

Introduction to Statistics and Data Visualisation with R

Lausanne, January 2026

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Multiple Regression



What happens if both,
age and weight variables
were included in the same model ?

One multiple regression with two variables

Call:

```
lm(formula = Height ~ Age + Weight, data = class)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.6248	-1.3016	-0.0176	0.8324	4.1019

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	132.1943	5.0823	26.011	1.61e-14 ***
Age	1.2267	0.5302	2.314	0.03431 *
Weight	0.2761	0.0695	3.973	0.00109 **

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 2.255 on 16 degrees of freedom

Multiple R-squared: 0.828, Adjusted R-squared: 0.8065

F-statistic: 38.52 on 2 and 16 DF, p-value: 7.646e-07

This model allows us to determine the respective contribution of each variable separately.

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	132.1943	5.0823	26.011	1.61e-14	***
Age	1.2267	0.5302	2.314	0.03431	*
Weight	0.2761	0.0695	3.973	0.00109	**

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1					

This is similar to the simple regression case.

Each test is conducted assuming that the tested parameter is the last one entering the model:

« If *weight* is already in the model, is the coefficient for *age* significantly different from 0 ? »

Two single regressions vs one multiple regression

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	142.57014	2.67989	53.200	< 2e-16 ***
Weight	0.39523	0.05231	7.555	7.89e-07 ***

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	125.2239	6.5217	19.201	5.82e-13 ***
Age	2.7871	0.4869	5.724	2.48e-05 ***

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	132.1943	5.0823	26.011	1.61e-14 ***
Age	1.2267	0.5302	2.314	0.03431 *
Weight	0.2761	0.0695	3.973	0.00109 **

While both age and weight seem significant by themselves, age is much less significant when weight is already included (see also the R²).

It is likely that a lot of the information provided by the age is also provided by the weight, so that there may be little need to have both terms in the model.

Multiple and adjusted R-squared

Multiple R-squared: 0.828,

Adjusted R-squared: 0.8065

As before, R^2 is the proportion of the total variance in the response data that is explained by the model.

Adding a new variable in the model will always increase R^2 , up to 1 when there the number of degrees of freedom is 0 (number of parameters to estimate = number of observations).

Multiple and adjusted R-squared

Multiple R-squared: 0.828,

Adjusted R-squared: 0.8065

The adjusted R-squared adjusts for the number of variables in the model, and does not necessarily increase when the number of variables increase; it can even be negative.

$$R^2 = 1 - \frac{SS_{\text{residuals}}}{SS_{\text{total}}}$$

It is always equal or below R^2 .

$$\text{Adjusted } R^2 = 1 - \frac{SS_{\text{residuals}} / (n - K)}{SS_{\text{total}} / (n - 1)}$$

Example

```
y <- rnorm(10)  
x1 <- rnorm(10); x2 <- rnorm(10); ... ; x9 <-  
rnorm(10)  
summary(lm(y ~ x1)); summary(lm(y ~ x1+x2));  
  
1: Multiple R-squared:  0.1419,    Adjusted R-squared:  0.03464  
2: Multiple R-squared:  0.5173,    Adjusted R-squared:  0.3794  
3: Multiple R-squared:  0.557 ,    Adjusted R-squared:  0.3355  
4: Multiple R-squared:  0.5577,    Adjusted R-squared:  0.2039  
5: Multiple R-squared:  0.7953,    Adjusted R-squared:  0.5395  
6: Multiple R-squared:  0.8321,    Adjusted R-squared:  0.4962  
7: Multiple R-squared:  0.984 ,    Adjusted R-squared:  0.9281  
8: Multiple R-squared:  0.9851,    Adjusted R-squared:  0.866  
9: Multiple R-squared:       1 ,    Adjusted R-squared:      NaN
```

The last regression from the example

Call:

