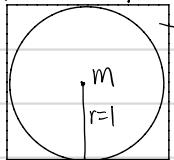


# Basic Monte carlo

L03 S01

1. What is def of MC?

2. Ex: Can you approximate  $\pi$ ?



suppose: ① I make  $n$ -shots to the rectangular, and  $m$ -shots fall in circle.

② shots are uniformly distributed, then  $\pi \approx \frac{m}{n} \cdot 4$

① Draw  $N$  random points uniformly in rectangular  $(-1, 1) \times (-1, 1)$  i.e.

$$\{z_i = (x_i, y_i) : x_i \sim U(-1, 1)$$

$$y_i \sim U(-1, 1) \text{ iid. } 1 \leq i \leq N\}$$

② Identify the number  $N_1$  of points inside the circle.

$$\hat{\pi} = 4 \cdot \frac{N_1}{N}$$

RK:  $\pi$  is deterministic value. But its approximation  $\hat{\pi}$  is random.

Def: A random estimator (approximation)  $\hat{x}$  to a deterministic value  $x$  is called MC (Monte carlo)

$$\text{① Bias} = E[\hat{x}] - x$$

$$\text{② MSE}(\hat{x}) = E[(\hat{x} - x)^2] = |\text{Bias}|^2 + \text{Var}(\hat{x})$$

$$\text{proof: } \text{MSE}(\hat{x}) = E[(\hat{x} - x)^2] = E(\beta^2), \text{ value } \beta = \hat{x} - x$$

$$= \text{Var}(\beta) + (E(\beta))^2 = \text{Var}(\hat{x} - x) + [E(\hat{x} - x)]^2 = \text{Var}(\hat{x}) + (E(\hat{x}) - E(x))^2$$

$$= \text{Var}(\hat{x}) + (E(\hat{x}) - x)^2 = \text{Var}(\hat{x}) + |\text{Bias}|^2$$

$$\text{draft: } \text{Var}(x) = E(x^2) - (E(x))^2$$

$$\text{Var}(X) = \text{Var}(X+C) \quad C \text{ is a constant.}$$

$$E(X+C) = E(X) + C$$

The smaller MSE, the better MC.

MC is called unbiased if  $\text{Bias} = 0$ . i.e.  $E(\hat{x}) = x$

Ex:

In approximation  $\pi$  by  $\hat{\pi}$  ① Find 1<sup>st</sup>, 2<sup>nd</sup> moments of  $\hat{\pi}$  ② Find Bias & MSE of  $\hat{\pi}$   
(Def  $m^{\text{th}}$  moment is  $E(x^m)$ )

$$\text{solv: } ① E(\hat{\pi}) = E(4 \cdot \frac{N_1}{N}) = \frac{4}{N} E(N_1) = \frac{4}{N} E\left(\sum_{i=1}^N I_{\{|z_i| < 1\}}\right) = \frac{4}{N} \sum_{i=1}^N E(I_{\{|z_i| < 1\}}) = \frac{4}{N} [E(I_{\{|z_1| < 1\}}) + E(I_{\{|z_2| < 1\}}) + \dots + E(I_{\{|z_N| < 1\}})] \\ = \frac{4}{N} \cdot N \cdot E(I_{\{|z_1| < 1\}}) = 4 \cdot P(|z_1| < 1) = 4 \cdot \frac{\text{Circle Area}}{\text{Rectangular Area.}} = 4 \cdot \frac{\pi \cdot 1^2}{2 \times 2} = \pi$$

$$\text{Recall: } \begin{cases} 1, & \text{if } A \text{ occur} \\ 0, & \text{if } A \text{ not occur} \end{cases} \quad E(X) + E(Y) = E(X+Y)$$

$$I_A = \begin{cases} 1, & \text{if } A \text{ occur} \\ 0, & \text{if } A \text{ not occur} \end{cases} \quad E(I_A) = P(A)$$

$$\text{Var}(X \cdot C) = C^2 \cdot \text{Var}(X)$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) \neq X+Y$$

$$E(\hat{\pi}^2) = [E(\hat{\pi})]^2 + \text{Var}(\hat{\pi}) = \pi^2 + \text{Var}(\hat{\pi})$$

