Probability Project

December 15, 2024

```
[2]: import math
     def log_likelihood(a, b, x):
         Compute the log-likelihood function for the Gamma distribution.
         Parameters:
             a (float): Shape parameter.
             b (float): Rate parameter.
             x (list): Observations.
         Returns:
             float: Log-likelihood value.
         n = len(x)
         S1 = sum(math.log(xi) for xi in x)
         S2 = sum(xi for xi in x)
         return n * a * math.log(b) - n * math.lgamma(a) + (a - 1) * S1 - b * S2
     # Example usage:
     x = [2.3, 1.7, 3.1, 0.8, 2.9] # Sample data
     a, b = 2.0, 3.0 # Parameters
     log_likelihood_value = log_likelihood(a, b, x)
     print("Log-Likelihood:", log_likelihood_value)
```

Log-Likelihood: -18.077370442152315

```
[3]: import numpy as np
from scipy.stats import gamma
import matplotlib.pyplot as plt

# Parameters
a_star = 2 # Shape parameter
b_star = 2 # Rate parameter
n = 20 # Sample size

# Generate a random sample
rng = np.random.default_rng(seed=42) # Set a seed for reproducibility
```

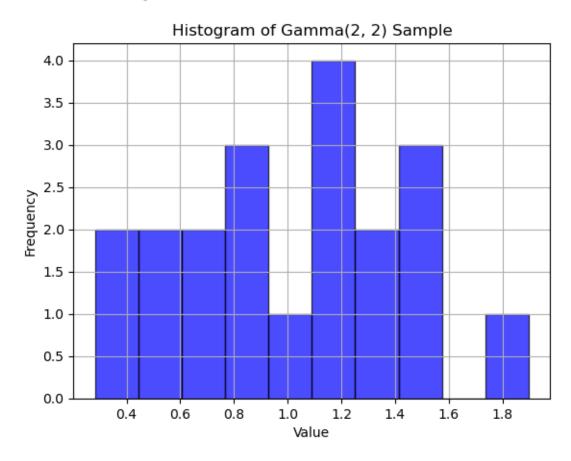
```
x = gamma.rvs(a=a_star, scale=1/b_star, size=n, random_state=rng)

# Print the sample
print("Generated Sample:", x)

# Plot the histogram of the sample
plt.hist(x, bins=10, alpha=0.7, color='blue', edgecolor='black')
plt.title(f"Histogram of Gamma({a_star}, {b_star}) Sample")
plt.xlabel("Value")
plt.ylabel("Frequency")
plt.grid(True)
plt.show()
```

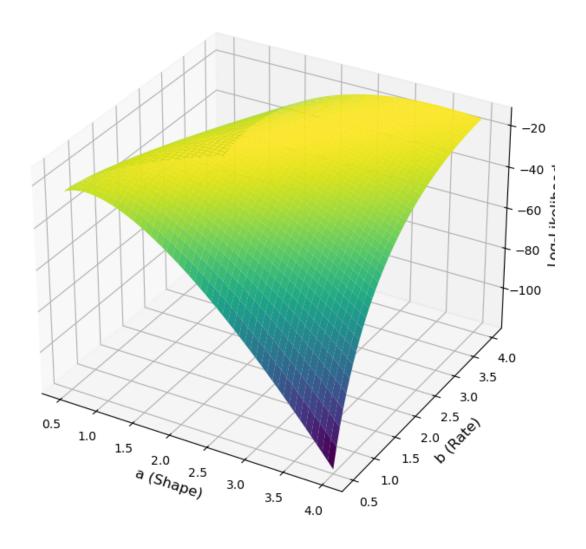
Generated Sample: [1.04590864 1.41767279 0.91860779 0.82253521 1.53962766 0.87668677

- $1.17300254\ 1.09474305\ 1.5387068\quad 0.71961025\ 1.89779193\ 0.58629914$
- 1.22632653 1.12915099 0.54448617 1.29753737 0.76192335 0.40638847
- 1.41106268 0.28290183]



```
[4]: import numpy as np
     import matplotlib.pyplot as plt
     from mpl_toolkits.mplot3d import Axes3D
     # Log-Likelihood function
     def log_likelihood(a, b, x):
        n = len(x)
         S1 = np.sum(np.log(x))
         S2 = np.sum(x)
         return n * a * np.log(b) - n * math.lgamma(a) + (a - 1) * S1 - b * S2
     # Parameters for the grid
     a_vals = np.linspace(0.5, 4, 50) # Range of 'a'
     b_vals = np.linspace(0.5, 4, 50) # Range of 'b'
     a_grid, b_grid = np.meshgrid(a_vals, b_vals)
     # Dataset (fixed sample)
     np.random.seed(42)
     x = np.random.gamma(shape=2, scale=1/2, size=20)
     # Compute Log-Likelihood for the grid
     z = np.array([[log_likelihood(a, b, x) for a, b in zip(a_row, b_row)]
                   for a_row, b_row in zip(a_grid, b_grid)])
     # 3D Visualization
     fig = plt.figure(figsize=(10, 8))
     ax = fig.add_subplot(111, projection='3d')
     ax.plot_surface(a_grid, b_grid, z, cmap='viridis', edgecolor='none')
     ax.set_title('Log-Likelihood Function', fontsize=16)
     ax.set_xlabel('a (Shape)', fontsize=12)
     ax.set_ylabel('b (Rate)', fontsize=12)
     ax.set_zlabel('Log-Likelihood', fontsize=12)
     plt.show()
```

Log-Likelihood Function

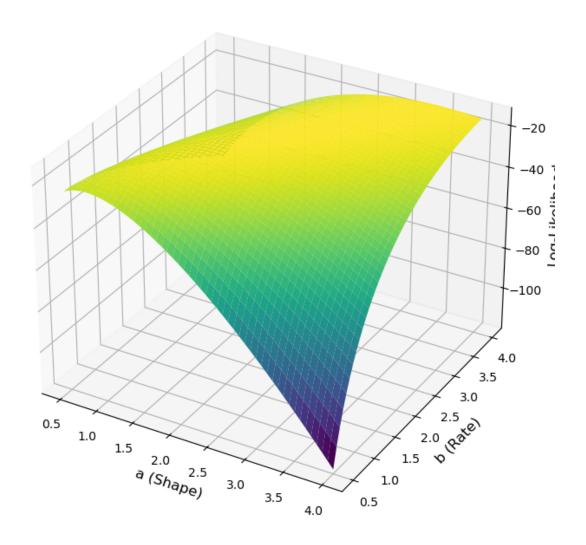


```
[5]: # Updated imports for flexibility
import numpy as np
import matplotlib.pyplot as plt
from mpl_toolkits.mplot3d import Axes3D
import math

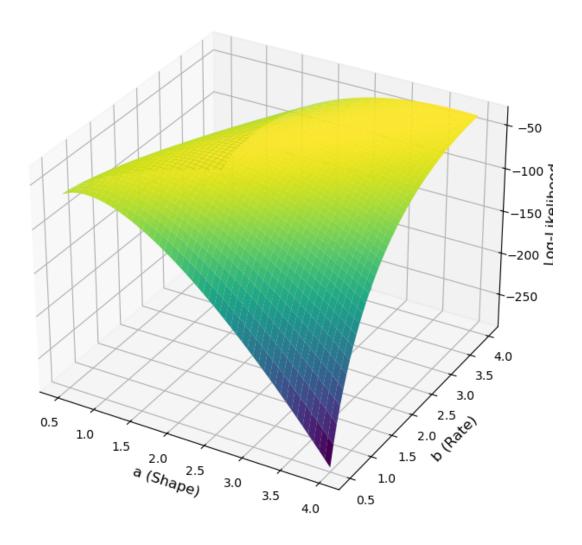
# Function to calculate log-likelihood
def log_likelihood(a, b, x):
    n = len(x)
    S1 = np.sum(np.log(x))
    S2 = np.sum(x)
    return n * a * np.log(b) - n * math.lgamma(a) + (a - 1) * S1 - b * S2
```

```
# Generate datasets with different parameters
def generate_gamma_data(a_star, b_star, n, seed=42):
    np.random.seed(seed)
    return np.random.gamma(shape=a_star, scale=1/b_star, size=n)
# Visualize log-likelihood for a dataset
def visualize_log_likelihood(x, a_range, b_range):
    a vals = np.linspace(*a range, 50)
    b_vals = np.linspace(*b_range, 50)
    a_grid, b_grid = np.meshgrid(a_vals, b_vals)
    z = np.array([[log_likelihood(a, b, x) for a, b in zip(a_row, b_row)]
                  for a_row, b_row in zip(a_grid, b_grid)])
    # 3D plot
    fig = plt.figure(figsize=(10, 8))
    ax = fig.add_subplot(111, projection='3d')
    ax.plot_surface(a_grid, b_grid, z, cmap='viridis', edgecolor='none')
    ax.set_title('Log-Likelihood Surface', fontsize=16)
    ax.set_xlabel('a (Shape)', fontsize=12)
    ax.set_ylabel('b (Rate)', fontsize=12)
    ax.set_zlabel('Log-Likelihood', fontsize=12)
    plt.show()
# Test cases with varying datasets
datasets = [
    (2, 2, 20), \quad \# \ a* = 2, \ b* = 2, \ n = 20
    (2, 2, 50), \quad \# \ a* = 2, \ b* = 2, \ n = 50
    (3, 1, 20), \# a* = 3, b* = 1, n = 20
    (1.5, 2.5, 100) # a* = 1.5, b* = 2.5, n = 100
]
# Generate and visualize log-likelihood for each dataset
for a_star, b_star, n in datasets:
    x = generate_gamma_data(a_star, b_star, n)
    print(f"Dataset: a*={a_star}, b*={b_star}, n={n}")
    visualize_log_likelihood(x, (0.5, 4), (0.5, 4))
```

