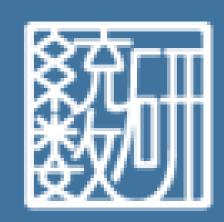
# STATISTICAL INVARIANCE OF BETTI NUMBERS

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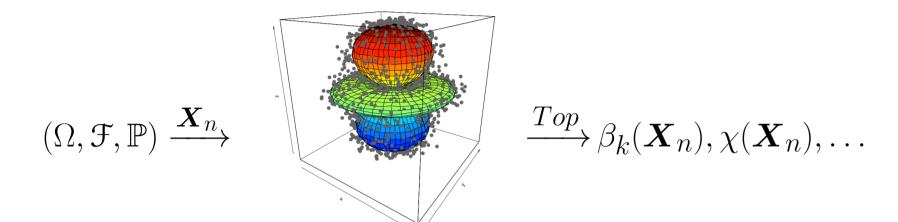
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## **Topology and Data**

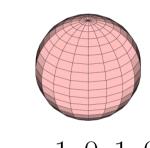
**Topological Data Analysis** (TDA) is a promising new paradigm comprising of *mathematical*, *statistical* and *algorithmic* tools to study the shape of data.

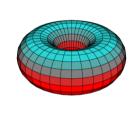
- Consider a "black-box" probability space  $(\Omega, \mathcal{F}, \mathbb{P})$
- $X_n = \{X_1 \dots X_n\}$  is a random collection of points in  $\mathbb{R}^d$
- Topological summaries  $Top(\boldsymbol{X}_n)$  are random variables pushed-forward to a summary space  $(S, \mathcal{B}(S), \mathbb{Q})$

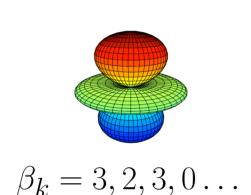


# **Betti Numbers and Topological Invariants**

- The homology  $\{H_k(\mathfrak{X})\}_{k\in\mathbb{N}}$  is a topological invariant of  $\mathfrak{X}$
- The Betti numbers are given by  $\beta_k(\mathfrak{X}) = \dim (H_k(\mathfrak{X}))$
- Informally,  $\beta_k(\mathfrak{X})$  counts the # of k-dimensional voids in  $\mathfrak{X}$







 $\beta_k = 1, 0, 1, 0 \dots$   $\beta_k = 1, 2, 1, 0 \dots$ 

# The TDA Pipeline

**Persistent homology** examines topological features across a wide spectrum of resolutions.

- 1. At resolution r > 0 form a thickening  $\{B_r(X_k)\}_{k=1}^n$
- 2. Next, construct a simplicial complex  $K_r$  (ex.  $\operatorname{\check{e}ech}$ , Rips, etc. )
- 3. Examine the simplicial homology for the filtration  $\{K_r\}_{r>0}$

