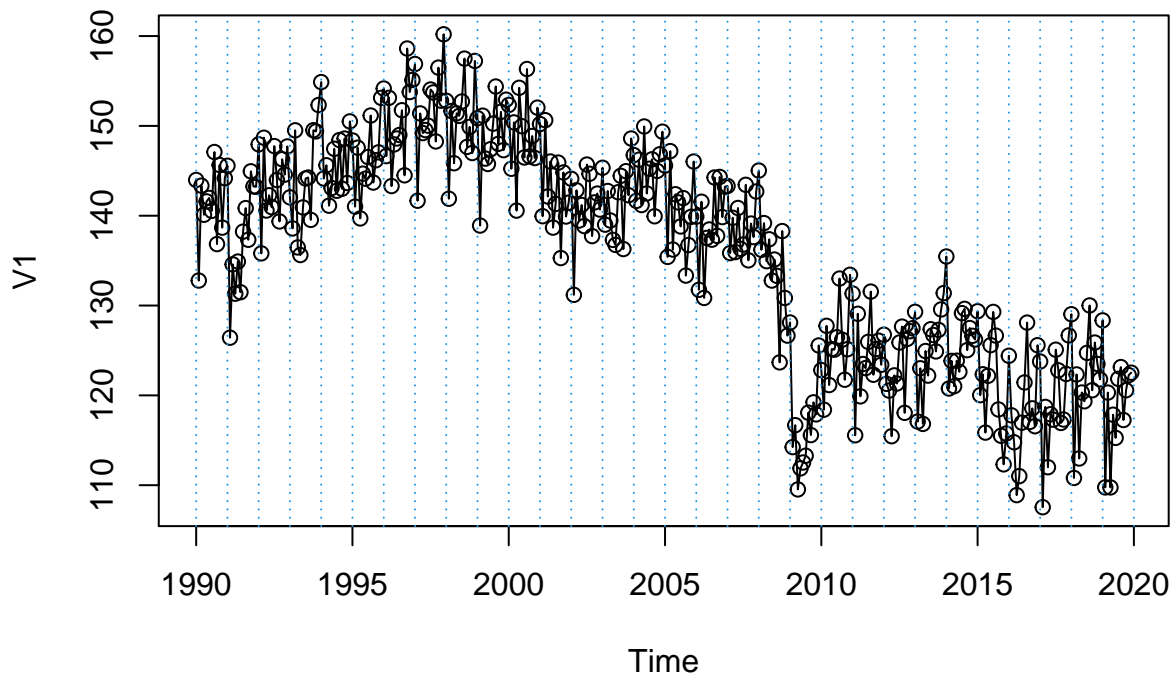


Project Time Series

Carolín Drenda, Silje Anfindsen, Jonathan Stålberg

3 5 2021

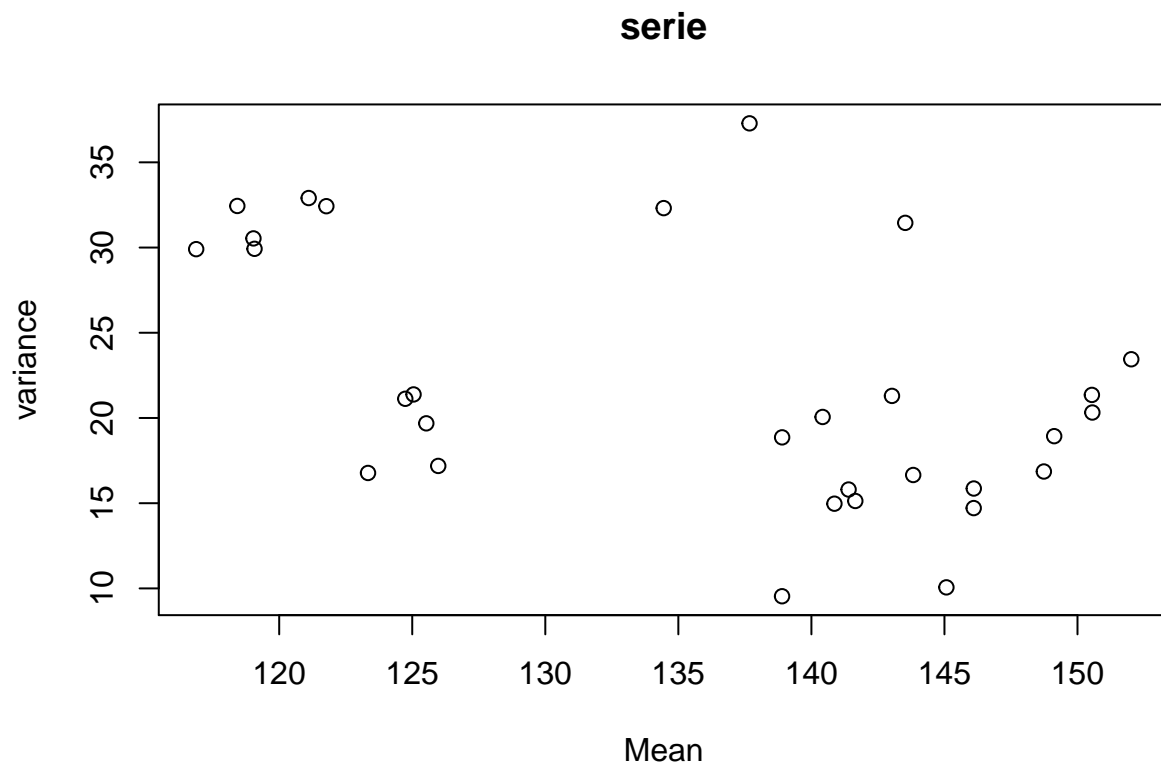
```
serie=ts(read.table("CO2IndUSA.dat"),start=1990,freq=12)  
plot(serie, type = "o")  
abline(v=1990:2020,col=4,lty=3)
```



Identification

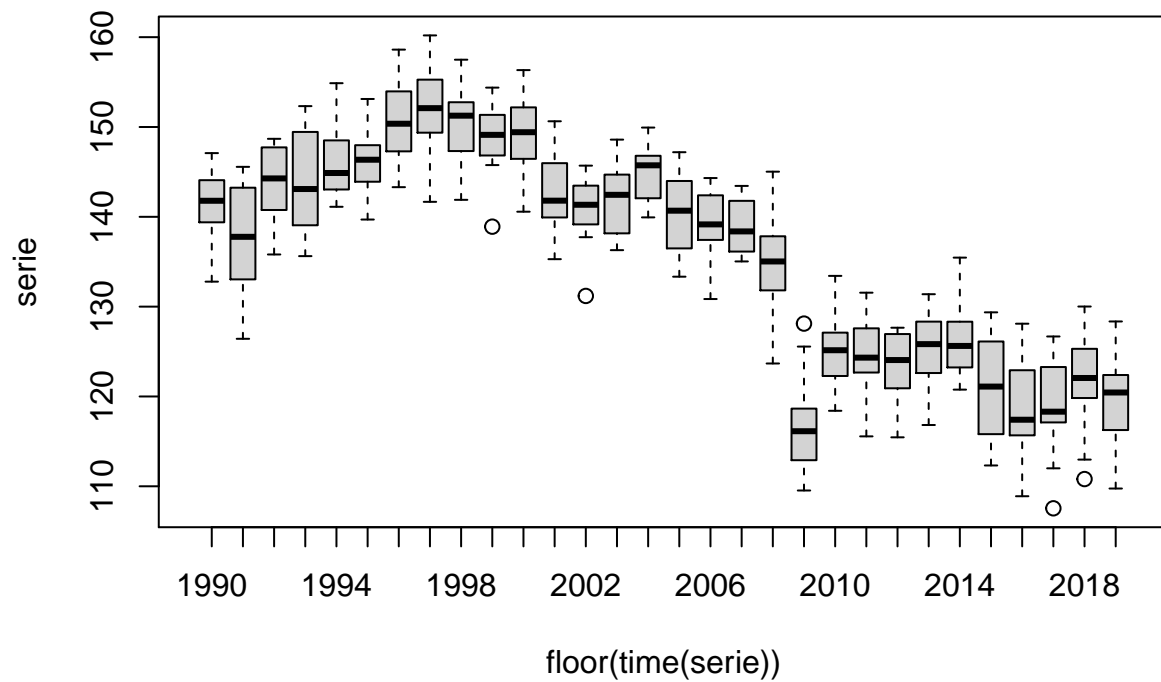
a)

```
m=apply(matrix(serie,nr=12),2,mean)  
v=apply(matrix(serie,nr=12),2,var)  
plot(m,v,xlab="Mean",ylab="variance",main="serie")
```

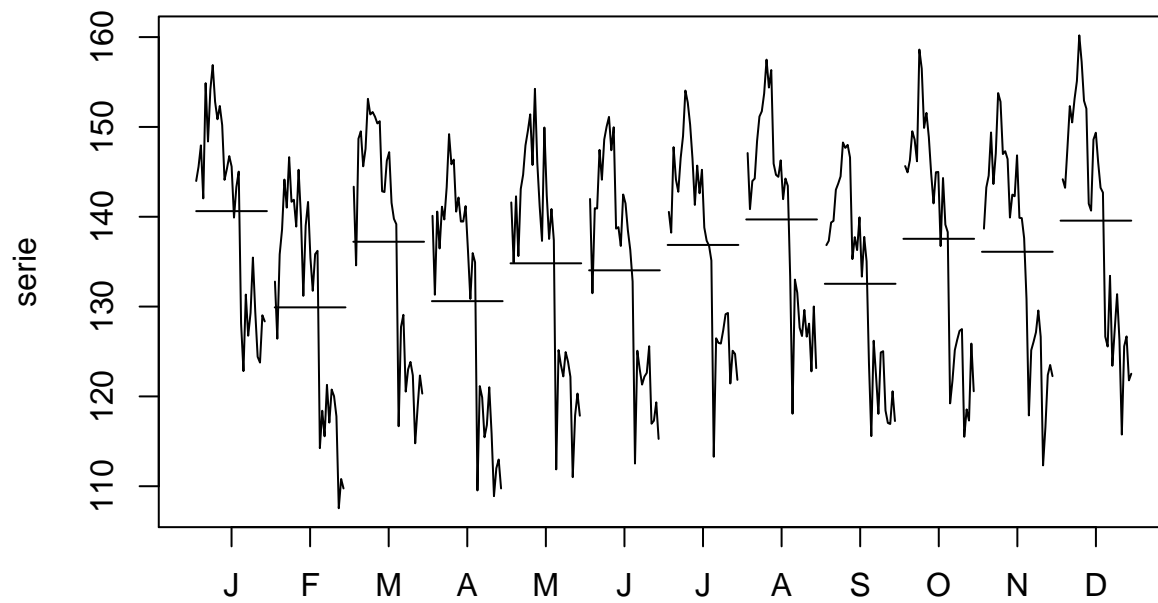


```
#abline(lm(v~m), col=2, lty=3,lwd=2)
```

```
boxplot(serie~floor(time(serie)))
```

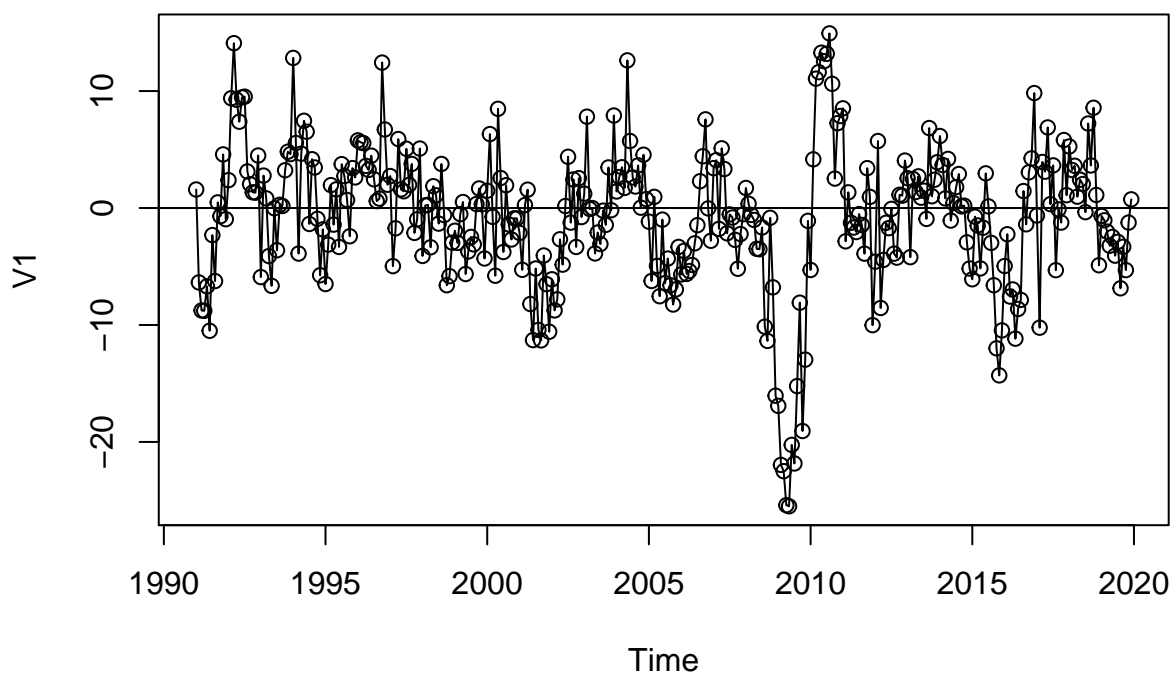


```
monthplot(serie)
```



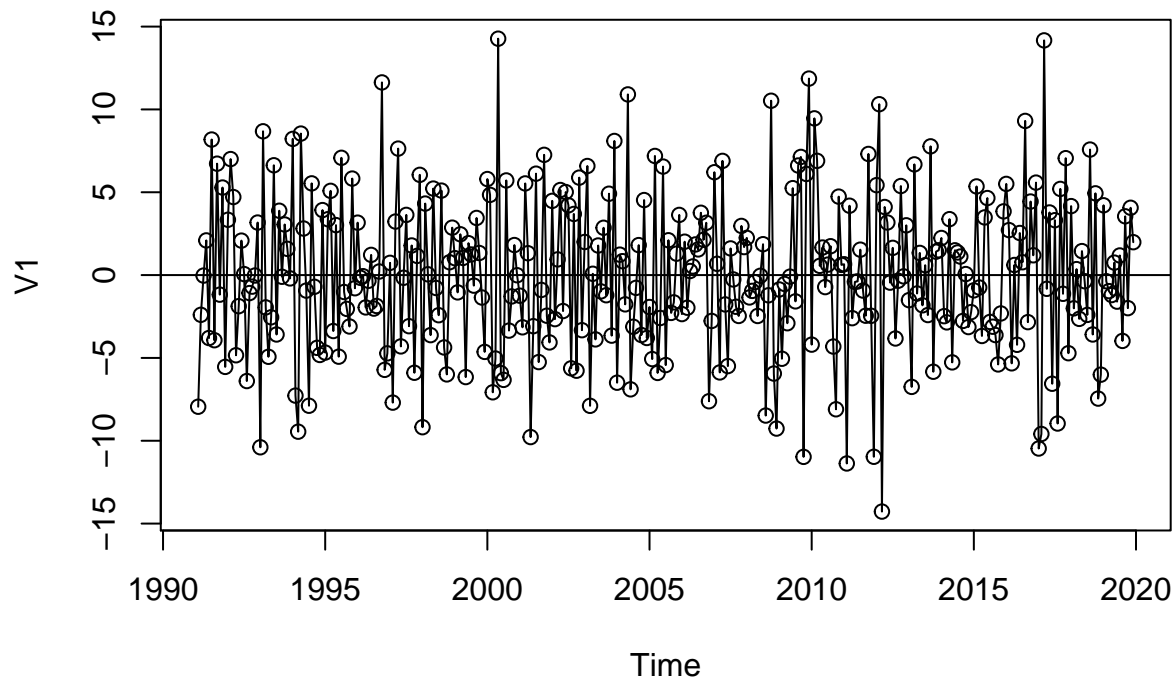
```
d12serie <- diff(serie, 12)
```

```
plot(d12serie, type = "o")
abline(h=0)
```



```
d1d12serie <- diff(d12serie)
```

```
plot(d1d12serie, type = "o")
abline(h=0)
```



```
var(serie)
```

```
##          V1
## V1 150.0063
```

```
var(d12serie)
```

```
##          V1
## V1 37.52583
```

```
var(d1d12serie)
```

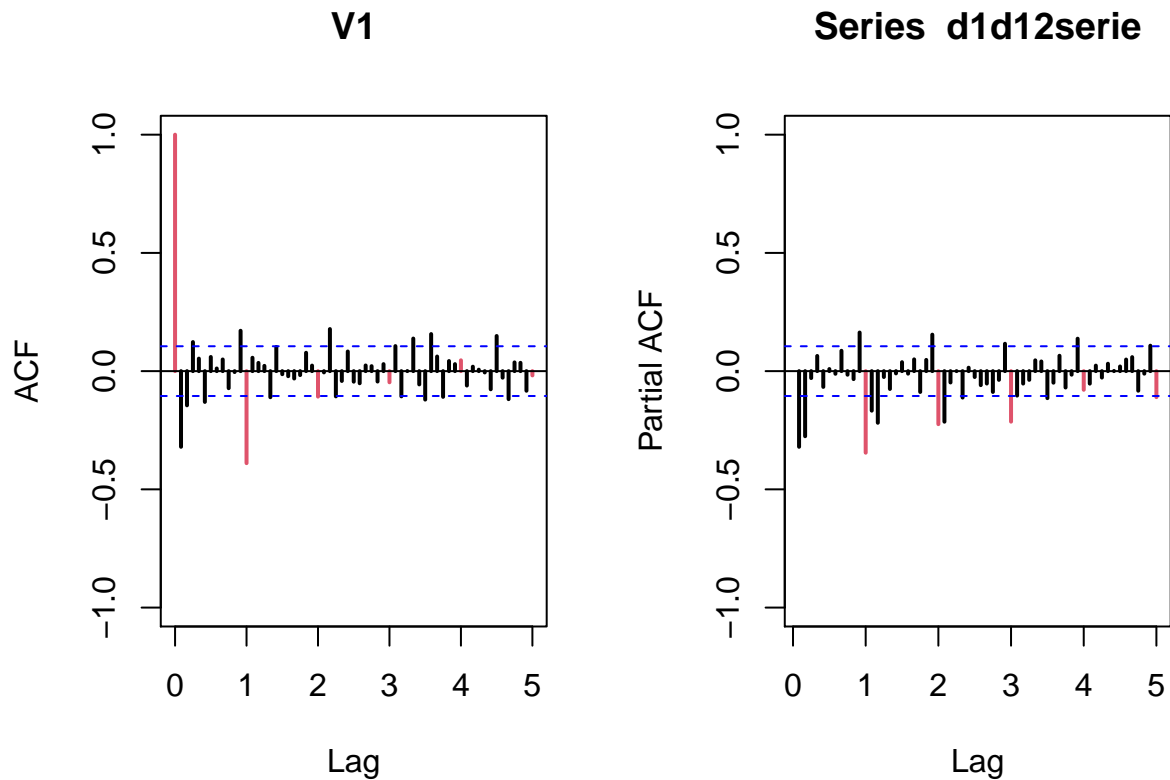
```
##          V1
## V1 22.08084
```

```
var(diff(d1d12serie)) # not needed
```

```
##          V1
## V1 58.26977
```

b)

```
par(mfrow = c(1,2))
acf(d1d12serie, ylim = c(-1,1), lag.max = 60, col = c(2,rep(1,11)), lwd = 2)
pacf(d1d12serie, ylim = c(-1,1), lag.max = 60, col = c(rep(1,11),2), lwd = 2)
```



Seasonal: MA(1),(AR(3)) Regular: ARMA(1,1), Ar(2), MA(1) (2,3)

Estimation

a)

Model 1

```
(mod=arima(d1d12serie, order=c(1,0,1),seasonal=list(order=c(0,0,1),period=12)))

##
## Call:
## arima(x = d1d12serie, order = c(1, 0, 1), seasonal = list(order = c(0, 0, 1),
##      period = 12))
##
## Coefficients:
##          ar1          ma1          sma1  intercept
##      0.0879   -0.5700   -0.8773    -0.0061
## s.e.  0.0969    0.0776    0.0350     0.0141
##
## sigma^2 estimated as 10.75:  log likelihood = -913.37,  aic = 1836.73

mean non-sign
# (mod1=arima(serie, order=c(1,1,1),seasonal=list(order=c(0,1,1),period=12)))
# ar 1 not sign
(mod1=arima(serie, order=c(0,1,1),seasonal=list(order=c(0,1,1),period=12)))

##
## Call:
## arima(x = serie, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1), period = 12))
```

```

##
## Coefficients:
##          ma1      sma1
##        -0.5085 -0.8743
## s.e.    0.0467  0.0348
##
## sigma^2 estimated as 10.79:  log likelihood = -913.87,  aic = 1833.74
validation=function(model,dades){
  s=frequency(get(model$series))
  resid=model$residuals
  par(mfrow=c(2,2),mar=c(3,3,3,3))
  #Residuals plot
  plot(resid,main="Residuals")
  abline(h=0)
  abline(h=c(-3*sd(resid),3*sd(resid)),lty=3,col=4)
  #Square Root of absolute values of residuals (Homocedasticity)
  scatter.smooth(sqrt(abs(resid)),main="Square Root of Absolute residuals",
    lpars=list(col=2))

  #Normal plot of residuals
  qqnorm(resid)
  qqline(resid,col=2,lwd=2)

  ##Histogram of residuals with normal curve
  hist(resid,breaks=20,freq=FALSE)
  curve(dnorm(x,mean=mean(resid),sd=sd(resid)),col=2,add=T)

  ## Individual Correlation Tests
  #ACF & PACF of residuals
  par(mfrow=c(1,2))
  acf(resid,ylim=c(-1,1),lag.max=60,col=c(2,rep(1,s-1)),lwd=1)
  pacf(resid,ylim=c(-1,1),lag.max=60,col=c(rep(1,s-1),2),lwd=1)
  par(mfrow=c(1,1))

  #ACF & PACF of square residuals
  par(mfrow=c(1,2))
  acf(resid^2,ylim=c(-1,1),lag.max=60,col=c(2,rep(1,s-1)),lwd=1)
  pacf(resid^2,ylim=c(-1,1),lag.max=60,col=c(rep(1,s-1),2),lwd=1)
  par(mfrow=c(1,1))

  #Global Correlation Test
  #Ljung-Box p-values
  par(mar=c(2,2,1,1))
  tsdiag(model,gof.lag=7*s)
  cat("\n-----\n")
  print(model)

  #Stationary and Invertible
  cat("\nModul of AR Characteristic polynomial Roots: ",
    Mod(polyroot(c(1,-model$model$phi))),"\n")
  cat("\nModul of MA Characteristic polynomial Roots: ",
    Mod(polyroot(c(1,model$model$theta))),"\n")

```

```

#Model expressed as an MA infinity (psi-weights)
psis=ARMAtoMA(ar=model$model$phi,ma=model$model$theta,lag.max=72)
names(psis)=paste("psi",1:72)
cat("\nPsi-weights (MA(inf))\n")
cat("\n-----\n")
print(psis[1:20])

plot(psis,type="h",main="Pesos Psis - MA infinito")

#Model expressed as an AR infinity (pi-weights)
pis=-ARMAtoMA(ar=-model$model$theta,ma=-model$model$phi,lag.max=72)
names(pis)=paste("pi",1:72)
cat("\nPi-weights (AR(inf))\n")
cat("\n-----\n")
print(pis[1:20])

plot(pis,type="h",main="Pesos Pis - AR infinito")

# #Some Complementary Tests
# cat("\nNormality tests\n")
# cat("\n-----\n")
# ##Shapiro-Wilks Normality test
# print(shapiro.test(resid(model)))
#
# suppressMessages(require(nortest,quietly=TRUE,warn.conflicts=FALSE))
# ##Anderson-Darling test: Normality
# print(ad.test(resid(model)))
#
# suppressMessages(require(tseries,quietly=TRUE,warn.conflicts=FALSE))
# ##Jarque-Bera test: Normality
# print(jarque.bera.test(resid(model)))
#
# cat("\nHomoscedasticity Test\n")
# cat("\n-----\n")
# suppressMessages(require(lmtest,quietly=TRUE,warn.conflicts=FALSE))
# ##Breusch-Pagan test
# obs=get(model$series)
# print(bptest(resid(model)~I(obs-resid(model))))
#
# cat("\nIndependence Tests\n")
# cat("\n-----\n")
#
# ##Durbin-Watson test
# print(dwtest(resid(model)~I(1:length(resid(model)))))
#
##Ljung-Box test
cat("\nLjung-Box test\n")
print(t(apply(matrix(c(1:4,(1:4)*s)),1,function(e1) {
  te=Box.test(resid(model),type="Ljung-Box",lag=e1)
  c(lag=(te$parameter),statistic=te$statistic[[1]],p.value=te$p.value)})))

#Sample ACF vs. Teoric ACF: similar?

