

1.(a) Let T be time interval between 1 and 1:30

let r.v. A is time taken by Alice, B is time taken by Bob and C be time taken by Carol to arrive.

$$A \sim (0, T) \quad B \sim (0, T) \quad C \sim (0, T)$$

We want to find

$$P(C < A < T, C < B < T)$$

so integrating joint PDF over given

$$\text{region} = \frac{1}{T^3} \int_0^T \int_C^T \int_C^T da db dc$$

$$= \frac{1}{T^3} \int_0^T \int_C^T (T-c) db dc = \frac{1}{T^3} \int_0^T (T-c)^2 dc$$

$$= -\frac{1}{T^3} \left[\frac{(T-c)^3}{3} \right]_0^T = \frac{1}{T^3} \times \frac{T^3}{3} = \frac{1}{3}$$

(b)

$$C = T/3 \quad \text{and}$$

~~$$(C+10) < B < T$$~~

$$\text{and } (C+10) < A < T$$

$$10 = T/3 \Rightarrow 2T/3 < B < T \quad \text{and} \quad 2T/3 < A < T$$

We want to find $P\left(\frac{2T}{3} < B < T, \frac{2T}{3} < A < T \mid C = T/3\right)$

$$= P\left(\frac{2T}{3} < B < T, \frac{2T}{3} < A < T\right) \quad (\text{as they are independent})$$

Integrating over given region

$$\frac{1}{T^2} \int_{2T/3}^T \int_{2T/3}^T da db = \frac{1}{T^2} \times \frac{T}{3} \times \frac{T}{3} = \frac{1}{9}$$

(C)

At least A or B arrives in the interval $[2T_3, T]$
and other one arrives in $[T_3, T]$

From LOTP and symmetry we want to
find

$$2 P(2T_3 < A < T, T_3 < B < T \mid C = T_3)$$

$$= 2 P(2T_3 < A < T, T_3 < B < T)$$

$$= \frac{2}{T^2} \int_{T_3}^T \int_{2T_3}^T da db = \frac{2}{T^2} \times \frac{2T}{3} \times \frac{T}{3} = \frac{4}{9}$$

2(a) Support of X is $[0, L]$

Given $X=x$, support for Y is $[0, \infty]$

$$f_{X,Y}(x,y) = f_{Y|X}(y|x) f_X(x)$$
$$= \frac{1}{x} \cdot \frac{1}{L} = \frac{1}{Lx}$$

(b) To get marginal PDF of X we integrate joint PDF over support of Y

$$= \int_0^{\infty} \frac{1}{Lx} = \frac{1}{L} \Rightarrow X \sim \text{Unif}(0, L)$$

(c)

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$
$$= \frac{\frac{1}{Lx}}{\frac{1}{L}} = \frac{1}{x}$$
$$\Rightarrow Y|x \sim \text{Unif}(0, \infty)$$

(d) For $Y=y$, X will be in $[y, L]$

$$\Rightarrow f_Y(y) = \int_y^L f_{X,Y}(x,y) dx$$
$$= \frac{1}{L} \int_y^L \frac{1}{x} dx$$
$$= \frac{1}{L} \log_e\left(\frac{L}{y}\right)$$

(e)

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$
$$= \frac{\frac{1}{Lx}}{\frac{1}{L} \log_e\left(\frac{L}{y}\right)}$$
$$= \frac{1}{x \log_e\left(\frac{L}{y}\right)}$$

3.(a) X, Y, Z have Multinomial distribution.

$$\begin{aligned} P(X=i, Y=j, Z=k) &= \frac{5!}{i!j!k!} \left(\frac{1}{13}\right)^i \left(\frac{1}{13}\right)^j \left(\frac{11}{13}\right)^k \\ &= \frac{5!}{i!j!k!} \frac{11^k}{(13)^{i+j+k}} = \frac{5! 11^k}{i!j!k! 13^5} \end{aligned}$$

(b) To get $P(X=i, Y=j)$ we marginalize over all possible $Z=k$. k varies from 0 to 5.

