

SM5000 : Foundation of Machine Learning  
Assignment – 1

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## Solution 1 : Linear Regression [Theory]

It is the multivariate regression case where  $(x_0, x_1, \dots, x_n)$  are independent variables and  $(y_0, y_1, \dots, y_n)$  are dependent variables.

Dependent variables are depending on some set of independent variables so independent noise is added to each dimension of the independent variable set.

Our Model represented as:

$$y(x, \omega) = \omega_o + \sum_{i=1}^D \omega_i x_i$$

if error(noise) is added to each observed x value, our function of y will change as given

$$y'(x, \omega) = \omega_o + \sum_{i=1}^D \omega_i (x_i + \varepsilon_i)$$

$$y'(x, \omega) = \omega_o + \sum_{i=1}^D \omega_i x_i + x_i \varepsilon_i$$

$$y'(x, \omega) = y(x, \omega) + \sum_{i=1}^D \omega_i \varepsilon_i$$

Our New error function will be represented as:

$$E'_D(x, \omega) = \frac{1}{2} \sum_{n=1}^N \{ y'(x_n, \omega) - t_n \}^2$$

$$E'_D(x, \omega) = \frac{1}{2} \sum_{n=1}^N \{ y'(x_n, \omega) - t_n \}^2$$

$$E'_D(x, \omega) = \frac{1}{2} \sum_{n=1}^N \left\{ y(x_n, \omega) + \sum_{i=1}^D \omega_i \varepsilon_{ni} - t_n \right\}^2 \quad \dots \dots \dots 1$$

Expanding equation 1:

$$E'_D(x, \omega) = \frac{1}{2} \sum_{n=1}^N \left\{ (y(x_n, \omega) - t_n)^2 + \left( \sum_{i=1}^D \omega_i \varepsilon_{ni} \right)^2 + 2(y(x_n, \omega) - t_n) \left( \sum_{i=1}^D \omega_i \varepsilon_{ni} \right) \right\} \dots 2$$

Taking expectation of equation 2:

$$E\{E'_D(x, \omega)\} = \frac{1}{2} E \left\{ \sum_{n=1}^N \left\{ (y(x_n, \omega) - t_n)^2 + \left( \sum_{i=1}^D \omega_i \varepsilon_{ni} \right)^2 + 2(y(x_n, \omega) - t_n) \left( \sum_{i=1}^D \omega_i \varepsilon_{ni} \right) \right\} \right\}$$

$$E\{E'_D(x, \omega)\} = \frac{1}{2} \sum_{n=1}^N \left\{ \underbrace{E\{(y(x_n, \omega) - t_n)^2\}}_{\text{I}} + \underbrace{E\left\{\left(\sum_{i=1}^D \omega_i \varepsilon_{ni}\right)^2\right\}}_{\text{II}} + 2(y(x_n, \omega) - t_n) \underbrace{\left\{E\left(\sum_{i=1}^D \omega_i \varepsilon_{ni}\right)\right\}}_{\text{III}} \right\}$$

I.  $E\{(y(x_n, \omega) - t_n)^2\} = (y(x_n, \omega) - t_n)^2$  **as**  $(y(x_n, \omega) - t_n)^2$  is free of noise

II.  $2(y(x_n, \omega) - t_n)\{E(\sum_{i=1}^D \omega_i \varepsilon_{ni})\} = 2(y(x_n, \omega) - t_n)[(\sum_{i=1}^D \omega_i E\{\varepsilon_{ni}\})] = 0$

**as**  $E\{\varepsilon_{ni}\} = 0$  is given to us.

III. Solving  $E\{(\sum_{i=1}^D \omega_i \varepsilon_{ni})^2\}$  below:

$$\begin{aligned} E \left\{ \left( \sum_{i=1}^D \omega_i \varepsilon_{ni} \right)^2 \right\} &= E \left\{ \sum_{i=1}^D \sum_{j=1}^D \omega_i \varepsilon_{ni} \omega_j \varepsilon_{nj} \right\} \\ &= E \left\{ \sum_{i=1}^D \sum_{j=1}^D \omega_i \omega_j \varepsilon_{ni} \varepsilon_{nj} \right\} \\ &= \sum_{i=1}^D \sum_{j=1}^D \omega_i \omega_j E\{\varepsilon_{ni} \varepsilon_{nj}\} \\ &= \sum_{i=1}^D \sum_{j=1}^D \omega_i \omega_j \delta_{ij} \sigma^2 \\ &= \sum_{i=1}^D \omega_i^2 \sigma^2 \end{aligned}$$

Substituting the above results in equation 3, we obtain:

$$E\{E'_D(x, \omega)\} = \frac{1}{2} \sum_{n=1}^N \left\{ (y(x_n, \omega) - t_n)^2 + \sum_{i=1}^D \omega_i^2 \sigma^2 \right\}$$

final equation will be:

$$\mathbf{E}\{\mathbf{E}'_D(\mathbf{x}, \boldsymbol{\omega})\} = \mathbf{E}_D(\boldsymbol{\omega}) + \frac{N}{2} \sum_{i=1}^D \omega_i^2 \sigma^2$$

$\therefore$  Here, we get a weight decay (L2 regularization term in which bias term ( $\omega_0$ ) is absent).

## Solution 2 : Multi – Output Regression [Theory]

**Solution 2 (Part-1): Provide the expression for the likelihood, and derive ML and MAP estimates of  $\mathbf{W}$  in the multi output regression case.**

### 1. a) For MLE:

Multiple outputs: In some applications, we may wish to predict  $K > 1$  target variables, which we denote collectively by the target vector  $\mathbf{t}$ . This could be done by introducing a different set of basis functions for each component of  $\mathbf{t}$ , leading to multiple, independent regression problems. However, a more interesting, and more common, approach is to use the same set of basis functions to model all of the components of the target vector so that

$$\mathbf{y}(\mathbf{x}, \mathbf{w}) = \mathbf{W}^T \boldsymbol{\phi}(\mathbf{x})$$

where  $\mathbf{y}$  is a  $K$ -dimensional column vector,  $\mathbf{W}$  is an  $M \times K$  matrix of parameters, and  $\boldsymbol{\phi}(\mathbf{x})$  is an  $M$ -dimensional column vector with elements  $\phi_j(\mathbf{x})$ , with  $\phi_0(\mathbf{x})=1$  as before. Suppose we take the conditional distribution of the target vector to be an isotropic Gaussian of the form

$$p(\mathbf{t}|\mathbf{x}, \mathbf{W}, \beta) = \mathcal{N}(\mathbf{t}|\mathbf{W}^T \boldsymbol{\phi}(\mathbf{x}), \beta^{-1} \mathbf{I})$$

Taking Log we get:-

$$\begin{aligned} \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) &= \sum_{n=1}^N \ln \mathcal{N}(t_n | \mathbf{W}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1} \mathbf{I}) \\ &= \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) - \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \boldsymbol{\phi}(\mathbf{x}_n)\|^2 \end{aligned}$$

equating  $\text{argmax}_{\mathbf{W}} \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) = 0$  and solving for  $\mathbf{W}$  will give us  $\mathbf{W}_{\text{ML}}$  (estimated)

$$\begin{aligned} \text{argmax}_{\mathbf{W}} \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) &= \text{argmax}_{\mathbf{W}} \left( \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) - \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \boldsymbol{\phi}(\mathbf{x}_n)\|^2 \right) \\ \text{argmax}_{\mathbf{W}} \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) &= \text{argmax}_{\mathbf{W}} \left( \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) \right) + \text{argmax}_{\mathbf{W}} \left( -\frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \boldsymbol{\phi}(\mathbf{x}_n)\|^2 \right) \end{aligned}$$

We know  $\operatorname{argmax}_W (X) = \operatorname{argmin}_W (-X)$

$$\operatorname{argmax}_W \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) = 0 + \operatorname{argmin}_W \left( \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \phi(x_n)\|^2 \right)$$

$$\begin{aligned} \operatorname{argmax}_W \ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, \beta) &= \frac{\beta}{2} \cdot 2 \sum_{n=1}^N (t_n - \mathbf{W}^T \phi(x_n)) (0 - \phi^T(x_n)) = 0 \\ &= \sum_{n=1}^N (-t_n \phi^T(x_n) + \mathbf{W}^T \phi(x_n) \phi^T(x_n)) = 0 \end{aligned}$$

Taking  $-t_n \phi^T(x_n)$  to right side of equal sign and taking transpose, we get:

$$\sum_{n=1}^N (\mathbf{W}^T \phi(x_n) \phi^T(x_n))^T = \sum_{n=1}^N (t_n \phi^T(x_n))^T$$

Solving transpose we get:-

$$\sum_{n=1}^N \phi(x_n) \phi^T(x_n) \cdot W = \sum_{n=1}^N t_n^T \cdot \phi(x_n)$$

Finally :-

$$W = \sum_{n=1}^N (\phi(x_n) \phi^T(x_n))^{-1} \sum_{n=1}^N t_n^T \cdot \phi(x_n)$$

This can be written as:-

$$W_{ML} = (\Phi^T \Phi)^{-1} \cdot (T^T \Phi)$$

**Hence, we get answer in closed form**

**1. b) For MAP:**

- Prior

Suppose we take the prior distribution to be an isotropic Gaussian of the form

$$\begin{aligned} P(W|\alpha) &= \mathcal{N}(W|0, \alpha^{-1}\mathbf{I}) \\ &= \frac{1}{(\sqrt{2\pi}\alpha^{-1})^k} \exp\left(-\frac{(W-0)^2}{2(\alpha^{-1})}\right) \\ &= \left(\frac{\alpha}{2\pi}\right)^{\frac{k}{2}} \exp\left(\frac{-\alpha\|W\|^2}{2}\right) \end{aligned}$$

- Posterior

Posterior can be written as

$$P(W|X, t, \alpha, \beta) = \frac{P(t|X, t, W, \beta) P(W|\alpha)}{P(t|X, \alpha, \beta)}$$

$$P(W|X, t, \alpha, \beta) \propto P(t|X, W, \beta) \cdot P(W|\alpha)$$

$$P(W|X, t, \alpha, \beta) = \left(\frac{\beta}{2\pi}\right)^{\frac{k}{2}} \exp\left(\frac{-\beta}{2}(t - \phi(x_n))^2\right) \cdot \left(\frac{\alpha}{2\pi}\right)^{\frac{k}{2}} \exp\left(\frac{-\alpha}{2}W^T W\right)$$

Taking log both sides we get-

$$\begin{aligned} \ln(P(W|X, t, \alpha, \beta)) &= \ln \sum_{n=1}^N \left(\frac{\beta}{2\pi}\right)^{\frac{k}{2}} \cdot \exp\left[\frac{-\beta}{2}(t - \phi(x_n))^2\right] + n \sum_{n=1}^N \left(\frac{\alpha}{2\pi}\right)^{\frac{k}{2}} \exp\left(\frac{-\alpha}{2}W^T W\right) \\ &= \frac{NK}{2} \ln\left(\frac{\beta}{2\pi}\right) - \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \phi(x_n)\|^2 + \frac{NK}{2} \ln\left(\frac{\alpha}{2\pi}\right) - \frac{\alpha}{2} \sum_{n=1}^N (W^T W) \end{aligned}$$

Equating  $\text{argmax}_W (\ln P(W|X, t, \alpha, \beta)) = 0$  and solving for W will give us W(MAP)

$$\text{argmax}_W \left( \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) - \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \phi(x_n)\|^2 + \frac{NK}{2} \ln \left( \frac{\alpha}{2\pi} \right) - \frac{\alpha}{2} \sum_{n=1}^N \|W\|^2 \right) = 0 \quad - 1$$

$$\bullet \quad \text{argmax}_W \left( \frac{NK}{2} \ln \left( \frac{\beta}{2\pi} \right) \right) = 0 \quad - 2$$

$$\bullet \quad \text{argmax}_W \left( \frac{NK}{2} \ln \left( \frac{\alpha}{2\pi} \right) \right) = 0 \quad - 3$$

Put equation 2 , and equation 3 in equation we get

$$\text{argmax}_W \left( -\frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \phi(x_n)\|^2 - \sum_{n=1}^N \frac{\alpha}{2} \|W\|^2 \right) = 0$$

