12)): 882.2776805554837,
          ((1, 1, 0), (1, 0, 0, 12)): 884.6501491722378,
          ((1, 1, 1), (1, 0, 0, 12)): 886.2868951863057,
          ((0, 1, 1), (0, 1, 0, 12)): 887.2936971142087,
          ((1, 1, 1), (0, 1, 0, 12)): 887.5809544179544,
          ((1, 0, 1), (1, 0, 0, 12)): 889.6927487827884,
          ((0, 1, 1), (1, 0, 0, 12)): 892.8874405001393,
          ((1, 1, 0), (0, 1, 0, 12)): 893.0941782749939,
          ((1, 0, 0), (1, 0, 0, 12)): 893.5742879078161,
          ((1, 0, 1), (0, 1, 0, 12)): 895.4819722663491,
          ((0, 1, 0), (0, 1, 0, 12)): 898.2577377990535,
          ((0, 1, 0), (1, 0, 0, 12)): 901.900047525677,
          ((1, 0, 0), (0, 1, 0, 12)): 905.752128898712,
          ((0, 0, 0), (1, 1, 1, 12)): 930.7565628459528,
          ((0, 0, 1), (1, 0, 1, 12)): 941.9746021501135,
          ((0, 0, 1), (1, 1, 0, 12)): 947.4386036577736,
          ((0, 0, 1), (0, 1, 1, 12)): 954.9732785958571,
          ((0, 0, 1), (1, 0, 0, 12)): 962.7908430127964,
          ((0, 0, 0), (1, 1, 0, 12)): 999.189649490497,
          ((0, 0, 0), (1, 0, 1, 12)): 1008.0082399524545,
          ((0, 0, 0), (1, 0, 0, 12)): 1018.1128824449078,
          ((0, 0, 0), (0, 1, 1, 12)): 1039.3684612710686,
          ((1, 1, 1), (0, 0, 1, 12)): 1047.047612871338,
          ((0, 1, 1), (0, 0, 1, 12)): 1053.9248192171872,
          ((1, 1, 0), (0, 0, 1, 12)): 1063.0736082639758,
          ((1, 0, 1), (0, 0, 1, 12)): 1064.6671829896356,
          ((0, 1, 0), (0, 0, 1, 12)): 1066.5001391578778,
```

```
((0, 0, 1), (0, 1, 0, 12)): 1071.9868359596408,
          ((1, 0, 0), (0, 0, 1, 12)): 1077.0858909333813,
          ((0, 0, 0), (0, 1, 0, 12)): 1183.7768361260032,
          ((1, 1, 1), (0, 0, 0, 12)): 1240.0096159726365,
          ((0, 1, 1), (0, 0, 0, 12)): 1245.6319844870782,
          ((1, 0, 1), (0, 0, 0, 12)): 1256.5507739583904,
          ((1, 1, 0), (0, 0, 0, 12)): 1257.9634107014663,
          ((0, 1, 0), (0, 0, 0, 12)): 1267.4912483529856,
          ((1, 0, 0), (0, 0, 0, 12)): 1278.2458443888524,
          ((0, 0, 1), (0, 0, 1, 12)): 1409.6308436189208,
          ((0, 0, 0), (0, 0, 1, 12)): 1575.2927100520965,
          ((0, 0, 1), (0, 0, 0, 12)): 1670.9731705571419,
          ((0, 0, 0), (0, 0, 0, 12)): 1853.867132204165}
In [62]: # (0,1,1) (1,1,1,12) - From Grid Search.
         # (2,2,2) (1,0,1,12) - From the acf and pacf of (1-L)^2 * Y_t series.
         \# (2,0,5) (0,1,0,12) - From the acf and pacf of (1-L^{12}) * Y t series.
         orders = [(0,1,1), (2,2,2), (2,0,5)]
         seasonal_orders = [(1,1,1,12), (1,0,1,12), (0,1,0,12)]
         results = []
         for i in range(len(orders)):
             sarima model = SARIMAX(train data,
                                             order=orders[i],
                                             seasonal order=seasonal orders[i],
                                             enforce stationarity=False,
                                             enforce invertibility=False
             result = sarima model.fit()
             results.append(result)
             print(f"{orders[i]}{seasonal orders[i]}")
             print(f"AIC: {result.aic}")
         (0, 1, 1)(1, 1, 1, 12)
         AIC: 802.8580194673696
```

(0, 1, 1)(1, 1, 1, 12) AIC: 802.8580194673696 (2, 2, 2)(1, 0, 1, 12) AIC: 860.4523841098804 (2, 0, 5)(0, 1, 0, 12) AIC: 866.6444026938915

ARIMA(0, 1, 1)(1, 1, 1)12 has the lowest AIC value out of the parameters we tried so we go with this model.