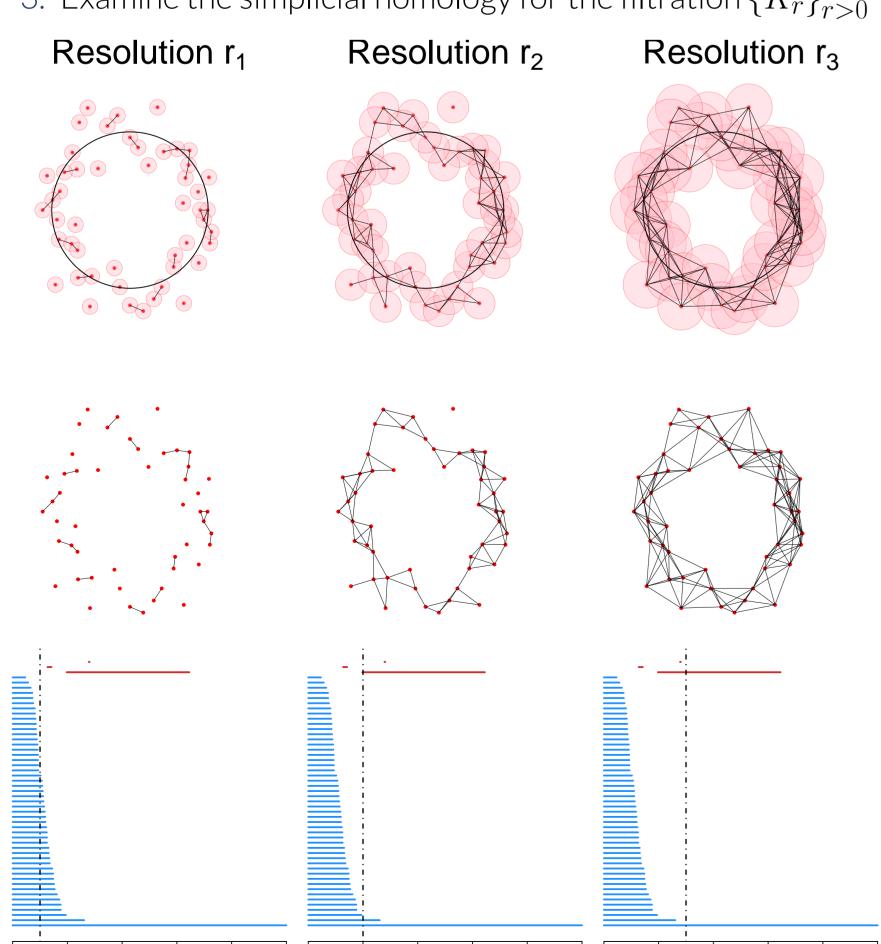


Figure 1. Data is uniformly sampled from a circle. The persistence barcode depicts the number of connected components (blue) and the number of loops (red) as the resolution r increases

#### **Motivation**

- $\mathcal{P} = \{\mathbb{P}_{\theta} : \theta \in \Theta\}$  is a parametric family of distributions
- Let  $X_n \sim \mathbb{P}_{\theta_1}$  and  $Y_n \sim \mathbb{P}_{\theta_2}$  be two collections of points which are from fundamentally different distributions
- Our aim is to examine the conditions under which they have identical asymptotic behaviour of the Betti numbers i.e.

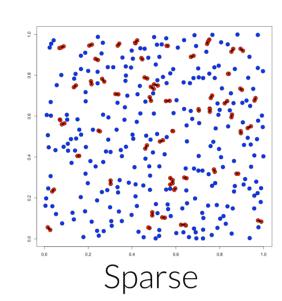
$$\lim_{n \to \infty} \frac{1}{n} \mathbb{E}\left(\beta_k \left(\check{\mathcal{C}}\left(\boldsymbol{X}_n, r_n\right)\right)\right) \stackrel{?}{=} \lim_{n \to \infty} \frac{1}{n} \mathbb{E}\left(\beta_k \left(\check{\mathcal{C}}\left(\boldsymbol{Y}_n, r_n\right)\right)\right)$$

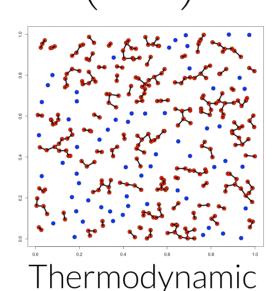
This outlines conditions when topological inference is possible

## **Asymptotic Regimes**

The asymptotic behaviour (as  $n \to \infty$ ) is qualitatively different as the behaviour of the resolution  $r_n \to 0$  varies. These are:

- Sparse regime :  $r_n = o\left(n^{-1/d}\right)$
- Thermodynamic regime:  $r_n = \Theta\left(n^{-1/d}\right)$
- Dense regime :  $r_n = \omega \left( n^{-1/d} \right)$





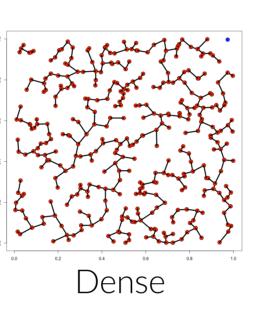


Table 1. Asymptotic regimes of  $\beta_0$  for 2D-Poisson Process with  $\lambda=700$ 

# **Minimal Spanning Trees**

Minimal spanning trees are intrinsically related to  $\beta_0(\boldsymbol{X}_n,r)$  (Lemma): Let  $\boldsymbol{X}_n=\{X_1,X_2\ldots X_n\}$  be a random collection of points in  $\mathbb{R}^d$ . Then the following hold true:

- 1. The Euclidean MST  $\mathcal{M}(\boldsymbol{X}_n)$  is unique a.s.  $\mathbb{P}$
- 2. The smallest edge  $e^*$  is an element of  $\mathcal{M}(\boldsymbol{X}_n)$
- 3. For each  $\boldsymbol{Y}, \boldsymbol{Z} \subseteq \boldsymbol{X}_n$  s.t.  $\boldsymbol{Y} \cap \boldsymbol{Z} = \varnothing$  and  $\boldsymbol{Y} \cup \boldsymbol{Z} = \boldsymbol{X}_n$ , the edge e defined by  $\|e\| = \min_{y \in \boldsymbol{Y}, z \in \boldsymbol{Z}} \|y z\|$  is in  $\mathfrak{M}(\boldsymbol{X}_n)$ .

This reveals the relationship between the Euclidean MST and the  $0^{th}$  persistence barcode.

(Theorem): Under the conditions of the previous Lemma:

- 1. The edges of  $\mathcal{M}(\boldsymbol{X}_n)$  generate the  $0^{th}$  persistence barcode.
- 2. The  $0^{th}$  persistent Betti number at resolution r is given by

$$\beta_0\left(\check{\mathcal{C}}(\boldsymbol{X}_n)\right) = n - \sum_{e \in \mathcal{M}(\boldsymbol{X}_n)} \mathbb{1}_{[0,r]}\left(\|e\|\right)$$

Table 2. Correspondence between  $\mathcal{M}(\boldsymbol{X}_n)$  and  $0^{th}$  Barcode

## Thermodynamic Behaviour

The thermodynamic regime exhibits interesting behaviour

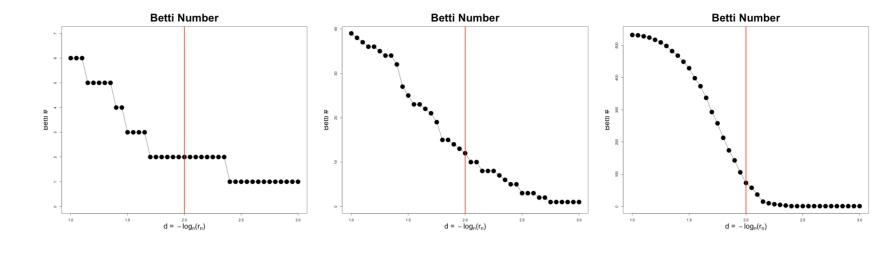


Table 3.  $\beta_0$  in the theormodynamic regime for PP with  $\lambda=10,50,500$ 

(Theorem): Let  $X_n \stackrel{iid}{\sim} f$ , where f(x) is bounded, Riemann integrable with compact support. When  $n^{1/d}r_n \to t \in (0, \infty)$ :

$$\frac{1}{n}\mathbb{E}\left(\beta_0\left(\check{\mathcal{C}}\left(\boldsymbol{X}_n,r_n\right)\right)\right) \to \int_{\mathbb{R}^d} \mathbb{E}\left(\sum_{e \in \mathcal{M}(\mathcal{P}_{1,0})} \mathbb{1}_{[0,t]}\left(f(\boldsymbol{x})^{\frac{-1}{d}} \|e\|\right)\right) f(\boldsymbol{x}) d\boldsymbol{x}$$

where  $M(\mathcal{P}_{1,0})$  is the MST for the unit intensity Poisson process with a point at the origin. Now, when we look at any  $\beta_k$ 

(Theorem 3.3, [3]) Under the conditions of the Theorem above there exist functions  $\hat{\beta}_k$  such that:

$$\frac{1}{n}\mathbb{E}\left(\beta_k\left(\check{\mathcal{C}}\left(\boldsymbol{X}_n,r_n\right)\right)\right)\xrightarrow[\mathbb{R}^d]{n\to\infty}\int\limits_{\mathbb{R}^d}\hat{\beta_k}\left(f(\boldsymbol{x})^{1/d}t\right)f(\boldsymbol{x})d\boldsymbol{x}:=\Psi_k(f,t)$$

### **Statistical Invariance : Characterization**

Define  $\mathcal{F}_k$  as a  $\Psi_k$ -invariant family of densities such that  $\Psi_k(f,t)=\Psi_k(g,t) \ \forall f,g\in\mathcal{F}_k$ . These densities admit identical behaviour for  $\beta_k$  in the thermodynamic regime.