```

```

par(mfrow=c(2,2),mar=c(3,3,3,3))
acf(dades, ylim=c(-1,1) ,lag.max=36,main="Sample ACF")

plot(ARMAacf(model$model$phi,model$model$theta,lag.max=36),ylim=c(-1,1),
     type="h",xlab="Lag", ylab="", main="ACF Teoric")
abline(h=0)

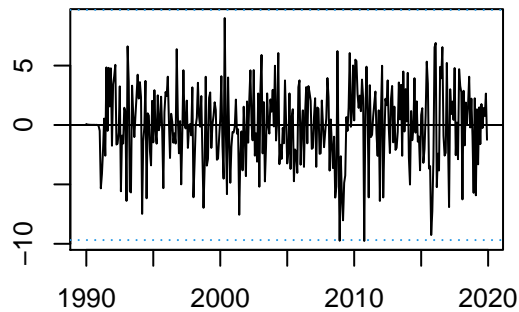
#Sample PACF vs. Teoric PACF
pacf(dades, ylim=c(-1,1) ,lag.max=36,main="Sample PACF")

plot(ARMAacf(model$model$phi,model$model$theta,lag.max=36, pacf=T),ylim=c(-1,1),
     type="h", xlab="Lag", ylab="", main="PACF Teoric")
abline(h=0)
par(mfrow=c(1,1))
}

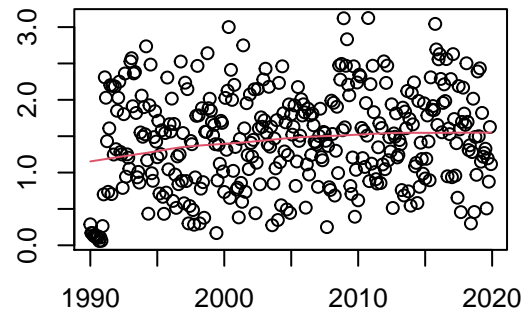
validation(mod1, d1d12serie)

```

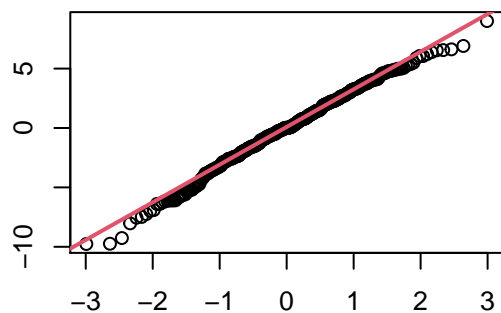
Residuals



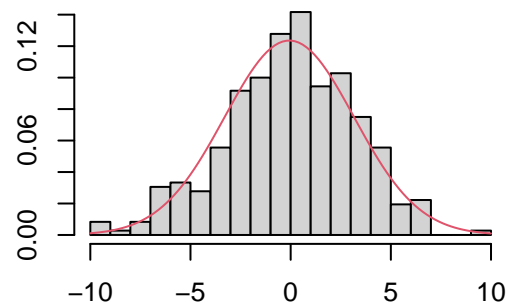
Square Root of Absolute residuals

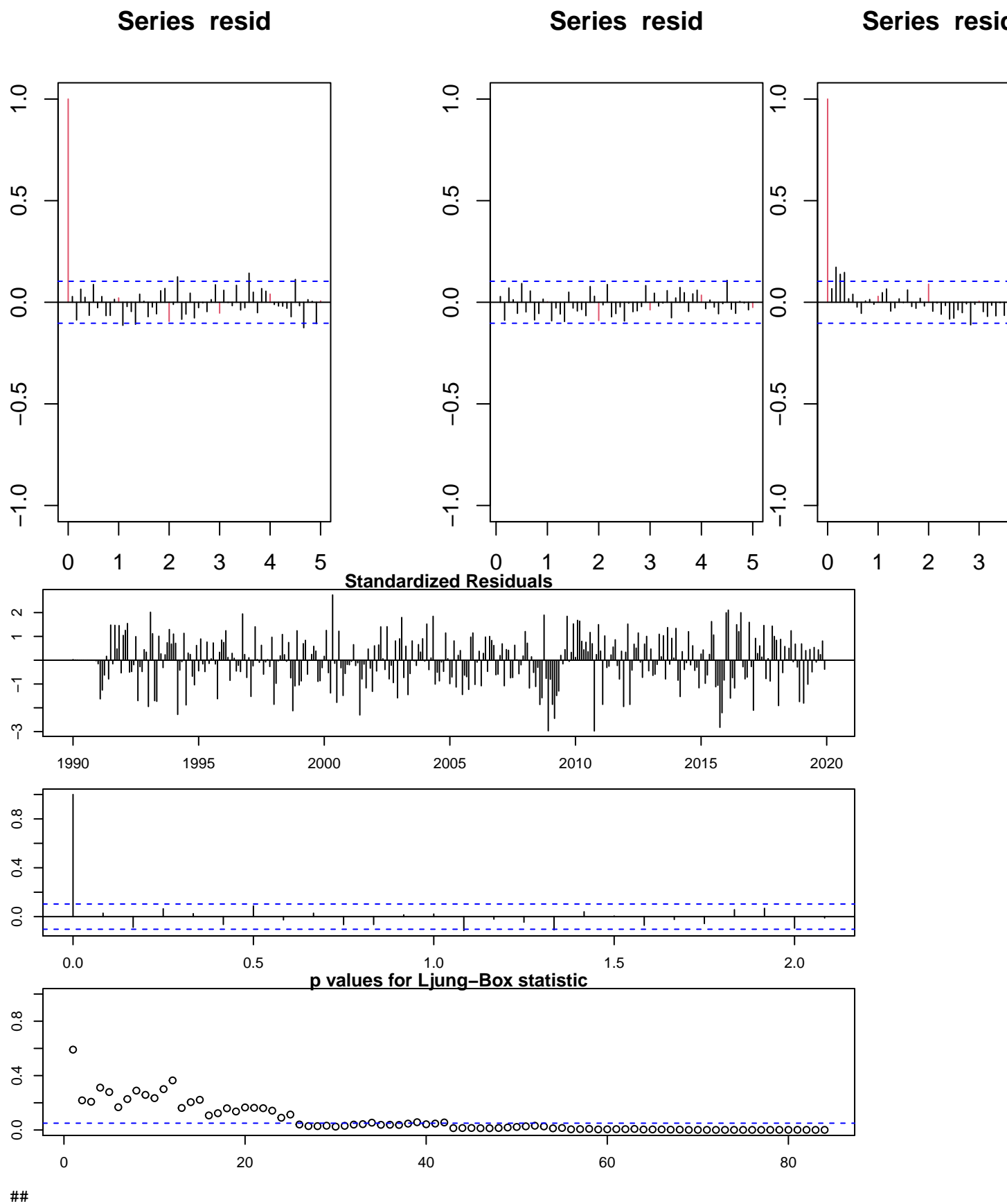


Normal Q-Q Plot

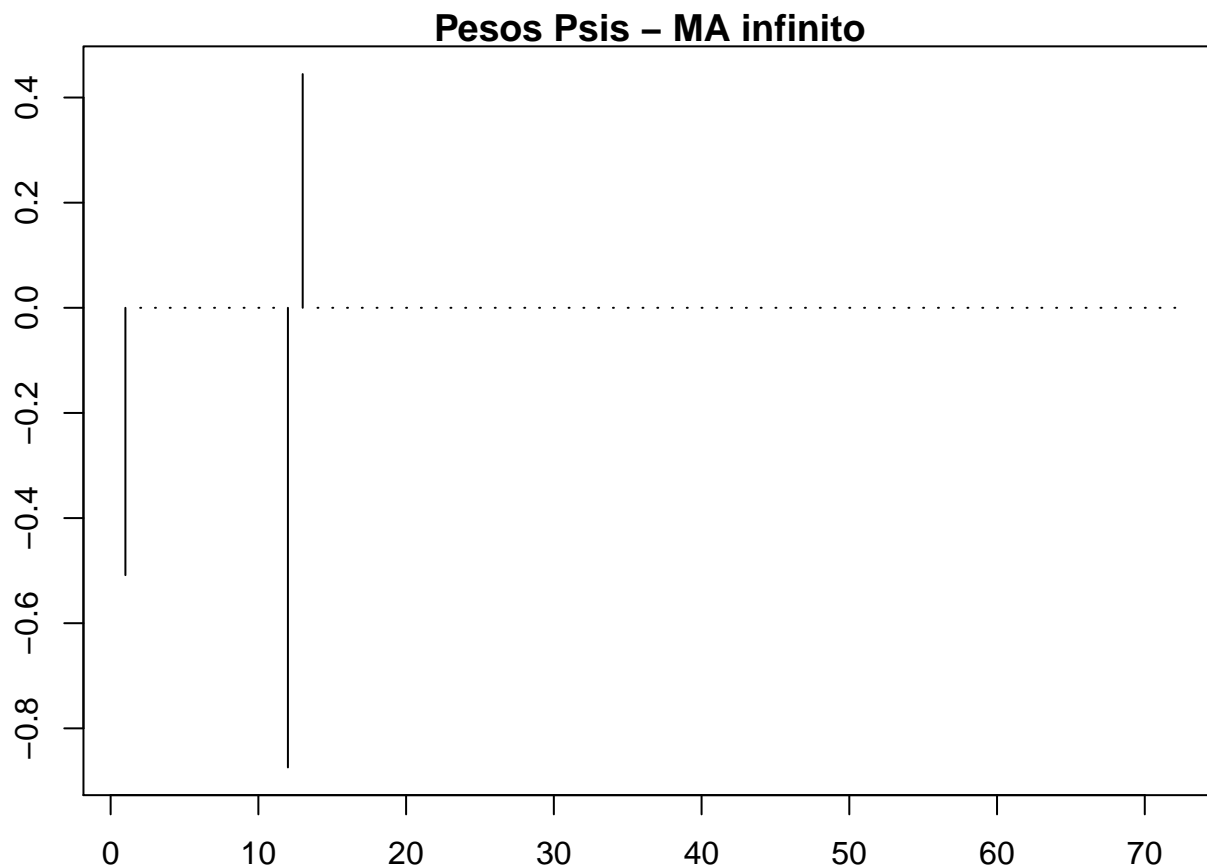


Histogram of resid





```
## -----
##
## Call:
## arima(x = serie, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1), period = 12))
##
## Coefficients:
##          ma1      sma1
##       -0.5085 -0.8743
## s.e.   0.0467  0.0348
##
## sigma^2 estimated as 10.79:  log likelihood = -913.87,  aic = 1833.74
##
## Modul of AR Characteristic polynomial Roots:
##
## Modul of MA Characteristic polynomial Roots:  1.011256 1.011256 1.011256 1.011256 1.011256 1.011256
##
## Psi-weights (MA(inf))
## -----
##      psi 1      psi 2      psi 3      psi 4      psi 5      psi 6      psi 7
## -0.5085041  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000
##      psi 8      psi 9      psi 10     psi 11     psi 12     psi 13     psi 14
##  0.0000000  0.0000000  0.0000000  0.0000000 -0.8743122  0.4445913  0.0000000
##      psi 15     psi 16     psi 17     psi 18     psi 19     psi 20
##  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000
```



```
##
```

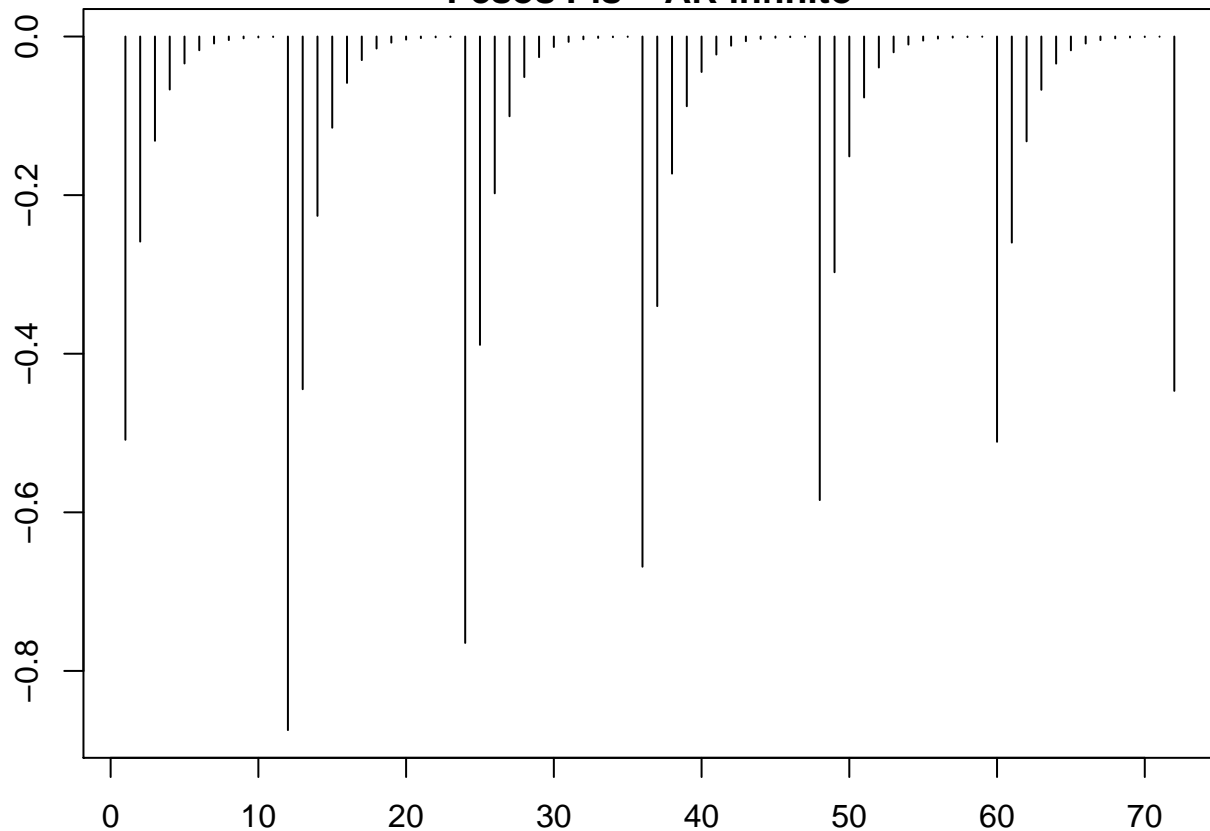
```
## Pi-weights (AR(inf))
```

```
##
```

```
## -----
```

```
##      pi 1      pi 2      pi 3      pi 4      pi 5
## -0.5085040680 -0.2585763871 -0.1314871447 -0.0668617480 -0.0339994708
##      pi 6      pi 7      pi 8      pi 9      pi 10
## -0.0172888692 -0.0087914603 -0.0044704933 -0.0022732641 -0.0011559640
##      pi 11     pi 12     pi 13     pi 14     pi 15
## -0.0005878124 -0.8746110823 -0.4447432932 -0.2261537738 -0.1150001140
##      pi 16     pi 17     pi 18     pi 19     pi 20
## -0.0584780258 -0.0297363140 -0.0151210366 -0.0076891086 -0.0039099430
```

Pesos Pis – AR infinito

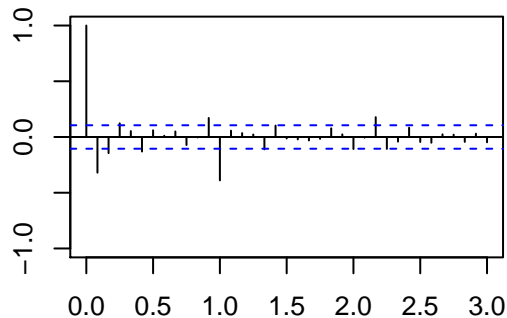


```
##
```

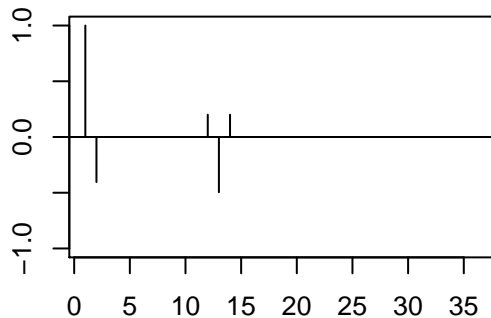
```
## Ljung-Box test
```

```
##      lag.df  statistic    p.value
## [1,]      1  0.2888776  0.59094066
## [2,]      2  3.0474834  0.21789506
## [3,]      3  4.5539267  0.20753030
## [4,]      4  4.7724261  0.31145487
## [5,]     12 13.0621625  0.36453869
## [6,]     24 33.6975781  0.09018904
## [7,]     36 52.3320458  0.03847726
## [8,]     48 71.8636274  0.01443847
```

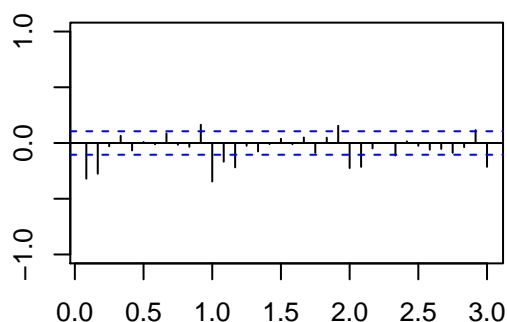
Sample ACF



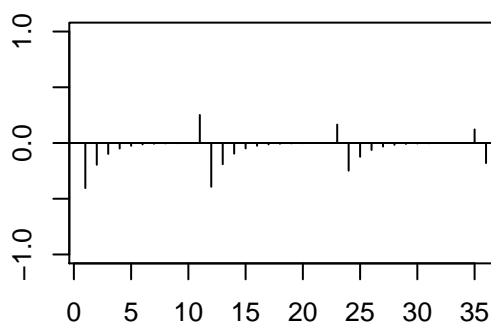
ACF Teoric



Sample PACF



PACF Teoric



Ljung box

is sign -> look for different model

```
(mod2 = arima(d1d12serie, order=c(2,0,0),seasonal=list(order=c(0,0,1),period=12)))
```

```
##
## Call:
## arima(x = d1d12serie, order = c(2, 0, 0), seasonal = list(order = c(0, 0, 1),
##   period = 12))
##
## Coefficients:
##      ar1      ar2      sma1  intercept
##    -0.4480 -0.2910 -0.8565   -0.0062
## s.e.    0.0513   0.0516   0.0351    0.0189
##
## sigma^2 estimated as 10.74:  log likelihood = -912.35,  aic = 1834.7
```

intercept not

#BEST MODEL:

```
(mod3 = arima(serie, order=c(2,1,0),seasonal=list(order=c(0,1,1),period=12)))
```

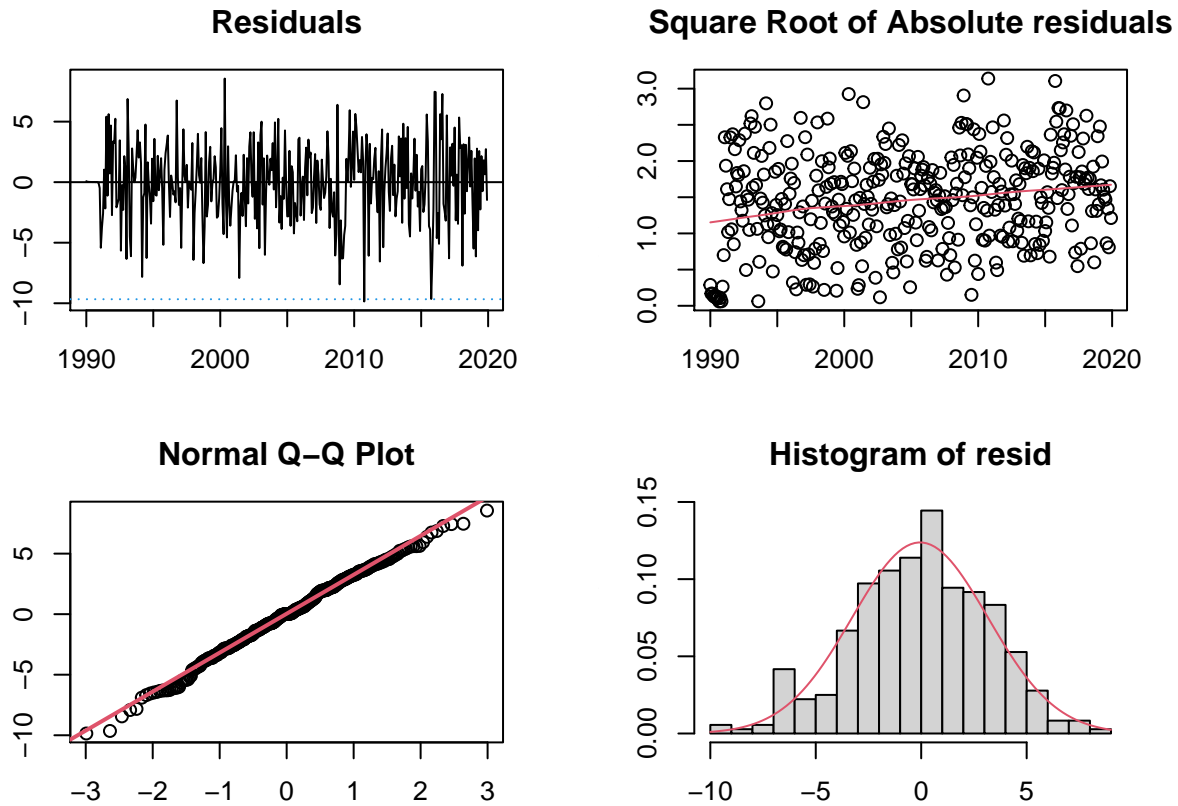
```
##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12))
##
## Coefficients:
##      ar1      ar2      sma1
##    -0.4478 -0.2909 -0.8557
## s.e.    0.0513   0.0516   0.0350
##
```

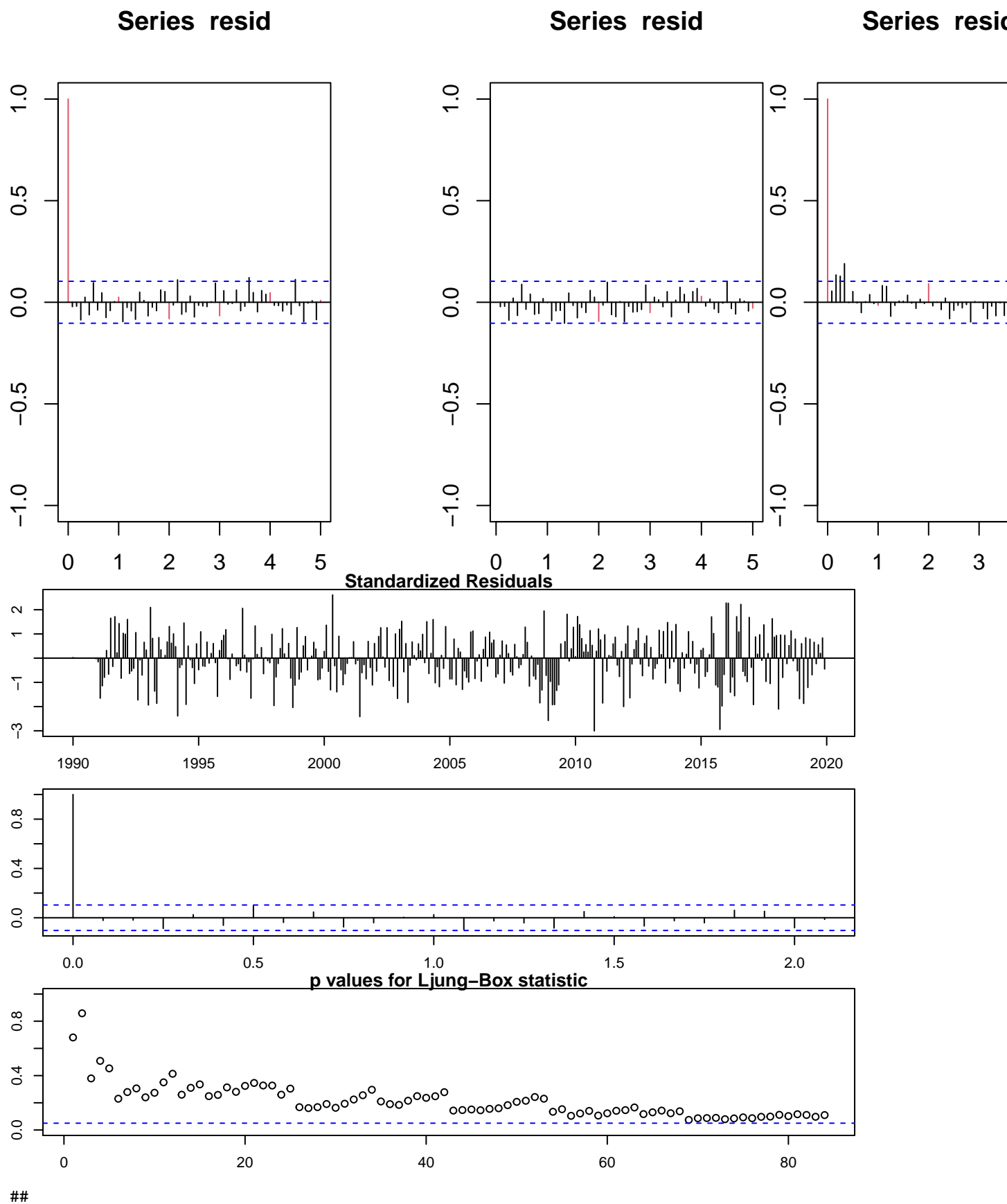
```
## sigma^2 estimated as 10.75: log likelihood = -912.4, aic = 1832.81
```