```
lm(formula = y ~ x1 + x2 + x3 + x4 + x5 + x6 + x7 + x8 + x9)
```

Residuals:

ALL 10 residuals are 0: no residual degrees of freedom!

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-0.02693	NA	NA	NA
x1	0.53886	NA	NA	NA
x2	-0.52227	NA	NA	NA
x3	0.51881	NA	NA	NA
x4	0.74757	NA	NA	NA
x5	0.14394	NA	NA	NA
x6	-0.65387	NA	NA	NA
x7	-0.48271	NA	NA	NA
x8	-0.62487	NA	NA	NA
x9	0.23759	NA	NA	NA

Residual standard error: NaN on 0 degrees of freedom

Multiple R-squared: 1, Adjusted R-squared: NaN

F-statistic: NaN on 9 and 0 DF, p-value: NA

F-statistic for significance of regression

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	81.77355	12.90896	6.335	9.92e-06	***
Age	3.11575	1.34668	2.314	0.03431	*
Weight	0.35064	0.08827	3.973	0.00109	**

F-statistic: 38.52 on 2 and 16 DF, p-value: 7.646e-07

Again, the F-statistic allows us to test if the whole regression (adding all variables vs having only the intercept in) is significant.

If any of the tests for the individual variables is significant, the F-test will generally be significant as well.

However, even if no individual variable is significant (e.g. $p < 0.05$), the F-test can still be significant.

Categorical variables, dummy variables and contrasts

Categorical variables

We'd like to use categorical variables in a linear model, as in:

$$\text{Height} = \mathbf{b_0} + \mathbf{b_1} \text{Age} + \mathbf{b_2} \ll \text{Gender} \gg + \text{error}$$

Intuitively, we want to estimate a « Male » and a « Female » effect.

Categorical variables

We'd like to use categorical variables in a linear model, as in:

$$\text{Height} = b_0 + b_1 \text{Age} + b_2 \text{« Gender »} + \text{error}$$

Intuitively, we want to estimate a « Male » and a « Female » effect.

In practice, categorical variables (factors in R) are turned (by default, based on alphabetical order) into **dummy variables** of the form

$$\text{Gender} = \begin{cases} 1 & \text{if Female} \\ 2 & \text{if Male} \end{cases}$$

Example of summary results of the lm command in R

Call:

```
lm(formula = Height ~ Age + Gender, data = class)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.483	-1.910	-0.319	1.326	5.317

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)		
(Intercept)	124.5241	5.8886	21.147	4.04e-13 ***		
Age	2.7276	0.4398	6.202	1.27e-05 ***		
GenderM	2.8362	1.2797	2.216	0.0415 *		

Signif. codes:	0 ‘***’	0.001 ‘**’	0.01 ‘*’	0.05 ‘.’	0.1 ‘ ’	1

Residual standard error: 2.78 on 16 degrees of freedom

Multiple R-squared: 0.7387, Adjusted R-squared: 0.706

F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

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baseline for
height among
Female

Coefficients:

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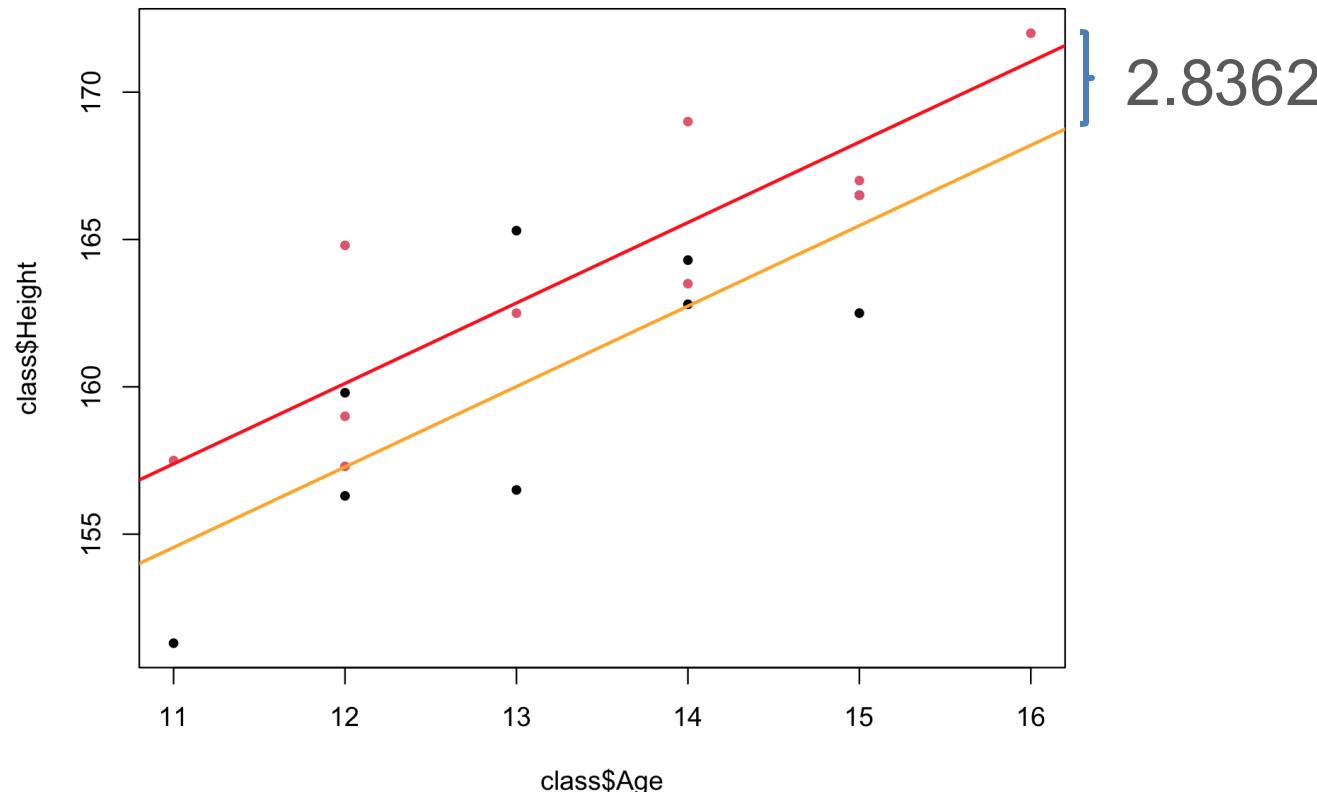
F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

The factor GenderM corresponds to the difference in baseline for Males compared to females

Graphical interpretation

The model specifies 2 straight lines, with the same slope but different y-intercepts:

- For women: Height = $124.52 + 2.72 \text{ Age}$ (in orange)
For men: Height = $127.3 + 2.72 \text{ Age}$ (in red)



What if we don't use a linear model ?

We could also compute the difference in means between males and females directly:

```
> tapply(class$Height, class$Gender, mean)
      F          M
 160.5889 163.9100
> means <- tapply(class$Height, class$Gender, mean)
> diff(means)
      M
 3.321111
```

This result is slightly different from the 2.8362 cm difference found with the linear model.

Where does the difference come from ?

Interactions

So far, we have assumed a difference between the lines, but the same slope; that is, for both men and women, the effect of age is the same.

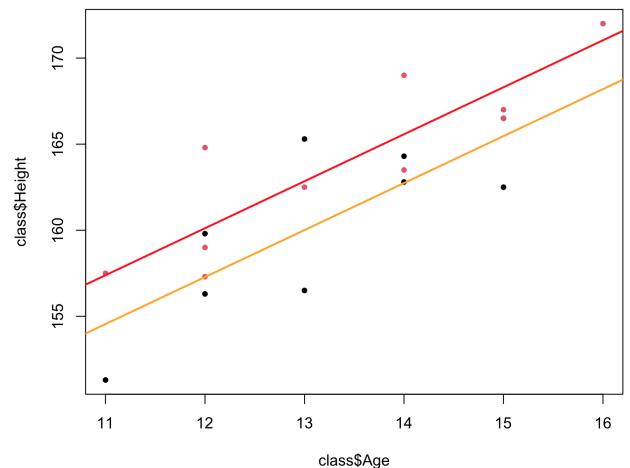
If this assumption is incorrect, it means that there is an *interaction* between the factors « age » and « gender », that is, the effect of age is different depending on the gender.

Interactions are modeled in R in the following way:

`lm(formula = Height ~ Age + Gender + Age:Gender)`

which is equivalent to

`lm(formula = Height ~ Age * Gender)`



Coefficients with an interaction

Call:

```
lm(formula = Height ~ Age * Gender, data = class)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.4429	-1.7844	-0.3648	1.3730	5.3571

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	122.1500	9.6409	12.670	2.05e-09 ***
Age	2.9071	0.7256	4.007	0.00114 **
GenderM	6.7443	12.4109	0.543	0.59483
Age:GenderM	-0.2940	0.9285	-0.317	0.75585

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 2.862 on 15 degrees of freedom

Multiple R-squared: 0.7404, Adjusted R-squared: 0.6885

F-statistic: 14.26 on 3 and 15 DF, p-value: 0.0001152

The coefficients can be interpreted as follows:

According to the model, the *height* is equal to

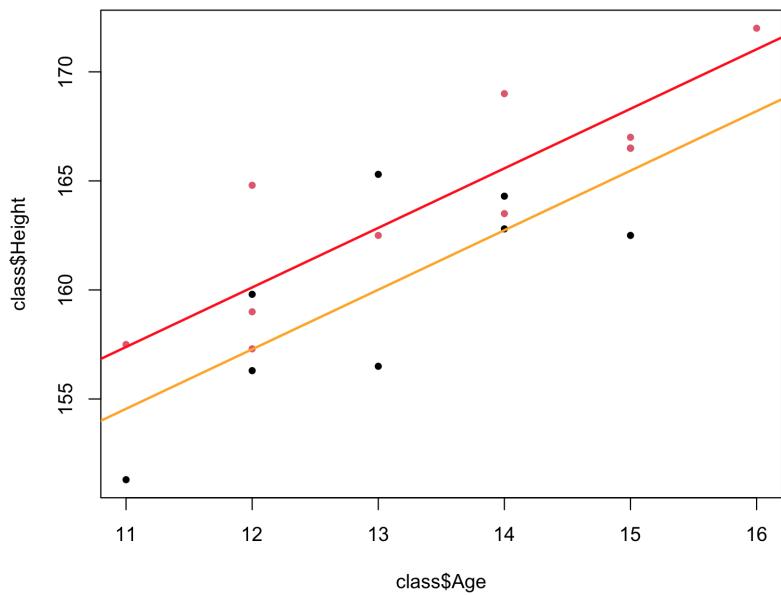
122.15 (the intercept)

plus 2.9071 times the person's age

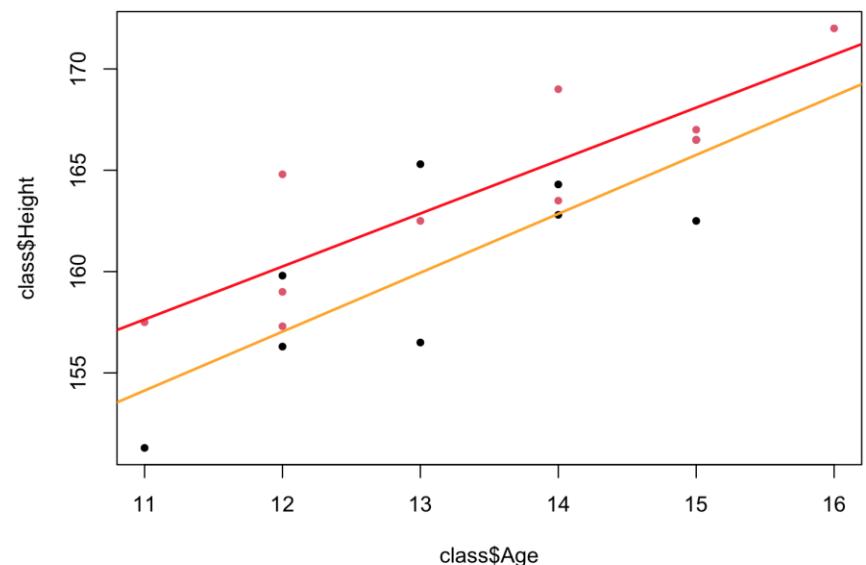
plus 6.7443, but only for males

-0.2940 times the person's age, but only for males.

Different slopes



No interaction



With interaction

```
> model <- lm( Height ~ Age+Gender1, data=class)
> summary(model)
```

Call:

```
lm(formula = Height ~ Age + Gender1, data = class)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.483	-1.910	-0.319	1.326	5.317

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	127.3603	5.9587	21.374	3.43e-13 ***
Age	2.7276	0.4398	6.202	1.27e-05 ***
Gender1F	-2.8362	1.2797	-2.216	0.0415 *

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 2.78 on 16 degrees of freedom
Multiple R-squared: 0.7387, Adjusted R-squared: 0.706
F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

```
> model <- lm( Height ~ Age+Gender, data=class)
> summary(model)
```

Call:

```
lm(formula = Height ~ Age + Gender, data = class)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.483	-1.910	-0.319	1.326	5.317

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	124.5241	5.8886	21.147	4.04e-13 ***
Age	2.7276	0.4398	6.202	1.27e-05 ***
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Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 2.78 on 16 degrees of freedom
Multiple R-squared: 0.7387, Adjusted R-squared: 0.706
F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

What if Males were the baseline ?

The two models are exactly the same; only the way we look at the coefficient changes.

```
Gender1 <- relevel(Gender, ref="M")
```

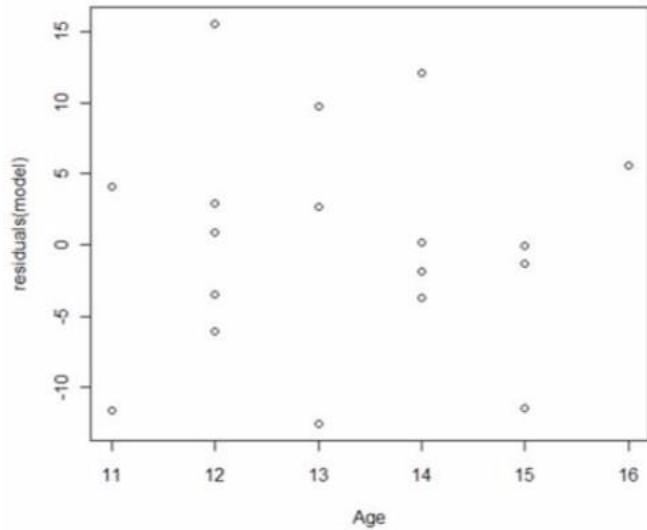
Diagnostic tools

**It is always possible to fit a linear model and find a slope and intercept
... but it does not mean that the model is meaningful !**

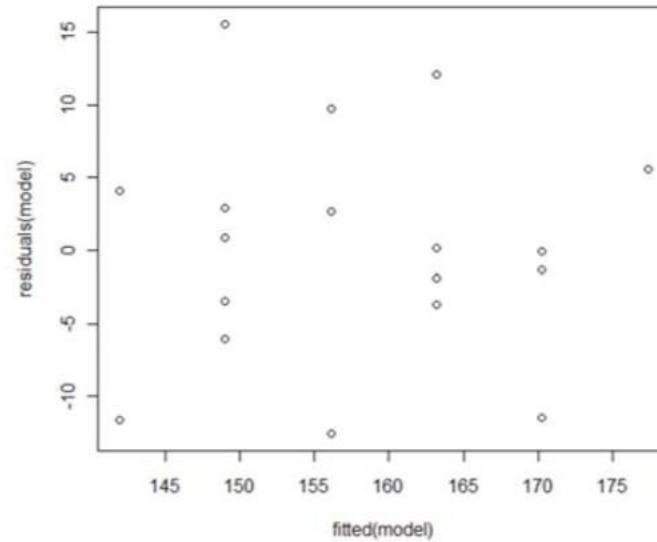
Examination of *residuals*: (which should show no obvious trend, since any systematic effect in the residuals should ideally be captured by the model):

- Normality
- Time effects
- Nonconstant variance – Curvature

Examination of *residuals*



```
plot( Age, residuals(model) )
```



```
plot( fitted(model), residuals(model) )
```

**Works only for simple regression
(only one variable on x axis)**

Works also for multiple regression

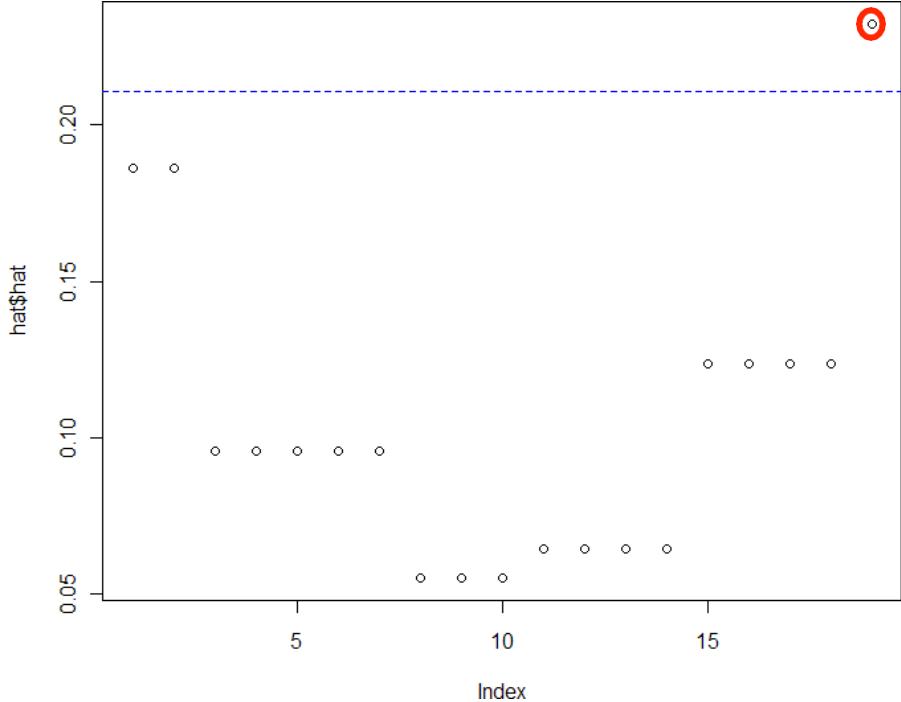
High leverage ('influential') points are far from the center, and have potentially greater influence

One way to assess points is through the *hat values* (obtained from the *hat matrix H*):

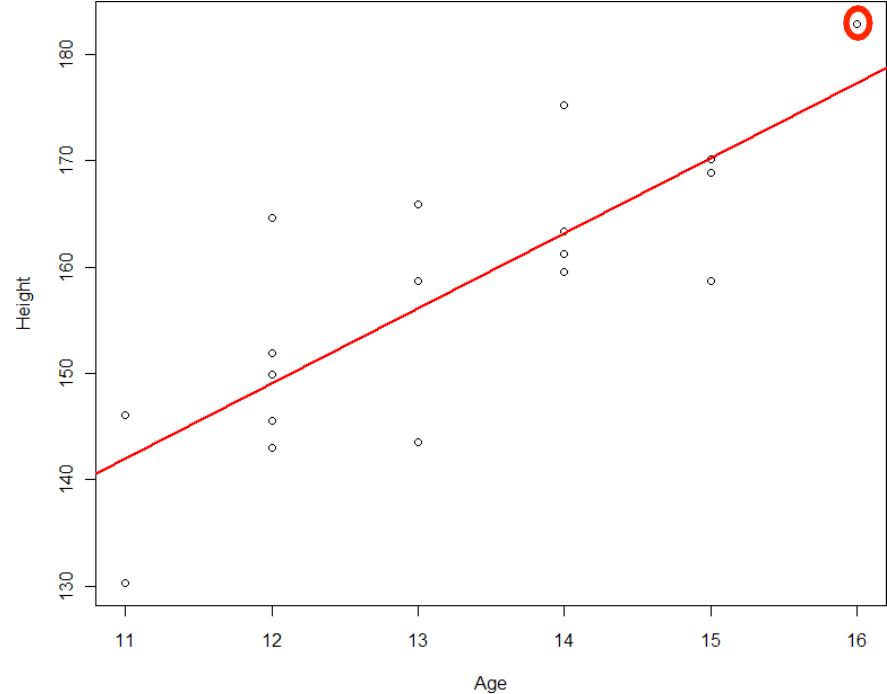
$$\begin{aligned}\hat{y} &= Xb = X(X'X)^{-1}X'y = Hy \\ h_i &= \sum_j h_{ij}\end{aligned}$$

Average value of $h = \text{number of coefficients}/n$
(including the intercept) = p/n

Cutoff typically $2p/n$ or $3p/n$

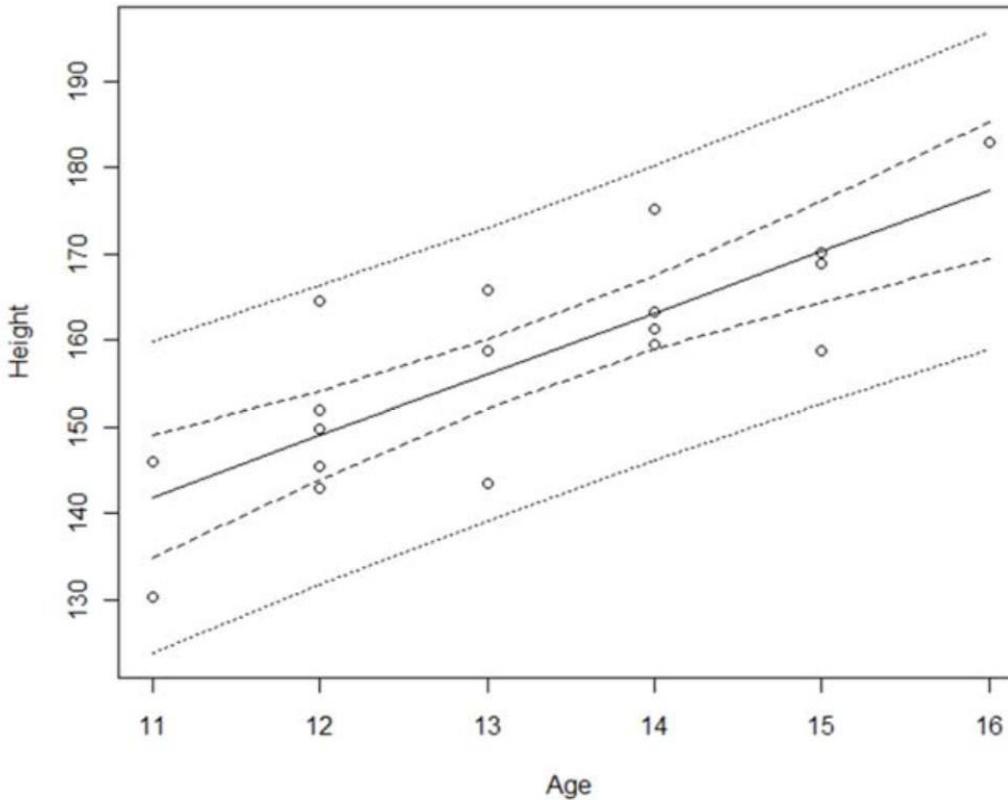


Hat values



Actual fit

```
>hat <- lm.influence( model )
>plot( hat$hat )
>abline(h=c(2,3)*2/19),lty=c(2,3),col=c("blue","red") )
```



Narrow bands: describe the uncertainty about the regression line

Wide bands: describe where most (95% by default) predictions would fall, assuming normality and constant variance.

In R: `?predict.lm`