

Assignment 1

Given,

$$P(\text{error} | x) = \min \{P(w_1 | x), P(w_2 | x)\}$$

let the min w_i be represented by w_m

a) $P(\text{error} | x) = P(w_m | x)$

for this, ~~we~~ let us consider $w_m = w_1$
i.e. $m=1$ $\therefore P(\text{error} | x) = P(w_1 | x)$

$$P(w_2 | x) = 1 - P(w_1 | x) \rightarrow (1)$$

$$P(w_1 | x) > P(w_2 | x) \rightarrow (2)$$

from (1) and (2)

$$P(w_1 | x) > \frac{1}{2} \quad \text{i.e. } 2P(w_1 | x) > 1$$

now multiplying by $P(w_2 | x)$

$$2P(w_1 | x) P(w_2 | x) > P(w_2 | x)$$

$$2P(w_1 | x) P(w_2 | x) > P(\text{error} | x)$$

by integrating this at every x

$$\int 2P(w_1 | x) P(w_2 | x) dx > \int P(\text{error} | x) dx$$



upper bound on full error

* we get same answer if we consider

$$P(\text{error} | x) = P(w_1 | x)$$

1(b) we know that $P(w_1|x) > \frac{1}{2}$

consider $\alpha = 1.2$

$$P(w_1|x) = 0.7$$

$$\therefore P(w_2|x) = 0.3$$

$$\therefore P(\text{error}|x) = 0.3$$

$$\begin{aligned} \alpha P(w_1|x) P(w_2|x) &= 1.2 \times 0.7 \times 0.3 \\ &= 1.2 \times 0.21 \\ &= 0.252 \end{aligned}$$

$$0.252 < P(\text{error}|x)$$

\therefore This cannot possibly give upper bound guarantee

1(c) Since $P(\text{error}|x) = P(w_2|x)$

we have

$$P(w_1|x) < P(w_2|x) \quad \begin{matrix} P(\text{error}|x) \\ \uparrow \end{matrix}$$

$$P(w_1|x) P(w_2|x) < P(w_2|x) P(w_2|x)$$

$$\int P(w_1|x) P(w_2|x) dx < \int P(w_2|x) P(\text{error}|x) dx$$

we know

$$\int P(w_2|x) P(\text{error}|x) dx < \int P(\text{error}|x) dx$$

$$\text{as } P(w_2|x) < 1 \quad \forall x$$

$$\therefore \int P(\text{error}|x) dx > \int P(w_1|x) P(w_2|x) dx$$

Hence we have lower bound.

1. (a) Let $\beta = 1.2$ $p(w_1|x) = 0.7$

} Same as 1(b)

2. Given,

$$p(x|w_i) \propto e^{-|x-a_i|/b_i} \rightarrow (1)$$

from (1) we can get-

$$p(x|w_i) = K e^{-|x-a_i|/b_i}$$

2a) To normalize, we integrate over full range and = 1

$$\int_{-\infty}^{\infty} p(x|w_i) dx = 1$$

$$\begin{aligned} x - a_i &< 0 \text{ if } x < a_i \\ x - a_i &\geq 0 \text{ if } x \geq a_i \end{aligned}$$

$$= \int_{-\infty}^{\infty} K e^{-|x-a_i|/b_i} dx = K \int_{-\infty}^{a_i} e^{(x-a_i)/b_i} dx$$

$$+ K \int_{a_i}^{\infty} e^{-(x-a_i)/b_i} dx$$

$$= \left[K b_i e^{(x-a_i)/b_i} \right]_{-\infty}^{a_i} + \left[-K b_i e^{-(x-a_i)/b_i} \right]_{a_i}^{\infty}$$

$$= b_1 k_0 - 0 + 0 - (-b_1 k)$$

$$= 2 b_1 k$$

$$2 b_1 k = 1$$

$$k = \frac{1}{2 b_1}$$

$$p(x|w_2) = \frac{1}{2 b_1} e^{-|x-a_1|/b_1}$$

2 a) Likelihood ratio is given by

$$\frac{p(x|w_2)}{p(x|w_1)} = \frac{\frac{1}{2 b_2} e^{-|x-a_2|/b_2}}{\frac{1}{2 b_1} e^{-|x-a_1|/b_1}}$$

