

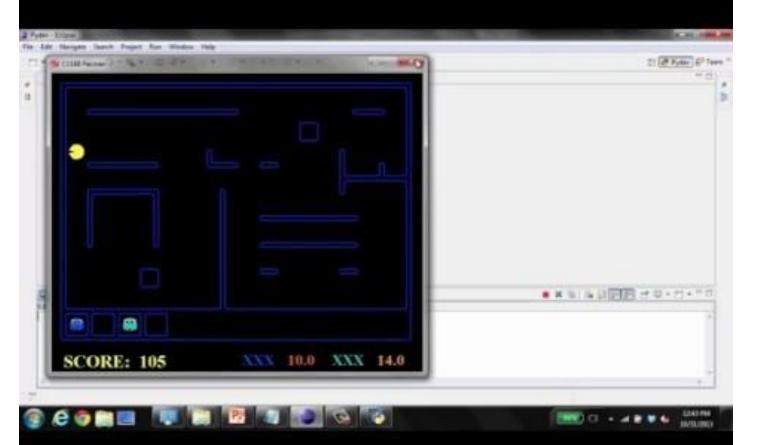
Chapter 5 Hidden Markov Models

COMP 3270 Artificial Intelligence

Dirk Schnieders

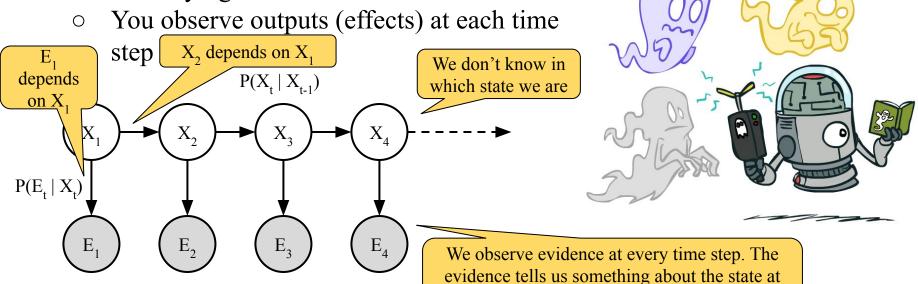
Hidden Markov Models





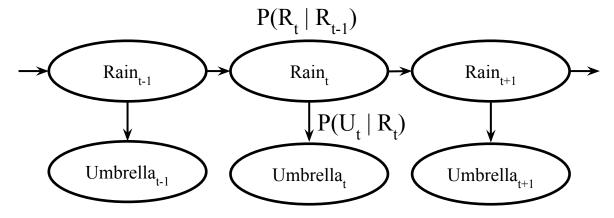
Hidden Markov Models

- Markov chains not so useful for most agents
 - Need observations to update your beliefs
- Hidden Markov models (HMMs)
 - Underlying Markov chain over states X



that time

Example: Weather HMM



- An HMM is defined by
 - \circ Initial distribution $P(R_1)$
 - \circ Transitions $P(R_t | R_{t-1})$
 - \circ Emissions $P(U_t | R_t)$





R_{t}	R_{t+1}	$P(R_{t+1} R_t)$
+r	+r	0.7
+r	-r	0.3
-r	+r	0.3
-r	-r	0.7

R_{t}	U_t	$P(U_t R_t)$
+r	+u	0.9
$+_{r}$	-u	0.1
-r	+u	0.2
-r	-u	0.8

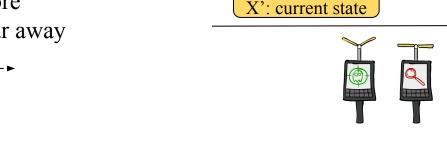
P(X|X') = usually move clockwise, but sometimes move in a random direction or stay in place

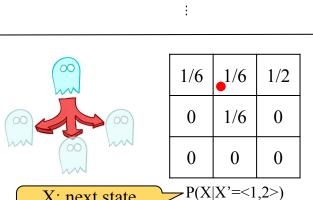
 $P(R_{ii}|X)$ = same sensor model as before

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Where is the ghost?

red means close, green means far away





X: next state

1/9

1/9

1/9

 $P(X_1)$

1/9

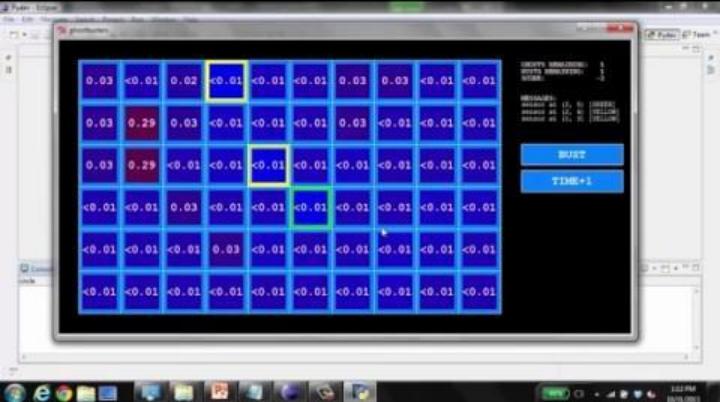
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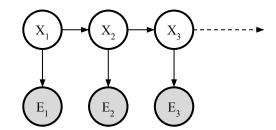
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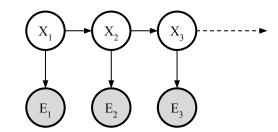


Joint Distribution of an HMM



- Joint distribution:
- $P(X_1, E_1, X_2, E_2, X_3, E_3) = ?$

Joint Distribution of an HMM

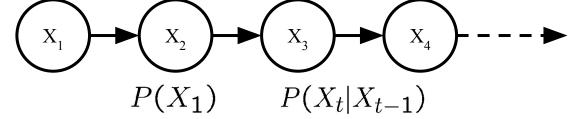


• Joint distribution:

$$P(X_1, E_1, X_2, E_2, X_3, E_3) = P(X_1)P(E_1|X_1)P(X_2|X_1)P(E_2|X_2)P(X_3|X_2)P(E_3|X_3)$$

• More generally:

$$P(X_1, E_1, \dots, X_T, E_T) = P(X_1)P(E_1|X_1)\prod^T P(X_t|X_{t-1})P(E_t|X_t)$$



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• Joint distribution:
$$D(Y \mid Y \mid Y) = D(Y \mid D(Y \mid Y) \mid$$

More generally:

oution:
$$P(X_1,X_2,X_3,X_4) = P(X_1)P(X_2|X_1)P(X_3|X_2)P(X_4|X_3) \label{eq:potential}$$
 rally:

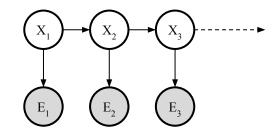
 $= P(X_1) \prod P(X_t|X_{t-1})$

ally:
$$P(X_1, X_2, \dots, X_T) = P(X_1)P(X_2|X_1)P(X_3|X_2)\dots P(X_T|X_{T-1})$$

- Questions to be resolved:

 - Does this indeed define a joint distribution? Can every joint distribution be factored this way, or are we making some assumptions about the joint distribution by using this factorization?

