

6.2 (a) $P(\text{Dave fails the quiz}) = p = \frac{1}{4}$

Probability that Dave fails exactly two of the next six quizzes will be a Bernoulli process with the probability of one event (Dave failing the quiz), $p = \frac{1}{4}$

\therefore Probability that Dave fails exactly two of the next six quizzes is a binomial with parameters p and n . as $\frac{1}{4}$ & 6 respectively.

$$\begin{aligned} \therefore P_S(k) &= \binom{n}{k} p^k (1-p)^{n-k} = \binom{6}{2} p^2 (1-p)^{6-2} = \frac{6!}{4!2!} \left(\frac{1}{4}\right)^2 \left(\frac{3}{4}\right)^4 \\ &= \frac{1215}{4096} = 0.2966 \end{aligned}$$

(b) Expected number of quizzes = Expected number of quizzes failed + Expected number of quizzes passed. \dots equ)

Number of quizzes upto third failure is defined by Pascal random variable, Y_k of order three and $p = \frac{1}{4}$.

$$\therefore E[Y_k] = \frac{k}{p} = \frac{3}{\frac{1}{4}} = 12$$

\therefore From equ),

$$\begin{aligned} \text{Expected number of quizzes passed} &= 12 - 3 \\ &= 9 \text{ Ans.} \end{aligned}$$

(c) Given question says Dave fails exactly 1 quiz in first seven quizzes (A), quiz eight is a failure for Dave (B) and quiz nine is also a failure (C).

\therefore Desired probability = $P(A \cap B \cap C)$
 Since A, B and C are independent

$$\begin{aligned}
 &= P(A) \cdot P(B) \cdot P(C) \\
 &= \binom{7}{1} \left(\frac{1}{4}\right)^1 \left(\frac{3}{4}\right)^{7-1} \cdot \left(\frac{1}{4}\right) \cdot \left(\frac{1}{4}\right) \\
 &= \frac{7!}{6!} \left(\frac{1}{4}\right)^3 \left(\frac{3}{4}\right)^6 \\
 &= \frac{5103}{262144} = 0.0195
 \end{aligned}$$

(d) If A is the event that Dave fails two quizzes in a row before he passes two quizzes in a row.

P: Dave passes the quiz $P(P) = \frac{3}{4}$

F: Dave fails the quiz $P(F) = \frac{1}{4}$

$\therefore P(A) = P(\text{FF} \cup \text{PFF} \cup \text{FPFF} \cup \text{PFPFF} \cup \text{FPFPFF} \cup \text{PFPFPFF} \cup \dots)$

$$\begin{aligned}
 &= P(\text{FF}) + P(\text{PFF}) + P(\text{FPFF}) + P(\text{PFPFF}) + P(\text{FPFPFF}) + \\
 &\quad P(\text{PFPFPFF}) + \dots
 \end{aligned}$$

$$\begin{aligned}
 &= \left(\frac{1}{4}\right)^2 + \left(\frac{3}{4}\right)\left(\frac{1}{4}\right)^2 + \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4}\right)^2 + \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4}\right)^2 + \frac{1}{4} \cdot \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4}\right)^2 + \\
 &\quad \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4}\right)^2 + \dots \\
 &= \left[\left(\frac{1}{4}\right)^2 + \left(\frac{1}{4}\right) \left(\frac{3}{4}\right) \left(\frac{1}{4}\right)^2 + \left(\frac{1}{4}\right) \left(\frac{3}{4}\right) \left(\frac{1}{4}\right) \left(\frac{3}{4}\right) \left(\frac{1}{4}\right)^2 + \dots \right]
 \end{aligned}$$

$$+ \left[\frac{3}{4} \left(\frac{1}{4} \right)^2 + \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4} \right)^2 + \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \left(\frac{1}{4} \right)^2 + \dots \right]$$

Above are infinite Geometric series with common ratio $\frac{3}{16}$

$$\therefore P(A) = \frac{\left(\frac{1}{4} \right)^2}{1 - \frac{3}{16}} + \frac{\frac{3}{4} \left(\frac{1}{4} \right)^2}{1 - \frac{3}{16}} = \frac{7}{52} \text{ Ans.}$$

6.3 (a) Probability that the task is from user 1 in a particular slot $P_1 = P_{1/B} P_B$ [Here the trials are associated with busy slots] $= \frac{2}{5} \cdot \frac{5}{6} = \frac{1}{3}$
 $\therefore P(\text{task from user 1 is executed for the first time during the } 4^{\text{th}} \text{ slot}) = P_1 (1 - P_1)^3 = \frac{1}{3} \left(\frac{2}{3} \right)^3 = 0.0988$

(b) This is a Bernoulli process which has a property of fresh start. We can consider a fresh start from slot 11 and ignore the conditioning information of the first 10 slots

