VARIATIONAL INFERENCE USING NORMALIZING FLOWS

SILAS BRACK AND JESPER MARTINHOFF KNUDSEN

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1. Introduction

In this project we will apply normalizing flows[1, 2] in the realm of variational inference. So to give some context to our endeavour we will start out by motivating variational inference and normalizing flows.¹

1.1. **Variational inference.** In variational inference we strive to approximate some distribution by an element in some distributional family.

The reason we might wish to do so is often that we have some implicit distribution which we do not have an analytical expression for but we might be able to sample form it by use of some sampling scheme like MCMC chains.

in our case this we have an implicit distribution which arises from forming the posterior form some non conjugate prior-likelihood pair:

$$p(z|\mathcal{D}) = \frac{p(\mathcal{D}|z)p(z)}{p(\mathcal{D})}$$

Here we have two annoyances, one being the intractability of the integral in $p(\mathcal{D}) = \int p(\mathcal{D}|z)p(z)dz$ and the other being annoyance of the non-parametric functional form of: $p(z|\mathcal{D}) \propto \prod_n p(d_n|z)p(z)$ which becomes excessively cumbersome for inference on large data sets. Thus this incentivises us to find a function which can approximate $p(z|\mathcal{D}) \approx q_{\theta}(z)$ as this would allow us to sample directly and independently from $q\theta$ and have q_{θ} be a parametric distribution and thus potentially much less cumbersome. However the effectiveness of this approach all depends on how well $q_{\theta}(z)$ approximates $p(z|\mathcal{D})$. This necessitates rich distributional families. and this is where the frame work of normalizing flows seems particular promising. That is normalizing flows gives us a general frame work for constructing arbitrarily complex distributional families. The general idea is to start out

Thanks to our supervisor Michael Riis Andersen.

¹https://github.com/silasbrack/normalizing-flows

with simple distribution in the sense that they are easy to sample from, then we apply a series of transformations to end up with a possibly much more complex distribution, but one which we can still sample independently from as we know its relation to a sampleable distribution.

With this we are ready to start developing normalizing flows further and start to look at how we can use them to approximate arbitrary distributions.

Letting $p(z|y) \approx q(z)$, we calculate the posterior predictive as:

$$p(y^*|y) = \int p(y^*|z)p(z|y)dz$$
$$p(y^*|y) \approx \int p(y^*|z)q(z)dz$$

2. NORMALISING FLOWS

As previously mentioned a normalizing flow can be defined by a sampleable base distribution along with a series of diffeomorphisms:

$$z_0 \sim q(z_0) \qquad \qquad f_{(n)} = f_n \circ f_{n-1} \circ \cdots \circ f_1$$

and well denote the z_0 s image through $f_{(n)}$ as z_n , ie $f_{(n)}(z_0) = z_n$. We can then recursively apply the change of variable formula:

$$p(z') = p(z) \left| \det(J(f^{-1})) \right| = p(z) \left| \det(J(f \circ z)) \right|^{-1} \quad \text{where} \quad z' = f(z)$$

to derive $q(z_n)$, here the last equality is due to the choice of diffeomorphisms. Through recursive application over all f_i we derive:

$$q(z_n) = q(z_0) \prod_k |\det J(f_k \circ z_{k-1})|^{-1}$$
.

This can be viewed as parametric distributional family parameterized by the choice of $q(z_0)$ as well as the choice of $f_{(n)}$ whose parameters we will denote $\theta = \{\theta_n, \theta_{n-1}, \cdots, \theta_1\}$

3. INVERTIBLE LINEAR-TIME TRANSFORMATIONS

A simple construction of a normalizing flow could be invertible linear-time transformations. Their general form is given by:

$$f(z) = \mathbf{z} + \mathbf{u} h_{\theta} (\mathbf{w}^{\mathsf{T}} \mathbf{z} + b)$$

with parameters $\lambda = \{\mathbf{w}, \mathbf{u}, b, \theta\}$ and h smooth nonliterary, NOTE[I don't know why as i would assume that it also has to be invertible for f to be invertible] by application of the theory discussed in section 2.