We know  $\text{argmax}_W (X) = \text{argmin}_W (-X)$

$$\text{argmin}_W \left( \frac{\beta}{2} \sum_{n=1}^N \|t_n - \mathbf{W}^T \phi(x_n)\|^2 + \frac{\alpha}{2} \sum_{n=1}^N \|W\|^2 \right) = 0$$

$$\frac{\beta}{2} \cdot 2 \sum_{n=1}^N (t_n - \mathbf{W}^T \phi(x_n)) (0 - \phi^T(x_n)) + \frac{\alpha}{2} \cdot 2 \sum (W^T) = 0$$

$$\sum_{n=1}^N \left( \beta (-t_n \phi^T(x_n) + \mathbf{W}^T \phi(x_n) \phi^T(x_n)) + \alpha \cdot W^T \right) = 0$$

$$\sum_{n=1}^N (-t_n \phi^T(x_n) + \mathbf{W}^T \phi(x_n) \phi^T(x_n) + \lambda \cdot W^T) = 0$$

where  $\lambda = \frac{\alpha}{\beta}$



Taking  $\sum_{n=1}^N t_n \phi^T(x_n)$  term to right side of equal sign

$$\sum_{n=1}^N \mathbf{w}^T \phi(x_n) \phi^T(x_n) + \lambda \cdot W^T = \sum_{n=1}^N t_n \phi^T(x_n)$$

$$\sum_{n=1}^N (W^T)(\phi(x_n) \phi^T(x_n) + \lambda) = \sum_{n=1}^N t_n \phi^T(x_n)$$

Taking Transpose, and solving for W.

$$\sum_{n=1}^N (\phi^T(x_n) \phi(x_n) + \lambda)(W) = \sum_{n=1}^N \phi(x_n) t_n^T$$

Finally we get-

$$W = \sum_{n=1}^N (\phi^T(x_n) \phi(x_n) + \lambda)^{-1} \phi(x_n) t_n^T$$

Which can be written as closed form

$$\mathbf{W}_{MAP} = (\Phi^T \Phi + \lambda \mathbf{I})^{-1} \Phi$$

Where  $\lambda = \frac{\alpha}{\beta}$

We see that this is somewhat similar to WMLE, the only difference is the  $\lambda \mathbf{I}$  is added to  $\Phi^T \Phi$ .

### Solution 2 (Part-2): Find the MLE for w1 and w2

$$\bullet \quad X = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \\ 0 & 1 \end{bmatrix}$$

$$\bullet \quad X^T = \begin{bmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix}$$

$$\bullet \quad Y = \begin{bmatrix} -1 & -1 \\ -1 & -2 \\ -2 & -1 \\ 1 & 1 \\ 1 & 2 \\ 2 & 1 \end{bmatrix}$$

We Derived above the closed form solution for w in case of multi output regression as-

$$W = (\varphi^T \varphi) \varphi^T Y$$

Here W will be 2X2 matrix

Finding values according to closed form solution-

$$\bullet \quad \varphi^T \varphi = \begin{bmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} = \begin{bmatrix} 3 & 0 \\ 0 & 3 \end{bmatrix}$$

$$\bullet \quad (\varphi^T \varphi)^{-1} = \frac{Adj(\varphi^T \varphi)}{|\varphi^T \varphi|} = \frac{\begin{bmatrix} 3 & 0 \\ 0 & 3 \end{bmatrix}}{9}$$

$$\bullet \quad \varphi^T Y = \begin{bmatrix} 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} -1 & -1 \\ -1 & -2 \\ -2 & -1 \\ 1 & 1 \\ 1 & 2 \\ 2 & 1 \end{bmatrix} = \begin{bmatrix} -4 & -4 \\ 4 & 4 \end{bmatrix}$$

Substituting above Matrix forms in our equation  $\mathbf{W} = (\boldsymbol{\varphi}^T \boldsymbol{\varphi}) \boldsymbol{\varphi}^T \mathbf{Y}$  we get

$$W_{2 \times 2} = \frac{1}{9} \begin{bmatrix} 3 & 0 \\ 0 & 3 \end{bmatrix} \begin{bmatrix} -4 & -4 \\ 4 & 4 \end{bmatrix}$$

Further Solving this we get-

$$\begin{bmatrix} \omega_1 & \omega_2 \\ \vdots & \vdots \end{bmatrix} = \begin{bmatrix} \frac{-4}{3} & \frac{-4}{3} \\ \frac{4}{3} & \frac{4}{3} \end{bmatrix}$$

$w_1$  and  $w_2$  are

$$\omega_1 = \begin{bmatrix} \frac{-4}{3} \\ \frac{4}{3} \end{bmatrix}$$

$$\omega_2 = \begin{bmatrix} \frac{-4}{3} \\ \frac{4}{3} \end{bmatrix}$$

# Answer 3 [Programming Question] : MAP and ML estimation of Poisson Distribution (Modeling the horse kick deaths)

## Importing Libraries

```
In [1]: import pandas as pd
from scipy.special import factorial
import scipy as scy
import scipy.stats as stats
from scipy.stats import poisson
import numpy as np
import matplotlib.pyplot as plt
```

Given data is divided in training data(13 years) and test data(7 years)

Training Data : 1875 to 1887 to model the Poisson distribtion. Test Data : 1888 to 1894 to test the model.

## Importing Training data

```
In [2]: training = pd.read_csv('training_set.csv', index_col = 'YEAR')
training
```

```
Out[2]:
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI	XIV	XV
YEAR														
1875	0	0	0	0	0	0	0	1	1	0	0	0	1	0
1876	2	0	0	0	1	0	0	0	0	0	0	0	1	1
1877	2	0	0	0	0	0	1	1	0	0	1	0	2	0
1878	1	2	2	1	1	0	0	0	0	0	1	0	1	0
1879	0	0	0	1	1	2	2	0	1	0	0	2	1	0
1880	0	3	2	1	1	1	0	0	0	2	1	4	3	0
1881	1	0	0	2	1	0	0	1	0	1	0	0	0	0
1882	1	2	0	0	0	0	1	0	1	1	2	1	4	1
1883	0	0	1	2	0	1	2	1	0	1	0	3	0	0
1884	3	0	1	0	0	0	0	1	0	0	2	0	1	1
1885	0	0	0	0	0	0	1	0	0	2	0	1	0	1
1886	2	1	0	0	1	1	1	0	0	1	0	1	3	0
1887	1	1	2	1	0	0	3	2	1	1	0	1	2	0

## Importing Test data

```
In [3]: test_data = pd.read_csv('test_set.csv', index_col = 'YEAR')
test_data
```

```
Out[3]:
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI	XIV	XV
YEAR														
1888	0	1	1	0	0	1	1	0	0	0	0	1	1	0
1889	0	0	1	1	0	1	1	0	0	1	2	2	0	2
1890	1	2	0	2	0	1	1	2	0	2	1	1	2	2
1891	0	0	0	1	1	1	0	1	1	0	3	3	1	0
1892	1	3	2	0	1	1	3	0	1	1	0	1	1	0
1893	0	1	0	0	0	1	0	2	0	0	1	3	0	0
1894	1	0	0	0	0	0	0	0	1	0	1	1	0	0

```
In [4]: headings = list(test_data)
```

```
print(headings)
```

```
['G', 'I', 'II', 'III', 'IV', 'V', 'VI', 'VII', 'VIII', 'IX', 'X', 'XI', 'XIV', 'XV']
```

## important stats of training data

```
In [5]: training.describe()
```

```
Out[5]:
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI
count	13.0	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000	13.000000
mean	1.0	0.692308	0.615385	0.615385	0.461538	0.384615	0.846154	0.538462	0.307692	0.692308	0.538462	1.000000
std	1.0	1.031553	0.869718	0.767948	0.518875	0.650444	0.987096	0.660225	0.480384	0.751068	0.776250	1.290994
min	0.0	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
25%	0.0	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
50%	1.0	0.000000	0.000000	0.000000	0.000000	0.000000	1.000000	0.000000	0.000000	1.000000	0.000000	1.000000
75%	2.0	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000
max	3.0	3.000000	2.000000	2.000000	1.000000	2.000000	3.000000	2.000000	1.000000	2.000000	2.000000	4.000000

The given data has been modeled using Poisson distribution.

## Poisson Distribution:

PDF for poissos is given as

$$p(k \text{ events in interval}) = \frac{e^{-\lambda} \lambda^k}{k!} \quad (1)$$

where  $\lambda$  is the average (mean) no of events per interval and  $k$  is no of events occurred independently.

## Part 1: MLE to learn the parameters

Let  $X_1, X_2, \dots, X_n$  are independent Poisson random variables each having mean  $\lambda$ , then maximum likelihood estimator(MLE) of  $\lambda$  can be given as

$$\hat{\lambda}_{ML} = \frac{\sum_1^n x_i}{n} \quad (2)$$

### Example:

For G-Corps  $\hat{\lambda}_{ML}$  estimate will be

$$\hat{\lambda}_{ML} = \frac{0 + 2 + 2 + 1 + 0 + 0 + 1 + 1 + 0 + 3 + 0 + 2 + 1}{13} = 1 \quad (3)$$

```
In [6]: list1 = list(training.mean())
lambdaForEachCorps= pd.DataFrame([training.mean()],index = ['MLE (λ)'],columns = headings)
display(lambdaForEachCorps)
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI	XIV	XV
MLE (λ)	1.0	0.692308	0.615385	0.615385	0.461538	0.384615	0.846154	0.538462	0.307692	0.692308	0.538462	1.0	1.461538	0.307692

we use MLE ( $\hat{\lambda}_{ML}$ ), to predict no of deaths for remaining 7 years.

### Example:

For G-Corps

$$p(0 \text{ deaths in 7 years}) = \frac{e^{-1} 1^0}{0!} * 7 = 3 \quad (4)$$

```
In [7]: summed1 = 0
for idx,i in enumerate(list1):
```

```

t = np.arange(0, 5, 1)
d = np.exp(-i)*np.power(i, t)/factorial(t)
pred = d*7

actual_for_thisYear=(list(test_data[headings[idx]]))

sum1=0

predictedList = []

actualList = []

for idx2, num in enumerate(pred):
    a=round(num)
    predictedList.append(a)

    b=actual_for_thisYear.count(idx2)
    actualList.append(b)

    mse= round(np.power((a-b),2)/7,2)

    sum1 = sum1 + mse

RMSE = round(np.power(sum1,0.5),2)
sum1 = pd.DataFrame([[sum1 , RMSE]], index = [f'For {headings[idx]} corp'], columns = ['MSE','RMSE'])
summed1 = summed1 +RMSE
prediction= pd.Series(predictedList)
actual= pd.Series(actualList)
df = pd.DataFrame([predictedList,actualList], index = [ f'PREDICTION for {headings[idx]} corp',
                                                    f'ACTUAL for {headings[idx]} corp'],
                  columns= ['yearsWithZeroDeath','yearsWithOneDeath',
                            'yearsWithTwoDeath','yearsWithThreeDeath',
                            'yearsWithFourDeath'])

display(df)
display(sum1)
print('-----')
print(summed1)

```

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
<b>PREDICTION for G corp</b>	3	3	1	0	0
<b>ACTUAL for G corp</b>	4	3	0	0	0

	MSE	RMSE
<b>For G corp</b>	0.28	0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
<b>PREDICTION for I corp</b>	4	2	1	0	0
<b>ACTUAL for I corp</b>	3	2	1	1	0

	MSE	RMSE
<b>For I corp</b>	0.28	0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
<b>PREDICTION for II corp</b>	4	2	1	0	0
<b>ACTUAL for II corp</b>	4	2	1	0	0

	MSE	RMSE
<b>For II corp</b>	0.0	0.0

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
<b>PREDICTION for III corp</b>	4	2	1	0	0
<b>ACTUAL for III corp</b>	4	2	1	0	0

	MSE	RMSE
For III corp	0.0	0.0

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for IV corp	4	2	0	0	0
ACTUAL for IV corp	5	2	0	0	0

	MSE	RMSE
For IV corp	0.14	0.37

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for V corp	5	2	0	0	0
ACTUAL for V corp	1	6	0	0	0

	MSE	RMSE
For V corp	4.58	2.14

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VI corp	3	3	1	0	0
ACTUAL for VI corp	3	3	0	1	0

	MSE	RMSE
For VI corp	0.28	0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VII corp	4	2	1	0	0
ACTUAL for VII corp	4	1	2	0	0

	MSE	RMSE
For VII corp	0.28	0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VIII corp	5	2	0	0	0
ACTUAL for VIII corp	4	3	0	0	0

	MSE	RMSE
For VIII corp	0.28	0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for IX corp	4	2	1	0	0
ACTUAL for IX corp	4	2	1	0	0

	MSE	RMSE
For IX corp	0.0	0.0

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for X corp	4	2	1	0	0
ACTUAL for X corp	2	3	1	1	0

	MSE	RMSE
For X corp	0.85	0.92

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XI corp	3	3	1	0	0
ACTUAL for XI corp	0	4	1	2	0

	MSE	RMSE
For XI corp	2.0	1.41

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XIV corp	2	2	2	1	0
ACTUAL for XIV corp	3	3	1	0	0

	MSE	RMSE
For XIV corp	0.56	0.75

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XV corp	5	2	0	0	0
ACTUAL for XV corp	5	0	2	0	0

	MSE	RMSE
For XV corp	1.14	1.07

9.310000000000002

## 2) Using maximum aposteriori estimation to learn parameters

**a) Prior Assumption:** Gamma Distribution with parameters  $\text{shape}(\alpha) = 2.5$  and  $\text{rate}(\beta) = 1$

$$\lambda \sim \Gamma(\alpha, \beta)$$

- Justification: given distribution is Poisson distribution which supports interval from 0 to  $\infty$  and the rate of deaths per year  $\approx 1$ .