$$\text{Var}(\hat{\pi}) = \text{Var}\left(4 \cdot \frac{N_1}{N}\right) = \frac{16}{N^2} \text{Var}(N_1) = \frac{16}{N^2} \text{Var}\left(\sum_{i=1}^N I_{\{|z_i| < 1\}}\right) = \frac{16}{N^2} \sum_{i=1}^N \text{Var}\left(I_{\{|z_i| < 1\}}\right) = \frac{16}{N^2} \cdot N \cdot \text{Var}\left(I_{\{|z_1| < 1\}}\right) = \frac{16}{N} \cdot P(|z_1| < 1) \cdot (1 - P(|z_1| < 1)) \\ \Rightarrow E(\hat{\pi}^2) = \pi^2 + \frac{\pi(4-\pi)}{N} = \frac{(N\pi - \pi^2)\pi}{N}$$

$$\text{② Bias} = E(\hat{\pi}) - \pi = 0$$

$$\text{MSE}(\hat{\pi}) = |\text{Bias}|^2 + \text{Var}(\hat{\pi}) = 0 + \frac{\pi(4-\pi)}{N} = \frac{\pi(4-\pi)}{N}$$

—

?

Def: Given a series of estimator of  $\alpha$ , say  $\{\hat{\alpha}_n, n \in \mathbb{N}\}$ , we say  $(\hat{\alpha}_n)$  is consistent if

$$P - \lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$$

prop: If  $MSE(\hat{\alpha}_n) \rightarrow 0$ , then  $\hat{\alpha}_n$  is consistent.

ex: If  $N$  is fixed, then  $Bias(\hat{\alpha}) = 0$

prop: Suppose  $n$  ppl. do the some unbiased approximation, their out comes are  $\{\hat{\alpha}^i : 1 \leq i \leq n\}$   
 $\sum_{i=1}^n \hat{\alpha}^i \rightarrow \alpha$  always surely.

Def: Law of large number (LLN)

ex: Let  $\hat{\pi}_N$  be estimator with  $N$  total points. Then  $\{\hat{\pi}_N : N \in \mathbb{N}\}$  is a consistent estimator of  $\pi$ . because  
 $MSE(\hat{\pi}_N) = \frac{\pi(1-\pi)}{N} \xrightarrow[N \rightarrow \infty]{} 0$

Def: ① "  $P - \lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$ " or " $\hat{\alpha}_n \xrightarrow{n \rightarrow \infty} \alpha$ " in prob" if  $\lim_{n \rightarrow \infty} P(|\hat{\alpha}_n - \alpha| > \epsilon) = 0$ ,  $\forall \epsilon > 0$ .  
 ② " $\lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$ " or " $\hat{\alpha}_n \xrightarrow{n \rightarrow \infty} \alpha$ " a.s if  $P(\lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha) = 1$   
 ③ " $L_2 - \lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$ " or " $\hat{\alpha}_n \xrightarrow{n \rightarrow \infty} \alpha$ " in  $L^2$  if  $\lim_{n \rightarrow \infty} E|\hat{\alpha}_n - \alpha|^2 = 0$

ex: Justify the statement " $P - \lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$ " implies " $L_2 - \lim_{n \rightarrow \infty} \hat{\alpha}_n = \alpha$ "

Ans: No. because following counter-example:

Let  $P = U([0, 1]) \rightarrow \text{probability}$

$\Omega = [0, 1] \rightarrow \text{sample space}$

$\hat{\alpha}_n : [0, 1] \rightarrow \mathbb{R}$  st.

$$\hat{\alpha}_n(\omega) = \begin{cases} n & \text{if } 0 < \omega < \frac{1}{n} \\ 0 & \text{otherwise} \end{cases}$$

$X = [0, 1] \rightarrow \mathbb{R}$ . st.

$$X(\omega) = 0$$

$$\Rightarrow \hat{\alpha}_2(\omega) = \begin{cases} \alpha & \text{if } 0 < \omega < \frac{1}{4} \\ 0 & \text{otherwise.} \end{cases}$$