$$q(z_n) = q(z_0) \prod_k \left| \det J(f_k \circ z_{k-1}) \right|^{-1}$$
$$= q(z_0) \prod_k \left| 1 + \det(\mathbf{u}h'(z_{k-1})^\mathsf{T}\mathbf{w}) \right|^{-1}$$
$$= q(z_0) \prod_k \left| 1 + \mathbf{u}^\mathsf{T}h'(z_{k-1})\mathbf{w} \right|^{-1}$$

a good pick here could be the SE kernel:

$$h_{\alpha}(x) = \exp(-\frac{x^2}{\alpha})$$
$$h'_{\alpha}(x) = -\exp(-\frac{x^2}{\alpha}) \cdot 2 \cdot \frac{x}{\alpha}$$

4. OPTIMIZATION

When we then to fit the posterior to the variational family we need an objective function. I.e., how do we determine what is a good fit. A common choice is the KL-divergence:

(1)
$$D_{KL}(q_{\theta} || p) = \int q_{\theta}(z) \ln \frac{q_{\theta}(z)}{p(z)} dz$$
$$= \ln p(x) - L[q_{\theta}]$$

Thus we get:

$$\ln p(x) = D_{\mathrm{KL}}(q_{\theta} \mid\mid p) + L[q_{\theta}]$$

Here $\ln p\left(x\right)$ is the marginal likelihood and thus does not does not depend on q_{θ} and is thus a constant w.r.t. θ as such maximizing $L\left[q_{\theta}\right]$ over θ , minimizes $D_{\mathrm{KL}}(q_{\theta}\mid\mid p)$ so we will resort to maximizing $L\left[q_{\theta}\right]$ due to their equivalence.

$$\begin{split} L\left[q_{\theta}\right] &= \mathbb{E}_{q_{\phi}(z_{n})}\left[\ln q_{\theta}\left(z_{n}\right)\right] - \mathbb{E}_{q_{\phi}(z_{n})}\left[\ln p\left(x, z_{n}\right)\right] \\ &= \mathbb{E}_{q_{0}(z_{0})}\left[\ln q_{\theta}\left(f_{(n)} \circ z_{0}\right)\right] - \mathbb{E}_{q_{0}(z_{0})}\left[\ln p\left(x, f_{(n)} \circ z_{0}\right)\right] \\ &= \mathbb{E}_{q_{0}(z_{0})}\left[\ln q\left(z_{0}\right)\right] - \sum_{k} \mathbb{E}_{q_{0}(z_{0})}\left[\ln \left|\det J\left(f_{k} \circ f_{(k-1)} \circ z_{0}\right)\right|\right] \\ &- \mathbb{E}_{q_{0}(z_{0})}\left[\ln p\left(x, f_{(n)} \circ z_{0}\right)\right] \end{split}$$

As $\mathbb{E}_{q_0(z_0)}[\ln q(z_0)] \perp \phi$ we can leave it out as it will not effect the optimization, we can further apply the reparameterization trick to $z_0 \sim q_0$, and then estimate the gradients using that:

$$\begin{split} L\left[q_{\theta}\right] &= -\sum_{k} \mathbb{E}_{q_{0}(z_{0})} \left[\ln\left|\det J\left(f_{k}\circ f_{(k-1)}\circ z_{0}\right)\right|\right] \\ &- \mathbb{E}_{q_{0}(z_{0})} \left[\ln p\left(x, f_{(n)}\circ z_{0}\right)\right] \\ &= -\sum_{k} \mathbb{E}_{q(\varepsilon)} \left[\ln\left|\det J\left(f_{k}\circ f_{(k-1)}\circ g\left(\lambda, \varepsilon\right)\right)\right|\right] \\ &- \mathbb{E}_{q(\varepsilon)} \left[\ln p\left(x, f_{(n)}\circ g\left(\lambda, \varepsilon\right)\right)\right] \\ \nabla_{\phi} L\left[q_{\theta}\right] &= -\sum_{k} \int q\left(\varepsilon\right) \nabla_{\phi} \ln\left|\det J\left(f_{k}^{\phi_{k}}\circ f_{k-1}^{\phi_{k-1}}\circ g\left(\lambda, \varepsilon\right)\right)\right| d\varepsilon \\ &- \int q\left(\varepsilon\right) \nabla_{\phi} \ln p\left(x, f_{(n)}^{\phi}\circ g\left(\lambda, \varepsilon\right)\right) d\varepsilon \\ &\approx -\sum_{k} \sum_{n \in N} \nabla_{\phi} \ln\left|\det J\left(f_{k}^{\phi_{k}}\circ f_{k-1}^{\phi_{k-1}}\circ g\left(\lambda, \varepsilon_{n}\right)\right)\right| \\ &- \sum_{n \in N} \nabla_{\phi} \ln p\left(x, f_{(n)}^{\phi}\circ g\left(\lambda, \varepsilon_{n}\right)\right), \quad \left\{\varepsilon_{n}\right\}^{|N|} \sim q\left(\varepsilon\right) \end{split}$$