Hence we have choosen gamma distribution as the prior distribution, which supports from interval 0 to  $\infty$  and the  $\text{rate}(\beta)$  parameter =1 and  $\text{shape}(\alpha) = 2$

The pdf of gamma distribution can be given as

$$f(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

**Maximum A-posteriori (MAP) Estimation:**

$$\text{prob}(\lambda|y) = \frac{\text{prob}(y|\lambda) \cdot \text{prob}(\lambda)}{\text{prob}(y)}$$

$$\hat{\lambda}_{MAP} = \underset{\lambda}{\operatorname{argmax}} (\prod \text{prob}(y|\lambda) \cdot \text{prob}(\lambda))$$

Using Loglokelihood

$$\hat{\lambda}_{MAP} = \underset{\lambda}{\operatorname{argmax}} (\sum \log \text{prob}(y|\lambda) + \log \text{prob}(\lambda))$$

by substituting likelihood and prior distributions

$$\hat{\lambda}_{MAP} = \underset{\lambda}{\operatorname{argmax}} (\sum \log(\lambda^{\sum y_i} e^{-n\lambda}) + \log(\beta^\alpha \lambda^{\alpha-1} e^{-\beta\lambda}))$$



$$\hat{\lambda}_{MAP} = \underset{\lambda}{argmax} (\log(\lambda^{\sum y_i} e^{-n\lambda}) + \log(\beta^\alpha \lambda^{\alpha-1} e^{-\beta\lambda}))$$

MAP estimate will be obtained by equating the gradient of above eq to 0.

$$\frac{\alpha - 1}{\lambda} - \beta + \frac{\sum y_i}{\lambda} - n = 0$$

$$\therefore \hat{\lambda}_{MAP} = \frac{\alpha + \sum y_i - 1}{\beta + n}$$

### Example:

For G-Corps

$$\hat{\lambda}_{MAP} = \frac{2.5 + (0 + 2 + 2 + 1 + 0 + 0 + 1 + 1 + 0 + 3 + 0 + 2 + 1) - 1}{1 + 13} = 1.035 \quad (5)$$

we use  $(\hat{\lambda}_{MAP})$ , to predict no of deaths for remaining 7 years.

$$p(k \text{ deaths in 7 years}) = \frac{e^{-\hat{\lambda}_{MAP}} \hat{\lambda}_{MAP}^k}{k!} * 7 \quad (6)$$

In [8]:

```
α = 2.5
β = 1
lambdaForMAP = []
for i in training:
    sum1 = np.sum(training[i])
    n = len(training)
    lambdaForMAP.append((α + sum1 - 1)/(β+n))
lambdaMAPForEachCorps = pd.DataFrame([lambdaForMAP], index = ['MAP (λ)'], columns = headings)
display(lambdaMAPForEachCorps)
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI	XIV	XV
MAP (λ)	1.035714	0.75	0.678571	0.678571	0.535714	0.464286	0.892857	0.607143	0.392857	0.75	0.607143	1.035714	1.464286	0.392857

In [9]:

```
summed = 0
for idx, i in enumerate(lambdaForMAP):
    t = np.arange(0, 5, 1)
    d = np.exp(-i)*np.power(i, t)/factorial(t)
    pred1 = d*7

    actual_for_thisYear=(list(test_data[headings[idx]]))

    sum11=0

    predictedListMAP = []

    actualList = []

    for idx2, num in enumerate(pred1):
        a=round(num)
        predictedListMAP.append(a)

        b=actual_for_thisYear.count(idx2)
        actualList.append(b)

        mse= round(np.power((a-b),2)/7,2)

        sum11 = sum11 + mse

    RMSE = round(np.power(sum11,0.5),2)
    summed = summed +RMSE
    sum11 = pd.DataFrame([[sum11 , RMSE]], index = [f'For {headings[idx]} corp'], columns = ['MSE','RMSE'])

    prediction= pd.Series(predictedListMAP)
    actual= pd.Series(actualList)
    df1 = pd.DataFrame([prediction,actual], index = [f'PREDICTION for {headings[idx]} corp',
                                                    f'ACTUAL for {headings[idx]} corp'],
                        columns= ['yearsWithZeroDeath','yearsWithOneDeath',
```

```
'yearsWithTwoDeath', 'yearsWithThreeDeath',
'yearsWithFourDeath']])
```

```
display(df1)
display(sum1)
print(summed)
```

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for G corp	2	3	1	0	0
ACTUAL for G corp	4	3	0	0	0

MSE RMSE

For G corp 0.71 0.84

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for I corp	3	2	1	0	0
ACTUAL for I corp	3	2	1	1	0

MSE RMSE

For I corp 0.14 0.37

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for II corp	4	2	1	0	0
ACTUAL for II corp	4	2	1	0	0

MSE RMSE

For II corp 0.0 0.0

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for III corp	4	2	1	0	0
ACTUAL for III corp	4	2	1	0	0

MSE RMSE

For III corp 0.0 0.0

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for IV corp	4	2	1	0	0
ACTUAL for IV corp	5	2	0	0	0

MSE RMSE

For IV corp 0.28 0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for V corp	4	2	0	0	0
ACTUAL for V corp	1	6	0	0	0

MSE RMSE

For V corp 3.58 1.89

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VI corp	3	3	1	0	0
ACTUAL for VI corp	3	3	0	1	0

MSE RMSE

For VI corp 0.28 0.53

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VII corp	4	2	1	0	0
ACTUAL for VII corp	4	1	2	0	0
	MSE	RMSE			
For VII corp	0.28	0.53			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for VIII corp	5	2	0	0	0
ACTUAL for VIII corp	4	3	0	0	0
	MSE	RMSE			
For VIII corp	0.28	0.53			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for IX corp	3	2	1	0	0
ACTUAL for IX corp	4	2	1	0	0
	MSE	RMSE			
For IX corp	0.14	0.37			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for X corp	4	2	1	0	0
ACTUAL for X corp	2	3	1	1	0
	MSE	RMSE			
For X corp	0.85	0.92			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XI corp	2	3	1	0	0
ACTUAL for XI corp	0	4	1	2	0
	MSE	RMSE			
For XI corp	1.28	1.13			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XIV corp	2	2	2	1	0
ACTUAL for XIV corp	3	3	1	0	0
	MSE	RMSE			
For XIV corp	0.56	0.75			

	yearsWithZeroDeath	yearsWithOneDeath	yearsWithTwoDeath	yearsWithThreeDeath	yearsWithFourDeath
PREDICTION for XV corp	5	2	0	0	0
ACTUAL for XV corp	5	0	2	0	0
	MSE	RMSE			
For XV corp	1.14	1.07			

9.46

Since the Posteriori will be a Gamma distribution

$$\lambda \sim \Gamma(\alpha + \sum y_i, \beta + n)$$

The values of shape and rate parameters calculated each Crop wise.

## Mode for MLE, Mode for Prior and Mode for MAP are calculated below

```
In [10]: modeForMLE = np.floor(list1)
# print(modeForMLE)
```

mode = For non-integer  $\lambda$ , it is the largest integer less than  $\lambda$ . For integer  $\lambda$ ,  $x = \lambda$  and  $x = \lambda - 1$  are both the mode.

```
In [11]: modeForPrior = []
for i in training:
    modePrior = round(( $\alpha$ -1) /  $\beta$  , 2)
    modeForPrior.append(modePrior)
# print(modeForPrior)
```

```
In [12]: modeForMAP = [ round(elem, 2) for elem in lambdaForMAP ]
# print(modeForMAP)
```

```
In [13]: mode = pd.DataFrame([modeForMLE, modeForPrior, modeForMAP], index = ['Mode For MLE', 'Mode For Prior', 'Mode For MAP'],
                             columns = headings)
display(mode)
```

	G	I	II	III	IV	V	VI	VII	VIII	IX	X	XI	XIV	XV
Mode For MLE	1.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	1.00	1.00	0.00
Mode For Prior	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50	1.50
Mode For MAP	1.04	0.75	0.68	0.68	0.54	0.46	0.89	0.61	0.39	0.75	0.61	1.04	1.46	0.39

## Part 2: Using maximum a posteriori estimation to learn parameters

- **For II - Corps:**

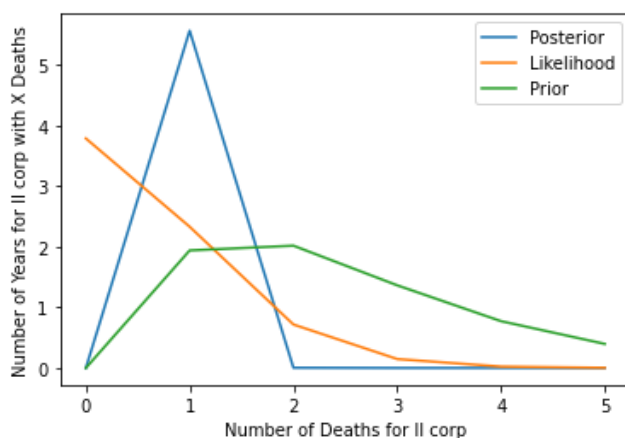
Likelihood mode = 0.00 prior mode = 1.5 Posteriori mode = 0.68

mode of Posteriori distribution is in between likelihood and prior distributions

Posterior, Likelihood and Prior for Number of Deaths for II corp

```
In [14]: x = np.arange(0.0, 6.0, 1)

# a = plt.plot(x, poisson.pmf(x, lambdaForMAP[2]), label='Posterior')
a = plt.plot(x, 7*stats.gamma.pdf(x, a =  $\alpha$  + np.sum(training["II"]), scale=1/( $\beta$ +13)), label='Posterior')
b = plt.plot(x, 7*poisson.pmf(x, list1[2]), label='Likelihood')
c = plt.plot(x, 7*stats.gamma.pdf(x, a= $\alpha$ , scale=1/ $\beta$ ), label='Prior')
plt.legend()
plt.xlabel('Number of Deaths for II corp')
plt.ylabel('Number of Years for II corp with X Deaths')
plt.show()
```



- **For IV - Corps:**

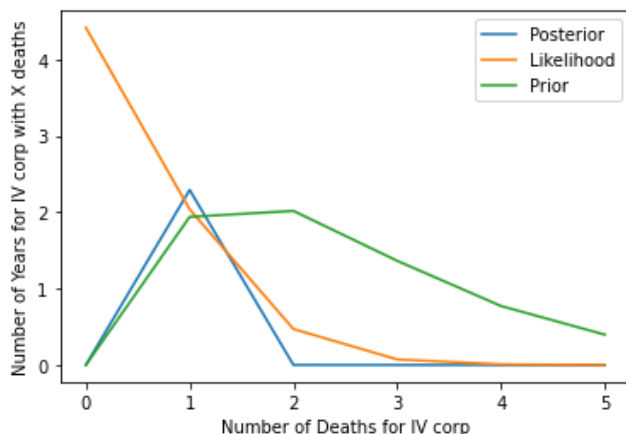
Likelihood mode = 0.00 prior mode = 1.5 Posteriori mode = 0.54

mode of Posteriori distribution is in between likelihood and prior distributions

### Posterior,Likelihood and Prior for Number of Deaths for IV corp