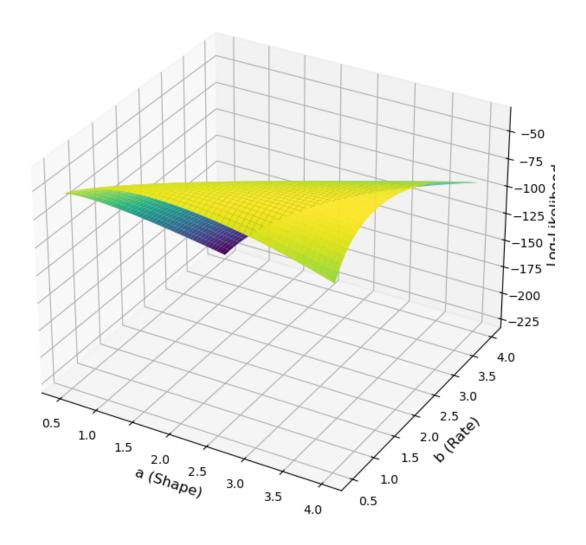
Dataset: a*=2, b*=2, n=20



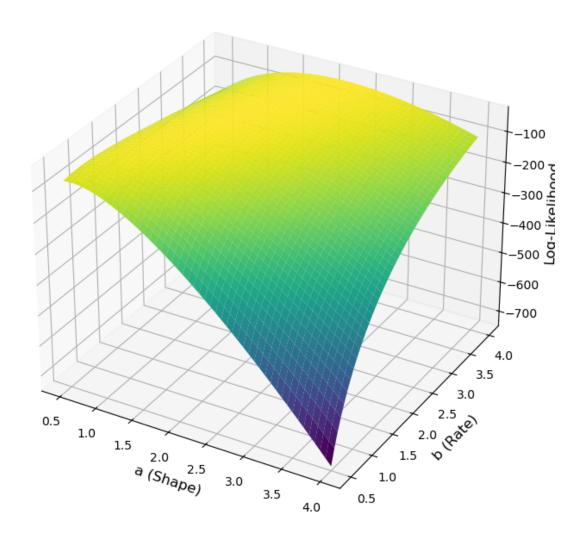
Dataset: a*=2, b*=2, n=50



Dataset: a*=3, b*=1, n=20



Dataset: a*=1.5, b*=2.5, n=100



```
[6]: from scipy.special import digamma, polygamma

def newton_raphson_gamma(x, tol=1e-6, max_iter=100):
    n = len(x)
    S1 = np.sum(np.log(x))
    S2 = np.sum(x)

# Initial guesses for a and b
    a = 1.0
    b = 1.0

for _ in range(max_iter):
```

```
# Gradient
        grad_a = n * np.log(b) - n * digamma(a) + S1
        grad_b = (n * a / b) - S2
        # Hessian
        hess_aa = -n * polygamma(1, a) # polygamma(1, a) is the trigamma_
 \hookrightarrow function
        hess_bb = -n * a / (b ** 2)
        hess_ab = n / b
        # Hessian matrix and its inverse
        hessian = np.array([[hess_aa, hess_ab], [hess_ab, hess_bb]])
        grad = np.array([grad_a, grad_b])
        try:
            step = np.linalg.solve(hessian, grad)
        except np.linalg.LinAlgError:
            print("Hessian is singular. Newton-Raphson failed.")
            break
        # Update parameters
        a -= step[0]
        b = step[1]
        # Convergence check
        if np.linalg.norm(step) < tol:</pre>
            break
    return a, b
# Example usage
x = generate_gamma_data(a_star=2, b_star=2, n=20)
a_est, b_est = newton_raphson_gamma(x)
print(f"Estimated parameters: a={a_est:.4f}, b={b_est:.4f}")
```