Where we will use autograd to evaluate $\nabla_{\phi} \ln \left| \det J \left(f_k^{\phi_k} \circ f_{k-1}^{\phi_{k-1}} \circ g \left(\lambda, \varepsilon_n \right) \right) \right|$ and $\nabla_{\phi} \ln p \left(x, f_{(n)}^{\phi} \circ g \left(\lambda, \varepsilon_n \right) \right)$ at ε_n , this will allow us to recursively approximate the direction of greatest ascent and we can then use something like the Adam algorithm to fit the flow our posterior.

Algorithm 1 Variational Inf. with Normalizing Flows

Parameters: ϕ variational, θ generative

while not converged do

$$egin{aligned} oldsymbol{x} &\leftarrow \{ ext{Get mini-batch} \} \ oldsymbol{z}_0 &\sim q_0\left(ullet|oldsymbol{x}
ight) \ oldsymbol{z}_k \leftarrow f_K \circ f_{K-1} \circ \ldots \circ f_1\left(oldsymbol{z}_0
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abla_{\phi} \mathcal{F}\left(oldsymbol{x}\right) \end{aligned}$$

end while

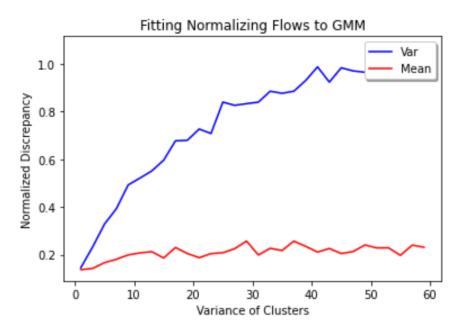


FIGURE 1. Experimentation on fitting planar flows on Gaussian Mixture Model with increasing variance and comparing mean (red) and variance (blue) of the two. Here the x-axis denotes the difference in variance $\sigma_{\rm NF}^2 - \sigma_{\rm gmm}^2$ and the y-axis denotes the difference in mean $|\mu_{\rm gmm} - \mu_{\rm NF}|$

- 4.1. **Optimization of planar flows.** from the derived formula above we can easily find the
- 4.2. **Training.** Furthermore, for Planar and Radial flows training is limited by a single entry point ... [3, 4]
- 4.3. **Metrics.** ELBO / KL divergence difference of means difference of variance k-hat

$$\left\| \Sigma^{-1} (\mu - \hat{\mu}) \right\|$$
$$\left\| I - \Sigma^{-1} \hat{\Sigma} \right\|_{F}$$

5. RESULTS

Pyro[5] was used to perform stochastic variational inference automatically and ArviZ[6] was used to generate simple statistics, summaries and visualizations for the inference results.

5.1. **Multivatiate normal.** Nulla malesuada portitior diam. Donec felis erat, congue non, volutpat at, tincidunt tristique, libero. Vivamus viverra fermentum felis. Donec nonummy pellentesque ante. Phasellus adipiscing semper elit. Proin fermentum massa ac quam. Sed diam turpis, molestie vitae, placerat a, molestie nec, leo. Maecenas lacinia. Nam ipsum ligula, eleifend at, accumsan nec, suscipit a, ipsum. Morbi blandit ligula feugiat magna. Nunc eleifend consequat lorem. Sed lacinia nulla vitae enim. Pellentesque tincidunt purus vel magna. Integer non enim. Praesent euismod nunc eu purus. Donec bibendum quam

FIGURE 2. Final ELBO of normalizing flow as a function of number of flows for correlated multivariate Normal posterior. The final ELBO was calculated as the mean of the ELBOs of the last 500 iterations from training.

