

Problems from the book:

Section 4.1: Problem 2

$$Y = e^X$$

$$\text{CDF of } Y: F_Y(y) = P(Y \leq y) = P(e^X \leq y) = \begin{cases} P(X \leq \ln y), & \text{if } y > 0 \\ 0 & , \text{ otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{d}{dy} F_X(\ln y), & \text{if } y > 0, \\ 0 & , \text{ otherwise} \end{cases} = \begin{cases} \frac{1}{y} f_X(\ln y) & , \text{if } y > 0 \\ 0 & , \text{ otherwise} \end{cases}$$

Ans.

Note:  $Y = e^X$  is a strictly monotonic function

$$\therefore f_Y(y) = f_X(h(y)) \left| \frac{dh(y)}{dy} \right|$$

$$x = h(y)$$

$$h(y) = \ln y \text{ (in this case)}$$

$$\text{If } X \sim U[0, 1]: f_X(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & , \text{ otherwise} \end{cases}$$

$$\therefore f_X(\ln y) = \begin{cases} 1 & , 0 \leq \ln y \leq 1 \\ 0 & , \text{ otherwise} \end{cases} = \begin{cases} 1 & , 1 \leq y \leq e \\ 0 & , \text{ otherwise} \end{cases}$$

$$\therefore f_Y(y) = \begin{cases} \frac{1}{y} & , 1 \leq y \leq e \\ 0 & , \text{ otherwise} \end{cases}$$

Ans.

Problem 5:  $X \sim U[0,1]$ ,  $Y \sim U[0,1]$

Let  $z = |X - Y|$

$$\text{CDF, } F_Z(z) = P(Z \leq z) = P(|X - Y| \leq z)$$

$$= P(-z \leq X - Y \leq z)$$

$$= P(-X - z \leq -Y \leq z - X)$$

$$= P(X - z \leq Y \leq X + z)$$

We can see that  $Y$  lies between the values from lines  $Y = X - z$  &  $Y = X + z$ . Since,  $X$  &  $Y \in [0,1]$  the CDF can be computed by considering the following figure:

For a particular value of  $z$ , the area of the shaded portion would give the CDF in the range  $0 \leq z \leq 1$

$$\text{Area} = 1 - \left( 2 \times \frac{1}{2} \times (1-z) \times (1-z) \right)$$

By symmetry

$$= 1 - (1-z)^2$$

which is basically the area between the lines  $Y = X + z$  &  $Y = X - z$ .

$$\therefore F_Z(z) = \begin{cases} 0 & z \leq 0 \\ 1 - (1-z)^2, & 0 \leq z \leq 1 \\ 1 & z > 1 \end{cases}$$

Ans.

$$f_Z(z) = \frac{d}{dz} F_Z(z) = \begin{cases} 2(1-z) & 0 \leq z \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Ans.

Problem 7: Let  $X$  and  $Y$  be the points which are chosen randomly in the interval  $[0, 1]$

Given:  $X \sim U[0, 1]$ ,  $Y \sim U[0, 1]$

Distance between the points is given by  $|X - Y|$

From ques. 5 we have CDF of  $z = |X - Y|$

$$\begin{aligned} \therefore E[z] &= \int_0^1 z f_z(z) dz \\ &= \int_0^1 z (2(1-z)) dz \\ &= 2 \left[ \left[ \frac{z^2}{2} \right]_0^1 - \left[ \frac{z^3}{3} \right]_0^1 \right] = 2 \left( \frac{1}{2} - \frac{1}{3} \right) = \frac{1}{3} \end{aligned}$$

Hence proved.