```
In [15]: x= np.arange(0.0,6.0,1)

#a=plt.plot(x, poisson.pmf(x, LambdaForMAP[4]),label='Posterior')
a=plt.plot(x, 7*stats.gamma.pdf(x,a=  $\alpha$  + np.sum(training["IV"]) , scale=1/( $\beta$ +13) ),label='Posterior')
b=plt.plot(x, 7*poisson.pmf(x,list1[4]),label='Likelihood')
c=plt.plot(x, 7*stats.gamma.pdf(x, a= $\alpha$ , scale=1/ $\beta$ ),label='Prior')
plt.legend()
plt.xlabel('Number of Deaths for IV corp')
plt.ylabel('Number of Years for IV corp with X deaths')
plt.show()
```



- For VI - Corps:

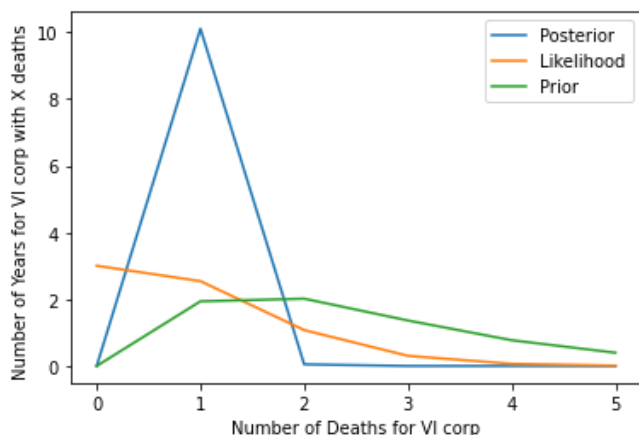
Likelihood mode = 0.00 prior mode = 1.5 Posteriori mode = 0.89

mode of Posteriori distribution is in between likelihood and prior distributions

### Posterior,Likelihood and Prior for Number of Deaths for VI corp

```
In [16]: x= np.arange(0.0,6.0,1)

#a=plt.plot(x, poisson.pmf(x, LambdaForMAP[6]),label='Posterior')
a=plt.plot(x, 7*stats.gamma.pdf(x,a=  $\alpha$  + np.sum(training["VI"]) , scale=1/( $\beta$ +13) ),label='Posterior')
b=plt.plot(x, 7*poisson.pmf(x,list1[6]),label='Likelihood')
c=plt.plot(x, 7*stats.gamma.pdf(x, a= $\alpha$ , scale=1/ $\beta$ ),label='Prior')
plt.legend()
plt.xlabel('Number of Deaths for VI corp')
plt.ylabel('Number of Years for VI corp with X deaths ')
plt.show()
```



## Question 4: Bike Sharing Demand

**Part 1: Explain maximum likelihood estimation in poisson regression and derive the loss function which is used to estimate the parameters.**

mean of poisson's equation:

$$\lambda = \mathbb{E}[Y|x] = e^{\theta^T x} \quad (1)$$

pdf for poisson's eqn is given as

$$p(y|x, \theta) = \frac{\lambda^y e^{-\lambda}}{y!} = \frac{e^{y\theta^T x} e^{-e^{\theta^T x}}}{y!} \quad (2)$$

the probability for observing  $y_1, \dots, y_n$  is given as

$$p(y_1, \dots, y_m | x_1, \dots, x_m; \theta) = \prod_{i=1}^m \frac{e^{y_i \theta^T x_i} e^{-e^{\theta^T x_i}}}{y_i!}. \quad (3)$$

the above equation can be represented as likelihood function

$$L(\theta | X, Y) = \prod_{i=1}^m \frac{e^{y_i \theta^T x_i} e^{-e^{\theta^T x_i}}}{y_i!}. \quad (4)$$

taking log of above likelihood function gives us log likelihood function

$$\ell(\theta | X, Y) = \log L(\theta | X, Y) = \sum_{i=1}^m \left( y_i \theta^T x_i - e^{\theta^T x_i} - \log(y_i!) \right). \quad (5)$$

$$\ell(\theta | X, Y) = \sum_{i=1}^m \left( y_i \theta^T x_i - e^{\theta^T x_i} \right). \quad (6)$$

minimising log likelihood and solving for theta

$$\underset{\theta}{\text{minimize}} \ell(\theta | X, Y) = \sum_{i=1}^m \left( e^{\theta^T x_i} - y_i \theta^T x_i \right). \quad (7)$$

$$\nabla \ell(\theta | X, Y) = \frac{\partial \ell(\theta | X, Y)}{\partial \theta} = \sum_{i=1}^m \left( e^{\theta^T x_i} - y_i \right) x_i \quad (8)$$

we get theta as:-

$$\theta := \theta - \alpha \nabla \ell(\theta | X, Y) \quad (9)$$

### Importing Libraries

```
In [1]: import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns
%matplotlib notebook
```

### Importing Data files

```
In [2]: train = pd.read_csv("train.csv")
test = pd.read_csv("test.csv")
```

```
In [3]: train.head(2)
display(test.head(2))
```

	datetime	season	holiday	workingday	weather	temp	atemp	humidity	windspeed
0	2011-01-20 00:00:00	1	0	1	1	10.66	11.365	56	26.0027
1	2011-01-20 01:00:00	1	0	1	1	10.66	13.635	56	0.0000

## Processing Given Data

```
In [4]: train['year'] = pd.DatetimeIndex(train['datetime']).year
train['month'] = pd.DatetimeIndex(train['datetime']).month
train['day'] = pd.DatetimeIndex(train['datetime']).day
train['dayOfWeek'] = pd.DatetimeIndex(train['datetime']).weekday
train['hour'] = pd.DatetimeIndex(train['datetime']).hour
```

```
In [5]: test['year'] = pd.DatetimeIndex(test['datetime']).year
test['month'] = pd.DatetimeIndex(test['datetime']).month
test['day'] = pd.DatetimeIndex(test['datetime']).day
test['dayOfWeek'] = pd.DatetimeIndex(test['datetime']).weekday
test['hour'] = pd.DatetimeIndex(test['datetime']).hour
```

```
In [6]: train = train[['year', 'season', 'month', 'day', 'dayOfWeek', 'hour', 'holiday', 'workingday',
'weather', 'temp', 'atemp', 'humidity', 'windspeed', 'count']]
train.head(2)
```

```
Out[6]:
```

	year	season	month	day	dayOfWeek	hour	holiday	workingday	weather	temp	atemp	humidity	winds
0	2011	1	1	1	5	0	0	0	1	9.84	14.395	81	
1	2011	1	1	1	5	1	0	0	1	9.02	13.635	80	

```
In [7]: test = test[['year', 'season', 'month', 'day', 'dayOfWeek', 'hour', 'holiday', 'workingday',
'weather', 'temp', 'atemp', 'humidity', 'windspeed']]
```

```
In [8]: test.isnull().sum()
train.isnull().sum()
```

```
Out[8]: year          0
season          0
month          0
day            0
dayOfWeek      0
hour           0
holiday        0
workingday     0
weather        0
temp           0
atemp          0
humidity       0
windspeed      0
count          0
dtype: int64
```

## Part 2: finding mean count per year, month, day, working day etc

```
In [9]: meanY = train.groupby('year', axis = 0).mean()
print(meanY['count'])
```

```
year
2011    144.223349
2012    238.560944
Name: count, dtype: float64
```

```
In [10]: meanS = train.groupby('season', axis = 0).mean()  
print(meanS['count'])
```

```
season  
1    116.343261  
2    215.251372  
3    234.417124  
4    198.988296  
Name: count, dtype: float64
```

```
In [11]: meanM = train.groupby('month', axis = 0).mean()  
print(meanM['count'])
```

```
month  
1      90.366516  
2     110.003330  
3     148.169811  
4     184.160616  
5     219.459430  
6     242.031798  
7     235.325658  
8     234.118421  
9     233.805281  
10    227.699232  
11    193.677278  
12    175.614035  
Name: count, dtype: float64
```

```
In [12]: meanD = train.groupby('day', axis = 0)['count'].mean()  
print(meanD)
```

```
day  
1      180.333913  
2      183.910995  
3      194.696335  
4      195.705575  
5      189.765217  
6      189.860140  
7      183.773519  
8      179.041812  
9      187.897391  
10     195.183566  
11     195.679577  
12     190.675393  
13     194.160279  
14     195.829268  
15     201.527875  
16     191.353659  
17     205.660870  
18     192.605684  
19     192.311847  
Name: count, dtype: float64
```

```
In [13]: meanDW = train.groupby('dayOfWeek', axis = 0).mean()  
print(meanDW['count'])
```

```
dayOfWeek  
0      190.390716  
1      189.723847  
2      188.411348  
3      197.296201  
4      197.844343  
5      196.665404  
6      180.839772  
Name: count, dtype: float64
```

```
In [14]: meanH = train.groupby('hour', axis = 0).mean()  
print(meanH['count'])
```

```
hour  
0      55.138462  
1      33.859031  
2      22.899554  
3      11.757506  
4       6.407240  
5      19.767699  
6      76.259341
```



```

7      213.116484
8      362.769231
9      221.780220
10     175.092308
11     210.674725
12     256.508772
13     257.787281
14     243.442982
15     254.298246
16     316.372807
17     468.765351
18     430.859649
19     315.278509
20     228.517544
21     173.370614
22     133.576754
23      89.508772
Name: count, dtype: float64

```

```
In [15]: meanHol = train.groupby('holiday', axis = 0).mean()
print(meanHol['count'])
```

```

holiday
0      191.741655
1      185.877814
Name: count, dtype: float64

```

```
In [16]: meanWD = train.groupby('workingday', axis = 0).mean()
print(meanWD['count'])
```

```

workingday
0      188.506621
1      193.011873
Name: count, dtype: float64

```

```
In [17]: meanW = train.groupby('weather', axis = 0).mean()
print(meanW['count'])
```

```

weather
1      205.236791
2      178.955540
3      118.846333
4      164.000000
Name: count, dtype: float64

```

## Part 3: Plot count against any 5 features

### a) Hypothesis Generation

Before exploring the data to understand the relationship between variables, it is recommended that we focus on hypothesis generation first. Here are some of the hypothesis which could influence the demand of bikes:

#### Hourly trend:

There must be high demand during office timings. Early morning and late evening can have different trend (cyclist) and low demand during 10:00 pm to 4:00 am.

#### Daily Trend:

demand of bike increases on weekdays than on weekends and holidays.

#### Season Trend:

Using of bike can be related with seasons, as people may tend not to use bike in rainy season and winter, but may prefer using bike in summer and spring.

#### Temperature:

People tends to use more bike when temprature rises.

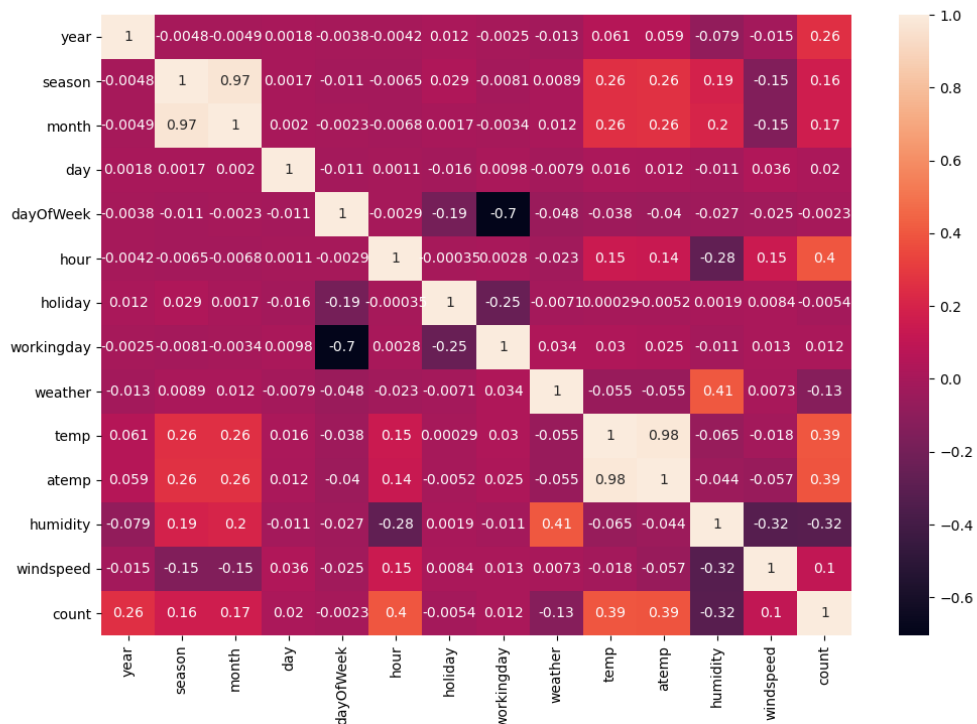
## Weather:

in good weather people are more likely to use bike compared to bad weather.

## b) Testing our Hypothesis with plots