Quiz 1: Joint Distribution of an HMM

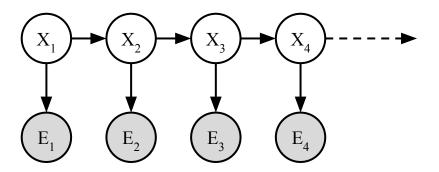


• Joint distribution: Why?

$$P(X_1, E_1, X_2, E_2, X_3, E_3) = P(X_1)P(E_1|X_1)P(X_2|X_1)P(E_2|X_2)P(X_3|X_2)P(E_3|X_3)$$

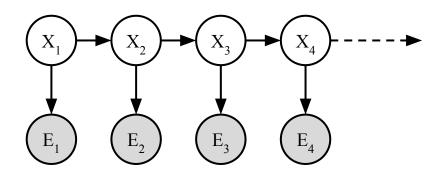
Conditional Independence

- HMMs have two important independence properties
 - Hidden Markov process: future depends on past via present
 - Current observation independent of all else given current state



Quiz 2: Conditional Independence

- HMMs have two important independence properties
 - Markov hidden process: future depends on past via present
 - Current observation independent of all else given current state

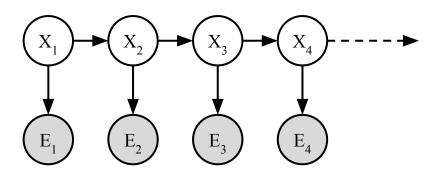


Quiz

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Ones this mean that evidence variables are guaranteed to be independent?

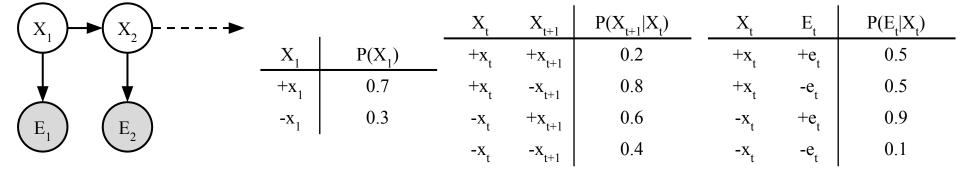
Implied Conditional Independencies



• Many implied conditional independencies, e.g.,

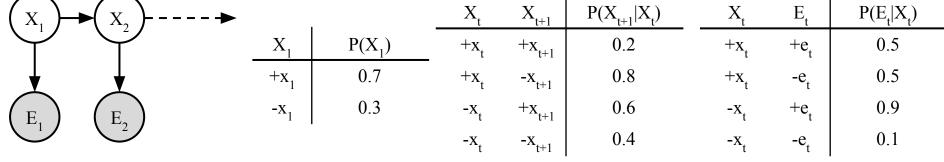
$$E_1 \perp \!\!\! \perp X_2, E_2, X_3, E_3 \mid X_1$$

Quiz 3: Joint Distribution of an HMM



•
$$P(X_1, X_2, E_1 = +e_1, E_2 = -e_2) = ?$$

Quiz 4: Joint Distribution of an HMM



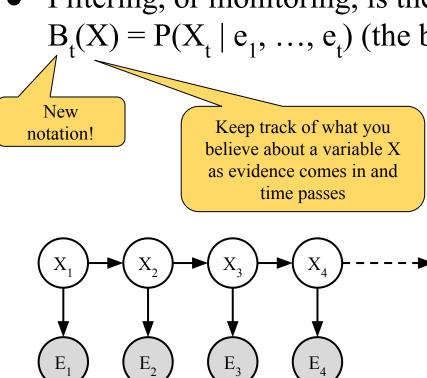
• $P(X_2, E_1 = +e_1, E_2 = -e_2) = ?$

HMM Examples

- Speech recognition
 - Observations are acoustic signals (continuous valued)
 - States are specific positions in specific words (so, tens of thousands)
- Machine translation
 - Observations are words (tens of thousands)
 - States are translation options
- Robot tracking
 - Observations are range readings (continuous)
 - States are positions on a map (continuous)

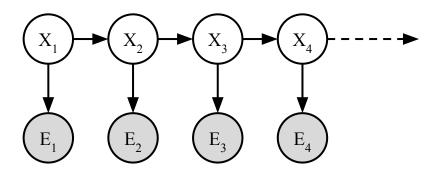
Filtering / Monitoring

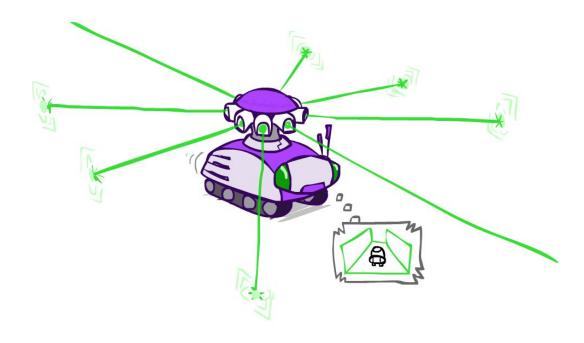
• Filtering, or monitoring, is the task of tracking the distribution $B_t(X) = P(X_t | e_1, ..., e_t)$ (the belief state) over time



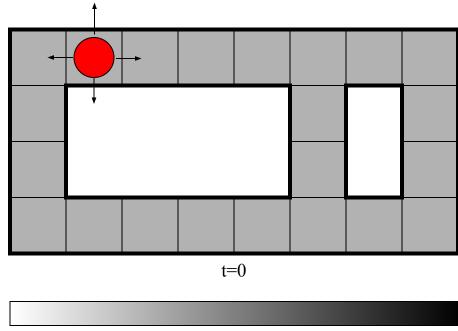
Filtering / Monitoring

- Filtering, or monitoring, is the task of tracking the distribution $B_t(X) = P(X_t \mid e_1, ..., e_t)$ (the belief state) over time
- We start with $B_1(X)$ in an initial setting, usually uniform
- \bullet As time passes, or we get observations, we update B(X)
- The <u>Kalman filter</u> was invented in the 60's and first implemented as a method of trajectory estimation for the Apollo program





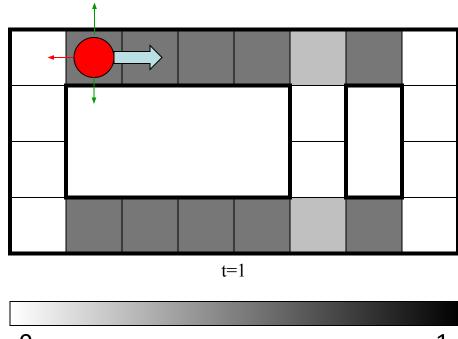
Sensor model: can read in which directions there is a wall, never more than 1 mistake Motion model: may not execute action with small prob.