$$= \frac{b_1}{b_2} e^{-[|x-a_2|/b_2 - |x-a_1|/b_1]}$$

$$2 c) \frac{p(x|w_1)}{p(x|w_2)} = \frac{b_2}{b_1} e^{-\frac{|x-a_1|}{b_1} + \frac{|x-a_2|}{b_2}}$$

$$a_1 = 0 \quad a_2 = 1 \quad b_1 = 1 \quad b_2 = 2$$

when $x < 0$

$$\frac{p(x|w_1)}{p(x|w_2)} = \frac{x+1}{2} \cdot \frac{1}{2} (x+1)$$

when $x \leq 1, x > 0$

$$-x + \frac{- (x-1)}{2}$$

$$-x - \frac{x}{2} + \frac{1}{2} = \frac{1}{2} (1 - 3x)$$

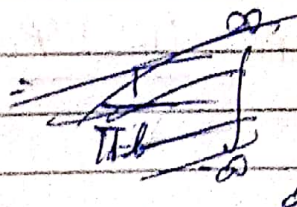
when $x > 1$

$$-x + \frac{x-1}{2} = -\frac{1}{2} (x+1)$$

$$\frac{p(x|w_1)}{p(x|w_2)} = \begin{cases} 2e^{(x+1)/2} & x \leq 0 \\ 2e^{(1-3x)/2} & 0 < x \leq 1 \\ 2e^{-(1+x)/2} & x > 1 \end{cases}$$

8. $p(x|w_i) = \frac{1}{\pi b} \frac{1}{1 + \frac{(x-a_i)^2}{b^2}} \quad i=1,2$

$$\int_{-\infty}^{\infty} p(x|w_i) dx = \int_{-\infty}^{\infty} \frac{1}{\pi b} \frac{1}{1 + \frac{(x-a_i)^2}{b^2}} dx$$



substituting $\frac{x-a_i}{b} = y$

$$dy = \frac{1}{b} dx \quad dx = b dy$$

$$\frac{1}{\pi b} \int_{-\infty}^{\infty} \frac{1}{1+y^2} b dy$$

$$= \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1}{1+y^2} dy$$

$$\frac{1}{\sqrt{1+y^2}} = \frac{1}{\sin^2 \theta} \quad \text{where } y^2 = 1 - \frac{1}{\sin^2 \theta}$$

$$d \cdot y = \frac{1}{\sin^2 \theta} d\theta$$

$$K = \frac{1}{\pi} \int_{\theta=-\pi}^{\theta=0} \frac{\sin^2 \theta}{\sin^2 \theta} d\theta$$

$$= \frac{1}{\pi} [0 - (-\pi)] = 1$$

Hence it is normalized.

Q6) Given, $P(\omega_1) = P(\omega_2)$

$$\text{TPT } P(\omega_1 | x) = P(\omega_2 | x) \text{ at } x = (a_1 + a_2)/2$$

We have to find the point of decision boundary

$$P(x | \omega_1) P(\omega_1) = P(x | \omega_2) P(\omega_2)$$

$$\cancel{P(x | \omega_1) P(\omega_1)} \neq \cancel{P(x | \omega_2) P(\omega_2)}$$

\neq

$$2 = \frac{P(x | \omega_1) P(\omega_1)}{P(x | \omega_2) P(\omega_2)}$$

$$\frac{1}{1 + \left(\frac{x - a_1}{b}\right)^2} = \frac{1}{1 + \left(\frac{x - a_2}{b}\right)^2}$$

$$x - a_1 = \pm (x - a_2)$$

$$a_1 \neq a_2 \quad x = \frac{a_1 + a_2}{2}$$

which is the midpoint.