Validation

a,b,d)

```
validation(mod3,d1d12serie)
```

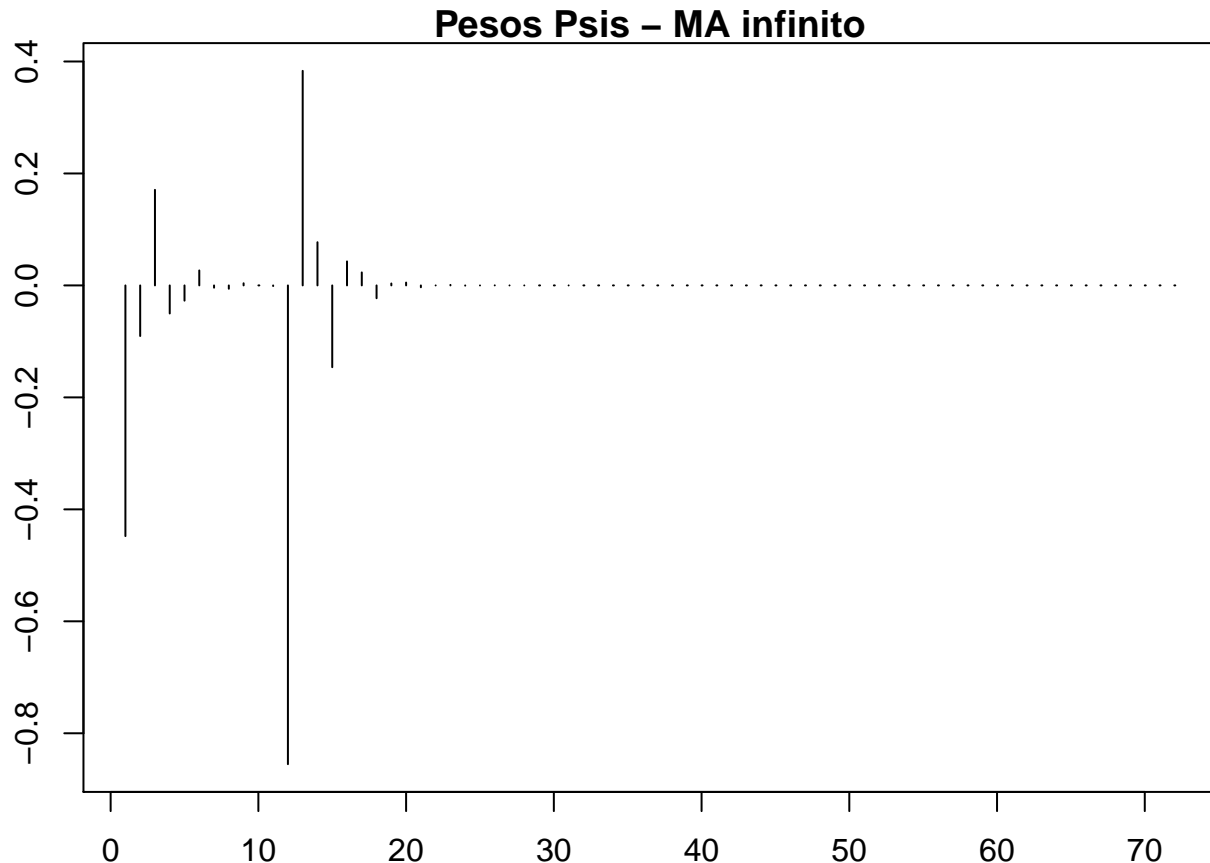




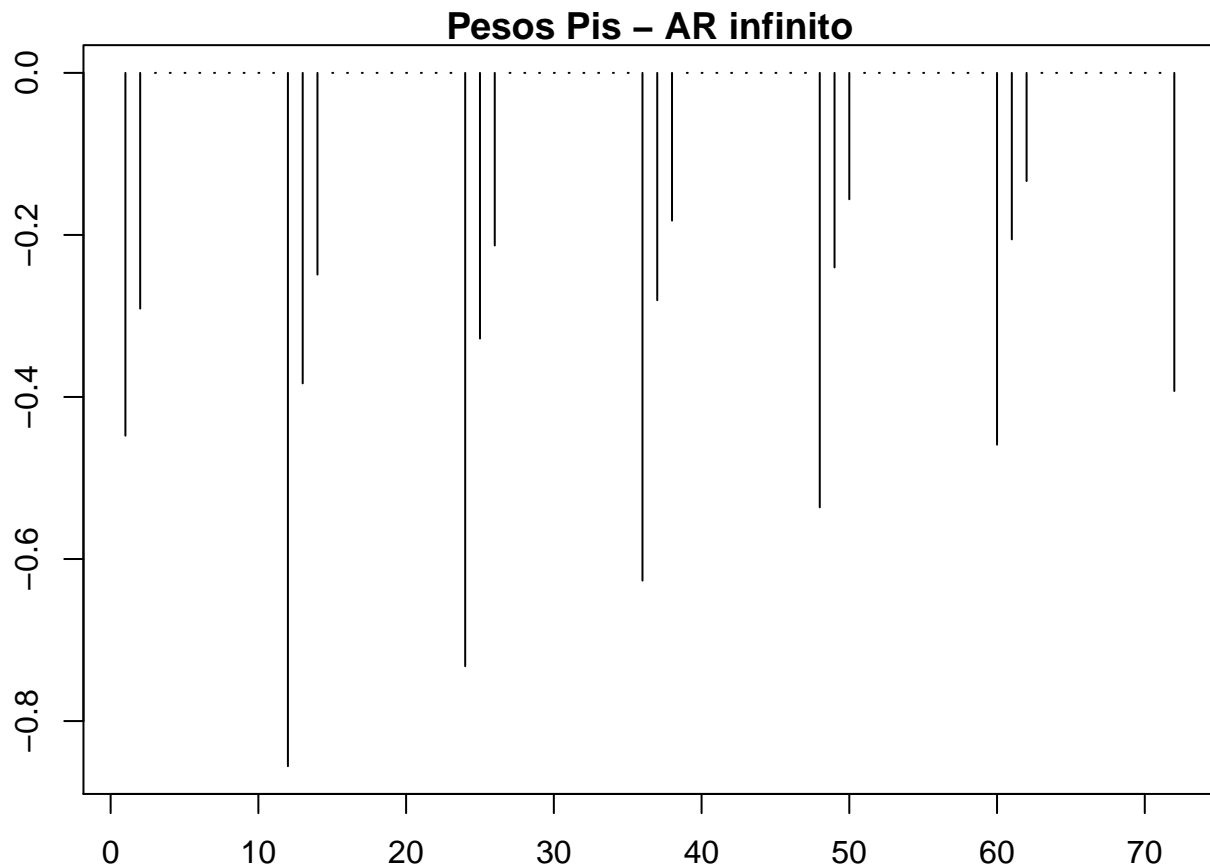
```

## -----
##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12))
##
## Coefficients:
##          ar1          ar2          sma1
##       -0.4478   -0.2909   -0.8557
## s.e.    0.0513    0.0516    0.0350
##
## sigma^2 estimated as 10.75:  log likelihood = -912.4,  aic = 1832.81
##
## Modul of AR Characteristic polynomial Roots:  1.854004 1.854004
##
## Modul of MA Characteristic polynomial Roots:  1.013068 1.013068 1.013068 1.013068 1.013068 1.013068
##
## Psi-weights (MA(inf))
## -----
##          psi 1          psi 2          psi 3          psi 4          psi 5
## -4.477948e-01 -9.040313e-02  1.707560e-01 -5.016326e-02 -2.721404e-02
##          psi 6          psi 7          psi 8          psi 9          psi 10
##  2.677997e-02 -4.074730e-03 -5.966273e-03  3.857100e-03  8.538685e-06
##          psi 11         psi 12         psi 13         psi 14         psi 15
## -1.125944e-03 -8.552243e-01  3.832925e-01  7.716827e-02 -1.460643e-01
##          psi 16         psi 17         psi 18         psi 19         psi 20
##  4.295677e-02  2.325768e-02 -2.291179e-02  3.493580e-03  5.101167e-03

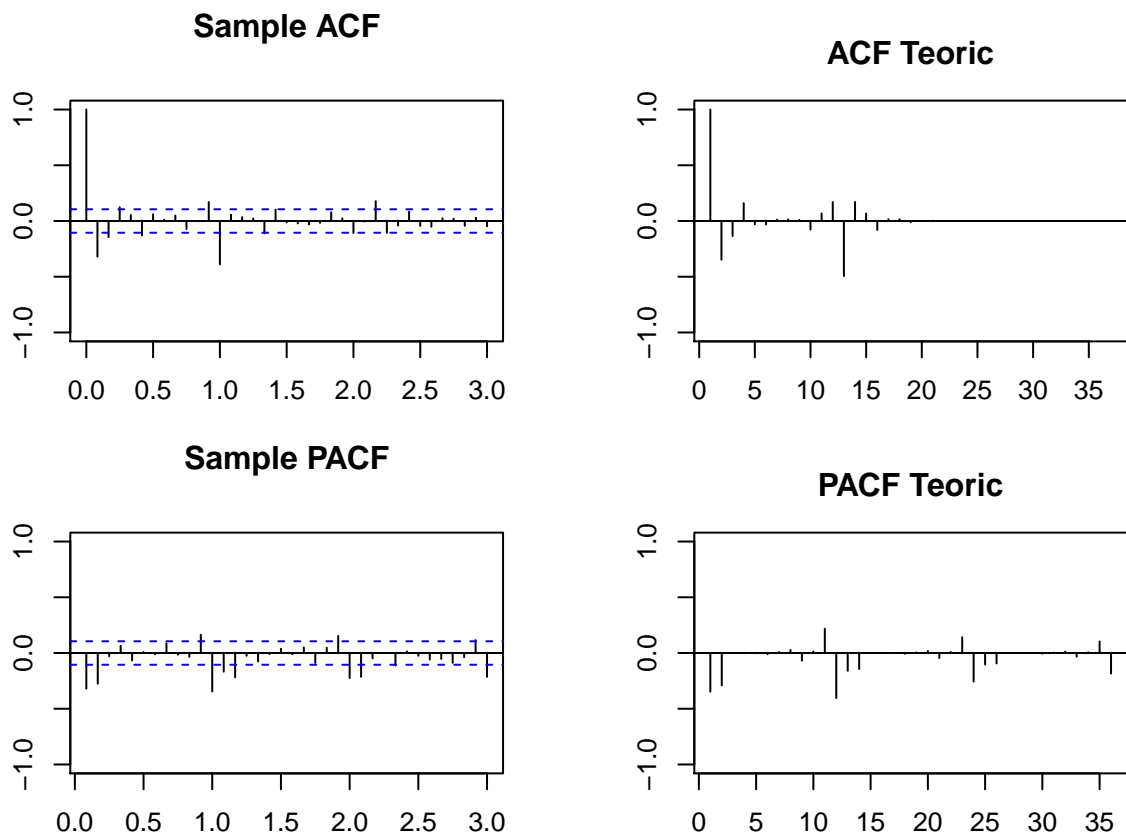
```



```
##
## Pi-weights (AR(inf))
##
## -----
##      pi 1      pi 2      pi 3      pi 4      pi 5      pi 6      pi 7
## -0.4477948 -0.2909233 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
##      pi 8      pi 9      pi 10     pi 11     pi 12     pi 13     pi 14
## 0.0000000 0.0000000 0.0000000 0.0000000 -0.8557260 -0.3831896 -0.2489506
##      pi 15     pi 16     pi 17     pi 18     pi 19     pi 20
## 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
```



```
##
## Ljung-Box test
##      lag.df  statistic  p.value
## [1,]      1  0.1695331 0.6805271
## [2,]      2  0.3058078 0.8582122
## [3,]      3  3.0820338 0.3791493
## [4,]      4  3.3053343 0.5080875
## [5,]     12 12.4093335 0.4133903
## [6,]     24 28.0117983 0.2595426
## [7,]     36 43.2220957 0.1901054
## [8,]     48 57.6881908 0.1595803
```

c) Stability

```

ultim = c(2018,12)
#pdq = c(2,1,0)
#PDQ = c(0,1,1)
serie1=window(serie)
serie2=window(serie,end=ultim)

(mod31= arima(serie1, order=c(2,1,0),seasonal=list(order=c(0,1,1),period=12)))

##
## Call:
## arima(x = serie1, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12))
##
## Coefficients:
##          ar1          ar2          sma1
##      -0.4478   -0.2909   -0.8557
## s.e.    0.0513    0.0516    0.0350
##
## sigma^2 estimated as 10.75:  log likelihood = -912.4,  aic = 1832.81
(mod32= arima(serie2, order=c(2,1,0),seasonal=list(order=c(0,1,1),period=12)))

##
## Call:
## arima(x = serie2, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12))
##

```

```
## Coefficients:
##          ar1      ar2      sma1
##      -0.4292  -0.3068  -0.8673
## s.e.   0.0523   0.0524   0.0356
##
## sigma^2 estimated as 10.8:  log likelihood = -882.47,  aic = 1772.95
```

Problem: shorter TS has lower aic

Prediction

a)

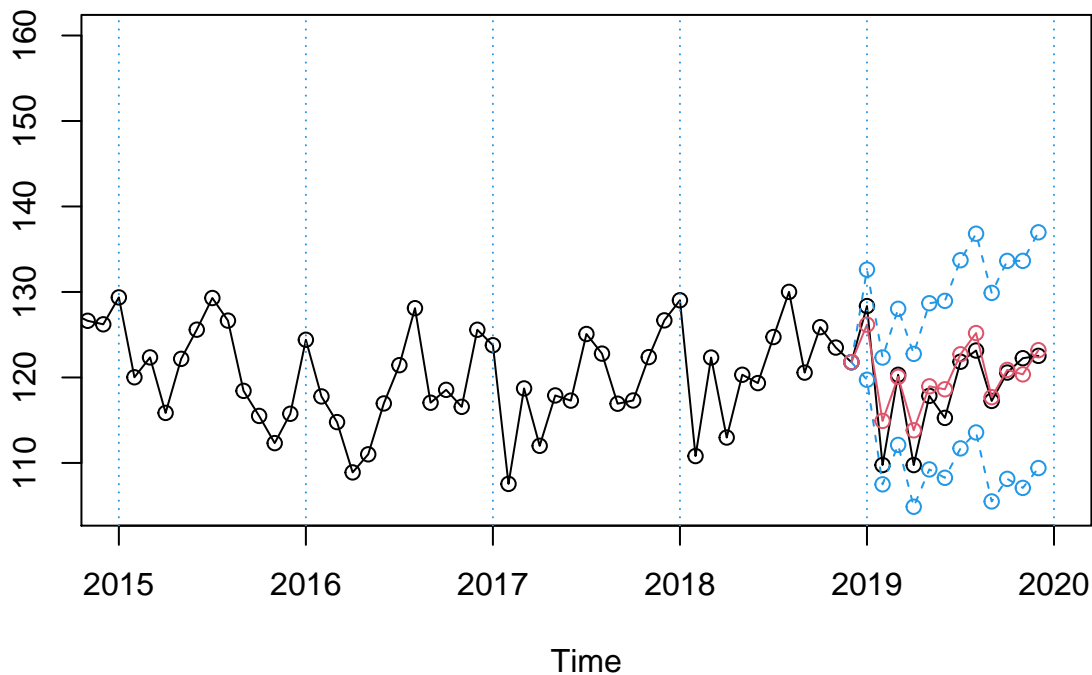
```
pred=predict(mod32,n.ahead=12)
pr<-ts(c(tail(serie2,1),pred$pred),start=ultim,freq=12)

se<-ts(c(0,pred$se),start=ultim,freq=12)

#Intervals
tl<-ts(pr-1.96*se,start=ultim,freq=12)
tu<-ts(pr+1.96*se,start=ultim,freq=12)
pr<-ts(pr,start=ultim,freq=12)

ts.plot(serie,tl,tu,pr,lty=c(1,2,2,1),col=c(1,4,4,2),xlim=ultim[1]+c(-3,+2),type="o",main="Model ARIMA(1,1,1)(0,1,1)12",
abline(v=(ultim[1]-3):(ultim[1]+2),lty=3,col=4))
```

Model ARIMA(1,1,1)(0,1,1)12



```
obs=window(serie,start=ultim)
(mod.EQM1=sqrt(sum(((obs-pr)/obs)^2)/12))
```

```
## [1] 0.02125953
```

```
(mod.EAM1=sum(abs(obs-pr)/obs)/12)
```

```
## [1] 0.01604008
```

```
pred <- predict(mod31,n.ahead=12)
```

```
pr<-ts(c(tail(serie1,1),pred$pred),start=ultim + c(1,0),freq=12)
```

```
se<-ts(c(0,pred$se),start=ultim + c(1,0),freq=12)
```

```
#Intervals
```

```
tl1<-ts(pr-1.96*se,start=ultim + c(1,0),freq=12)
```

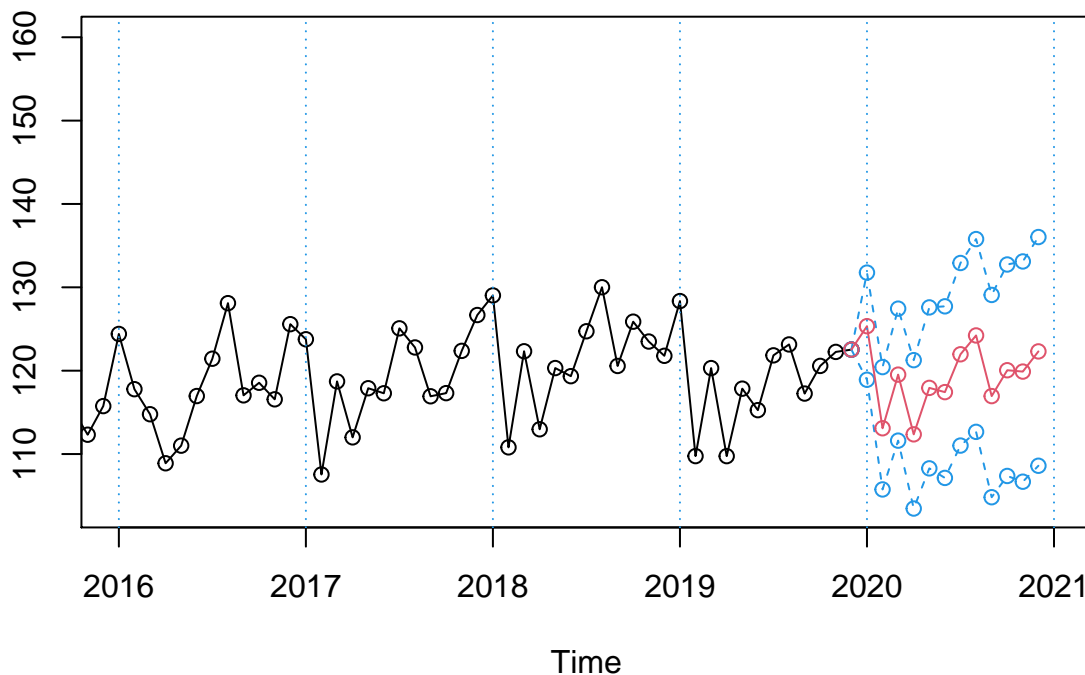
```
tu1<-ts(pr+1.96*se,start=ultim + c(1,0),freq=12)
```

```
pr1<-ts(pr,start=ultim + c(1,0),freq=12)
```

```
ts.plot(serie,tl1,tu1,pr1,lty=c(1,2,2,1),col=c(1,4,4,2),xlim=c(ultim[1]-2,ultim[1]+3),type="o",main="Mo
```

```
abline(v=(ultim[1]-2):(ultim[1]+3),lty=3,col=4)
```

Model ARIMA(2,1,0)(0,1,1)12



```
previs1=window(cbind(tl1,pr1,tu1),start=ultim+c(1,0))
```

Outlier treatment

```
source("CalendarEffects.r")
```

```
source("atipics2.r")
```

a) Calendar effect:

```
data=c(start(serie)[1],start(serie)[2], length(serie)) #starting year, month, series size  
(wTradDays=Wtrad(data)) #creates auxiliary variable for trading days configurations (5/2 the ideal prop
```

##	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
## 1990	3.0	0.0	-0.5	-1.5	3.0	-1.5	-0.5	3.0	-5.0	3.0	2.0	-4.0
## 1991	3.0	0.0	-4.0	2.0	3.0	-5.0	3.0	-0.5	-1.5	3.0	-1.5	-0.5
## 1992	3.0	-2.5	-0.5	2.0	-4.0	2.0	3.0	-4.0	2.0	-0.5	-1.5	3.0
## 1993	-4.0	0.0	3.0	2.0	-4.0	2.0	-0.5	-0.5	2.0	-4.0	2.0	3.0
## 1994	-4.0	0.0	3.0	-1.5	-0.5	2.0	-4.0	3.0	2.0	-4.0	2.0	-0.5
## 1995	-0.5	0.0	3.0	-5.0	3.0	2.0	-4.0	3.0	-1.5	-0.5	2.0	-4.0
## 1996	3.0	1.0	-4.0	2.0	3.0	-5.0	3.0	-0.5	-1.5	3.0	-1.5	-0.5
## 1997	3.0	0.0	-4.0	2.0	-0.5	-1.5	3.0	-4.0	2.0	3.0	-5.0	3.0
## 1998	-0.5	0.0	-0.5	2.0	-4.0	2.0	3.0	-4.0	2.0	-0.5	-1.5	3.0
## 1999	-4.0	0.0	3.0	2.0	-4.0	2.0	-0.5	-0.5	2.0	-4.0	2.0	3.0
## 2000	-4.0	1.0	3.0	-5.0	3.0	2.0	-4.0	3.0	-1.5	-0.5	2.0	-4.0
## 2001	3.0	0.0	-0.5	-1.5	3.0	-1.5	-0.5	3.0	-5.0	3.0	2.0	-4.0
## 2002	3.0	0.0	-4.0	2.0	3.0	-5.0	3.0	-0.5	-1.5	3.0	-1.5	-0.5
## 2003	3.0	0.0	-4.0	2.0	-0.5	-1.5	3.0	-4.0	2.0	3.0	-5.0	3.0
## 2004	-0.5	-2.5	3.0	2.0	-4.0	2.0	-0.5	-0.5	2.0	-4.0	2.0	3.0
## 2005	-4.0	0.0	3.0	-1.5	-0.5	2.0	-4.0	3.0	2.0	-4.0	2.0	-0.5
## 2006	-0.5	0.0	3.0	-5.0	3.0	2.0	-4.0	3.0	-1.5	-0.5	2.0	-4.0
## 2007	3.0	0.0	-0.5	-1.5	3.0	-1.5	-0.5	3.0	-5.0	3.0	2.0	-4.0
## 2008	3.0	1.0	-4.0	2.0	-0.5	-1.5	3.0	-4.0	2.0	3.0	-5.0	3.0
## 2009	-0.5	0.0	-0.5	2.0	-4.0	2.0	3.0	-4.0	2.0	-0.5	-1.5	3.0
## 2010	-4.0	0.0	3.0	2.0	-4.0	2.0	-0.5	-0.5	2.0	-4.0	2.0	3.0
## 2011	-4.0	0.0	3.0	-1.5	-0.5	2.0	-4.0	3.0	2.0	-4.0	2.0	-0.5
## 2012	-0.5	1.0	-0.5	-1.5	3.0	-1.5	-0.5	3.0	-5.0	3.0	2.0	-4.0
## 2013	3.0	0.0	-4.0	2.0	3.0	-5.0	3.0	-0.5	-1.5	3.0	-1.5	-0.5
## 2014	3.0	0.0	-4.0	2.0	-0.5	-1.5	3.0	-4.0	2.0	3.0	-5.0	3.0
## 2015	-0.5	0.0	-0.5	2.0	-4.0	2.0	3.0	-4.0	2.0	-0.5	-1.5	3.0
## 2016	-4.0	1.0	3.0	-1.5	-0.5	2.0	-4.0	3.0	2.0	-4.0	2.0	-0.5
## 2017	-0.5	0.0	3.0	-5.0	3.0	2.0	-4.0	3.0	-1.5	-0.5	2.0	-4.0
## 2018	3.0	0.0	-0.5	-1.5	3.0	-1.5	-0.5	3.0	-5.0	3.0	2.0	-4.0
## 2019	3.0	0.0	-4.0	2.0	3.0	-5.0	3.0	-0.5	-1.5	3.0	-1.5	-0.5

```
(wEast=Weaster(data))
```