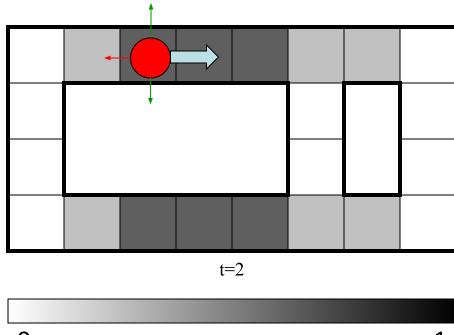


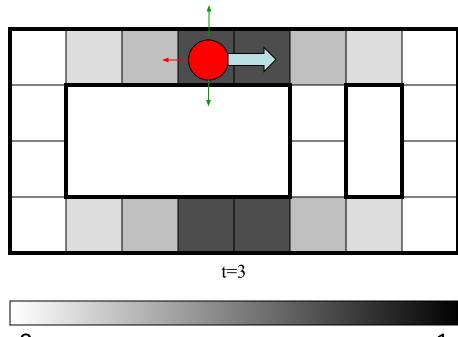
Prob

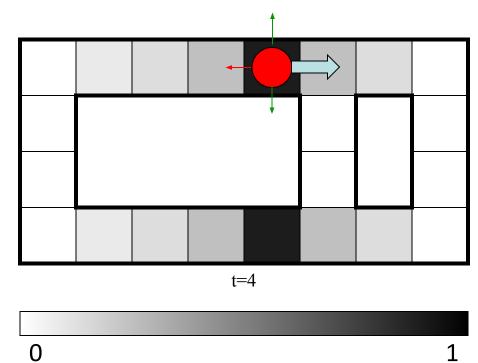
0

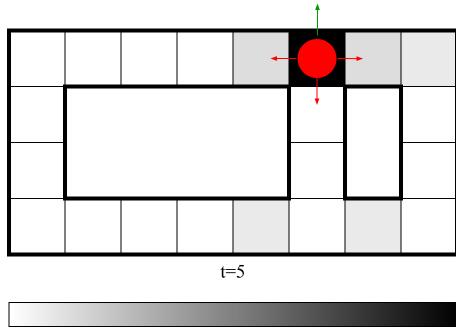
1



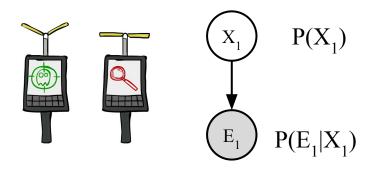






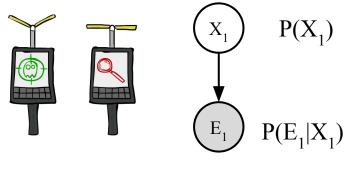


Seeing evidence:



 $P(X_1|e_1) = ?$

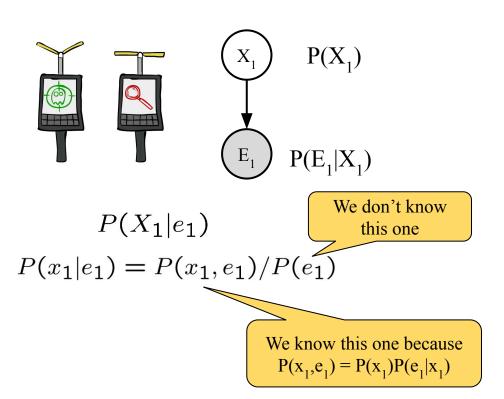
Seeing evidence:



$$P(X_1|e_1)$$

$$P(x_1|e_1) = P(x_1,e_1)/P(e_1)$$
Definition of conditional probability

Seeing evidence:



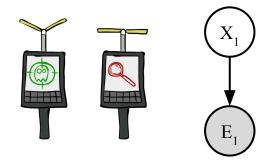
Let's say we have two distributions:

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- Prior distribution over ghost location: P(G)
 - Let this be uniform
 - Sensor reading model: P(R|G)
 - Given: we know what our sensors do
 - \blacksquare R = reading color measured at (1,1)
 - E.g. P(R = yellow | G=(1,1)) = 0.1
- We can calculate the posterior distribution P(G|r) over ghost locations given a reading using Bayes' rule:

$$P(g|r) \propto P(r|g)P(g)$$

Seeing evidence:



$$P(X_1|e_1)$$

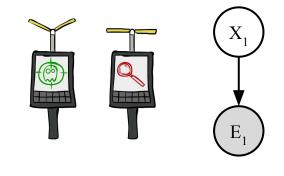
$$P(x_1|e_1) = P(x_1, e_1)/P(e_1)$$

$$\propto_{X_1} P(x_1, e_1)$$

$$= P(x_1)P(e_1|x_1)$$

We take our current propabilities and multiply with the evidence probability. Then renormalize

Seeing evidence:



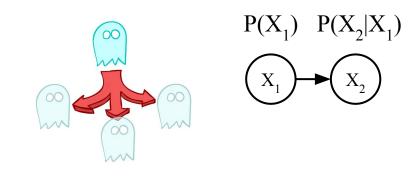
$$P(X_1|e_1)$$

$$P(x_1|e_1) = P(x_1, e_1)/P(e_1)$$

$$\propto_{X_1} P(x_1, e_1)$$

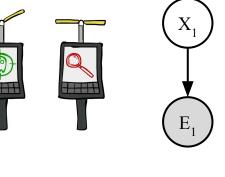
$$= P(x_1)P(e_1|x_1)$$

Time passes:



$$P(X_2)$$

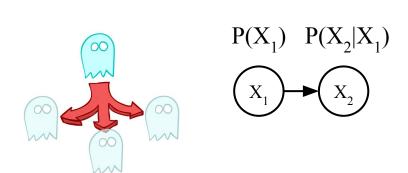
Seeing evidence:



$$P(X_1|e_1)$$

 $P(x_1|e_1) = P(x_1, e_1)/P(e_1)$
 $\propto_{X_1} P(x_1, e_1)$

$$= P(x_1)P(e_1|x_1)$$



 $P(X_2)$

 $= \sum P(x_1)P(x_2|x_1)$

 $P(x_2) = \sum_{x_1} P(x_1, x_2)$

Time passes:

Passage of Time

• Assume we have current believe P(X | evidence to date)

$$B(X_t) = P(X_t|e_{1:t})$$

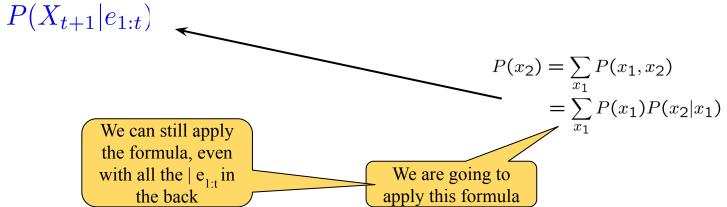
$$e_{1:t} = e_1, e_2, ..., e_t$$

Passage of Time

• Assume we have current believe P(X | evidence to date)

$$B(X_t) = P(X_t|e_{1:t})$$

• Then, after one time step passes:



Passage of Time

• Assume we have current believe P(X | evidence to date)

$$B(X_t) = P(X_t|e_{1:t})$$

• Then, after one time step passes:

$$P(X_{t+1}|e_{1:t}) = \sum_{x_t} P(X_{t+1}, x_t|e_{1:t})$$

$$= \sum_{x_t} P(X_{t+1}|x_t, e_{1:t}) P(x_t|e_{1:t}) \qquad P(x_2) = \sum_{x_1} P(x_1, x_2)$$

$$= \sum_{x_1} P(x_1) P(x_2|x_1)$$
Use conditional independence assumption to get rid of this

Passage of Time

• Assume we have current believe P(X | evidence to date)

$$B(X_t) = P(X_t|e_{1:t})$$

• Then, after one time step passes:

$$P(X_{t+1}|e_{1:t}) = \sum_{x_t} P(X_{t+1}, x_t|e_{1:t})$$

$$= \sum_{x_t} P(X_{t+1}|x_t, e_{1:t}) P(x_t|e_{1:t})$$

$$= \sum_{x_t} P(X_{t+1}|x_t) P(x_t|e_{1:t})$$

Or, compactly

$$B'(X_{t+1}) = \sum P(X'|x_t)B(x_t)$$

New notation! Note that $B'(X_{t+1}) = P(X_{t+1} | e_{1:t})$ is different from

We have not seen e_{t+1} yet

Passage of Time

• Assume we have current believe P(X | evidence to date)

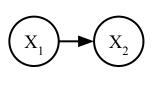
$$B(X_t) = P(X_t|e_{1:t})$$

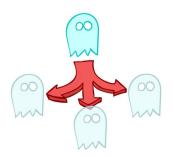
• Then, after one time step passes:

$$P(X_{t+1}|e_{1:t}) = \sum_{x_t} P(X_{t+1}, x_t|e_{1:t})$$

$$= \sum_{x_t} P(X_{t+1}|x_t, e_{1:t}) P(x_t|e_{1:t})$$

$$= \sum_{x_t} P(X_{t+1}|x_t, e_{1:t}) P(x_t|e_{1:t})$$





Or, compactly

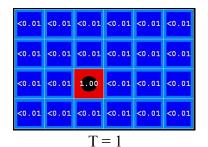
$$B'(X_{t+1}) = \sum_{x} P(X'|x_t)B(x_t)$$

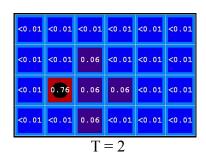
Basic idea

Beliefs get "pushed" through the transitions

Example: Passage of Time

• As time passes, uncertainty "accumulates"





.07 0.03 0.05 <0.01 0.03 <0.01
.03 0.03 <0.01 <0.01 <0.01 <0.01 <0.01

T = 5

0.01 0.05 <0.01 <0.01 <0.0

(Transition model: ghosts usually go clockwise)

0.35 < 0.01 < 0.01



Observation

• Assume we have current belief P(X | previous evidence):

$$B'(X_{t+1}) = P(X_{t+1}|e_{1:t})$$

• Then, after evidence comes in

$$P(X_{t+1}|e_{1:t+1}) = P(X_{t+1},e_{t+1}|e_{1:t})/P(e_{t+1}|e_{1:t})$$

We are going to apply this formula
$$P(x_1|e_1) = P(x_1,e_1)/P(e_1)$$

$$\propto_{X_1} P(x_1,e_1)$$

$$= P(x_1)P(e_1|x_1)$$

Observation

Assume we have current belief $P(X \mid previous evidence)$:

$$B'(X_{t+1}) = P(X_{t+1}|e_{1:t})$$

Then, after evidence comes in

Then, after evidence comes in
$$P(X_{t+1}|e_{1:t+1}) = P(X_{t+1},e_{t+1}|e_{1:t})/P(e_{t+1}|e_{1:t})$$

$$\propto_{X_{t+1}} P(X_{t+1},e_{t+1}|e_{1:t})$$

$$= P(e_{t+1}|e_{1:t},X_{t+1})P(X_{t+1}|e_{1:t})$$

$$= P(e_{t+1}|X_{t+1})P(X_{t+1}|e_{1:t})$$

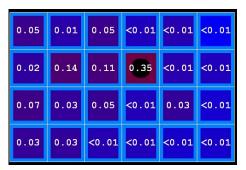


$$B(X_{t+1}) \propto_{X_{t+1}} P(e_{t+1}|X_{t+1})B'(X_{t+1})$$

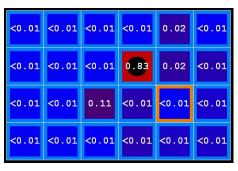
Basic idea: beliefs "reweighted" by likelihood of evidence Unlike passage of time, we have to renormalize

Example: Observation

• As we get observations, beliefs get reweighted, uncertainty "decreases"



Before observation



After observation



 $B(X) \propto P(e|X)B'(X)$



B(-r) = 0.5	B(-r) = ?	B(-r) = ?		-1 (0	111 011	
			_	X_{t}	X_{t+1}	$P(X_{t+1} X_t)$
				+r	+r	0.7
$\left(\begin{array}{c} \operatorname{Rain}_0 \end{array} \right)$	Rain ₁)→(Rain,) —	$+_{\Gamma}$	-r	0.3
				-r	+r	0.3
				-r	-r	0.7
	Umbrella ₁	Umbrella,		X_{t}	\mathbf{E}_{t}	$P(E_t X_t)$
	Ombrena ₁	Offibrella ₂		+r	+u	0.9
				+r	-u	0.1
				-r	+u	0.2
				-r	-u	0.8

0.8

-r

-u

0.8

-r

-u

B(+r) = 0.5 B(-r) = 0.5	b(+r) = ? $B(-r) = ?$	B(+r) = ? $B(-r) = ?$	B(X)