$\mathbb{R}^2 = \mathbb{R} \times \mathbb{R} = \{(x, y) : x \in \mathbb{R}, y \in \mathbb{R}\}$ .  $z \in \mathbb{R}^2$  means a point on  $\mathbb{R}^2$ -plane  $|z| = \sqrt{|zx|^2 + |zy|^2}$

$$I_{\{|z|=1\}} = \begin{cases} 1 & \text{if } |z| < 1 \\ 0 & \text{if } |z| \geq 1 \end{cases}$$

$$\text{if } x = \begin{pmatrix} a_1 & p_1 \\ a_2 & p_2 \\ \vdots & \vdots \\ a_n & p_n \end{pmatrix} \text{ Then } E(x) = \sum_{i=1}^n a_i p_i; \quad \text{if } X = I_A = \begin{cases} 1 & P(A) \\ 0 & 1 - P(A) \end{cases}$$

$$E(I_A) = E(X) = 1 \cdot P(A) + 0 \cdot [1 - P(A)] = P(A)$$

$$E(I_A^2) = 1^2 \cdot P(A) + 0^2 \cdot (1 - P(A)) = P(A)$$

$$\text{Var}(I_A) = E(I_A^2) - [E(I_A)]^2 = P(A) - P^2(A) = P(A) \cdot [1 - P(A)]$$

Sep 20. L03 S02

Recall prop1: If  $MSE(\hat{x}_n) \rightarrow 0$ , then  $\hat{x}_n$  is consistent. (prop1 is the consequence of the following fact.)

prop2:  $L_2 - \lim_{n \rightarrow \infty} x_n = \alpha$  implies  $p - \lim_{n \rightarrow \infty} x_n = \alpha$

Pf:  $\forall \varepsilon > 0, P(|x_n - \alpha| > \varepsilon) = P\left(\frac{|x_n - \alpha|}{\varepsilon} > 1\right) \leq E\left(\frac{|x_n - \alpha|^2}{\varepsilon^2}\right) (\rightarrow 0)$  by Chebyshev inequality, b/c  $L_2 - \lim_{n \rightarrow \infty} x_n = \alpha$ .

RK Chebyshov inequality is used above i.e.  $P(|X| > 1) \leq E(|X|)$

Pf RHS =  $\int_{\mathbb{R}} |x| p(x) dx$ , when  $p \sim X \geq \int_{-\infty}^{-1} p(x) dx + \int_1^{\infty} p(x) dx = LHS$

prop If  $F$  is a strictly increasing CDF.  $U \sim U(0,1)$ . uniform on r.v. on  $[0,1]$

then  $X = F^{-1}(U)$  has its EDF  $F$ .

$$\begin{aligned} Pf \quad & P(X < a) \stackrel{?}{=} F(a) \\ &= P(F^{-1}(U) < a) \\ &= P(U < F(a)) \\ &= F(a) \end{aligned} \quad \left| \begin{array}{l} U \sim U(0,1) \\ P(U < \frac{1}{2}) = \frac{1}{2} \\ P(U < \tilde{a}) = \tilde{a} \quad \text{if } \tilde{a} \in (0,1) \end{array} \right.$$

A company sells option.

If ICC crashes within 10 yrs, Buyer will obtain \$100, otherwise \$0

Suppose  $P(\text{ICC crashes within 10 yr}) = 1\%$ , then what is fair price of the option?

## L04 S01 Black Sholes formula

BS model assumes the distribution of stock as lognormal. In particular, it writes:

$$\ln \frac{S(T)}{S(0)} \sim N\left(r - \frac{1}{2}\sigma^2 T, \sigma^2 T\right)$$

The call & put price with maturity  $T$  &  $K$  will be known as  $C_0$  &  $P_0$  given as below:

$$\begin{cases} C_0 = E\left[e^{-rT} \cdot (S(T) - K)^+\right] = S_0 \phi(d_1) - K e^{-rT} \phi(d_2) \\ P_0 = E\left[e^{-rT} \cdot (S(T) - K)^-\right] = K e^{-rT} \phi(-d_2) - S_0 \phi(-d_1) \end{cases} \quad \begin{cases} d_1 = \frac{(r + \frac{1}{2}\sigma^2)T - \ln \frac{K}{S_0}}{\sigma \sqrt{T}} \\ d_2 = \frac{(r - \frac{1}{2}\sigma^2)T - \ln \frac{K}{S_0}}{\sigma \sqrt{T}} = d_1 - \sigma \sqrt{T} \end{cases}$$

Verify put-call parity:  $C_0 - P_0 = S(0) - e^{-rT} K$

$\Rightarrow B-S$ -valuation (rewrite code)

## L04 S02 Calibration of BSM Volatility

In simple terms, problem of calibration is to find parameters for the proposed model such that observed market quotes of liquidly traded plain vanilla options are replicated as closely as possible. The general approach is that, one defines an error function that is to be minimized.