$$\begin{aligned} P(X=i, Y=j) &= \sum_{k=0}^5 \frac{5! 11^k}{i!j!k! 13^5} \\ &= \frac{5!}{i!j!13^5} \sum_{k=0}^5 \frac{11^k}{k!} \approx \frac{5!}{i!j!13^5} e^{11} \end{aligned}$$

(c) i kings can be chosen in $\binom{4}{i}$ ways
 j queen can be chosen in $\binom{4}{j}$ ways
 k other cards can be chosen in $\binom{44}{k}$ ways
 5 cards can be chosen in $\binom{52}{5}$ ways

$$\Rightarrow P(X=i, Y=j, Z=k) = \frac{\binom{4}{i} \binom{4}{j} \binom{44}{k}}{\binom{52}{5}}$$

(4)

Let x be r.v. for total number of matches and I_j be indicator r.v. for j th person match.

$$\Rightarrow x = I_1 + I_2 + \dots + I_n = \sum_{j=1}^n I_j$$

$$\Rightarrow \text{Var}(x) = \text{Var}\left(\sum_{j=1}^n I_j\right)$$

$$= \text{Var}(I_1) + \text{Var}(I_2) + \dots + \text{Var}(I_n) + 2 \sum_{i < j} \text{Cov}(I_i, I_j)$$

from symmetry,

$$\text{Var}(I_1) = \text{Var}(I_2) = \dots = \text{Var}(I_n)$$

and, $\text{Cov}(I_1, I_2) = \text{Cov}(I_i, I_j)$ for all $i \neq j$ and $i \neq j$

$$\Rightarrow \text{Var}(x) = n \text{Var}(I_1) + 2 \times \binom{n}{2} \text{Cov}(I_1, I_2)$$

$$\text{Var}(I_1) = E(I_1^2) - (E I_1)^2$$

$$= E(I_1) - (E I_1)^2 = \frac{1}{n} - \frac{1}{n^2} = \frac{n-1}{n^2}$$

$$\text{Cov}(I_1, I_2) = E(I_1 I_2) - E(I_1) E(I_2)$$

$$= \frac{1}{n} \cdot \frac{1}{n-1} - \frac{1}{n^2} = \frac{n-n+1}{n^2(n-1)} = \frac{1}{n^2(n-1)}$$

$$\Rightarrow \text{Var}(x) = \frac{n \times (n-1)}{n^2} + 2 \times \frac{n(n-1)}{2} \times \frac{1}{n^2(n-1)}$$

$$= \frac{n-1}{n} + \frac{1}{n} = 1$$

$$\Rightarrow \text{Standard deviation of } x = \sqrt{\text{Var}(x)} = 1$$

(5) Let I_1 be indicator r.v. that ship of length 2 is missed, I_2 for ship of length 3 missed and I_3 for ship of length 4 missed.

X is r.v. for # ships missed.

$$\Rightarrow X = I_1 + I_2 + I_3$$

$$\Rightarrow EX = E(I_1 + I_2 + I_3) = EI_1 + EI_2 + EI_3$$

EI_1 = Probability that ship of length 2 is missed.

We can choose 5 grids from 98 remaining grids

where ship of length 2 does not reside in order to miss the ship.

$$\Rightarrow EI_1 = \frac{\binom{98}{5}}{\binom{100}{5}} = \frac{98!}{93!5!} \times \frac{95!5!}{100!}$$

$$= \frac{94 \times 95}{99 \times 100} \approx 0.902$$

$$\text{Similarly, } EI_2 = \frac{\binom{97}{5}}{\binom{100}{5}} = \frac{97!}{92!} \times \frac{95!}{100!} \approx 0.856$$

$$EI_3 = \frac{\binom{96}{5}}{\binom{100}{5}} = \frac{96!}{91!} \times \frac{95!}{100!} \approx 0.812$$

$$\Rightarrow EX = EI_1 + EI_2 + EI_3 = 0.902 + 0.856 + 0.812 \approx 2.57$$

Let Y is r.v. for # ships hit

$$\Rightarrow X+Y=3 \Rightarrow EY = E(3)-EX = 3-2.57 \approx 0.43$$

$$\text{Now, } \text{Var}(X) = \text{Var}(I_1 + I_2 + I_3)$$

$$= \text{Var}(I_1) + \text{Var}(I_2) + \text{Var}(I_3) + 2(\text{Cov}(I_1, I_2) + \text{Cov}(I_2, I_3) + \text{Cov}(I_1, I_3))$$