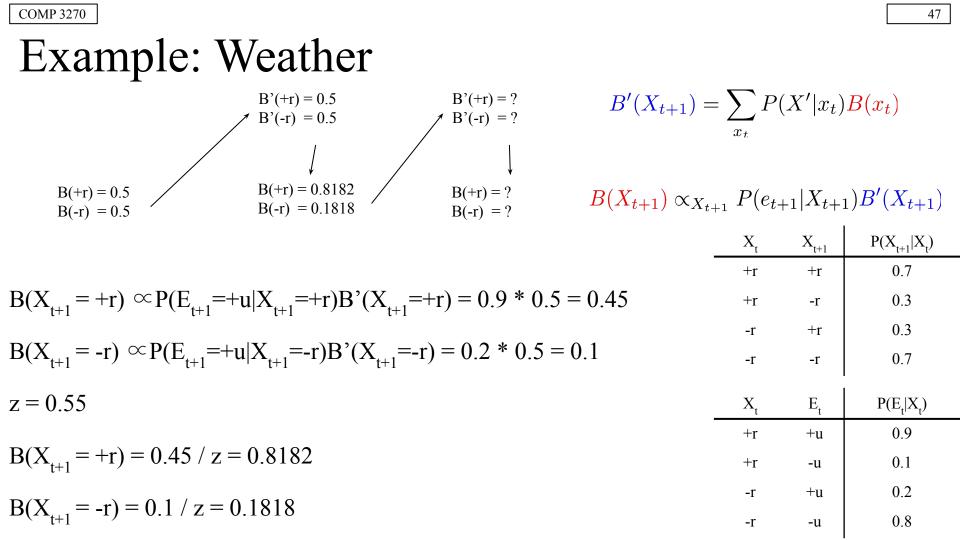
$B(X_{t+1}) \propto_{X_{t+1}}$	$P(e_{t-1})$	$+1 X_{t+1} $	$(X_{t+1})B'(X_{t+1})$
	X_{t}	X_{t+1}	$P(X_{t+1} X_t)$
•	+r	+r	0.7

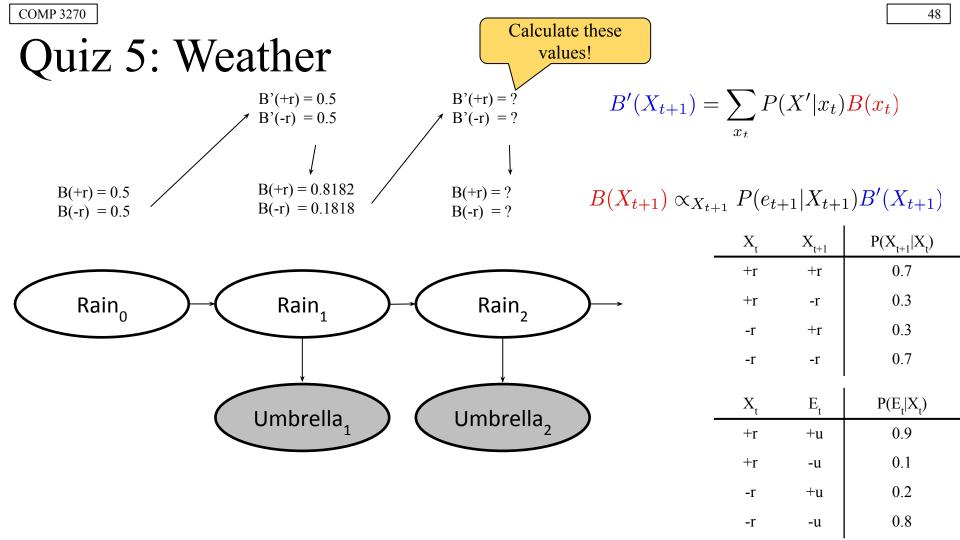
-r

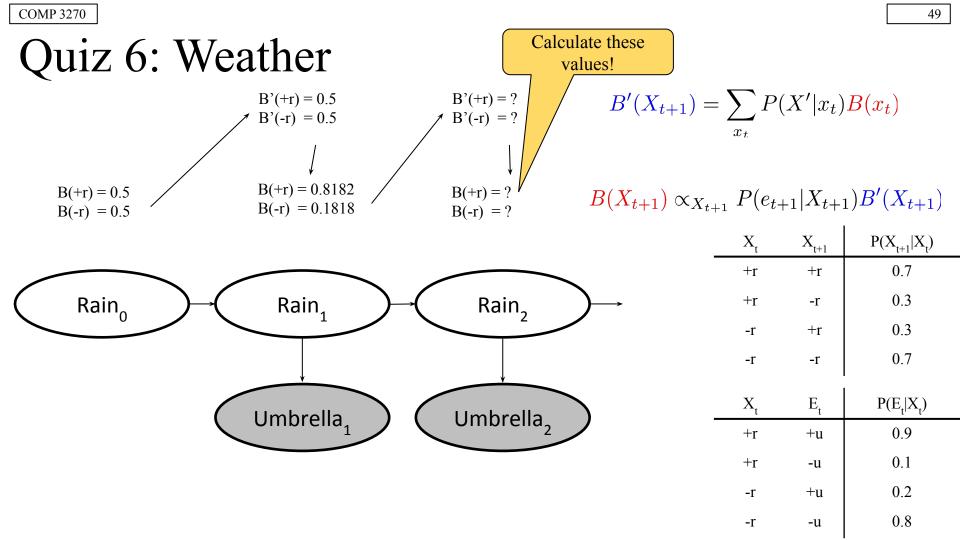
0.3 +r-r 0.3 -r +r0.7 -r -r E_{t} X_{t}

 $P(E_t|X_t)$ 0.9 +r+u0.1 +r-u 0.2 +u-r 0.8

-u







We can normalize as we go if we want to

the end...

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- Alternatively, we can also just do a single update
- We are given evidence at each time and want to know

$$B_t(X) = P(X_t|e_{1:t})$$

- We can derive the following updates
 - have P(x|e) at each time step, or just once at $P(x_t|e_{1:t}) \propto_X P(x_t,e_{1:t})$ $=\sum P(x_{t-1},x_t,e_{1:t})$

$$= \sum_{x_{t-1}} P(x_{t-1}, e_{1:t-1}) P(x_t | x_{t-1}) P(e_t | x_t)$$

$$= P(e_t|x_t) \sum_{x_{t-1}} P(x_t|x_{t-1}) P(x_{t-1}, e_{1:t-1})$$

observation update

Time update

Online Belief Updates

- \bullet Every time step, we start with current P(X | evidence)
- We update for time:

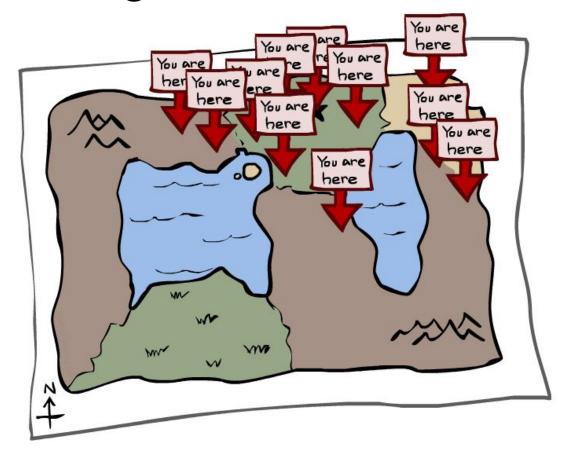
$$P(x_t|e_{1:t-1}) = \sum_{x_{t-1}} P(x_{t-1}|e_{1:t-1}) \cdot P(x_t|x_{t-1})$$

• We update for evidence:

$$P(x_t|e_{1:t}) \propto_X P(x_t|e_{1:t-1}) \cdot P(e_t|x_t)$$

• The forward algorithm does both at once (and doesn't normalize)