```
In [18]: #pd.options.display.float_format = '{3,.2f}'.format
```

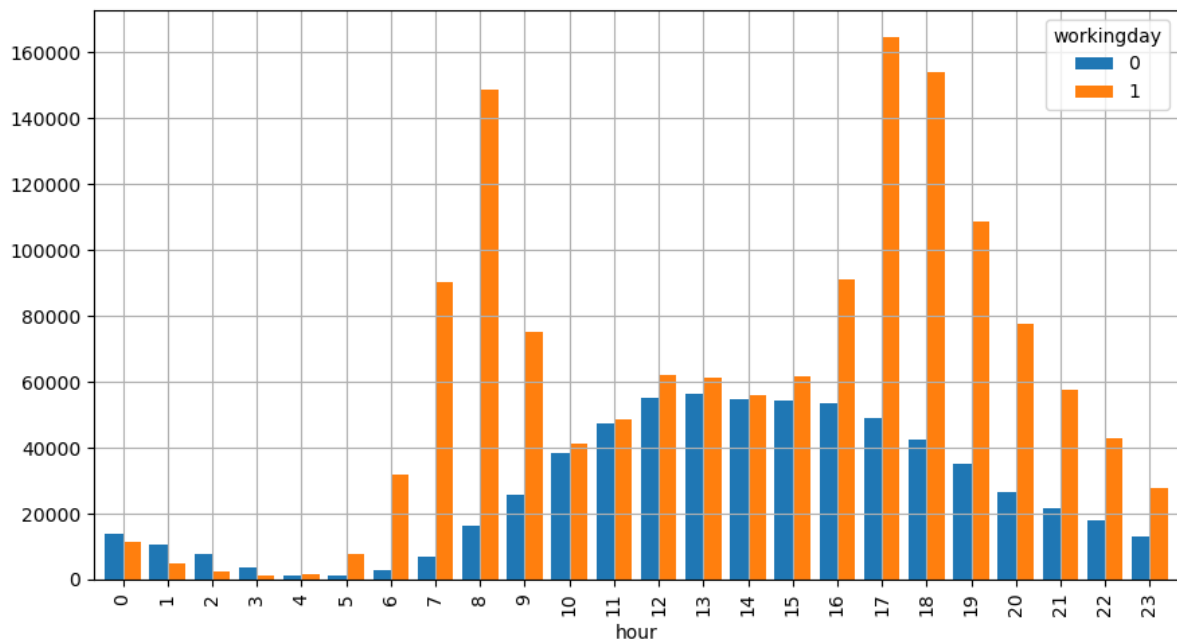
```
In [19]: correlation = train.corr()
plt.figure(figsize = (12, 8))
sns.heatmap(correlation, annot = True)
plt.show()
```



```
In [20]: by_hour = train.groupby(['hour', 'workingday'])['count'].agg('sum').unstack()
by_hour.head(3)
```

```
Out[20]: workingday    0    1
          hour
0  13701  11387
1  10427   4945
2   7686   2573
```

```
In [21]: by_hour.plot(kind='bar', figsize=(9,5), width=0.8);
plt.grid(True)
plt.tight_layout()
```



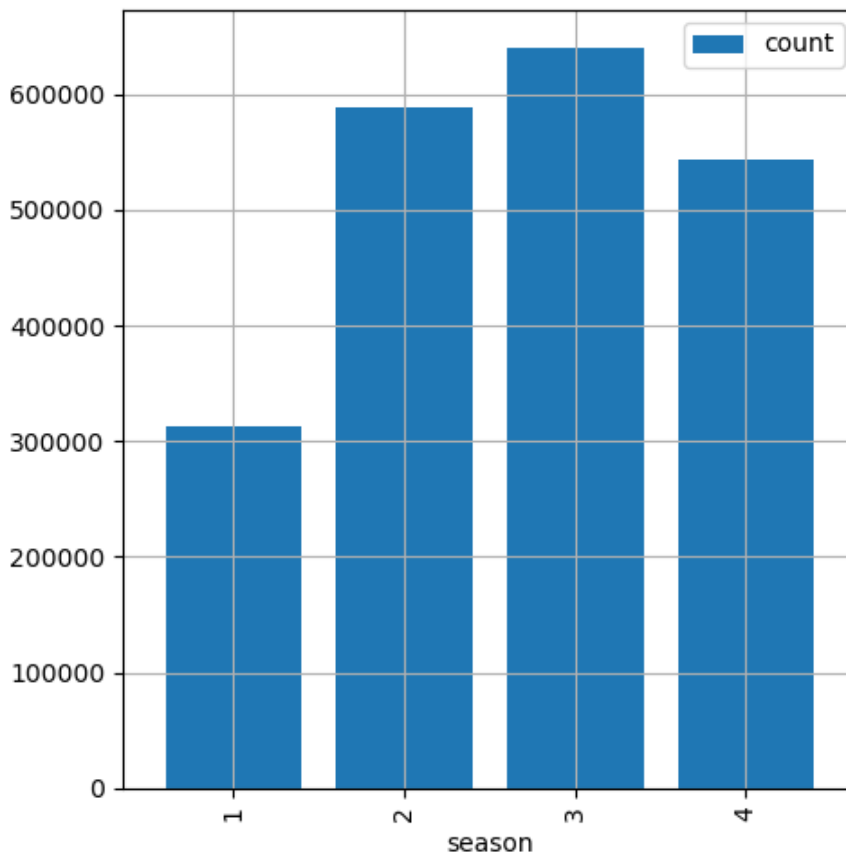
we can clearly see

- on working days peak is at 7-9 in morning and 4 to 8 in evening.
- on non-working days(Holidays, Weekends) peak is very less and occurs 12-4 in noon

```
In [22]: by_season = train.groupby(['season'])[['count']].agg(sum)
display(by_season)
```

```
count
season
1  312498
2  588282
3  640662
4  544034
```

```
In [23]: by_season.plot(kind='bar', figsize=(5,5), width=0.8);
plt.grid(True)
plt.tight_layout()
```

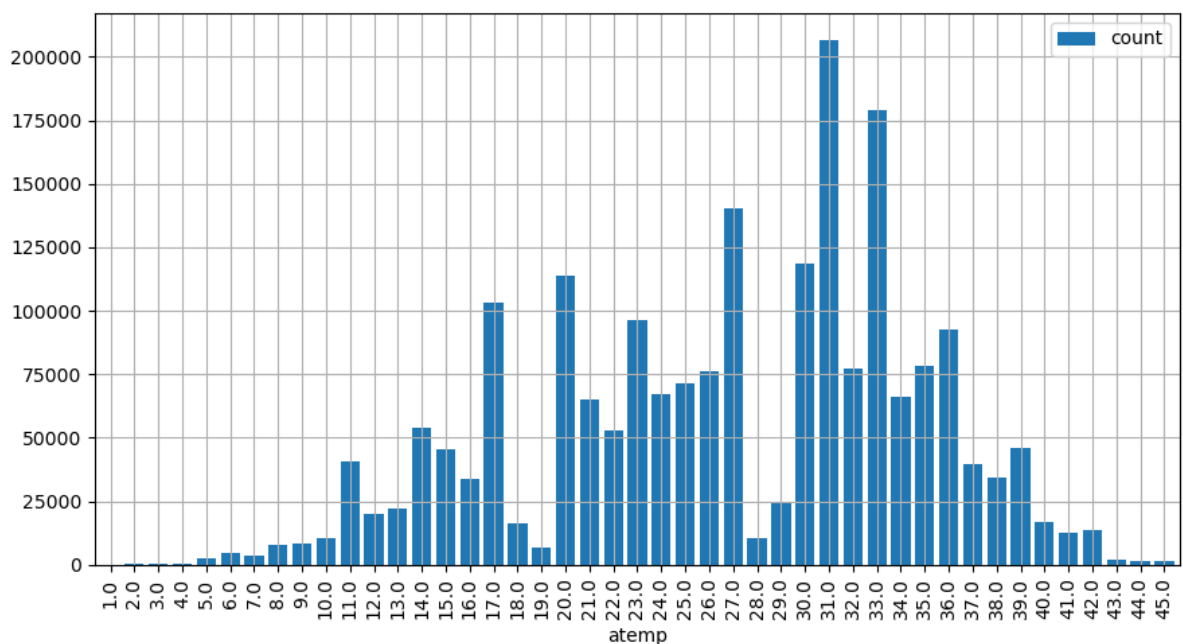


we can clearly see

- according to hypothesis season 1 must be rainy season and the bike demand is fairly less for that season.
- season 2, and season 3 will be spring and summer respectively
- season 4 must be winter

```
In [24]: x = train.round()
by_temp = x.groupby(['atemp'])['count'].agg(sum)
```

```
In [25]: by_temp.plot(kind='bar', figsize=(9,5), width=0.8);
plt.grid(True)
plt.tight_layout()
```

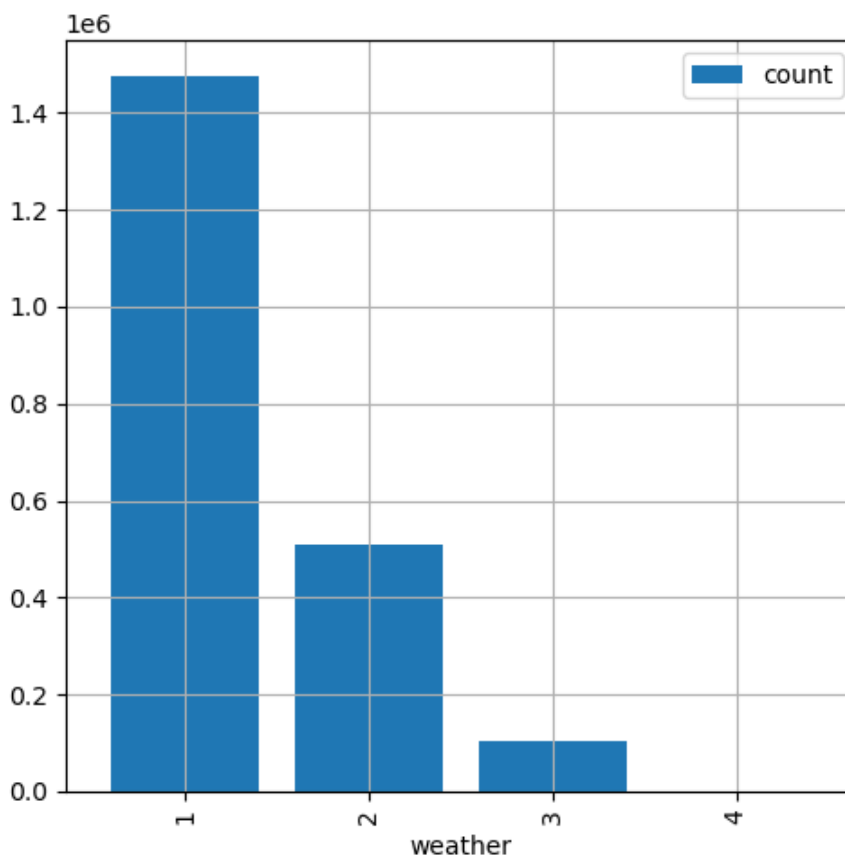


we see here the vehicle rented count increase with increase in temprature till 35 C and than start to decrease.