$$\text{Var}(I_1) = EI_1^2 - (EI_1)^2 = EI_1 - (EI_1)^2 = (0.902) - (0.902)^2 \approx 0.089$$

$$\text{Similarly, } \text{Var}(I_2) = \cancel{(0.856)} \quad 0.856 - (0.856)^2 \approx 0.123$$

$$\text{Var}(I_3) = 0.812 - (0.812)^2 \approx 0.152$$

$$\text{Cov}(I_1, I_2) = E(I_1 I_2) - EI_1 EI_2$$

$$E(I_1 I_2) = \text{Probability that both ships of length 2 and 3 are missed} = \frac{\binom{95}{5}}{\binom{100}{5}} \quad \begin{array}{l} \text{(choose 5 grids} \\ \text{from remaining} \\ \text{95 grids)} \end{array}$$

$$= \frac{95!}{90!} \times \frac{95!}{100!} \approx 0.77$$

$$\text{Similarly, } E(I_1 I_3) = \frac{\binom{94}{5}}{\binom{100}{5}} \approx 0.73$$

$$E(I_2 I_3) = \frac{\binom{93}{5}}{\binom{100}{5}} \approx 0.69$$

$$\Rightarrow \text{Cov}(I_1, I_2) = 0.77 - 0.902 \times 0.856 = -0.0021$$

$$\text{Cov}(I_1, I_3) = 0.73 - 0.902 \times 0.812 = -0.0024$$

$$\text{Cov}(I_2, I_3) = 0.69 - 0.856 \times 0.812 = -0.005$$

$$\Rightarrow \text{Var}(X) = 0.089 + 0.123 + 0.152 - 2(0.0021 + 0.0024 + 0.005) \approx 0.345$$

$$\Rightarrow \text{Var}(Y) = \text{Var}(3-X) = \text{Var}(X) = 0.345$$

(6) Probability that a person makes purchase

$$\text{purchase} = p = \frac{1}{2} \times \frac{1}{10} + \frac{1}{3} \times \frac{9}{10} = 0.35$$

\Rightarrow Probability that a person does not make purchase

$$\text{purchase} = q = 1 - 0.35 = 0.65$$

$X = \# \text{ People make purchase}$

$Y = \# \text{ People}^{\text{who}} \text{ does not make purchase}$

$$\Rightarrow (X+Y) \sim \text{Pois}(100)$$

As X and Y are independent and from chicken-egg problem,

$$P(X=n | Y=42) = P(X=n)$$

$$\text{and } X \sim \text{Pois}(100 \times 0.35) \sim \text{Pois}(35)$$

$$\text{Hence, } P(X=n | Y=42) = P(X=n) = \frac{35^n e^{-35}}{n!}$$

7.(a) Let X be r.v. for # choices of 4 courses

$$\begin{aligned}
 \text{From LOTUS, } EX &= \sum_{k=0}^{\infty} \binom{k}{4} P(X=k) = \sum_{k=0}^{\infty} \frac{k!}{k!(k-4)! 4!} \times \frac{e^{-\lambda} \lambda^k}{k!} \\
 &= \frac{e^{-\lambda}}{4!} \sum_{k=0}^{\infty} \frac{\lambda^k}{(k-4)!} = \frac{e^{-\lambda} \lambda^4}{4!} \sum_{k=0}^{\infty} \frac{\lambda^{(k-4)}}{(k-4)!} \\
 &= \frac{e^{-\lambda} \lambda^4}{4!} \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} \\
 &= \frac{e^{-\lambda} \lambda^4}{4!} e^{\lambda} = \frac{\lambda^4}{4!} = \frac{\lambda^4}{24}
 \end{aligned}$$

$$(b) P(X_{am}=a, X_{pm}=b) = \sum_n P(X_{am}=a, X_{pm}=b | N=n) P(N=n)$$

$$(b) P(X_{am}=a, X_{pm}=b) = \sum_n P(X_{am}=a, X_{pm}=b | N=n) P(N=n) \quad (\text{from LOTP})$$