We define  $\mathcal{F}^*$  as the family of densities such that for each  $t \geq 0$   $\mathbb{E}\left(\mathbb{1}\left(f(\boldsymbol{X}) \geq t\right)\right) = \mathbb{E}\left(\mathbb{1}\left(g(\boldsymbol{Y}) \geq t\right)\right) \ \forall \boldsymbol{X} \sim f, \boldsymbol{Y} \sim g.$ 

For families indexed by  $\Theta$  we denote them  $\mathcal{F}^*(\Theta)$  and  $\mathcal{F}_k(\Theta)$ 

(Lemma): 
$$\mathfrak{F}^* \subset \bigcap_{k=0}^{\infty} \mathfrak{F}_k$$

Thus,  $\mathcal{F}^*$  admits identical behaviour for each Betti number  $\beta_k$ 

### Statistical Invariance : I

We employ groups to characterize strong invariance properties.

(Theorem): Suppose  $\mathcal G$  is a group of Borel-measurable isometries acting on  $\mathfrak X\subseteq\mathbb R^d$ , and  $T:\mathfrak X\to\mathfrak T$  is  $\mathcal G$ -maximal invariant. If  $\boldsymbol X_\theta\sim f_\theta(\boldsymbol x)$  where

$$f_{\theta}(\boldsymbol{x}) = \phi\left(g_{\theta} \circ \Psi\left(\boldsymbol{x}\right)\right)$$

where  $\Psi \in \mathcal{C}^1(\mathfrak{X})$ ; and,  $\phi : \mathfrak{X} \to \mathbb{R}_{\geq 0}$  is some function which ensures that  $f_\theta$  is a valid density.

Then,  $\{f_{\theta}: \theta \in \Theta\}$  admits  $\mathcal{F}^*(\Theta)$ -invariance if and only if  $\det \left( \boldsymbol{J}_{\Psi^{-1}}(x) \right) = \zeta \left( T\left( \boldsymbol{x} \right) \right)$  for some function  $\zeta: \mathcal{T} \to \mathbb{R}$ 

This gives us necessary and sufficient conditions for identifying  $f_{\theta}$  upto isometry from the asymptotic behaviour of Betti numbers.

**(Theorem)**: Let  $\mathcal{P}$  be a family of distributions such that, for each  $f_{\theta} \in \mathcal{P}$ ,  $f_{\theta}$  statisfies stochastic regularity conditions. Then,  $\mathcal{P}$  admits  $\mathcal{F}^*$ -invariance if and only if

$$\langle f_{\theta}^k, S_{\theta} \rangle_{L^2} = 0 \ \forall k \in \mathbb{N}$$

where,  $S_{\theta}$  is the score-function given by  $S_{\theta}(\boldsymbol{x}) = \nabla_{\theta} \log(f_{\theta}(\boldsymbol{x}))$ 

(Example 1): Consider  $\mathfrak{X} = \mathbb{R}^2$  and  $R_{\theta} = \begin{pmatrix} \cos(\theta) & \sin(\theta) \\ -\sin(\theta) & \cos(\theta) \end{pmatrix}$  is the 2D-rotation matrix with  $v_{\theta} := (\cos(\theta), \sin(\theta))^{\top}$ .

Denote  $\Phi(\boldsymbol{x})$  as the CDF of a  $\mathcal{N}(\mathbf{0}, I_2)$  distribution. Then,

$$f_{\theta}(x,y) = \left(v_{\theta}^{\top} \Phi^{-1}(x,y)\right)^2 = \left(\cos(\theta)\Phi^{-1}(x) + \sin(\theta)\Phi^{-1}(y)\right)^2$$

admits invariance for each 2D-rotation  $R_{\theta}$ .

In the general case,  $\mathfrak{X}=\mathbb{R}^d$  and  $\mathcal{G}=S\mathcal{O}(d)$  with  $v_{\theta}\in S^{d-1}$ 

#### **Statistical Invariance : Non-isometric Cases**

We illustrate conditions where fundamentally different distributions admit asymptotic invariance.