```
In [63]: for result in results:
    print(result.summary())
    print("\n\n")
```

SARIMAX Results

B	======================================	=======		N.	Ob		42
Dep. Variable:			Passengers No. Observations:				132
Model:	SARI			., 12) Log	Likelihood		-397.429
Date:			•	2023 AIC			802.85
Time:			11:	19:59 BIC			813.47
Sample:			01-01	1949 HQIC			807.16
			- 12-01	1959			
Covariance 1	Гуре:			opg			
========	coef	std err	z	P> z	[0.025	0.975]	
ma.L1	-0.2380	0.087	-2.735	0.006	-0.409	-0.067	
ar.S.L12	-0.6332	0.166	-3.825	0.000	-0.958	-0.309	
ma.S.L12	0.5695	0.252	2.263	0.024	0.076	1.063	
sigma2	109.1523	17.692	6.170	0.000	74.477	143.828	
Ljung-Box (l		=======	0.00	========= Jarque-Bera	:======= : (JB):	 9	.04
Prob(Q):			0.95 Prob(JB):		0.98		
Heteroskedasticity (H):			1.60 Skew:			-0.05	
Prob(H) (two-sided):			0.17 Kurtosis:			3.01	

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

SARIMAX Results

========	.=======		JANINA	========			.=======
Dep. Variab	ole:		Pas	sengers No	o. Observatio	ns:	132
Model:	SAR	IMAX(2, 2,	2)x(1, 0, [1], 12) Lo	og Likelihood		-423.226
Date:			Fri, 29 D	ec 2023 A	IC		860.452
Time:			1	1:19:59 B	IC		879.667
Sample:			01-	01-1949 H	SIC		868.251
			- 12-	01-1959			
Covariance	Type:			opg			
========	========		========	========		=======	
	coef	std err	Z	P> z	[0.025	0.975]	
ar.L1	-0.3159			0.799	·		
ar.L2	-0.0009	0.432	-0.002	0.998	-0.847	0.845	
ma.L1	-1.0464	189.126	-0.006	0.996	-371.726	369.633	
ma.L2	0.0464	8.608	0.005	0.996	-16.825	16.918	
ar.S.L12	1.1394	0.016	69.528	0.000	1.107	1.172	
ma.S.L12	-0.4882	0.112	-4.362	0.000	-0.708	-0.269	
sigma2	86.1357	1.63e+04	0.005	0.996	-3.19e+04	3.2e+04	
Ljung-Box (11) (0):		0.00	Jarque-Bera	========= a (]B):	 1.3	== RA
Prob(Q):	/ (4).		0.98	Prob(JB):	. (55).	0.5	
,	sticity (H)	•	1.40	Skew:		-0.2	
cc. oskeda	(11)	•	0	J		0.2	

```
Prob(H) (two-sided): 0.30 Kurtosis: 3.05
```

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

SARIMAX Results

______ Dep. Variable: Passengers No. Observations: 132 SARIMAX(2, 0, 5)x(0, 1, [], 12) Log Likelihood Model: -425.322 Date: Fri, 29 Dec 2023 AIC 866,644 Time: 11:19:59 BIC 888.534 Sample: 01-01-1949 HOIC 875.528

- 12-01-1959

Covariance Type:

========	=========		========	=======	.=======	========
	coef	std err	z	P> z	[0.025	0.975]
ar.L1	0.1711	0.121	1.417	0.156	-0.065	0.408
ar.L2	0.8223	0.119	6.918	0.000	0.589	1.055
ma.L1	0.6620	495.205	0.001	0.999	-969.922	971.246
ma.L2	-0.0906	167.371	-0.001	1.000	-328.132	327.950
ma.L3	-0.0887	122.493	-0.001	0.999	-240.170	239.993
ma.L4	-0.3282	166.385	-0.002	0.998	-326.438	325.781
ma.L5	0.0078	3.835	0.002	0.998	-7.508	7.523
sigma2	99.3506	4.92e+04	0.002	0.998	-9.63e+04	9.65e+04
========	=========		========	========	=========	========

Ljung-Box (L1) (Q): 0.03 Jarque-Bera (JB): 0.61
Prob(Q): 0.86 Prob(JB): 0.74
Heteroskedasticity (H): 1.24 Skew: -0.12
Prob(H) (two-sided): 0.52 Kurtosis: 3.27

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