TABLE 1. Final ELBO for different potential energy posteriors from Rezende and Mohamed [1] (table 2).

		Planar			Radial	
	2	8	32	2	8	32
$U_1(z)$	73.50	256.24	279.24	-61.84	236.32	272.28
$U_2(z)$	271.71	501.16	659.48	-94.02	184.09	550.71
$U_3(z)$	418.16	666.00	967.83	211.54	576.92	761.59
$U_4(z)$	473.85	638.39	750.76	229.20	462.36	648.31

FIGURE 3. Average ELBO of the last 1000 iterations for ...

FIGURE 4. 360.0pt

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With covariance matrix

5.2. **Energy functions.** Lorem ipsum dolor sit amet, consectetuer adipiscing elit. Ut purus elit, vestibulum ut, placerat ac, adipiscing vitae, felis. Curabitur dictum gravida mauris. Nam arcu libero, nonummy eget, consectetuer id, vulputate a, magna. Donec vehicula augue eu neque. Pellentesque habitant morbi tristique senectus et netus et malesuada fames ac turpis egestas. Mauris ut leo. Cras viverra metus rhoncus sem. Nulla et lectus vestibulum urna fringilla ultrices. Phasellus eu tellus sit amet tortor gravida placerat. Integer sapien est, iaculis in, pretium quis, viverra ac, nunc. Praesent eget sem vel leo ultrices bibendum. Aenean faucibus. Morbi dolor nulla, malesuada eu, pulvinar at, mollis ac, nulla. Curabitur auctor semper nulla. Donec varius orci eget risus. Duis nibh mi, congue eu, accumsan eleifend, sagittis quis, diam. Duis eget orci sit amet orci dignissim rutrum.

The normalizing flows were trained for 10000 epochs with an ADAM learning rate of 0.005, 16 Markov Chain gradient estimation samples and 256 base distribution samples $(z_0 \sim q_0 \, (\bullet | x))$.

Rezende and Mohamed [1] used normalizing flows for variational inference on four target posterior distributions, defined as in table 2.

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TABLE 2. Test energy functions.

$$\begin{split} & \frac{\textbf{Potential }U(\mathbf{z})}{\textbf{1:} \ \frac{1}{2} \left(\frac{\|\mathbf{z}\|-2}{0.4}\right)^2 - \ln \left(e^{-\frac{1}{2}\left[\frac{\mathbf{z}_1-2}{0.6}\right]^2} + e^{-\frac{1}{2}\left[\frac{\mathbf{z}_1+2}{0.6}\right]^2}\right)}{\textbf{2:} \ \frac{1}{2} \left[\frac{\mathbf{z}_2-w_1(\mathbf{z})}{0.4}\right]^2 \\ \textbf{3:} - \ln \left(e^{-\frac{1}{2}\left[\frac{\mathbf{z}_2-w_1(\mathbf{z})}{0.35}\right]^2} + e^{-\frac{1}{2}\left[\frac{\mathbf{z}_2-w_1(\mathbf{z})+w_2(\mathbf{z})}{0.35}\right]^2\right)} \\ \textbf{4:} - \ln \left(e^{-\frac{1}{2}\left[\frac{\mathbf{z}_2-w_1(\mathbf{z})}{0.4}\right]^2} + e^{-\frac{1}{2}\left[\frac{\mathbf{z}_2-w_1(\mathbf{z})+w_3(\mathbf{z})}{0.35}\right]^2\right)} \\ & \text{with } w_1(\mathbf{z}) = \sin \left(\frac{2\pi\mathbf{z}_1}{4}\right), w_2(\mathbf{z}) = 3e^{-\frac{1}{2}\left[\frac{(\mathbf{z}_1-1)}{0.6}\right]^2}, \\ & w_3(\mathbf{z}) = 3\sigma \left(\frac{\mathbf{z}_1-1}{0.3}\right) \text{ and } \sigma(x) = 1/(1+e^{-x}). \end{split}$$

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TABLE 3. Metrics

Number of flows		2			8			32	
ELBO samples	32	128	512	32	128	512	32	128	512
ELBO									
$\mu - \hat{\mu}$ $\sigma^2 - \hat{\sigma}^2$									
$\sigma^2 - \hat{\sigma}^2$ \hat{k}									
<u>κ</u>									