Problem 13: Given:  $f_x(x) = f_y(y) \geq 0 \quad \forall x, y \in [a, b]$

$$f_x\left(x + \frac{a+b}{2}\right) = f_x\left(-x + \frac{a+b}{2}\right)$$

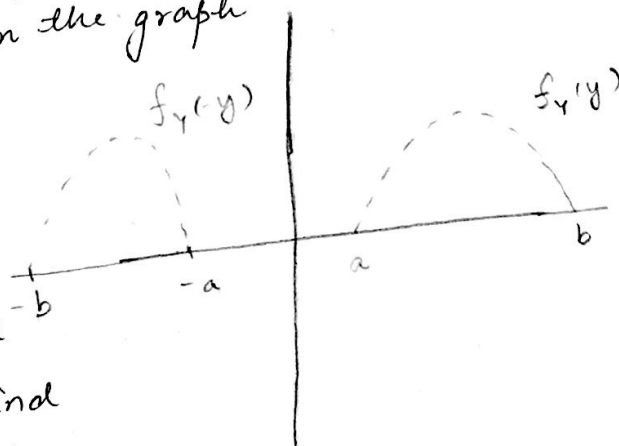
$$f_y\left(y + \frac{a+b}{2}\right) = f_y\left(-y + \frac{a+b}{2}\right)$$

Since, PDFs are symmetric about the mean,  $\frac{a+b}{2}$

$$\text{Also, } f_y(y) = f_y(-y + a+b)$$

which is also clear from the graph

$\therefore$  PDF of  $X+Y$  is same as  $X-Y + (a+b)$



Thus, if we have calculated the PDF of  $X+Y$  we can find the PDF of  $X-Y$  easily by

shifting the PDF of  $X+Y$  to the left by a quantity equal to  $a+b$ .

## Section 4.2 : Problem 18

We know,

$$\rho_{X,Y} = \frac{\text{cov}(X,Y)}{\sqrt{\text{var}(X) \text{var}(Y)}}$$

$$\text{cov}(X_1 + X_2, Y_1 + Y_2) = \text{cov}(X_1, Y_1) + \text{cov}(X_1, Y_2) + \text{cov}(X_2, Y_1) + \text{cov}(X_2, Y_2)$$

$$\text{cov}(X, X) = \text{var}(X)$$

$$\begin{aligned} \therefore \rho_{(W+X), (X+Y)} &= \frac{\text{cov}(W+X, X+Y)}{\sqrt{\text{var}(W+X) \text{var}(X+Y)}} \\ &= \frac{\text{cov}(W, X) + \text{cov}(W, Y) + \text{cov}(X, X) + \text{cov}(X, Y)}{\sqrt{(\text{var}(W) + \text{var}(X) + 2\text{cov}(W, X))(\text{var}(X) + \text{var}(Y) + 2\text{cov}(X, Y))}} \\ &= \frac{\text{var}(X)}{\sqrt{1+1+0} \sqrt{1+1+0}} = \frac{1}{2} \text{ Ans.} \end{aligned}$$

... Since  $W, X, Y, Z$  are pairwise uncorrelated  
thus  $\text{cov}(W, X) = \text{cov}(W, Y) = \dots \text{etc.} = 0$ .

Similarly,

$$\rho_{(R,T)} = \frac{\text{cov}(W+X, Y+Z)}{\sqrt{\text{var}(W+X) \text{var}(Y+Z)}} = 0 \text{ Ans.}$$

## Section 4.3 : Problem 24

(a) Let  $X$  be the random variable representing the amount of time that the professor devotes to the task and  $Y$  be the random variable that represents the length of the time interval between 9 a.m. and the time of his arrival. Given:  $X$  is a function of  $Y$ . It is an exponential random variable with  $\lambda(y) = \frac{1}{5-y}$

$$Y \sim U[0, 4]$$

$\therefore$  Expected amount of time that the professor devotes to the task =  $E[X] = E[E[X|Y]]$  ... by law of iterated expectations

We know, that the expected value of an exponential random variable is  $\frac{1}{\lambda}$ .

$$\therefore E[X|Y=y] = \frac{1}{1/5-y} = 5-y$$

$$\Rightarrow E[X|Y] = 5-Y$$

$$\therefore E[X] = E[E[X|Y]] = E[5-Y] = 5 - E[Y] = 5 - 2 = 3$$

$\because Y \sim U[0,4]$

Ans. 3 hours

(b) Time at which the task is completed will be defined by the random variable  $X+Y$

$$\therefore E[X+Y] = E[X] + E[Y] = 3 + 2 = 5 \text{ hours}$$

Expected time at which the task is completed is 2:00 p.m. i.e. 5 hours after 9 a.m.

(c) Let  $Z$  be the random variable defining the length of time interval between 9 a.m. and the time of arrival of the Ph.D. student,  $A$  be the random variable defining the amount of time the student will spend with the professor, if he meets the professor and  $B$  be the random variable defining the amount of time the professor will spend with the student.