In this below, we will use Sep-30-2014 market quote on call options underlying Euro Stock 50-BSM model -RMSE as an error function.

### Calibration by minimizing RMSE

Suppose there are  $N$  options available in the market quote with various strike and maturities. We denote the market quote of  $N$  options by  $\{C_n^* : n=1, 2, \dots, N\}$

Correspondingly, there are  $N$  BSM theoretical prices available as a function of the volatility. We denote it by:  $\{C_n(\sigma) : n=1, 2, \dots, N\}$

The error function RMSE is defined by:  $RMSE(\sigma) = \sqrt{\frac{1}{N} \sum_{n=1}^N |C_n^* - C_n(\sigma)|^2}$

Our goal is to find out the calibrated volatility  $\hat{\sigma}$ , which is given by:  $\hat{\sigma} = \arg \min_{\sigma} RMSE(\sigma)$

Sep 27

Zero Coupon Bond ( $T$ )  $\rightarrow$  maturity payoff( $T$ ) = 1

suppose  $r$  is rate, today's price is  $P(0, T)$  ( $0$  for today,  $T$  for maturity)

then  $P(0, T) \cdot e^{rT} = 1$

$$P(0, T) = e^{-rT}$$

stock  $\sim BSM(\sigma)$

suppose  $S_t$  is price at  $t$

$W_t$  is Brownian Motion

$$S_t = S_0 \cdot \exp \left\{ (r - \frac{1}{2}\sigma^2)t + \sigma W_t \right\} \rightarrow ENM \otimes$$

$$W_t = (W_{t_1} - W_{t_0}) + (W_{t_2} - W_{t_1}) + \dots + (W_{t_n} - W_{t_{n-1}}) = \Delta W_{t_1} + \Delta W_{t_2} + \dots + \Delta W_{t_n}$$

proposition ①  $\{\Delta W_{t_1}, \Delta W_{t_2}, \dots, \Delta W_{t_n}\}$  are independent.

$$\textcircled{2} \Delta W_{t_n} = W_{t_n} - W_{t_{n-1}} \sim N(0, \sigma t_n) \quad (\sigma t_n = t_n - t_{n-1})$$

### Asian Option (Geometric)

Payoff for  $\downarrow GAC \downarrow \downarrow$  ( $T, K, n$ )

Geom Asian call

$$\Pi_T^C = (A_T - K)^+$$

$$\text{where } \Delta T = \frac{T}{n}$$

$$A_T = (S_{t_1} \cdot S_{t_2} \cdots S_{t_n})^{\frac{1}{n}}$$

payoff for GAP ( $T, K, n$ )

$$\Pi_T^P = (A_T - K)^-$$

Goal what's the price of GAC & GAP

$$\text{Ans } \Pi_0^C = \mathbb{E}^Q \left[ e^{-rT} \Pi_T^C \right] \quad \Pi_0^P = \mathbb{E}^Q \left[ e^{-rT} \Pi_T^P \right]$$

$$0 \xrightarrow{\Delta t} t_1 \xrightarrow{\Delta t} t_2 \xrightarrow{\dots} t_{n-1} \xrightarrow{\Delta t} t_n = T$$

$$S(t_i) = S_0 \cdot \exp \left\{ i\mu \Delta t + \sigma \sqrt{\Delta t} \sum_{j=1}^i \hat{Z}_j \right\}$$

$$\prod_{i=1}^n S(t_i) = S_0^n \cdot \exp \left\{ \sum_{i=1}^n \left( i\mu \Delta t + \sigma \sqrt{\Delta t} \sum_{j=1}^i \hat{Z}_j \right) \right\} = S_0^n \cdot \exp \left\{ \frac{n(n+1)}{2} \mu \Delta t + \sigma \sqrt{\Delta t} \sum_{i=1}^n \sum_{j=1}^i \hat{Z}_j \right\}$$

$$= S_0^n \cdot \exp \left\{ \frac{n(n+1)}{2} \mu \Delta t + \sigma \sqrt{\Delta t} \cdot \underbrace{\sum_{i=1}^n (n+1-i) \hat{Z}_i}_{a \text{ is sum of normal random variables, thus is another normal random variable. Its mean is.}} \right\}$$

a is sum of normal random variables, thus is another normal random variable. Its mean is.