```
In [26]: by_weather = train.groupby(['weather'])[['count']].agg(sum)
display(by_season)
```

	count
season	
1	312498
2	588282
3	640662
4	544034

```
In [27]: by_weather.plot(kind='bar', figsize=(5,5), width=0.8);
plt.grid(True)
plt.tight_layout()
```



we can assume :-

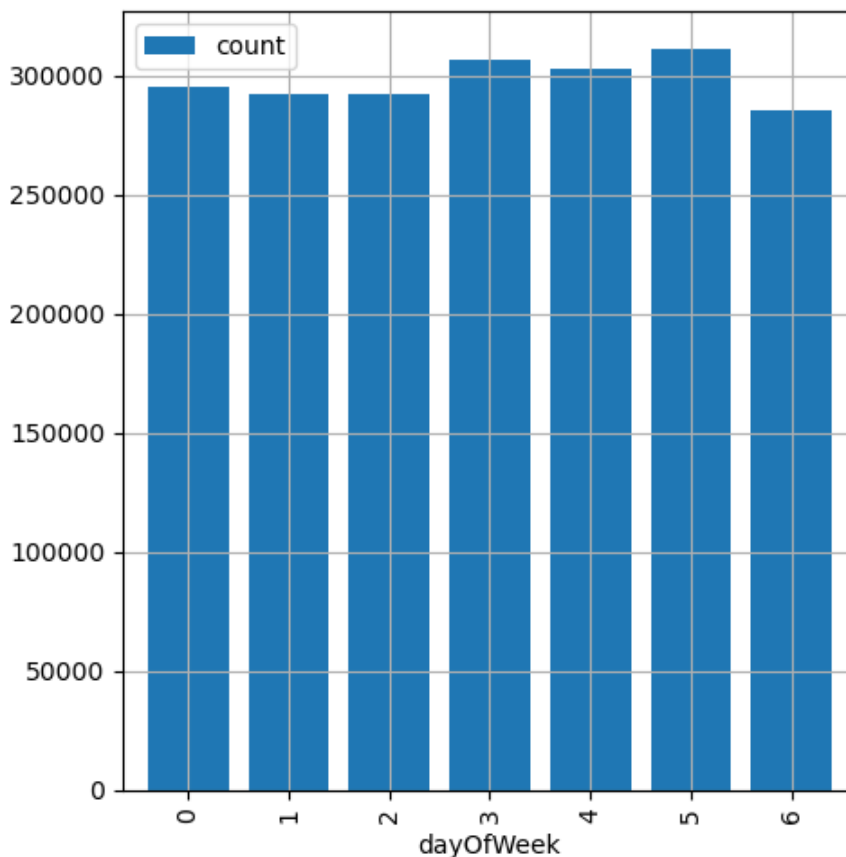
- weather 1 corresponds to: great weather
- weather 2 corresponds to: good weather
- weather 3 corresponds to: bad weather
- weather 4 corresponds to: worse weather

```
In [28]: by_dayOfWeek = train.groupby(['dayOfWeek'])[['count']].agg(sum)
display(by_dayOfWeek)
```

	count
dayOfWeek	
0	295296

	count
dayOfWeek	
1	291985
2	292226
3	306401
4	302504
5	311518
6	285546

```
In [29]: by_dayOfWeek.plot(kind='bar', figsize=(5,5), width=0.8);
plt.grid(True)
plt.tight_layout()
```



Clearly we could not establish any clear relationship between day of week and bike rent count

**Part 4: Apply L1 and L2 norm regularization over weight vectors, and find the best hyper-parameter settings for the mentioned problem using validation data and report the accuracy on test data for no regularization, L1 norm regularization and L2 norm regularization.**

```
In [30]: train.head(2)
```

```
Out[30]:
```

	year	season	month	day	dayOfWeek	hour	holiday	workingday	weather	temp	atemp	humidity	winds
0	2011	1	1	1	5	0	0	0	1	9.84	14.395	81	
1	2011	1	1	1	5	1	0	0	1	9.02	13.635	80	

```
In [31]: # Separating the input features and output parameter from the given training data
X_data1 = train[['year','hour','season','weather','temp',
                'atemp','humidity','windspeed']]
y_data1 = train[['count']]
day = train[['day']]
X_data1.head()
```

```
Out[31]:
```

	year	hour	season	weather	temp	atemp	humidity	windspeed
0	2011	0	1	1	9.84	14.395	81	0.0
1	2011	1	1	1	9.02	13.635	80	0.0
2	2011	2	1	1	9.02	13.635	80	0.0
3	2011	3	1	1	9.84	14.395	75	0.0
4	2011	4	1	1	9.84	14.395	75	0.0

```
In [32]: def isCategory(df):
NonCategorical = df.loc[:, 'temp': 'windspeed']
Categorical = df.loc[:, 'year': 'weather']
return NonCategorical, Categorical
```

```
In [33]: def normalize(df):
result = df.copy()
for feature_name in df.columns:
    max_value = df[feature_name].max()
    min_value = df[feature_name].min()
    result[feature_name] = (df[feature_name] - min_value) / (max_value - min_value)
return result
```

```
In [34]: Xdf = normalize(X_data1)
ydf = normalize(y_data1)
Xdf.head()
```

```
Out[34]:
```

	year	hour	season	weather	temp	atemp	humidity	windspeed
0	0.0	0.000000	0.0	0.0	0.224490	0.305068	0.81	0.0
1	0.0	0.043478	0.0	0.0	0.204082	0.288064	0.80	0.0
2	0.0	0.086957	0.0	0.0	0.204082	0.288064	0.80	0.0
3	0.0	0.130435	0.0	0.0	0.224490	0.305068	0.75	0.0
4	0.0	0.173913	0.0	0.0	0.224490	0.305068	0.75	0.0

```
In [35]: Xdf = pd.concat([Xdf, day], axis=1)
Xdf.head(2)
```

```
Out[35]:
```

	year	hour	season	weather	temp	atemp	humidity	windspeed	day
0	0.0	0.000000	0.0	0.0	0.224490	0.305068	0.81	0.0	1
1	0.0	0.043478	0.0	0.0	0.204082	0.288064	0.80	0.0	1

```
In [36]: ydf = pd.concat([ydf, day], axis=1)
ydf.head(2)
```

```
Out[36]:
```

	count	day
0	0.015369	1
1	0.039959	1

```
In [37]: def train_valid_split(data, cutoff_day=15):
train = data[data['day'] <= cutoff_day]
```

```
valid = data[data['day'] > cutoff_day]

return train, valid
```

```
In [38]: trainX, valX = train_valid_split(Xdf)
         trainY, valY = train_valid_split(ydf)
```

```
In [39]: trainX.pop('day') ; valX.pop('day') ; trainY.pop('day'); valY.pop('day')
```

```
Out[39]: 348      16
         349      16
         350      16
         351      16
         352      16
         ..
        10881     19
        10882     19
        10883     19
        10884     19
        10885     19
        Name: day, Length: 2286, dtype: int64
```

```
In [40]: print(trainX.shape)
         print(valX.shape)
         print(trainY.shape)
         print(valY.shape)
```

```
(8600, 8)
(2286, 8)
(8600, 1)
(2286, 1)
```

```
In [41]: # Test data set
         testX = test[['year', 'hour', 'season', 'weather',
                       'temp', 'atemp', 'humidity', 'windspeed']]
```

```
In [42]: testX = normalize (testX)
         testX.head()
```

```
Out[42]:
```

	year	hour	season	weather	temp	atemp	humidity	windspeed
0	0.0	0.000000	0.0	0.0	0.25	0.2273	0.47619	0.464346
1	0.0	0.043478	0.0	0.0	0.25	0.2727	0.47619	0.000000
2	0.0	0.086957	0.0	0.0	0.25	0.2727	0.47619	0.000000
3	0.0	0.130435	0.0	0.0	0.25	0.2576	0.47619	0.196458
4	0.0	0.173913	0.0	0.0	0.25	0.2576	0.47619	0.196458

```
In [43]: def prediction(w, X):
         yCap = np.exp(np.matmul(X, w))
         return yCap
```

```
In [44]: def gradient(X, y, yCap):
         gradient = np.divide(np.matmul(np.transpose(X), np.subtract(yCap, y)), len(y))
         return gradient

         def G_l1(X, y, yCap, w, reg_const):
             G_l1 = gradient(X, y, yCap)
             n = len(w)
             for i in range(len(w)):
                 if w[i,0] > 0:
                     G_l1[i,0] += (reg_const / n)
                 else:
                     G_l1[i,0] -= (reg_const / n)
             return G_l1

         def G_l2(X, y, yCap, w, reg_const):
             gradient_values_l2 = gradient(X, y, yCap) + reg_const * w
             return gradient_values_l2 # L2 norm regularized weight parameters
```



```
In [45]: def gradient_descent(X,y,alpha=0.1,iterations=50000,norm=0,reg_const=0):

    w = np.zeros((len(X[1,:]), 1)) #np.zeros((X.shape[1], 1))

    Loss = []

    for i in range(iterations):
        yCap = prediction(w, X)
        if norm == 1: # L1 Regularization
            gradient_value = G_l1(X, y, yCap, w, reg_const)
            L = float(np.sum(np.subtract(yCap,np.multiply(y,np.matmul(X, w))))+
                      ((reg_const/2)*np.sum(gradient_value)))
        elif norm == 2: # L2 Regularization
            gradient_value = G_l2(X, y, yCap, w, reg_const)
            L = float(np.sum(np.subtract(yCap,np.multiply(y,np.matmul(X, w))))+
                      ((reg_const/2)*np.matmul(gradient_value.T,gradient_value)))
        else:
            gradient_value = gradient(X, y, yCap)
            L = float(np.sum(np.subtract(yCap,np.multiply(y,np.matmul(X, w)))))
        Loss.append(L)

        w = np.subtract(w, alpha * gradient_value)

    return w, Loss
```

```
In [46]: trainX = trainX.to_numpy()
trainY = trainY.to_numpy()
valX = valX.to_numpy()
valY = valY.to_numpy()
testX = testX.to_numpy()
```

```
In [47]: w, Loss = gradient_descent(trainX, trainY, iterations = 50000, norm=0)
print('Estimated weight paramters with out regularization are:')
print(w)
```

Estimated weight paramters with out regularization are:

```
[[ 0.21404355]
 [ 0.24316723]
 [ 0.42599567]
 [ 0.22431102]
 [ 1.01124393]
 [-0.85963286]
 [-3.1850456 ]
 [-1.35657141]]
```

```
In [48]: valY_cap = prediction(w, valX)
valY_cap=pd.DataFrame(valY_cap, columns=['count'])
valY_cap.head(2)
```

```
Out[48]:
```

	count
0	0.163206
1	0.137080

```
In [49]: valY=pd.DataFrame(valY, columns=['count'])
valY.head(2)
```

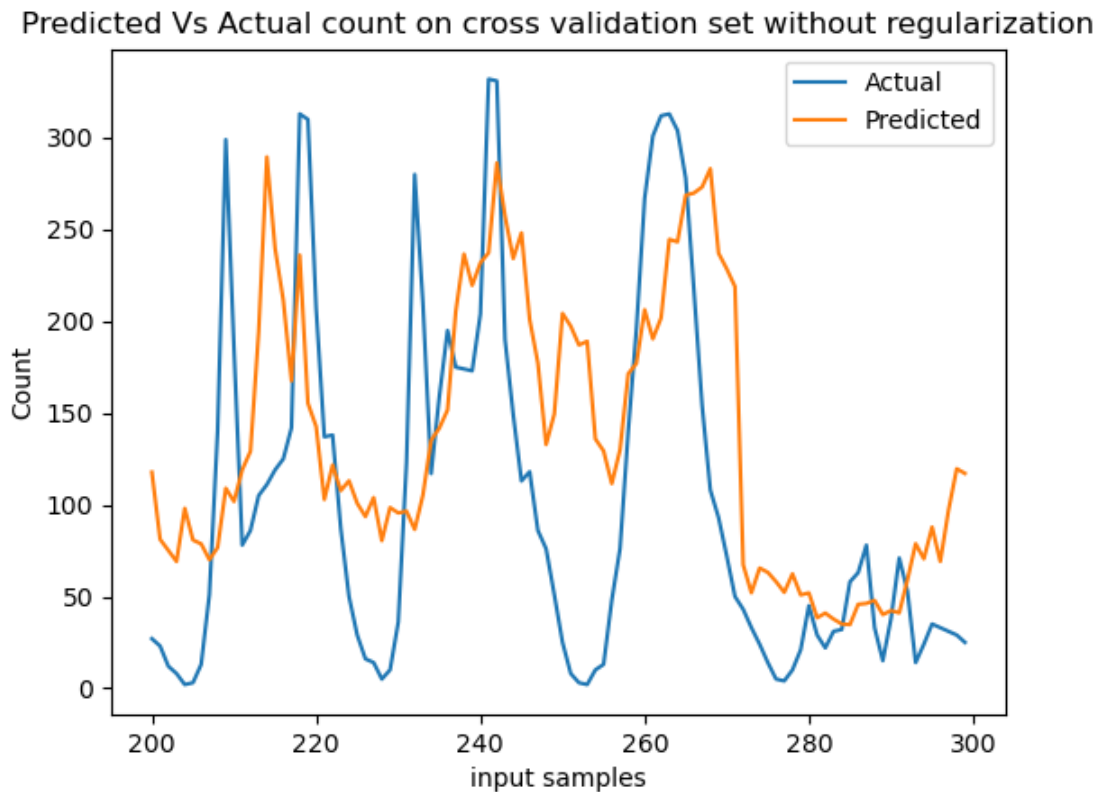
```
Out[49]:
```

	count
0	0.038934
1	0.022541