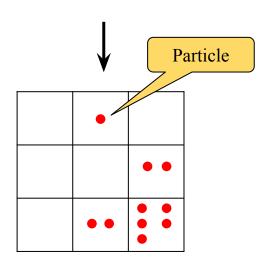
Particle Filtering



Particle Filtering

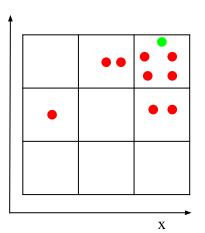
- Filtering: approximate solution
- Sometimes |X| is too big to use exact inference
 - \circ |X| may be too big to even store B(X)
 - E.g. X is continuous
- Solution: approximate inference
 - Track samples of X, not all values
 - Samples are called particles
 - Time per step is linear in the number of samples
 - But: number needed may be large
 - In memory: list of particles, not states
- This is how robot localization works in practice
 - Particle is just new name for sample

0.0	0.1	0.0
0.0	0.0	0.2
0.0	0.2	0.5



Representation: Particles

- Our representation of P(X) is now a list of N particles (samples)
 - \circ Generally, N << |X|
 - Storing map from X to counts would defeat the point
- P(x) approximated by number of particles with value x
 - So, many x may have P(x) = 0!
 - o More particles, more accuracy
- For now, all particles have a weight of 1



Particles (x,y):
(3,3)
(2,3)
(3,3)
(3,2)
(3,3)
(3,2)
(1,2)
(3,3)

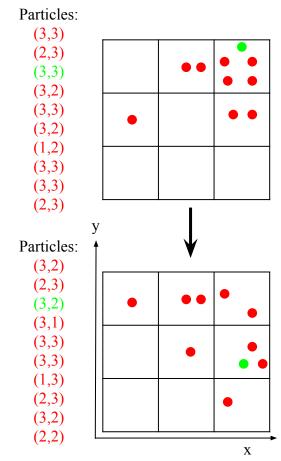
(3,3) (2,3)

Particle Filtering: Elapse of Time

• Each particle is moved by sampling its next position from the transition model

$$x' = \operatorname{sample}(P(X'|x))$$

- This is like prior sampling samples' frequencies reflect the transition probabilities
- Here, most samples move clockwise, but some move in another direction or stay in place
- This captures the passage of time
 - If enough samples, close to exact values before and after (consistent)



Particle Filtering: Observe

- Slightly trickier:
 - o Don't sample observation, fix it
 - Similar to likelihood weighting, downweight samples based on the evidence

$$w(x) = P(e|x)$$

$$B(X) \propto P(e|X)B'(X)$$

• As before, the probabilities don't sum to one, since all have been downweighted (in fact they now sum to (N times) an approximation of P(e))

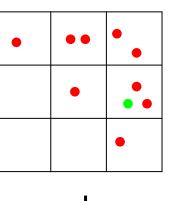
(3,3)

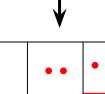
Particles:

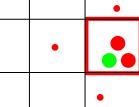
- (2,3) (3,3)
- (3,2)
- (3,2) (3,3)
- (3,3)
- (1,2)
- (3,3) (3,3)
- (2,3)



- (3,2) w=.9 (2,3) w=.2
- (3,2) w=.9
- (3,2) W=.9 (3,1) w=.4
- (3,3) w=.4
- (3,3) w=.4
- (1,3) w=.1
- (2,3) w=.2
- (3.2) w=.9
- 5,2) W-.9
- (2.2) w=.4





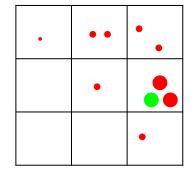


Particle Filtering: Resample

- Rather than tracking weighted samples, we resample
 - N times, we choose from our weighted sample distribution (i.e. draw with replacement)
 - This is equivalent to renormalizing the distribution
 - Now the update is complete for this time step, continue with the next one

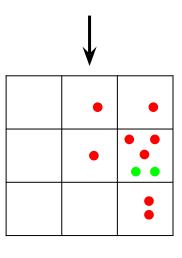


- (3,2) w=.9 (2,3) w=.2
- (3.2) w=.9
- (3,1) w=.4
- (3.3) w=.4
- (3.3) w=.4
- (1.3) w=.1
- (2,3) w=.2
- (3,2) w=.9
- (2.2) w=.4

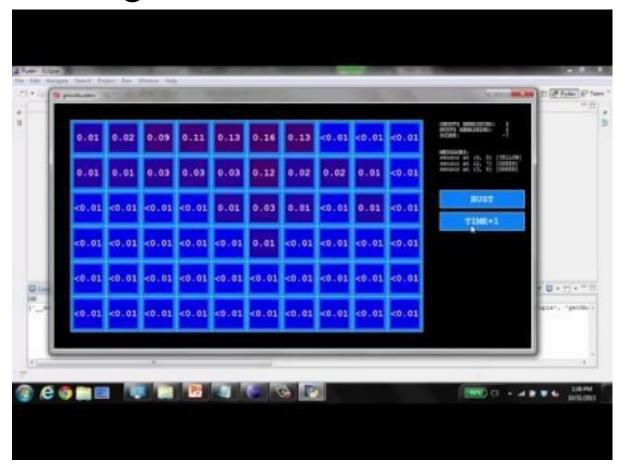


(New) Particles:

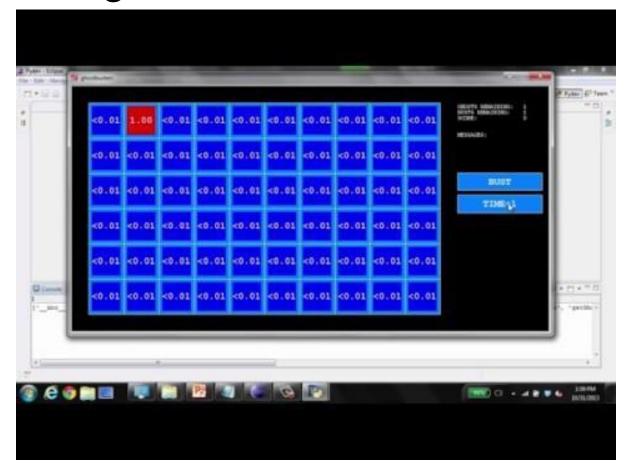
- (3,2)(2,2)
- (3,2)
- (3,1)(3,3)
- (3,2)
- (3,1)
- (2,3)
- (3,2)
- (3,2)