##	Jan	Feb	Mar	Apr	May	Jun
## 1990	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1991	0.0000000	0.0000000	0.5000000	-0.5000000	0.0000000	0.0000000
## 1992	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1993	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1994	0.0000000	0.0000000	0.1666667	-0.1666667	0.0000000	0.0000000
## 1995	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1996	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1997	0.0000000	0.0000000	0.5000000	-0.5000000	0.0000000	0.0000000
## 1998	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 1999	0.0000000	0.0000000	0.0000000	0.0000000	0.0000000	0.0000000
## 2000	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2001	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2002	0.0000000	0.0000000	0.5000000	-0.5000000	0.0000000	0.0000000
## 2003	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2004	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2005	0.0000000	0.0000000	0.5000000	-0.5000000	0.0000000	0.0000000
## 2006	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2007	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000
## 2008	0.0000000	0.0000000	0.5000000	-0.5000000	0.0000000	0.0000000
## 2009	0.0000000	0.0000000	-0.5000000	0.5000000	0.0000000	0.0000000

```
## 2010 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2011 0.0000000 0.0000000 -0.5000000 0.5000000 0.0000000 0.0000000
## 2012 0.0000000 0.0000000 -0.5000000 0.5000000 0.0000000 0.0000000
## 2013 0.0000000 0.0000000 0.5000000 -0.5000000 0.0000000 0.0000000
## 2014 0.0000000 0.0000000 -0.5000000 0.5000000 0.0000000 0.0000000
## 2015 0.0000000 0.0000000 -0.1666667 0.1666667 0.0000000 0.0000000
## 2016 0.0000000 0.0000000 0.5000000 -0.5000000 0.0000000 0.0000000
## 2017 0.0000000 0.0000000 -0.5000000 0.5000000 0.0000000 0.0000000
## 2018 0.0000000 0.0000000 0.5000000 -0.5000000 0.0000000 0.0000000
## 2019 0.0000000 0.0000000 -0.5000000 0.5000000 0.0000000 0.0000000
##      Jul      Aug      Sep      Oct      Nov      Dec
## 1990 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1991 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1992 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1993 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1994 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1995 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1996 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1997 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1998 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 1999 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2001 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2002 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2003 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2004 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2005 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2006 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2007 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2008 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2009 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2010 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2011 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2012 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2013 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2014 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2015 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2016 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2017 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2018 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
## 2019 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
```

```
(mod3EC=arima(serie,order=c(2,1,0),seasonal=list(order=c(1,1,0),period=12),xreg=data.frame(wTradDays,wE
```

```
##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(1, 1, 0), period = 12),
##      xreg = data.frame(wTradDays, wEast))
##
## Coefficients:
##      ar1      ar2      sar1 wTradDays      wEast
##    -0.3804 -0.3156 -0.4787     0.2429    -0.3014
## s.e.   0.0510   0.0512   0.0474     0.0424   0.5557
##
## sigma^2 estimated as 13.35:  log likelihood = -943.76,  aic = 1899.52
```

```

(mod3Ea=arima(serie,order=c(2,1,0),seasonal=list(order=c(1,1,0),period=12),xreg=data.frame(wEast)))

##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(1, 1, 0), period = 12),
##       xreg = data.frame(wEast))
##
## Coefficients:
##          ar1          ar2          sar1          wEast
##      -0.4416   -0.3382   -0.4473   -0.5613
## s.e.    0.0507    0.0511    0.0487    0.5946
##
## sigma^2 estimated as 14.55:  log likelihood = -958.41,  aic = 1926.82
(mod3TD=arima(serie,order=c(2,1,0),seasonal=list(order=c(1,1,0),period=12),xreg=data.frame(wTradDays)))

##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(1, 1, 0), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##          ar1          ar2          sar1  wTradDays
##      -0.3802   -0.3174   -0.4786    0.2448
## s.e.    0.0510    0.0511    0.0474    0.0423
##
## sigma^2 estimated as 13.37:  log likelihood = -943.91,  aic = 1897.81

```

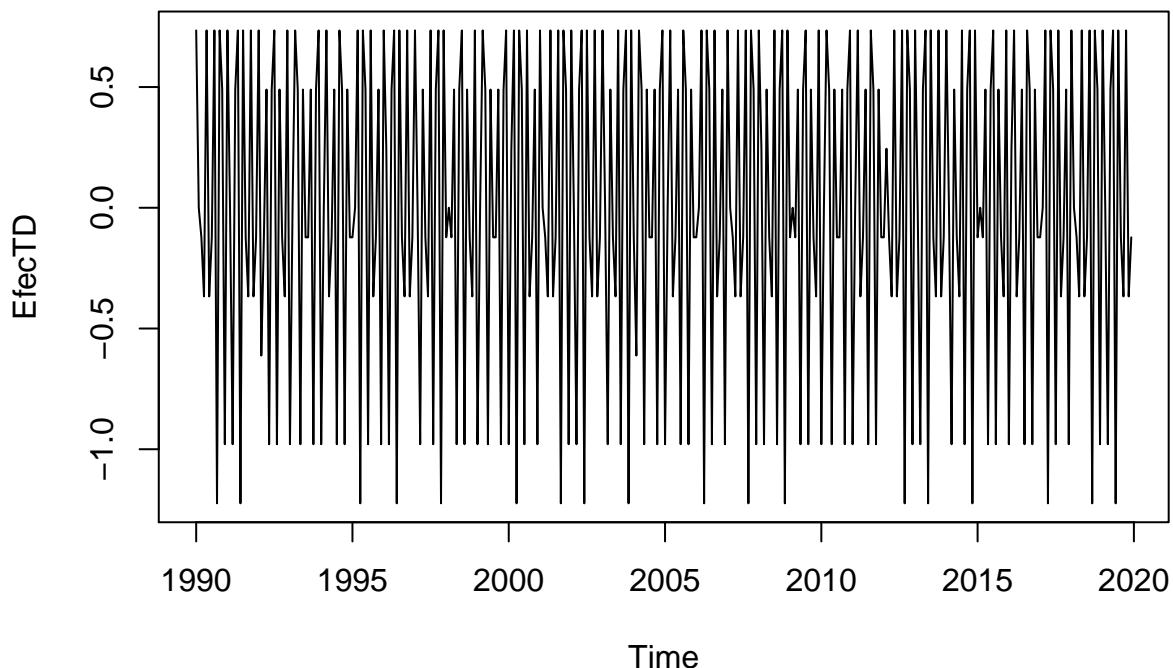
Only use trading days not easter, because lowest aic and other parameters are not significant

Calculate Trading day effect:

```

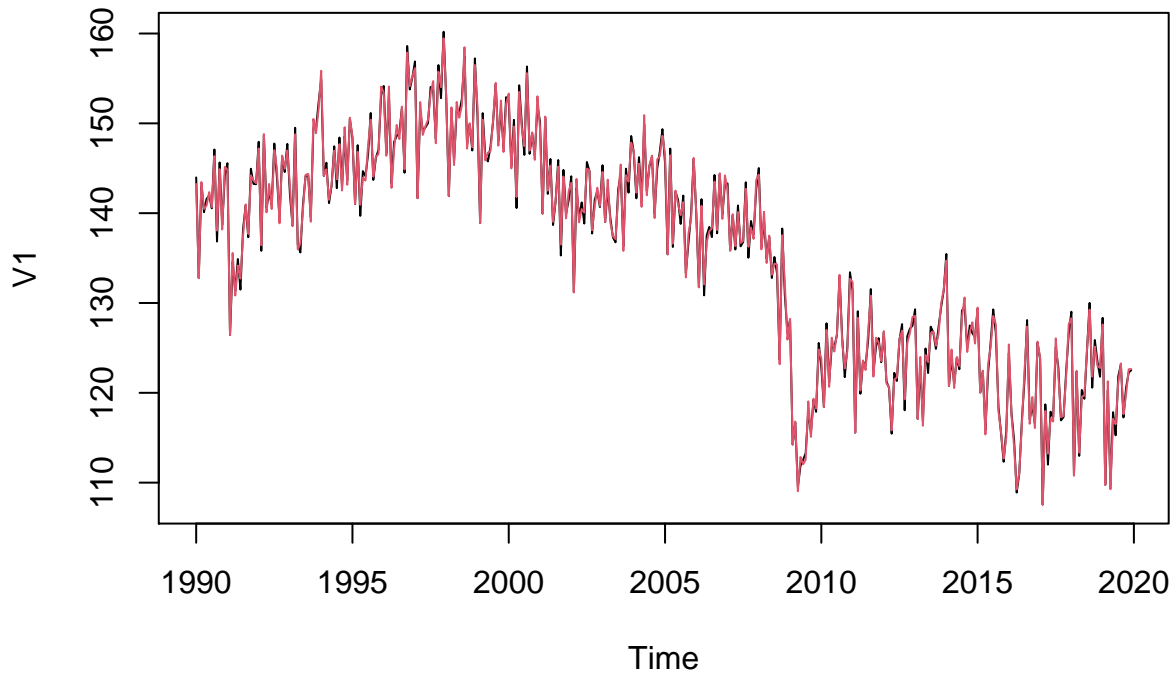
EfecTD=coef(mod3TD)["wTradDays"]*wTradDays
plot(EfecTD)

```



corrected series for trading day effect and compare with original serie:

```
serieTD=serie-EfecTD
plot(serie)
lines(serieTD,col=2)
```



Transform into stationarity:

```
d12serieTD <- diff(serieTD, 12)
d1d12serieTD <- diff(d12serieTD)
```

Variance

```
var(serieTD)
```

```
##          serie
## serie 149.4372
```

```
var(d12serieTD)
```

```
##          serie
## serie 37.02932
```

```
var(d1d12serieTD)
```

```
##          serie
## serie 20.63326
```

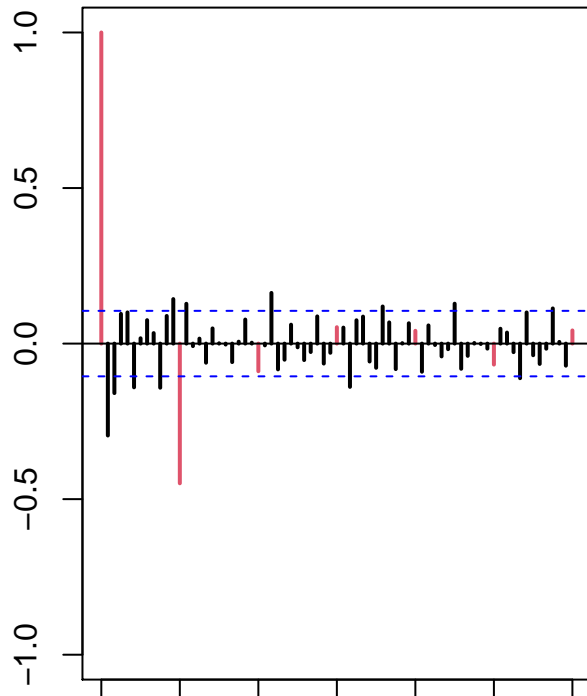
```
var(diff(d1d12serieTD)) # not necessary
```

```
##          serie
## serie 53.43205
```

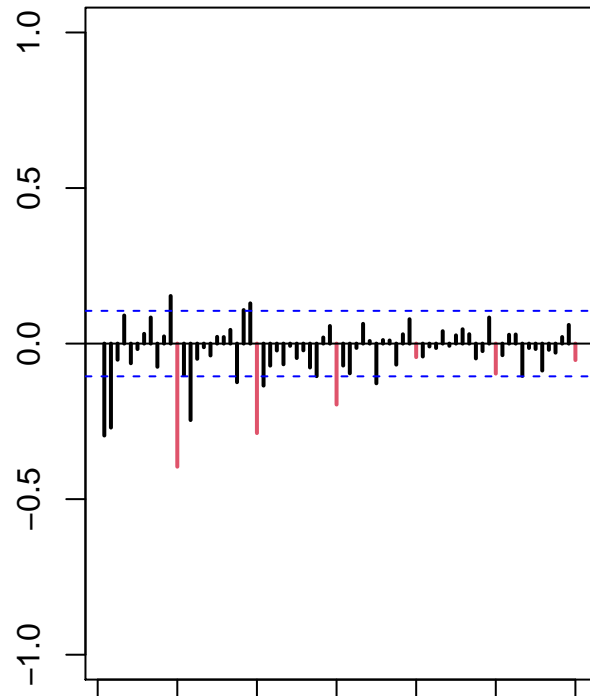
Same Transformation as before.

```
par(mfrow=c(1,2), mar =c(1,2,4,1))
acf(d1d12serieTD,ylim=c(-1,1),lag.max=72,col=c(2,rep(1,11)),lwd=2)
pacf(d1d12serieTD,ylim=c(-1,1),lag.max=72,col=c(rep(1,11),2),lwd=2)
```

serie



Series d1d12serieTD



Modelidentification: seasonal: MA(1) (AR(3)) Regular: MA(2), AR(2)

Mdoel estimation

```
#(moder <- arima(serie, order=c(0,1,1),seasonal=list(order=c(0,1,2),period=12), xreg=data.frame(wTradDays))
# sma 2 not sign
(moder <- arima(serie, order=c(0,1,1),seasonal=list(order=c(0,1,1),period=12), xreg=data.frame(wTradDays))
```

```
##
## Call:
## arima(x = serie, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##          ma1          sma1    wTradDays
##       -0.4677   -0.8652      0.2530
## s.e.    0.0502    0.0343    0.0438
##
## sigma^2 estimated as 9.889:  log likelihood = -898.33,  aic = 1804.66
```

```
# same parameters as before
(moder2 <- arima(serie, order=c(2,1,0),seasonal=list(order=c(1,1,0),period=12), xreg=data.frame(wTradDays))
```

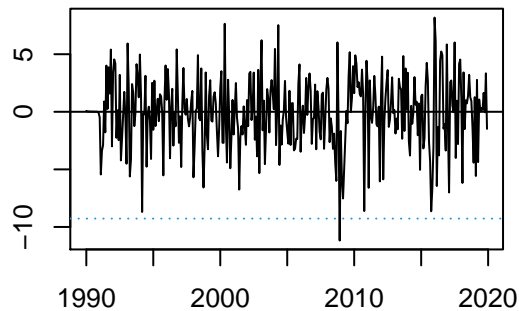
```
##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(1, 1, 0), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##          ar1          ar2          sar1    wTradDays
##       -0.3802   -0.3174   -0.4786      0.2448
```



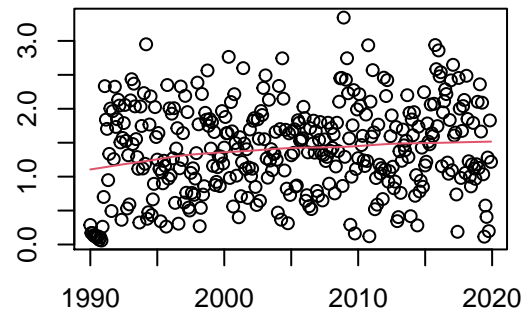
```
## s.e.    0.0510    0.0511    0.0474    0.0423
##
## sigma^2 estimated as 13.37:  log likelihood = -943.91,  aic = 1897.81
(modec3 <- arima(serie, order=c(2,1,0),seasonal=list(order=c(0,1,1),period=12), xreg=data.frame(wTradDays))

##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##          ar1          ar2          sma1  wTradDays
##        -0.3850   -0.2614   -0.8494    0.2556
## s.e.    0.0518    0.0521    0.0344    0.0445
##
## sigma^2 estimated as 9.872:  log likelihood = -897.43,  aic = 1804.86
Prefer modec ARIMA(0,1,1)(0,1,1)_12 + wTradDays since aic is smaller, modec3 with ARIMA(2,1,0)(0,1,1)_12
+ wTradDays also possible
dades=d1d12serieTD
model=modec
validation(model,dades)
```

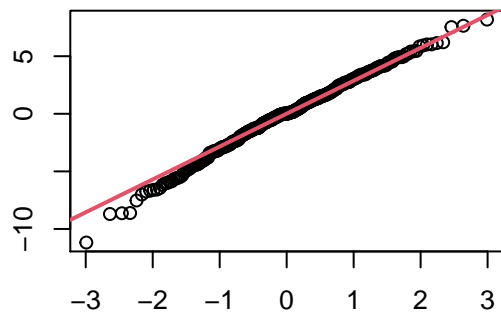
Residuals



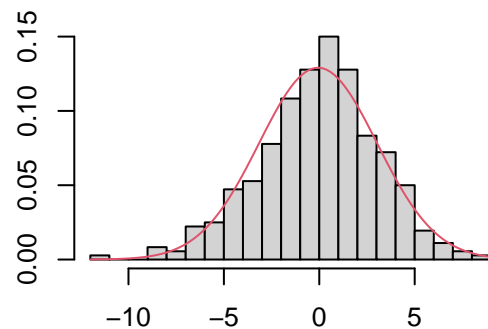
Square Root of Absolute residuals

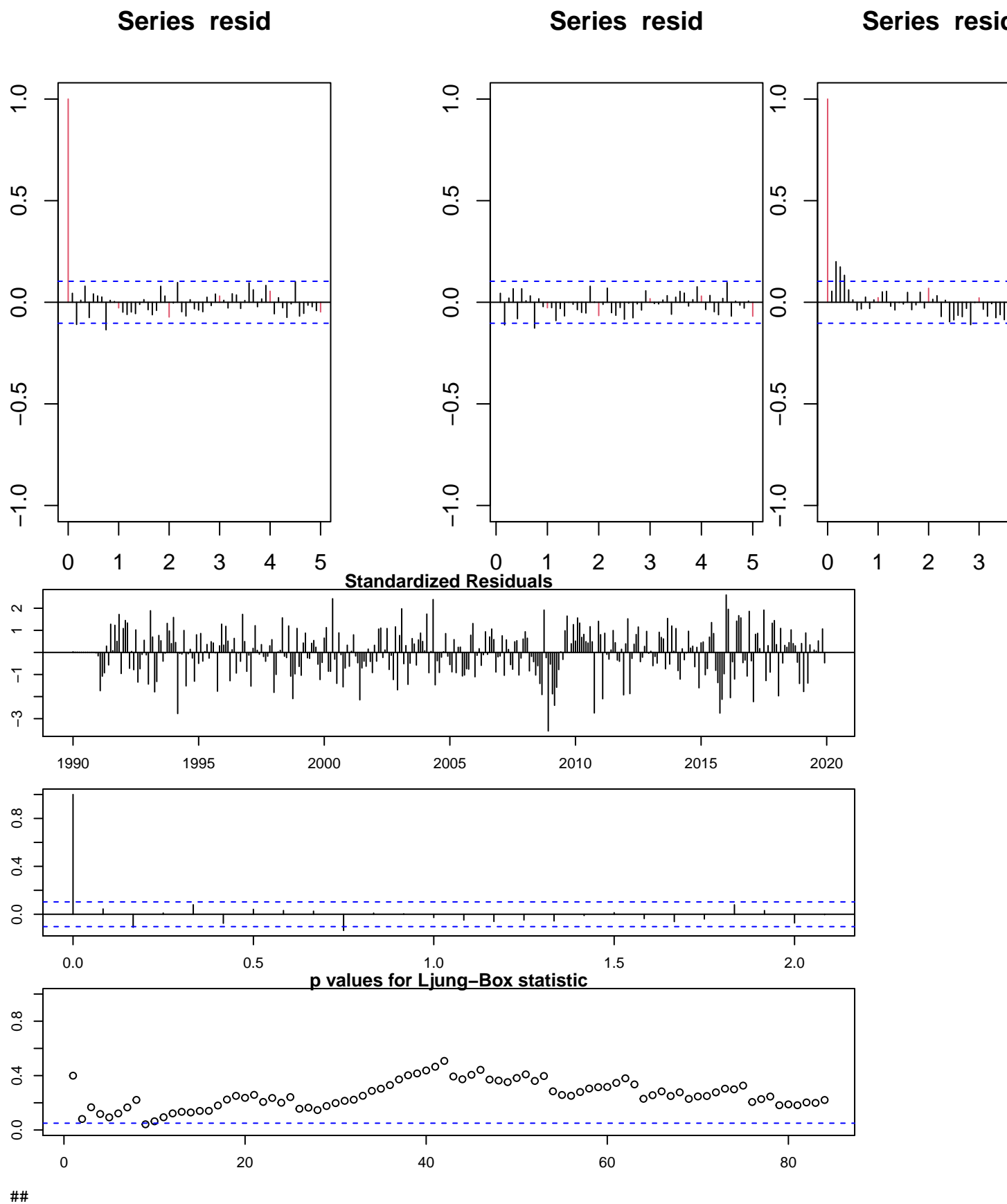


Normal Q-Q Plot

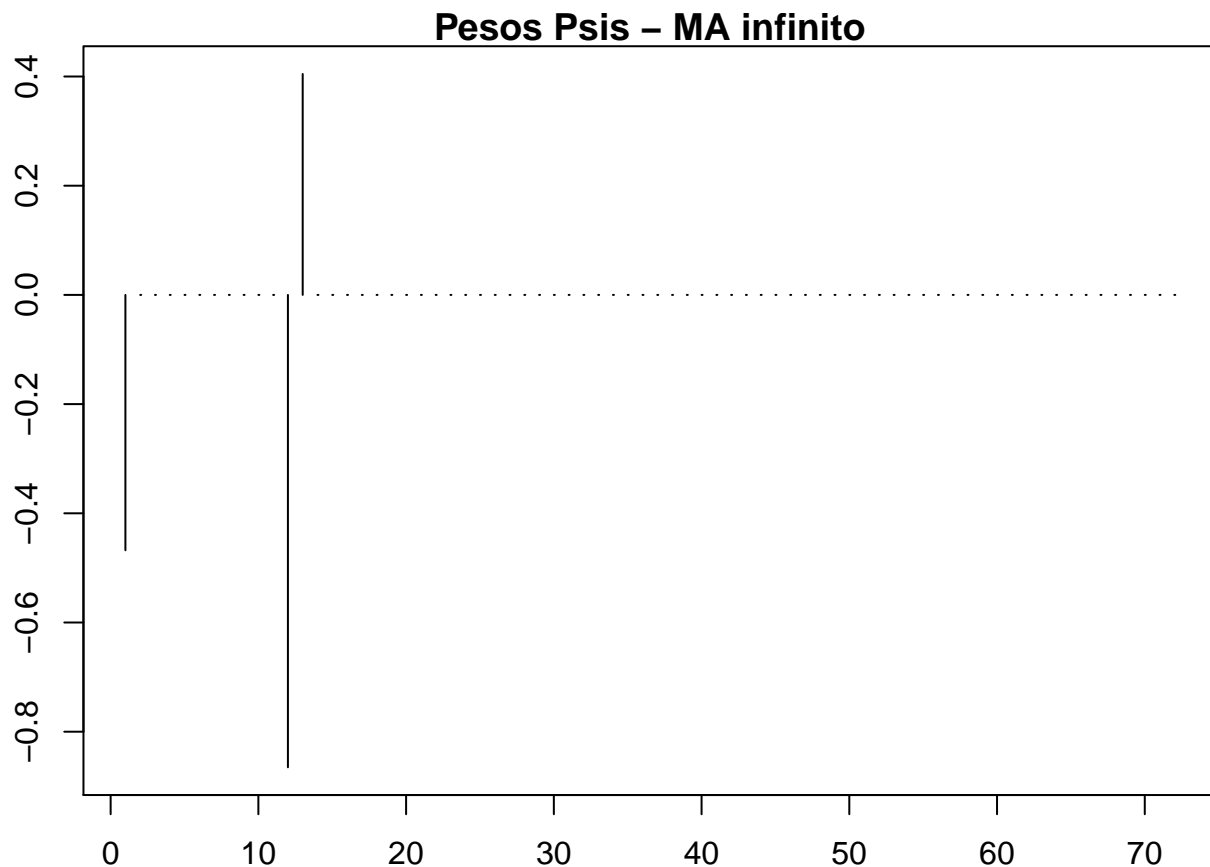


Histogram of resid

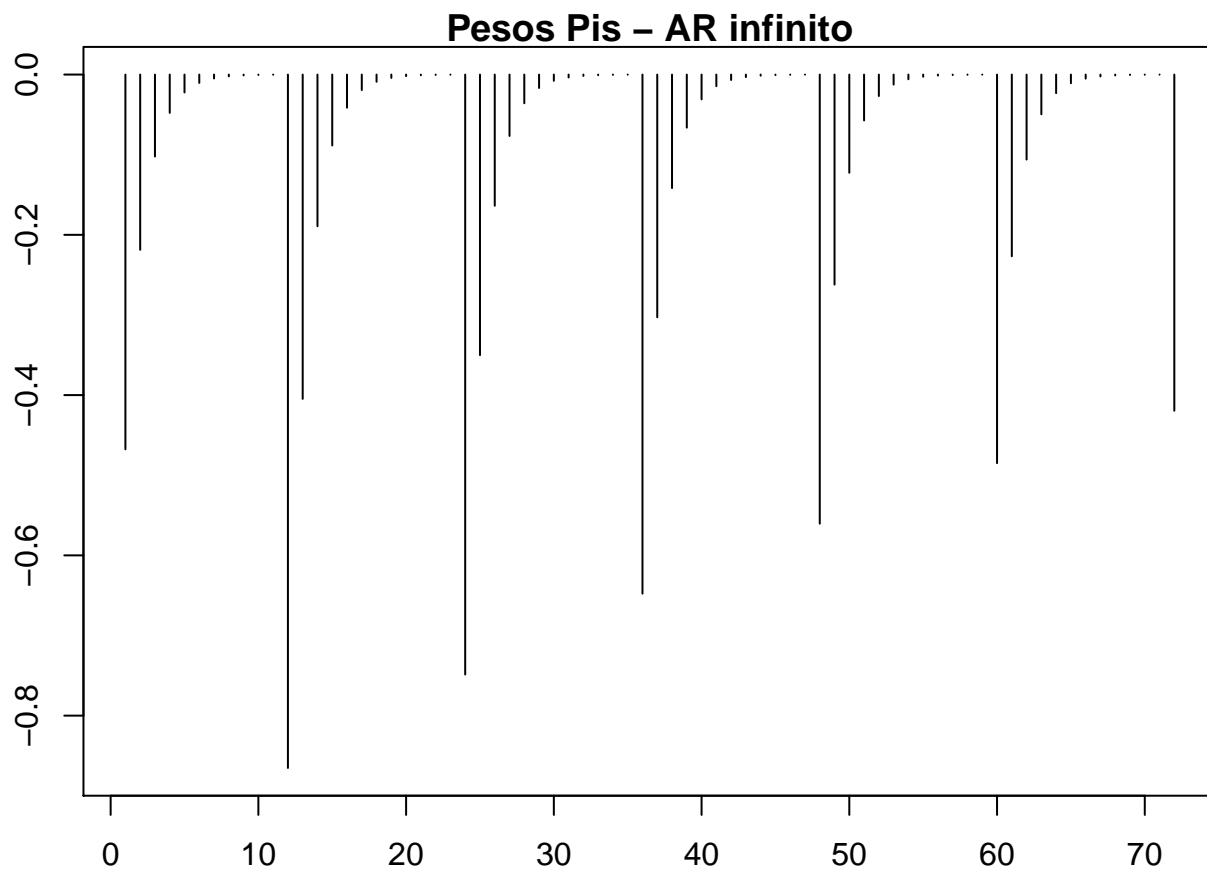




```
## -----
##
## Call:
## arima(x = serie, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##           ma1      sma1  wTradDays
##      -0.4677  -0.8652    0.2530
## s.e.   0.0502   0.0343    0.0438
##
## sigma^2 estimated as 9.889:  log likelihood = -898.33,  aic = 1804.66
##
## Modul of AR Characteristic polynomial Roots:
##
## Modul of MA Characteristic polynomial Roots:  1.012139 1.012139 1.012139 1.012139 1.012139 1.012139
##
## Psi-weights (MA(inf))
##
## -----
##      psi 1      psi 2      psi 3      psi 4      psi 5      psi 6      psi 7
## -0.4676774  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000
##      psi 8      psi 9      psi 10     psi 11     psi 12     psi 13     psi 14
##  0.0000000  0.0000000  0.0000000  0.0000000 -0.8652049  0.4046368  0.0000000
##      psi 15     psi 16     psi 17     psi 18     psi 19     psi 20
##  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000
```