Estimated parameters: a=2.7056, b=3.0059

```
[7]: import matplotlib.pyplot as plt

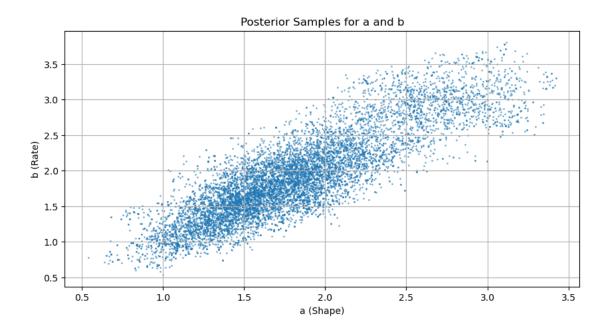
# Visualization of MLE vs MME performance
def visualize_comparison(a_star, b_star, n, num_simulations=100):
    mle_results = []
    mme_results = []

# Simulate datasets and compute estimates
for _ in range(num_simulations):
    x = generate_gamma_data(a_star, b_star, n)
    a_mle, b_mle = newton_raphson_gamma(x)
```

```
a_mme, b_mme = method_of_moments(x)
        mle_results.append((a_mle, b_mle))
        mme_results.append((a_mme, b_mme))
    mle_results = np.array(mle_results)
    mme_results = np.array(mme_results)
    # True values for reference
    true_params = np.array([a_star, b_star])
    # Bias and variance
    mle_bias = np.mean(mle_results, axis=0) - true_params
    mme_bias = np.mean(mme_results, axis=0) - true_params
    mle_var = np.var(mle_results, axis=0)
    mme_var = np.var(mme_results, axis=0)
    # Plotting
    labels = ['a', 'b']
    x_pos = np.arange(len(labels))
    # Bias plot
    plt.figure(figsize=(10, 5))
    plt.bar(x_pos - 0.2, mle_bias, width=0.4, label='MLE Bias', color='blue')
    plt.bar(x_pos + 0.2, mme_bias, width=0.4, label='MME_Bias', color='orange')
    plt.xticks(x_pos, labels)
    plt.title('Bias of MLE vs MME')
    plt.ylabel('Bias')
    plt.legend()
    plt.grid(True)
    plt.show()
    # Variance plot
    plt.figure(figsize=(10, 5))
    plt.bar(x_pos - 0.2, mle_var, width=0.4, label='MLE Variance', color='blue')
    plt.bar(x_pos + 0.2, mme_var, width=0.4, label='MME Variance', u
 ⇔color='orange')
    plt.xticks(x_pos, labels)
    plt.title('Variance of MLE vs MME')
    plt.ylabel('Variance')
    plt.legend()
    plt.grid(True)
    plt.show()
# Run and visualize
visualize_comparison(a_star=2, b_star=2, n=20, num_simulations=100)
```

```
[8]: import scipy.stats as stats
     # Log-posterior function
     def log_posterior(a, b, x, alpha_a, beta_a, alpha_b, beta_b):
        n = len(x)
         S1 = np.sum(np.log(x))
         S2 = np.sum(x)
         # Log-likelihood
         log_likelihood = (
             n * a * np.log(b) - n * math.lgamma(a) + (a - 1) * S1 - b * S2
         )
         # Log-priors
         log_prior_a = stats.gamma.logpdf(a, alpha_a, scale=1/beta_a)
         log_prior_b = stats.gamma.logpdf(b, alpha_b, scale=1/beta_b)
         return log_likelihood + log_prior_a + log_prior_b
     # MCMC using Metropolis-Hastings
     def metropolis hastings(x, alpha a, beta a, alpha b, beta b, num samples=10000):
         samples = []
         current_a, current_b = 1.0, 1.0 # Initial guesses
         for _ in range(num_samples):
             # Propose new values
             proposed_a = np.random.normal(current_a, 0.1)
             proposed_b = np.random.normal(current_b, 0.1)
             # Ensure positivity
             if proposed_a <= 0 or proposed_b <= 0:</pre>
```

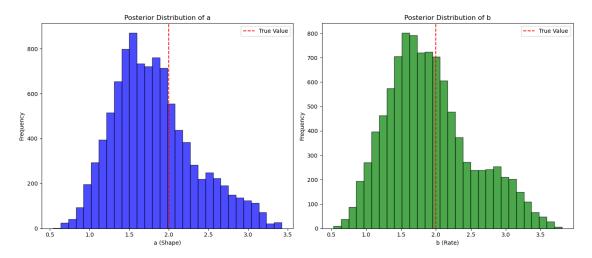
```
continue
        # Compute log-posterior for current and proposed values
        current_lp = log_posterior(current_a, current_b, x, alpha_a, beta_a,__
 ⇒alpha_b, beta_b)
        proposed_lp = log_posterior(proposed_a, proposed_b, x, alpha_a, beta_a,__
 →alpha_b, beta_b)
        # Metropolis acceptance criterion
        accept_prob = np.exp(proposed_lp - current_lp)
        if np.random.rand() < accept_prob:</pre>
            current_a, current_b = proposed_a, proposed_b
        samples.append((current_a, current_b))
    return np.array(samples)
# Example usage
x = generate_gamma_data(a_star=2, b_star=2, n=20)
alpha_a, beta_a = 2, 2 # Prior parameters for a
alpha_b, beta_b = 2, 2 # Prior parameters for b
samples = metropolis_hastings(x, alpha_a, beta_a, alpha_b, beta_b)
# Plot the posterior samples
import matplotlib.pyplot as plt
plt.figure(figsize=(10, 5))
plt.scatter(samples[:, 0], samples[:, 1], alpha=0.5, s=1)
plt.title("Posterior Samples for a and b")
plt.xlabel("a (Shape)")
plt.ylabel("b (Rate)")
plt.grid(True)
plt.show()
```



```
[9]: # Analyze posterior samples
     def analyze_posterior(samples, true_a, true_b):
         a_samples = samples[:, 0]
         b_samples = samples[:, 1]
         # Posterior statistics
         a_mean = np.mean(a_samples)
         b_mean = np.mean(b_samples)
         a_var = np.var(a_samples)
         b_var = np.var(b_samples)
         a_ci = np.percentile(a_samples, [2.5, 97.5])
         b_ci = np.percentile(b_samples, [2.5, 97.5])
         print("Posterior Analysis:")
         print(f"True a: {true_a}, True b: {true_b}")
         print(f"Posterior Mean (a): {a_mean:.4f}, Variance (a): {a_var:.4f}, 95% CI:
      → {a_ci}")
         print(f"Posterior Mean (b): {b_mean:.4f}, Variance (b): {b_var:.4f}, 95% CI:
      → {b_ci}")
         # Visualization of posterior distributions
         plt.figure(figsize=(14, 6))
         # Histogram for 'a'
         plt.subplot(1, 2, 1)
         plt.hist(a_samples, bins=30, color='blue', alpha=0.7, edgecolor='black')
```

```
plt.axvline(true_a, color='red', linestyle='--', label='True Value')
   plt.title("Posterior Distribution of a")
   plt.xlabel("a (Shape)")
   plt.ylabel("Frequency")
   plt.legend()
   # Histogram for 'b'
   plt.subplot(1, 2, 2)
   plt.hist(b_samples, bins=30, color='green', alpha=0.7, edgecolor='black')
   plt.axvline(true_b, color='red', linestyle='--', label='True Value')
   plt.title("Posterior Distribution of b")
   plt.xlabel("b (Rate)")
   plt.ylabel("Frequency")
   plt.legend()
   plt.tight_layout()
   plt.show()
# Analyze the posterior samples
true_a, true_b = 2, 2
analyze_posterior(samples, true_a, true_b)
```

```
Posterior Analysis:
True a: 2, True b: 2
Posterior Mean (a): 1.8130, Variance (a): 0.2719, 95% CI: [0.97499517 3.02102337]
Posterior Mean (b): 1.9161, Variance (b): 0.3733, 95% CI: [0.9380117 3.26418381]
```



```
[13]: def compare_all_estimators(samples, x, true_a, true_b):
    # Bayesian Estimates
```

```
a_bayes = np.mean(samples[:, 0])
b_bayes = np.mean(samples[:, 1])

# MLE Estimates
a_mle, b_mle = newton_raphson_gamma(x)

# MME Estimates
a_mme, b_mme = method_of_moments(x)

print("\nComparison of Estimates:")
print(f"True Values: a*={true_a}, b*={true_b}")
print(f"Bayesian Estimates: a={a_bayes:.4f}, b={b_bayes:.4f}")
print(f"MLE Estimates: a={a_mle:.4f}, b={b_mle:.4f}")
print(f"MME Estimates: a={a_mme:.4f}, b={b_mme:.4f}")

# Compare all estimators
compare_all_estimators(samples, x, true_a, true_b)
```

```
[14]: # Risk evaluation: Mean Squared Error (MSE)
def evaluate_risks(samples, x, true_a, true_b):
    # True values
    true_params = np.array([true_a, true_b])

# Bayesian Estimates (posterior mean)
bayes_mean = np.mean(samples, axis=0)
bayes_risk = np.mean((bayes_mean - true_params) ** 2)

# MLE Estimates
a_mle, b_mle = newton_raphson_gamma(x)
mle_risk = np.mean((np.array([a_mle, b_mle]) - true_params) ** 2)
```

```
# MME Estimates
a_mme, b_mme = method_of_moments(x)
mme_risk = np.mean((np.array([a_mme, b_mme]) - true_params) ** 2)

print("Risk Evaluation (MSE):")
print(f"Bayesian Risk: {bayes_risk:.4f}")
print(f"MLE Risk: {mle_risk:.4f}")
print(f"MME Risk: {mme_risk:.4f}")

# Run risk evaluation
evaluate_risks(samples, x, true_a, true_b)
```