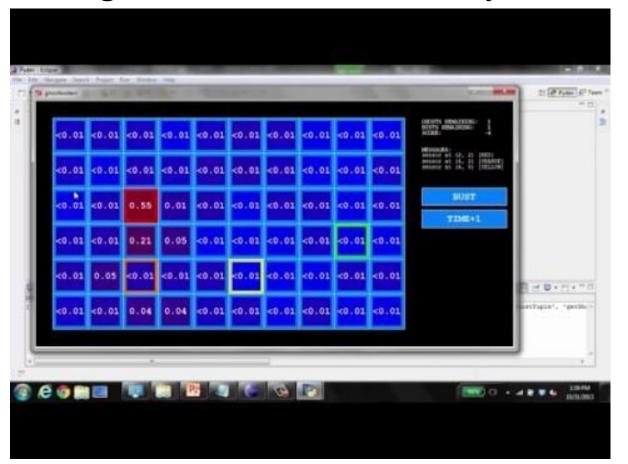
Particle Filtering in Ghostbusters - Few Particles



Particle Filtering in Ghostbusters - One Particle

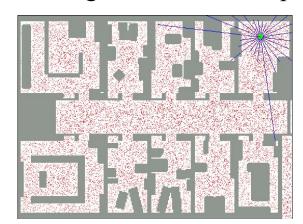


Particle Filtering in Ghostbusters - Many Particles

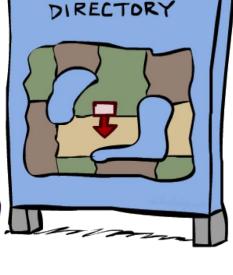


Robot Localization

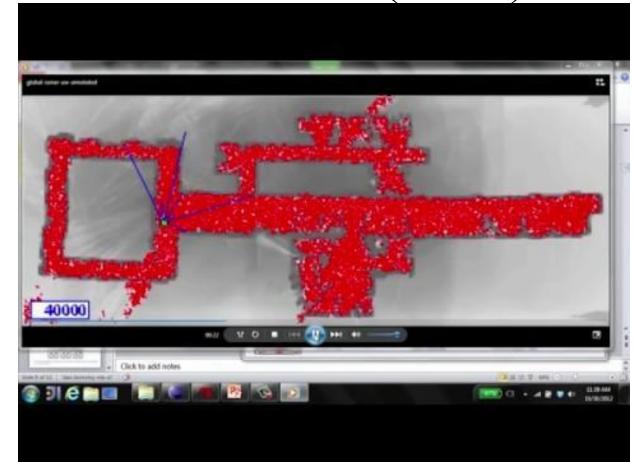
- In robot localization:
 - We know the map, but not the robot's position
 - Observations may be vectors of range finder readings
 - State space and readings are typically continuous (works basically like a very fine grid) and so we cannot store B(X)
 - Particle filtering is a main technique



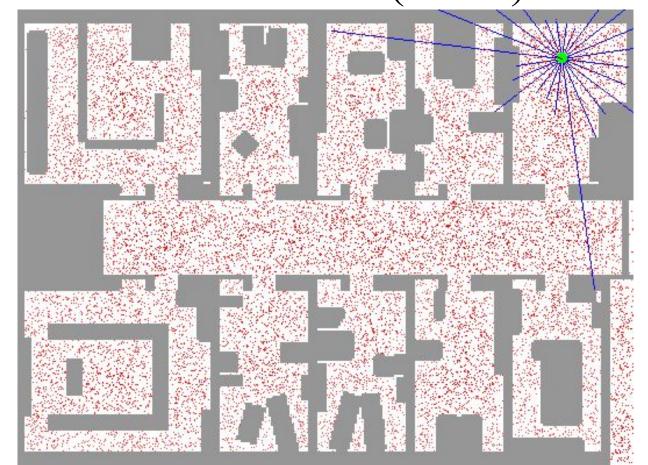




Particle Filter Localization (Sonar)

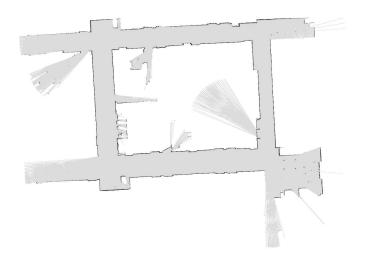


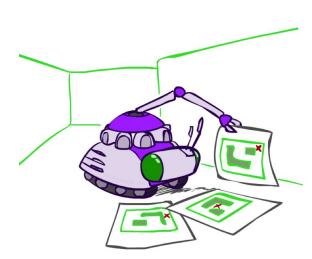
Particle Filter Localization (Laser)



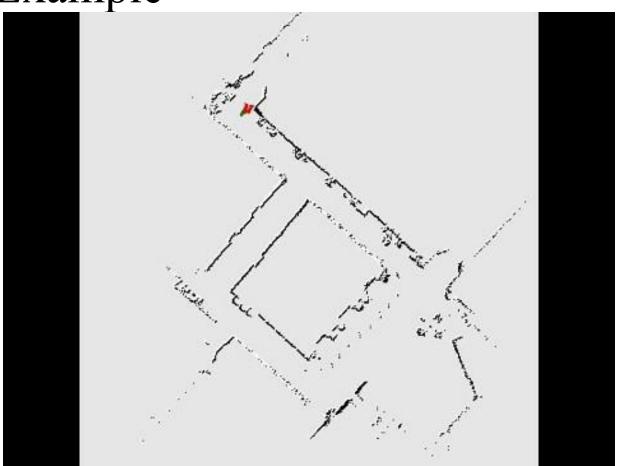
Robot Mapping

- SLAM: Simultaneous Localization And Mapping
 - We do not know the map or our location
 - State consists of position AND map!
 - Main techniques: Kalman filtering (Gaussian HMMs) and particle methods