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From the log term in eq. (1) it is clear that minimizing the KL-divergence heavily penalizes candidate distributions q for which the probability at z, q(z), is high and p(z|x) is

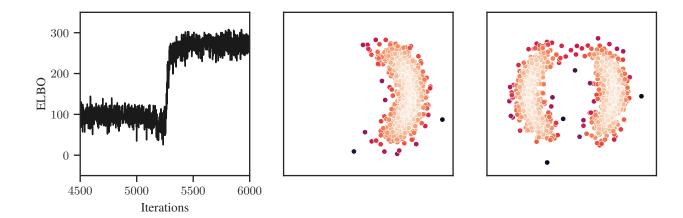


FIGURE 5. Training curve for a normalizing flow with 4 planar flows. When variational inference is performed on this bi-modal target distribution, an 'elbow' forms in the ELBO as the variational approximation changes shape to 'snap' to the new mode[7].

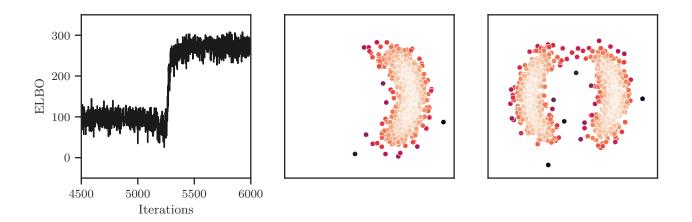


FIGURE 6. Training curve for a normalizing flow with 4 planar flows. When variational inference is performed on this bi-modal target distribution, an 'elbow' forms in the ELBO as the variational approximation changes shape to 'snap' to the new mode[7].

FIGURE 7. Probabilistic graphical model for Poisson regression problem.

low. This is why it takes some time for the second mode to be fitted — the space between the two modes has a low p(z), leading to a penalization on the ELBO for approximations which traverse this space. Indeed, it can be seen in $\ref{eq:condition}$ that in iteration 10000, the variational approximation deviates from a uni-modal approximation, and this iteration corresponds to the area in the training curve in fig. 5 where the ELBO drops, before in the second mode is captured and the ELBO rises again in iteration 11000.

FIGURE 8. Posterior predictive distribution $t_*|\mathbf{t}$

FIGURE 9. Probabilistic graphical model for eight schools problem.

TABLE 4. ELBO and \hat{k} -statistic[12] for different variational inference algorithms trained on the eight schools test problem. It can be seen that there is a loosely monotonic relationship between the ELBO and the \hat{k} -statistic.

		ELBO	\hat{k}	
Mean-fi	eld	-33.406	0.8710	
Full-ran	ık	-32.598	0.8226	
Planar	4	-33.297	0.8329	
	8	-32.728	0.8066	
	16	-32.412	0.7700	
	32	-31.791	0.6379	
Radial	4	-37.756	0.9209	
	8	-37.683	0.8979	
	16	-35.904	0.8760	
	32	-33.292	0.8203	

FIGURE 10. Probabilistic graphical model for non-negative matrix factorization.

5.3. Poisson regression.

5.4. **Eight schools.** The eight schools problem[8] is a classic test problem in Bayesian statistics[9, 10].

The results in table 4 were generated for different variational approximations for 10k iterations, an ADAM [11] learning rate of 10^{-2} , and 256 Monte Carlo ELBO gradient estimation samples.

A normalizing flow variational approximation with 32 planar flows which was trained for 1000 iterations achieved a final ELBO of -32.746 and a \hat{k} -statistic of 0.7723, which means this approximation performs worse than a fully-converged full-rank family. This suggests that a more complex model whose training is terminated before convergence generally performs worse than a simpler model which has been trained to convergence.

5.5. Non-negative matrix factorization.

5.6. **Radon.**

5.7. Bayesian neural network.

6. CONCLUSION

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APPENDIX A. APPENDIX

A.1. The Reparameterization Trick. To elaborate a bit on the reparameterization trick used in (4): if we let $q_0 = \mathcal{N}(\mu, \sigma)$ and thus let $\lambda = \{\mu, \sigma\}$ we can express $z_0 \sim q_0$ as

$$z_0 = \mu + \sigma \varepsilon, \quad \varepsilon \sim \mathcal{N}(0, 1)$$

where