```
In [50]: df_ds = train.describe()
```

```
In [51]: valY_act = valY*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
valY_cap_act = valY_cap*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
plt.figure()
```

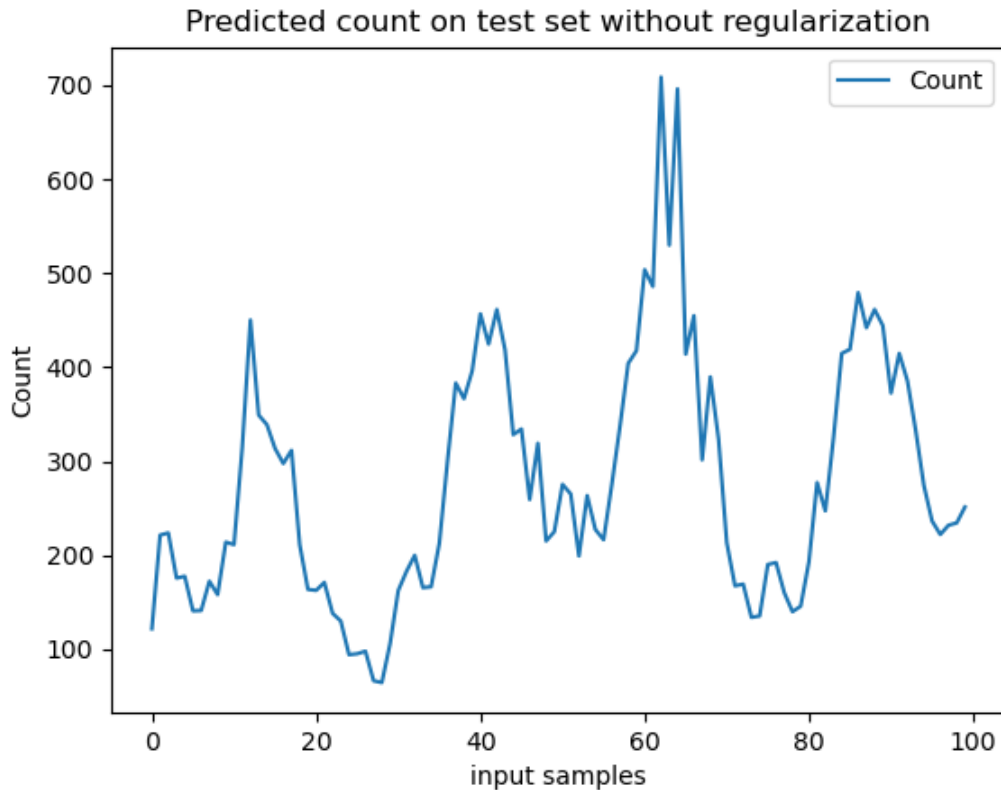
```
plt.plot(valY_act[200:300],label='Actual')
plt.plot(valY_cap_act[200:300],label='Predicted')
plt.title('Predicted Vs Actual count on cross validation set without regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()
```



```
In [52]: rmse_no_reg = np.sqrt(np.mean(np.square(valY_act - valY_cap_act)))
print('RMSE for the cross validation data is: {}'.format(rmse_no_reg))
```

RMSE for the cross validation data is: count 167.237206  
dtype: float64

```
In [53]: testY_cap = prediction(w, testX)
testY_cap_act = testY_cap*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
plt.figure()
plt.plot(testY_cap_act[:100],label='Count')
plt.title('Predicted count on test set without regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()
```



```
In [54]: y_test_cap = prediction(w, testX)
y_test_cap=pd.DataFrame(valY_cap, columns=['count'])
y_test_cap.head(3)
```

```
Out[54]:
```

	count
0	0.163206
1	0.137080
2	0.146613

```
In [55]: reg_const = [ 0.001,0.003,0.01,0.03,0.1]
l1_w = []
for i in reg_const:
    w, Loss = gradient_descent(trainX, trainY ,alpha=5, iterations = 10000, norm=1,reg_const=i)
    print('weight parameters with L1 norm for reg_const {0} are:'.format(i))
    print(w)
    l1_w.append(w)
    valY_cap_l1 = prediction(w, valX)
    rmse_l1_reg = np.sqrt(np.mean(np.square(valY - valY_cap_l1)))
    print('RMSE for reg_const {0} is: {1}'.format(i,rmse_l1_reg))
```

weight parameters with L1 norm for reg\_const 0.001 are:

```
[[ 0.13756084]
 [ 0.17326245]
 [ 0.35380469]
 [ 0.18502593]
 [ 1.17136302]
 [-1.14567907]
 [-3.20065862]
 [-1.34837289]]
```

RMSE for reg\_const 0.001 is: count 0.175745

dtype: float64

weight parameters with L1 norm for reg\_const 0.003 are:

```
[[ 0.12967499]
 [ 0.14056389]
 [ 0.34428037]
 [ 0.13566448]
 [ 0.14575541]
 [-0.10610593]
 [-3.21282666]]
```

```

[-1.27404936]]
RMSE for reg_const 0.003 is: count    0.176241
dtype: float64
weight parameters with L1 norm for reg_const 0.01 are:
[[ 1.13724981e-01]
 [ 8.16261962e-02]
 [ 3.08978550e-01]
 [ 7.20648982e-03]
 [-1.86425156e-03]
 [ 1.56404057e-03]
 [-3.08147557e+00]
 [-1.06623545e+00]]
RMSE for reg_const 0.01 is: count    0.17687
dtype: float64
weight parameters with L1 norm for reg_const 0.03 are:
[[-0.0178845 ]
 [-0.06247101]
 [ 0.07311545]
 [-0.00929195]
 [-0.00760109]
 [-0.21243629]
 [-2.96979858]
 [-0.73719493]]
RMSE for reg_const 0.03 is: count    0.194238
dtype: float64
weight parameters with L1 norm for reg_const 0.1 are:
[[ 1.59828800e-01]
 [-4.45385697e-03]
 [ 2.61974792e-01]
 [ 8.06420203e-04]
 [ 9.34871926e-02]
 [ 3.09488507e-02]
 [-2.31354227e+00]
 [ 1.46025150e-02]]
RMSE for reg_const 0.1 is: count    0.231962
dtype: float64

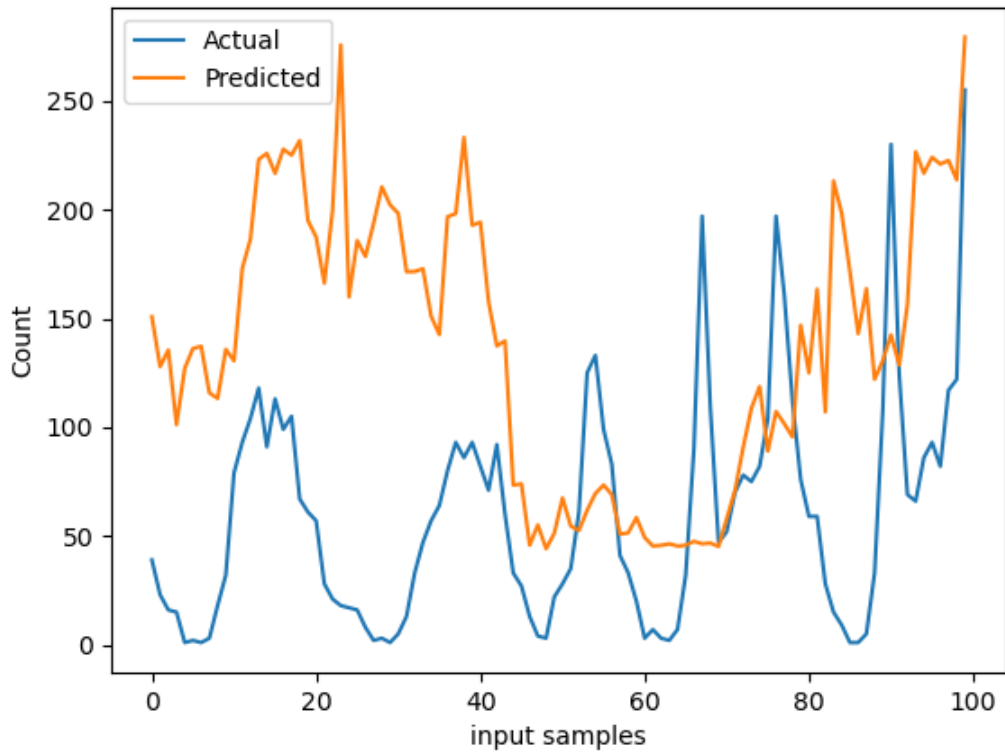
```

```

In [56]: valY_cap_l1 = prediction(l1_w[0], valX)
valY_act_l1 = valY*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
valY_cap_act_l1 = valY_cap_l1*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['co
plt.figure()
plt.plot(valY_act_l1[:100],label='Actual')
plt.plot(valY_cap_act_l1[:100],label='Predicted')
plt.title('Predicted Vs Actual count on cross validation set with L1 regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()

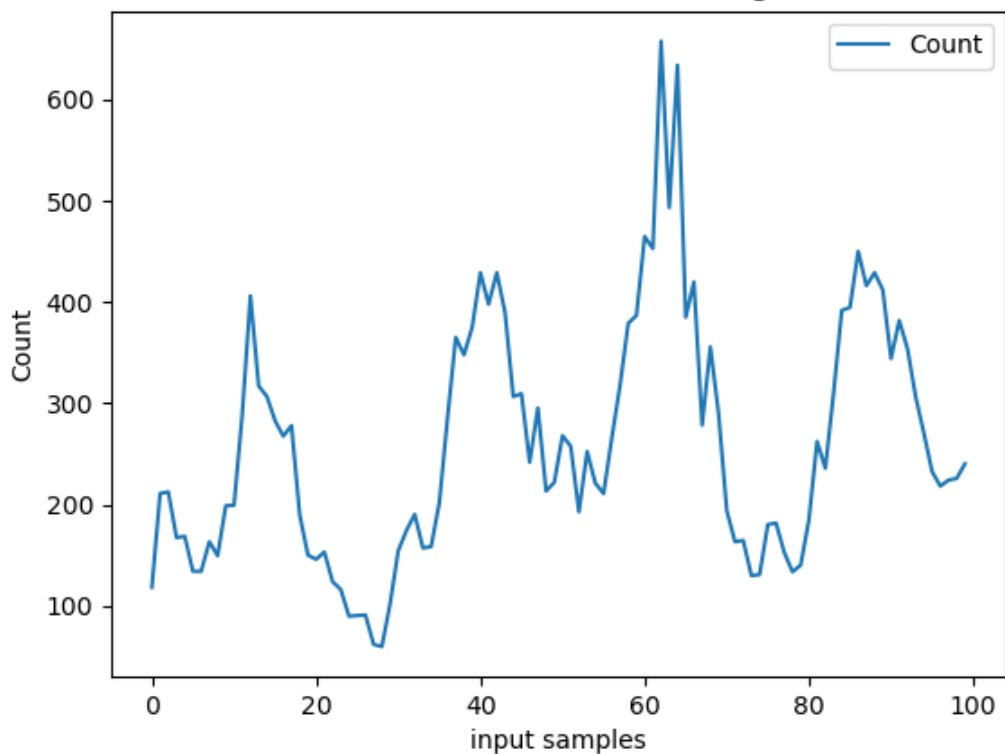
```

Predicted Vs Actual count on cross validation set with L1 regularization