where $N = \# \text{ of courses}$

$$\Rightarrow X_{am} + X_{pm} = N$$

$$\text{Hence, } P(X_{am}=a, X_{pm}=b | N=n) = 0 \quad \forall n \neq a+b$$

So we are left with

$$\begin{aligned}
 &P(X_{am}=a, X_{pm}=b | N=a+b) P(N=a+b) \\
 &= P(X_{am}=a | N=a+b) P(N=a+b) \quad \left(\text{If } X_{am}=a \text{ and given } N=a+b, X_{pm}=b \text{ is implied} \right)
 \end{aligned}$$

Now for fixed N , $X_{am} \sim \text{Bin}(N=n, 1/2)$

$$\text{i.e. } (X_{am} | N=n) \sim \text{Bin}(n, 1/2)$$

$$\textcircled{*} \quad P(N=a+b) = \frac{\lambda^{a+b} e^{-\lambda}}{(a+b)!}$$

$$\Rightarrow P(X_{am}=a, X_{pm}=b) = \binom{a+b}{a} \left(\frac{1}{2}\right)^{a+b} \lambda^{a+b} e^{-\lambda} \frac{e^{-\lambda}}{(a+b)!}$$

$$= \frac{1}{a! b!} \left(\frac{1}{2}\right)^a \left(\frac{1}{2}\right)^b \lambda^a \lambda^b e^{-\lambda} e^{-\lambda} \frac{e^{-\lambda}}{\left(\frac{1}{2}\right)^a \left(\frac{1}{2}\right)^b}$$

$$= \frac{\left(\frac{\lambda}{2}\right)^a e^{-\frac{\lambda}{2}}}{a!} \times \frac{\left(\frac{\lambda}{2}\right)^b e^{-\frac{\lambda}{2}}}{b!}$$

$$= \frac{\left(\frac{\lambda}{2}\right)^{a+b} e^{-\lambda}}{a! b!}$$

(C) In part (b) we observe that

$$X_{am} \sim \text{Pois}\left(\frac{\lambda}{2}\right)$$

$$X_{pm} \sim \text{Pois}\left(\frac{\lambda}{2}\right)$$

(Also analogous to
chicken-egg problem)

$$\text{And, } X_1 + X_2 = X_{am}$$

$$X_3 + X_4 = X_{pm}$$

Since X_1, X_2, X_3, X_4 are independent r.v.s

$$\Rightarrow X_1 \sim \text{Pois}\left(\frac{\lambda}{4}\right), X_2 \sim \text{Pois}\left(\frac{\lambda}{4}\right), X_3 \sim \text{Pois}\left(\frac{\lambda}{4}\right)$$

$$X_4 \sim \text{Pois}\left(\frac{\lambda}{4}\right)$$

$$P(X_1=a, X_2=b, X_3=c, X_4=d) = \frac{\left(\frac{\lambda}{4}\right)^a e^{-\frac{\lambda}{4}}}{a!} \times \frac{\left(\frac{\lambda}{4}\right)^b e^{-\frac{\lambda}{4}}}{b!} \times \frac{\left(\frac{\lambda}{4}\right)^c e^{-\frac{\lambda}{4}}}{c!} \times \frac{\left(\frac{\lambda}{4}\right)^d e^{-\frac{\lambda}{4}}}{d!}$$

$$= \frac{\left(\frac{\lambda}{4}\right)^{a+b+c+d} e^{-\lambda}}{a! b! c! d!}$$

Expected # of non-conflicting courses

$$= E(X_1 + X_2 + X_3 + X_4)$$

$$= EX_1 + EX_2 + EX_3 + EX_4 = 4 \times \frac{\lambda}{4} = \lambda$$

$$\frac{EX(a)}{EX(c)} = \frac{\frac{\lambda^4}{24 \times \lambda}}{\frac{\lambda^3}{24}} = \frac{\lambda^3}{24}$$