```
In [64]: # Getting the start and end dates for prediction.
    pred_start_date = test_data.index[0]
    pred_end_date = test_data.index[-1]
    print(f'pred_start_date: {pred_start_date}')
    print(f'pred_end_date: {pred_end_date}')
```

pred_start_date: 1960-01-01 00:00:00
pred_end_date: 1960-12-01 00:00:00

```
In [65]: # predictions is a pandas Series object.
         predictions = results[0].predict(start=pred_start_date, end=pred_end_date)
         residuals = test_data['Passengers'] - predictions
In [66]: predictions
Out[66]: 1960-01-01
                       422.522544
         1960-02-01
                       404.089924
         1960-03-01
                       466.170683
         1960-04-01
                       455.600327
         1960-05-01
                       479.380371
         1960-06-01
                       532.876304
```

Freq: MS, Name: predicted_mean, dtype: float64

607.846748

619.393009

521.772779

466.984277

421.747837

463.563553

1960-07-01

1960-08-01

1960-09-01

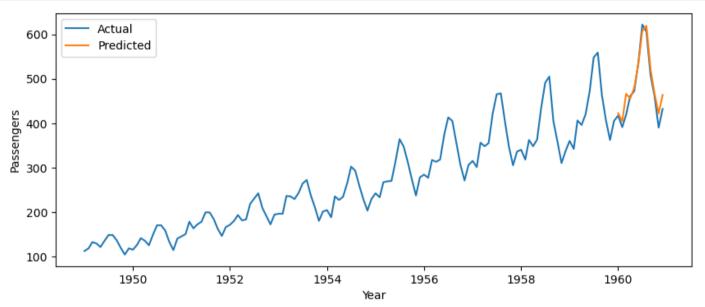
1960-10-01

1960-11-01

1960-12-01

```
In [67]: plt.figure(figsize=(10,4))
    plt.plot(df_airline['Passengers'], label="Actual")
    # plt.plot(test_data['Passengers'], label="Actual")
    plt.plot(predictions, label="Predicted")
    plt.xlabel('Year')
    plt.ylabel('Passengers')
    plt.legend()

plt.show()
```



```
In [68]: # References:
# https://online.stat.psu.edu/stat510/lesson/4/4.1
# https://www.statsmodels.org/stable/generated/statsmodels.tsa.statespace.sarimax.SARIMAX.html#statsmodels.tsa.statespace.sarimax.SARIMAX
```

Evaluating the model performance:

In [69]: from sklearn.metrics import mean_absolute_error, mean_absolute_percentage_error, mean_squared_error

```
In [70]: mean_absolute_error(test_data['Passengers'], predictions)

Out[70]: 15.941799825319455

In [71]: mean_absolute_percentage_error(test_data['Passengers'], predictions)

Out[71]: 0.0358363393098052

In [72]: np.sqrt(mean_squared_error(test_data['Passengers'], predictions))

Out[72]: 20.6648365631335
```

Conclusion:

ARIMA(0, 1, 1)(1, 1, 1)12 model is the best to capture the trend and seasonality present in the data.

Metric used to select best models are:

- AIC
- MAD(Mean Absolute Deviation)

The best parameters obtained for SARIMA while keeping the complexity of the model simple in concern are:

From the grid search method:

- Model: ARIMA(0, 1, 1)(1, 1, 1)12
- AIC: 802.86

From approach 1 using ACF and PACF of (1-L)^2 * Y_t:

- Model: ARIMA(2, 2, 2)(1, 0, 1)12
- AIC: 860.45

From approach 2 using ACF and PACF of (1-L^12) * Y_t:

- Model: ARIMA(2, 0, 5)(0, 1, 0)12
- AIC: 866.64