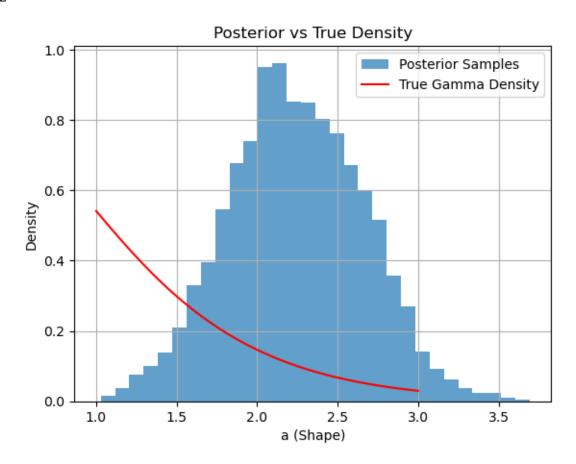
```
bayes_mean = np.mean(samples, axis=0)
            bayes_mse.append(np.mean((bayes_mean - np.array([true_a, true_b]))_u
 →** 2))
            # MLE Estimates
            a mle, b mle = newton raphson gamma(x)
            mle_mse.append(np.mean((np.array([a_mle, b_mle]) - np.
 →array([true_a, true_b])) ** 2))
            # MME Estimates
            a_mme, b_mme = method_of_moments(x)
            mme mse.append(np.mean((np.array([a mme, b mme]) - np.
 →array([true_a, true_b])) ** 2))
        # Average risks for this sample size
       bayes_risks.append(np.mean(bayes_mse))
        mle_risks.append(np.mean(mle_mse))
        mme_risks.append(np.mean(mme_mse))
    # Plot results
   plt.figure(figsize=(10, 6))
   plt.plot(sample_sizes, bayes_risks, label='Bayesian Risk', marker='o')
   plt.plot(sample sizes, mle risks, label='MLE Risk', marker='o')
   plt.plot(sample_sizes, mme_risks, label='MME Risk', marker='o')
   plt.title('Risk vs Sample Size')
   plt.xlabel('Sample Size (n)')
   plt.ylabel('Mean Squared Error (Risk)')
   plt.legend()
   plt.grid(True)
   plt.show()
# Run sensitivity analysis
sensitivity analysis(true_a=2, true_b=2, sample_sizes=[10, 20, 50, 100, 200])
```

```
---> 25 a_mme, b_mme = method_of_moments(x)
26 mme_mse.append(np.mean((np.array([a_mme, b_mme]) - np.array([true_a_true_b])) ** 2))
28 # Average risks for this sample size

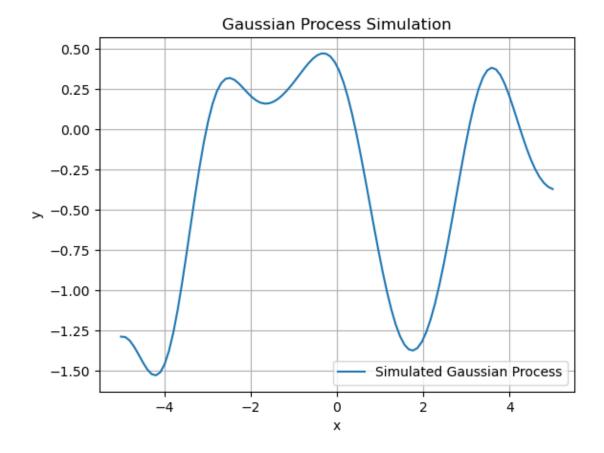
NameError: name 'method_of_moments' is not defined
```

```
[16]: import numpy as np
      import matplotlib.pyplot as plt
      from scipy.stats import gamma
      # Log posterior function for Gamma
      def log_posterior_gamma(a, b, x, alpha_a, beta_a, alpha_b, beta_b):
          n = len(x)
          S1 = np.sum(np.log(x))
          S2 = np.sum(x)
          log_likelihood = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - 10
       →b * S2
          log_prior_a = gamma.logpdf(a, alpha_a, scale=1/beta_a)
          log_prior_b = gamma.logpdf(b, alpha_b, scale=1/beta_b)
          return log_likelihood + log_prior_a + log_prior_b
      # MCMC for Gamma posterior
      def mcmc_gamma(x, alpha_a, beta_a, alpha_b, beta_b, num_samples=10000):
          samples = []
          current_a, current_b = 1.0, 1.0
          for _ in range(num_samples):
              proposed_a = np.random.normal(current_a, 0.1)
              proposed_b = np.random.normal(current_b, 0.1)
              if proposed_a <= 0 or proposed_b <= 0:</pre>
                  continue
              current lp = log posterior_gamma(current a, current b, x, alpha a, u
       ⇔beta_a, alpha_b, beta_b)
              proposed_lp = log_posterior_gamma(proposed_a, proposed_b, x, alpha_a,_
       ⇔beta_a, alpha_b, beta_b)
              accept_prob = np.exp(proposed_lp - current_lp)
              if np.random.rand() < accept_prob:</pre>
                  current_a, current_b = proposed_a, proposed_b
              samples.append((current_a, current_b))
          return np.array(samples)
      # Gamma posterior analysis
      x = np.random.gamma(shape=2, scale=0.5, size=50) # Simulated data
      alpha_a, beta_a = 2, 2 # Priors for a
      alpha_b, beta_b = 2, 2 # Priors for b
      samples = mcmc_gamma(x, alpha_a, beta_a, alpha_b, beta_b)
```

C:\Users\sshubhankar\AppData\Local\Temp\ipykernel_21920\1690873188.py:10:
DeprecationWarning: `np.math` is a deprecated alias for the standard library
`math` module (Deprecated Numpy 1.25). Replace usages of `np.math` with `math`
 log_likelihood = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - b
* S2



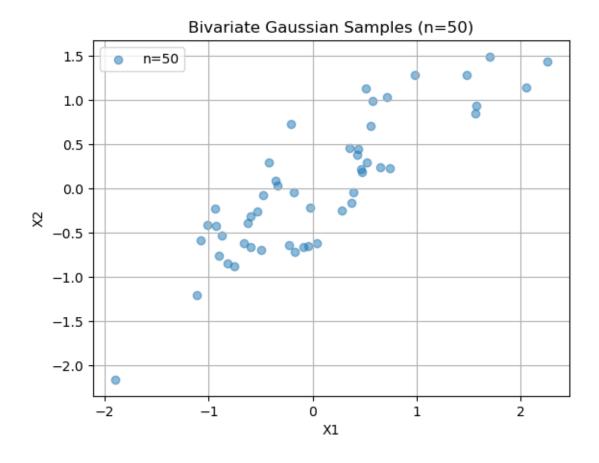
```
[17]: import numpy as np
      import matplotlib.pyplot as plt
      # Squared Exponential Kernel
      def squared_exponential_kernel(x1, x2, sigma_f, 1):
          x1 = np.atleast_2d(x1).T
          x2 = np.atleast_2d(x2).T
          sqdist = np.sum(x1**2, 1).reshape(-1, 1) + np.sum(x2**2, 1) - 2 * np.
       \rightarrowdot(x1, x2.T)
          return sigma_f**2 * np.exp(-0.5 / 1**2 * sqdist)
      # Gaussian Process Simulation
      def simulate_gaussian_process(n_points, sigma_f, 1):
          x = np.linspace(-5, 5, n_points)
          cov_matrix = squared_exponential_kernel(x, x, sigma_f, 1)
          y = np.random.multivariate_normal(np.zeros(n_points), cov_matrix)
          return x, y
      # Simulate and Plot
      np.random.seed(42)
      x, y = simulate_gaussian_process(100, sigma_f=1.0, l=1.0)
      plt.plot(x, y, label='Simulated Gaussian Process')
      plt.title('Gaussian Process Simulation')
      plt.xlabel('x')
      plt.ylabel('y')
      plt.grid()
      plt.legend()
      plt.show()
```