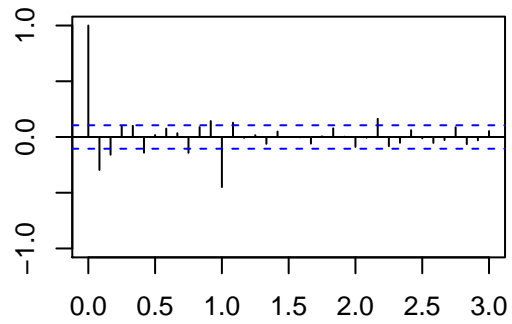


```
##
## Pi-weights (AR(inf))
##
## -----
##      pi 1      pi 2      pi 3      pi 4      pi 5
## -0.4676774094 -0.2187221593 -0.1022914128 -0.0478393830 -0.0223733987
##      pi 6      pi 7      pi 8      pi 9      pi 10
## -0.0104635331 -0.0048935581 -0.0022886066 -0.0010703296 -0.0005005690
##      pi 11     pi 12     pi 13     pi 14     pi 15
## -0.0002341048 -0.8653143873 -0.4046879910 -0.1892634313 -0.0885142312
##      pi 16     pi 17     pi 18     pi 19     pi 20
## -0.0413961064 -0.0193600238 -0.0090542458 -0.0042344662 -0.0019803642
```

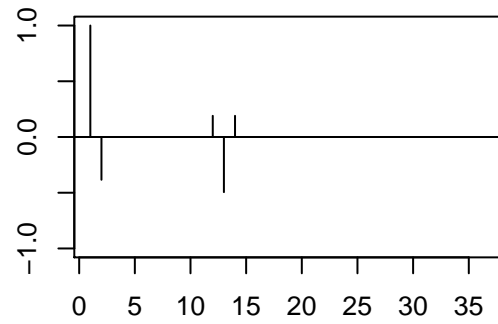


```
##
## Ljung-Box test
##      lag.df  statistic    p.value
## [1,]      1  0.7099452 0.39946200
## [2,]      2  5.0290265 0.08090228
## [3,]      3  5.0653621 0.16707310
## [4,]      4  7.3680593 0.11767004
## [5,]     12 17.7926561 0.12213277
## [6,]     24 29.5411654 0.20042721
## [7,]     36 39.1538578 0.33016082
## [8,]     48 50.8110983 0.36344213
```

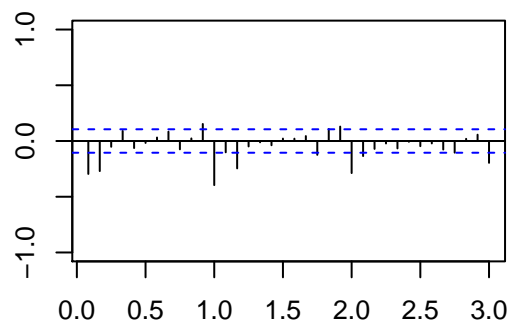
Sample ACF



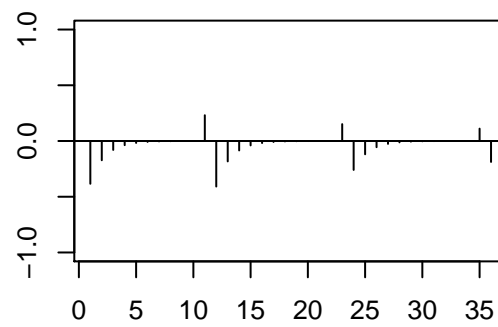
ACF Teoric



Sample PACF

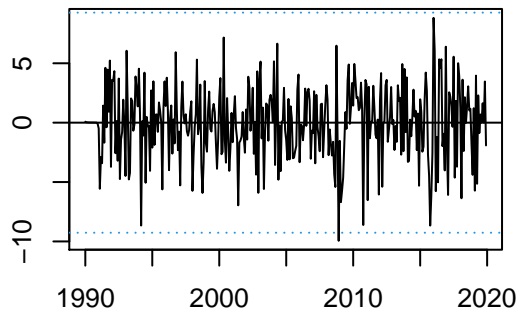


PACF Teoric

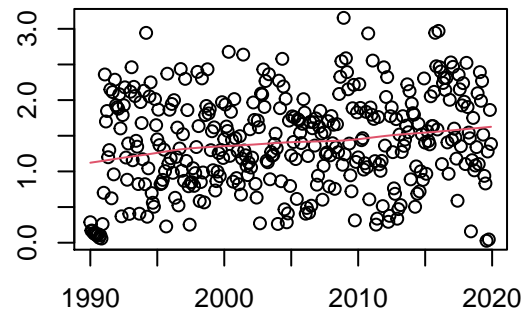


```
dades=d1d12serieTD  
model=moder3  
validation(model,dades)
```

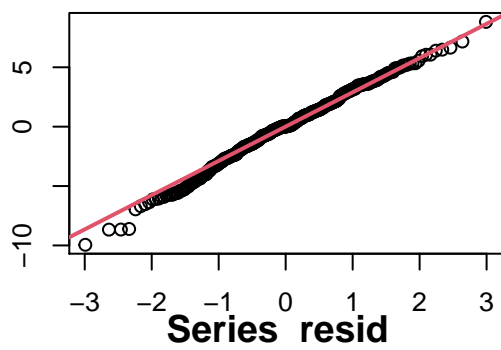
Residuals



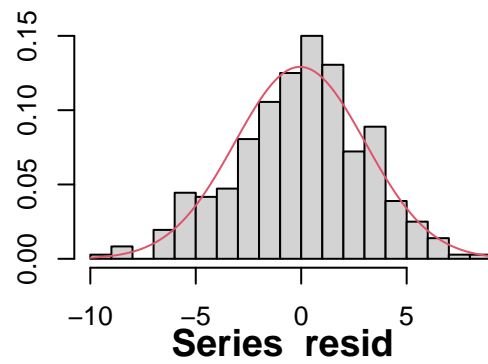
Square Root of Absolute residuals



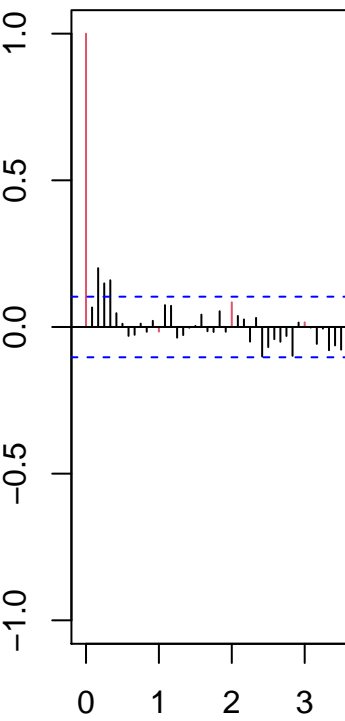
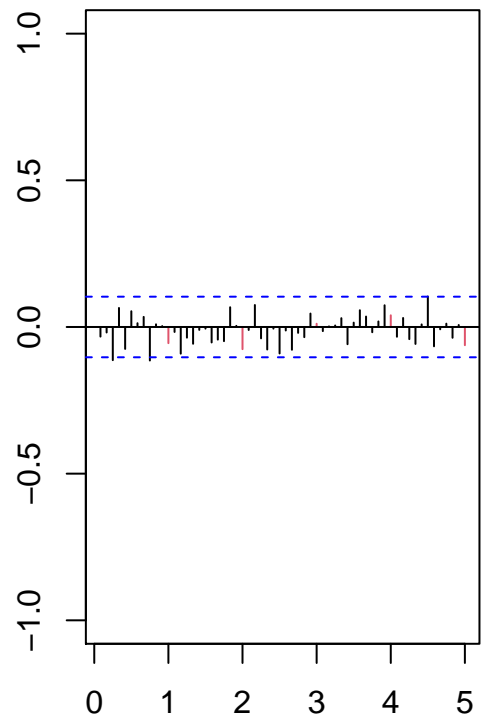
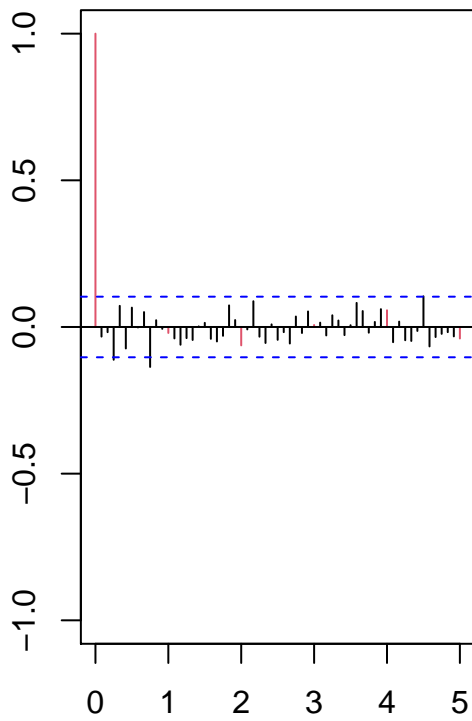
Normal Q-Q Plot

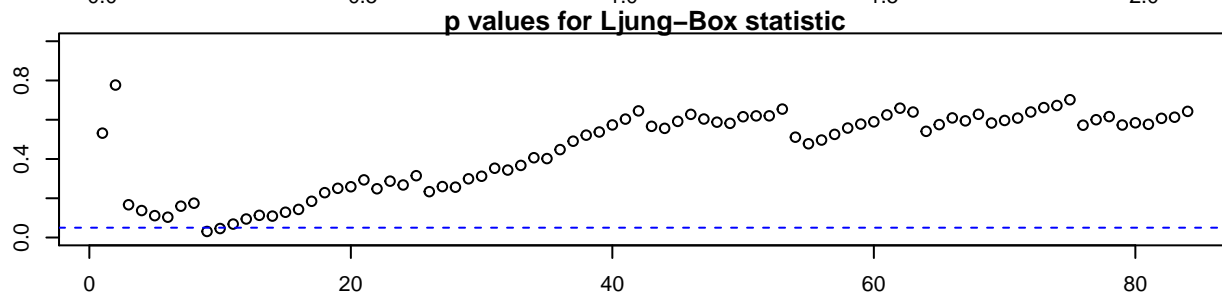
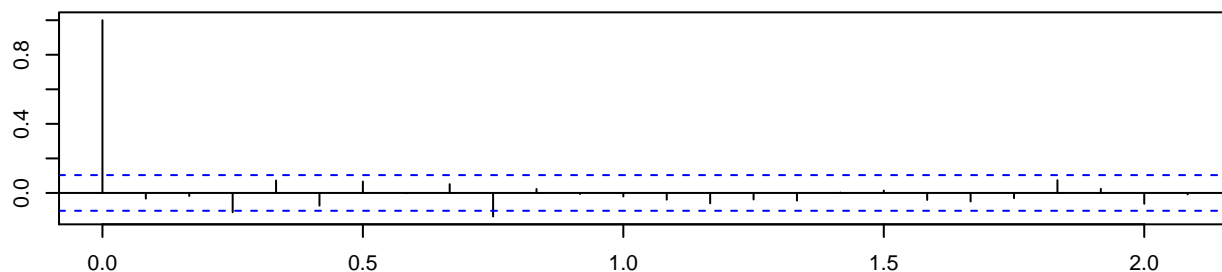
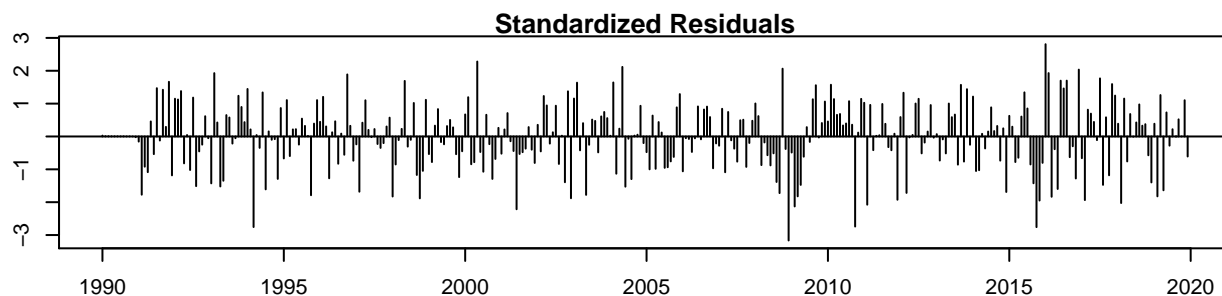


Histogram of resid



Series resid

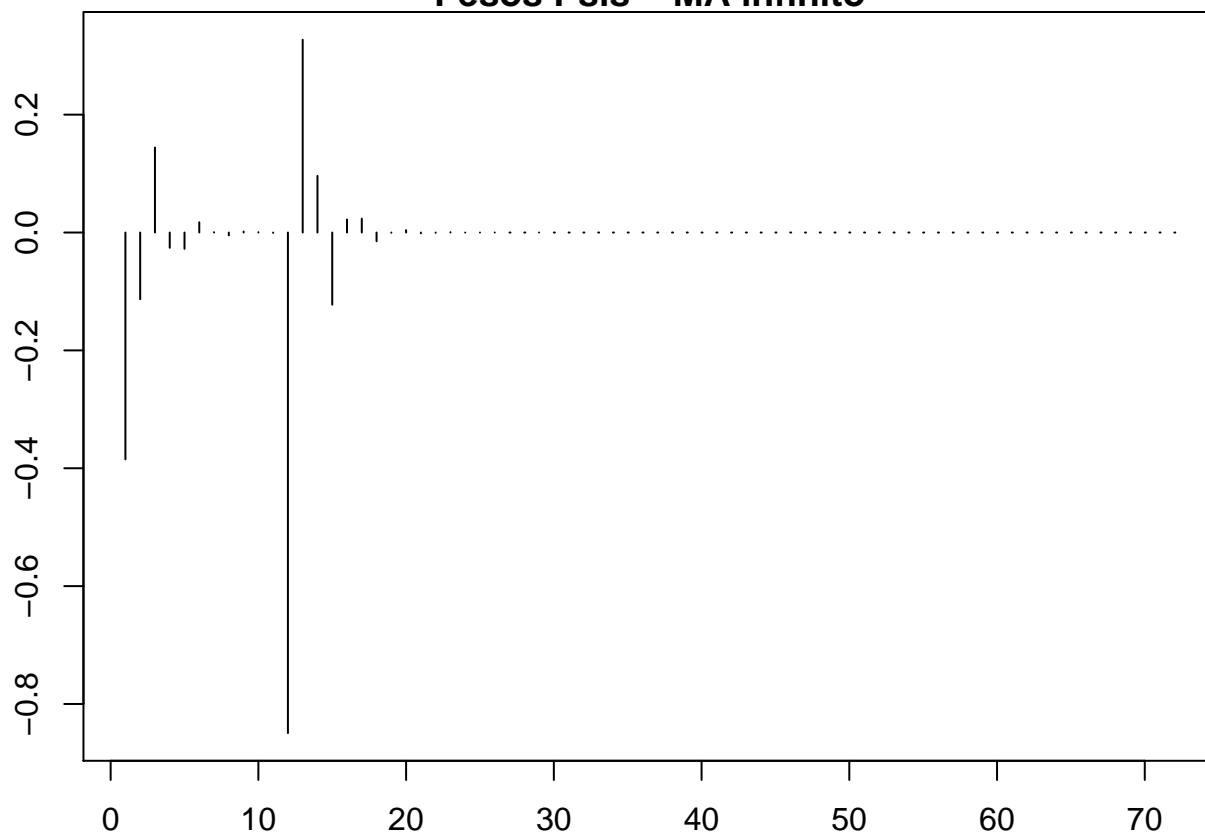




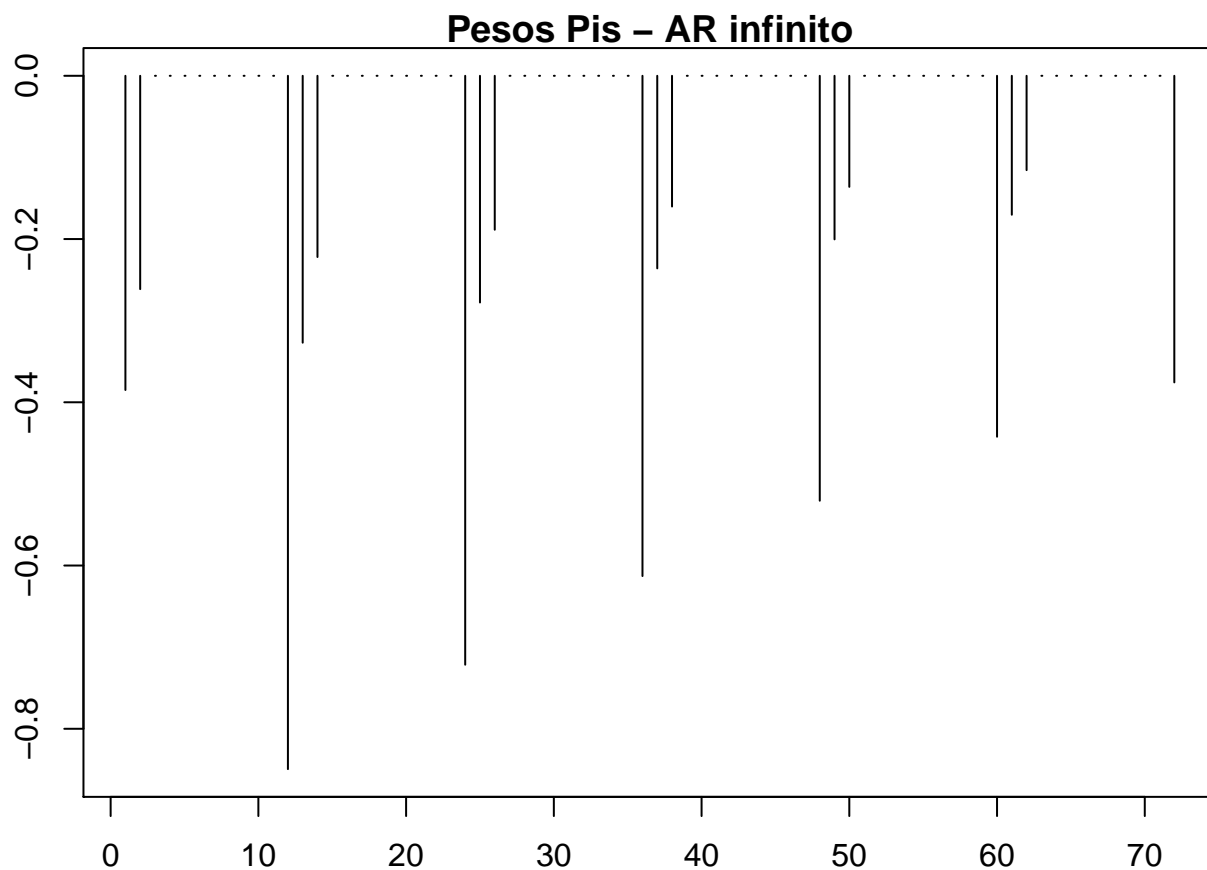
```
##
## -----
##
## Call:
## arima(x = serie, order = c(2, 1, 0), seasonal = list(order = c(0, 1, 1), period = 12),
##       xreg = data.frame(wTradDays))
##
## Coefficients:
##          ar1      ar2      sma1  wTradDays
##      -0.3850  -0.2614  -0.8494    0.2556
## s.e.   0.0518   0.0521   0.0344   0.0445
##
## sigma^2 estimated as 9.872:  log likelihood = -897.43,  aic = 1804.86
##
## Modul of AR Characteristic polynomial Roots:  1.955997 1.955997
##
## Modul of MA Characteristic polynomial Roots:  1.013692 1.013692 1.013692 1.013692 1.013692 1.013692
##
## Psi-weights (MA(inf))
##
## -----
##          psi 1      psi 2      psi 3      psi 4      psi 5
## -0.3849552064 -0.1131843451  0.1441885146 -0.0259225775 -0.0277082211
##          psi 6      psi 7      psi 8      psi 9      psi 10
##  0.0174419339  0.0005278690 -0.0047620889  0.0016952192  0.0005921068
##          psi 11     psi 12     psi 13     psi 14     psi 15
```

```
## -0.0006710223 -0.8493274127 0.3271283978 0.0960630503 -0.1224831093
##      psi 16      psi 17      psi 18      psi 19      psi 20
## 0.0220420447 0.0235288052 -0.0148187723 -0.0004452745 0.0040446652
```

