## Verweis von Bosker & Snijder 2012

## 2. Analysis of colour morphs (Binomial mixed models)

```
# Clear memory
rm(list = ls())
# Read colour morph data (Binary, available for males only)
library(AzureML)
ws <- workspace()
Data <- download.datasets(
  dataset = ws.
         = "BeetlesMale.csv")
  name
# Fit null model without fixed effects (but including all random effects)
m0 <- glmer(Colour ~ 1 + (1 | Population) + (1 | Container),
            family = "binomial", data = Data)
summary(m0)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
## Family: binomial (logit)
## Formula: Colour ~ 1 + (1 | Population) + (1 | Container)
     Data: Data
##
##
##
        AIC
                BIC
                       logLik deviance df.resid
##
      602.4
               614.9
                     -298.2
                                 596.4
                                            477
##
## Scaled residuals:
      Min
              1Q Median
                                3Q
## -1.5040 -0.7071 -0.3306 0.8486 3.0248
##
## Random effects:
## Groups
              Name
                           Variance Std.Dev.
## Container (Intercept) 0.0000
                                    0.0000
## Population (Intercept) 0.9458
                                    0.9725
## Number of obs: 480, groups: Container, 60; Population, 12
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
                            0.2997 -1.261
                                              0.207
## (Intercept) -0.3779
# Fit alternative model including fixed and all random effects
mF <- glmer(Colour ~ Treatment + Habitat + (1 | Population) + (1 | Container),
            family = "binomial", data = Data)
summary(mF)
## Generalized linear mixed model fit by maximum likelihood (Laplace
```

```
Approximation) [glmerMod]
## Family: binomial (logit)
## Formula: Colour ~ Treatment + Habitat + (1 | Population) + (1 | Container)
     Data: Data
##
##
##
       AIC
                BIC
                       logLik deviance df.resid
##
      573.1
               594.0
                      -281.6
                                 563.1
##
## Scaled residuals:
##
      Min
                1Q Median
                                3Q
                                       Max
## -2.3762 -0.7204 -0.3274 0.7662 3.2282
##
## Random effects:
                           Variance Std.Dev.
## Groups
               Name
## Container (Intercept) 0.005793 0.07611
## Population (Intercept) 1.110213 1.05367
## Number of obs: 480, groups: Container, 60; Population, 12
##
## Fixed effects:
               Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                -1.2511
                             0.3641 -3.436 0.00059 ***
## TreatmentExp
                1.0118
                             0.2147
                                    4.712 2.45e-06 ***
                                    3.192 0.00141 **
## HabitatB
                 0.6760
                             0.2118
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
               (Intr) TrtmnE
## TreatmntExp -0.343
              -0.332 0.079
## HabitatB
# Extraction of fitted value for the alternative model
# fixef() extracts coefficents for fixed effects
# mF@X returns fixed effect design matrix
Fixed <- fixef(mF)[2] * getME(mF,"X")[, 2] + fixef(mF)[3] * getME(mF,"X")[, 3]
fixef(mF)[2]
## TreatmentExp
##
      1.011793
fixef(mF)[3]
## HabitatB
## 0.676039
# Calculation of the variance in fitted values
VarF <- var(Fixed)</pre>
VarF
## [1] 0.3709611
```

```
# An alternative way for getting the same result
VarF <- var(as.vector(fixef(mF) %*% t(getME(mF,"X"))))</pre>
```

$$R_{\text{GLMM}(m)}^{2} = \frac{\sigma_f^2}{\sigma_f^2 + \sum_{l=1}^{u} \sigma_l^2 + \sigma_e^2 + \sigma_d^2},$$
 eqn 29

Figure 1:

$$R_{\text{GLMM}(c)}^{2} = \frac{\sigma_f^2 + \sum_{l=1}^{u} \sigma_l^2}{\sigma_f^2 + \sum_{l=1}^{u} \sigma_l^2 + \sigma_e^2 + \sigma_d^2}.$$
 eqn 30

Figure 2:

```
# R2GLMM(m) - marginal R2GLMM
# see Equ. 29 and 30 and Table 2
VarF/(VarF + VarCorr(mF)$Container[1] + VarCorr(mF)$Population[1] + pi^2/3)
## [1] 0.07765834

VarCorr(mF)$Container[1]
## [1] 0.00579262
VarCorr(mF)$Population[1]
## [1] 1.110213

# R2GLMM(c) - conditional R2GLMM for full model
# Equ. XXX, XXX
(VarF + VarCorr(mF)$Container[1] + VarCorr(mF)$Population[1])/
    (VarF + VarCorr(mF)$Container[1] + VarCorr(mF)$Population[1] + pi^2/3)
## [1] 0.311287
```

## Literatur

Nakagawa, Shinichi, Holger Schielzeth, and Robert B. O'Hara. 2013. "A General and Simple Method for Obtaining R2 from Generalized Linear Mixed-Effects Models." *Methods in Ecology and Evolution* 4 (2): 133–42. doi:10.1111/j.2041-210x.2012.00261.x.

	Binary and proportion data	
Link function Distribution-specific variance	Logit link $\pi^2/3$	Probit link 1
Model specification	$Y_{ijk} = \text{Binomial}(m_{ijk}, p_{ijk}),$	
	$\operatorname{link}(p_{ijk}) = \beta_0 + \sum_{h=1}^{p} \beta_h x_{hijk} + \gamma_k + \alpha_{jk} + e_{ijk}$	
	$\gamma_k \sim \text{Gaussian}(0, \sigma_{\gamma}^2)$	
	$\alpha_{\not r} \sim \operatorname{Gaussian}(0, \sigma_{\alpha}^2)$	
	$e_{ijk} \sim \text{Gaussian}(0, \sigma_e^2)$	
Description	$Y_{ijk}$ is the number of 'successes' in $m_{ijk}$ trials by the jth individual in the kth group at the ith occasion (for binary data, $m_{ijk}$ is 1), $p_{ijk}$ is the underlying (latent) probability of success for the jth individual in the kth group at the ith occasion (for binary data, $\sigma_e^2$ is 0).	
Marginal $R^2$ $(R^2_{GLMM(m)})$	$R_{\text{logit}(m)}^2 = \frac{\sigma_f^2}{\sigma_f^2 + \sigma_\gamma^2 + \sigma_\alpha^2 + \sigma_e^2 + \pi^2/3}$	$R_{\mathrm{probit}(m)}^2 = \frac{\sigma_f^2}{\sigma_f^2 + \sigma_\gamma^2 + \sigma_\alpha^2 + \sigma_e^2 + 1}$
Conditional $R^2$ $(R^2_{GLMM(c)})$	$R_{\text{logit}(c)}^2 = \frac{\sigma_f^2 + \sigma_\gamma^2 + \sigma_\alpha^2}{\sigma_f^2 + \sigma_\gamma^2 + \sigma_\alpha^2 + \sigma_e^2 + \pi^2/3}$	$R_{\mathrm{probit}(c)}^{2} = \frac{\sigma_{f}^{2} + \sigma_{\gamma}^{2} + \sigma_{\alpha}^{2}}{\sigma_{f}^{2} + \sigma_{\gamma}^{2} + \sigma_{\alpha}^{2} + \sigma_{e}^{2} + 1}$

Figure 3: