

14.

$$a) g \sim p(x, \theta) = \frac{1}{\theta} \{ [0; 2\theta] \} \quad \theta \in [0; 2\theta]$$

$$\vec{x}_n = \text{выборка} \quad F(x) = \frac{x}{\theta} - 1$$

Метод моментов:

$$\mu_k = M[g^k] = \int_{-\infty}^{2\theta} x^k p(x, \theta) dx \quad \bar{\mu}_k = \frac{1}{n} \sum_{i=1}^n x_i^k$$

$$M[g] = \int_{\theta}^{2\theta} x \frac{1}{\theta} dx = \frac{3}{2} \theta \quad \bar{\mu}_1 = \bar{x} \quad \frac{3}{2} \theta = \bar{x} \Rightarrow \hat{\theta}_1 = \frac{2}{3} \bar{x}$$

Метод максимального правдоподобия:

$$L(x, \theta) = \prod_{i=1}^n p(x_i; \theta) = \left( \frac{1}{\theta} \right)^n \quad \ln L = n \ln(\theta)$$

$$\frac{\partial \ln L}{\partial \theta} = \frac{-n}{\theta} \quad \sup \text{ на } [0; 2\theta] \text{ при } \theta \rightarrow \theta_{\max}$$

$$\begin{cases} 2\theta \geq x_{\max} \\ \theta \geq x_{\max} \end{cases} \rightarrow \frac{1}{2} x_{\max} \leq \theta \leq x_{\min} \rightarrow \sup \text{ при } \hat{\theta}_2 = \frac{1}{2} x_{\max}$$

$$b) M[\hat{\theta}_1] = M\left[\frac{2}{3n} \sum_{i=1}^n x_i\right] = \frac{2}{3} M[g] = \frac{2}{3} \cdot \frac{3}{2} \theta = \theta - \text{несмещен}$$

$$D[\hat{\theta}_1] = \frac{4}{9n^2} D[g] \quad M[g^2] = \int_{\theta}^{2\theta} x^2 \frac{1}{\theta} dx = \frac{7\theta^2}{3}$$

$$D[g] = \left( \frac{7}{3} \theta^2 - \frac{9}{4} \theta^2 \right) = \frac{\theta^2}{12}$$

$$D[\hat{\theta}_1] = \frac{4}{9n^2} \cdot n \cdot \frac{\theta^2}{12} = \frac{\theta^2}{27n} \xrightarrow{n \rightarrow \infty} 0 - \text{усреднение}$$

$$M[\hat{\theta}_2] = M\left[\frac{1}{2} x_{\max}\right] = \frac{1}{2} \int_{\theta}^{2\theta} x \cdot n \left( \frac{x}{\theta} - 1 \right)^{n-1} \frac{1}{\theta} dx =$$

$$= \frac{1}{2} \int_{\theta}^{2\theta} \left( t - \frac{x}{\theta} + 1 \right) = \theta \int_0^1 n(t+1) \cdot t^{n-1} dt = \frac{2n+1}{2n+2} \theta$$

$$q(y) = n(F(y))^{n-1} F'(y) = n \left( \frac{x}{\theta} - 1 \right)^{n-1} \frac{1}{\theta} \text{ усреднение}$$



$$\tilde{\theta}^* = \frac{1}{2} x_{\max} \cdot \frac{2n+2}{2n+1} - \frac{n+1}{2n+1} x_{\max} - \text{не минимизирует}$$

$$M[x_{\max}] = \int_0^{2\theta} x n \left( \frac{x}{\theta} - 1 \right)^{n-1} \frac{1}{\theta} dx = \frac{2n+1}{n+1} \theta$$

$$M[x_{\max}^2] = \int_0^{2\theta} x^2 n \left( \frac{x}{\theta} - 1 \right)^{n-1} \frac{1}{\theta} dx = \frac{n(n^2+8n+2)}{n(n+1)(n+2)} \theta^2$$

$$D[\tilde{\theta}_2^*] = \frac{(n+1)^2}{(2n+1)^2} \cdot D[x_{\max}] = \frac{(n+1)^2}{(2n+1)^2} \left[ \frac{4n^2+8n+2}{(n+1)(n+2)} - \frac{(2n+1)^2}{(n+1)^2} \right] \theta^2$$

$$= \frac{\theta^2 n}{4n^3+12n^2+9n+2} = \frac{\theta^2 n}{(2n+1)^2(n+2)} \xrightarrow{n \rightarrow \infty} 0 \quad \tilde{\theta}_2^* - \text{состоит.}$$

$$c) I(\theta) = M\left[\left(\frac{1}{\theta}\right)^2\right] = \frac{1}{\theta^2} > 0 \quad \forall \theta \in [\theta; 2\theta]$$

$$g'(\theta) = -\frac{1}{\theta^2} \quad \text{f. вып. гуга по } \theta \text{ на } \Pi$$

$$0 = \int_0^{2\theta} \frac{\partial}{\partial \theta} g(x, \theta) dx = \int_0^{2\theta} -\frac{1}{\theta^2} dx = -\frac{2}{\theta} \neq 0 \quad \text{неограниченно}$$

$$n D[\tilde{\theta}_1] = \frac{\theta^2}{2n} \xrightarrow{n \rightarrow \infty} \frac{\theta^2}{2n}$$

$$n D[\tilde{\theta}_2^*] = \frac{\theta^2 n^2}{(2n+1)^2(n+2)} \xrightarrow{n \rightarrow \infty} 0$$

$$d) P(t_1 < f(x_n; \theta) < t_2) = \beta$$

$$t_2 - t_1 \rightarrow \min$$

$$f(x_n, \theta) = \frac{x_{\max}}{\theta}$$

$$c) F_f(x) = (F(\theta; x))^n = (x-1)^n$$

$$t_1 = q \sqrt{\frac{1-p}{2}}$$

$$t_2 = q \sqrt{\frac{1+p}{2}}$$

$$t_1 < \frac{x_{\max}}{\theta} < t_2$$



$$\frac{x_{\max}}{t_1} < \theta < \frac{x_{\max}}{t_2}$$

$$t_2 = \sqrt{\frac{1+\beta}{2}} + 1 \quad t_1 = \sqrt{\frac{1-\beta}{2}} + 1$$

$t_2 = \theta < t_1$  - точный гвер. интер.

е) По ОШП не можем (можем не переписать)

Ищем по ОШП

$$\sqrt{n} \frac{g(\bar{x}) - g(d)}{G(d)} \sim N(0, 1)$$

$$f(\bar{x}) = \hat{\theta} = \frac{2}{3} \bar{x} \quad \text{и } f = \frac{2}{3}$$

$$f(d) = \theta = \frac{2}{3} d_1 \quad d_1 = \bar{x} \quad d_2 = \bar{x}^2$$

$$K = \bar{x}^2 - \bar{x}^2 \quad G(d) = \frac{2}{3} \sqrt{\bar{x}^2 - \bar{x}^2}$$

$$\frac{\hat{\theta} - \theta}{\frac{2}{3} \sqrt{\bar{x}^2 - \bar{x}^2}} \sim N(0, 1) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

$$u_{\frac{1-\beta}{2}} = -u_{\frac{\beta}{2}}$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{u_{\frac{1-\beta}{2}}} e^{-\frac{x^2}{2}} dx = \frac{1-\beta}{2} = \frac{1}{2} \left( \text{erf}\left(\frac{u_2}{u_1}\right) + 1 \right) = \frac{1-\beta}{2}$$

$$\text{erf}\left(\frac{u_{\frac{1-\beta}{2}}}{\sqrt{2}}\right) = -\beta \rightarrow \text{erf}^{-1}(-\beta) = \frac{u_{\frac{1-\beta}{2}}}{\sqrt{2}}$$