8. (a) From chicken-egg problem,

$$X = Y + Z$$

$$Y \sim \text{Pois}(p\lambda)$$

$$Z \sim \text{Pois}((1-p)\lambda)$$

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}}$$

$$= \frac{\text{Cov}(Y+Z, Y)}{\sqrt{\lambda p}}$$

$$= \frac{\text{Cov}(Y, Y) + \text{Cov}(Y, Z)}{\lambda \sqrt{p}}$$

$$= \frac{\lambda p + 0}{\lambda \sqrt{p}} = \sqrt{p} \quad (\text{Cov}(Y, Z) = 0 \text{ and } Y \text{ and } Z \text{ are independent and } \text{Cov}(Y, Y) = \text{Var}(Y) = \lambda p)$$

(b)

Let S be r.v. for # non-statistics major

$$S \sim \text{Bin}(n, 1-r)$$

$$P(Y=y, Z=z, S=s) = P(Y=y, Z=z | S=s) P(S=s)$$

$$= P(Y=y | S=s) P(S=s) \quad (\text{Because given } S=s, S=s \text{ and } Z=n-s-y)$$

Now, $(Y=y | S=s) \sim \text{Bin}(n-s, p)$ is implied

$$\Rightarrow P(Y=y, Z=z, S=s) = \binom{n-s}{y} p^y (1-p)^{n-s-y} (1-r)^s r^{n-s}$$

$$= \frac{n!}{s! y! (n-s-y)!} \times P^y (1-p)^{n-y-s} (1-r)^s r^{n-s}$$

$$= \frac{n!}{y! z! s!} p^y (1-p)^z (1-r)^s r^{y+z}$$

$$Y \sim \text{Bin}(n, p\lambda) \Rightarrow P(Y=y) = \binom{n}{y} (p\lambda)^y (1-p\lambda)^{n-y}$$

$$Z \sim \text{Bin}(n, (1-p)\lambda) \Rightarrow P(Z=z) = \binom{n}{z} ((1-p)\lambda)^z (1-\lambda + p\lambda)^{n-z}$$

$$S \sim \text{Bin}(n, 1-r) \Rightarrow P(S=s) = \binom{n}{s} (1-r)^s r^{n-s}$$

(C) $X \sim \text{Bin}(n, r)$

$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

~~etc corr(2+X, Y)~~

$$= \frac{\text{cov}(2+X, Y)}{\sqrt{n r(1-r) \times n p r(1-p r)}}$$

$$= \frac{\text{cov}(Y, Y) + \text{cov}(2, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

($\text{cov}(2, Y) = 0$
because 2 and
 Y are independent)

$$= \frac{\text{Var} Y}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

$$= \sqrt{\frac{\text{Var}(Y)}{\text{Var}(X)}} = \sqrt{\frac{p r(1-p r)}{P(1-P)}} = \sqrt{\frac{r(1-p r)}{1-P}}$$

$$(9) \quad R^2 = x^2 + y^2, \quad X = \sqrt{R^2} \cos \theta$$

$$Y = \sqrt{R^2} \sin \theta$$

$$\theta = \cos^{-1}\left(\frac{x}{\sqrt{x^2+y^2}}\right) = \sin^{-1}\left(\frac{y}{\sqrt{x^2+y^2}}\right)$$

$$f_{R^2, \theta}(r^2, \theta) = \frac{\partial(x, y)}{\partial(r^2, \theta)} f_{x, y}(x, y)$$

$$\frac{\partial(x, y)}{\partial(r^2, \theta)} = \begin{pmatrix} \frac{\cos \theta}{2R} & -R \sin \theta \\ \frac{\sin \theta}{2R} & R \cos \theta \end{pmatrix}$$

$$\left| \frac{\partial(x, y)}{\partial(r^2, \theta)} \right| = \left| \frac{\cos^2 \theta}{2} + \frac{\sin^2 \theta}{2} \right| = 1/2$$

$$\Rightarrow f_{R^2, \theta}(r^2, \theta) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \times \frac{1}{\sqrt{2\pi}} e^{-y^2/2} \times \frac{1}{2} \\ = \frac{1}{4\pi} e^{-\frac{(x^2+y^2)}{2}} \\ = \frac{1}{4\pi} e^{-r^2/2}$$