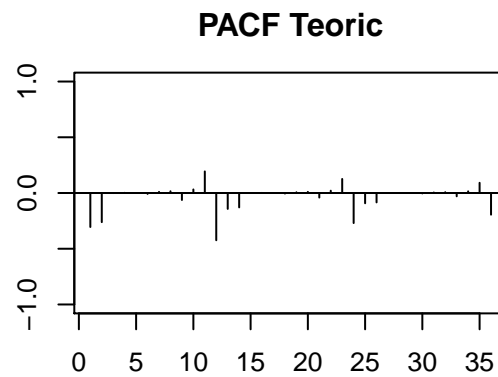
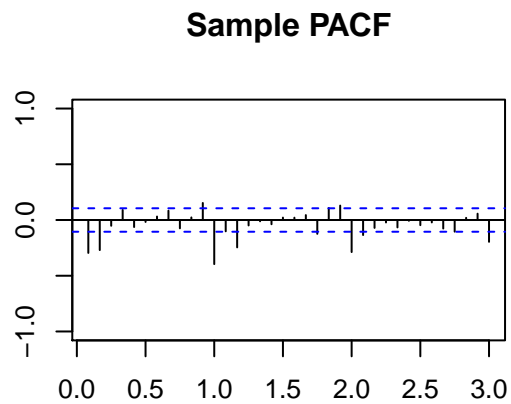
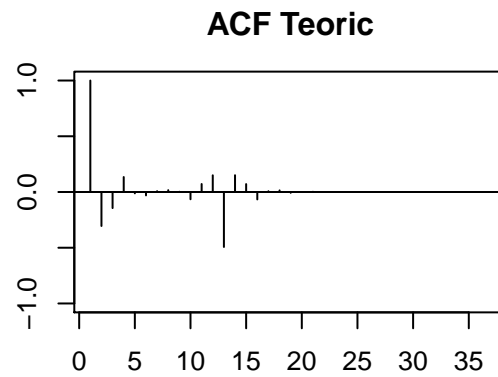
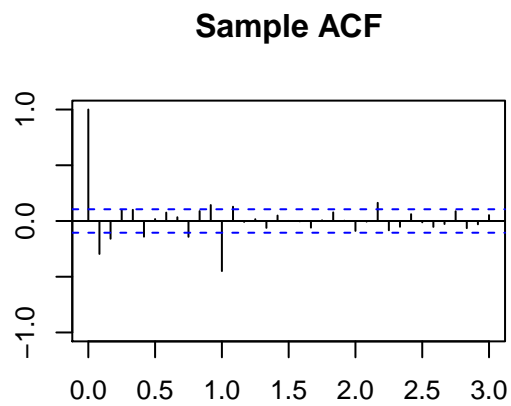
Pesos Psis – MA infinito



```
##
## Pi-weights (AR(inf))
##
## -----
##      pi 1      pi 2      pi 3      pi 4      pi 5      pi 6      pi 7
## -0.3849552 -0.2613749 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
##      pi 8      pi 9      pi 10     pi 11     pi 12     pi 13     pi 14
## 0.0000000 0.0000000 0.0000000 0.0000000 -0.8494310 -0.3269929 -0.2220199
##      pi 15     pi 16     pi 17     pi 18     pi 19     pi 20
## 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
```

```
##
## Ljung-Box test
##      lag.df  statistic    p.value
## [1,]      1  0.3911560 0.53169239
## [2,]      2  0.5053738 0.77671102
## [3,]      3  5.0697845 0.16675792
## [4,]      4  6.9728113 0.13733202
## [5,]     12 18.7677730 0.09428901
## [6,]     24 27.8216525 0.26763427
## [7,]     36 36.4468952 0.44787184
## [8,]     48 45.2163861 0.58758822
```

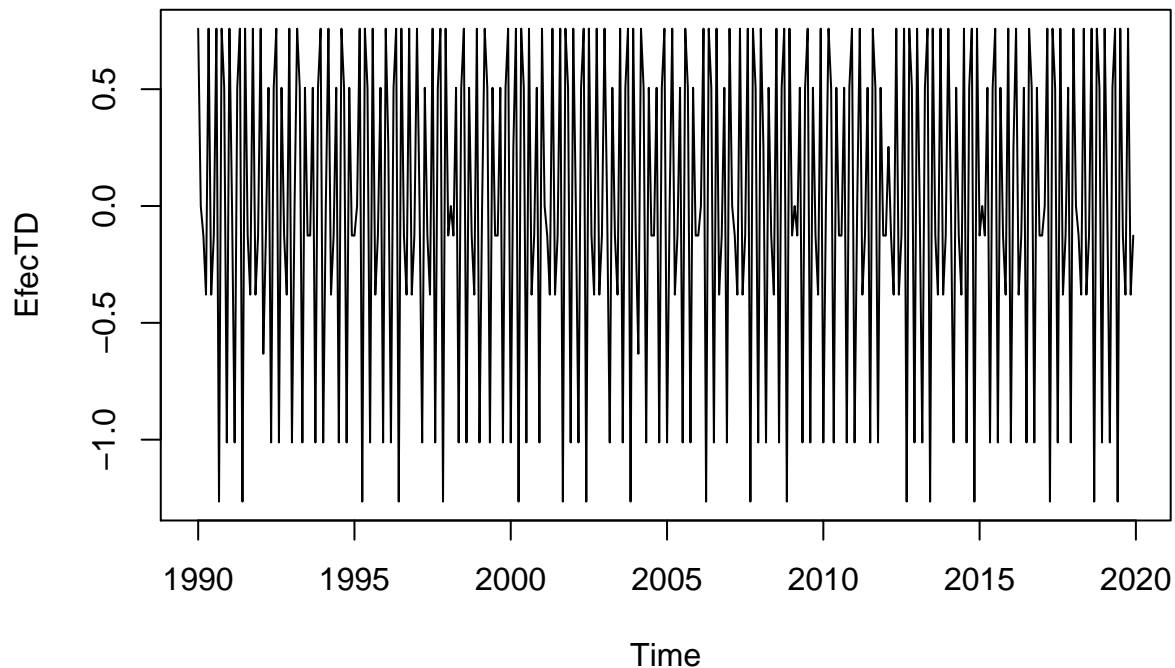


Prefer modec3 (arima(2,1,0)(0,1,1)) because there are only 2 sign Ljung Box

Maybe change model and use modec instead of modec3

Effect of trading days:

```
EfecTD=coef(modec)["wTradDays"]*wTradDays
plot(EfecTD)
```



b) Outlier

(session 5:)

```
# automatic detection of outliers
mod.atip=outdetec(modec,dif=c(1,12),crit=2.8,LS=T)
```

```
# estimated variance after outlier detection and treatment
mod.atip$sigma2
```

```
## [1] 6.530993
```

```
modec$sigma2
```

```
## [1] 9.888563
```

Variance after detection is smaller

```
#Table with detected outliers, their types, magnitud, statistic values and chronology
atipics=mod.atip$atip[order(mod.atip$atip[,1]),]
```

```
meses=c("Jan","Feb","Mar","Apr","May","Jun","Jul","Aug","Sep","Oct","Nov","Dec")
```

```
data.frame(atipics,Fecha=paste(meses[(atipics[,1]-1)%%12+1],start(serie)[1]+((atipics[,1]-1)%/%12)))
```

##	Obs	type_detected	W_coeff	ABS_L_Ratio	Fecha
## 9	14	TC	-6.867291	2.869206	Feb 1991
## 13	40	TC	-6.658615	2.913762	Apr 1993
## 12	51	LS	-6.464667	2.895178	Mar 1994
## 14	82	TC	6.577298	2.911797	Oct 1996
## 6	125	TC	7.149050	2.886624	May 2000
## 4	173	A0	7.442664	3.131095	May 2004
## 5	225	A0	-7.331548	3.122850	Sep 2008
## 1	228	LS	-11.060704	4.296762	Dec 2008
## 10	231	LS	-6.587483	2.877064	Mar 2009

```
## 7 242      LS  6.881041    2.903492 Feb 2010
## 8 266      AO  6.485825    2.849909 Feb 2012
## 3 310      LS -8.688132    3.459232 Oct 2015
## 11 313     AO  6.495974    2.925275 Jan 2016
## 2 314      AO  8.895463    3.582324 Feb 2016
```

#additional column: percentage variation

```
data.frame(atipics,Fecha=paste(meses[(atipics[,1]-1)%%12+1],start(serie)[1]+((atipics[,1]-1)%/%12)),perc
```

```
##   Obs type_detected   W_coeff ABS_L_Ratio   Fecha   perc.Obs
## 9   14              TC  -6.867291    2.869206 Feb 1991 1.041294e-01
## 13  40              TC  -6.658615    2.913762 Apr 1993 1.282921e-01
## 12  51              LS  -6.464667    2.895178 Mar 1994 1.557510e-01
## 14  82              TC   6.577298    2.911797 Oct 1996 7.185952e+04
## 6  125              TC   7.149050    2.886624 May 2000 1.272897e+05
## 4  173              AO   7.442664    3.131095 May 2004 1.707293e+05
## 5  225              AO  -7.331548    3.122850 Sep 2008 6.545593e-02
## 1  228              LS -11.060704    4.296762 Dec 2008 1.571801e-03
## 10 231              LS  -6.587483    2.877064 Mar 2009 1.377503e-01
## 7  242              LS   6.881041    2.903492 Feb 2010 9.736395e+04
## 8  266              AO   6.485825    2.849909 Feb 2012 6.557794e+04
## 3  310              LS  -8.688132    3.459232 Oct 2015 1.685746e-02
## 11 313              AO   6.495974    2.925275 Jan 2016 6.624693e+04
## 2  314              AO   8.895463    3.582324 Feb 2016 7.298785e+05
```

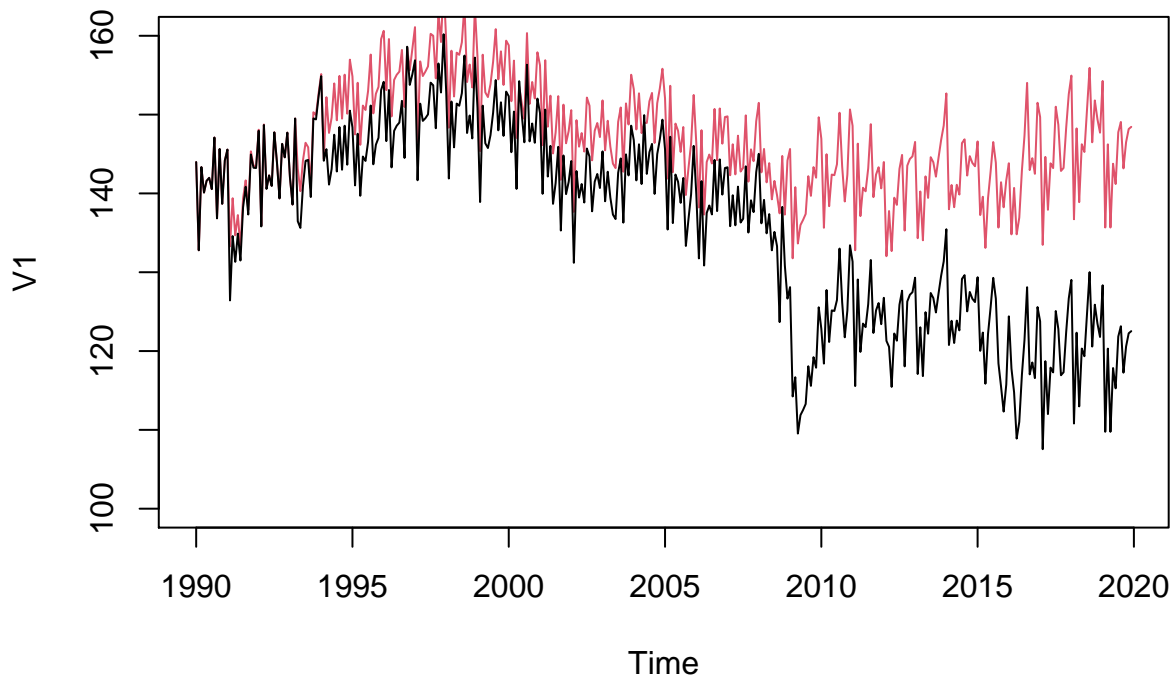
Linearize serie, (without outliers)

```
serie.lin=lineal(serie,mod.atip$atip)
```

#serieEC.lin=serie.lin-EfecTD

```
plot(serie.lin,col=2, ylim = c(100,160))
```

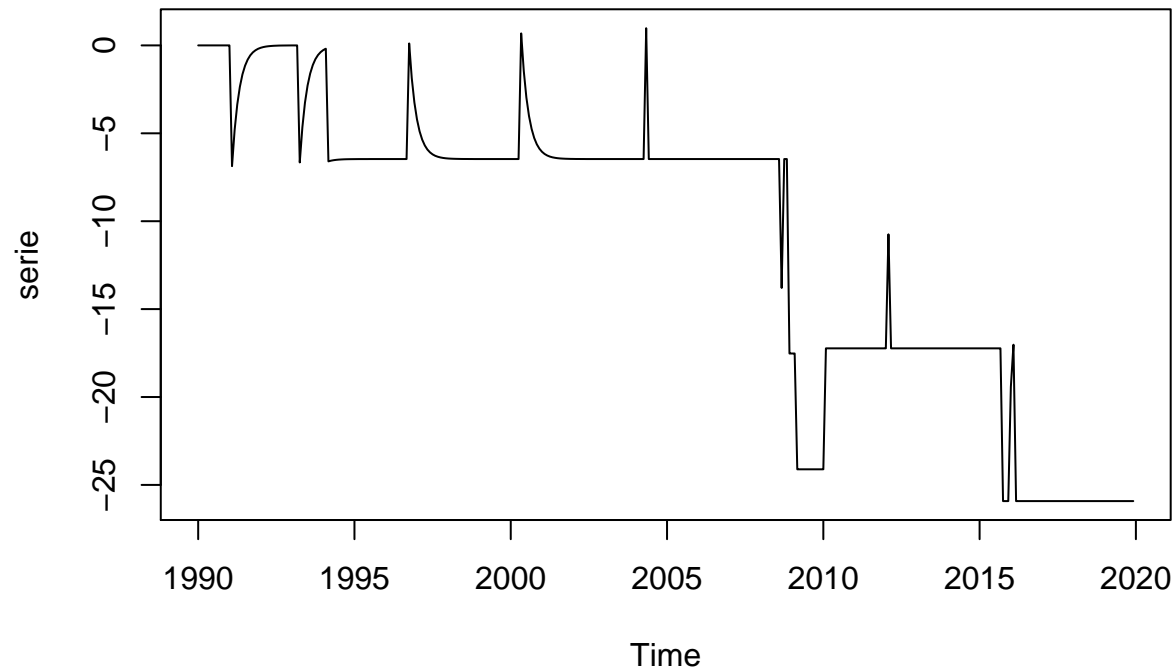
```
lines(serie)
```



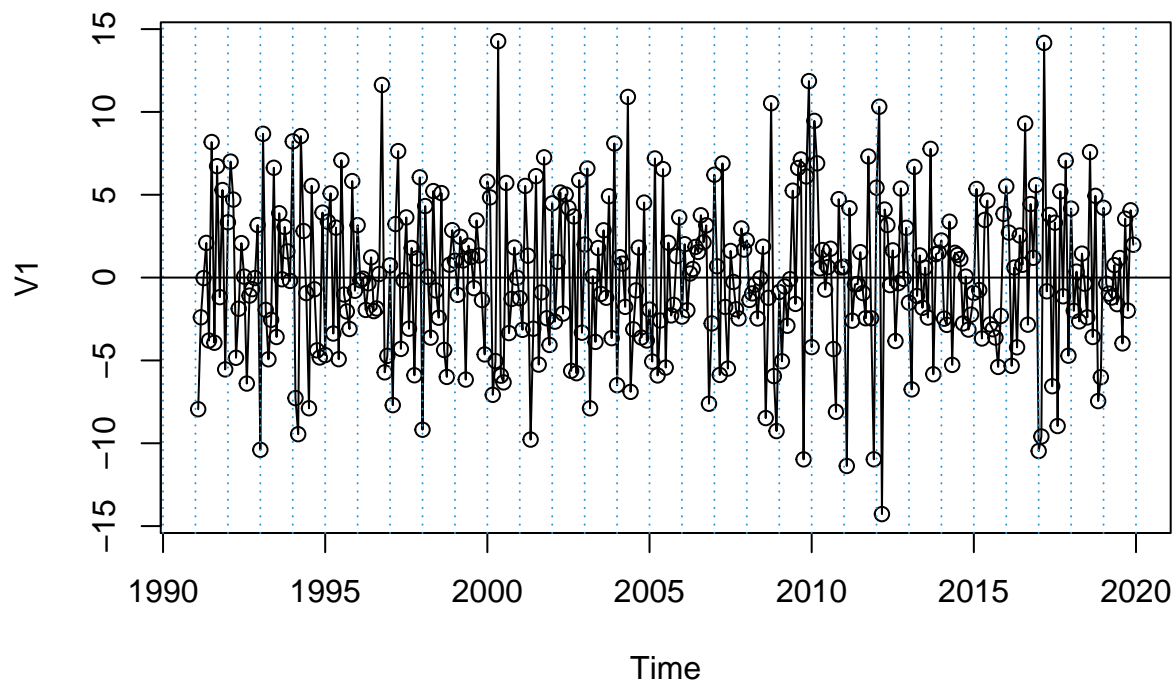
outlier profile

Plot of

```
plot(series-serie.lin)
```



```
d12serie.lin=diff(series.lin,12)
d1d12serie.lin=diff(d12serie.lin)
plot(d1d12serie, col=1, type="o")
abline(v=1990:2020, lty=3, col=4)
abline(h=0)
```

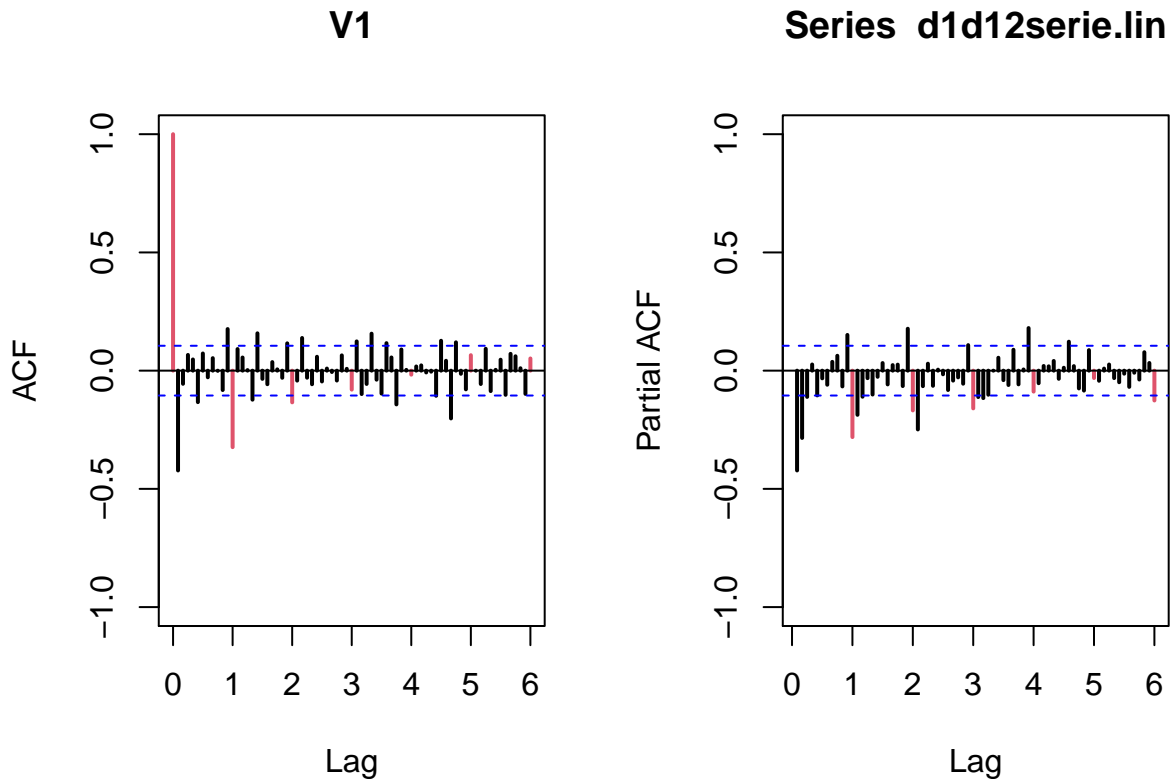


tify model:

```
par(mfrow=c(1,2))
acf(d1d12serie.lin, ylim=c(-1,1), lag.max=72, col=c(2, rep(1,11)), lwd=2)
```

Iden-

```
pacf(d1d12serie.lin,ylim=c(-1,1),lag.max=72,col=c(rep(1,11),2),lwd=2)
```



sonal: MA(1) MA(2) (AR(3)) Regular: AR(1), AR(2), MA(1)

```
#(mod.lin2=arima(serie.lin,order=c(2,1,0),seasonal=list(order=c(0,1,2),period=12),xreg=data.frame(wTradDays))
#ma 2 not sign
```

```
(mod.lin2=arima(serie.lin,order=c(0,1,1),seasonal=list(order=c(0,1,1),period=12),xreg=data.frame(wTradDays))
```

```
##
```

```
## Call:
```

```
## arima(x = serie.lin, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1),
##      period = 12), xreg = data.frame(wTradDays))
```

```
##
```

```
## Coefficients:
```

```
##      ma1      sma1  wTradDays
```

```
##      -0.5794 -0.8232    0.2667
```

```
## s.e.   0.0474   0.0366    0.0377
```

```
##
```

```
## sigma^2 estimated as 6.641:  log likelihood = -827.85,  aic = 1663.69
```

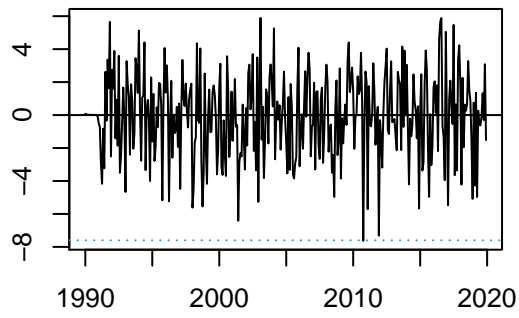
Looks okay:

```
dades=d1d12serie.lin
```

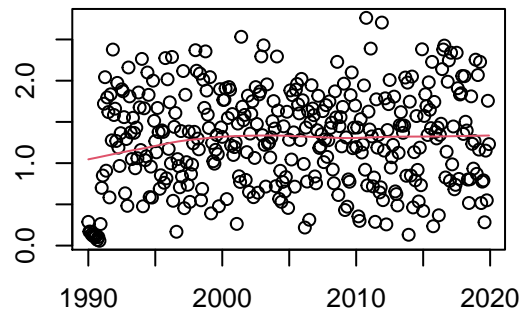
```
model=mod.lin2
```

```
validation(model,dades)
```

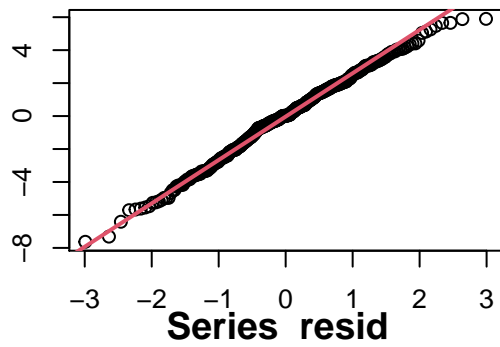
Residuals



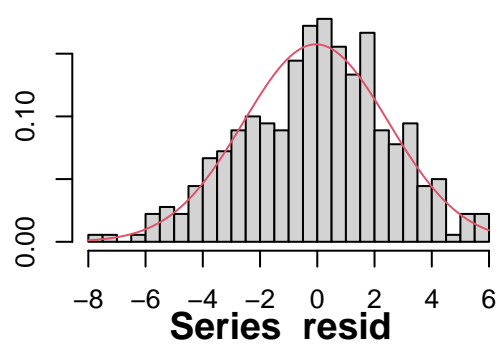
Square Root of Absolute residuals



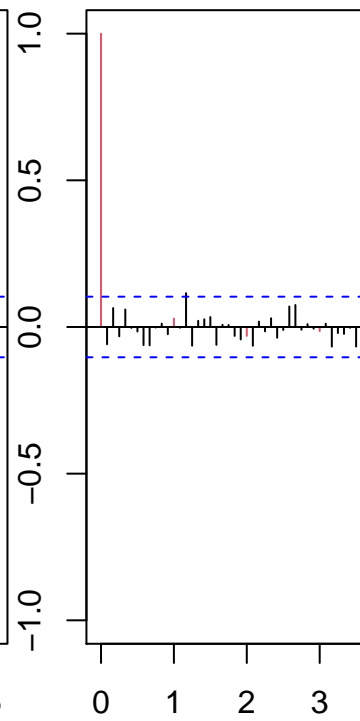
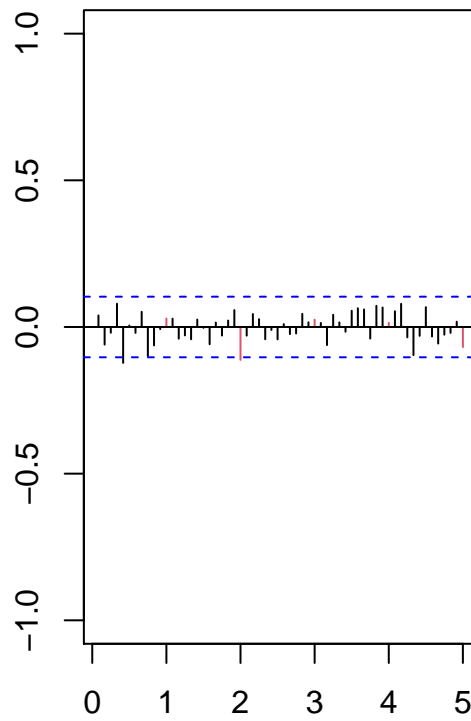
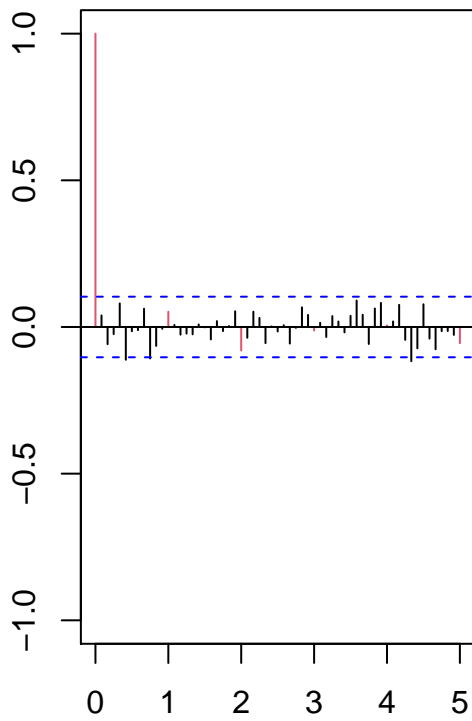
Normal Q-Q Plot