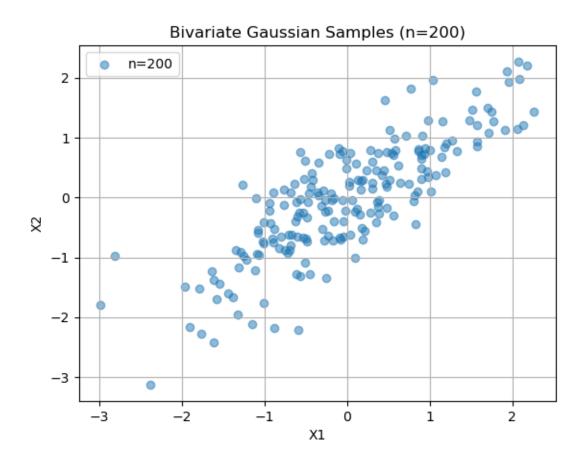


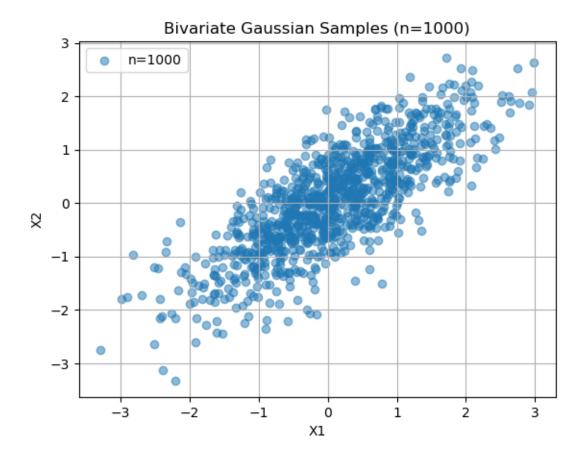
```
[18]: from scipy.optimize import minimize
      # Negative Log Marginal Likelihood
      def negative_log_marginal_likelihood(params, x, y):
          sigma_f, l, noise = params
          K = squared_exponential_kernel(x, x, sigma_f, 1) + noise**2 * np.eye(len(x))
          K_inv = np.linalg.inv(K)
          return 0.5 * y.T @ K_inv @ y + 0.5 * np.log(np.linalg.det(K)) + 0.5 *__
       \rightarrowlen(x) * np.log(2 * np.pi)
      # Optimize Hyperparameters
      def optimize_hyperparameters(x, y):
          initial_params = [1.0, 1.0, 0.1] # Initial guesses for sigma_f, l, noise
          bounds = [(1e-3, None), (1e-3, None), (1e-3, None)]
          result = minimize(negative_log_marginal_likelihood, initial_params,__
       ⇒args=(x, y), bounds=bounds)
          return result.x
      # Simulate Data and Optimize
      x, y = simulate_gaussian_process(50, sigma_f=1.0, l=2.0)
```

Optimized Parameters: Sigma_f: 0.8515, Length Scale: 1.9663, Noise: 0.0010

```
[19]: import numpy as np
      import matplotlib.pyplot as plt
      from scipy.stats import multivariate_normal
      # Generate bivariate Gaussian samples
      def generate_bivariate_gaussian(mean, cov, n_samples, seed=None):
          rng = np.random.default_rng(seed)
          return rng.multivariate_normal(mean, cov, n_samples)
      # Parameters
      mean = [0, 0]
      cov = [[1, 0.8], [0.8, 1]] # High correlation
      sample_sizes = [50, 200, 1000] # Different sample sizes
      # Generate and visualize samples
      for n in sample_sizes:
          samples = generate_bivariate_gaussian(mean, cov, n, seed=42)
          plt.scatter(samples[:, 0], samples[:, 1], alpha=0.5, label=f'n={n}')
          plt.title(f'Bivariate Gaussian Samples (n={n})')
          plt.xlabel('X1')
          plt.ylabel('X2')
          plt.legend()
          plt.grid(True)
          plt.show()
```







```
[21]: import numpy as np

def analytical_mutual_information(cov_matrix):
    rho = cov_matrix[0, 1] / np.sqrt(cov_matrix[0, 0] * cov_matrix[1, 1])
    return -0.5 * np.log(1 - rho**2)

# Covariance matrix and mutual information
cov = np.array([[1, 0.8], [0.8, 1]]) # High correlation
mi_analytical = analytical_mutual_information(cov)
print(f"Analytical Mutual Information: {mi_analytical:.4f}")
```

Analytical Mutual Information: 0.5108

```
⇔grid_size),
          )
          points = np.vstack([x.ravel(), y.ravel()])
          p_xy = kde_joint(points).reshape(grid_size, grid_size)
          # Marginal KDEs
          kde_x = gaussian_kde(samples[:, 0])
          kde_y = gaussian_kde(samples[:, 1])
          p_x = kde_x(x[:, 0])
          p_y = kde_y(y[0, :])
          # Compute mutual information
          mi_estimate = 0
          dx = (x[0, 1] - x[0, 0])
          dy = (y[1, 0] - y[0, 0])
          for i in range(grid_size):
              for j in range(grid_size):
                  if p_xy[i, j] > 0: # Avoid log(0)
                      mi_estimate += p_xy[i, j] * np.log(p_xy[i, j] / (p_x[i] *_u)
       \rightarrow p_y[j])) * dx * dy
          return mi_estimate
      # Generate samples and compute plug-in mutual information
      samples = generate_bivariate_gaussian(mean, cov, 1000, seed=42)
      mi_plugin = kde_mutual_information(samples)
      print(f"Plug-in Mutual Information: {mi plugin:.4f}")
     Plug-in Mutual Information: 26.5129
[24]: print(f"Analytical Mutual Information: {mi_analytical:.4f}")
      print(f"Plug-in Mutual Information: {mi_plugin:.4f}")
     Analytical Mutual Information: 0.5108
     Plug-in Mutual Information: 26.5129
 [1]: import numpy as np
      def mme_gamma(x):
         n = len(x)
          mean_x = np.mean(x)
          var_x = np.var(x, ddof=0) # Population variance
          a_mme = mean_x**2 / var_x
          b_mme = mean_x / var_x
          return a_mme, b_mme
      x = np.random.gamma(shape=2, scale=0.5, size=20)
      a_mme, b_mme = mme_gamma(x)
```

np.linspace(samples[:, 1].min() - 1, samples[:, 1].max() + 1, \Box

```
print(f"MME Estimates: a={a_mme:.4f}, b={b_mme:.4f}")
```