$$- \text{erf}^{-1}(-\beta) = \sqrt{n} \frac{\hat{\theta} - \theta}{\frac{2}{3} \sqrt{\bar{x}^2 - \bar{x}^2}} > \sqrt{2} \text{erf}^{-1}(-\beta) \quad \hat{\theta} = \frac{2}{3} \bar{x}$$

$$- \frac{1}{\sqrt{n}} \text{erf}^{-1}(-\beta) \cdot \frac{2}{3} \sqrt{\bar{x}^2 - \bar{x}^2} + \frac{2}{3} \bar{x} > \theta > \frac{\sqrt{2}}{\sqrt{n}} \text{erf}^{-1}(-\beta) \cdot \frac{2}{3} \sqrt{\bar{x}^2 - \bar{x}^2} + \frac{2}{3} \bar{x}$$

асимпт. гвер. интервал.



TS.

a) MLE MLE

$$L(x, \theta) = \prod_{i=1}^n p(x_i, \theta) = \frac{(\theta-1)^n}{\prod_{i=1}^n x_i^\theta} \quad \ln L = n \ln(\theta-1) - \theta \sum_{i=1}^n \ln x_i$$

$$\frac{\partial \ln L}{\partial \theta} = \frac{n}{\theta-1} - \sum_{i=1}^n \ln x_i = 0 \quad \sum_{i=1}^n \ln x_i = \frac{n}{\theta-1}$$

$$\hat{\theta} = \frac{n}{\sum_{i=1}^n \ln x_i} + 1$$

b) med - median

$$\int_{-\infty}^{\infty} p(x) dx = \frac{1}{2}$$

$$\int_0^{\text{med}} (\theta-1) \frac{1}{x^\theta} dx = -x^{-\theta+1} \Big|_0^{\text{med}} = \frac{1}{2}$$

$$\text{med} = 2^{\frac{1}{\theta-1}}$$

$$\frac{f(\hat{\theta}) - f(\theta)}{G(\theta)} \sqrt{n} \sim N(0, 1)$$

$$f(\theta) = 2^{\frac{1}{\theta-1}} \rightarrow \nabla f = 2^{\frac{1}{\theta-1}} \ln 2 \cdot \left( \frac{1}{(\theta-1)^2} \right)$$

$$f(\hat{\theta}) = 2^{\frac{1}{\hat{\theta}-1}} \quad \hat{\theta} = \frac{n}{\sum_{i=1}^n \ln x_i} + 1$$

$$I(\theta) = \mathbb{E} \left[ \left( \frac{\partial \ln p}{\partial \theta} \right)^2 \right] = \int_0^{\infty} \left( \frac{1}{\theta-1} - \ln x \right)^2 \frac{\theta-1}{x^\theta} dx = \frac{1}{(\theta-1)^2}$$

$$G(\theta) = \left[ \frac{1}{2^{\theta-1}} \ln 2 \cdot \left( \frac{1}{(\hat{\theta}-1)^2} \right) \cdot (\hat{\theta}-1)^2 \cdot 2^{\frac{1}{\hat{\theta}-1}} \ln 2 \cdot \frac{1}{(\hat{\theta}-1)^2} \right]^{\frac{1}{2}} \\ = 2^{\frac{1}{\theta-1}} \ln 2 \cdot \frac{1}{(\theta-1)^2}$$



$$-\sqrt{2} \operatorname{erf}^{-1}(-\beta) > \frac{2^{\frac{1}{\theta-1}} - \operatorname{med}}{2^{\frac{1}{\theta-1}} \ln 2 \cdot \frac{1}{(\theta-1)^2}} \sqrt{n} > \sqrt{2} \operatorname{erf}^{-1}(-\beta)$$

$$\frac{\sqrt{2}}{\sqrt{n}} \operatorname{erf}^{-1}(-\beta) \cdot 2^{\frac{1}{\theta-1}} \cdot \ln 2 \cdot \frac{1}{(\theta-1)^2} = EP$$

$$EP + \frac{1}{2^{\theta-1}} \leftarrow \operatorname{med} \leftarrow -EP + \frac{1}{2^{\theta-1}} \quad \text{где } \tilde{\theta} = \frac{n}{\sum_{i=1}^n \ln x_i} + 1$$