Let's assume that the probability of the student meeting the professor be  $p$  and the event be  $C$ .

$\therefore$  By Total Expectation theorem,

$$E[B] = p E[B|C] + (1-p) E[B|C^c] \quad \dots \text{eqn}$$

Clearly,

$E[B|c] = E[A] = \frac{1}{2}$  [since  $A \sim U[0,1]$ ]  
 $A$  &  $B$  are equivalent events when the student meets the professor.

$E[B|c^c] = 0$  [since time spent with professor will be 0 if he doesn't find the professor ( $c^c$ ) & leaves]

$\therefore$  From eqn, we have

$$E[B] = \frac{P}{2}$$

If student needs to meet the professor, he should arrive after the professor has arrived and before the professor leaves.

$$\begin{aligned} \therefore P &= P(Y \leq Z \leq X+Y) \\ Z &\sim U[0,8] \\ X+Y &\geq 0 \end{aligned} \quad f_Z(z) = \begin{cases} \frac{1}{8} & 0 \leq z \leq 8 \\ 0 & \text{otherwise} \end{cases}$$

$$f_{Y,Y}(y) = \begin{cases} \frac{1}{4} & 0 \leq y \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

$$\therefore P = 1 - [P(Z < Y) + P(Z > X+Y)] \quad \dots \text{eqn (i)}$$

$$P(Z < Y) = \int_0^4 \int_0^y f_{Y,Z}(y,z) dy dz$$

$$= \int_0^4 \frac{1}{4} \int_0^y \frac{1}{8} dz dy = \frac{1}{8 \times 4} \left[ \frac{y^2}{2} \right]_0^4 = \frac{16}{2 \times 8 \times 4} = \frac{1}{4}$$

$$P(Z > X+Y) = \int_0^4 P(Z > X+Y | Y=y) f_Y(y) dy$$

$$= \int_0^4 P(X < Z-Y | Y=y) f_Y(y) dy$$

$$= \int_0^4 \int_y^8 F_{X|Y}(z-y) f_Z(z) f_Y(y) dz dy$$

$$\begin{aligned}
&= \int_0^4 \frac{1}{4} \int_y^8 \frac{1}{8} \int_0^{z-y} \frac{1}{5-y} e^{-\frac{x}{5-y}} dx dz dy \quad \left[ \because F_x(x) = \int_0^x f_x(x) dx \right] \\
&= \frac{1}{32} \int_0^4 \int_y^8 \frac{5-y}{5-y} \left[ e^{-\frac{x}{5-y}} \right]_0^{z-y} dz dy \\
&= \frac{1}{32} \int_0^4 \int_y^8 1 - e^{-\frac{z-y}{5-y}} dz dy \\
&= \frac{1}{32} \int_0^4 \left[ z + (5-y) e^{-\frac{z-y}{5-y}} \right]_y^8 dy \\
&= \frac{1}{32} \int_0^4 \left( (8-y) + (5-y) \left[ e^{-\frac{8-y}{5-y}} - 1 \right] \right) dy \\
&= \frac{12}{32} + \frac{1}{32} \int_0^4 (5-y) e^{-\frac{(8-y)}{5-y}} dy
\end{aligned}$$

Integrating numerically, we get

$$\int_0^4 (5-y) e^{-\frac{(8-y)}{5-y}} dy = 1.7584$$

$\therefore$  From eq(ii), we have

$$\begin{aligned}
p &= 1 - \frac{1}{4} - \frac{12}{32} - 0.05495 \\
&= 0.32
\end{aligned}$$

$$\therefore E[B] = \frac{p}{2} = 0.16 \text{ hours or } 9.6 \text{ mins}$$

If  $Z$  is the random variable defining the length of the time interval measured from 9 a.m. until he leaves his office.

We can write,

$$\begin{aligned}\therefore E[z] &= p E[z|c] + (1-p) E[z|c^c] \\ &= p E[X+Y+A] + (1-p) E[X+Y] \\ &= 0.32 \left(5 + \frac{1}{2}\right) + 0.68 \times 5 \\ &= 5.16 \text{ hours.}\end{aligned}$$

Ans. Expected amount of time that the professor will spend with the student is 9.6 mins and the expected time at which he will leave his office is 5.16 hours after 9 a.m.