```
In [68]: testY_cap_l1 = prediction(l1_w[0], testX)
testY_cap_act_l1 = testY_cap_l1*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
plt.figure()
plt.plot(testY_cap_act_l1[:100],label='Count')
plt.title('Predicted count on test set with L1 regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()
```

Predicted count on test set with L1 regularization



```
In [58]: reg_const = [0.0001,0.001,0.003,0.01]
l2_w = []
for i in reg_const:
    w, Loss = gradient_descent(trainX, trainY, iterations = 10000, norm=2, reg_const=i)
    print('weight parameters with L1 norm for reg_const {0} are:'.format(i))
    print(w)
    l2_w.append(w)
    plt.plot(Loss, label=i)
    valY_cap = prediction(w, valX)
    rmse_l2_reg = np.sqrt(np.mean(np.square(valY - valY_cap)))
    print('RMSE for reg_const {0} is {1}'.format(i, rmse_l2_reg))
```

weight parameters with L1 norm for reg\_const 0.0001 are:

```
[[ 0.20657005]
 [ 0.22163597]
 [ 0.41675576]
 [ 0.20410437]
 [ 0.33957369]
 [-0.1831982 ]
 [-3.1917347 ]
 [-1.31855227]]
```

RMSE for reg\_const 0.0001 is count 0.171402

dtype: float64

weight parameters with L1 norm for reg\_const 0.001 are:

```
[[ 0.17920981]
 [ 0.16270002]
 [ 0.34999295]
 [ 0.08596776]
 [ 0.22775037]
 [-0.16433602]
 [-2.96201196]
 [-1.15797432]]
```

RMSE for reg\_const 0.001 is count 0.171923

dtype: float64

weight parameters with L1 norm for reg\_const 0.003 are:

```
[[ 0.13277592]
 [ 0.07143345]
 [ 0.24156871]
 [-0.06464141]
 [ 0.08056864]
 [-0.17058134]
 [-2.58996833]
 [-0.93241396]]
```

RMSE for reg\_const 0.003 is count 0.174865

dtype: float64

weight parameters with L1 norm for reg\_const 0.01 are:

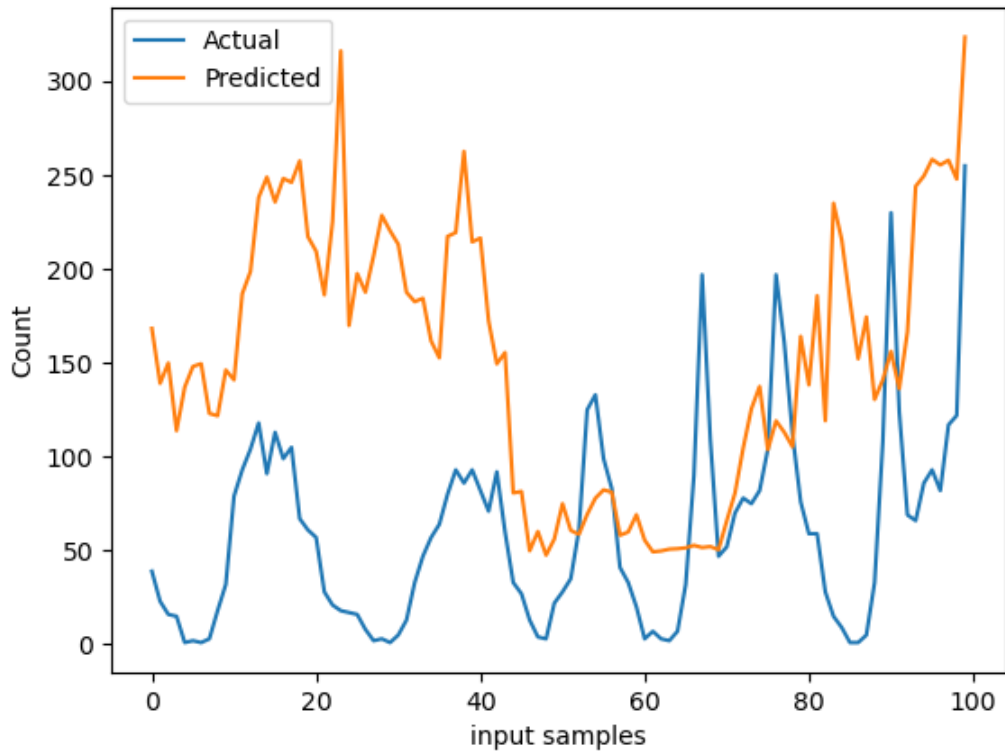
```
[[ 0.03567657]
 [-0.08657984]
 [ 0.04097541]
 [-0.21003999]
 [-0.11706019]
 [-0.24646499]
 [-1.91250394]
 [-0.61252694]]
```

RMSE for reg\_const 0.01 is count 0.186601

dtype: float64

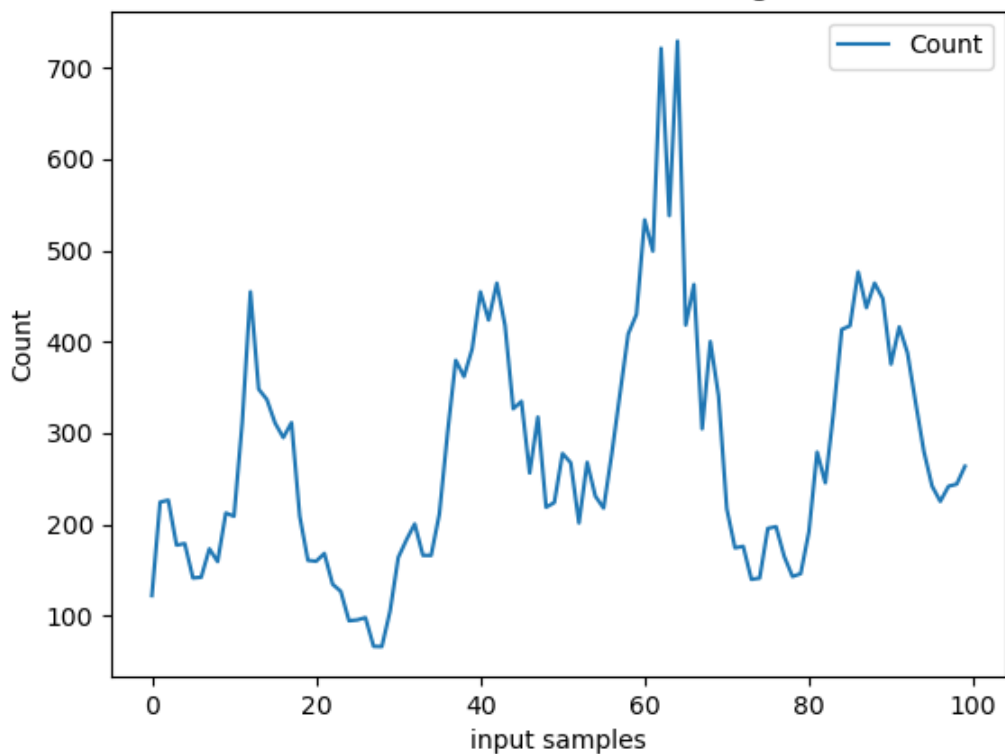
```
In [59]: valY_cap_l2 = prediction(l2_w[0], valX)
valY_act_l2 = valY*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
valY_cap_act_l2 = valY_cap_l2*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['co
plt.figure()
plt.plot(valY_act_l2[:100], label='Actual')
plt.plot(valY_cap_act_l2[:100], label='Predicted')
plt.title('Predicted Vs Actual count on cross validation set with L2 regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()
```

Predicted Vs Actual count on cross validation set with L2 regularization



```
In [60]: testY_cap = prediction(l2_w[0], testX)
testY_cap_act = testY_cap*(df_ds.iloc[7]['count']-df_ds.iloc[3]['count'])+df_ds.iloc[3]['count']
plt.figure()
plt.plot(testY_cap_act[:100],label='Count')
plt.title('Predicted count on test set with L2 regularization')
plt.xlabel('input samples')
plt.ylabel('Count')
plt.legend(loc='best')
plt.show()
```

Predicted count on test set with L2 regularization



## Part 5: most important features determining count of bikes rented

we can see this by plotting correlation of predicted values for rented bikes with other features.

```
In [61]: testX_pd = pd.DataFrame(testX, columns=['year', 'hour', 'season', 'weather', 'temp', 'atemp', 'humidi',
testX_pd.head(3)
```

```
Out[61]:
```

	year	hour	season	weather	temp	atemp	humidity	windspeed
0	0.0	0.000000	0.0	0.0	0.25	0.2273	0.47619	0.464346
1	0.0	0.043478	0.0	0.0	0.25	0.2727	0.47619	0.000000
2	0.0	0.086957	0.0	0.0	0.25	0.2727	0.47619	0.000000

```
In [62]: test_pred = pd.DataFrame(testY_cap_act, columns=['count'])
test_pred.head(3)
```

```
Out[62]:
```

	count
0	121.853314
1	224.220604
2	226.382033

```
In [63]: test_pred.shape
```

```
Out[63]: (6493, 1)
```

```
In [64]: testX_pd.shape
```

```
Out[64]: (6493, 8)
```

```
In [65]: test_pred = pd.concat([test_pred ,testX_pd], axis=1)
test_pred.head(2)
```

```
Out[65]:
```

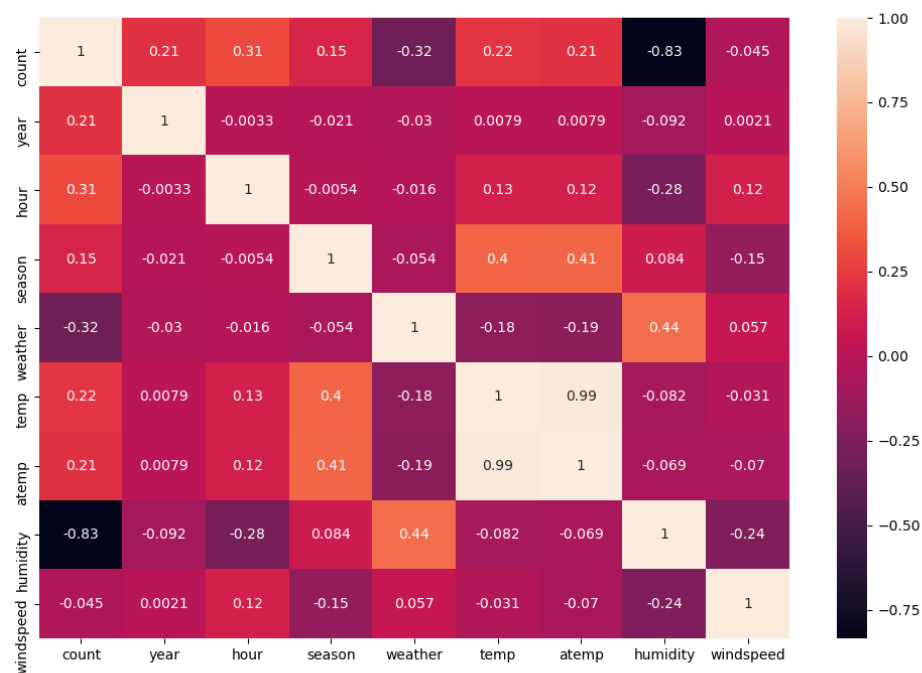
	count	year	hour	season	weather	temp	atemp	humidity	windspeed
0	121.853314	0.0	0.000000	0.0	0.0	0.25	0.2273	0.47619	0.464346
1	224.220604	0.0	0.043478	0.0	0.0	0.25	0.2727	0.47619	0.000000

```
In [66]: test_pred.isnull().sum()
```

```
Out[66]: count      0
year      0
hour      0
season    0
weather   0
temp      0
atemp     0
humidity  0
windspeed 0
dtype: int64
```

```
In [67]: corr = test_pred.corr()
plt.figure(figsize = (12, 8))
sns.heatmap(corr, annot = True)
plt.show()
```





here we clearly see that features that are most usefull in determining bike rented count are humidity, weather, hour

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