Thus,

$$P(\text{slot 11 is busy and slot 12 is idle}) = \frac{5}{6} \cdot \frac{1}{6} = 0.139$$

(c) $P_1 = \frac{1}{3}$

Time of 5th slot task from user 1 is a Pascal random variable, Y_k with $k=5$ and $p_1 = \frac{1}{3}$.

$$\therefore E[Y_k] = \frac{k}{P_1} = \frac{5}{1/3} = 15$$

(d) In this case, we consider only busy slots upto and including 5th task from user 1. Thus, here kth arrival time will be defined by the Pascal's PMF Y_k with $P_{1|B} = \frac{2}{5}$

$$\therefore E[Y_5] = \frac{k}{P} = \frac{5}{2/5} = \frac{25}{2} = 12.5$$

(e) If we subtract the number of busy slots occupied by user 1 (i.e 5) ^{until the 5th task from user 1} from the number of busy slots until the 5th task from user 1, we get the number of busy slots occupied by user 2 until the 5th task from user 1. This will be the number of tasks from user 2 until the 5th task from user 1 (A). we can write this as a Pascal Random variable, with $k=5$ and $P_{1|B} = 2/5$.

Thus, PMF is given by,

$$P_{Y_k}(t) = \binom{t-1}{k-1} P_{1|B}^k (1-P_{1|B})^{t-k} \quad t = k, k+1, \dots$$

$$\text{Putting } k=5 \Rightarrow P_B(t) = \binom{t-1}{4} \left(\frac{2}{5}\right)^5 \left(1-\frac{2}{5}\right)^{t-5}, \quad t = 5, 6, \dots$$

$$A = B - 5 \quad \therefore P_A(t) = P_B(t+5)$$

$$\therefore P_A(t) = \binom{t+4}{4} \left(\frac{2}{5}\right)^5 \left(1-\frac{2}{5}\right)^t, \quad t = 0, 1, \dots$$

Using the formulas for mean and variance of the Pascal random variable B, we have

$$E[A] = E[B] - 5 = \frac{25}{2} - 5 = 7.5$$

$$\text{var}[A] = \text{var}[B] = \frac{5(1-2/5)}{(2/5)^2} = \frac{3}{0.16} = 18.75$$

6.10 (a) This is a Poisson process with 0 arrivals in 2 hours
 $\lambda = 0.6$ per hour.

$$\therefore P_{N_T}(k) = P(k, \tau) = e^{-\lambda\tau} \frac{(\lambda\tau)^k}{k!}$$

$$\therefore k=0, \tau=2$$

$$= e^{-0.6 \times 2} \frac{(0.6 \times 2)^0}{0!}$$

$$= e^{-0.6 \times 2} = 0.301$$

(b) For this to happen, he should not catch fish in first 2 hours (0 successes in first two hours) and at least one fish between 2 and 5 hours. These are disjoint events and the probability that the total time he spends fishing is between two and five hours is given by,

$$P(0, 2) (1 - P(0, 3)) = e^{-0.6 \times 2} (1 - e^{-0.6 \times 3}) = 0.251$$

(c) The only case when the fisherman catches at least 2 fish is when he fishes for exactly 2 hours. Thus, the probability that he catches at least two fish is same as the probability that the number of fish caught in first two hours is at least two, i.e.,

$$\sum_{k=2}^{\infty} P(k, 2) = 1 - P(0, 2) - P(1, 2) = 1 - e^{-0.6 \times 2} - (0.6 \times 2) e^{-0.6 \times 2} = 0.337$$

(d) Expected number of fish caught = Expected number of fish caught in first two hours + Expected number of fish caught after first two hours if 0 fish is caught in first two hours.

$$\text{Expected no. of fish caught in first two hours} = \lambda T = 2\lambda = 1.2$$

There are two possibilities after two hours -

1. he quits in which case 0 fish is caught after 2 hours.
2. he catches 1 fish in which case 0 fish is caught in first two hours.

$$\therefore E[\text{event 2}] = P[0 \text{ fish caught in 1st two hours}] = P(0, 2) = 0.301$$

$$\therefore \text{Ans. } 1.2 + 0.301 = 1.501$$

(e) We know, Poisson process is memoryless. Thus, expected time spent to catch 1 fish after 4 hours, can be found by using the exponential random variable with parameter λ .