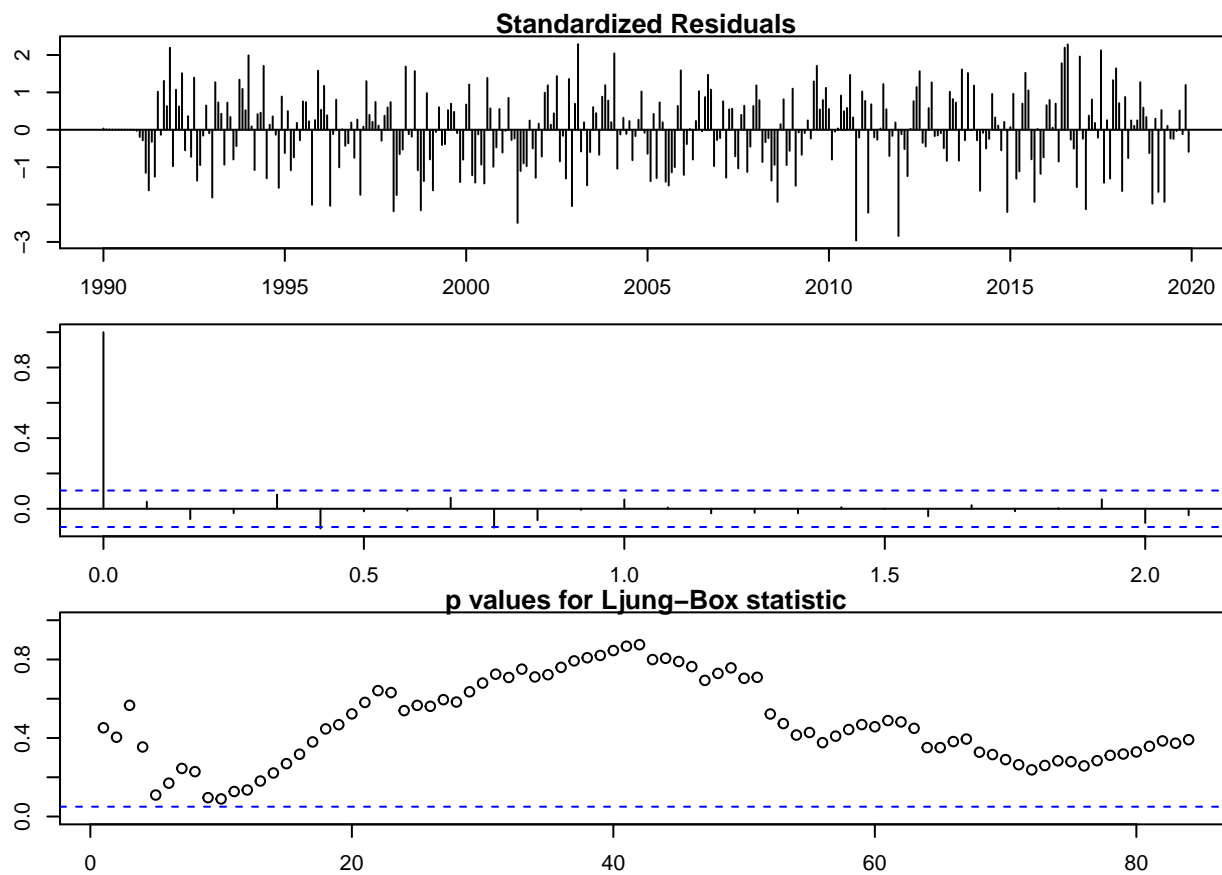


Histogram of resid



Series resid





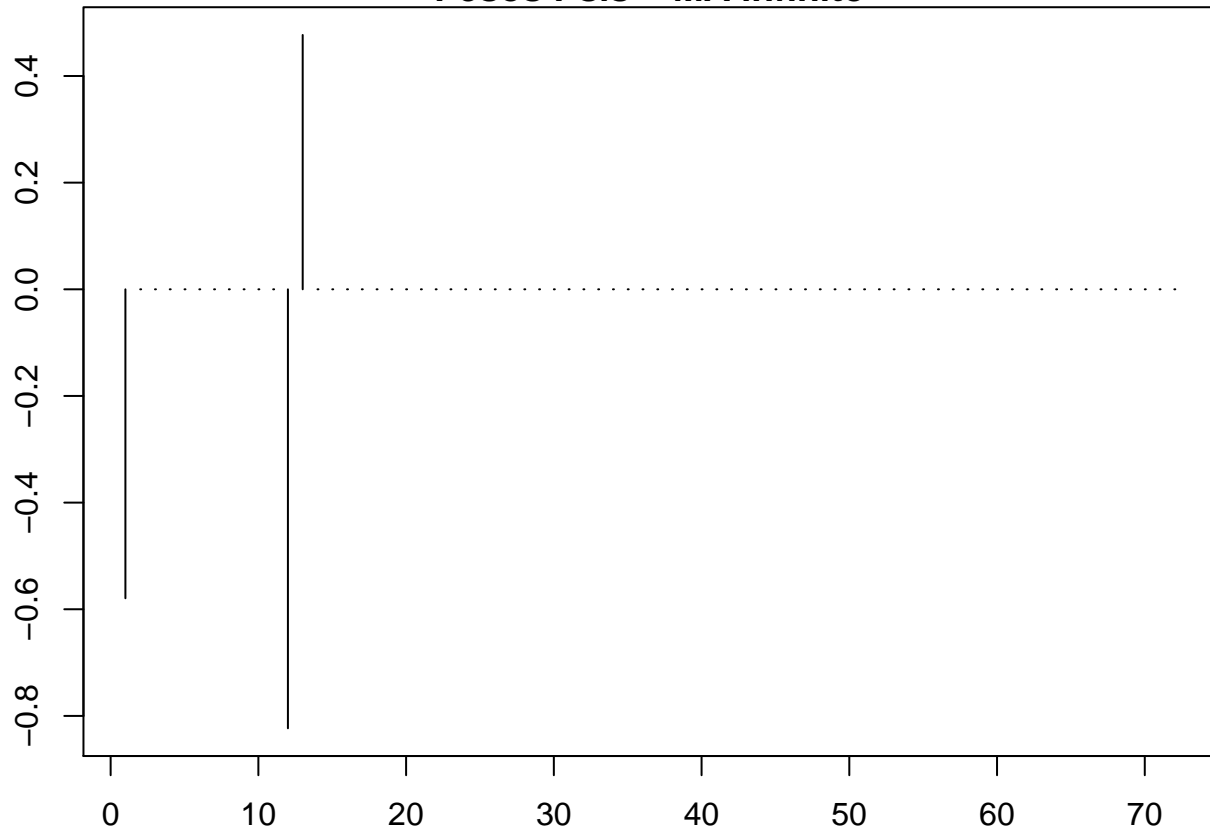
```
##
## -----
##
## Call:
## arima(x = serie.lin, order = c(0, 1, 1), seasonal = list(order = c(0, 1, 1),
##   period = 12), xreg = data.frame(wTradDays))
##
## Coefficients:
##          ma1      sma1  wTradDays
##      -0.5794  -0.8232    0.2667
## s.e.   0.0474   0.0366   0.0377
##
## sigma^2 estimated as 6.641:  log likelihood = -827.85,  aic = 1663.69
##
## Modul of AR Characteristic polynomial Roots:
##
## Modul of MA Characteristic polynomial Roots:  1.016343 1.016343 1.016343 1.016343 1.016343 1.016343
##
## Psi-weights (MA(inf))
##
## -----
##      psi 1      psi 2      psi 3      psi 4      psi 5      psi 6      psi 7
## -0.5793860  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000  0.0000000
##      psi 8      psi 9      psi 10     psi 11     psi 12     psi 13     psi 14
##  0.0000000  0.0000000  0.0000000  0.0000000 -0.8232220  0.4769633  0.0000000
##      psi 15     psi 16     psi 17     psi 18     psi 19     psi 20

```

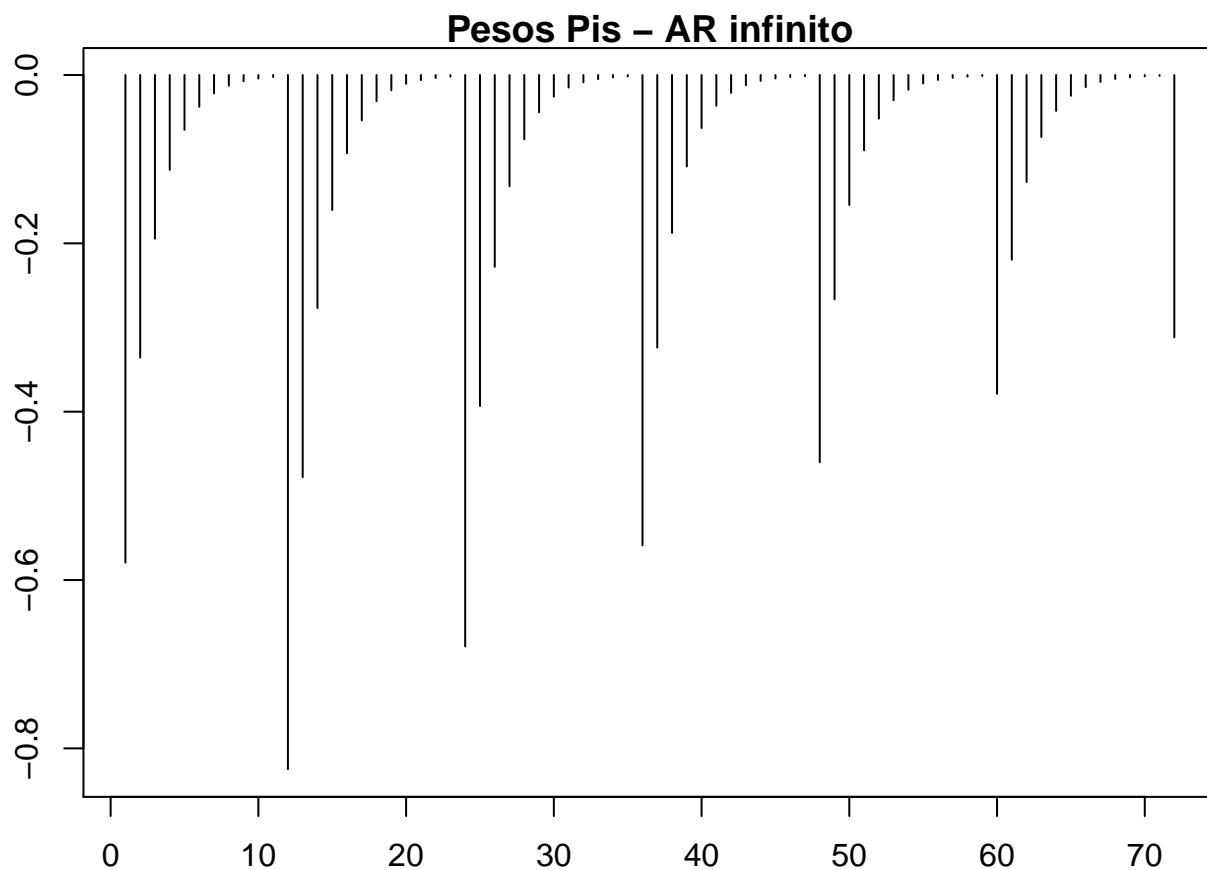


```
## 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
```

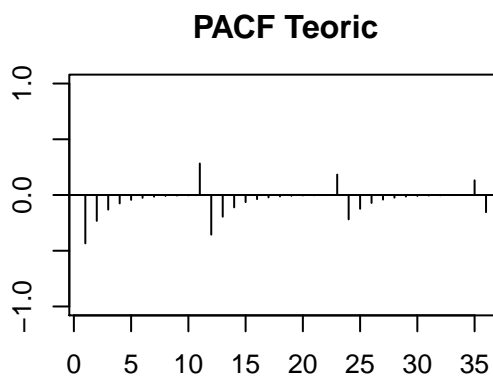
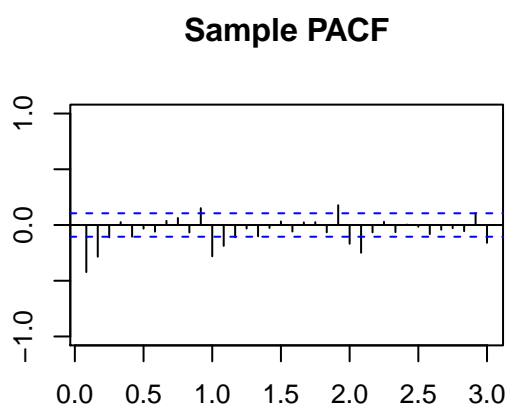
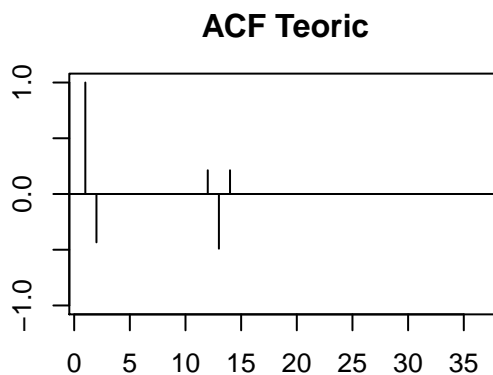
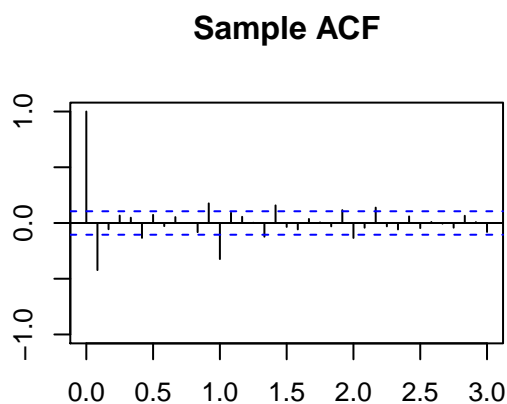
Pesos Psis – MA infinito



```
##
## Pi-weights (AR(inf))
##
## -----
##      pi 1      pi 2      pi 3      pi 4      pi 5      pi 6
## -0.579386033 -0.335688175 -0.194493040 -0.112686551 -0.065289014 -0.037827543
##      pi 7      pi 8      pi 9      pi 10     pi 11     pi 12
## -0.021916750 -0.012698259 -0.007357194 -0.004262655 -0.002469723 -0.824652922
##      pi 13     pi 14     pi 15     pi 16     pi 17     pi 18
## -0.477792385 -0.276826234 -0.160389254 -0.092927293 -0.053840776 -0.031194594
##      pi 19     pi 20
## -0.018073712 -0.010471656
```



```
##
## Ljung-Box test
##      lag.df statistic  p.value
## [1,]      1  0.5656864 0.4519781
## [2,]      2  1.8129189 0.4039519
## [3,]      3  2.0289656 0.5664166
## [4,]      4  4.4028800 0.3542192
## [5,]     12 17.4040374 0.1350204
## [6,]     24 22.6682688 0.5394448
## [7,]     36 29.7382710 0.7598491
## [8,]     48 41.6542403 0.7289835
```

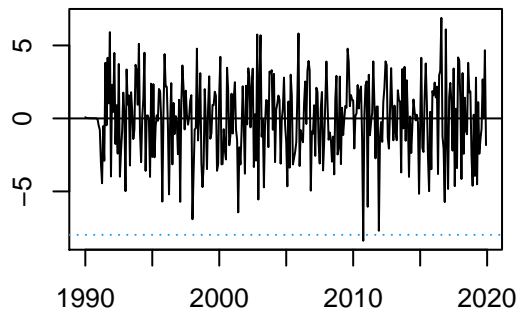


```
(mod.lin=arima(serie.lin,order=c(2,1,0),seasonal=list(order=c(3,1,0),period=12),xreg=data.frame(wTradDays
```

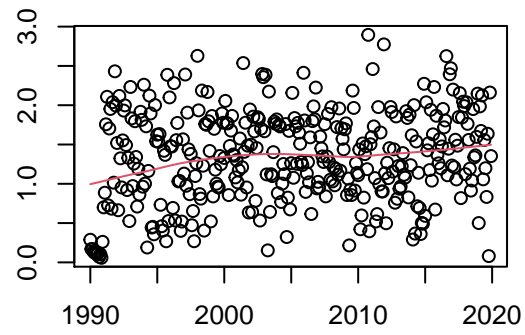
```
##
## Call:
## arima(x = serie.lin, order = c(2, 1, 0), seasonal = list(order = c(3, 1, 0),
##   period = 12), xreg = data.frame(wTradDays))
##
## Coefficients:
##      ar1      ar2      sar1      sar2      sar3 wTradDays
##    -0.4922 -0.2882 -0.6502 -0.5021 -0.2731    0.2706
## s.e.   0.0515   0.0520   0.0539   0.0586   0.0565   0.0371
##
## sigma^2 estimated as 7.312:  log likelihood = -841.8,  aic = 1697.6
```

```
dades=d1d12serie.lin
model=mod.lin
validation(model,dades)
```

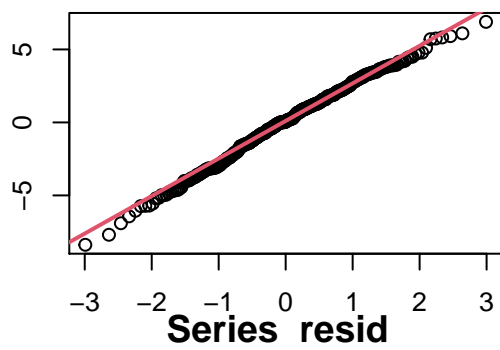
Residuals



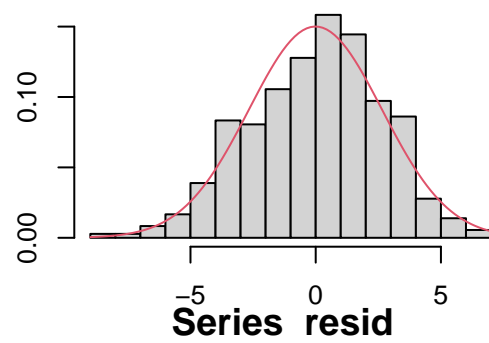
Square Root of Absolute residuals



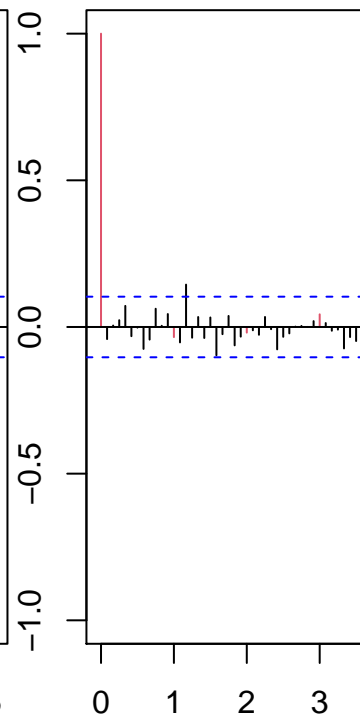
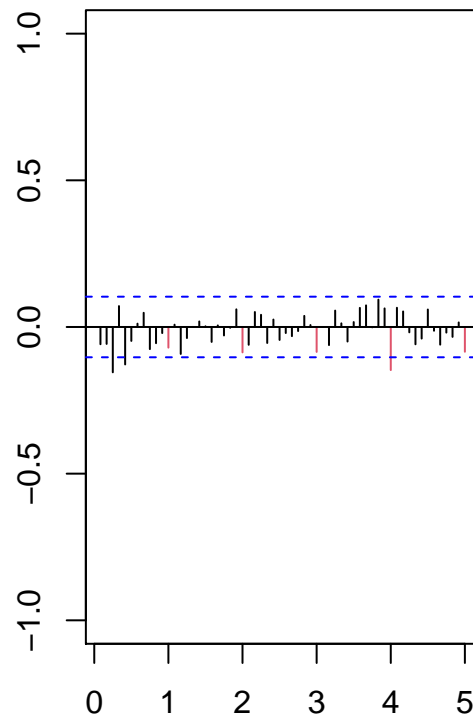
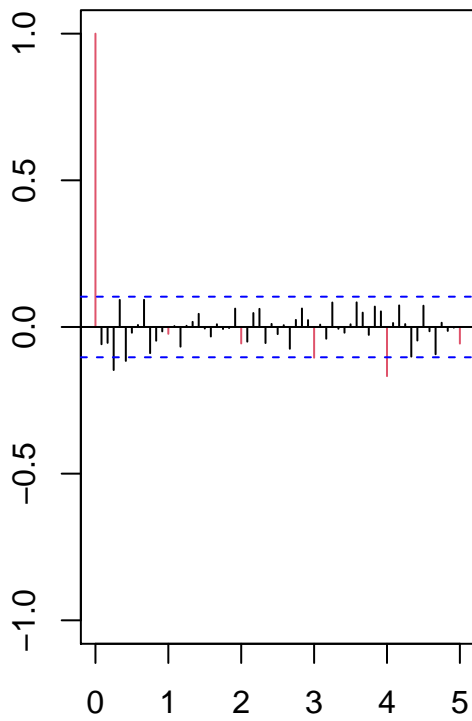
Normal Q-Q Plot