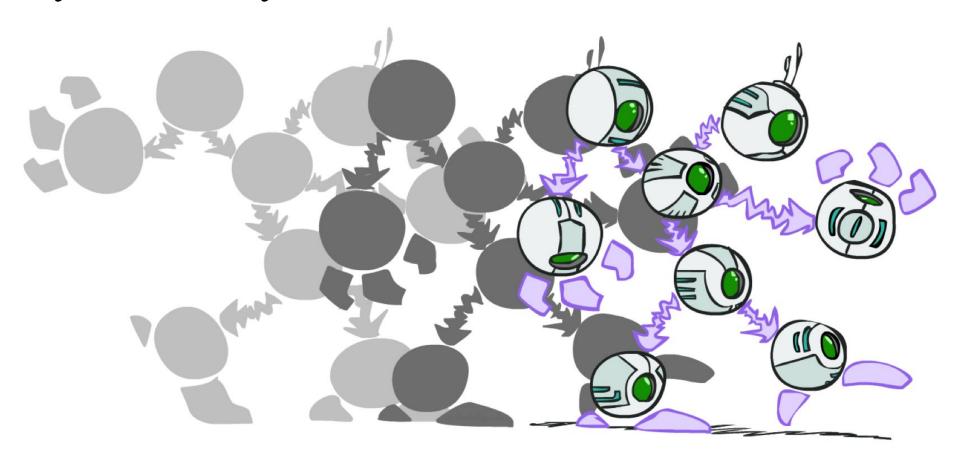




SLAM Example

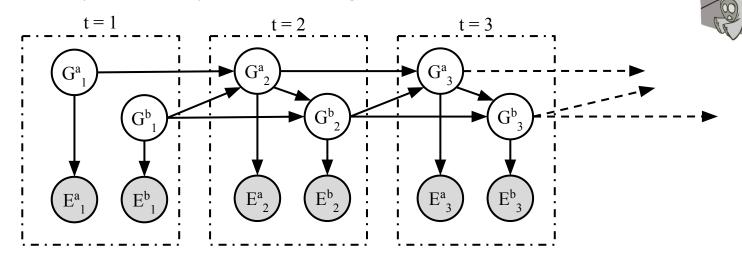


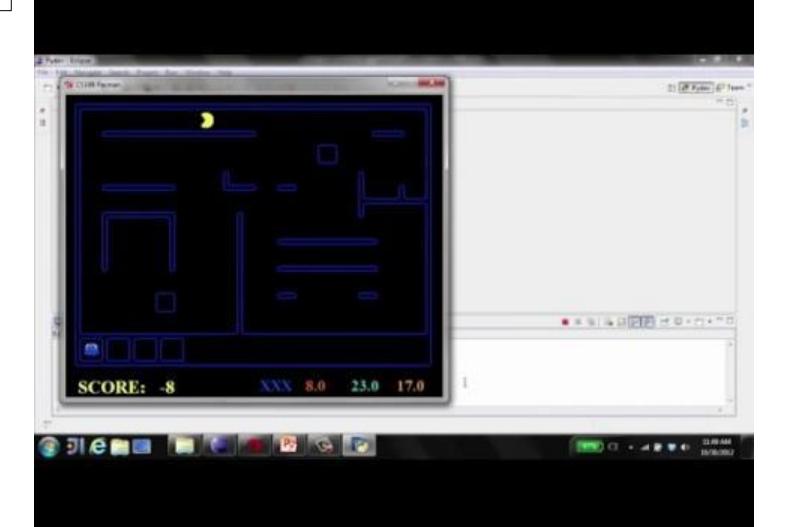
Dynamic Bayes Nets



Dynamic Bayes Nets (DBNs)

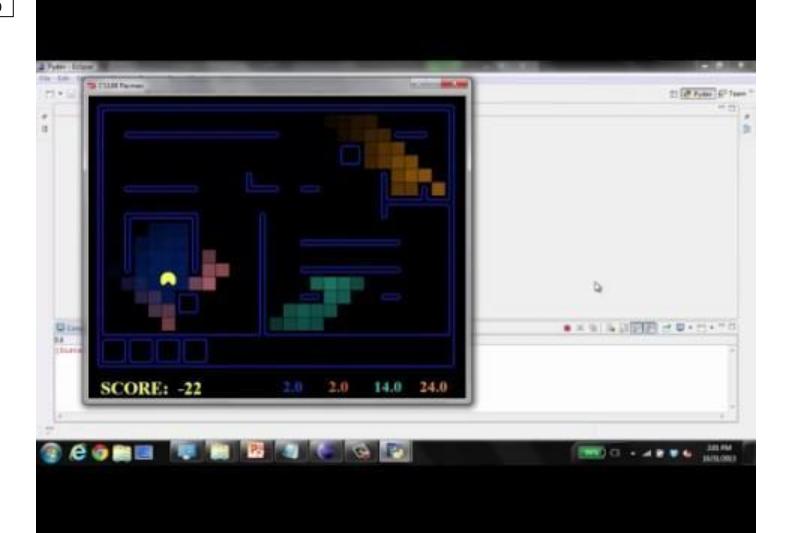
- We want to track multiple variables over time, using multiple sources of evidence
- Idea: Repeat a fixed Bayes net (BN) structure at each time
 - More on BNs later in the course
- Variables from time t can condition on those from t-1
- Dynamic Bayes nets are a generalization of HMMs





DBN Particle Filter

- A particle is a complete sample for a time step
- Initialize: Generate prior samples for the t=1 Bayes net
 - Example particle: $G_1^a = (3,3) G_1^b = (5,3)$
- Elapse time: Sample a successor for each particle
 - Example successor: $G_2^a = (2,3) G_2^b = (6,3)$
- Observe: Weight each entire sample by the likelihood of the evidence conditioned on the sample
 - \circ Likelihood: $P(E_1^a|G_1^a) * P(E_1^b|G_1^b)$
- Resample: Select prior samples (tuples of values) in proportion to their likelihood

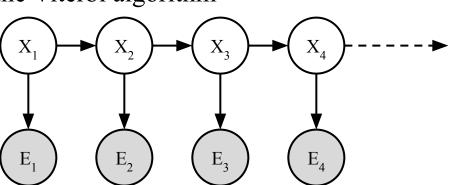


Most Likely Explanation

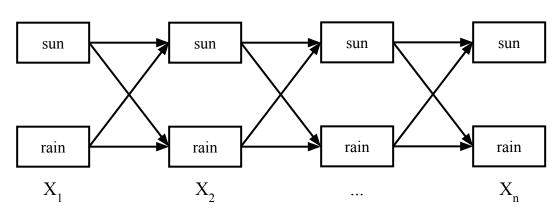


HMMs: MLE Queries

- HMMs defined by
 - States X
 - Observations E
 - Initial distribution: $P(X_1)$
 - Transitions: $P(X|X_{-1})$
 - \circ Emissions: P(E|X)
- New query: most likely explanation: $\underset{x_{1:t}}{\operatorname{arg max}} P(x_{1:t}|e_{1:t})$
- New method: the Viterbi algorithm

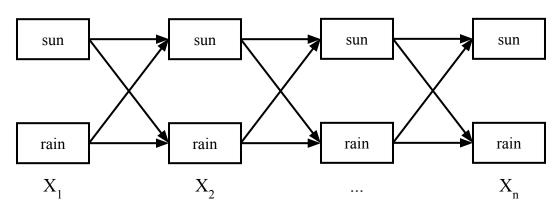


State Trellis



- State trellis: graph of states and transitions over time
- Each arc represents some transition $x_{t-1} \rightarrow x_t$
- Each arc has weight $P(x_t|x_{t-1})P(e_t|x_t)$
- Each path is a sequence of states
- The product of weights on a path is that sequence's probability along with the evidence
- Forward algorithm computes sums of paths, Viterbi computes best paths

Forward / Viterbi Algorithms



Forward Algorithm (Sum)

Viterbi Algorithm (Max)

$$f_t[x_t] = P(x_t, e_{1:t})$$

$$= P(e_t|x_t) \sum_{x_{t-1}} P(x_t|x_{t-1}) f_{t-1}[x_{t-1}]$$

$$m_t[x_t] = \max_{x_1:t-1} P(x_{1:t-1}, x_t, e_{1:t})$$

$$= P(e_t|x_t) \max_{x_{t-1}} P(x_t|x_{t-1}) m_{t-1}[x_{t-1}]$$

Quiz 7

• Solve the problem on the <u>handout sheet 1</u>

Quiz 8

• Solve the problem on the <u>handout sheet 2</u>