MME Estimates: a=3.3195, b=5.5485

```
[5]: from scipy.special import digamma, polygamma
     def newton raphson mle(x, tol=1e-6, max iter=100, init="mme"):
         n = len(x)
         S1 = np.sum(np.log(x))
         S2 = np.sum(x)
         if init == "mme":
             a, b = mme_gamma(x)
         else:
             a, b = 1.0, 1.0 # Default initialization
         for _ in range(max_iter):
             if a <= 0 or b <= 0: # Handle invalid parameters</pre>
                 print("Reinitializing due to invalid parameters.")
                 a, b = 1.0, 1.0
                 continue
             grad_a = n * np.log(b) - n * digamma(a) + S1
             grad_b = (n * a / b) - S2
             hess_aa = -n * polygamma(1, a)
             hess_bb = -n * a / (b**2)
             hess_ab = n / b
             hessian = np.array([[hess_aa, hess_ab], [hess_ab, hess_bb]])
             grad = np.array([grad_a, grad_b])
             try:
                 step = np.linalg.solve(hessian, grad)
             except np.linalg.LinAlgError:
                 print("Hessian is singular. Skipping this step.")
                 break
             a -= step[0]
             b = step[1]
             # Ensure parameters remain positive
             if a <= 0 or b <= 0:</pre>
                 print("Invalid parameters encountered, resetting.")
                 a, b = 1.0, 1.0
                 continue
```

```
if np.linalg.norm(step) < tol:</pre>
            break
    return a, b
# Test the updated function
x = np.random.gamma(shape=2, scale=0.5, size=20)
a_mle, b_mle = newton_raphson_mle(x, init="mme")
print(f"MLE Estimates with Improved Error Handling: a={a mle:.4f}, b={b mle:.

4f}")
111
def\ newton\ raphson\ mle(x,\ tol=1e-6,\ max\ iter=100,\ init="mme"):
    n = len(x)
    S1 = np.sum(np.log(x))
    S2 = np.sum(x)
    if init == "mme":
        a, b = mme\_gamma(x)
    else:
        a, b = 1.0, 1.0
    for _ in range(max_iter):
        grad_a = n * np.log(b) - n * digamma(a) + S1
        qrad_b = (n * a / b) - S2
        hess_aa = -n * polygamma(1, a)
        hess\_bb = -n * a / (b**2)
        hess_ab = n / b
        hessian = np.array([[hess_aa, hess_ab], [hess_ab, hess_bb]])
        grad = np.array([grad_a, grad_b])
        try:
            step = np.linalq.solve(hessian, grad)
        except np.linalg.LinAlgError:
            break
        a = step[0]
        b -= step[1]
        if np.linalg.norm(step) < tol:</pre>
            break
    return a, b
a_mle, b_mle = newton_raphson_mle(x, init="mme")
print(f"MLE\ Estimates\ with\ MME\ Initialization:\ a=\{a\_mle:.4f\},\ b=\{b\_mle:.4f\}")
111
```

MLE Estimates with Improved Error Handling: a=1.7310, b=2.3040

```
[5]: '\ndef newton_raphson_mle(x, tol=1e-6, max_iter=100, init="mme"):\n n = len(x)\n S1 = np.sum(np.log(x))\n S2 = np.sum(x)\n if init == "mme":\n
```

```
a, b = mme_gamma(x) \n
                             else:\n
                                             a, b = 1.0, 1.0\n
     range(max_iter):\n
                               grad_a = n * np.log(b) - n * digamma(a) + S1\n
     grad_b = (n * a / b) - S2\n
                                        hess_aa = -n * polygamma(1, a) \n
    hess_bb = -n * a / (b**2) \n
                                        hess_ab = n / b \ n
                                                                  hessian =
    np.array([[hess_aa, hess_ab], [hess_ab, hess_bb]])\n
                                                                  grad =
    np.array([grad_a, grad_b])\n
                                         try:\n
                                                            step =
    np.linalg.solve(hessian, grad)\n
                                             except np.linalg.LinAlgError:\n
    break\n
                    a -= step[0] \n
                                          b = step[1] \n
                                                                 if
    np.linalg.norm(step) < tol:\n</pre>
                                              break\n
                                                          return a, b\n\na mle, b mle
     = newton_raphson_mle(x, init="mme")\nprint(f"MLE Estimates with MME
     Initialization: a=\{a mle:.4f\}, b=\{b mle:.4f\}"\}\n'
[6]: def compare_estimators(n_simulations, sample_size, a_star, b_star):
         mle risks = []
         mme_risks = []
         for _ in range(n_simulations):
             # Generate dataset
             x = np.random.gamma(shape=a_star, scale=1/b_star, size=sample_size)
             # MME estimates
             a mme, b mme = mme gamma(x)
             # MLE estimates
             a_mle, b_mle = newton_raphson_mle(x, init="mme")
             # Squared errors
             mle_error = (a_mle - a_star)**2 + (b_mle - b_star)**2
             mme_error = (a_mme - a_star)**2 + (b_mme - b_star)**2
             mle risks.append(mle error)
             mme_risks.append(mme_error)
         # Calculate average risks
         mle risk = np.mean(mle risks)
         mme_risk = np.mean(mme_risks)
         print(f"MLE Risk: {mle_risk:.4f}")
         print(f"MME Risk: {mme_risk:.4f}")
     # Simulate and compare
     compare_estimators(n_simulations=1000, sample_size=20, a_star=2, b_star=2)
```

MLE Risk: 1.4856 MME Risk: 2.2218

```
[7]: #Implement a Markov Chain Monte Carlo (MCMC) algorithm to sample from the
      \hookrightarrowposterior distribution of the Gamma parameters a and b, and compare the
      ⇔estimated posterior density to the true Gamma posterior.
     import numpy as np
     from scipy.stats import gamma
     # Log posterior
     def log_posterior(a, b, x, alpha_a, beta_a, alpha_b, beta_b):
         n = len(x)
         S1 = np.sum(np.log(x))
         S2 = np.sum(x)
         # Log-likelihood
         log_likelihood = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - 10
      →b * S2
         # Log priors
         log_prior_a = gamma.logpdf(a, alpha_a, scale=1/beta_a)
         log_prior_b = gamma.logpdf(b, alpha_b, scale=1/beta_b)
         return log_likelihood + log_prior_a + log_prior_b
     # Metropolis-Hastings algorithm
     def mcmc_gamma(x, alpha_a, beta_a, alpha_b, beta_b, n_samples=5000):
         samples = []
         current_a, current_b = 1.0, 1.0
         for _ in range(n_samples):
             proposed_a = np.random.normal(current_a, 0.1)
             proposed_b = np.random.normal(current_b, 0.1)
             # Ensure positivity
             if proposed_a <= 0 or proposed_b <= 0:</pre>
                 continue
             # Compute log-posterior for current and proposed values
             current_lp = log_posterior(current_a, current_b, x, alpha_a, beta_a,__
      ⇒alpha_b, beta_b)
             proposed_lp = log_posterior(proposed_a, proposed_b, x, alpha_a, beta_a,__
      →alpha_b, beta_b)
             # Metropolis acceptance criterion
             accept_prob = np.exp(proposed_lp - current_lp)
             if np.random.rand() < accept_prob:</pre>
                 current_a, current_b = proposed_a, proposed_b
```

```
samples.append((current_a, current_b))

return np.array(samples)

# Generate data and run MCMC

x = np.random.gamma(shape=2, scale=0.5, size=20)
alpha_a, beta_a = 2, 2
alpha_b, beta_b = 2, 2

samples = mcmc_gamma(x, alpha_a, beta_a, alpha_b, beta_b)
print(f"Generated {len(samples)} posterior samples.")
```

C:\Users\sshubhankar\AppData\Local\Temp\ipykernel_22764\3312317694.py:13:
DeprecationWarning: `np.math` is a deprecated alias for the standard library
`math` module (Deprecated Numpy 1.25). Replace usages of `np.math` with `math`
 log_likelihood = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - b
* S2

Generated 5000 posterior samples.