#### Section 4.4: Problem 30

$$X \sim N(0, 1)$$

We know, for  $X \sim N(\mu, \sigma^2)$

$$M_X(s) = e^{(\sigma^2 s^2 / 2) + \mu s}$$

$$\begin{aligned}M_X(s) &= E[e^{sX}] = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-(x^2/2) + sx} dx \\ &= \frac{e^{s^2/2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(x^2/2) + sx - \frac{s^2}{2}} dx \\ &= \frac{e^{s^2/2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(x-s)^2/2} dx = e^{s^2/2}\end{aligned}$$

We know,  $\left. \frac{d^n}{ds^n} M_X(s) \right|_{s=0} = E[X^n]$



$$\begin{aligned}
 \therefore E[X^3] &= \left. \frac{d^3}{ds^3} M_X(s) \right|_{s=0} \\
 &= \left. \frac{d^3}{ds^3} e^{s^2/2} \right|_{s=0} = \\
 \frac{d}{ds} e^{s^2/2} &= \frac{2s}{2} e^{s^2/2} = s e^{s^2/2} \\
 \frac{d^2}{ds^2} e^{s^2/2} &= e^{s^2/2} + s \cdot s e^{s^2/2} \\
 \frac{d^3}{ds^3} e^{s^2/2} &= s e^{s^2/2} + 2s e^{s^2/2} + s^2 \cdot s e^{s^2/2} \quad \dots \text{eq(ii)} \\
 \frac{d^4}{ds^4} e^{s^2/2} &= 3(e^{s^2/2} + s^2 e^{s^2/2}) + 3s^2 e^{s^2/2} + s^3 \cdot s e^{s^2/2} \quad \dots \text{eq(iii)}
 \end{aligned}$$

$$\begin{aligned}
 \text{Ans. } \therefore E[X^3] &= 0 \quad \dots \text{from eq(ii)} \\
 E[X^4] &= 3 \quad \dots \text{from eq(iii)}
 \end{aligned}$$

Problem 34:  $X_i$  is a Bernoulli random variable that takes value 1 if the  $i^{\text{th}}$  player is successful.  $i = 1, 2, 3$

Considering convolutions of PMF of  $X_1$  &  $X_2$ . ( $Z = X_1 + X_2$ )

$$p_z(z) = \begin{cases} (1-p_1)(1-p_2) & , z=0 \\ (1-p_1)p_2 + p_1(1-p_2) & , z=1 \\ p_1 p_2 & , z=2 \\ 0 & , \text{otherwise} \end{cases}$$

Convolution of the PMF of  $Z$  and  $X_3$ ,

Ans.

$$P_X(x) = \begin{cases} (1-p_1)(1-p_2)(1-p_3) & x=0 \\ p_1(1-p_2)(1-p_3) + (1-p_1)p_2(1-p_3) + (1-p_1)(1-p_2)p_3 & x=1 \\ (1-p_1)p_2p_3 + p_1(1-p_2)p_3 + p_1p_2(1-p_3) & x=2 \\ p_1p_2p_3 & x=3 \\ 0 & \text{otherwise} \end{cases}$$

$$M_X(s) = (1-p_1 + p_1e^s)(1-p_2 + p_2e^s)(1-p_3 + p_3e^s)$$

In the above equation, the coefficients of the terms  $e^{ks}$  gives the probability that  $x$  takes the value  $k$ .  
 Ex.  $e^{s(0)}$  has the coefficient  $(1-p_1)(1-p_2)(1-p_3)$  which is same which we got using convolution.  
 Hence verified.