**(Example 2)**: Let g be a density on  $\mathbb{R}_+$  with  $\phi_a(x) = ax$  and  $\phi_b(x) = -bx$  with the condition that  $\frac{1}{a} + \frac{1}{b} = 1$ . Then,

 $f(x) = \{g(ax)\mathbb{1} (x \ge 0) + g(-bx)\mathbb{1} (x \le 0)\} \text{ admits invariance }$ 

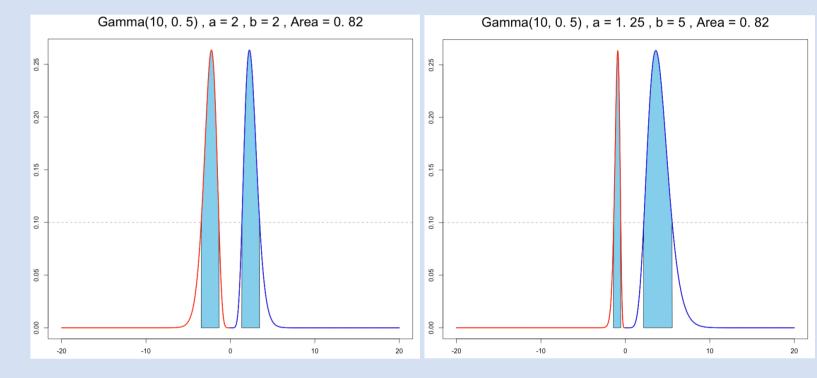


Table 4. Illustration of invariance for Gamma(10, 0.5) distribution

(Abridged Theorem): Under some technical conditions on  $\nu$ ,  $\mu$  and  $\Theta$ ; Suppose g is a density with respect to  $\nu$  which satisfies  $\nu(d\boldsymbol{y}) = \Psi\left(\left|\det\left(J_{\phi}\right)\right|\right)\nu(d\boldsymbol{x})$ , where  $\phi_{\boldsymbol{\theta}}: \mathbb{R}^d \to \mathbb{R}^d$  is a full-rank linear bijection for each  $\boldsymbol{\theta} \in (\Theta, \Xi, \mu)$ . Define,

$$f(\boldsymbol{x}, \boldsymbol{\theta}) = g(\phi_{\boldsymbol{\theta}}(\boldsymbol{x}))$$

Then f admits  $\mathcal{F}^*$ -invariance for each  $(\phi_{\pmb{\theta}}, \mu)$  if

$$\int\limits_{\Theta} \Psi\left(\left|\det\left(J_{\phi_{\pmb{\theta}}^{-1}}\right)\right|\right) \mu(d\pmb{\theta}) = 1$$

**(Example 3)**: Suppose g(r) be a density with  $supp(g) \subseteq \mathbb{R}_+$  w.r.t. the measure  $\nu$  such that  $\nu(dr) = d(r^d) = r^{d-1}dr$ .

Let  $\Theta = S^{d-1}$  and  $a: S^{d-1} \to \mathbb{R}_+$  be a non-negative function.

For each  $\pmb{\theta} \in S^{d-1}$  define the mapping  $\phi_{\pmb{\theta}}(r) = ra(\pmb{\theta})$  such that

$$\int_{S^{d-1}} \frac{\mu(d\boldsymbol{\theta})}{a(\boldsymbol{\theta})^d} = 1$$

Then for each such  $(a(\boldsymbol{\theta}), \mu)$  we have that

 $f(\boldsymbol{x}) = f(r, \boldsymbol{\theta}) = g(ra(\boldsymbol{\theta}))$  admits invariance

#### References

- [1] Morris L Eaton. Group invariance applications in statistics. In Regional conference series in Probability and Statistics, pages i–133. JSTOR, 1989.
- [2] Mathew D Penrose, Joseph E Yukich, et al. Weak laws of large numbers in geometric probability. *The Annals of Applied Probability*, 13(1):277–303, 2003.
- [3] Khanh Duy Trinh. A remark on the convergence of betti numbers in the thermodynamic regime. *Pacific Journal of Mathematics for Industry*, 9(1):4, 2017.
- [4] Robert A Wijsman. Invariant measures on groups and their use in statistics. IMS, 1990.