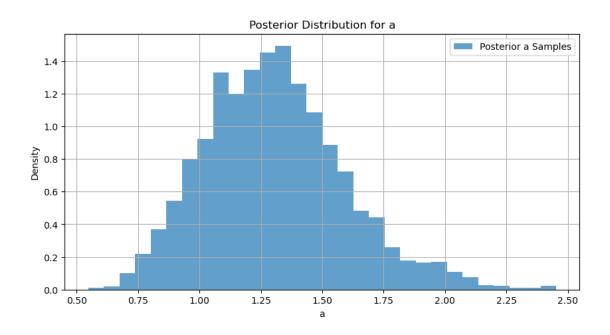
```
[8]: import matplotlib.pyplot as plt
     # Visualize posterior samples
     posterior_a = samples[:, 0]
     posterior_b = samples[:, 1]
     plt.figure(figsize=(10, 5))
     plt.hist(posterior_a, bins=30, alpha=0.7, label="Posterior a Samples", __

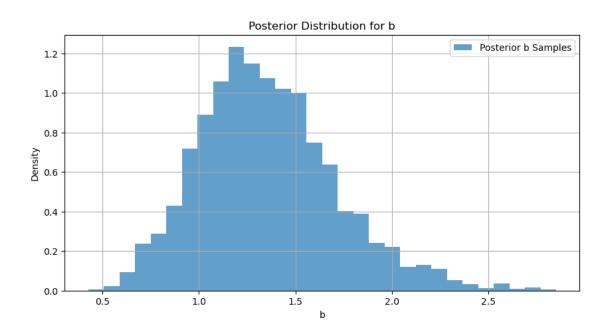
density=True)

     plt.title("Posterior Distribution for a")
     plt.xlabel("a")
     plt.ylabel("Density")
     plt.grid()
     plt.legend()
     plt.show()
     plt.figure(figsize=(10, 5))
     plt.hist(posterior_b, bins=30, alpha=0.7, label="Posterior b Samples", u

density=True)

     plt.title("Posterior Distribution for b")
     plt.xlabel("b")
     plt.ylabel("Density")
     plt.grid()
     plt.legend()
     plt.show()
     #to do? evaluate intervals
```





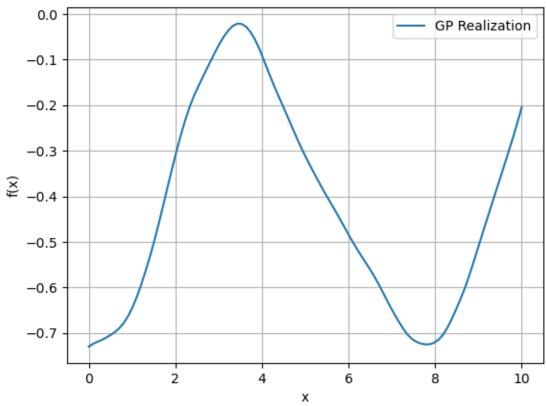
```
[9]: #Simulate realizations of a Gaussian process (GP) with a given covariance function and estimate the hyperparameter () using maximum likelihood.

import numpy as np
import matplotlib.pyplot as plt

# Covariance function
```

```
def covariance_function(x1, x2, lam):
   x1 = np.atleast_2d(x1).T
    x2 = np.atleast_2d(x2).T
    diff = np.abs(x1 - x2.T)
    return (1 + diff / lam + (diff**2) / (3 * lam**2)) * np.exp(-diff / lam)
# Simulate a Gaussian process
def simulate_gaussian_process(x, lam, sigma=1.0, seed=None):
    np.random.seed(seed)
    cov_matrix = sigma**2 * covariance_function(x, x, lam)
    mean = np.zeros(len(x))
    return np.random.multivariate_normal(mean, cov_matrix)
# Generate and visualize GP realizations
x = np.linspace(0, 10, 100)
lam = 2.0
gp_realization = simulate_gaussian_process(x, lam, seed=42)
plt.plot(x, gp_realization, label="GP Realization")
plt.title("Gaussian Process Realization")
plt.xlabel("x")
plt.ylabel("f(x)")
plt.legend()
plt.grid()
plt.show()
```





```
[10]: #Hyperparameter Estimation
      from scipy.optimize import minimize
      # Negative log-likelihood function
      def negative_log_likelihood(lam, x, y, sigma=1.0):
          cov_matrix = sigma**2 * covariance_function(x, x, lam) + 1e-6 * np.
       \hookrightarroweye(len(x))
          try:
              K_inv = np.linalg.inv(cov_matrix)
          except np.linalg.LinAlgError:
              return np.inf
          log_det = np.linalg.slogdet(cov_matrix)[1]
          return 0.5 * y.T @ K_inv @ y + 0.5 * log_det + 0.5 * len(x) * np.log(2 * np.
       ⊶pi)
      # Optimize lambda
      y = gp_realization # Simulated data
      result = minimize(
          lambda lam: negative_log_likelihood(lam, x, y, sigma=1.0),
```

```
x0=[1.0],
    bounds=[(0.1, 10)]
)
optimal_lambda = result.x[0]
print(f"Optimal lambda: {optimal_lambda:.4f}")
```

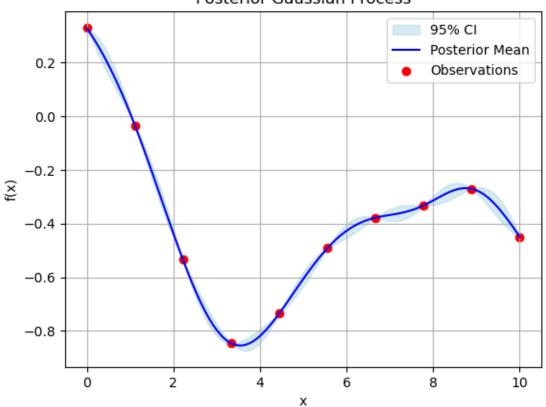
Optimal lambda: 2.1142

```
[11]: #Simulate a posterior Gaussian process conditioned on observed data 1,..., x n
      →and plot the posterior mean and confidence intervals.
      #Gaussian Process Posterior Simulation
      #x_pred spans the range where the GP is estimated.
      def gp_posterior(x_obs, y_obs, x_pred, lam, sigma=1.0):
          K_obs = covariance_function(x_obs, x_obs, lam) + 1e-6 * np.eye(len(x_obs))
          K_pred = covariance_function(x_pred, x_obs, lam)
          K_pred_pred = covariance_function(x_pred, x_pred, lam)
          K_obs_inv = np.linalg.inv(K_obs)
          # Posterior mean
          mu_pred = K_pred @ K_obs_inv @ y_obs
          # Posterior covariance
          cov_pred = K_pred_pred - K_pred @ K_obs_inv @ K_pred.T
          return mu_pred, cov_pred
      # Simulate observed data
      x_{obs} = np.linspace(0, 10, 10)
      y_obs = simulate_gaussian_process(x_obs, lam=2.0, seed=42)
      # Prediction points
      x_pred = np.linspace(0, 10, 100)
      # Compute posterior
      mu_pred, cov_pred = gp_posterior(x_obs, y_obs, x_pred, lam=2.0)
      # Plot posterior mean and confidence intervals
      std_pred = np.sqrt(np.diag(cov_pred))
      plt.fill_between(x_pred, mu_pred - 2 * std_pred, mu_pred + 2 * std_pred,__

color="lightblue", alpha=0.5, label="95% CI")
      plt.plot(x_pred, mu_pred, label="Posterior Mean", color="blue")
      plt.scatter(x_obs, y_obs, color="red", label="Observations")
      plt.title("Posterior Gaussian Process")
      plt.xlabel("x")
```

```
plt.ylabel("f(x)")
plt.legend()
plt.grid()
plt.show()
```

Posterior Gaussian Process



```
#Mutual Information Estimation Using Kernel Density Estimation (KDE)
#Compute the mutual information for a bivariate Gaussian distribution and_
approximate it using KDE.

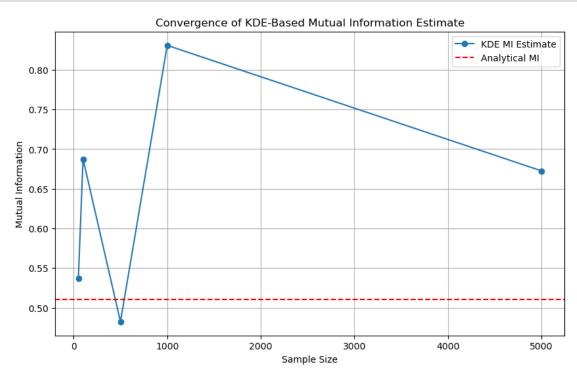
def analytical_mutual_information(cov_matrix):
    rho = cov_matrix[0, 1] / np.sqrt(cov_matrix[0, 0] * cov_matrix[1, 1])
    return -0.5 * np.log(1 - rho**2)

# Example covariance matrix
cov_matrix = np.array([[1.0, 0.8], [0.8, 1.0]]) # High correlation
mi_analytical = analytical_mutual_information(cov_matrix)
print(f"Analytical Mutual Information: {mi_analytical:.4f}")
```