Проверим, что модель регуляриза :  $\rho$  непрерывна  
по  $\theta$  в  $\Pi$

$$0 = \int_1^{\infty} \frac{\partial}{\partial \theta} \rho(x, \theta) dx = \int_1^{\infty} \frac{x^{\theta} - x^{\theta} \ln x (\theta - 1)}{x^{2\theta}} dx = x^{1-\theta} \ln \Big|_1^{\infty} = 0$$

модель регуляриза, все хорошо!

а) По ОМП  $f(\theta) = \theta \quad \forall f = 1$

$$f(\tilde{\theta}) = \tilde{\theta} \quad I^{-1} = (\theta - 1)^2 \quad G = \theta - 1$$

$$G(\tilde{\theta}) = \tilde{\theta} - 1$$

$$-\sqrt{2} \operatorname{erf}^{-1}(-\beta) > \frac{\tilde{\theta} - \theta}{\tilde{\theta} - 1} \sqrt{n} > \sqrt{2} \operatorname{erf}^{-1}(-\beta)$$

$$\frac{\sqrt{2}}{\sqrt{n}} (\tilde{\theta} - 1) \operatorname{erf}^{-1}(-\beta) + \tilde{\theta} < \theta < -\frac{\sqrt{2}}{\sqrt{n}} \operatorname{erf}^{-1}(-\beta) (\tilde{\theta} - 1) + \tilde{\theta},$$

$$\text{где } \tilde{\theta} = \frac{n}{\sum_{i=1}^n \ln x_i} + 1$$

Асимпт. гом. интер. по ОМП



по Olliv

оценка

$$L_1 = \mathbb{E} \int_1^{\infty} x \frac{\theta-1}{x^\theta} dx = \frac{\theta-1}{\theta-2} \quad \text{где } \theta > 2$$

$$\tilde{L}_1 = \bar{x}$$

$$1 + \frac{1}{\theta-2} = \tilde{x}^3 \rightarrow \tilde{\theta} = \frac{1}{\bar{x}-1} + 2 \quad \text{где } \theta > 2$$

$$f(\tilde{L}_1) = \tilde{\theta} = \frac{1}{\bar{x}-1} + 2 \quad \nabla f(\tilde{L}_1) = \frac{-1}{(\bar{x}-1)^2} \quad \cdot k = \tilde{L}_2 - \tilde{L}_1 = \bar{x}^2 - \bar{x}^2 = 0$$

$$-\sqrt{2} \operatorname{erf}^{-1}(-\beta) > \frac{\frac{1}{\bar{x}-1} + 2 - \theta}{\frac{1}{(\bar{x}-1)^2} \sqrt{(\bar{x}^2 - \bar{x}^2)}} \quad \sqrt{n} > \sqrt{2} \operatorname{erf}^{-1}(-\beta)$$

$$\frac{\sqrt{2}}{\sqrt{n}} \operatorname{erf}^{-1}(-\beta) \frac{\sqrt{\bar{x}^2 - \bar{x}^2}}{(\bar{x}-1)^2} + \frac{1}{\bar{x}-1} + 2 - \theta < -\frac{\sqrt{2}}{\sqrt{n}} \operatorname{erf}^{-1}(-\beta) \frac{\sqrt{\bar{x}^2 - \bar{x}^2}}{(\bar{x}-1)^2} + \frac{1}{\bar{x}-1} + 2$$

асимпт. гом. интер. по Olliv.

Проверим, что модель сепаратизируема

$$0 = \frac{\delta^2}{\partial \theta^2} \int_1^{\infty} p(x, \theta) dx = \int_1^{\infty} \frac{\delta^2}{\partial \theta^2} p(x, \theta) = \int_1^{\infty} \frac{\ln x ((\theta-1) \ln x - 2)}{x^\theta} dx = 0$$

все работает!