$$\text{Thus, } E[T] = \frac{1}{\lambda} = \frac{1}{0.6} = 1.667$$

$$\therefore \text{Ans. } 4 + 1.667 = 5.667$$

6.14(a) T : time until the first bulb failure
 A : First bulb is of type A $P(A) = \frac{1}{2}$
 B : First bulb is of type B $P(B) = \frac{1}{2}$

using total expectation theorem,

$$\begin{aligned} E[T] &= E[T|A]P(A) + E[T|B]P(B) = \frac{1}{\lambda_A} \cdot \frac{1}{2} + \frac{1}{\lambda_B} \cdot \frac{1}{2} \\ &= 1 \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2} = \frac{2}{3} \text{ Ans.} \end{aligned}$$

- (b) If C is the event of no bulb failure before time t .
We can use total probability theorem as,

$$P(C) = P(C|A)P(A) + P(C|B)P(B)$$

$$P(C|A) = e^{-t}$$

$$P(C|B) = e^{-3t}$$

$$\therefore \text{Ans. } \frac{1}{2}e^{-t} + \frac{1}{2}e^{-3t}$$

- (c) Using Bayes' Rule,

$$P(A|C) = \frac{P(C|A)P(A)}{P(C)}$$

$$= \frac{\frac{1}{2}e^{-t}}{\frac{1}{2}e^{-t} + \frac{1}{2}e^{-3t}} = \frac{1}{1+e^{-2t}}$$

- (d) $E[T^2] = E[T^2] + \text{var}[T]$. Because ^{this is} second moment of an exponential random variable with parameter λ .

$$= \frac{1}{\lambda^2} + \frac{1}{\lambda^2} = \frac{2}{\lambda^2}$$

Conditional form of total expectation theorem,

$$E[T^2] = E[T^2|A]P(A) + E[T^2|B]P(B) = 2 \cdot \frac{1}{2} + \frac{2}{9} \cdot \frac{1}{2} = \frac{10}{9}$$

$$\therefore \text{var}(T) = E[T^2] - (E[T])^2$$

$$= \frac{10}{9} - \left(\frac{2}{3}\right)^2 = \frac{2}{9} \text{ Ans.}$$

- (e) This can be understood as the event that out of first 11 bulbs, exactly 3 were of type A as well as 12th bulb was of type A.

$$\text{Thus, } P = \binom{11}{3} \left(\frac{1}{2}\right)^{11} \cdot \left(\frac{1}{2}\right) \\ = \binom{11}{3} \left(\frac{1}{2}\right)^{12}$$

(f) Probability that exactly 4 type-A bulbs have failed upto and including the 12th bulb failure = probability that out of first 12 bulbs, exactly 4 were type A = $\binom{12}{4} \left(\frac{1}{2}\right)^{12}$

(g) PDF of the time between failures is $\frac{e^{-x} + 3e^{-3x}}{2}, x \geq 0$
Associated transform is given by,

$$\frac{1}{2} \left(\frac{1}{1-s} + \frac{3}{3-s} \right)$$

The times between successive failures are independent
 \therefore the transform associated with the time until the 12th failure is given by

$$\left[\frac{1}{2} \left(\frac{1}{1-s} + \frac{3}{3-s} \right) \right]^{12}$$

(h) Y: total period of illumination provided by the first two type-B bulbs.

This has an Erlang distribution of order 2,

PDF:

$$f_Y(y) = 9ye^{-3y}, y \geq 0$$

If T is the period of illumination provided by first type-A bulb. Its PDF is

$$f_T(t) = e^{-t}, t \geq 0$$

For $T < 4$

$$P(T < 4 | Y = y) = 1 - e^{-y} \quad y \geq 0$$

$$\begin{aligned} \therefore P(T < 4) &= \int_0^{\infty} f_Y(y) P(T < 4 | Y = y) dy \\ &= \int_0^{\infty} 9ye^{-3y} (1 - e^{-y}) dy = \frac{7}{16} \end{aligned}$$

Let T_1^A be period of illumination of first Type A bulb

Let T_1^B & T_2^B be period of illumination of 1st and 2nd type B bulb respectively.

$$\text{for } \{ T_1^A < T_1^B + T_2^B \}$$

$$\begin{aligned} P(T_1^A < T_1^B + T_2^B) &= P(T_1^A < T_1^B) + P(T_1^A \geq T_1^B) P(T_1^A < T_1^B + T_2^B | T_1^A \geq T_1^B) \\ &= \frac{1}{1+3} + P(T_1^A \geq T_1^B) P(T_1^A - T_1^B < T_2^B | T_1^A \geq T_1^B) \\ &= \frac{1}{4} + \frac{3}{4} P(T_1^A - T_1^B < T_2^B | T_1^A \geq T_1^B) \end{aligned}$$