$$\mathbb{E}(a) = \mathbb{E} \left( \sum_{i=1}^n (n+1-i) \hat{Z}_i \right) = \sum_{i=1}^n \mathbb{E}(n+1-i) \hat{Z}_i = 0$$

$$\text{Var}(a) = \text{Var} \left( \sum_{i=1}^n (n+1-i) \hat{Z}_i \right) = \sum_{i=1}^n \text{Var}(n+1-i) \hat{Z}_i = \sum_{i=1}^n (n+1-i)^2 \cdot \frac{\text{Var}(\hat{Z}_i)}{n} = n^2 + (n-1)^2 + \dots + 2^2 + 1^2$$

$$\therefore a \sim N(0, \frac{n(n+1)(2n+1)}{6}) = \frac{n(n+1)(2n+1)}{6} \hat{Z}, \text{ where } \hat{Z} \sim N(0, 1)$$

$$A_T = (S_{t_1} \cdot S_{t_2} \cdots S_{t_n})^{\frac{1}{n}} = \left[ \prod_{i=1}^n (S_{t_i}) \right]^{\frac{1}{n}} = S_0 \cdot \exp \left\{ \frac{1}{n} \cdot \left( \frac{n(n+1)}{2} \mu \Delta t + \sigma \sqrt{\Delta t} \cdot \frac{n(n+1)}{6} \hat{Z} \right) \right\}$$

$$= S_0 \cdot \exp \left\{ \frac{n(n+1)}{2n} T + \frac{\sigma \sqrt{T}}{n} \cdot \sqrt{\frac{n(n+1)(2n+1)}{6}} \hat{Z} \right\}$$

Recall BSM-Call ( $S_0, K, 0, T, r, \sigma$ )

$$\text{If } S_t = S_0 \cdot \exp \left\{ (r - \frac{1}{2}\sigma^2)T + \sigma \sqrt{T} \hat{Z} \right\}, \text{ then } \mathbb{E} \left[ e^{-rT} (S_T - K)^+ \right] = \text{BSM-Call } (S_0, K, 0, T, r, \sigma)$$

Now I want  $\Pi_0^C = \mathbb{E} [e^{-rT} (A_T - k)^+]$  where  $\begin{cases} A_T = S_0 \cdot \exp \left\{ (\hat{r} - \frac{1}{2} \hat{\sigma}^2) T + \hat{\sigma} \sqrt{T} \hat{Z} \right\} \\ \hat{\sigma} = \frac{\sigma}{n} \cdot \sqrt{\frac{(n+1)(2n+1)}{6}} \\ \hat{r} - \frac{1}{2} \hat{\sigma}^2 = \frac{(n+1)\mu}{2n} \end{cases}$

$$\therefore \Pi_0^C = e^{\hat{r}T - rT} \mathbb{E} [e^{-\hat{r}T} (A_T - k)^+] = e^{(\hat{r}-r)T} \cdot \text{BSM-Call}(S_0, k, 0, T, \hat{r}, \hat{\sigma})$$

## L05 Approximation to BSM Option Valuation by CRR

We denote CRR Model by  $M^{CRR} = \{\Omega, \mathcal{F}, \mathbb{P}, \mathcal{A}\}, T, (S, B)\}$

the time horizon is  $t \in \{0, \Delta t, \dots, T\} := \bar{T}$

Bond price follows binomial tree:  $B_t = e^{-r(t-\Delta t)}$

Stock price follows  $S_{t+\Delta t} = S_t \cdot m_t$

where  $m_t \in \{u, d\}$  is a random variable with 2 possible values with:

$$u = e^{\sqrt{\Delta t}} \quad d = e^{-\sqrt{\Delta t}}$$

[A1]  $\sigma > r$ , under [A1], there exists unique risk-neutral probability  $\alpha$  given by:

$$\alpha = \alpha(m_t=u) = \frac{e^{\sigma \sqrt{\Delta t}} - d}{u - d}$$

$$1 - \alpha = \alpha(m_t=d) = \frac{u - e^{\sigma \sqrt{\Delta t}}}{u - d}$$

[Q] if [A1] is violated, then does it still make the model arbitrage free?