As θ does not appear in joint PDF, it must be uniform and $\theta \in [0, 2\pi] \Rightarrow f_\theta(\theta) = \frac{1}{2\pi}$

$$\Rightarrow f_{R^2}(r^2) = \frac{1}{2} e^{-r^2/2}$$

~~According to product of marginal PDF~~

$$\textcircled{1} \quad R^2 \sim \text{Exp}(1/2)$$

$$\theta \sim \text{Unif}(0, 2\pi)$$

$$10.(a) \text{ Let } A = \frac{T_1}{T_1+T_2} \text{ and } B = \frac{T_2}{T_1+T_2}$$

$$A+B=1 \Rightarrow B=1-A$$

$$\Rightarrow \frac{A}{B} = \frac{A}{1-A}$$

From Post office - Bank Problem,

A and (T_1+T_2) are independent r.v.s

$\Rightarrow \frac{A}{1-A}$ and (T_1+T_2) are also independent

$$(b) P(T_1 < T_2) = P(T_1 - T_2 < 0) \Rightarrow T_1 \in [0, \infty], T_2 \in [0, \infty]$$

$$\begin{aligned} &= \int_0^\infty \int_0^{t_2} (\lambda_1 e^{-\lambda_1 t_1} dt_1) \lambda_2 e^{-\lambda_2 t_2} dt_2 \\ &= \lambda_1 \lambda_2 \int_0^\infty e^{-\lambda_2 t_2} dt_2 \int_0^{t_2} e^{-\lambda_1 t_1} dt_1 \\ &= \lambda_2 \int_0^\infty e^{-\lambda_2 t_2} (1 - e^{-\lambda_1 t_2}) dt_2 \\ &= 1 - \lambda_2 \int_0^\infty e^{-(\lambda_1 + \lambda_2)t_2} dt_2 \\ &= 1 - \frac{\lambda_2}{\lambda_1 + \lambda_2} = \frac{\lambda_1}{\lambda_1 + \lambda_2} \end{aligned}$$

if $(\lambda_1 = \lambda_2) \Rightarrow P(T_1 < T_2) = \frac{1}{2}$ which makes sense from symmetry

(c) We want to find,

$$EX = E(\min(T_1, T_2)) + E(T_1)P(T_1 < T_2) + E(T_2)P(T_2 < T_1)$$

We know that $\min(T_1, T_2) \sim \text{Exp}(0)(\lambda_1 + \lambda_2)$

$$\begin{aligned} \Rightarrow EX &= \frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_1} \times \frac{\lambda_1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_2} \times \frac{\lambda_2}{\lambda_1 + \lambda_2} \\ &= \frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_1 + \lambda_2} + \frac{1}{\lambda_1 + \lambda_2} = \frac{3}{\lambda_1 + \lambda_2} \end{aligned}$$

(11) From the question,

$$P(P>r | X=n) = c$$

where X be r.v. for # successful trials

$$\Rightarrow P(P>r | X=n) = \frac{P(X=n | P>r) P(P>r)}{P(X=n)}$$

$$P(X=k | P>r) \sim \text{Bin}(n, p)$$

$$\Rightarrow P(X=n | P) = p^n$$

$$P(P>r) = (1-\alpha)$$

$$P(X=n) = \int_0^n p^n dp = \left(\frac{p^{n+1}}{n+1} \right)_0^1 = \frac{1}{n+1}$$

$$\Rightarrow P(P>r | X=n) = \frac{p^n (1-\alpha)}{\frac{1}{n+1}}$$
$$= p^n (n+1) (1-\alpha)$$

$$\Rightarrow c = p^n (n+1) (1-\alpha)$$

12. (a) There are $(n+1)$ possible spots where Y_{new} can go, and to be in $[Y_{(j)}, Y_{(k)}]$ there are $(k-j)$ possibilities.

Hence, as all positions are equally likely
 $\Rightarrow P(Y_{\text{new}} \in [Y_{(j)}, Y_{(k)}]) = \frac{k-j}{n+1}$

(b) We want a interval of $[Y_{(j)}, Y_{(j+95)}]$

And since there are $n=59$ r.v.s

We want 3 r.v.s out of this interval.