Analytical Mutual Information: 0.5108

```
[13]: from scipy.stats import gaussian_kde
      def kde_mutual_information(samples, grid_size=100):
          kde_joint = gaussian_kde(samples.T)
          kde_x = gaussian_kde(samples[:, 0])
          kde_y = gaussian_kde(samples[:, 1])
          x_grid = np.linspace(samples[:, 0].min(), samples[:, 0].max(), grid_size)
          y_grid = np.linspace(samples[:, 1].min(), samples[:, 1].max(), grid_size)
          x, y = np.meshgrid(x_grid, y_grid)
          joint_density = kde_joint(np.vstack([x.ravel(), y.ravel()])).
       ⇒reshape(grid size, grid size)
          marginal_x = kde_x(x_grid)
          marginal_y = kde_y(y_grid)
          dx = x grid[1] - x grid[0]
          dy = y_grid[1] - y_grid[0]
          mi = stimate = 0
          for i in range(grid_size):
              for j in range(grid_size):
                  if joint_density[i, j] > 0: # Avoid log(0)
                      mi_estimate += joint_density[i, j] * np.log(
                          joint_density[i, j] / (marginal_x[i] * marginal_y[j])
                      ) * dx * dy
          return mi_estimate
      # Simulate samples and compute KDE-based MI
      samples = np.random.multivariate_normal(mean=[0, 0], cov=cov_matrix, size=1000)
      mi_plugin = kde_mutual_information(samples)
      print(f"KDE-Based Mutual Information: {mi plugin:.4f}")
     KDE-Based Mutual Information: 0.4921
[14]: print(f"Analytical Mutual Information: {mi_analytical:.4f}")
      print(f"KDE-Based Mutual Information: {mi_plugin:.4f}")
     Analytical Mutual Information: 0.5108
     KDE-Based Mutual Information: 0.4921
[15]: def evaluate_convergence(cov_matrix, sample_sizes, grid_size=100):
          analytical_mi = analytical_mutual_information(cov_matrix)
          kde_estimates = []
          for n in sample_sizes:
              samples = np.random.multivariate_normal(mean=[0, 0], cov=cov_matrix,_
       ⇔size=n)
```

```
kde_mi = kde_mutual_information(samples, grid_size=grid_size)
       kde_estimates.append(kde_mi)
   return analytical_mi, kde_estimates
# Sample sizes to test
sample_sizes = [50, 100, 500, 1000, 5000]
# Evaluate convergence
cov_matrix = np.array([[1.0, 0.8], [0.8, 1.0]])
analytical_mi, kde_estimates = evaluate_convergence(cov_matrix, sample_sizes)
# Plot convergence
plt.figure(figsize=(10, 6))
plt.plot(sample_sizes, kde_estimates, label="KDE MI Estimate", marker="o")
plt.axhline(analytical_mi, color="red", linestyle="--", label="Analytical MI")
plt.title("Convergence of KDE-Based Mutual Information Estimate")
plt.xlabel("Sample Size")
plt.ylabel("Mutual Information")
plt.legend()
plt.grid()
plt.show()
# robustness of KDE-based methods for estimating mutual information
```



```
[18]: #Compare the performance of the Maximum Likelihood Estimator (MLE), Method of
       →Moments Estimator (MME), and Bayesian Estimator using quadratic risk.
      def compare_risks(n_simulations, sample_size, a_star, b_star, alpha_a, beta_a,_
       →alpha_b, beta_b):
          mle risks = []
          mme_risks = []
          bayes_risks = []
          for _ in range(n_simulations):
              x = np.random.gamma(shape=a_star, scale=1/b_star, size=sample_size)
              # MME
              a_mme, b_mme = mme_gamma(x)
              mme_error = (a_mme - a_star)**2 + (b_mme - b_star)**2
              # MLE
              a_mle, b_mle = newton_raphson_mle(x, init="mme")
              mle_error = (a_mle - a_star)**2 + (b_mle - b_star)**2
              # Bayesian Estimator
              samples = mcmc_gamma(x, alpha_a, beta_a, alpha_b, beta_b, n_samples=500)
              bayes_mean_a = np.mean(samples[:, 0])
              bayes_mean_b = np.mean(samples[:, 1])
              bayes_error = (bayes_mean_a - a_star)**2 + (bayes_mean_b - b_star)**2
              mle_risks.append(mle_error)
              mme risks.append(mme error)
              bayes_risks.append(bayes_error)
          # Average risks
          mle_risk = np.mean(mle_risks)
          mme_risk = np.mean(mme_risks)
          bayes_risk = np.mean(bayes_risks)
          return mle_risk, mme_risk, bayes_risk
      # Compare risks
      sample_size = 20
      n_simulations = 100
      a_star, b_star = 2, 2
      alpha_a, beta_a = 2, 2
      alpha_b, beta_b = 2, 2
      mle_risk, mme_risk, bayes_risk = compare_risks(
          n_simulations, sample_size, a_star, b_star, alpha_a, beta_a, alpha_b, beta_b
      )
```

```
print(f"MLE Risk: {mle_risk:.4f}")
      print(f"MME Risk: {mme_risk:.4f}")
      print(f"Bayesian Risk: {bayes_risk:.4f}")
     C:\Users\sshubhankar\AppData\Local\Temp\ipykernel_22764\3312317694.py:13:
     DeprecationWarning: `np.math` is a deprecated alias for the standard library
     `math` module (Deprecated Numpy 1.25). Replace usages of `np.math` with `math`
       \log_{\text{likelihood}} = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - b
     * S2
     MLE Risk: 1.2294
     MME Risk: 2.5194
     Bayesian Risk: 0.7145
[20]: sample_sizes = [10, 20, 50, 100]
      risks = []
      for n in sample_sizes:
          risks.append(
              compare risks(n_simulations, n, a_star, b_star, alpha_a, beta_a,_
       →alpha_b, beta_b)
          )
      mle_risks, mme_risks, bayes_risks = zip(*risks)
      # Plot risks
      plt.figure(figsize=(10, 6))
      plt.plot(sample_sizes, mle_risks, label="MLE Risk", marker="o")
      plt.plot(sample_sizes, mme_risks, label="MME Risk", marker="o")
      plt.plot(sample_sizes, bayes risks, label="Bayesian Risk", marker="o")
      plt.title("Risk Comparison Across Sample Sizes")
      plt.xlabel("Sample Size")
      plt.ylabel("Quadratic Risk")
      plt.legend()
      plt.grid()
      plt.show()
     C:\Users\sshubhankar\AppData\Local\Temp\ipykernel_22764\3312317694.py:13:
     DeprecationWarning: `np.math` is a deprecated alias for the standard library
```

C:\Users\sshubhankar\AppData\Local\Temp\ipykernel_22764\3312317694.py:13:
DeprecationWarning: `np.math` is a deprecated alias for the standard library
`math` module (Deprecated Numpy 1.25). Replace usages of `np.math` with `math`
 log_likelihood = n * a * np.log(b) - n * np.math.lgamma(a) + (a - 1) * S1 - b
* S2

Invalid parameters encountered, resetting.

