

Shucen Liu

Assn 2

1. (a) $\hat{c} = \arg \min_{c \in C_k} R_k(c)$.
 where $R_k(c) = \frac{1}{n} \sum_{i=1}^n \min_{1 \leq j \leq k} \|x_i - c_j\|^2$
 $c^* = \arg \min_{c \in C_k} R(c)$
 where $R(c) = E \min_{1 \leq j \leq k} \|X - c_j\|^2$

$R(c^*)$ is the minimum of $R(c)$.

If $R(\hat{c})$ is close to $R(c^*)$,
 since the distribution of X may not be
 equal (the probability of X may not be
 the same, $P(X_1) \neq P(X_2) \neq P(X_3)$.)

In certain case that $P(X_i) \neq \frac{1}{n}$,
 \hat{c} will be far away from c^* , even if
 $R(\hat{c})$ is close to $R(c^*)$

(b) $R^{(k)} = \min R(c)$
 $= \min E \min_{1 \leq j \leq k} \|X - c_j\|^2$

As k increases, there will be more cluster centers.

Thus for the best clusters,
 the difference between the points
 and the center will not increase
 as k increases.

(c) Taking derivative w.r.t λ_i , we'll get
 $\sum \lambda_i \cdot (X_i - \mu - V_k \lambda_i)^T V_k (X_i - \mu - V_k \lambda_i) = 0$.
 Thus we must

have $X_i = \mu + V_k \lambda_i$.

If we have $\bar{\mu} = \bar{X}_i$.

and $\lambda_i = V_k^T (X_i - \bar{X}_n)$

$$\Rightarrow X_i = \bar{X}_i + V_k (V_k^T (X_i - \bar{X}_n))$$

$$= \bar{X}_i + X_i - \bar{X}_n \quad (V_k \text{ orthogonal})$$

$$= X_i$$

Thus, the formula is minimized.

$$[V_k] : - \sum (X_i - \mu - V_k \lambda_i)^T V_k (X_i - \mu - V_k \lambda_i) = V_k \cdot \sum ([X_i])$$

$$[\mu] : - \sum (X_i - \mu - V_k \lambda_i) \mu (X_i - \mu - V_k \lambda_i) = \mu \cdot \sum ([X_i])$$

As long as the derivative of $\lambda_i = 0$, the other
 two first order conditions are satisfied.

Thus, μ is not unique, as long as

$$\mu = X_i - V_k \lambda_i \quad (\text{we can adjust } \lambda \text{ when } \mu \text{ changes})$$