Given event, $T_1^A \geq T_1^B$ and using memorylessness remaining time $T_1^A - T_1^B$ is exponentially distributed

$$\begin{aligned} P(T_1^A - T_1^B < T_2^B | T_1^A \geq T_1^B) &= P(T_1^A < T_2^B) \\ &= P(T_1^A < T_2^B) = \frac{1}{4} \end{aligned}$$

$$\text{Ans. } \frac{1}{4} + \frac{3}{4} \cdot \frac{1}{4} = \frac{7}{16}$$

(i) V : total period of illumination provided by type B bulbs

N : No. of light bulbs

x_i : period of illumination from the i th type B bulb

$$V = Y_1 + \dots + Y_N$$

N - binomial Random variable with $n=12, p=\frac{1}{2}$

$$E[N] = 6, \text{Var}(N) = 12 \cdot \frac{1}{2} \cdot \frac{1}{2} = 3$$

$$E[x_i] = \frac{1}{3}, \text{Var}(x_i) = \frac{1}{9}$$

$$E[V] = E[N] \cdot E[x_i] = 2$$

$$\text{Var}(V) = \text{Var}(x_i) \cdot E[N] + E[x_i]^2 \text{Var}(N)$$

$$= \frac{1}{9} \cdot 6 + \frac{1}{9} \cdot 3 = 1$$

$$(j) E[T|D] = t + E[T-t | D \cap A] P[A|D] + E[T-t | D \cap B] P[B|D]$$

$$= t + 1 \cdot \frac{1}{1+e^{-2t}} + \frac{1}{3} \left(1 - \frac{1}{1+e^{-2t}} \right)$$

$$= t + \frac{1}{3} + \frac{2}{3} \cdot \frac{1}{1+e^{-2t}}$$

Other Problems

1.

$$\lambda = 0.4$$

$$\{N(t), t \in [0, \infty)\}$$

(a) $P(\text{No. of arrival in } (2, 4])$

$$= P(N(4) - N(2) = 0)$$

$$= P(N(2) = 0) = \frac{e^{-0.4 \times 2} \times (0.4 \times 2)^0}{0!} = e^{-0.8} = 0.449$$

(b) $P(1 \text{ arrival in } (0,1], (3,5])$

$$P(N(1) - N(0) = 1) \cdot P(N(5) - N(3) = 1)$$

$$= \frac{e^{-0.4} \times (0.4 \times 1)^1}{1!} \times \frac{e^{-0.8} \times 0.8}{1}$$

$$= 0.26812 \times 0.35948 = 0.09638$$

(c) $P(1 \text{ arrival in } (0,1] \text{ \& \; } 3 \text{ arrival in } (0,5])$

$$(0,1] \cap (0,5] = (0,1]$$

Let X, Y be no. of arrival in $(0,1]$ & $(1,5]$ respectively

$$X \sim \text{Poisson}(\lambda, 1)$$

$$Y \sim \text{Poisson}(\lambda, 4)$$

$$P(A) = P(X=1 \text{ \& \; } X+Y=3)$$

$$= P(X=1) \text{ \& \; } P(Y=2)$$

$$= P(N(1)=1) \cdot P(N(4)=2)$$

$$= \frac{e^{-0.4 \times 1} \times 0.4}{1!} \times \frac{e^{-0.4 \times 4} \times (0.4 \times 4)^2}{2!}$$

$$= 0.06925$$

$$2. \quad N(t) = N_1(t) + N_2(t)$$

$$N(t) \sim \text{Poisson}(\lambda_1 t + \lambda_2 t)$$

$$N(t) \sim \text{Poisson}(3t)$$

$$\therefore \lambda = 3$$

$P(2 \text{ arrivals in } [0,1] \text{ and } 3 \text{ arrivals in } (1,2])$

$$= \frac{(3 \times 1)^2 e^{-3 \times 1}}{2!} \times \frac{(3 \times 1)^3 e^{-3 \times 1}}{3!}$$

$$= \frac{9}{2} e^{-3} \times \frac{9}{2} e^{-3} = \frac{81}{4} e^{-6} = 0.0501 \text{ Ans.}$$

$$(b) \quad N(1) = 2$$

$$P(N_1(1) = 1 \mid N(1) = 2)$$

$$= \frac{P(N_1(1) = 1, N(1) = 2)}{P(N(1) = 2)}$$

(By conditional probability)

$$= \frac{P(N_1(1) = 1, N_2(1) = 1)}{P(N(1) = 2)}$$

(if $N_1(1) = 1$, & $N(1) = 2$
then $N_2(1) = 1$)

$$= \frac{P(N_1(1) = 1) \times P(N_2(1) = 1)}{P(N(1) = 2)}$$

N_1 & N_2 are
independent

$$= \frac{\frac{(1 \times 1)^1 e^{-1}}{1!} \times \frac{(2 \times 1)^1 e^{-2}}{1!}}{\frac{(3 \times 1)^2 \times e^{-3}}{2!}} = \frac{4}{9} \text{ Ans.}$$