$$\text{sol)} \quad \lim_{x \rightarrow \infty} p(x|w_i) = \frac{1}{\pi b} \frac{1}{1 + \left(\frac{x - a_i}{b}\right)^2}$$

$$\text{sol)} \quad p(w_i|x) = \frac{p(x|w_i) p(w_i)}{Z}$$

$$Z = \sum_i p(x|w_i) p(w_i)$$

$$\therefore \lim_{x \rightarrow \infty} p(w_i|x) = \lim_{x \rightarrow \infty} \frac{\frac{1}{\pi b} \cdot \frac{1}{1 + \left(\frac{x - a_i}{b}\right)^2}}{Z \text{ (samples to num but } i=1, 2)}$$

$$\lim_{x \rightarrow \infty} \frac{b^2 + (x - a_i)^2}{b^2 + (x - a_1)^2 + b^2 + (x - a_2)^2}$$

$$= \frac{x^2}{x^2 + x^2} = \frac{1}{2}$$

g 9a) Consider $a_2 > a_1 \rightarrow$ we can do this without loss of generality from 8 we know that the decision boundary is at $(a_1 + a_2)/2$.

$$P(\text{error}) = \int_{-\infty}^{(a_1+a_2)/2} p(w_2|x) dx + \int_{(a_1+a_2)/2}^{\infty} p(w_1|x) dx$$

$$= \frac{1}{\pi a} \int_{-\infty}^{(a_1+a_2)/2} \frac{1}{1 + \left(\frac{x-a_2}{a}\right)^2} dx + \frac{1}{\pi a} \int_{(a_1+a_2)/2}^{\infty} \frac{1}{1 + \left(\frac{x-a_1}{a}\right)^2} dx$$

$y = \frac{x-a_2}{a}$

$$\therefore \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1}{1+y^2} dy$$

$$\int \frac{1}{1+y^2} dy = \tan^{-1} y$$

$$\therefore P(\text{error}) = \frac{1}{\pi} \left[\tan^{-1} \left(\frac{a_1 - a_2}{2a} \right) - \tan^{-1} (-\infty) \right]$$

$$= \frac{1}{2} - \frac{1}{\pi} \tan^{-1} \left| \frac{a_2 - a_1}{2a} \right|$$

9c) $P(\text{error}) = \frac{1}{2}$

occurs for $\frac{|a_2 - a_1|}{2b} = 0$

i.e. if $a_1 = a_2$ ~~or~~ \rightarrow when both dist or
same

~~or~~
 $b = \infty \rightarrow$ when it is flat

~~less~~

12 Given,

$w_{\max}(x)$

$P(w_{\max}(x)|x) \geq P(w_0|x)$

12a) T S T $P(w_{\max}|x) \geq \frac{1}{c}$

if $P(w_0|x) = P(w_j|x) \forall j$

then $P(w_0|x) = \frac{1}{c}$

$\therefore P(w_{\max}|x) = \frac{1}{c}$

if $\exists i$ st $P(w_0|x) < \frac{1}{c}$

then $\exists j$, $P(w_j|x) > \frac{1}{c}$

~~that is~~ which is not possible here

$P(w_{\max}|x) = \frac{1}{c}$

$$12b) \quad p(\text{error}) = 1 - \int p(w_{\max} | x)$$

$$12b) \quad p(\text{error}) = 1 - p(\text{correct})$$

$$\therefore p(\text{error}) =$$

$$p(\text{correct}) = \int \max_c p(w_c | x) dx$$

$$= \int p(w_{\max} | x) dx$$

$$\therefore p(\text{error}) = 1 - \int p(w_{\max} | x) p(x) dx$$

12c) consider $\alpha \geq 1/c$

$$p(\text{error}) \leq 1 - \int \alpha p(x) dx$$

$$\leq 1 - \alpha \int p(x) dx = 1 - \alpha$$

$$\therefore p(\text{error}) \leq 1 - \frac{1}{c} = \frac{c-1}{c}$$

12d) This is possible only if all classes have same priors and identical distributions