CRR call value can be evaluated backwardly from

$$C_0 = e^{-rt} E^\alpha_0(G)$$

L06

$$W(t_{i+1}) = W(t_i) + \sqrt{t_{i+1} - t_i} Z_{i+1}$$

$$\Rightarrow \hat{W}(t_{i+1}) = \hat{W}(t_i) + \frac{1}{\sqrt{n}} Z_{i+1}, \text{ for } i=0, 1, \dots, n-1$$

Generating GBM path.

GBM is given by

$$X(t) = X_0 \cdot \exp \left\{ (r - \frac{1}{2}\sigma^2)t + \sigma W(t) \right\}$$

$$\Rightarrow \hat{X}(t) = X_0 \cdot \exp \left\{ (r - \frac{1}{2}\sigma^2)t + \sigma \hat{W}(t) \right\}$$

Application to Arithmetic Asian option price.

call option price:

$$C(T) = (A(T) - K)^+$$

$$\Rightarrow C_0 = \mathbb{E} \left[ e^{-rT} (A(T) - K)^+ \right]$$

$$A(T) = \frac{1}{n} \sum_{i=1}^n S(t_i)$$

## Oct 04 § Binomial option pricing

Recall.

BSM.  $S_0 = 100$ ,  $\sigma = 15\%$ , rate =  $r$

Then it implies stock under EMM &

$$\frac{dS_t}{S_t} = r dt + \sigma dW_t$$

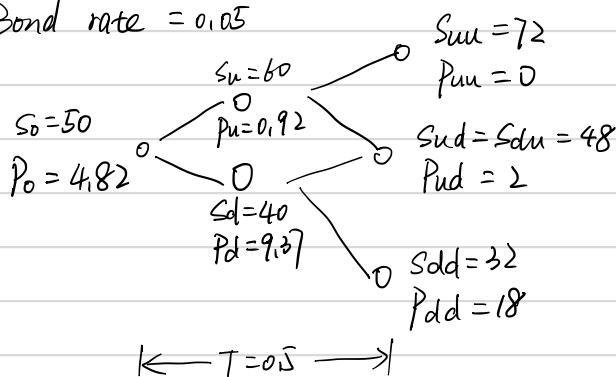
$$\text{or } S_t = S_0 \cdot \exp \left\{ (r - \frac{1}{2}\sigma^2)t + \sigma W_t \right\}.$$

We consider Binomial Tree model, denoted by

★ Bin Tree ( $S_0 = 50$ ,  $N=2$ ,  $T=0.5$ ,  $u=1.2$ ,  $d=0.8$ ,  $r=0.05$ ,  $K=50$ )

① Bond rate = 0.05

②



\* option types: ① EuCall ( $T=0.5$ ,  $K=50$ ) Payoff  $|_T = (S_T - K)^+$

② EuPut ( $T$ ,  $K$ ) Payoff  $|_T = (K - S_T)^-$

③ AmCall ( $T$ ,  $K$ ) Payoff  $|_T = (S_T - K)^+, \quad t \leq T$

④ AmPut ( $T$ ,  $K$ ) Payoff  $|_T = (K - S_T)^-, \quad t \leq T$

Q. St  $\frac{q}{u-d}$  vs  $\frac{d}{u-d}$  what's the EMM?  $St \cdot e^{r_{st}} = q \cdot u_{st} + (1-q) d_{st}$

$$\Rightarrow q = \frac{e^{r_{st}} - d}{u - d}$$

Recall Martingale

$\{X_t : t \in T\}$  is a mtgl process if  $E[X_{t+h} | X_t] = X_t$

Recall mtgl (discrete)

$\{X_t : t \in N\}$  is a mtgl process if  $E[X_{t+1} | X_t] = X_t$

ex1  $X_0 = 100$   $\frac{q=1}{1-q=\frac{1}{2}}$   $1000 = X_1(u)$

① If  $Q = (q, 1-q) = (\frac{1}{2}, \frac{1}{2})$ , then it's not mtgl. b/c

$$E[X_1] = 1000 \cdot \frac{1}{2} + 99 \cdot \frac{1}{2} = \frac{1099}{2} \neq X_0$$

② Is there mtgl measure? i.e. Find Q s.t.  $E^Q(X_1) = X_0$

$$1000 q + 99 (1-q) = 100 \Rightarrow q = \frac{1}{99}$$

ex2  $X_0 = 100$   $\frac{q}{1-q}$   $1000 = X_1(u)$

Find mtgl measure?