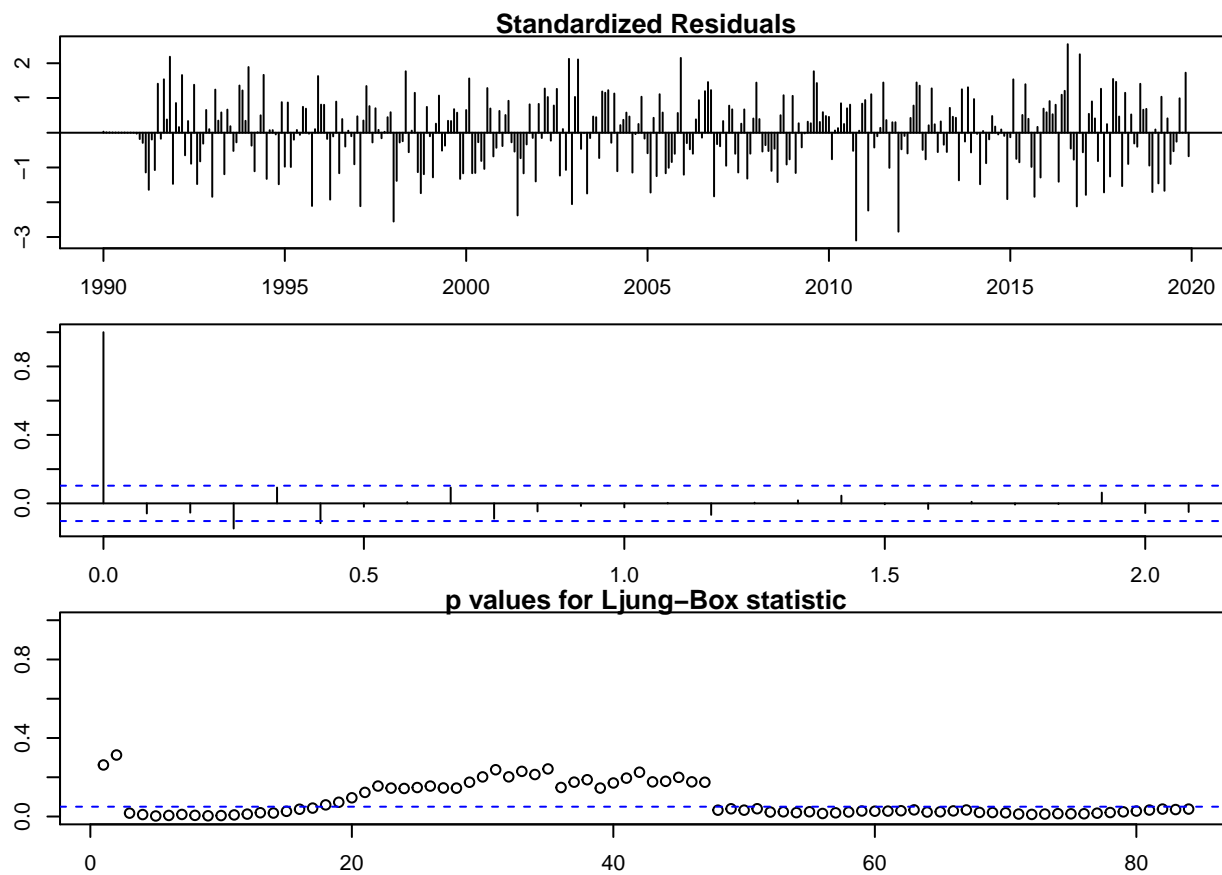


Histogram of resid



Series resid

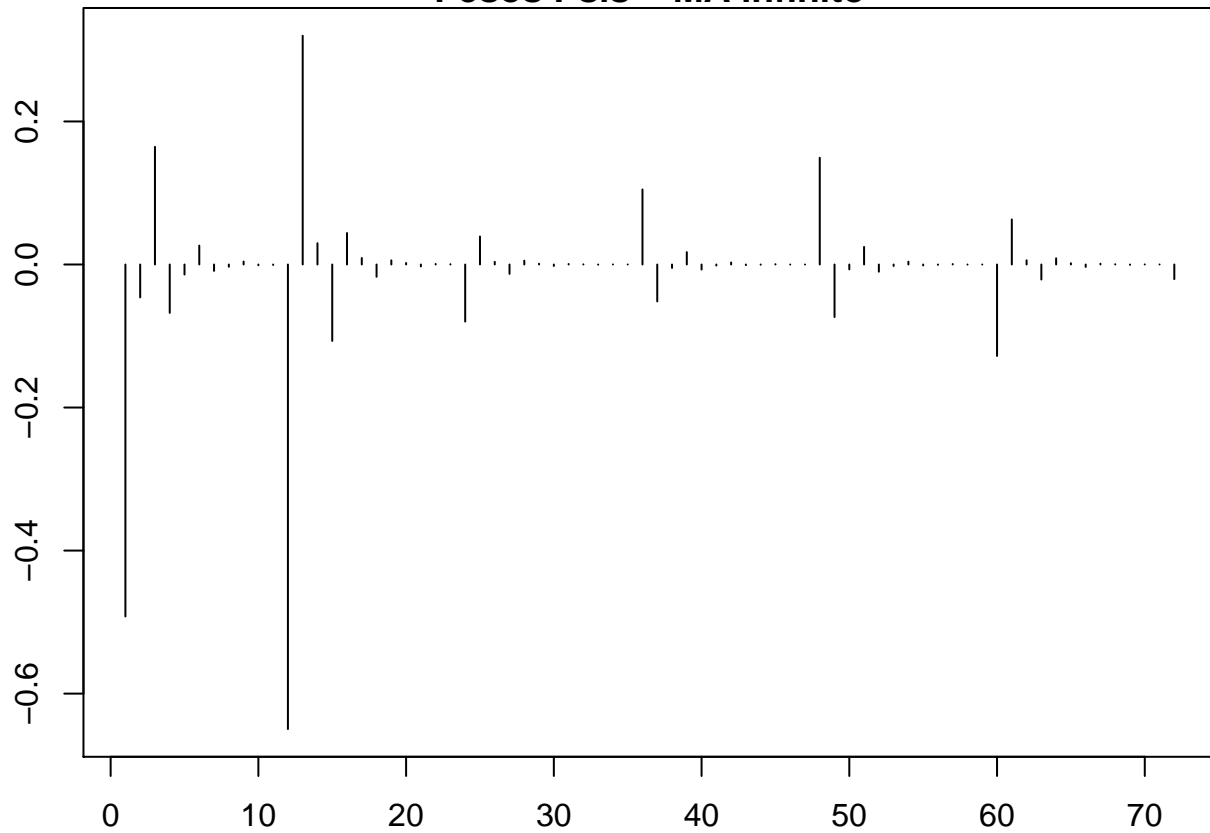




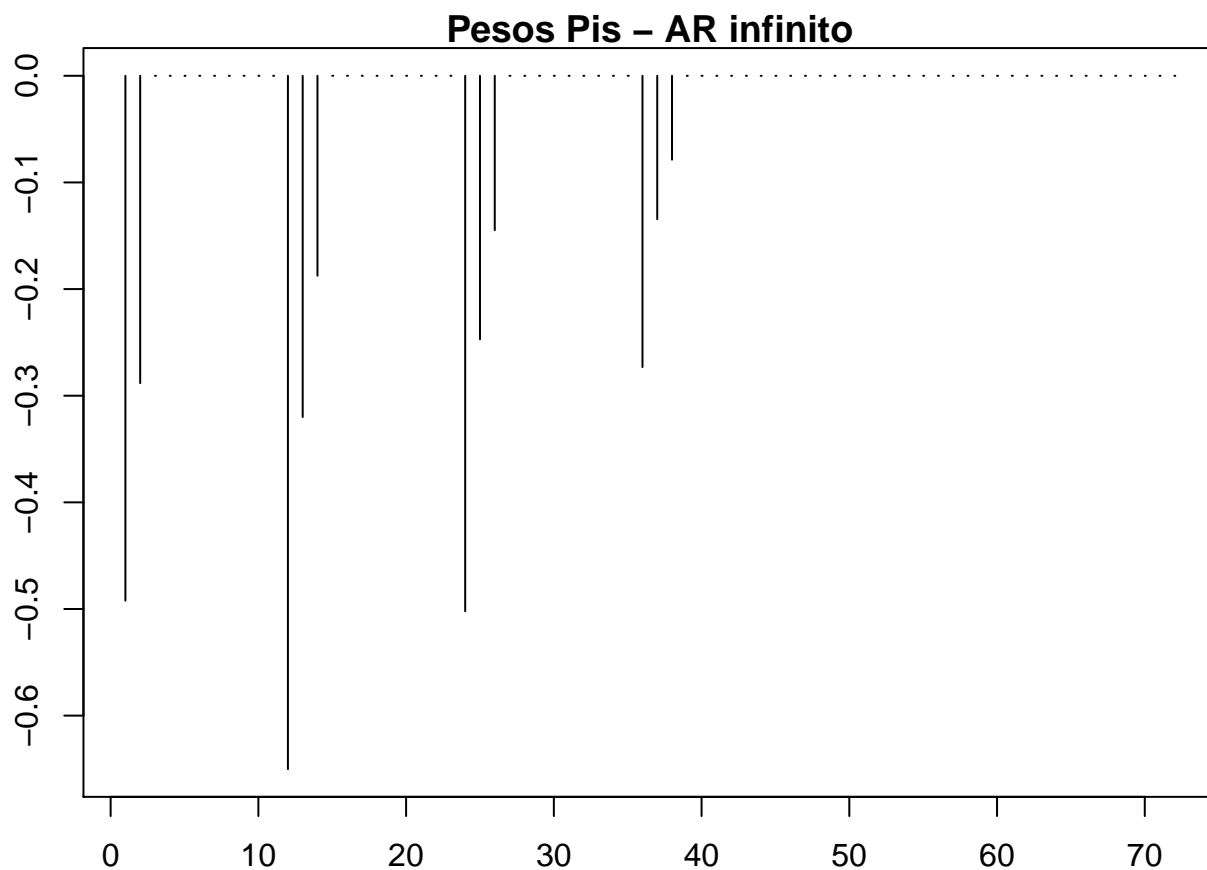
```
##
## -----
##
## Call:
## arima(x = serie.lin, order = c(2, 1, 0), seasonal = list(order = c(3, 1, 0),
##   period = 12), xreg = data.frame(wTradDays))
##
## Coefficients:
##          ar1      ar2      sar1      sar2      sar3  wTradDays
##      -0.4922 -0.2882 -0.6502 -0.5021 -0.2731      0.2706
## s.e.   0.0515   0.0520   0.0539   0.0586   0.0565   0.0371
##
## sigma^2 estimated as 7.312:  log likelihood = -841.8,  aic = 1697.6
##
## Modul of AR Characteristic polynomial Roots:  1.032404 1.032404 1.032404 1.032404 1.032404 1.032404
##
## Modul of MA Characteristic polynomial Roots:
##
## Psi-weights (MA(inf))
##
## -----
##          psi 1      psi 2      psi 3      psi 4      psi 5
## -0.4921530242 -0.0460053020  0.1644899445 -0.0676945800 -0.0140931833
##          psi 6      psi 7      psi 8      psi 9      psi 10
##  0.0264469279 -0.0089539997 -0.0032157929  0.0041633831 -0.0011221661
##          psi 11      psi 12      psi 13      psi 14      psi 15
```

```
## -0.0006476924 -0.6495356030 0.3198575892 0.0297902075 -0.1068506634
##      psi 16      psi 17      psi 18      psi 19      psi 20
## 0.0440007465 0.0091413872 -0.0171808522 0.0058208786 0.0020871005
```

Pesos Psis – MA infinito

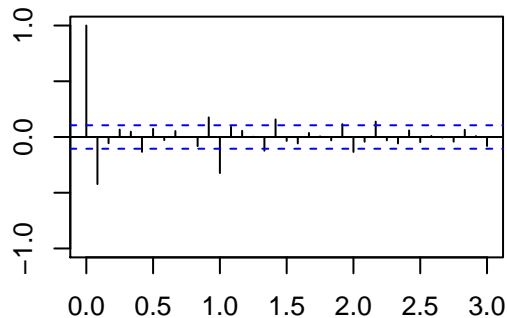


```
##
## Pi-weights (AR(inf))
##
## -----
##      pi 1      pi 2      pi 3      pi 4      pi 5      pi 6      pi 7
## -0.4921530 -0.2882199 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
##      pi 8      pi 9      pi 10     pi 11     pi 12     pi 13     pi 14
## 0.0000000 0.0000000 0.0000000 0.0000000 -0.6501778 -0.3199870 -0.1873942
##      pi 15     pi 16     pi 17     pi 18     pi 19     pi 20
## 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000 0.0000000
```

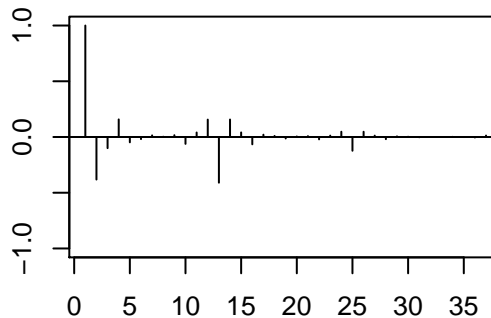


```
##
## Ljung-Box test
##      lag.df statistic      p.value
## [1,]      1  1.255102 0.262580229
## [2,]      2  2.320745 0.313369490
## [3,]      3 10.180004 0.017096401
## [4,]      4 13.298904 0.009903984
## [5,]     12 25.597919 0.012230271
## [6,]     24 31.393390 0.142760822
## [7,]     36 44.855073 0.147849733
## [8,]     48 67.650740 0.032223555
```

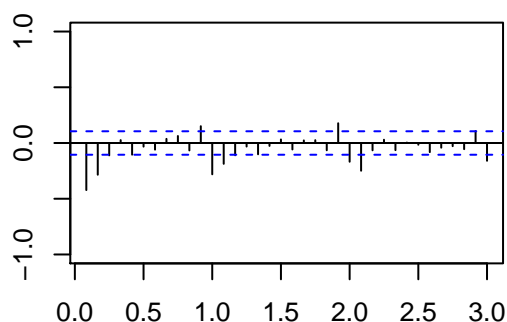
Sample ACF



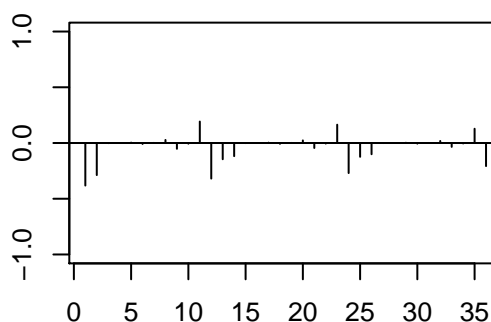
ACF Teoric



Sample PACF



PACF Teoric



This

model does not look goof

Stability

```
ultim=c(2018,12)
pdq=c(0,1,1)
PDQ=c(0,1,1)

serie1=window(serie.lin,end=ultim+c(1,0))
serie2=window(serie.lin,end=ultim)

wTradDays2=window(wTradDays,end=ultim)

(mod=arima(serie1,order=pdq,seasonal=list(order=PDQ,period=12),xreg=data.frame(wTradDays)))

##
## Call:
## arima(x = serie1, order = pdq, seasonal = list(order = PDQ, period = 12), xreg = data.frame(wTradDays)
##
## Coefficients:
##          ma1      sma1  wTradDays
##      -0.5794  -0.8232    0.2667
## s.e.   0.0474   0.0366    0.0377
##
## sigma^2 estimated as 6.641:  log likelihood = -827.85,  aic = 1663.69
(mod2=arima(serie2,order=pdq,seasonal=list(order=PDQ,period=12),xreg=data.frame(wTradDays2)))

##
```



```
## Call:
## arima(x = serie2, order = pdq, seasonal = list(order = PDQ, period = 12), xreg = data.frame(wTradDays2))
##
## Coefficients:
##          ma1      sma1  wTradDays2
##      -0.5798  -0.8323    0.2736
## s.e.    0.0496   0.0383    0.0386
##
## sigma^2 estimated as 6.687:  log likelihood = -800.9,  aic = 1609.8
model is stable
```

Forecast:

Insample prediction: (not in assignment?)

```
pred=predict(mod.lin2,n.ahead=12,newxreg=window(wTradDays,start=c(ultim[1]+1,1)))
predic=pred$pr

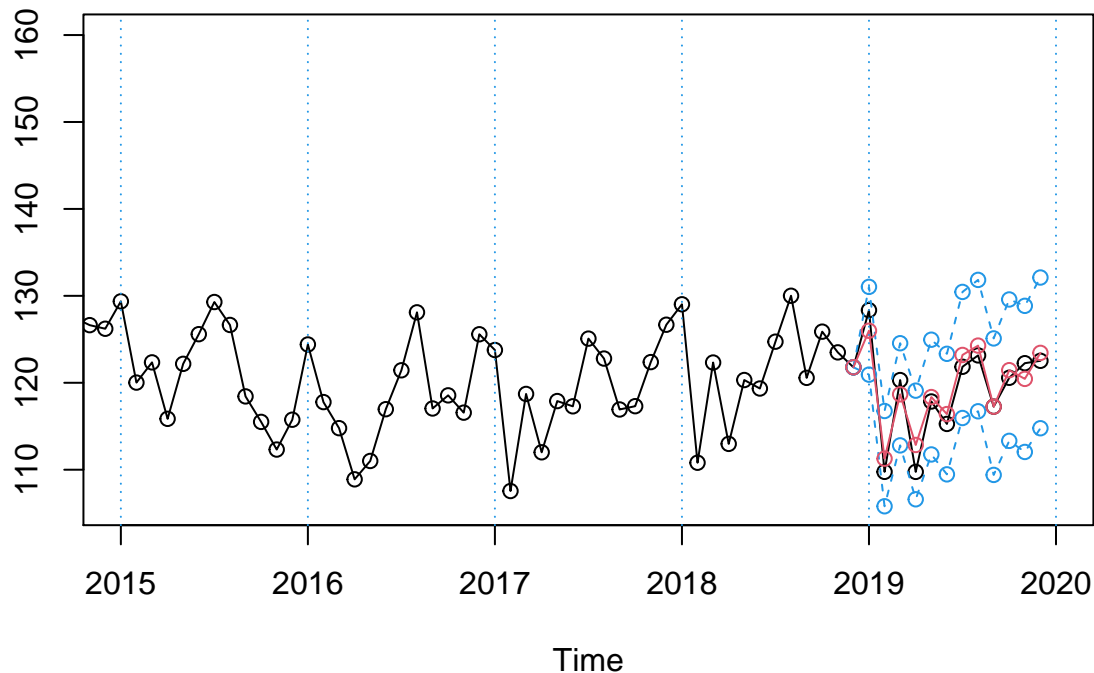
wLS=sum(mod.atip$atip[mod.atip$atip$type_detected=="LS" & mod.atip$atip$Obs<=length(serie)-12,3])
pr<-ts(c(tail(serie2,1),predic)+wLS,start=ultim,freq=12)

se<-ts(c(0,pred$se),start=ultim,freq=12)

#Intervals
tl<-ts(pr-1.96*se,start=ultim,freq=12)
tu<-ts(pr+1.96*se,start=ultim,freq=12)
pr<-ts(pr,start=ultim,freq=12)

ts.plot(serie,tl,tu,pr,lty=c(1,2,2,1),col=c(1,4,4,2),xlim=ultim[1]+c(-3,2),type="o",main="Model ARIMA(2
abline(v=(ultim[1]-3):(ultim[1]+2),lty=3,col=4)
```

Model ARIMA(2,1,0)(3,1,0)12+CE+Atip



2019 is weird, besides that: looks good,

red obs in

```
obs=window(serie, start=ultim)
(mod.EQM3=sqrt(sum(((obs-pr)/obs)^2)/12))
```

```
## [1] 0.01342432
```

```
(mod.EAM3=sum(abs(obs-pr)/obs)/12)
```

```
## [1] 0.01150628
```

```
data3=c(ultim[1]+2, 1, 12)
```

```
wTradDays3=Wtrad(data3)
```

```
pred=predict(mod.lin2, n.ahead=12, newxreg=data.frame(wTradDays3)) ##wEast3))
```

```
predic=pred$pr
```

```
wLS=sum(mod.atip$atip[mod.atip$atip$type_detected=="LS", 3])
```

```
pr<-ts(c(serie[length(serie)], predic+wLS), start=ultim+c(1,0), freq=12)
```

```
se<-ts(c(0, pred$se), start=ultim+c(1,0), freq=12)
```

```
#Intervals
```

```
tl3<-ts(pr-1.96*se, start=ultim+c(1,0), freq=12)
```

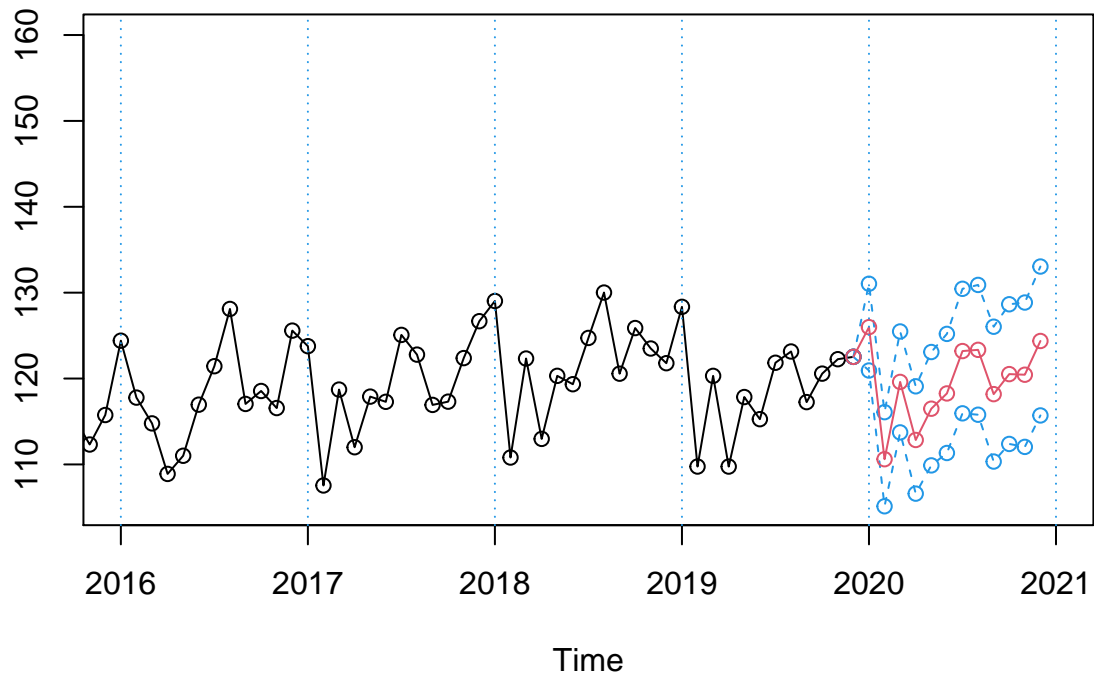
```
tu3<-ts(pr+1.96*se, start=ultim+c(1,0), freq=12)
```

```
pr3<-ts(pr, start=ultim+c(1,0), freq=12)
```

```
ts.plot(serie, tl3, tu3, pr3, lty=c(1,2,2,1), col=c(1,4,4,2), xlim=ultim[1]+c(-2,3), type="o", main="Model ARIMA(2,1,0)(3,1,0)12+CE+Atip")
```

```
abline(v=(ultim[1]-2):(ultim[1]+3), lty=3, col=4)
```

Model ARIMA(0,1,1)(0,1,1)12+CE+Atip

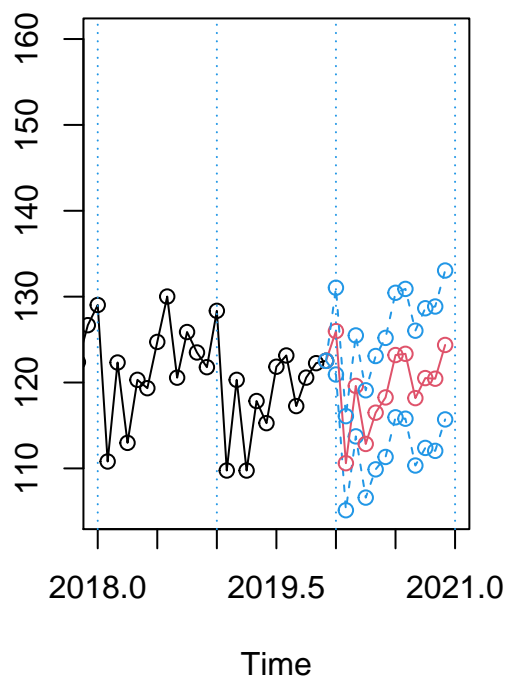
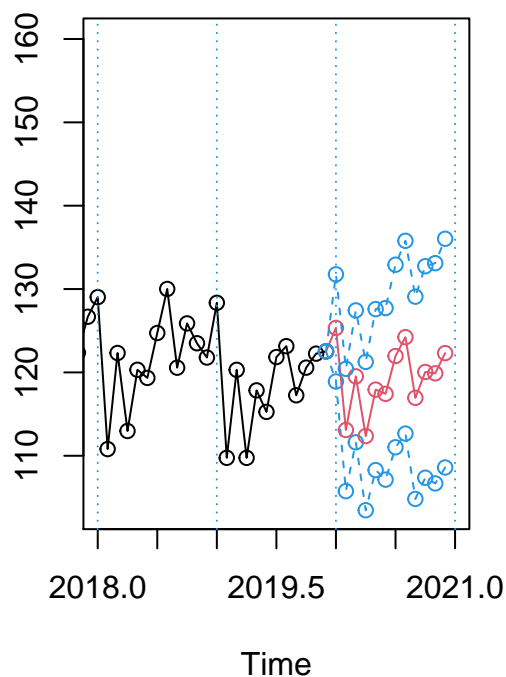


```
previs3=window(cbind(tl3,pr3,tu3),start=ultim+c(1,0))
```

Compare models

```
par(mfrow = c(1,2))
ts.plot(serie,previs1,lty=c(1,2,1,2),col=c(1,4,2,4),xlim=c(2018,2021),type="o",main="Model ARIMA(2,1,1)
abline(v=2016:2021,lty=3,col=4,ylim=c(15,280))
ts.plot(serie,previs3,lty=c(1,2,1,2),col=c(1,4,2,4),xlim=c(2018,2021),type="o",main="Model ARIMA(3,1,0)
abline(v=2016:2021,lty=3,col=4,ylim=c(15,280))
```

Model ARIMA(2,1,1)(0,1,1)12 Model ARIMA(3,1,0)(2,1,0)12+EC+I



```

resul=data.frame(
  par=c(length(coef(mod3)),length(coef(mod.lin2))+nrow(mod.atip$atip)),

  Sigma2Z=c(mod3$sigma2, mod.lin2$sigma2),
  AIC=c(AIC(mod3), AIC(mod.lin2) + 2 * nrow(mod.atip$atip)),
  BIC=c(BIC(mod3), BIC(mod.lin2) + log(length(serie)-13) * nrow(mod.atip$atip)),
  RMSPE=c(mod.EQM1,mod.EQM3),
  MAPE=c(mod.EAM1,mod.EAM3),
  meanLength=c(sum(previs1[,3]-previs1[,1]),sum(previs3[,3]-previs3[,1]))/12)
row.names(resul)=c("ARIMA(2,1,0)(0,1,1)12","ARIMA(0,1,1)(0,1,1)12+EC+IA+Atip")

```

resul

##	par	Sigma2Z	AIC	BIC	RMSPE
## ARIMA(2,1,0)(0,1,1)12	3	10.745501	1832.807	1848.204	0.02125953
## ARIMA(0,1,1)(0,1,1)12+EC+IA+Atip	17	6.640788	1691.693	1760.981	0.01342432
##	MAPE	meanLength			
## ARIMA(2,1,0)(0,1,1)12	0.01604008	20.79713			
## ARIMA(0,1,1)(0,1,1)12+EC+IA+Atip	0.01150628	14.00928			