Extra problems:

1.  $X \sim U[0,1]$

$$1.1 \quad F_X(x) = \begin{cases} x & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} Y &= X^2 \\ F_Y(y) &= P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) \\ &= F_X(\sqrt{y}) - F_X(-\sqrt{y}) \\ &= \begin{cases} \sqrt{y} - 0 & 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

$$\begin{aligned} \therefore f_Y(y) &= \frac{d}{dy} F_Y(y) \\ &= \begin{cases} \frac{1}{2\sqrt{y}} & 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases} \quad \text{Ans.} \end{aligned}$$

$$1.2 \quad f(x, y) = \frac{\text{cov}(x, y)}{\sqrt{\text{var}(x) \cdot \text{var}(x^2)}} \\ = \frac{E[x \cdot x^2] - E[x] \cdot E[x^2]}{\sqrt{(E[x^2] - E[x]^2)(E[x^4] - E[x^2]^2)}}$$

$$E[x] = \frac{1}{2} \int_0^1 x^2 \cdot 1 dx = \left[ \frac{x^3}{3} \right]_0^1 = \frac{1}{3} \\ E[x^2] = \int_0^1 x^3 \cdot 1 dx = \left[ \frac{x^4}{4} \right]_0^1 = \frac{1}{4} \\ E[x^3] = \int_0^1 x^4 \cdot 1 dx = \left[ \frac{x^5}{5} \right]_0^1 = \frac{1}{5}$$

$$\therefore f(x, y) = \frac{\frac{1}{4} - \frac{1}{2} \cdot \frac{1}{3}}{\sqrt{\left(\frac{1}{3} - \frac{1}{4}\right)\left(\frac{1}{5} - \frac{1}{9}\right)}} \\ = \frac{\frac{1}{12}}{\sqrt{\frac{1}{12} \cdot \frac{4}{45}}} = \frac{\sqrt{15}}{4} \quad \text{Ans.}$$

2. Using convolution,

$$f_z(z) = \int_{-\infty}^{\infty} f_x(x) f_y(z-x) dx$$

\*

$$y = z - x$$

$$f_y(y) = \begin{cases} \frac{1 - (z-x)}{2} \\ 0 \end{cases}$$

$$x-1 \leq z \leq 1+x$$

\* is non-zero if  $-1 \leq x \leq 1$  &  $x-1 \leq z \leq 1+x$ , otherwise

$$\Rightarrow \underbrace{-2 \leq x-1 \leq z}_{R_1} \leq \underbrace{x+1 \leq 2}_{R_2}$$

Case 1 :  $x \in R_1$ ,  $-1 \leq x \leq z+1$

$$\therefore f_z(z) = \int_{-1}^{z+1} \left( \frac{1+x}{2} \right) \left( \frac{1-z+x}{2} \right) dx \quad ; \quad z < 0$$

$$-2 \leq z \leq 0$$

$$= \frac{1}{4} \int_{-1}^{z+1} (1-z+2x-xz+x^2) dx$$

$$= \frac{1}{4} \int_{-1}^{1+z} [(1-z) + (2-z)x + x^2] dx$$

$$= \frac{1}{4} \left[ (1-z)(1+z-(-1)) + \left[ \frac{(2-z)x^2}{2} \right]_{-1}^{z+1} \right.$$

$$\left. + \frac{1}{3} (1+z)^3 - (-1)^3 \right]$$

$$= \frac{1}{4} \left[ (1-z)(2+z) + \frac{(2-z)}{2} (1+z^2+2z-1) + \frac{1}{3} (1+z^3+3z^2+3z+1) - (-1) \right]$$

$$= \frac{1}{4 \times 6} (-z^3 + 12z + 16) \quad ; \quad -2 \leq z < 0$$

Case 2 :  $x \in R_2$ ,  $0 \leq z \leq 2 \quad \Leftarrow \quad z-1 \leq x \leq 1$

$$\therefore f_z(z) = \frac{1}{4} \int_{z-1}^1 (1-z+x)(1+x) dx$$

$$= \frac{1}{4} \left[ \left( \frac{x+1}{2} \right)^3 - \frac{z}{2} (x+1)^2 \right]_{z-1}^1$$

$$= \frac{1}{4} \left[ \frac{8}{3} - 2z - \frac{z^3}{3} + \frac{z^3}{2} \right]$$

$$= \frac{z^3}{24} - \frac{z}{2} + \frac{2}{3} = \frac{1}{24} (z^3 - 12z + 16)$$

$$\text{Ans. } f_z(z) = \begin{cases} \frac{1}{24}(-z^3 + 12z + 16) & -2 \leq z \leq 0 \\ \frac{1}{24}(z^3 - 12z + 16) & 0 \leq z \leq 2 \\ 0 & \text{otherwise} \end{cases}$$