$$\frac{q}{1-q} = X_1(d)$$

$$\text{soln: } E(X_1) = 1000q + 101(1-q) = 100 \Rightarrow q \text{ does not exist.}$$

ex3  $X_0 = 100$   $\frac{q_1}{q_2}$   $101 = X_1(u)$  Find mtgl measure?

$$\frac{q_1}{q_2} = 100 \quad \text{soln: } 101q_1 + 100q_2 + 99(1-q_1 - q_2) = 100 \Rightarrow 2q_1 + q_2 = 1$$

$$\frac{1-q_1}{1-q_2} = X_1(d)$$

$$(Q = (0, 1, 0) \text{ or } (\frac{1}{2}, 0, \frac{1}{2}) \text{ or } (q_1, q_2, 1))$$

$$q_i \in [0, \frac{1}{2}]$$

Def  $T = \{0, \Delta t, 2\Delta t, \dots, (n-1)\Delta t, n\Delta t = T\} = \{n\Delta t : n=0, 1, \dots, n\}$

Stock price:  $S_{n\Delta t}$

Bond rate:  $r$

$Q$  is EMM if  $E[S_{(n+1)\Delta t} | S_{n\Delta t}] = e^{r\Delta t} S_{n\Delta t}$  or  $S_{n\Delta t} = \underbrace{E[e^{-r\Delta t} S_{(n+1)\Delta t} | S_{n\Delta t}]}_{\text{discounted Stock Price}}$

R.K.

Discounted stock price is a mtgl

$$S_t \frac{q}{u-d} u_{st} \text{ mtgl: } q = \frac{e^{r\Delta t} - d}{u - d} \quad \Delta t = \frac{T}{n}$$

constructing BinEuOpt

set-up-parameters:  $M = N + 1$  # of terminal nodes.

$$u = 1 + p_u \quad d = 1 - p_d \quad \left. \begin{array}{l} \# \text{up/down factor} \end{array} \right.$$

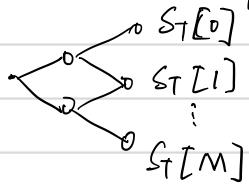
$$z_u = (e^{r\Delta t} - d) / (u - d)$$

$$z_d = 1 - z_u$$

$$\Delta t = \frac{T}{N} \quad \# \text{Time length of one period}$$

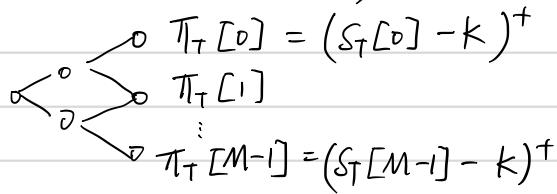
$$df = e^{-r\Delta t} \quad \# \text{discount factor}$$

Init - stock - tree: compute:  $\{S_T[i], i=0, 1, \dots, M-1\}$



Init - payoff - tree:

ex: EuCall ( $T=0.5, K=30$ )



Traverse - tree (Backward)

$$\Pi_N[i] \rightarrow \Pi_{N-1}[i] \rightarrow \Pi_{N-2}[i] \dots \rightarrow \Pi_0[0]$$

$$\Pi_{i+1}[j] = e^{-rt} \left( \varrho \Pi_i[j] + (1-\varrho) \Pi_i[j+1] \right)$$

Q. Given BSM( $r$ ), rate= $r$ , we have BSM formula for pricing EuCall( $T, K$ ), EuPut( $T, K$ )

Arithmetic Asian option } has no formula  
Barrier option }

$\star$  CRR ( $S_0=50, N=2, T=0.5, \sigma=0.3, r=0.05$ ) = Bin Tree ( $S_0=50, N=2, T=0.5, u=e^{\sqrt{rt}}, d=\frac{1}{u}, r=0.05$ )  
 $\xrightarrow{N \rightarrow \infty}$  BSM ( $S_0=50, T=0.5, \sigma=0.3, r=0.05$ ) where  $\Delta t = T/N$

$$S_{t+\Delta t}(u) = e^{\sqrt{\Delta t}} S_t$$

$$S_{t+\Delta t}(d) = e^{-\sqrt{\Delta t}} S_t$$

Ex: what is price of Put ( $T=0.5, K=30$ )

Q1: for BSM ( $S_0=50, r=0.05, \sigma=0.3$ ) = ?

Q2: for CRR ( $S_0=50, N=200, T=0.5, \sigma=0.3, r=0.05$ ) = ?



